

STAT5303HW2

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E4.4

This experiment has six treatments. There are 3 treatments about low moisture: particle board at low moisture, plywood at low moisture and solid fir at low moisture. There are 2 treatments about medium moisture: particle board at medium moisture and plywood at medium moisture. And one particle board at high moisture.

Because we are only interested in comparing low moisture to medium moisture, I will set up a linear contrast low moisture vs medium moisture with coefficients of these six treatments $(1/3, 1/3, 1/3, -1/2, -1/2, 0)$ with respect to particle board at low moisture, plywood at low moisture, solid fir at low moisture, particle board at medium moisture, plywood at medium moisture and particle board at high moisture.

P4.2

```
library(cfcdae)
```

```
## Warning: package 'cfcdae' was built under R version 4.0.2
```

```
## Registered S3 method overwritten by 'DoE.base':
```

```
##   method          from
```

```
##   factorize.factor conf.design
```

```
data("FruitFlyLifespan")
```

```
attach(FruitFlyLifespan)
```

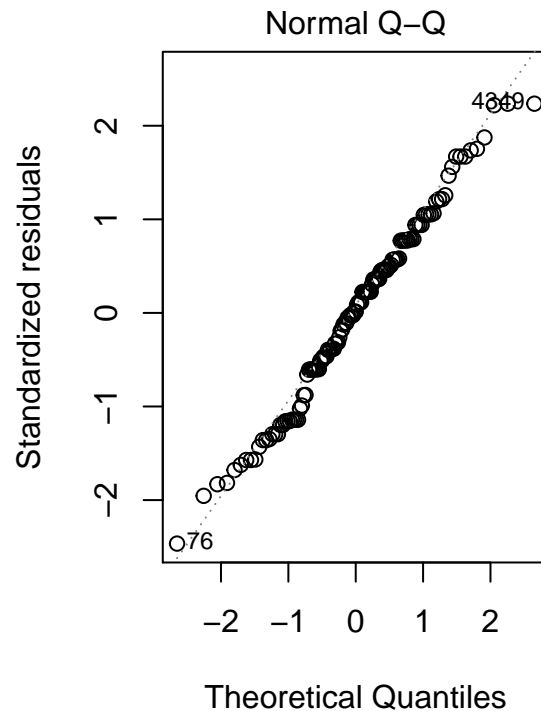
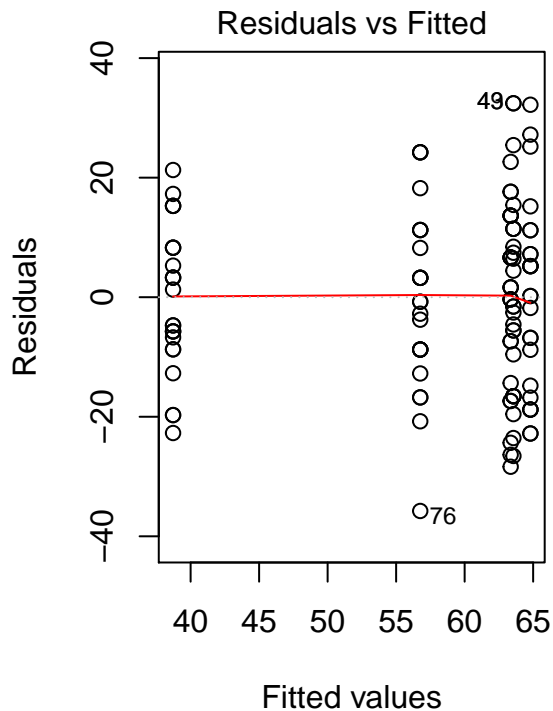
```
FruitFlyLifespan$companions<-as.factor(FruitFlyLifespan$companions)
```

```
fly.separate<-lm(longevity~companions,data = FruitFlyLifespan)
```

```
##We check the model assumptions.
```

```
par(mfrow = c(1, 2))
```

```
plot(fly.separate, which = c(1, 2)) ##assumptions are satisfied
```



```
summary(fly.separate);model.effects(fly.separate,"companions")
```

```
##
## Call:
## lm(formula = longevity ~ companions, data = FruitFlyLifespan)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -35.76  -8.76   0.20  11.20  32.44
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    57.440     1.325  43.368 < 2e-16 ***
