

H4

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Oehlert Problem 8.4.

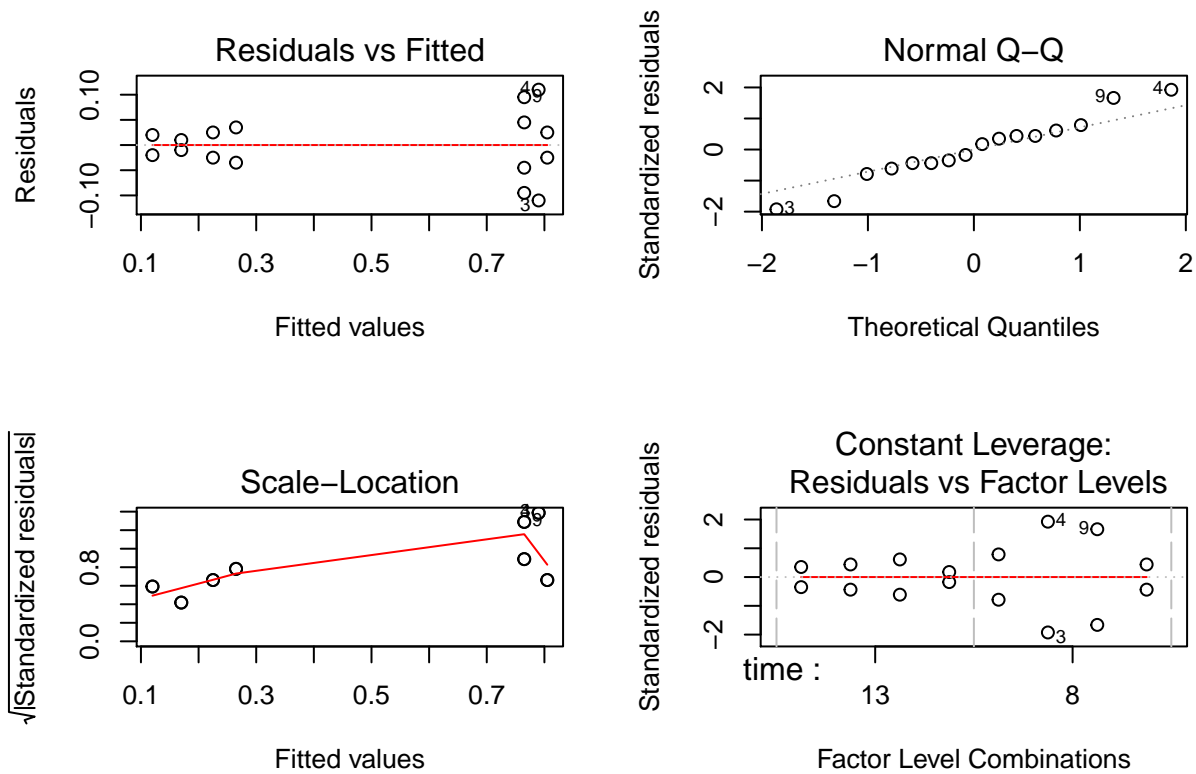
```
c1<-c(.83,.78)
c2<-c(.68,.90)
c3<-c(.18,.16)
c4<-c(.25,.20)
c5<-c(.86,.67)
c6<-c(.72,.81)
c7<-c(.30,.23)
c8<-c(.10,.14)

delamination<-data.frame(substrates=c(c1,c2,c3,c4,c5,c6,c7,c8),
  time=c(rep("8",4),rep("13",4),rep("8",4),rep("13",4)),
  airflow=c( rep(c(rep("low",2),rep("high",2)),2)),
  laser=c(rep("old",8),rep("new",8)) )

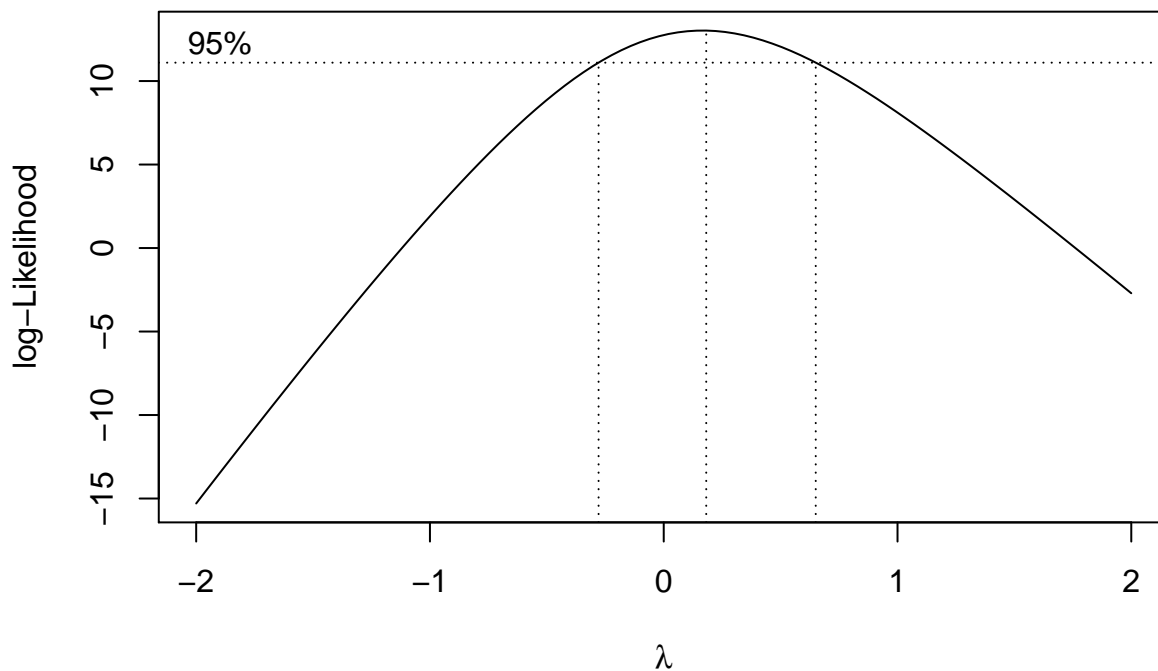
attach(delamination)
model_1=aov(substrates~time*airflow*laser,data = delamination)
summary(model_1)
```

