

Question 1: Multiple Choice (20 marks, 2 marks per question)

Answer Question 1 using the multiple choice form provided.

1. Which of the following statements about intercept in the single linear regression model is correct?
 - a. The intercept corresponds to the point where the regression line crosses the x-axis
 - b. If the model is correctly specified, the value of the intercept is always equal to zero
 - c. The intercept corresponds to the average value of outcome Y for the case in which the regressor X equals zero
 - d. Statements a. and c. are both correct
 - e. None of the above
2. In a regression model with two correlated regressors ($\rho_{X_1X_2} \neq 0$), if you exclude one of the regressors then...
 - a. It is no longer reasonable to assume that the errors are homoskedastic
 - b. The OLS estimator becomes biased
 - c. You are no longer controlling for the influence of the excluded variable
 - d. Statements b. and c. are both correct
 - e. None of the above applies
3. You estimate a single linear regression:

$$Y_i = \beta_0 + \beta_1 X_i + u_i$$

and obtain a 90% confidence interval for β_1 of $[-0.50, -0.10]$. Which of the following statements is necessarily true?

- a. The value of R^2 for the model is greater than zero
 - b. We reject the null that β_1 equals 0 in a two-sided test at the 1% level
 - c. The sample average of Y is equal to -0.30
 - d. The slope of the OLS regression line is positive
 - e. None of the above is necessarily true
4. The standard deviation of $\hat{\beta}_1$ increases...
 - a. With more observations in the sample
 - b. With greater dispersion of the regressor X
 - c. With greater dispersion of the error term u
 - d. With greater value of R^2
 - e. With none of the options above

5. What constitutes the dummy variable trap?
- a. Adjusted R^2 for our regression falls below zero
 - b. A linear combination of one or more dummies in our regression has no explanatory power
 - c. A linear combination of one or more dummies in our regression is equal to the constant
 - d. Exclusion of one or more relevant dummies from the regression causes omitted variable bias
 - e. None of the above

6. Consider the following regression model:

$$Y_i = \beta_0 + \beta_1 D_i + u_i$$

Assuming that the zero conditional mean assumption holds, which is the correct interpretation of β_0 ?

- a. It is the difference in the population means between group 1 ($D_i = 0$) and group 2 ($D_i = 1$) and group 2
 - b. It is the difference in the sample averages between group 1 ($D_i = 0$) and group 2 ($D_i = 1$) and group 2
 - c. It is the population mean for group 1 ($D_i = 0$)
 - d. It is the sample average for group 1 ($D_i = 0$)
 - e. None of the above
7. Which of these statements is a direct consequence of heteroskedasticity?
- a. The OLS estimators are biased
 - b. The OLS estimators are not BLUE
 - c. The OLS estimators do not have unique values
 - d. All of the above
 - e. None of the above
8. What problem does the classical measurement error in a regressor create for regression analysis?
- a. It causes regression coefficients to be smaller in magnitude than the population parameters
 - b. It causes regression coefficients to be larger in magnitude than the population parameters
 - c. It increases variance of the error term
 - d. It decreases variance of the error term
 - e. None of the above

9. Consider the estimation results from the following regression (standard errors are in brackets):

$$Y_i = 420.21 + 129.31X_{1i} - 30.34X_{2i} + 0.01X_{3i}; \quad \bar{R}^2 = 0.53, SER = 10.11$$

(12.702) (80.503) (7.929) (0.002)

What is the conclusion that can be drawn from these regression results?

- a. Variable X_3 can be omitted from the regression model since its OLS coefficient is close to zero
 - b. The model which omits variable X_3 is likely to have higher adjusted R^2
 - c. Omitting variable X_1 would lead to the steepest drop of adjusted R^2
 - d. The model is likely to be affected by omitted variable bias
 - e. None of the above
10. In order to evaluate whether Australia is becoming subject to more extreme heat-waves, you were given daily temperature data for years 1999 and 2019. Which descriptive statistic should you pay special attention to in your analysis of the yearly temperature distribution?
- a. 1st percentile
 - b. Variance
 - c. Skewness
 - d. Kurtosis
 - e. None of the above

Question 2: Short Answer Questions (20 marks)

Answer Questions 2-5 using exam booklets. You do not have to answer the questions in the order in which they are asked.

The answers to questions 2b, 2d, 2e and 2f should not be longer than three sentences.

