

MAST30025_2021_SM1 MAST30025 assignment 1

Michael Le

TOTAL POINTS

38 / 40

QUESTION 1 **5.5 2 / 2**

7 pts

5.6 2 / 2

1.1 5 / 5

5.7 3 / 3

1.2 1 / 2

QUESTION 2

2 4 / 4

QUESTION 3

3 4 / 4

QUESTION 4

10 pts

4.1 2 / 2

4.2 2 / 2

4.3 3 / 3

4.4 2 / 3

- >You must use theorem 3.5 directly. i.e. what is the dist. of x/σ^2 ?

QUESTION 5

15 pts

5.1 2 / 2

5.2 2 / 2

5.3 2 / 2

5.4 2 / 2

MAST30025: Linear Statistical Models

Assignment 1 S1 2021

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Question 1 Solution:

Part a:

$$A^2 = A^3$$

Suppose A is a square matrix is (real and) symmetric then its eigenvalues are all real, and its eigenvalues are orthogonal.

Theorem 2.3

Proof:

Take A to be a square matrix, n x n. First we diagonalise A,i.e., find P such that.

=

$$D = P^T AP$$

$$= \begin{bmatrix} \lambda_1 & \dots & 0 \\ \dots & \dots & \dots \\ 0 & \dots & \lambda_k \end{bmatrix}$$

where $\lambda_1, \lambda_2, \dots, \lambda_k$ are the eigenvalues of A.

Since P is orthogonal both P and P^T are non - singular,

$$r(P^T AP) = r(P^T A) = r(A)$$

Because $P^T AP$ is diagonal $r(P^T AP)$ is the number of non zero eigenvalues of A.

But we wanted to prove **Theorem 2.2**

that A any symmetric matrix is idempotent. Which has eigenvalues of $\lambda = 0$ or $\lambda = 1$.

The eigenvalues of idempotent matrices are always either

$$\lambda = 0 \text{ or } = 1.$$

$$A^2 = \lambda^2 x$$

Multiplying by A!!!

$$A^3 x = A^2 \lambda x = \lambda A^2 x = \lambda^3 x$$

$$(\lambda^3 - \lambda^2)x = 0$$

By definition, $x \neq 0$,

$$\begin{aligned}\lambda^3 - \lambda^2 &= 0 \\ \lambda^2(\lambda - 1) &= 0\end{aligned}$$

Therefore there are two values with eigenvalues of 0 and one eigenvalue of 1! satisfies this theorem that A is idempotent!

Part b:

$$A = A^3$$

$$A^3 x = A \lambda x = \lambda A x = \lambda^3 x$$

Using the same theorem from the previous it has eigenvalues of 0,1 and -1. Since we care that A has to be positive semi-definite. Which has an eigenvalue of -1. Which does not satisfy Theorem 2.2! A is not idempotent!

Question 2 Solution:

Theorem 2.4

There exists a matrix P which diagonalises A_1, \dots, A_m .

$$P^T A_i P = D_i$$

and

$$P^T A_j P = D_j$$

We take A_i and A_j to be $k \times k$ matrices first we diagonalizes A_i , A_j , i.e. find P such that,

$$D_i = P^T A_i P = \begin{bmatrix} \lambda_1 & \dots & 0 \\ \dots & \lambda_2 & \dots \\ 0 & \dots & \lambda_k \end{bmatrix}$$

for $i = 1, \dots, k$

$$D_j = P^T A_j P = \begin{bmatrix} \lambda_1 & \dots & 0 \\ \dots & \lambda_2 & \dots \\ 0 & \dots & \lambda_m \end{bmatrix}$$

for $j = 1, \dots, m$

Proof:

$$P^T A_i A_j P = (P^T A_i P)(P^T A_j P) = (P^T A_j P)(P^T A_i P) = P^T A_j A_i P$$

Pre-multiply by P and post-multiply by P^T to get $A_i A_j = A_j A_i$.

