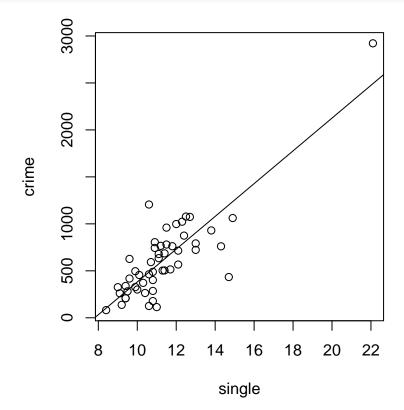
MAST20005/MAST90058: Week 6 Lab Solutions

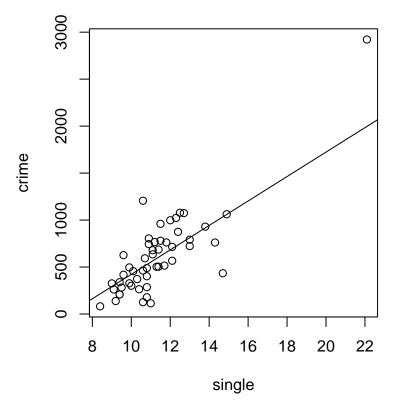
```
library(foreign)
cdata <- read.dta("crime.dta")</pre>
```

```
1. (a) m1 <- lm(crime ~ single, cdata)
    par(mar = c(4, 4, 1, 1)) # tighter figure margins
    plot(crime ~ single, cdata)
    abline(m1)</pre>
```



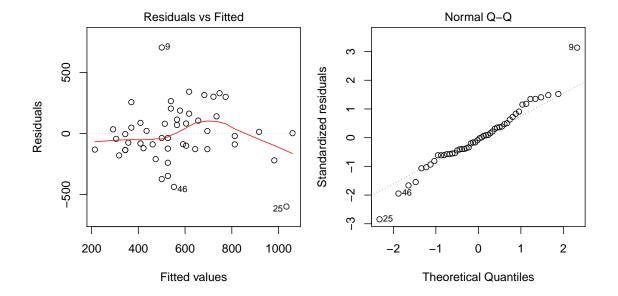
- (b) The point in the top-right corner of the plot. It corresponds to the state dc (Washington, D.C.).
- (c) An informal geometric argument: it seems like it might be quite influential because it is all by itself and will pull the best-fit line to be closer to itself while 'pivoting' it around the cluster of points in the bottom-left. (Indeed, this point does indeed have a strong influence on the parameter estimates. You will learn more about this in more advanced subjects that cover regression.)
- (d) We could do this by copying the data frame and removing the row corresponding to this outlying point. A neater way is to use the subset argument in lm():

```
m1a <- lm(crime ~ single, cdata, subset = -51)
par(mar = c(4, 4, 1, 1)) # tighter figure margins
plot(crime ~ single, cdata)
abline(m1a)</pre>
```



$$par(mfrow = c(1, 2), mar = c(4, 4, 2, 1))$$

 $plot(m1a, 1:2)$



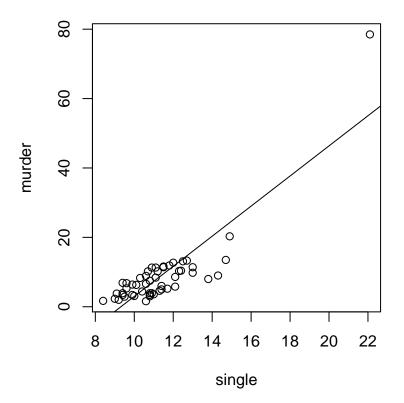
```
coef(summary(m1))
##
                 Estimate Std. Error
                                       t value
                                                    Pr(>|t|)
## (Intercept) -1362.5324 186.23306 -7.316276 2.150037e-09
                 174.4186
                            16.16796 10.787910 1.529137e-14
## single
coef(summary(m1a))
                Estimate Std. Error
                                      t value
                                                   Pr(>|t|)
## (Intercept) -878.8612 246.89537 -3.559650 8.491666e-04
## single
                130.1099
                           22.03329 5.905152 3.498831e-07
confint(m1)
##
                    2.5 %
                             97.5 %
## (Intercept) -1736.7818 -988.2831
## single
                 141.9278 206.9093
confint(m1a)
                     2.5 %
                              97.5 %
## (Intercept) -1375.27761 -382.4448
## single
                  85.80902
                           174.4108
```

The estimates have shifted substantially!

