

# Computational Numerical Statistics

## Assignment 11

Migla Miskinyte Reis

26.12.20

**Exercise 4.** (0.1 points) **Points:** \_\_\_\_\_

Let  $X_1, \dots, X_n \sim \text{Pois}(\theta)$ . Show that  $\text{Gamma}(a, b)$  is a natural conjugate prior for  $\theta$ .

Assuming a parameter  $\alpha \in (0, \infty)$  from  $\text{Pois}$  and prior distribution for  $\alpha$  is  $\text{Gamma}(a, b)$  distribution, we can write a prior for  $\alpha$  :

$$f_{\alpha}(\theta) = \frac{b^a}{\Gamma(a)} \theta^{(a-1)} e^{-b\theta}$$

and a likelihood function:

$$f_{X|\alpha}(x_1, \dots, x_n|\alpha) = \prod_{i=1}^n \frac{\theta^{x_i} e^{-\theta}}{x_i!}$$

If we ignore constants that do not depend on  $\theta$ , the posterior distribution of  $\alpha$  parameter given that:

$$X_1=x_1, \dots, X_n=x_n$$

is

$$\begin{aligned} f_{\alpha|X}(\alpha|x_1, \dots, x_n) &= f_{X|\alpha}(x_1, \dots, x_n|\alpha) f_{\alpha}(\theta) = \\ &= \prod_{i=1}^n \theta^{x_i} e^{-\theta} \theta^{(a-1)} e^{-b\theta} \\ &= \theta^{c+a-1} e^{-(n+b)\theta}, \end{aligned}$$

where  $c$  is equal to  $x_1 + \dots + x_n$ . We can see that this equation is the same as probability density function of the gamma distribution  $\text{Gamma}(c+a, n+b)$ , so that means that the posterior distribution of parameter  $\alpha$  must be  $\text{Gamma}(c+a, n+b)$ . Because the prior and posterior are both gamma distributions, the gamma distribution is a natural conjugate prior in the Poisson model.

**Exercise 6.** (0.1 points) **Points:** \_\_\_\_\_

Consider a drug to be given for the relief of chronic pain and that experience with similar compounds has suggested that response rates, say  $\theta$ , between 0.2 and 0.6 could be feasible. Let  $Y$  be the number of patients that experienced pain relief (positive response) in a sample of  $n$  treated patients. One has  $Y \sim \text{Bin}(n, \theta)$ .

According to the prior information above, a prior for  $\theta$  should be chosen such that it has mean  $\mu = \frac{0.2 + 0.6}{2} = 0.4$  and such that the standard deviation  $\sigma$  is so that  $\mu \pm 2 \times \sigma$  gives us the interval  $[0.2, 0.6]$ , i.e.,  $\sigma = 0.1$ . Assume that  $n = 20$  volunteers were treated with the compound and that out of those we observed  $y = 15$  positive responses to treatment.

**6.1** Given that the Beta distribution is a conjugate prior for the Binomial model, identify the prior distribution's hyperparameters.

Given  $\theta \sim \text{Beta}(a, b)$  is our prior for Binomial model, we know that prior mean is:

$$E(\theta) = \frac{a}{a+b} = 0.4$$

From here:

$$b = 1.5a$$

We also know that standard deviation  $\sigma = 0.1$  and variance for prior Beta distribution is:

$$\text{Var}(\theta) = \frac{ab}{(a+b)^2(a+b+1)} = SD^2 = 0.01$$

We substitute  $b = 1.5a$  and solve the previous equation, where we find  $a = \frac{46}{5} = 9.2$ , so  $b = 13.8$ . From here we can see that our prior  $\text{Beta}(a, b)$  is approximately  $\text{Beta}(9, 14)$ .