Question 3 Solution:

Pre Proof Using Theorem 2.3

For any matrix A

$$r(A) = r(A^T) = r(A^T A) = \text{tr}(A)$$

$$A = \begin{bmatrix} | & | & & | & & | \\ a_1 & a_2 & & a_p & & a_n \\ | & | & & | & & | \end{bmatrix}$$

Given A matrix with dimensions n x p with p independent columns.

Let $x_1, x_2, x_3, \dots, x_k$ the basis for column space of A.

Definition of basis every column vector of A is a linear combination of the column vectors of x.

$$a_1 = b_1 x_1 + b_2 x_2 + \dots + b_k x_k$$

Definition of linear combination

where b is scalar

$$B = \begin{bmatrix} -- b_1 -- \\ -- b_2 -- \\ \vdots \\ -- b_p -- \end{bmatrix}$$

$$\begin{bmatrix} | & | & & | & & | \\ a_1 & a_2 & & a_p & & a_n \\ | & | & & | & & | \end{bmatrix} = \begin{bmatrix} | & | & & | & & | \\ x_1 & x_2 & & x_p & & x_n \\ | & | & & | & & | \end{bmatrix} \begin{bmatrix} -- b_1 -- \\ -- b_2 -- \\ \vdots \\ -- b_p -- \end{bmatrix}$$

$$A = XB$$

$$A^T = (XB)^T = B^T X^T$$

$r(A) \leq r(A^T)$ or $r(A) \geq r(A^T)$ to satisfy!

$$r(A) = r(A^T) = r(A^T A) = \text{tr}(A) = p$$

Since P is orthogonal both P and P^T are non-singular. Therefore we need to sum up the diagonal elements of $P^T A P$, so we need to sum up its trace!

$$r(A) = r(P^T A P) = \text{tr}(P^T A P) = \text{tr}(P P^T A) = \text{tr}(A) = p$$

Because $D = P^T A P$ is diagonal $r(P^T A P)$ is the number of nonzero values of A!

But A is idempotent so its takes eigenvalues between 0 or 1. To Prove

Theorem 2.7! We need only the identity matrix to allow $A^T A$ to be positive definite!

Using Theorem 2.7

Proof (\Leftarrow):

We want $A^T A$ to be symmetric

and have all the eigenvalues to be strictly positive to prove $A^T A$ is a positive definite matrix!

we know $r(A^T A) = p$ is a $p \times p$ matrix so it has to be a full rank matrix, p !

Let, $\lambda_1, \lambda_2, \dots, \lambda_p > 0$ be the eigenvalues of $A^T A$ for every x and for each eigenvalue has to have a value of 1.

for $z = P^T x = (z_1, \dots, z_p)^T$

$$x^T (A^T A)x = x^T P D P^T x = z^T D z = \sum_{i=1}^p z_i^2 \lambda_i$$

since $\lambda_i = 1$!

$$= \sum_{i=1}^p z_i^2$$

> 0

Thus $A^T A$ is positive definite as required!

Proof (\Rightarrow):

Suppose $A^T A$ is positive definite let x_i be its normalised i-th eigenvector then,

$$x_i^T (A^T A)x_i = \lambda_i x_i^T x_i = \lambda_i$$

From theorem 2.3 we want $A^T A$ to be symmetric and idempotent. We want the eigenvalues to be 0 or 1. This case all of the eigenvalues must equal to 1.

$$\lambda_i = 1 > 0$$

So, the eigenvalues of $A^T A$ are strictly positive as required!!