Which model we prefer to use will depend on to what extent we think that dc is representative or different to the other states. Given its special status within the USA and its very small size, there is a good argument for excluding it.

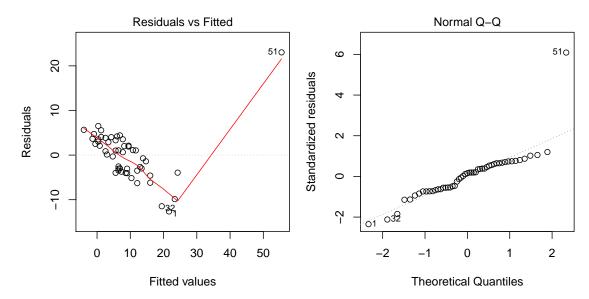
2. (a) m2 <- lm(murder ~ single, cdata)

```
(b) par(mar = c(4, 4, 1, 1)) # tighter figure margins
plot(murder ~ single, cdata)
abline(m2)
```



$$par(mfrow = c(1, 2), mar = c(4, 4, 2, 1))$$

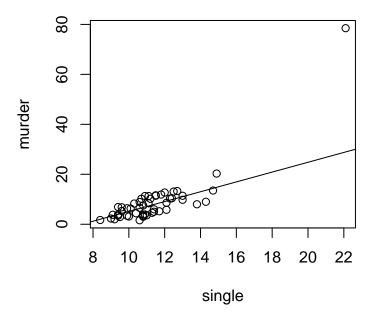
 $plot(m2, 1:2)$



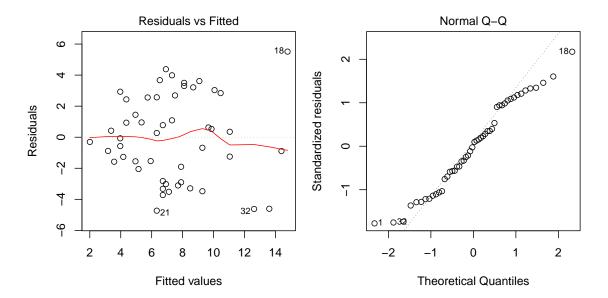
Outlying point (dc) is causing a poor model fit.

(c) Removing dc on the same basis as before:

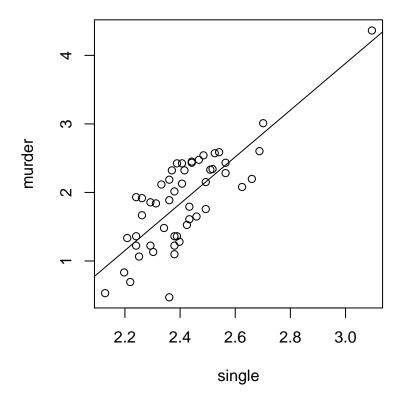
```
m2a <- lm(murder ~ single, cdata, subset = -51)
plot(murder ~ single, cdata)
abline(m2a)</pre>
```



```
par(mfrow = c(1, 2), mar = c(4, 4, 2, 1))
plot(m2a, 1:2)
```