- List the Least Squares assumptions corresponding to the multiple linear regression model. (3 marks)
- Explain the difference between the probability distribution and the cumulative probability distribution of a discrete random variable. (3 marks)
- Consider the estimation results from the following regression (standard errors are in brackets):

$$Y_i = \underset{(0.50)}{5.59} + \underset{(0.33)}{4.53}X_{1i} - \underset{(0.35)}{3.34}X_{2i}$$

and the following statistics:

$$n = 100, s_Y = 3.28, SER = 1.73$$

Using this information, compute R^2 , \bar{R}^2 , and the overall F-statistic corresponding to this regression. (5 marks)

- What is the consequence of independence of two random variables X and Y for the relationship between their joint and marginal probability distributions? (3 marks)
- What is the interpretation of the standard error of the regression? (3 marks)
- Explain the issue of simultaneous causality (3 marks)

Question 3: Malawi Trial (20 marks)

You are a researcher evaluating the results of a randomized control trial in Malawi which aims at identifying the effects of insecticide-treated bed nets on the incidence of malaria among Malawian families. The experiment involved 2000 families who were followed over the period of one year. The experimenters randomly distributed bed nets to half of the families and tracked the outbreaks of malaria in all families over the period of observation. The resulting dataset contains 2000 observations of two variables: $BN_i = 1$ if the family received a bed net and 0 otherwise; and $M_i = 1$ if the family members contracted Malaria within the period of observation and 0 otherwise.

Below is a table of descriptive statistics corresponding to the two groups and their outcomes

	Average contraction rate of malaria in the sample, \bar{M}	Sample standard deviation of malaria contractions, s_M
Families without bed nets	0.37	0.48
Families with bed nets	0.32	0.47

- Compute the 99% Confidence Interval for the population mean of malaria contraction rate among families with and without bed nets (4 marks)
- To evaluate whether the intervention influenced the incidence of malaria outbreaks, conduct a formal three-step hypothesis test using an appropriate null hypothesis and a two-sided alternative hypothesis. Formulate the null and the alternative hypotheses, list each step of the test, and decide what is the outcome of this test (present the formulas and calculations of the relevant statistics, use 1% significance level). (6 marks)
- The same evaluation can be also done using a single linear regression model. Write down the specification of this regression model and outline this alternative testing procedure. (4 marks)
- Do you expect the error terms corresponding to this particular regression model to be homoskedastic, or heteroskedastic? Define the concept of heteroskedasticity (in words or using a formula), answer the question, and motivate your answer. (6 marks)

Question 4: Smoking in Australia (20 Marks)

The Australian Institute for Health and Welfare (AIHW) has approached you to study the determinants of smoking in Australia. You have been given a dataset which contains responses to a telephone survey which AIHW has been collecting over the last five years. Each survey respondent has been interviewed once, and the dataset contains the following information for a total sample of $n = 2421$ respondents:

$cigs_i$: number of cigarettes smoked by individual i per week

$educ_i$: number of years of education of individual i

$cigpric_i$: local price of cigarettes per pack (at the time of the interview)

$lcigpric_i$: natural logarithm of price of cigarettes per pack.

$cauca_i$: dummy variable equaling 1 if individual i 's race is Caucasian

age_i : age of individual i in years

$agesq_i$: age of individual i in years, squared

$income_i$: annual income of individual i in AUD\$

$lincome_i$: natural logarithm of $income_i$

The AIHW's leadership wants to know two things - first, they want to understand which personal attributes are associated with higher incidence of smoking. Second, they want to know whether people smoke less if they are facing higher prices of cigarettes. Figure 1 on the next page produces summary statistics for these data, and regression output from R for three different regressions that explore the two associations (the dependent variable in these regressions is $cigs_i$). Based on this output, answer the following questions. Throughout assume a 5% level of confidence for assessing statistical significance.

- Suppose you ranked all people in the dataset by their income. What is the nominal value of income for the person who is in the exact middle of this ranking? (2 marks)
- Provide an interpretation of the regression coefficient estimate on $lincome_i$ in Regression 1. Comment on its sign, magnitude, and statistical significance (3 marks)
- The regression coefficient on age_i changes substantially between Regressions 1 and 2. Carefully explain what drives this large change. (4 marks)
- Interpret the sign, magnitude and statistical significance of the regression coefficient estimate on $cauca_i$ in Regression 3. (3 marks)
- Based on Regression 3, quantify the expected cigarette consumption among 17-year-olds with median values of each of income, race, and years of education who are subject to the median cigarette price. Work with the OLS coefficients rounded to the third decimal. At which age will the adult respondents with the same characteristics reach the same expected cigarette consumption as the 17-year-olds? (5 marks)
- What is your response to the AIHW's second question? Is there an association between cigarette prices and demand for cigarettes? Would you say that this regression captures the causal effect of cigarette prices on cigarette demand? Motivate your answer! (3 marks)