## companions1     6.120     2.649   2.310  0.02258 *
## companions2     7.360     2.649   2.778  0.00634 **
## companions3    -0.680     2.649  -0.257  0.79785
## companions4   -18.720     2.649  -7.067 1.13e-10 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 14.81 on 120 degrees of freedom
## Multiple R-squared:  0.3121, Adjusted R-squared:  0.2892
## F-statistic: 13.61 on 4 and 120 DF, p-value: 3.516e-09
##
## 1pregnant  1virgin 8pregnant  8virgin      none
##      6.12     7.36    -0.68   -18.72     5.92
```

```
anova(fly.separate)##Since the p-value is so small, there is significant difference between these treatments
```

```
## Analysis of Variance Table
##
## Response: longevity
```

```
##           Df Sum Sq Mean Sq F value    Pr(>F)
## companions    4  11939  2984.82  13.612 3.516e-09 ***
## Residuals  120   26314   219.28
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##contrast 1:
## I want to test if the males were supplied with receptive virgin female fruit flies per day have a sh
linear.contrast(fly.separate,companions,c(0,0.5,0,0.5,-1))

##   estimates      se  t-value    p-value lower-ci upper-ci
## 1      -11.6 3.627225 -3.198037 0.001769626 -18.78165 -4.418347
##Since p-value is 0.001769626, it shows a statistically sinificant difference so I can reject the null
##contrast 2:
## I want to test if the males were supplied with 8 receptive virgin female fruit flies per day have a s
linear.contrast(fly.separate,companions,c(0,1,0,-1,0))

##   estimates      se  t-value    p-value lower-ci upper-ci
## 1       26.08 4.188358 6.226783 7.268917e-09 17.78734 34.37266
##Since p-value is 7.268917e-09 which is so small, it shows a statistically sinificant difference between
##contrast 3:
## I want to test if the males were supplied with receptive virgin female fruit flies per day have a sh
linear.contrast(fly.separate,companions,c(0.5,-0.5,0.5,-0.5,0))

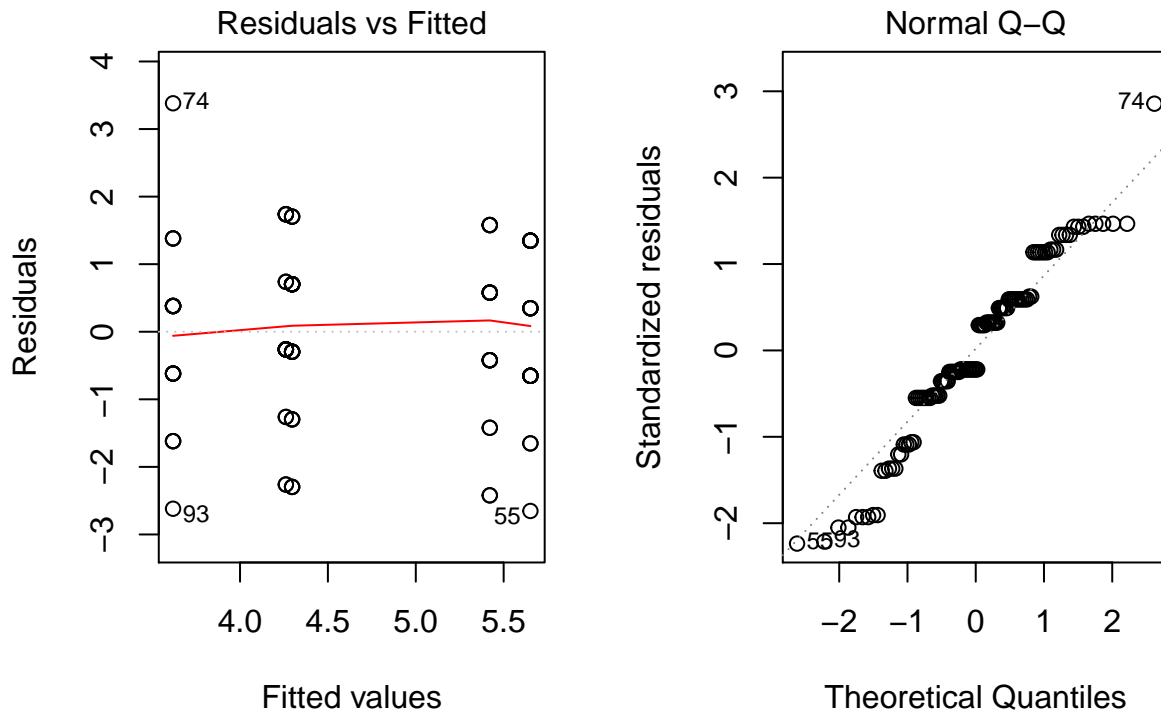
##   estimates      se  t-value    p-value lower-ci upper-ci
## 1         8.4 2.961617 2.836289 0.005358879 2.536205 14.26379
##Since p-value is 0.005358879,it cannt explian there is a statistically sinificant difference between
```