Question 4 Solution:

Part a:

Given information:

Let,

$x_1, x_2, x_3 \sim (MVN(\mu, \sigma^2))$ be a sequence of independent normal random variables,

$$\mu = \bar{x} = \frac{x_1 + x_2 + x_3}{3}$$

$$\mathbf{x}^T = (x_1, x_2, x_3)^T$$

Supposed to be x^T as noted!

$$\mathbf{y} = (x_1 - \bar{x}, x_2 - \bar{x}, x_3 - \bar{x})^T$$

To solve A from:

$$\mathbf{y} = A\mathbf{x}$$

$$\begin{aligned} \mathbf{x} &\sim N(\mu \begin{bmatrix} \mu \\ \mu \\ \mu \end{bmatrix}, \begin{bmatrix} \sigma^2 & 0 & 0 \\ 0 & \sigma^2 & 0 \\ 0 & 0 & \sigma^2 \end{bmatrix}) = \\ \begin{bmatrix} x_1 - \bar{x} \\ x_2 - \bar{x} \\ x_3 - \bar{x} \end{bmatrix} &= A \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} \\ \bar{x} &= \frac{x_1 + x_2 + x_3}{3} \end{aligned}$$

$$= \frac{1}{3} [1 \ 1 \ 1] \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix}$$

$$= \frac{1}{3} [1 \ 1 \ 1] \mathbf{x}$$

Then,

$$\begin{aligned} \mathbf{y} &= \begin{bmatrix} x_1 - \bar{x} \\ x_2 - \bar{x} \\ x_3 - \bar{x} \end{bmatrix} = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} - \bar{x} \begin{bmatrix} 1 \\ 1 \\ 1 \end{bmatrix} = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} - \begin{bmatrix} \frac{1}{3}[1, 1, 1]x \\ \frac{1}{3}[1, 1, 1]x \\ \frac{1}{3}[1, 1, 1]x \end{bmatrix} = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} - \frac{1}{3} \begin{bmatrix} [1, 1, 1] \\ [1, 1, 1] \\ [1, 1, 1] \end{bmatrix} \mathbf{x} \\ &= \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} - \frac{1}{3} \begin{bmatrix} [1, 1, 1] \\ [1, 1, 1] \\ [1, 1, 1] \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} = \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} - \frac{1}{3} \begin{bmatrix} [1, 1, 1] \\ [1, 1, 1] \\ [1, 1, 1] \end{bmatrix} \begin{bmatrix} x_1 \\ x_2 \\ x_3 \end{bmatrix} = \mathbf{x} [\mathbf{I} - \frac{1}{3} \begin{bmatrix} [1, 1, 1] \\ [1, 1, 1] \\ [1, 1, 1] \end{bmatrix}] \\ \mathbf{A} &= [\mathbf{I} - \frac{1}{3} \begin{bmatrix} [1, 1, 1] \\ [1, 1, 1] \\ [1, 1, 1] \end{bmatrix}] = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix} - \frac{1}{3} \begin{bmatrix} [1, 1, 1] \\ [1, 1, 1] \\ [1, 1, 1] \end{bmatrix} = \frac{1}{3} \begin{bmatrix} 2 & -1 & -1 \\ -1 & 2 & -1 \\ -1 & -1 & 2 \end{bmatrix} \end{aligned}$$

Part b: Finding the rank of A

Proof: that there is a linear combination for any columns?

Where A is symmetric and idempotent!

$$A^2 = \frac{1}{9} \begin{bmatrix} 6 & -3 & -3 \\ -3 & 6 & -3 \\ -3 & -3 & 6 \end{bmatrix} = \frac{1}{3} \begin{bmatrix} 2 & -1 & -1 \\ -1 & 2 & -1 \\ -1 & -1 & 2 \end{bmatrix} = A$$

```
````{r}
```

```
A = matrix(c(2,-1,-1,-1,2,-1,-1,-1,2)/3,3,3)
```

```
````
```

```
 [,1]      [,2]      [,3]
[1,] 0.6666667 -0.3333333 -0.3333333
[2,] -0.3333333 0.6666667 -0.3333333
[3,] -0.3333333 -0.3333333 0.6666667
```

Finding rank of A

```
...{r}
rankMatrix(A)[1]
```

```
[1] 2
```

Each column all added up together gives us 0!.

$$x_1 + x_2 + x_3 = 0$$

Can be written as,

$$x_1 = -x_2 - x_3$$

That are linearly dependant and similar for x_2 and x_3 .