Figure 1: Summary Statistics and Estimation Results for the Cigarette Regressions

```

                                q4_cig_output.txt
educ          cigpric          cauca          age          income
Min.   : 6.00   Min.   :22.00   Min.   :0.0000   Min.   :17.00   Min.   : 2000
1st Qu.:10.00   1st Qu.:29.07   1st Qu.:1.0000   1st Qu.:28.00   1st Qu.: 50000
Median :12.00   Median :30.53   Median :1.0000   Median :38.00   Median : 80000
Mean   :12.47   Mean   :30.15   Mean   :0.8786   Mean   :41.24   Mean   : 77219
3rd Qu.:13.50   3rd Qu.:31.59   3rd Qu.:1.0000   3rd Qu.:54.00   3rd Qu.:120000
Max.   :18.00   Max.   :35.06   Max.   :1.0000   Max.   :88.00   Max.   :120000

cigs          lincome          agesq          lcigpric
Min.   : 0.000   Min.   : 7.601   Min.   : 289   Min.   :3.091
1st Qu.: 0.000   1st Qu.:10.820   1st Qu.: 784   1st Qu.:3.370
Median : 0.000   Median :11.290   Median :1444   Median :3.419
Mean   : 8.686   Mean   :11.074   Mean   :1990   Mean   :3.403
3rd Qu.:20.000   3rd Qu.:11.695   3rd Qu.:2916   3rd Qu.:3.453
Max.   :80.000   Max.   :11.695   Max.   :7744   Max.   :3.557

```

Regression 1

t test of coefficients:

```

              Estimate Std. Error t value Pr(>|t|)
(Intercept) -3.746343   3.605091 -1.0392 0.2988244
age          -0.031315   0.013334 -2.3485 0.0189313 *
lincome       1.239360   0.324421  3.8202 0.0001367 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

Regression 2

t test of coefficients:

```

              Estimate Std. Error t value Pr(>|t|)
(Intercept) -6.38273256  3.61262217 -1.7668 0.07739 .
age          0.72575406   0.08092398  8.9683 < 2e-16 ***
agesq       -0.00833493   0.00085303 -9.7710 < 2e-16 ***
lincome      0.15607113   0.33058031  0.4721 0.63689
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