P4.3

```
data("PoliticalAds")
attach(PoliticalAds)
condition<-as.factor(condition)
levels(condition)

## [1] "Neg ed."      "Pos ed."      "Trans & neg ed." "Trans & pos ed."
## [5] "Trans only"

pads<-lm(response~condition)
##We check the model assumptions.
par(mfrow = c(1, 2))
plot(pads, which = c(1, 2)) ##assumptions are satisfied
```



```
anova(pads);model.effects(pads,"condition")
```

```
## Analysis of Variance Table
##
## Response: response
##           Df Sum Sq Mean Sq F value    Pr(>F)
## condition   4  63.524   15.881   10.796 2.123e-07 ***
## Residuals 108 158.866    1.471
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##           Neg ed.          Pos ed. Trans & neg ed. Trans & pos ed.      Trans only
##           0.7711646        -1.0308404        1.0022859        -0.3535917        -0.3890184
```

From the ANOVA table, the p-value 2.123e-07 is so small so it shows that there is a statistically significant difference between these 5 treatments. So, it implies that the media coverage affects perception.

```
## I will set up a linear contrast to compare Trans vs no Trans
linear.contrast(pads,condition,c(-0.5,-0.5,1/3,1/3,1/3))
```

```
##   estimates      se t-value    p-value  lower-ci upper-ci
## 1 0.2163965 0.2390239 0.905334 0.3673037 -0.2573904 0.6901833
```

There is not enough evidence that shows the actual transcript changes the response to media reporting.

```
## I will set up a linear contrast to compare pos.ad vs neg.ad
linear.contrast(pads,condition,c(-0.5,0.5,-0.5,0.5,0))
```

```
##   estimates      se t-value    p-value  lower-ci upper-ci
## 1 -1.578941 0.2578289 -6.123989 1.505443e-08 -2.090003 -1.06788
```

Because the p-value 1.505443e-08 is so small so it has enough evidence to show that there is a statistically significant difference between positive reporting and negative reporting in improving perception.