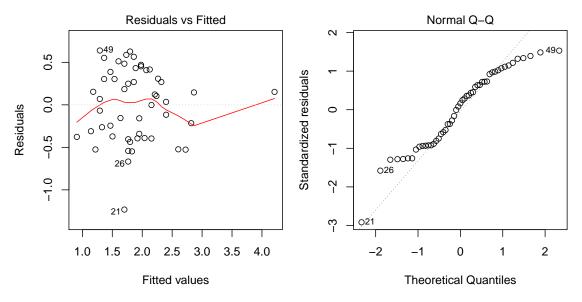


```
coef(summary(m2))
                  Estimate Std. Error t value
   ##
   ## (Intercept) -40.41533 4.257303 -9.493177 1.099312e-12
                   4.33913 0.369601 11.740041 7.536045e-16
   ## single
   coef(summary(m2a))
                   Estimate Std. Error t value Pr(>|t|)
   ## (Intercept) -14.514349 2.994406 -4.847155 1.354009e-05
   ## single
                   confint(m2)
   ##
                      2.5 %
                                97.5 %
   ## (Intercept) -48.970698 -31.859958
   ## single
                   3.596389 5.081871
   confint(m2a)
   ##
                      2.5 %
                              97.5 %
   ## (Intercept) -20.535005 -8.493693
              1.429076 2.503660
   ## single
(d) data12 <- data.frame(single = 12)
   predict(m2a, newdata = data12, interval = "prediction", level = 0.9)
   ##
            fit
                    lwr
                             upr
   ## 1 9.082068 4.39098 13.77316
(e) data8 <- data.frame(single = 8)
   predict(m2a, newdata = data8, interval = "prediction", level = 0.9)
   ##
             fit
                      lwr
                               upr
   ## 1 1.216595 -3.660916 6.094107
   Negative estimates (lower endpoint of interval)!
(f) cdata2 <- log(cdata[, c("murder", "single")])</pre>
   str(cdata2)
   ## 'data.frame': 51 obs. of 2 variables:
   ## $ murder: num 2.2 2.45 2.32 2.15 2.57 ...
   ## $ single: num 2.66 2.44 2.37 2.49 2.53 ...
   m3 <- lm(murder ~ single, cdata2)
   par(mar = c(4, 4, 1, 1)) # tighter figure margins
   plot(murder ~ single, cdata2)
   abline(m3)
```



```
par(mfrow = c(1, 2), mar = c(4, 4, 2, 1))

plot(m3, 1:2)
```



```
data3 <- data.frame(single = log(8))
predint <- predict(m3, newdata = data3, interval = "prediction", level = 0.9)
predint

## fit lwr upr
## 1 0.741214 -0.01148196 1.49391

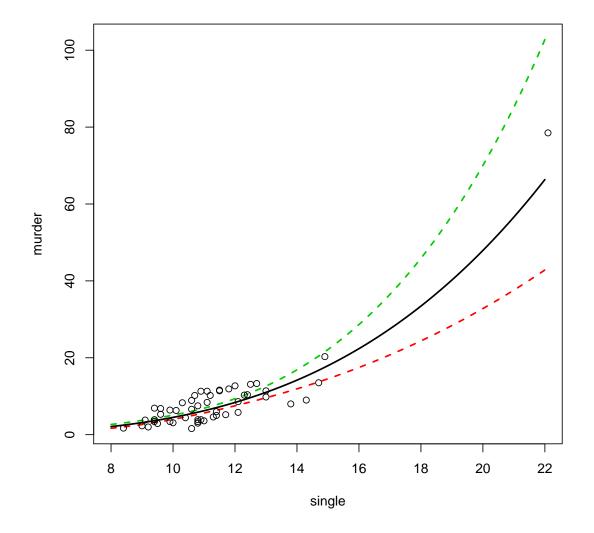
exp(predint)

## fit lwr upr
## 1 2.098482 0.9885837 4.454479</pre>
```

(g) The model equation is $\log(\text{murder}) = \alpha + \beta \log(\text{single})$, which can be expressed as:

$$murder = A \times single^{\beta}$$

The 'slope' parameter β has become an exponent. The model defines a power-law between the two variables, with β measuring the strength of this relationship.



```
3. x <- c(8, 8, 8, 11, 17, 17, 20, 20, 23, 26, 26, 26)

y <- c(14.8, 9.0, 11.0, 17.3, 20.8, 23.7, 24.4,

28.9, 27.8, 29.3, 35.0, 33.4, 37.8)

m1 <- lm(y ~ x)

summary(m1)
```

```
(a) coef(m1)

## (Intercept) x
## 1.438676 1.280423

sigma(m1)

## [1] 2.160964
```

```
(b) coef(summary(m1))

## Estimate Std. Error t value Pr(>|t|)

## (Intercept) 1.438676 1.69846317 0.8470459 4.150186e-01

## x 1.280423 0.08982449 14.2547169 1.946385e-08
```

```
(c) data2 <- data.frame(x = 18)
    predict(m1, data2, interval = "confidence")

## fit lwr upr
## 1 24.48628 23.16574 25.80683</pre>
```

```
(d) predict(m1, data2, interval = "prediction")
## fit lwr upr
## 1 24.48628 19.55012 29.42245
```