Regression 3

t test of coefficients:

```

              Estimate Std. Error t value Pr(>|t|)
(Intercept)  2.71157331 12.59356936  0.2153 0.82954
age          0.78208284  0.08058579  9.7050 < 2.2e-16 ***
agesq       -0.00912285  0.00085147 -10.7143 < 2.2e-16 ***
lincome      0.75352761  0.34595379  2.1781 0.02949 *
cauca       -0.20488617  0.79582724 -0.2575 0.79685
educ        -0.51430203  0.09373105 -5.4870 4.515e-08 ***
lcigpric    -2.90086752  3.46614254 -0.8369 0.40272
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

Question 5: Forecasting crypto currency prices (20 Marks)

Westpac bank has hired you to develop time series models for crypto currency prices. They provide you with a dataset that has the daily prices in dollars for two crypto currencies – Bitcoin $bitcoin_t$ and Ethereum eth_t – over the last year and thus a total of $T = 365$ observations. For all parts of this question, only conduct hypothesis tests based on regressions with heteroskedasticity-robust standard errors.

- a Westpac asks you to compute the elasticity of Ethereum with respect to Bitcoin because they want to know by how much the Ethereum price changes in percentage terms on a given day for a 1% change in the Bitcoin price that day. A Westpac analyst speculates that the elasticity is 1 and Westpac wants you to test this hypothesis.

Starting from the raw data in `crypto_daily.csv`, write down the pseudo-code in R you would develop to estimate this elasticity and test this hypothesis assuming a 5% level of significance. List the steps you would include in your code, and if it helps in describing your answer, you may state explicit R code though this is not necessary for obtaining full marks. Be precise and explicitly describe any variable transformation you would perform, regressions run in your code, hypothesis tests required, test statistics used, or any other calculations necessary. (4 marks)

- b Figure 2 presents time series plots of the Ethereum prices eth_t and their first difference $\Delta eth_t = eth_t - eth_{t-1}$. What does it tell you about the stationarity of Ethereum prices and their daily changes? (4 marks)
- c To estimate the relationship between Ethereum and Bitcoin prices you estimate the following ADL(1,2) model:

$$eth_t = \beta_0 + \beta_1 eth_{t-1} + \beta_2 bitcoin_{t-1} + \beta_3 bitcoin_{t-2} + u_t$$

The regression results are reported in Figure 3 below. How many observations are used in estimating this model? Briefly explain why this is the number of observations used in estimation. (2 marks)

- d Interpret the regression coefficients on 1st and 2nd lags of the Bitcoin price. Comment on their statistical significance at the 5% level. (4 marks)
- e The Ethereum price today is \$220.59, the Bitcoin price today is \$6,517.18, and was \$6,281.2 yesterday and \$6,371.3 the day before yesterday. Based on the ADL(1,2) model from Figure 3, what is your forecast for the Ethereum price tomorrow? Compute the 95% forecast interval assuming IID normal errors in the regression equation. Round to 3 digits after the decimal in conducting your calculations. (4 marks)
- f The results of 3 tests are reported in Figure 4 based on the ADL(1,2) model from question c. Choose the appropriate set of results and describe the results of the Granger Causality test determining whether $bitcoin_t$ Granger Causes eth_t . Assume a 5% level of significance. (2 marks)

Figure 2: Time Series Plots of Ethereum Prices and Changes of Ethereum Prices by Day

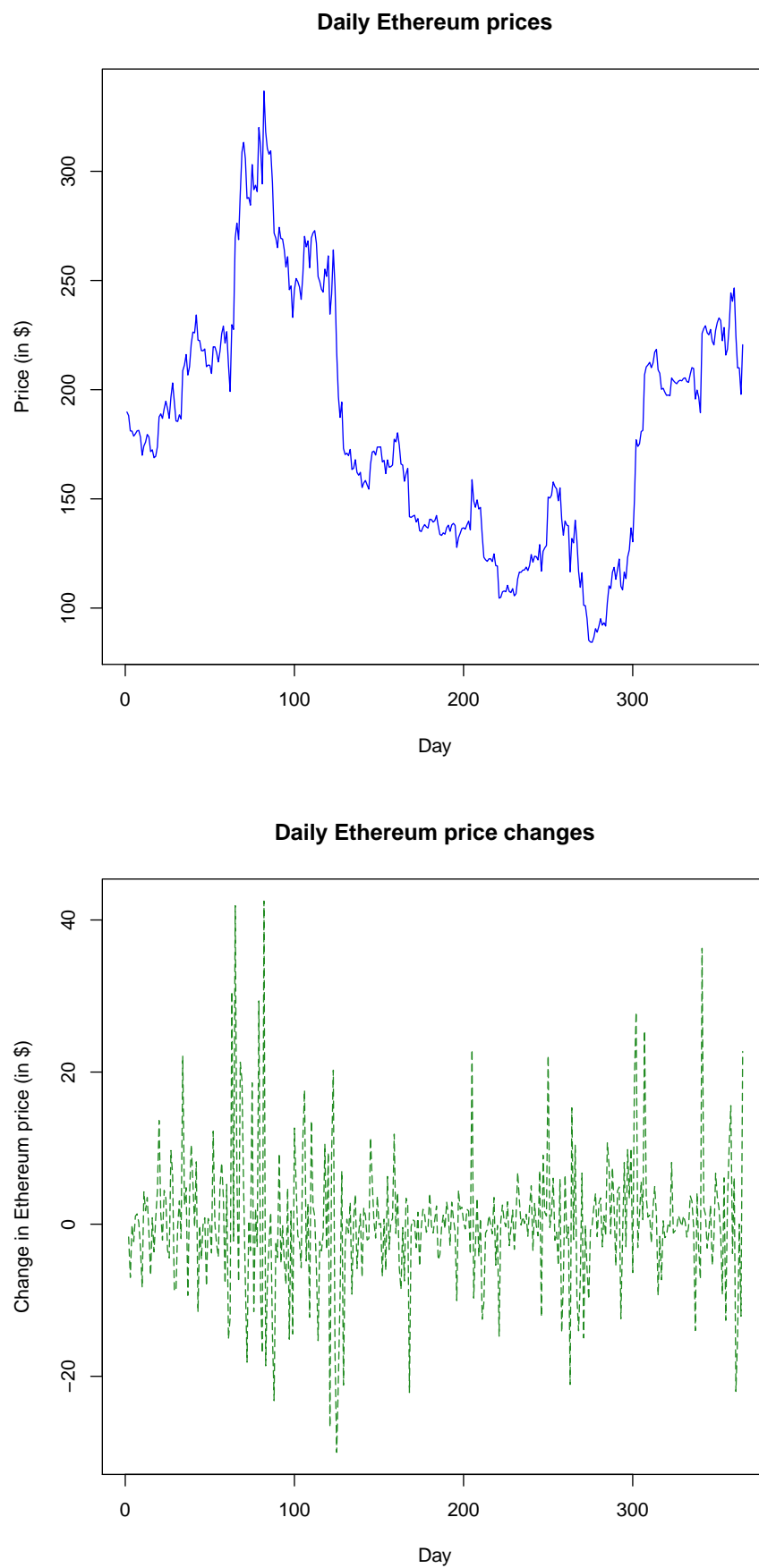


Figure 3: Ethereum price regression results