P4.6

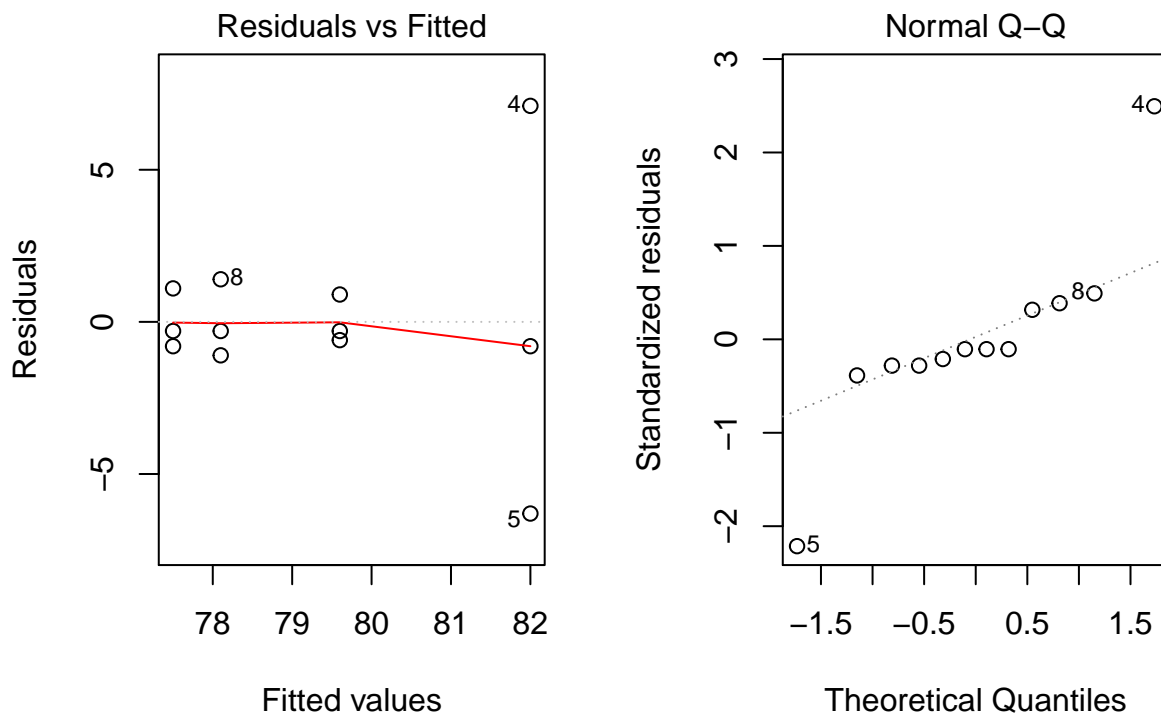
```
2*0+-1*1+-1*-1
```

```
## [1] 0
```

The result is 0 so these two contrast are orthogonol. These two contrasts are statistically independent. In this experiment,there are 3 group.Then by Pythagorean theorem, these two contrasts' sum squares partition the sum of squares for treatments with 2 df.So the sum of squares for the contrast with coefficients (0, 1, -1) is $100-80=20$.

E5.5

```
library(cfcdae)
data("OrangePulpSilage")
attach(OrangePulpSilage)
treatment<-as.factor(treatment)
orange.seperate<-lm(moisture~treatment)
##We check the model assumptions.
par(mfrow = c(1, 2))
plot(orange.seperate, which = c(1, 2)) ##assumptions are satisfied
```



```
levels(treatment)
```

```
## [1] "BeetPulp" "Control" "FormicAcid" "NaCl"
```

```
compare.to.control(orange.seperate,treatment,control=2,conf=0.95)
```

```
##           difference      lower      upper
## BeetPulp - Control      0.6 -7.582884  8.782884
## NaCl - Control          2.1 -6.082884 10.282884
## FormicAcid - Control    4.5 -3.682884 12.682884
```