Using Theorem 2.3:

$$\text{Hence } r(A) = \text{tr}(A) = 2$$

Part c: Computing $E[y^T y]$

Using Theorem 3.5:

$$E[y^T A y] = \text{tr}(AV) + \mu^T A \mu$$

since $A = I$,

$$= \text{tr}(V) + \mu^T \mu$$

$$V = \text{var}y = \text{var}Ax = A \text{var}x A^T$$

since A is symmetric and idempotent!!

$$\text{var}(\mathbf{x}_i) = \begin{bmatrix} \sigma^2 & 0 & 0 \\ 0 & \sigma^2 & 0 \\ 0 & 0 & \sigma^2 \end{bmatrix}$$

$$\mathbf{V} = \frac{1}{3} \begin{bmatrix} 2\sigma^2 & -1 & -1 \\ -1 & 2\sigma^2 & -1 \\ -1 & -1 & 2\sigma^2 \end{bmatrix}$$

$$\mu = E[y] = E[Ax] = AE[x]$$

$$= \frac{1}{3} \begin{bmatrix} 2 & -1 & -1 \\ -1 & 2 & -1 \\ -1 & -1 & 2 \end{bmatrix} \begin{bmatrix} \mu \\ \mu \\ \mu \end{bmatrix} = 0$$

$$\mathbb{E}[y^T y] = \text{tr}\left(\frac{1}{3} \begin{bmatrix} 2\sigma^2 & -1 & -1 \\ -1 & 2\sigma^2 & -1 \\ -1 & -1 & 2\sigma^2 \end{bmatrix}\right) + 0$$

$$= 2\sigma^2$$

Part d:

Using Theorem 3.5:

Proof:

Assuming that A is idempotent and has rank k. Because it is symmetric, it can be diagonalised. Let the (orthogonal) diagonalising matrix be P.

$$\mathbf{D} = \mathbf{P}^T \mathbf{A} \mathbf{P} = \begin{bmatrix} \lambda_1 & \dots & 0 \\ \dots & \lambda_2 & \dots \\ 0 & \dots & \lambda_k \end{bmatrix}$$

since A is symmetric and idempotent, all eigenvalues are either 0 or 1. We know from definition:

$$\text{tr}(A) = r(A) = k$$

$$\mathbf{A} = \frac{1}{3} \begin{bmatrix} 2 & -1 & -1 \\ -1 & 2 & -1 \\ -1 & -1 & 2 \end{bmatrix}$$

$$\mathbf{A}^2 = \mathbf{A} = \frac{1}{3} \begin{bmatrix} 2 & -1 & -1 \\ -1 & 2 & -1 \\ -1 & -1 & 2 \end{bmatrix}$$

from Part 4b, we find out the rank and trace of matrix A we found in Part 4a. Is also is the same number of degrees of freedom for the chi squared distribution.

$$\text{tr}(A) = r(A) = 2$$

Therefore, A must have two eigenvalues of 1 and one eigenvalue of 0.

Using Theorem 3.5 and Corollary 3.7:

with our non central parameter λ !

$$\lambda = \frac{1}{2} \mu^T A \mu$$

$$= \frac{1}{2} [\mu \ \mu \ \mu] \frac{1}{3} \begin{bmatrix} 2 & -1 & -1 \\ -1 & 2 & -1 \\ -1 & -1 & 2 \end{bmatrix} \begin{bmatrix} \mu \\ \mu \\ \mu \end{bmatrix}$$

$$= 0$$

\iff if and only if

$$E[y] = E\left[\begin{bmatrix} x_1 - \mu \\ x_2 - \mu \\ x_3 - \mu \end{bmatrix}\right]$$

Since x_1, x_2 and x_3 is identically independently distributed! and taking the expectation of the expectation is the expectation itself!

$$E[y] = E\left[\begin{bmatrix} \mu - \mu \\ \mu - \mu \\ \mu - \mu \end{bmatrix}\right] = 0$$

NOTE: $\mu = \bar{x}$

In which case,

$$\frac{y^T y}{\sigma^2}$$

is just the sum of two independent standard normal's. This is just an ordinary (central) chi squared distribution χ_2^2 .

with expectation of 2 and variance of 4 with 2 degrees of freedom. In which A is symmetric and idempotent!

