

Exercises week 5

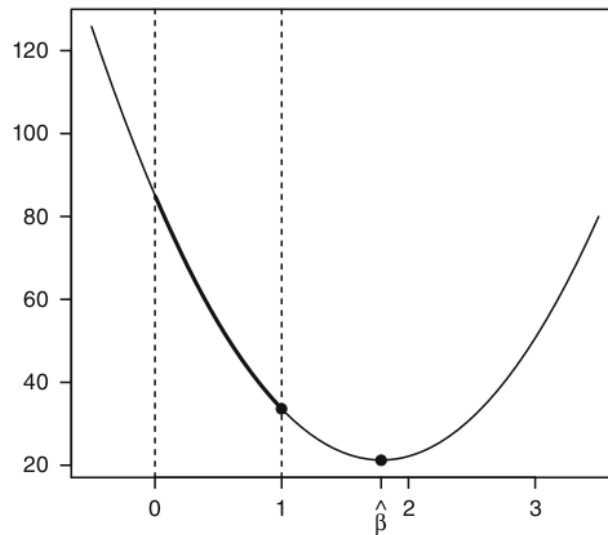
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1.

$$\begin{aligned} i) \quad \mathbf{C} &= (0, 5, -7, 0), & d &= -2 \\ ii) \quad \mathbf{C} &= \begin{pmatrix} 0 & 1 & -1 & 0 & 0 \\ 0 & 0 & 1 & -1 & 0 \end{pmatrix}, & d &= \begin{pmatrix} 0 \\ 0 \end{pmatrix} \end{aligned}$$

2. This plot shows the SSE in function of β :



The F-test score is $\frac{SSE_{H_0} - SSE}{SSE}$ where $SSE = SSE(\hat{\beta})$ that is the y in the plot where $x = \hat{\beta}$ and $SSE_{H_0} = SSE(\beta = 1)$

The plot for the numerator of Wald test is the same of the previous because the numerator is the same except for a square difference instead of normal difference.

The main difference between F-test and Wald test is the denominator in the first is $\hat{\theta}$ while in the second is $var(\hat{\theta})$. There is also a relation between test statistics F and W : $W = rF$

3. The confidence interval for μ_0 is $\mathbf{x}'_0 \hat{\beta} \pm t_{n-p}(1 - \alpha/2) \hat{\sigma}(\mathbf{x}'_0(\mathbf{X}'\mathbf{X})^{-1}\mathbf{x}_0)^{1/2}$ while for prediction is $\mathbf{x}'_0 \hat{\beta} \pm t_{n-p}(1 - \alpha/2) \hat{\sigma}(1 + \mathbf{x}'_0(\mathbf{X}'\mathbf{X})^{-1}\mathbf{x}_0)^{1/2}$. The

difference between the two formulas is the "1 +" near x_0 that become from the fact that the estimation try to give an confidence interval of a parameter , so with no variance, while the prediction try to estimate a aleatory variable with a expected value and a variance.

$$4. f(x_0) = \mathbf{x}'_0 \hat{\beta} \pm t_{n-2}(1-\alpha/2) \hat{\sigma} (1 + \mathbf{x}'_0 (\mathbf{X}'\mathbf{X})^{-1} \mathbf{x}_0)^{1/2} \text{ where } X = \begin{pmatrix} 1 & x_{11} \\ \vdots & \vdots \\ 1 & x_{n1} \end{pmatrix}$$

and $p = 2$.

The length of interval is minimum when $\mathbf{x}_0 = \mu$ because we can write the product $\mathbf{x}'_0 (\mathbf{X}'\mathbf{X})^{-1} \mathbf{x}_0$ as $\frac{\sum (x_i - x_0)^2}{dev(x)}$ then the numerator $\sum (x_i - x_0)^2$ is minimized when $x_0 = \mu$

5. Increasing the number of parameters is not always a good idea: we can ex-

press the SPSE (expected squared prediction error) as: $\underbrace{n\sigma^2}_{\text{irreducible part}} + \underbrace{|M|\sigma^2}_{\text{Variance error}} + \underbrace{\sum_{i=1}^n (u_{iM} - u_i)^2}_{\text{BIAS}^2}$

6. If we consider only \mathbf{X}_1 we have $\mathbf{X} = \mathbf{X}_1$ then

$$\begin{aligned} E[\tilde{\beta}_1] &= E[(\mathbf{X}'_1 \mathbf{X}_1)^{-1} \mathbf{X}'_1 \mathbf{y}] \\ &= (\mathbf{X}'_1 \mathbf{X}_1)^{-1} \mathbf{X}'_1 E[\mathbf{y}] \\ &= (\mathbf{X}'_1 \mathbf{X}_1)^{-1} \mathbf{X}'_1 (\mathbf{X}_1 \beta_1 + \mathbf{X}_2 \beta_2) \\ &= \beta_1 + (\mathbf{X}'_1 \mathbf{X}_1)^{-1} \mathbf{X}'_1 \mathbf{X}_2 \beta_2 \end{aligned}$$