P5.1

```
library(cfcdae)
data("FruitFlyLifespan")
```

```
attach(FruitFlyLifespan)
```

```
## The following objects are masked from FruitFlyLifespan (pos = 5):
```

```
##
```

```
##      companions, longevity
```

```
FruitFlyLifespan$companions<-as.factor(FruitFlyLifespan$companions)
```

```
fly.seperate<-lm(longevity~companions,data = FruitFlyLifespan)
```

```
fly.tukey<-pairwise(fly.seperate,companions)
```

```
fly.tukey
```

```
##
```

```
## Pairwise comparisons ( hsd ) of companions
```

	estimate	signif	diff	lower	upper
## 1pregnant - 1virgin	-1.24	11.60047	-12.840468	10.360468	
## 1pregnant - 8pregnant	6.80	11.60047	-4.800468	18.400468	
## * 1pregnant - 8virgin	24.84	11.60047	13.239532	36.440468	
## 1pregnant - none	0.20	11.60047	-11.400468	11.800468	
## 1virgin - 8pregnant	8.04	11.60047	-3.560468	19.640468	
## * 1virgin - 8virgin	26.08	11.60047	14.479532	37.680468	
## 1virgin - none	1.44	11.60047	-10.160468	13.040468	
## * 8pregnant - 8virgin	18.04	11.60047	6.439532	29.640468	
## 8pregnant - none	-6.60	11.60047	-18.200468	5.000468	
## * 8virgin - none	-24.64	11.60047	-36.240468	-13.039532	

```
sidelines(fly.tukey)
```

```
##
```

```
## 8virgin -18.72
```

```
## 8pregnant -0.68 |
```

```
## none 5.92 |
```

```
## 1pregnant 6.12 |
```

```
## 1virgin 7.36 |
```

So, the results shows that there is no significant difference between the treatments 8pregnant,1pregnant,1virgin and none. However,it also shows that 8virgin has significant difference with other treatment groups.

P5.2

```
library(cfcdae)
```

```
data("Autoclaving")
```

```
attach(Autoclaving)
```

```
## The following object is masked _by_ .GlobalEnv:
```

```
##
```

```
##      treatment
```

```
## The following object is masked from OrangePulpSilage:
```

```
##
```

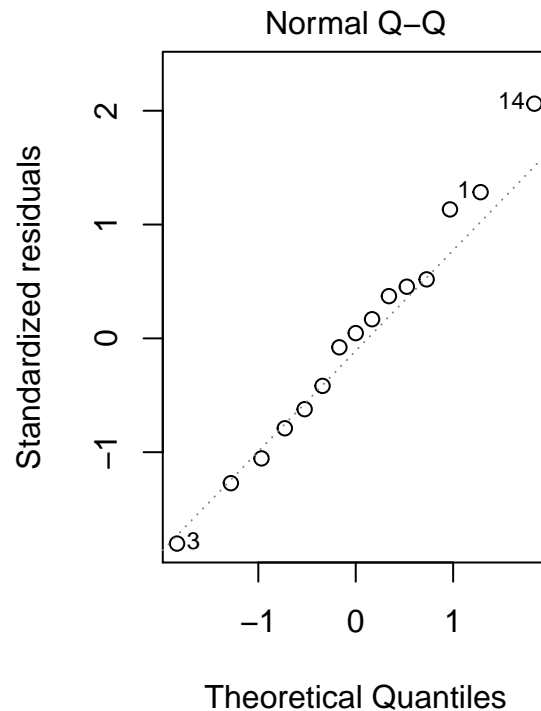
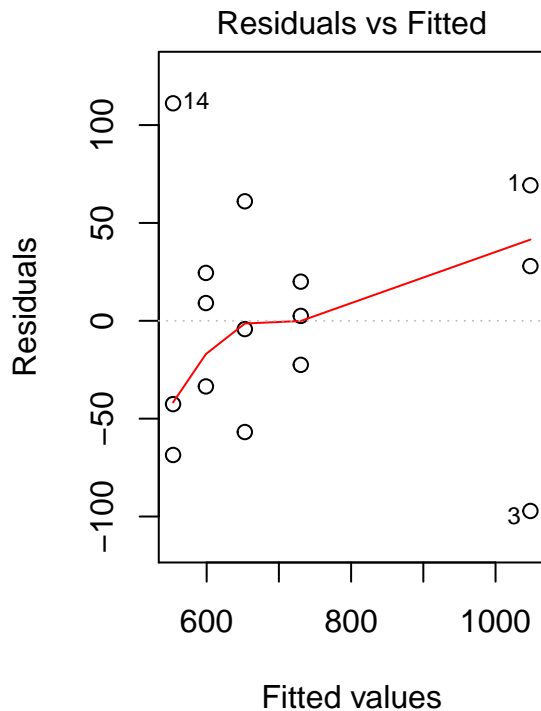
```
##      treatment
```

```
auto.seperate<-lm(modulus~treatment,data = Autoclaving)
```

```
##We check the model assumptions.
```

```
par(mfrow = c(1, 2))
```

```
plot(auto.seperate, which = c(1, 2)) ##assumptions are satisfied
```



```
anova(auto.seperate);model.effects(auto.seperate,treatment)
```

```
## Analysis of Variance Table
##
## Response: modulus
##           Df Sum Sq Mean Sq F value    Pr(>F)
## treatment  4 464143  116036    26.63 2.594e-05 ***
## Residuals 10  43574     4357
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

##      121_10      121_20      135_10      135_20      None
##      13.54000  -63.92667 -117.82667 -163.29333  331.50667
```