	##	Df	Sum Sq	Mean Sq	F value	Pr(>F)
## time		1	1.3748	1.3748	210.489	4.99e-07 ***
## airflow		1	0.0028	0.0028	0.422	0.534
## laser		1	0.0014	0.0014	0.215	0.655
## time:airflow		1	0.0014	0.0014	0.215	0.655
## time:laser		1	0.0008	0.0008	0.116	0.742
## airflow:laser		1	0.0086	0.0086	1.310	0.285
## time:airflow:laser		1	0.0116	0.0116	1.769	0.220
## Residuals		8	0.0523	0.0065		
## ---						
## Signif. codes:		0 '***'	0.001 '**'	0.01 '*'	0.05 '.'	0.1 ' ' 1

```
##examine the statistical assumptions:
##these observations and groups are independent
par(mfrow=c(2,2))
plot(model_1)
```



```
par(mfrow=c(1,1))
##In the residuals vs fitted values plot,it is not a ideal pattern.So, the assumption of cantant varian
library(MASS)
par(mfrow=c(1,1))
b<-boxcox(model_1)
```

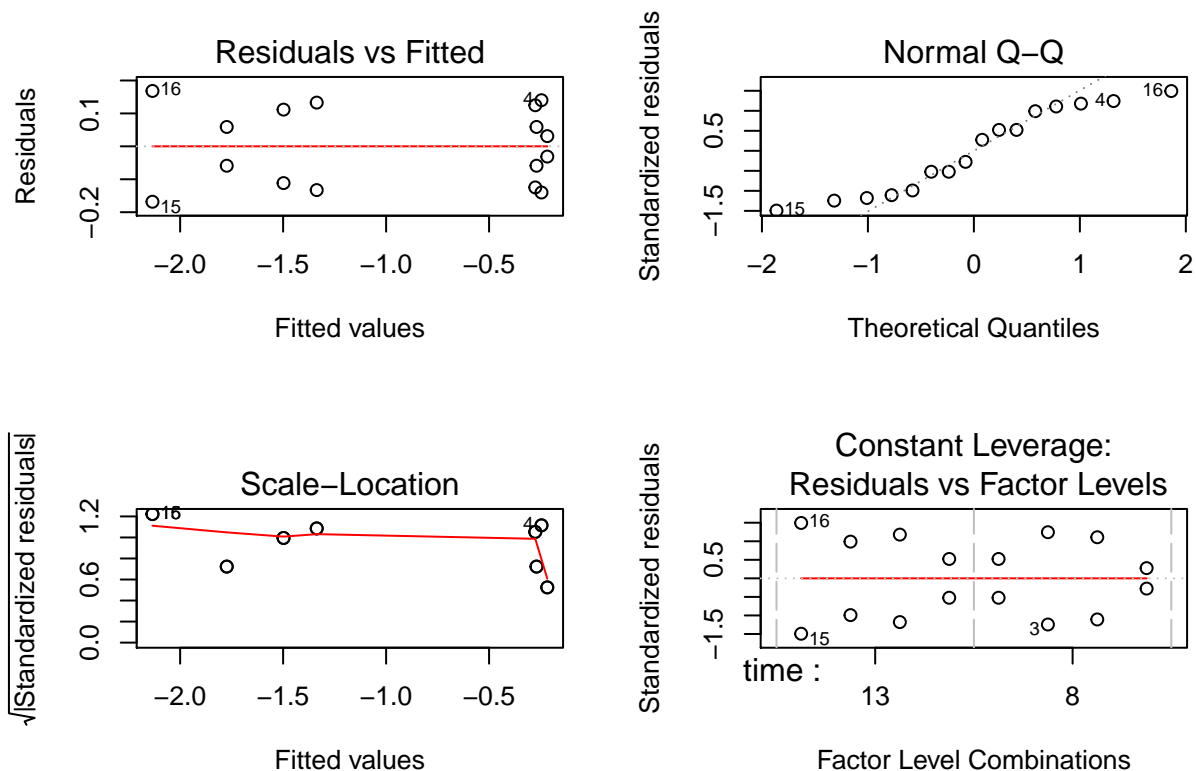


```
lamb<-b$x[which.max(b$y)]
lamb
```

```
## [1] 0.1818182
##Take lambda is 0
model_11=aov(log(substrates)~time*airflow*laser,data = delamination)
summary(model_11)
```

```
##              Df Sum Sq Mean Sq F value    Pr(>F)
## time          1  8.221    8.221 323.906 9.32e-08 ***
## airflow       1  0.074    0.074   2.912  0.12628
## laser        1  0.020    0.020   0.783  0.40200
## time:airflow  1  0.062    0.062   2.459  0.15551
## time:laser    1  0.003    0.003   0.135  0.72241
## airflow:laser  1  0.270    0.270  10.637  0.01150 *
## time:airflow:laser 1  0.307    0.307  12.081  0.00837 **
## Residuals     8  0.203    0.025
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##examine the statistical assumptions:
##these observations and groups are independent
par(mfrow=c(2,2))
plot(model_11)
```



Before boxcox transformation, based on the p-value, the term that has significant difference is time.

After boxcox transformation, based on the p-value, the terms that have significant difference are time, airflow:laser and time:airflow:laser. So, the treatments time, airflow and laser exist significant difference.