```
> reg1=lm(eth~eth_lag1+bitcoin_lag1+bitcoin_lag2,data=mydata)
> coeftest(reg1, vcov = vcovHC(reg1, "HC1"))

t test of coefficients:

              Estimate Std. Error t value Pr(>|t|)
(Intercept)   2.701750   1.619441   1.668   0.0961 .
eth_lag1       0.965649   0.014392  67.095 <2e-16 ***
bitcoin_lag1  -0.002552   0.001571  -1.625   0.1051
bitcoin_lag2   0.003105   0.001556   1.995   0.0468 *
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

> summary(reg1)
# (note: regression output for reg1 with homoscedastic standard
# errors are omitted)

Residual standard error: 8.929 on 359 degrees of freedom
(2 observations deleted due to missingness)
Multiple R-squared:  0.9747, Adjusted R-squared:  0.9745
F-statistic: 4608 on 3 and 359 DF, p-value: < 2.2e-16
```

Figure 4: Ethereum price regression tests

```
> linearHypothesis(reg1,c("eth_lag1=0"),vcov = vcovHC(reg1, "HC1"))
Linear hypothesis test

Hypothesis:
eth_lag1 = 0

Model 1: restricted model
Model 2: eth ~ eth_lag1 + bitcoin_lag1 + bitcoin_lag2

Note: Coefficient covariance matrix supplied.

   Res.Df Df      F    Pr(>F)
1      360
2      359  1 4530.4 < 2.2e-16 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

> linearHypothesis(reg1,c("eth_lag1=0","bitcoin_lag1=0"),vcov = vcovHC(reg1, "HC1"))
Linear hypothesis test

Hypothesis:
eth_lag1 = 0
bitcoin_lag1 = 0

Model 1: restricted model
Model 2: eth ~ eth_lag1 + bitcoin_lag1 + bitcoin_lag2

Note: Coefficient covariance matrix supplied.

   Res.Df Df      F    Pr(>F)
1      361
2      359  2 2270.3 < 2.2e-16 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

> linearHypothesis(reg1,c("bitcoin_lag1=0","bitcoin_lag2=0"),vcov = vcovHC(reg1, "HC1"))
Linear hypothesis test

Hypothesis:
bitcoin_lag1 = 0
bitcoin_lag2 = 0

Model 1: restricted model
Model 2: eth ~ eth_lag1 + bitcoin_lag1 + bitcoin_lag2

Note: Coefficient covariance matrix supplied.

   Res.Df Df      F Pr(>F)
1      361
2      359  2 2.4573 0.0871 .
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

END OF EXAMINATION

Formula Sheet

Expected Values, Variances, Correlation

$$E(c) = c$$

$$E(cx) = cE(x)$$

$$E(a + cx) = a + cE(x)$$

$$E(x + y) = E(x) + E(y)$$

$$E(c_1x + c_2y) = c_1E(x) + c_2E(y)$$

$$\text{var}(x) = \sigma^2 = E[(x - E(x))^2]$$

$$\text{std}(x) = \sigma = \sqrt{E[(x - E(x))^2]}$$

$$\text{var}(a + cx) = c^2\text{var}(x)$$

$$\text{cov}(x, y) = E[(x - E(x))(y - E(y))]$$

$$\text{corr}(x, y) = \rho = \frac{\text{cov}(x, y)}{\sqrt{\text{var}(x)\text{var}(y)}}$$

$$P(y = y_1 | x = x_1) = \frac{P(x=x_1, y=y_1)}{p(x=x_1)}$$

$$\bar{y} = \frac{\sum_{i=1}^n y_i}{n}$$

$$\text{var}(\bar{Y}) = \frac{\sigma_Y^2}{n}$$

$$\text{std}(\bar{Y}) = \frac{\sigma_Y}{\sqrt{n}}$$

$$s_y^2 = \frac{1}{n-1} \sum_{i=1}^N (y_i - \bar{y})^2$$

$$s_y = \sqrt{\frac{1}{n-1} \sum_{i=1}^N (y_i - \bar{y})^2}$$

$$SE(\bar{y}) = \frac{s_y}{\sqrt{n}}$$

$$s_{xy} = \frac{1}{n-1} \sum_{i=1}^n \sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})$$

$$r_{xy} = \frac{s_{xy}}{s_x s_y}$$

Logarithms

$$x = \ln(e^x)$$

$$\frac{d \ln(x)}{dx} = \frac{1}{x}$$

$$\ln(1/x) = -\ln(x)$$

$$\ln(ax) = \ln(a) + \ln(x)$$

$$\ln(x/a) = \ln(x) - \ln(a)$$

$$\ln(x^a) = a \ln(x)$$

$$\ln(x + \Delta x) - \ln(x) \approx \frac{\Delta x}{x} \text{ (approximately equal for small } \Delta x \text{)}$$