Question 5 Solution:

Part a: Computing y, X, β and ϵ

$$\mathbf{y} = \begin{bmatrix} 27.3 \\ 42.7. \\ 38.7 \\ 4.5 \\ 23.0 \\ 166.3 \\ 109.7 \\ 80.1 \\ 150.7 \\ 20.3 \\ 189.7 \\ 131.3 \\ 404.2 \\ 149 \end{bmatrix}, \quad \mathbf{X} = \begin{bmatrix} 1 & 13.1 \\ 1 & 15.3 \\ 1 & 25.8 \\ 1 & 1.8 \\ 1 & 4.9 \\ 1 & 55.4 \\ 1 & 39.3 \\ 1 & 26.7 \\ 1 & 47.5 \\ 1 & 6.6 \\ 1 & 94.7 \\ 1 & 61.1 \\ 1 & 135.6 \\ 1 & 47.6 \end{bmatrix}, \quad \boldsymbol{\beta} = \begin{bmatrix} \beta_0 \\ \beta_1 \end{bmatrix}, \quad \boldsymbol{\epsilon} = \begin{bmatrix} \epsilon_0 \\ \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \epsilon_4 \\ \epsilon_5 \\ \epsilon_6 \\ \epsilon_7 \\ \epsilon_8 \\ \epsilon_9 \\ \epsilon_{10} \\ \epsilon_{11} \\ \epsilon_{12} \\ \epsilon_{13} \end{bmatrix}$$

$y = X\beta + \epsilon$ becomes,

$$\begin{bmatrix} 27.3 \\ 42.7. \\ 38.7 \\ 4.5 \\ 23.0 \\ 166.3 \\ 109.7 \\ 80.1 \\ 150.7 \\ 20.3 \\ 189.7 \\ 131.3 \\ 404.2 \\ 149 \end{bmatrix} = \begin{bmatrix} 1 & 13.1 \\ 1 & 15.3 \\ 1 & 25.8 \\ 1 & 1.8 \\ 1 & 4.9 \\ 1 & 55.4 \\ 1 & 39.3 \\ 1 & 26.7 \\ 1 & 47.5 \\ 1 & 6.6 \\ 1 & 94.7 \\ 1 & 61.1 \\ 1 & 135.6 \\ 1 & 47.6 \end{bmatrix} \begin{bmatrix} \beta_0 \\ \beta_1 \end{bmatrix} + \begin{bmatrix} \epsilon_0 \\ \epsilon_1 \\ \epsilon_2 \\ \epsilon_3 \\ \epsilon_4 \\ \epsilon_5 \\ \epsilon_6 \\ \epsilon_7 \\ \epsilon_8 \\ \epsilon_9 \\ \epsilon_{10} \\ \epsilon_{11} \\ \epsilon_{12} \\ \epsilon_{13} \end{bmatrix}$$

Part b: Solving the least squares estimator

$$\hat{b} = (X^T X)^{-1} X^T y$$

```
{r}
b = solve(t(X) %*% X, t(X) %*% y)
b
```

```
[,1]
[1,] -1.233836
[2,] 2.701553
```

Part c:

$$s^2 = \frac{SS_{Res}}{n-p}$$

```
{r}
e = y - X %*% b
e #Residual errors
```

[,1]

[1,] -6.8565106

[2,] 2.3000724

[3,] -29.7662361

[4,] 0.8710405

[5,] 10.9962256

[6,] 17.8677893

[7,] 4.7627957

[8,] 9.2023660

[9,] 23.6100596

[10,] 3.7035852

[11,] -64.9032511

[12,] -32.5310639

[13,] ³ 39.1032233

[14,] 21.6399042

```
```{r}
n = 14 #sample size
p = 2 #number of parameters
SSRes = sum(e^z)
ssquared = SSRes/(n-p)
ssquared
```