Hothorn and Everitt Ex. 5.2

```
library(HSAUR3)
```

```
## Loading required package: tools
```

```
data(schooldays,package="HSAUR3")
head(schooldays)
```

```
##           race gender school learner absent
## 1 aboriginal  male     F0    slow      2
## 2 aboriginal  male     F0    slow     11
## 3 aboriginal  male     F0    slow     14
## 4 aboriginal  male     F0 average    5
## 5 aboriginal  male     F0 average    5
## 6 aboriginal  male     F0 average    13
```

```
attach(schooldays)
```

a.

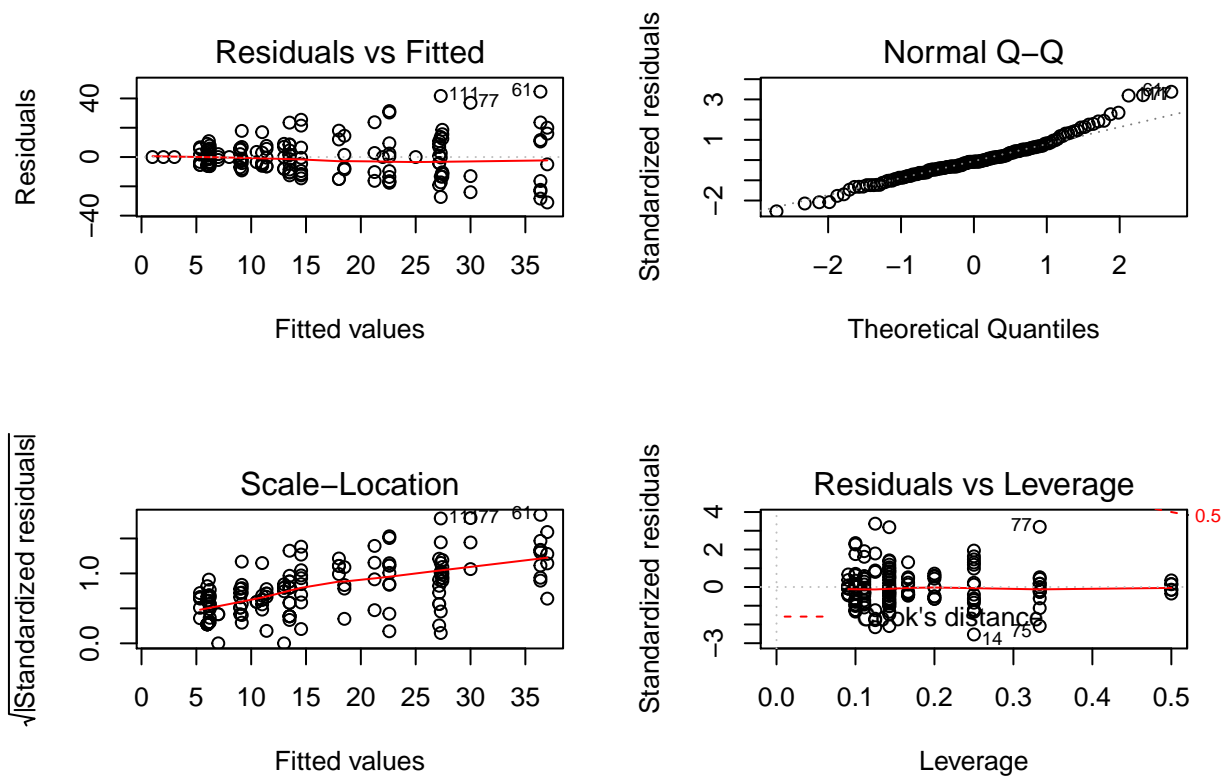
```
model2_1=aov(absent~race*gender*school*learner,data = schooldays)
summary(model2_1)
```

```
##               Df Sum Sq Mean Sq F value    Pr(>F)
## race           1   2646   2645.7   13.258 0.000400 ***
## gender         1    339    338.9    1.698 0.194985
## school         3   1222    407.3    2.041 0.111672
## learner        1     17     17.3    0.087 0.769087
## race:gender     1    173    173.3    0.868 0.353295
## race:school     3   3628   1209.2    6.059 0.000702 ***
## gender:school   3   1504    501.3    2.512 0.061747 .
## race:learner    1     67     67.4    0.338 0.562291
## gender:learner  1      9      9.4    0.047 0.828522
## school:learner  3   1931    643.5    3.225 0.025011 *
## race:gender:school 3    206     68.7    0.344 0.793404
## race:gender:learner 1    388    387.6    1.942 0.165957
## race:school:learner 3   1419    473.0    2.370 0.073836 .
## gender:school:learner 3    609    203.1    1.018 0.387326
## race:gender:school:learner 3    202     67.4    0.338 0.798014
## Residuals      122  24346    199.6
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

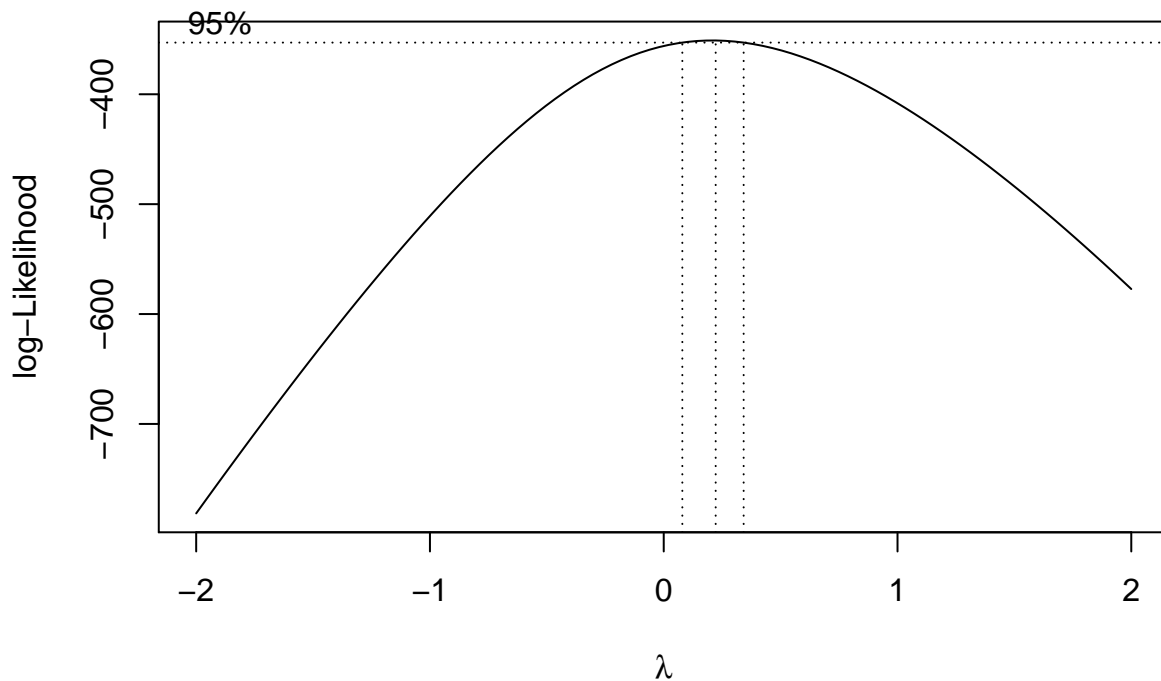
```
##examine the statistical assumptions:
##these observations and groups are independent
par(mfrow=c(2,2))
plot(model2_1)
```