## 6.2

Derive the posterior distribution  $p(\theta|y)$ . We know that in order to calculate posterior distribution for the Beta-Binomial case, we can derive a very general relationship between the likelihood, prior, and posterior:  $p(\theta|X) \propto p(\theta)p(X|\theta)$ . Given the Binomial likelihood up to proportionality (ignoring the constant):  $\theta^y(1-\theta)^{n-y}$  and given the prior, also up to proportionality,  $\theta^{a-1}(1-\theta)^{b-1}$ , we can write:

$$\theta^y(1-\theta)^{n-y}\theta^{a-1}(1-\theta)^{b-1} = \theta^{a+y-1}(1-\theta)^{b+n-y-1}$$

From this given Binomial likelihood  $(n, y|\theta)$  and prior  $\text{Beta}(a, b)$ , our posterior will be:

$$\text{Beta}(a+y, b+n-y)$$

So, in is simply  $\text{Beta}(9+15, 14+20-15) = \text{Beta}(24, 19)$ .

**6.3** Given the observed data, plot the likelihood versus the prior and the posterior as in the class **Example 2**.

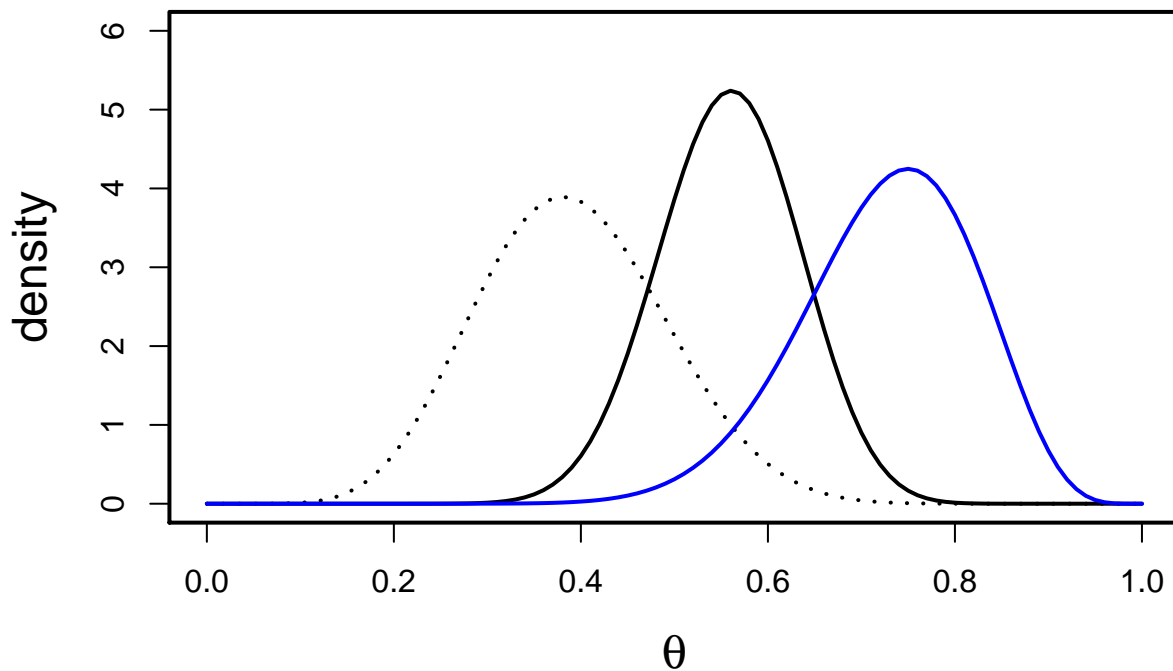
```
n = 20; y = 15; a = 9; b = 14

# prior Beta(9,14)
curve(dbeta(x,a,b),from=0,to=1,lwd=2,
type="l",ylab="density",
xlab=expression(theta),
cex.lab=1.5,cex.main=2,
lty=3,ylim=c(0,6))
# posterior Beta(24,19)
curve(dbeta(x,a+y,b+n-y),add=T,lwd=2,
type="l",ylab="density",cex.lab=1.5,
```

```

cex.main=2)
# proportional likelihood Beta(16,6)
curve(dbeta(x,y+1,n-y+1),add=T,col=4,lwd=2)
legend(0.25, 14,
legend=c("prior","posterior","likelihood"),
col=c(1,1,4),lty=c(3,1,1),cex=1,
lwd=c(1,2,2))
box(lwd=2)