Quadratic Formula

The solution to the quadratic equation:

$$ax + bx^2 + c = 0$$

where a , b , and c are constants can be computed by the quadratic formula:

$$x = \frac{-b \pm \sqrt{b^2 - 4ac}}{2a}$$

Calculus

x^* that maximizes (minimizes) a strictly concave (convex) function, $f(x)$, solves $\frac{df(x)}{dx} = 0$

OLS Estimator for Single Linear Regression

$$\begin{aligned}\hat{\beta}_1 &= \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{s_{XY}}{s_X^2} \\ \hat{\beta}_0 &= \bar{Y} - \hat{\beta}_1 \bar{X} \\ \sigma_{\hat{\beta}_1}^2 &= \frac{1}{n} \frac{\text{var}((X_i - \mu_X)u_i)}{(\text{var}(X_i))^2} \\ \sigma_{\hat{\beta}_0}^2 &= \frac{1}{n} \frac{\text{var}(H_i u_i)}{(E(H_i^2))^2}; \text{ where } H_i = 1 - (\frac{\mu_X}{E(X_i^2)})X_i \\ \hat{\beta}_1 &\rightarrow \beta_1 + \rho_{Xu} \frac{\sigma_u}{\sigma_X}\end{aligned}$$

Testing Differences in Means

$$\begin{aligned}H_0 : \mu_w - \mu_m &= d_0; \text{ vs. } H_1 : \mu_w - \mu_m \neq d_0 \\ SE(\bar{Y}_w - \bar{Y}_m) &= \sqrt{s_w^2/n_w + s_m^2/n_m} \\ t^{act} &= \frac{(\bar{Y}_w - \bar{Y}_m) - d_0}{SE(\bar{Y}_w - \bar{Y}_m)}\end{aligned}$$

Single Hypothesis Testing in Regression Models

$$\begin{aligned}t^{act} &= \frac{\hat{\beta}_1 - \beta_{1,0}}{SE(\hat{\beta}_1)} \\ H_0 : \beta_1 &= \beta_{1,0} \text{ vs. } H_1 : \beta_1 \neq \beta_{1,0}, \text{ p-value} = 2\Phi(-|t^{act}|) \\ H_0 : \beta_1 &= \beta_{1,0} \text{ vs. } H_1 : \beta_1 < \beta_{1,0}, \text{ p-value} = \Phi(t^{act}) \\ H_0 : \beta_1 &= \beta_{1,0} \text{ vs. } H_1 : \beta_1 > \beta_{1,0}, \text{ p-value} = 1 - \Phi(t^{act}) \\ t^\alpha &\text{ is the critical value for a two-sided test with } \alpha \text{ significance level} \\ \alpha &= 2\Phi(-|t^\alpha|) \\ (1 - \alpha) \text{ CI: } &[\hat{\beta}_1 - t^\alpha SE(\hat{\beta}_1), \hat{\beta}_1 + t^\alpha SE(\hat{\beta}_1)] \\ \text{For testing means, replace } \beta &\text{ with } \mu_X \text{ and } \hat{\beta} \text{ with } \bar{X}\end{aligned}$$

Joint Hypothesis Testing in Regression Models

$H_0 : \beta_j = \beta_{j,0}, \beta_m = \beta_{m,0}, \dots$ for a total of q restrictions

H_1 : one or more of the q restrictions under H_0 does not hold

F-statistic for the test is distributed $F_{q,n-k-1}$

$p\text{-value} = \Pr[F_{q,n-k-1} > F^{act}] = 1 - G(F^{act}; q, n - k - 1)$

For $q = 2$ restrictions, relationship between F-statistic and individual t-statistics for testing coefficients jointly equal 0:

$$F = \frac{1}{2} \left(\frac{(t_1^{act})^2 + (t_2^{act})^2 - 2\hat{\rho}_{t_1^{act}, t_2^{act}} t_1^{act} t_2^{act}}{1 - \hat{\rho}_{t_1^{act}, t_2^{act}}^2} \right)$$