```
## Warning: not plotting observations with leverage one:
## 34, 62, 93, 112, 143, 144
```

```
## Warning: not plotting observations with leverage one:
## 34, 62, 93, 112, 143, 144
```



```
par(mfrow=c(1,1))
##In the residuals vs fitted values plot,these points have a tendency to gradually spread along the x-a
schooldays$absent=schooldays$absent+1
library(MASS)
par(mfrow=c(1,1))
b<-boxcox(model12_1)
```



```
lam<-b$x[which.max(b$y)]
lam
```

```
## [1] 0.2222222
```

```
##Take lambda=0
```

```
model2_11=aov(log(absent)~race*gender*school*learner,data = schooldays)
summary(model2_11)
```

```
##              Df Sum Sq Mean Sq F value    Pr(>F)
## race              1  14.89   14.889    17.820 4.7e-05 ***
## gender            1   0.90    0.902     1.080 0.30072
## school            3   2.27    0.758     0.907 0.44003
## learner           1   0.00    0.004     0.005 0.94349
## race:gender        1   0.21    0.205     0.245 0.62119
## race:school        3  11.07    3.691     4.417 0.00551 **
## gender:school      3   8.33    2.778     3.325 0.02203 *
## race:learner       1   0.07    0.073     0.087 0.76813
## gender:learner     1   0.01    0.012     0.014 0.90634
## school:learner     3   5.70    1.900     2.274 0.08342 .
## race:gender:school  3   2.36    0.788     0.943 0.42193
## race:gender:learner 1   2.62    2.617     3.132 0.07928 .
## race:school:learner 3   6.14    2.048     2.451 0.06670 .
## gender:school:learner 3   3.21    1.070     1.281 0.28406
## race:gender:school:learner 3   0.97    0.324     0.387 0.76224
## Residuals        122 101.93    0.836
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##examine the statistical assumptions:
```

```
##these observations and groups are independent
```

```
par(mfrow=c(2,2))
```

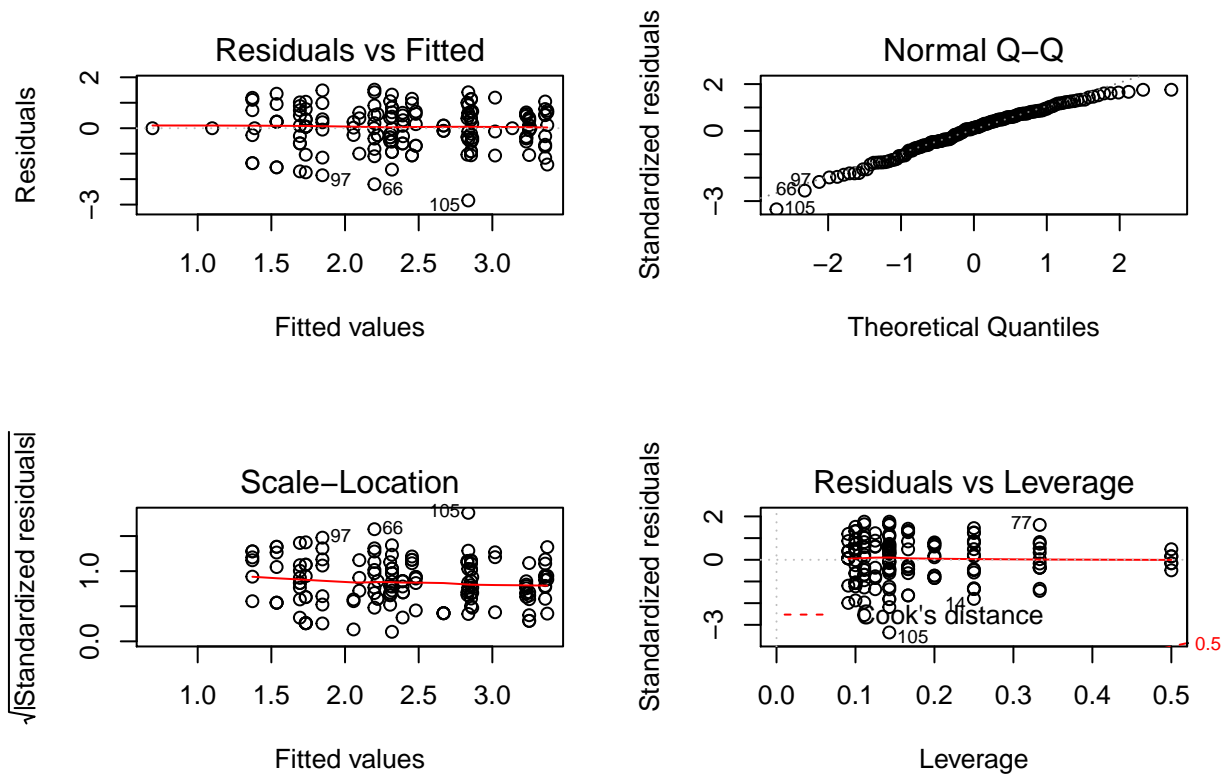
```
plot(model2_11)
```

