Exercises Week 4

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- 1. The advantage of OLS to absolute deviation is the derivability because the second is not derivable in some points
- 2. Normal equation: $\mathbf{X}'\mathbf{X}\hat{\beta} = \mathbf{X}'y$. OLS estimation: $\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'y$ The solution is unique since \mathbf{X} is definite positive, linearly independent $(rk(\mathbf{X}) = \#rows)$ then invertible. If \mathbf{X} is not linearly independent we can't calculate the term $(\mathbf{X}'\mathbf{X})^{-1}$
- 3. MLE require assuming a distribution of the error because to fit into non-linear data distributions. If the errors are distributed normally then MLE converges to OLS
- 4. $H = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'$ Symmetric property:

$$\begin{split} H &= H' \\ \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}' &= (\mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}')' \\ \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}' &= \mathbf{X}((\mathbf{X}'\mathbf{X})^{-1})'\mathbf{X}' \\ \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}' &= \mathbf{X}((\mathbf{X}'\mathbf{X})')^{-1}\mathbf{X}' \\ \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}' &= \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}' \end{split}$$

Idempotent:

$$\begin{split} H \cdot H &= \\ \mathbf{X} (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' \cdot \mathbf{X} (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' &= \\ \mathbf{X} \underbrace{((\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' \mathbf{X})}_{1} (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X} &= \mathbf{X} (\mathbf{X}' \mathbf{X})^{-1} \mathbf{X}' &= H \end{split}$$

- 5. With MLE of a linear model we have a variance $\sigma_{ML}^2 = \frac{\hat{\varepsilon}'\hat{\varepsilon}}{n}$. Since $\mathbb{E}\left[\hat{\varepsilon}'\hat{\varepsilon}\right] = (n-p)\cdot\sigma^2$ we have $\mathbb{E}\left[\sigma_{ML}^2\right] = \frac{n-p}{n}\cdot\sigma^2$ which is a biased estimator. From $\mathbb{E}\left[\hat{\varepsilon}'\hat{\varepsilon}\right]$ we can easily find an unbiased estimator $\hat{\sigma} = \frac{1}{n-p}\varepsilon'\varepsilon$
- 6. $\mathbf{X}' \varepsilon = \mathbf{X}' (\mathbf{I} \mathbf{H}) \mathbf{y} = \mathbf{X}' \mathbf{y} \mathbf{X}' \mathbf{H} \mathbf{y} = \mathbf{X}' \mathbf{y} \underbrace{\mathbf{X}' \mathbf{X} (\mathbf{X}' \mathbf{X})^{-1}}_{\mathbf{1}} \mathbf{X}' \mathbf{y} = \mathbf{0}.$

With the last result and assuming that the model has intercept we can assert that: $\sum_{i} \mathbf{1}\hat{\boldsymbol{\varepsilon}}_{i} = \sum_{i} \varepsilon_{i} = 0$

- 7. The coefficient of determination $R^2 = \frac{s_{\hat{y}}^2}{s_y^2} = \frac{\sum_i (\hat{y}_i \overline{y})^2}{\sum_i (y_i \overline{y})^2}$ vary between [0,1] and explain the goodness of the fit of the regression. If $R^2 = 1$ we have a perfect fit while with $R^2 = 0$ we have $\forall i \ \hat{y}_i = \overline{y}$ so it means that every point \hat{y} is in the mean. $R^2 = 0$ doesn't necessarily mean that the response is unrelated with the variables because there could be a non-linear relationship
- 8. For model M2 we can say that $R_{M2}^2 \geq R_{M1}^2$. Instead for R_{M3}^3 we can't say anything except if the data are distributed logarithmically then $R_{M3}^2 \geq R_{M1}^2$
- 9. The three condition for compare model by \mathbb{R}^2 are:
 - Same response variable
 - Same number of parameters
 - Must include intercept
- 10. Zero mean: $\mathbb{E}\left[\varepsilon\right] = 0$

Homoscedasticity: $Cov\left(\varepsilon\right)=I\cdot\sigma^{2}$ No correlation: Corr(x,y)=0

Normality: $\varepsilon N(0, \sigma^2 I)$

11. $\mathbb{E}\left[\hat{\beta}\right] = \mathbb{E}\left[(X'X)^{-1}X'y\right] = (X'X)^{-1}X'\,\mathbb{E}\left[y\right] = (X'X)^{-1}X'X\beta = \beta$