Homoskedasticity-Only F-statistic with q restrictions

$$F^{act} = \frac{(SSR_{restricted} - SSR_{unrestricted})/q}{SSR_{unrestricted}/(n - k - 1)} = \frac{(R_{unrestricted}^2 - R_{restricted}^2)/q}{(1 - R_{unrestricted}^2)/(n - k - 1)}$$

$F_{1,n-k-1}^{act} = (t^{act})^2$ if $q = 1$ restriction

Goodness of Fit in Regression Models

$$SSR = \sum_{i=1}^n \hat{u}_i^2; \quad ESS = \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2$$

$$TSS = \sum_{i=1}^n (Y_i - \bar{Y})^2 = (n - 1)s_Y^2$$

$$R^2 = \frac{ESS}{TSS} = 1 - \frac{SSR}{TSS}$$

$$SER = s_{\hat{u}} = \sqrt{s_{\hat{u}}^2}, \quad s_{\hat{u}}^2 = \frac{SSR}{n-k-1}$$

$$\bar{R}^2 = 1 - \frac{n-1}{n-k-1} \frac{SSR}{TSS} = 1 - \frac{s_{\hat{u}}^2}{s_Y^2}$$

Nonlinear Regression Partial Effects, Standard Errors, and CIs

$$E[Y|X_1, X_2, \dots, X_k] = f(X_1, X_2, \dots, X_k)$$

$$\Delta \hat{Y} = \hat{f}(X_1 + \Delta X_1, X_2, \dots, X_k) - \hat{f}(X_1, X_2, \dots, X_k); \quad SE(\Delta \hat{Y}) = \frac{|\Delta \hat{Y}|}{\sqrt{F}}$$

$$(1 - \alpha) \text{ CI: } [\Delta \hat{Y} - t^\alpha SE(\Delta \hat{Y}), \Delta \hat{Y} + t^\alpha SE(\Delta \hat{Y})]$$

Time Series Regression

$$\text{RMSFE} = \sqrt{E[(Y_{T+1} - \hat{Y}_{T+1|T})^2]}$$

$$SE(Y_{T+1} - \hat{Y}_{T+1|T}) = \widehat{RMSE} = \sqrt{\text{var}(\hat{u}_t)} = SER$$

$$(1 - \alpha) \text{ CI: } [\hat{Y}_{T+1|T} - t^\alpha \times SE(Y_{T+1} - \hat{Y}_{T+1|T}), \hat{Y}_{T+1|T} + t^\alpha \times SE(Y_{T+1} - \hat{Y}_{T+1|T})]$$

$$\text{BIC}(K) = \ln \left[\frac{SSR(K)}{T} \right] + K \frac{\ln(T)}{T}$$

$$\text{AIC}(K) = \ln \left[\frac{SSR(K)}{T} \right] + K \frac{2}{T}$$

F-statistic for the Granger Causality test has q and $T - \ell - p - 1$ degrees of freedom, where q is the number of restrictions imposed under the null, T is sample size, ℓ is the maximum lag length included in the time series model, p is the number of parameters in the time series model excluding the constant.

Statistical Distribution Tables

Critical Values of the t Distribution

<i>Significance Level</i>						
	<i>1- Tailed:</i>	<i>.10</i>	<i>.05</i>	<i>.025</i>	<i>.01</i>	<i>.005</i>
	<i>2- Tailed:</i>	<i>.20</i>	<i>.10</i>	<i>.05</i>	<i>.02</i>	<i>.01</i>
<i>Degrees of Freedom</i>	1	3.078	6.314	12.706	31.821	63.657
	2	1.886	2.920	4.303	6.965	9.925
	3	1.638	2.353	3.182	4.541	5.841
	4	1.533	2.132	2.776	3.747	4.604
	5	1.476	2.015	2.571	3.365	4.032
	6	1.440	1.943	2.447	3.143	3.707
	7	1.415	1.895	2.365	2.998	3.499
	8	1.397	1.860	2.306	2.896	3.355
	9	1.383	1.833	2.262	2.821	3.250
	10	1.372	1.812	2.228	2.764	3.169
	11	1.363	1.796	2.201	2.718	3.106
	12	1.356	1.782	2.179	2.681	3.055
	13	1.350	1.771	2.160	2.650	3.012
	14	1.345	1.761	2.145	2.624	2.977
	15	1.341	1.753	2.131	2.602	2.947
	16	1.337	1.746	2.120	2.583	2.921
	17	1.333	1.740	2.110	2.567	2.898
	18	1.330	1.734	2.101	2.552	2.878
	19	1.328	1.729	2.093	2.539	2.861
	20	1.325	1.725	2.086	2.528	2.845
	21	1.323	1.721	2.080	2.518	2.831
	22	1.321	1.717	2.074	2.508	2.819
	23	1.319	1.714	2.069	2.500	2.807
	24	1.318	1.711	2.064	2.492	2.797
	25	1.316	1.708	2.060	2.485	2.787
	26	1.315	1.706	2.056	2.479	2.779
	27	1.314	1.703	2.052	2.473	2.771
	28	1.313	1.701	2.048	2.467	2.763
	29	1.311	1.699	2.045	2.462	2.756
	30	1.310	1.697	2.042	2.457	2.750
	35	1.306	1.690	2.030	2.438	2.724
	36	1.306	1.688	2.028	2.434	2.719
	37	1.305	1.687	2.026	2.431	2.715
	38	1.304	1.686	2.024	2.429	2.712
	39	1.304	1.685	2.023	2.426	2.708
	40	1.303	1.684	2.021	2.423	2.704
	60	1.296	1.671	2.000	2.390	2.660
	90	1.291	1.662	1.987	2.368	2.632
	120	1.289	1.658	1.980	2.358	2.617
	∞	1.282	1.645	1.960	2.326	2.576