```
## Warning: not plotting observations with leverage one:
```

```
## 34, 62, 93, 112, 143, 144
```

```
## Warning: not plotting observations with leverage one:
```

```
## 34, 62, 93, 112, 143, 144
```



```
par(mfrow=c(1,1))
```

```
##In the residuals vs fitted values plot, the most of the dots are scattered around the red line which
##In the Q-Q plot, the most of points are on the line. So, the assumption of normality distributed has
```

Based on the p-value, the items which have significant difference are race, race:school, school:learner

So, the factors that affect absenteeism are race, school and learner.

b.

```
model2_2=aov(log(absent)~gender*school*race*learner,data = schooldays)
summary(model2_2)
```

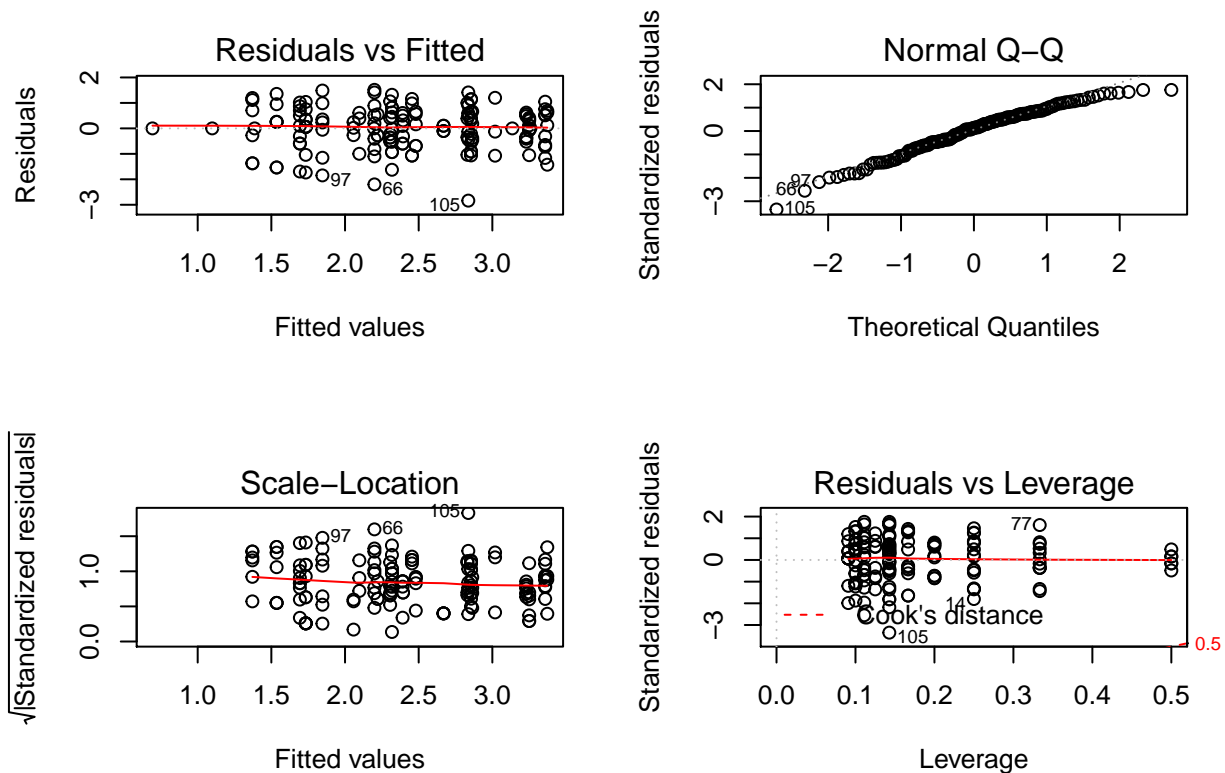
	Df	Sum Sq	Mean Sq	F value	Pr(>F)
gender	1	0.78	0.785	0.939	0.33444
school	3	3.38	1.127	1.349	0.26187
race	1	13.90	13.899	16.635	8.11e-05 ***
learner	1	0.00	0.004	0.005	0.94349
gender:school	3	8.98	2.993	3.582	0.01589 *
gender:race	1	0.06	0.061	0.073	0.78744
school:race	3	10.57	3.524	4.218	0.00709 **
gender:learner	1	0.01	0.014	0.016	0.89896
school:learner	3	5.77	1.923	2.301	0.08056 .
race:learner	1	0.00	0.002	0.002	0.96561
gender:school:race	3	2.36	0.788	0.943	0.42193
gender:school:learner	3	3.60	1.201	1.437	0.23532
gender:race:learner	1	2.60	2.596	3.107	0.08045 .
school:race:learner	3	5.77	1.924	2.303	0.08039 .
gender:school:race:learner	3	0.97	0.324	0.387	0.76224
Residuals	122	101.93	0.836		

```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
par(mfrow=c(2,2))
plot(model2_2)
```

```
## Warning: not plotting observations with leverage one:
## 34, 62, 93, 112, 143, 144
```

```
## Warning: not plotting observations with leverage one:
## 34, 62, 93, 112, 143, 144
```



```
par(mfrow=c(1,1))## the assumptions have been met
```

Each term's p-value has changed but gender:school:race:learner p-value is same (a).

c.

