

# Tut5

August 28, 2023

## 1 Tutorial 5

### 1.1 6.1

(a)

$$\begin{aligned}
 & P(\theta_A, \theta_B) \\
 &= P(\theta, \gamma) \left| \frac{\partial(\theta_A, \theta_B)}{\partial(\theta, \gamma)} \right| \\
 &= P(\theta) P(\gamma) \theta \\
 &= P_\theta(\theta_A) P_\gamma\left(\frac{\theta_B}{\theta_A}\right) \theta_A
 \end{aligned}$$

$$\begin{aligned}
 & P(\theta_A) P(\theta_B) \\
 &= P_\theta(\theta_A) \int_0^\infty \frac{1}{\theta} P_\gamma\left(\frac{\theta_B}{\theta}\right) P_\theta(\theta) d\theta
 \end{aligned}$$

In most circumstance, the part  $P_\gamma\left(\frac{\theta_B}{\theta_A}\right)\theta_A$  is a function of  $\theta_A$  and  $\theta_B$  since the  $P_\gamma\left(\frac{\theta_B}{\theta_A}\right)$  is the density of Gamma distribution. Thus the part  $P_\gamma\left(\frac{\theta_B}{\theta_A}\right)\theta_A$  is unequal to the part  $P(\theta_B)$  which is only function of  $\theta_B$

So in most circumstance

$$P(\theta_A, \theta_B) \neq P(\theta_A) P(\theta_B)$$

which means the  $\theta_A$  and  $\theta_B$  are dependent. But with some restrictions to hyperparameters we may get the independent  $\theta_A$  and  $\theta_B$ . The restriction can be derive as below.

The kernel of  $P_\gamma\left(\frac{\theta_B}{\theta_A}\right)\theta_A$  is  $\left(\frac{\theta_B}{\theta_A}\right)^{(a_\gamma-1)} e^{-(b_\gamma \frac{\theta_B}{\theta_A})} \theta_A$ . To make it become only a function of  $\theta_B$ , we need  $a_\gamma = 2, b_\gamma = 0$ .

(b)

$$\begin{aligned}
 & P(\theta | \mathbf{y}_A, \mathbf{y}_B, \gamma) \\
 & \propto P(\theta) P(\mathbf{y}_A | \theta) P(\mathbf{y}_B | \theta, \gamma) \\
 & \propto \theta^{a_\theta-1} e^{-(b_\theta \theta)} \theta^{\sum_A y_A + \sum_B y_B} e^{-(n_A \theta + n_B \theta \gamma)} \\
 & \propto \theta^{[(a_\theta + \sum_A y_A + \sum_B y_B) - 1]} e^{-\theta(b_\theta + n_A + n_B \gamma)}
 \end{aligned}$$

So the form is

$$P(\theta|\mathbf{y}_A, \mathbf{y}_B, \gamma) \sim \Gamma(a_\theta + \sum_{i=1}^{58} y_{Ai} + \sum_{j=1}^{218} y_{Bj}, b_\theta + 58 + 218\gamma)$$

```
[ ]: y_A<-read.table('menchild30bach.txt')
      y_B<-read.table('menchild30nobach.txt')
      nA<-nrow(y_A)
      nB<-nrow(y_B)
      c(nA,nB)
```

1. 58 2. 218

(c)

$$\begin{aligned} P(\gamma|\mathbf{y}_A, \mathbf{y}_B, \theta) \\ &\propto P(\gamma)P(\mathbf{y}_B|\theta, \gamma) \\ &\propto \gamma^{a_\gamma-1} e^{-(b_\gamma\gamma)} \gamma^{\sum_B y_B} e^{-n_B\theta\gamma} \\ &\propto \gamma^{a_\gamma+\sum_{j=1}^{218} y_{Bj}-1} e^{-(b_\gamma+n_B\theta)\gamma} \end{aligned}$$

So the form is

$$P(\gamma|\mathbf{y}_A, \mathbf{y}_B, \theta) \sim \Gamma(a_\gamma + \sum_{j=1}^{218} y_{Bj}, b_\gamma + 218\theta)$$

(d)

The bigger parameters are, the smaller  $E[\theta_B - \theta_A|\mathbf{y}_A, \mathbf{y}_B]$  is.

```
[ ]: a_theta<-2
      b_theta<-1
      a_gamma<-c(8,16,32,64,128)
      b_gamma<-a_gamma
      sum_yA<-sum(y_A)
      sum_yB<-sum(y_B)
      n_ite<-5000
      par_initial<-1
      theta<-par_initial
      gamma<-par_initial
      Theta<-matrix(nrow=n_ite,ncol=length(a_gamma))
      Gamma<-Theta
      E<-NULL
      for (j in 1:length(a_gamma)){
        a_g<-a_gamma[j]
        b_g<-a_g
        for (i in 1:n_ite){
          theta<-rgamma(1,shape=a_theta+sum_yA+sum_yB,rate=b_theta+nA+nB*gamma)
          gamma<-rgamma(1,shape=a_g+sum_yB,rate=b_g+nB*theta)
          Theta[i,j]<-theta
          Gamma[i,j]<-gamma
        }
      }
```

```

    }
  }
  # E<-rowSums(Theta*Gamma-Theta)
  E<-colSums(Theta*Gamma-Theta)/n_ite
  E

```

1. 0.37764605403732   2. 0.334444596751819   3. 0.273652883952121   4. 0.20051808232011  
 5. 0.133719958562493

```

[ ]: par(bg='white')
for (i in 1:5){
  plot(Theta[,i],Gamma[,i])
}

```