95th Percentile for the F-distribution F_{v_1, v_2}

		Numerator v_1											
D e n o m i n a t o r v_2	v_2/v_1	1	2	3	4	5	7	9	10	15	20	60	∞
	1	161.45	199.50	215.71	224.58	230.16	236.77	240.54	241.88	245.95	248.01	252.2	254.31
	2	18.51	19.00	19.16	19.25	19.30	19.35	19.41	19.40	19.43	19.45	19.48	19.50
	3	10.13	9.55	9.28	9.12	9.01	8.89	8.81	8.79	8.70	8.66	8.57	8.53
	4	7.71	6.94	6.59	6.39	6.26	6.09	6.00	5.96	5.86	5.80	5.69	5.63
	5	6.61	5.79	5.41	5.19	5.05	4.88	4.77	4.74	4.62	4.56	4.43	4.37
	6	5.99	5.14	4.76	4.53	4.39	4.21	4.10	4.06	3.94	3.87	3.74	3.67
	7	5.59	4.74	4.35	4.12	3.97	3.79	3.68	3.64	3.51	3.44	3.30	3.23
	8	5.32	4.46	4.07	3.84	3.69	3.50	3.39	3.35	3.22	3.15	3.01	2.93
	9	5.12	4.26	3.86	3.63	3.48	3.29	3.18	3.14	3.01	2.94	2.79	2.71
	10	4.96	4.10	3.71	3.48	3.33	3.14	3.02	2.98	2.85	2.77	2.62	2.54
	15	4.54	3.68	3.29	3.06	2.90	2.71	2.59	2.54	2.40	2.33	2.16	2.07
	20	4.35	3.49	3.10	2.87	2.71	2.51	2.39	2.35	2.20	2.12	1.92	1.84
	30	4.17	3.32	2.92	2.69	2.53	2.33	2.21	2.16	2.01	1.93	1.74	1.62
	40	4.08	3.23	2.84	2.61	2.45	2.25	2.12	2.08	1.92	1.84	1.64	1.51
	50	4.03	3.18	2.79	2.56	2.40	2.20	2.07	2.03	1.87	1.78	1.58	1.44
	60	4.00	3.15	2.76	2.53	2.37	2.17	2.04	1.99	1.84	1.75	1.53	1.39
	120	3.92	3.07	2.68	2.45	2.29	2.09	1.95	1.91	1.75	1.66	1.43	1.25
	∞	3.84	3.00	2.60	2.37	2.21	2.01	1.88	1.83	1.67	1.57	1.32	1.00

Critical Values for the Chi-Squared Distribution

Degrees of Freedom	Critical Values		
	1%	5%	10%
1	6.64	3.84	2.71
2	9.21	5.99	4.61
3	11.35	7.81	6.25
4	13.28	9.49	7.78
5	15.09	11.07	9.24
6	16.81	12.59	10.65
7	18.48	14.07	12.02
8	20.09	15.51	13.36
9	21.67	16.92	14.68
10	23.21	18.31	15.99
11	24.73	19.68	17.28
12	26.22	21.0	18.55
13	27.69	22.4	19.81
14	29.14	23.7	21.06
15	30.58	25.0	22.31
16	32.00	26.3	23.54
17	33.41	27.6	24.77
18	34.81	28.9	25.99
19	36.19	30.1	27.20
20	37.57	31.4	28.41