```
[1] 777.1528
```

Part d:

$$t^T b = [1, 28]b = b_0 + 28b_1$$

```
{r}
c(1, 28) %*% b
```

```
[,1]
```

```
[1,] 74.40965
```

Part e:

$$z_i = \frac{e_i}{\sqrt{(s^2(1-H_{ii}))}}$$

```
* * * {r}
a = solve(t(X)%%X)
```

```
a
```

```
* * *
```

```
 [,1] [,2]
[1,] 0.163081936 -2.230009e-03
[2,] -0.002230009 5.425812e-05
```

```
* * * {r}
```

```
H = X%*%a%*%t(X)
```

```
H
```

```
* * *
```

```
```{r}
z = e/sqrt(ssquared * (1 - diag(H)))
z[13]
```

```

```
[1] 2.104999
```

Part f:

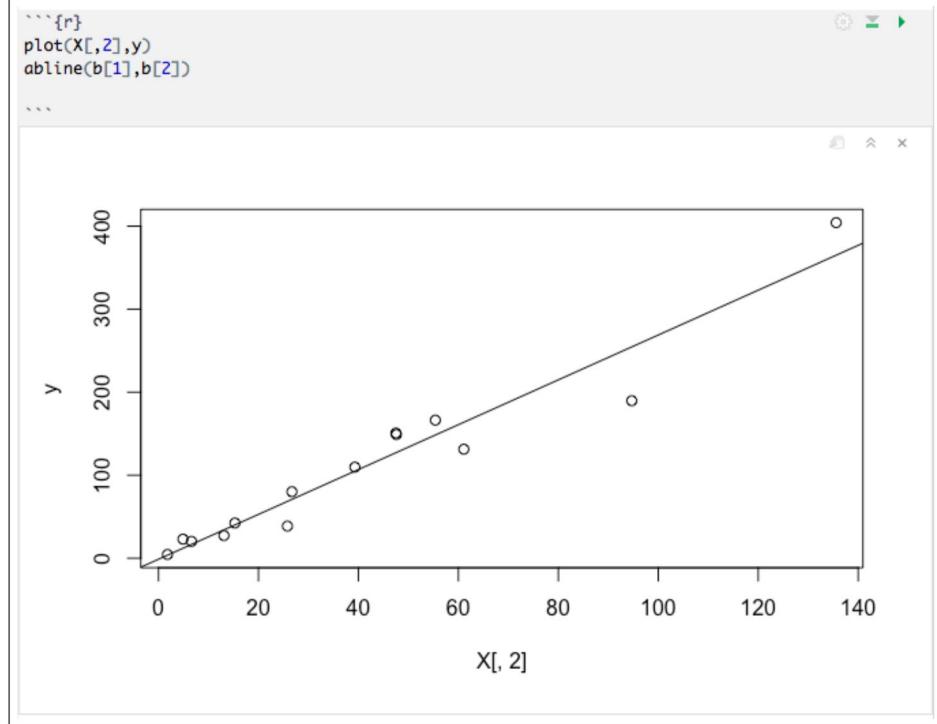
$$D_i = \frac{z_i^2 H_{ii}}{(k+1)(1-H_{ii})}$$

```
```{r}
k = 1
D = z^2 * (diag(H)/(1-diag(H))) * 1/(k+1)
D[13]
```

```

```
[1] 2.774008
```

Part g:



Full explanation: The Cook's distance certainly indicates it should be of some concern; however looking at the plot, it seems that the fit is actually okay. There is considerable evidence for heteroskedasticity — the variance increases with x (the design variable). Sea scallops has (by far) the largest x and so may be prone to a larger variance than the remaining points. The high Cook's distance therefore comes primarily from a very high leverage, rather than a bad fit to the model.

**END OF ASSIGNMENT!!**