```
library(car)
```

```
## Loading required package: carData
```

```
Anova(lm(log(absent)~race*gender*school*learner,data = schooldays),type=2)
```

```
## Anova Table (Type II tests)
```

```
##
```

```
## Response: log(absent)
```

```
##
```

	Sum Sq	Df	F value	Pr(>F)	
## race	12.070	1	14.4461	0.0002265	***
## gender	0.977	1	1.1693	0.2816804	
## school	2.007	3	0.8007	0.4958067	
## learner	0.062	1	0.0747	0.7851147	
## race:gender	0.296	1	0.3542	0.5528227	
## race:school	4.840	3	1.9311	0.1280931	


```
## gender:school          11.000    3  4.3887 0.0057087 **
## race:learner           0.086    1  0.1034 0.7483367
## gender:learner         0.357    1  0.4267 0.5148267
## school:learner        7.359    3  2.9358 0.0361005 *
## race:gender:school     1.457    3  0.5814 0.6283416
## race:gender:learner    0.676    1  0.8086 0.3703125
## race:school:learner    5.773    3  2.3030 0.0803907 .
## gender:school:learner  3.210    3  1.2808 0.2840598
## race:gender:school:learner 0.971    3  0.3874 0.7622400
## Residuals             101.933 122
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
summary(aov(log(absent)~race*gender*school*learner,data = schooldays))
```

```
##              Df Sum Sq Mean Sq F value    Pr(>F)
## race              1  14.89   14.889   17.820 4.7e-05 ***
## gender            1   0.90    0.902    1.080 0.30072
## school            3   2.27    0.758    0.907 0.44003
## learner           1   0.00    0.004    0.005 0.94349
## race:gender        1   0.21    0.205    0.245 0.62119
## race:school         3  11.07    3.691    4.417 0.00551 **
## gender:school       3   8.33    2.778    3.325 0.02203 *
## race:learner        1   0.07    0.073    0.087 0.76813
## gender:learner      1   0.01    0.012    0.014 0.90634
## school:learner      3   5.70    1.900    2.274 0.08342 .
## race:gender:school   3   2.36    0.788    0.943 0.42193
## race:gender:learner  1   2.62    2.617    3.132 0.07928 .
## race:school:learner  3   6.14    2.048    2.451 0.06670 .
## gender:school:learner 3   3.21    1.070    1.281 0.28406
## race:gender:school:learner 3   0.97    0.324    0.387 0.76224
## Residuals         122 101.93    0.836
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

In these two models, the Residuals' Sum Sq and df keep the same but other terms' Sum Sq changed.

Beside, in the model using Type I sum of squares, the terms that have significant difference are race, race:school,gender:school. In the model using Type II sum of squares, the terms that have significant difference are race,gender:school and school:learner.

d.

For factor race, if I put race into model first

Type I sum of squares: SS(race)

Type II sum of squares: SS(race|1,gender,school,learner,gender:school,gender:learner,school:learner,gender:school:learner)

e.

For race:gender:school:learner

Type I sum of squares: SS(race:gender:school:learner|race,gender,school,learnerr,race:gender,race:school,gender:school,race:learnerr)

Type II sum of squares:SS(race:gender:school:learner|1,race,gender,school,learnerr,race:gender,race:school,gender:school,race:learnerr)

When we calculate Type I sum of squares of race:gender:school:learner, it is actually an incremental improvement given other items added to the model. So, they are same.

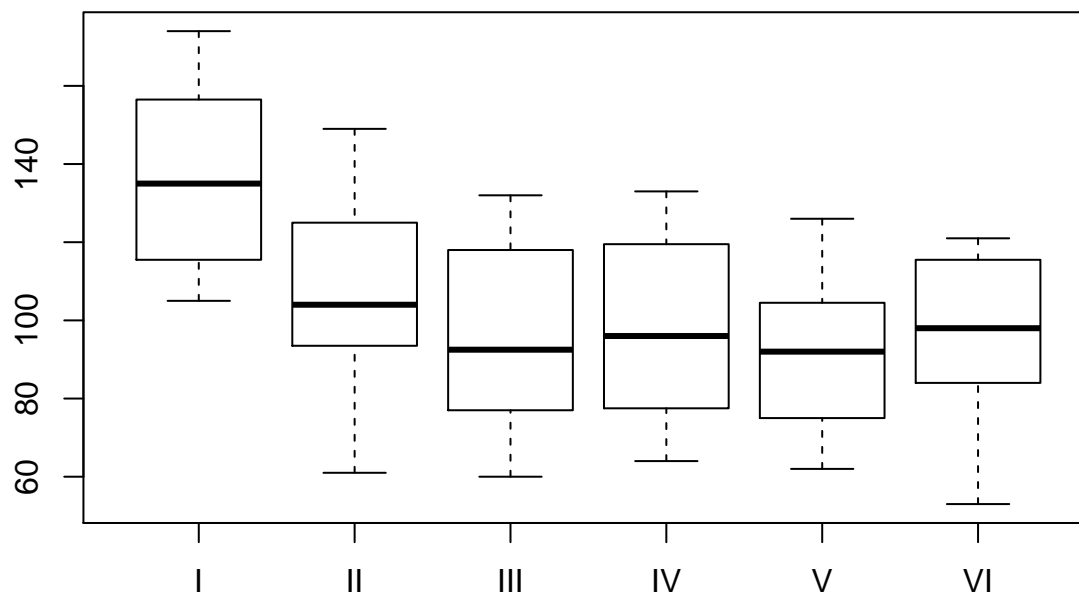
Oat Problem

```
library(MASS)
data(oats, package="MASS")
attach(oats)
head(oats)
```