The bias vanish if $\mathbf{X}'_1 \mathbf{X}_2 = \mathbf{0}$ that in statcal words mean that \mathbf{X}_1 and \mathbf{X}_2 are uncorrelated or when β_2 is irrelevant

7. If we omit a relevant covariate, we can have a decrease the variance of model such that the increase of $bias^2$ is overall convenient

8. $SPSE = \sum (y_{n+i} - \hat{y}_{iM})^2$. See point 5

9.

$$\begin{aligned} E(SSE) &= E\left(\sum (y_i - \hat{y}_{iM})^2\right) \\ &= E(\varepsilon' \varepsilon) \\ &= (n - p) \sigma^2 \\ &= E(SPSE) - 2|M| \sigma^2 \\ &= n\sigma^2 + |M| \sigma^2 - 2|M| \sigma^2 \\ &= n\sigma^2 - |M| \sigma^2 \\ &= (n - p) \sigma^2 \end{aligned}$$

SSE underestimate the $SPSE$ because $\hat{SPSE} = \hat{SSE} + 2|M| \hat{\sigma}^2$.

This bias is more severe for complex models because the bias is proportional to model complexity

10. $AIC = n \log(\hat{\sigma}^2) + 2(|M| + 1)$
 $BIC = n \log(\hat{\sigma}^2) + \log(n)(|M| + 1)$.
 Since $SPSE$ is not observable the two latter indicators are used to evaluate the fit of the model.
 The difference between AIC and BIC is the term before $(|M| + 1)$: in AIC is equal to 2 while in BIC is $\log(n)$. This mean that BIC have a major penalty when the model complexity increase in big dataset.
 Next we can see the tables with the results:

n = 100	AIC	BIC
$l = -300, M = 5$	612	627.63
$l = -290, M = 9$	600	626.05

n = 200	AIC	BIC
$l = -300, M = 5$	612	631.79
$l = -290, M = 9$	600	632.98

Between $n = 100$ and $n = 200$ we can note the phenomenon described before with BIC , because with $n = 100$ there is a decrement of BIC between the two configurations while for $n = 200$ the BIC increase between the two configurations

11. $VIF_j = \frac{1}{1 - R_j^2}$.

It measures the correlation between the covariate x_j with the other covariates. The greater is $R_j^2 \in [0, 1]$ the greater is VIF_j . An empiric alert for variance is when $VIF_j > 10$

12. The ridge regression add a penalty term λ to the linear regression. The $\text{PLS}(\beta) = (\mathbf{y} - \mathbf{X}\beta)'(\mathbf{y} - \mathbf{X}\beta) + \lambda\beta'\beta$.

The ridge regression is biased respect to linear regression:

$$E_{LS}(\hat{\beta}) = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{X}\beta = \beta$$

$$E_{PLS}(\hat{\beta}) = (\mathbf{X}'\mathbf{X} + \lambda\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{X}\beta \text{ (when penalty} = 0 \text{ } E_{PLS}(\hat{\beta}) = E_{LS}(\hat{\beta}))$$

When λ increases the coefficients β tends to zero because $\lambda\mathbf{K} = (\mathbf{X}'\mathbf{X} + \lambda\mathbf{K}) - \mathbf{X}'\mathbf{X}$.

If we multiply two definite semipositive matrices $(\mathbf{X}'\mathbf{X} + \lambda\mathbf{K})^{-1}$ with $(\mathbf{X}'\mathbf{X} + \lambda\mathbf{K}) - \mathbf{X}'\mathbf{X}$ also the result is definite semipositive that is $\mathbf{I}_p - (\mathbf{X}'\mathbf{X} + \lambda\mathbf{K})^{-1}\mathbf{X}'\mathbf{X}$.

This imply that $(\mathbf{X}'\mathbf{X} + \lambda\mathbf{K})^{-1}\mathbf{X}'\mathbf{X}$ have the diagonal elements between $[0, 1]$ and this term appears when we express $\hat{\beta}_{PLS}$ in function of $\hat{\beta}_{LS}$:

$$\hat{\beta}_{PLS} = (\mathbf{X}'\mathbf{X} + \lambda\mathbf{K})^{-1}\mathbf{X}'\mathbf{X}\beta_{LS}$$

The covariance matrix of $\hat{\beta}_{PLS}$ for the same reason is lower than variance of $\hat{\beta}_{LS}$ so in some cases can be convenient use a ridge regression when the bias error is less relevant than variance error

Stata code

```
clear
set seed 1032299
set obs 150
gen x1 = runiform()
gen x2 = x1 + runiform()
```

```
gen x3 = runiform()  
gen y = rnormal(-1 + 0.3 * x1 + 0.2*x3, 0.2^2)  
regress y x1 x2 x3  
regress y x1 x3
```