```



#### 6.4

Compute the Bayes estimate of  $\theta$  for the quadratic loss and compare it with the ML estimate.

```
theta.mle = y/n; theta.mle; #calculating ML estimate
```

```
## [1] 0.75
```

```
#Prior mean
```

```
e.theta.prior = a/(a+b); e.theta.prior #calculating theta prior
```

```
## [1] 0.3913043
```

```
e.theta.post = (a+y)/(a+b+n); e.theta.post #theta postoreior
```

```
## [1] 0.5581395
```

Now we will compute Bayes estimate of  $\theta$  for the quadratic loss. Our prior for  $\theta \sim \text{Beta}(9,14)$  and posterior  $\theta|y \sim \text{Beta}(24,19)$ .

$$\hat{\theta}_B = E(\theta|y) = \frac{24}{24 + 19} = 0.058$$

Or, calculating in R:

```
a=24; b=19
# Bayes estimator with quadratic loss
e.theta = a/(a+b); e.theta
```

```
## [1] 0.5581395
```

Bayes estimate of  $\theta$  for the quadratic loss is 0.558, while ML estimate is 0.75.

**6.5** Compute both the 90% CC and 90%-HPD intervals for  $\theta$ . We could also display the 90% credible interval, the range over which we are 90% certain the true value of  $\theta$  lies, given the data and model.

```
qbeta(c(0.05,0.95),shape1=24,shape2=19)
```

```
## [1] 0.4332790 0.6798878
```

```
library(HDIInterval)
HDP = hdi(qbeta, 0.90, shape1=24, shape2=19); HDP
```

```
##      lower      upper
## 0.4351697 0.6817033
## attr(,"credMass")
## [1] 0.9
```

## 6.6

Test the researcher's hypothesis that the percentage of treated individuals that experience a positive result is less or equal than 35%.

We want to test the following hypothesis:

$$H_0: Y \leq 0.35 \text{ vs } H_1: Y > 0.35.$$

For this we need to calculate the prior odds:

```
prior.odds = pbeta(0.35,shape1=9, shape2=14)/(1-pbeta(0.35,shape1=9, shape2=14)); prior.odds
```

```
## [1] 0.5465314
```

And then compute posterior odds:

```
posterior.odds = pbeta(0.35,shape1=24, shape2=19)/(1-pbeta(0.35,shape1=24, shape2=19)); posterior.odds

## [1] 0.002752952
```

We then calculate the Bayes factor for  $H_0$  hypothesis:

```
BF = posterior.odds/prior.odds; BF

## [1] 0.005037134
```

We reject our null hypothesis assuming that we test at 95% significance level.

## 6.7

What's the expected number of treated patients that experience pain relief in a future sample of  $m = 5$ ?

Given the posterior Beta(9,14) distribution as our prior for now, one now wishes to predict the number of treated patients that experience pain relief with  $m=5$ , for this we need to compute  $E(z|y)$ . We will use the R library LearnBayes to compute these probabilities at once via function `pbetap()`.

```
library(LearnBayes)
ab = c(24,19) # posterior (c,d) values are now the prior (a,b) values for pbetap()
m = 5 # new sample size
z = 0:m # values of z for which we want to compute the probabilities
pred = pbetap(ab,m,z) # call the R function
pred
```

```
## [1] 0.02193634 0.11445044 0.26011465 0.32204670 0.21738153 0.06407034
```

```
e.Z = sum(z*pred); e.Z
```

```
## [1] 2.790698
```

The expected mean of the number treated patients that experience pain relief in a future sample of  $m = 5$  inspected individuals is  $E(Z) = 2.8$ .