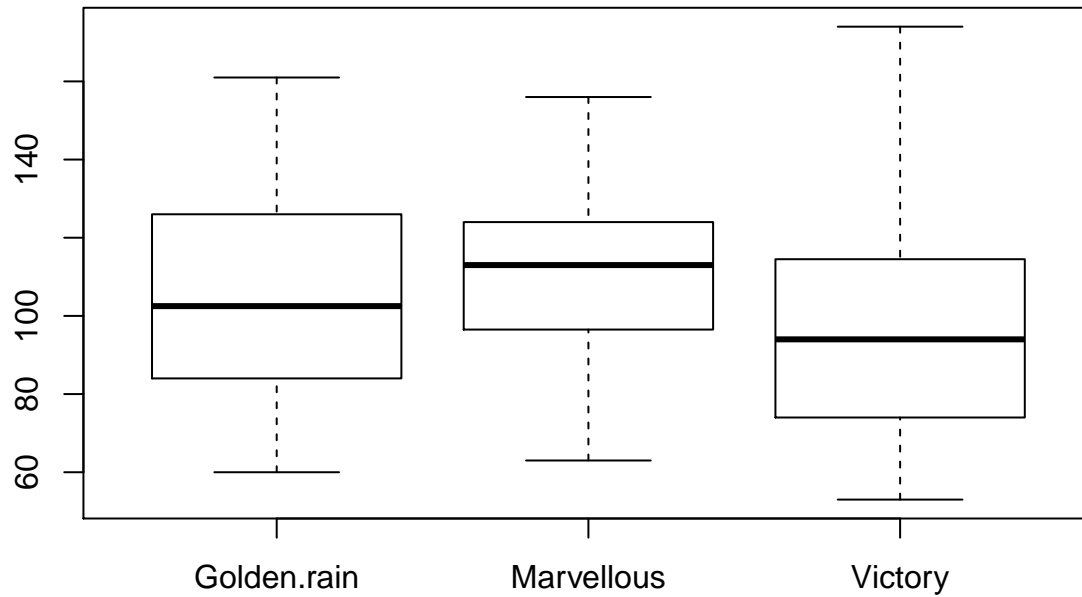
```
##   B      V      N   Y
## 1 I  Victory 0.0cwt 111
## 2 I  Victory 0.2cwt 130
## 3 I  Victory 0.4cwt 157
## 4 I  Victory 0.6cwt 174
## 5 I Golden.rain 0.0cwt 117
## 6 I Golden.rain 0.2cwt 114
```

a.

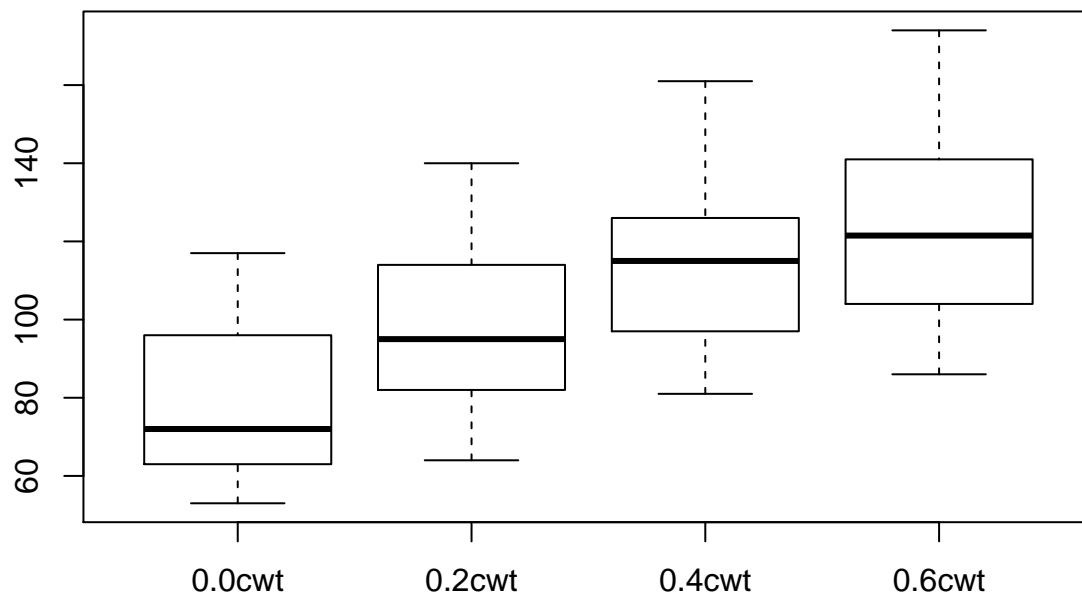
```
boxplot(Y~B)
```



```
boxplot(Y~V)
```



```
boxplot(Y~N)
```



From the plot, the treatment Nitrogen concentration appears different. And, the block appears different.

b.

Hypothesis test:

Nitrogen concentration:

H_0 : All nitrogen concentrations are the same H_a : At least one nitrogen concentration is different

Variety of oats:

H_0 : All variety of oats are the same H_a : At least one kind of oats is different

Nitrogen concentration by Variety of oats interaction:

H_0 : All interactions are the same H_a : All interactions are different

```
model3_1=aov(Y~V*N,data=oats)
summary(model3_1)
```

```
##              Df Sum Sq Mean Sq F value    Pr(>F)
## V              2   1786      893   1.795    0.175
## N              3  20020     6673  13.411 8.37e-07 ***
## V:N            6    322       54   0.108    0.995
## Residuals     60  29857      498
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Because the p-value of term N is less than 0.05, we reject the null hypothesis. So, there is significant difference between nitrogen concentrations.

Because the p-value of term V and term V:N is greater than 0.05 or 0.1, we fail to reject the null hypothesis. So, there is not significant difference between variety of oats, and also nitrogen concentration by variety of oats interaction.

c.

```
model3_2=aov(Y~V+N+V:N+B,data=oats)
summary(model3_2)
```

```
##              Df Sum Sq Mean Sq F value    Pr(>F)
## V              2   1786      893   3.513   0.0366 *
## N              3  20020     6673  26.251 1.13e-10 ***
## B              5  15875     3175  12.489 4.09e-08 ***
## V:N            6    322       54   0.211   0.9719
## Residuals     55  13982      254
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Because the p-value of term V and term N is less than 0.05, we reject the null hypothesis. So, there is significant difference between variety of oats, and also nitrogen concentration. Also, the p-value of block is less than 0.05, the block has significant difference.

Because the p-value of term V:N is greater than 0.05 or 0.1, we fail to reject the null hypothesis. So, there is not significant difference between nitrogen concentration by variety of oats interaction.

By including block in my model, it causes the reduction of sum squares of residuals, loss of df for residuals and the reduction of MSE. Because of the reduction of MSE, F value becomes larger and p-value becomes smaller of the terms N and V. Hence, we get that variety of oats have significant difference which is opposite to part (b).

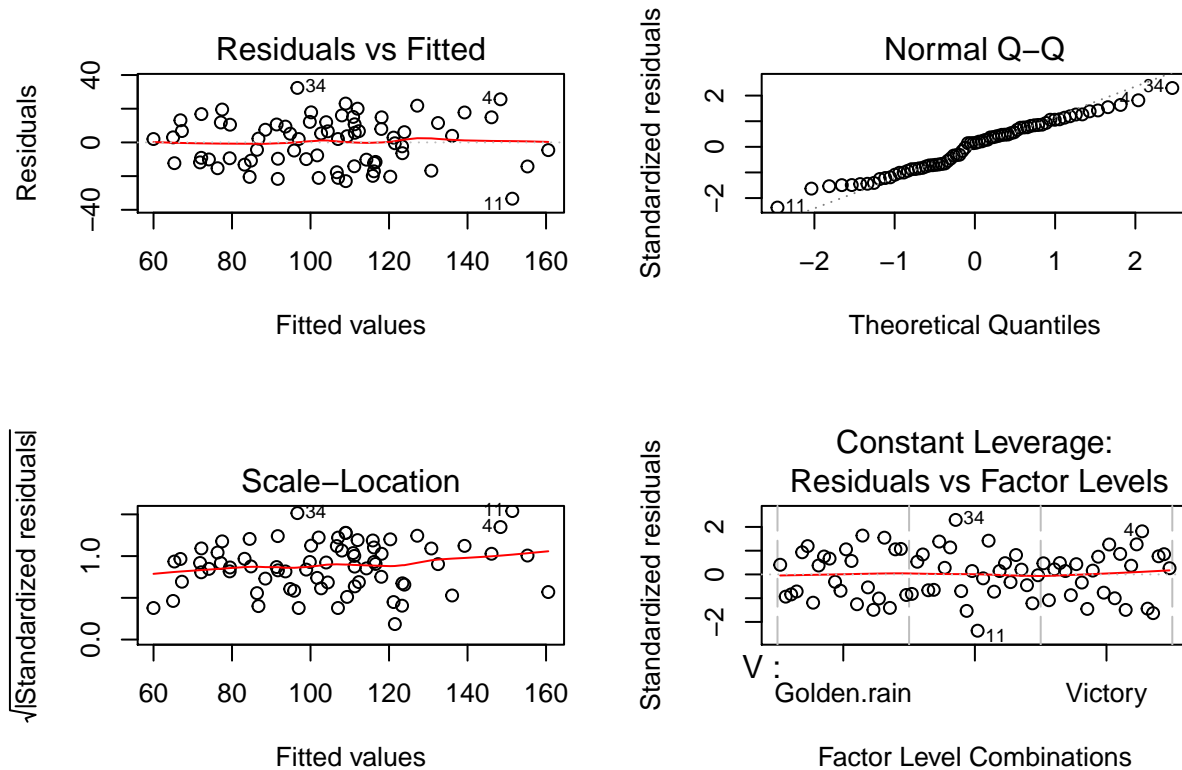
d.

```
##my final model: from part(c)
model3_3=aov(Y~V+N+B,data=oats)
summary(model3_3)
```

```
##              Df Sum Sq Mean Sq F value    Pr(>F)
## V              2   1786      893   3.809   0.0276 *
## N              3  20020     6673  28.460 1.24e-11 ***
## B              5  15875     3175  13.540 6.91e-09 ***
## Residuals     61  14304      234
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##examine the statistical assumptions:
##The observations and groups are independent
```

```
par(mfrow=c(2,2))
plot(model3_3)
```



##In the residuals vs fitted values plot, the most of the dots are scattered around the red line which

##In the Q-Q plot, the most of points are on the line. So, the assumption of normality distributed has

e.

```
TukeyHSD(aov(Y~V+N,data=oats))
```

```
## Tukey multiple comparisons of means
## 95% family-wise confidence level
##
## Fit: aov(formula = Y ~ V + N, data = oats)
##
## $V
##          diff          lwr          upr      p adj
## Marvellous-Golden.rain  5.291667 -9.509153 20.092486 0.6689436
## Victory-Golden.rain    -6.875000 -21.675819  7.925819 0.5090858
## Victory-Marvellous    -12.166667 -26.967486  2.634153 0.1274623
##
## $N
##          diff          lwr          upr      p adj
## 0.2cwt-0.0cwt 19.500000  0.712970 38.28703 0.0389760
## 0.4cwt-0.0cwt 34.833333 16.046303 53.62036 0.0000400
## 0.6cwt-0.0cwt 44.000000 25.212970 62.78703 0.0000003
## 0.4cwt-0.2cwt 15.333333 -3.453697 34.12036 0.1479244
## 0.6cwt-0.2cwt 24.500000  5.712970 43.28703 0.0055113
## 0.6cwt-0.4cwt  9.166667 -9.620363 27.95370 0.5749975
```

By Tukey's test without block, we could see that the treatment nitrogen concentrations have significant difference from the p-value.