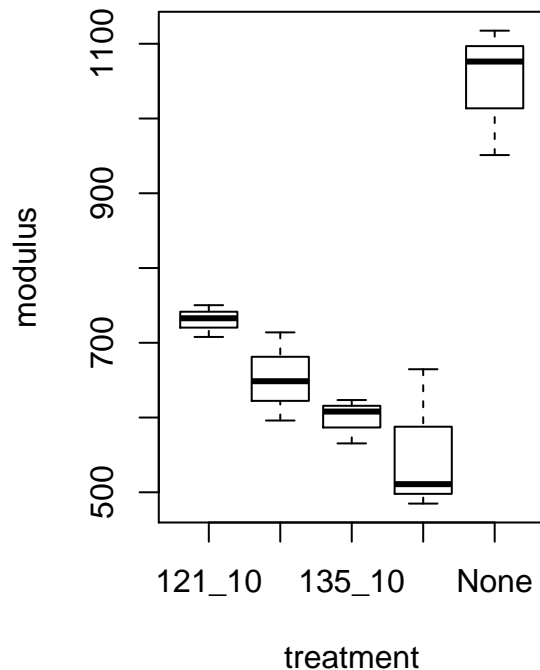
##By the anova table, the small p-value explains that there is significant difference between these treatments

```
boxplot(modulus~treatment,data = Autoclaving)
```

##I will choose the per comparison error rate.

```
pairwise(auto.seperate,treatment,confidence = 0.99,type="lsd")
```

```
##
## Pairwise comparisons ( lsd ) of treatment
##           estimate signif diff      lower      upper
##      121_10 - 121_20    77.46667    170.8151  -93.348421  248.2818
##      121_10 - 135_10   131.36667    170.8151  -39.448421  302.1818
## * 121_10 - 135_20   176.83333    170.8151    6.018245  347.6484
## * 121_10 - None   -317.96667    170.8151 -488.781755 -147.1516
##      121_20 - 135_10    53.90000    170.8151 -116.915088  224.7151
##      121_20 - 135_20    99.36667    170.8151  -71.448421  270.1818
## * 121_20 - None   -395.43333    170.8151 -566.248421 -224.6182
##      135_10 - 135_20    45.46667    170.8151 -125.348421  216.2818
## * 135_10 - None   -449.33333    170.8151 -620.148421 -278.5182
## * 135_20 - None   -494.80000    170.8151 -665.615088 -323.9849
```



P5.6

pros:

per comparison error rate: For these 10,000 locations, this error rate just does individual test for each null hypothesis ignoring the multiple testing. Each individual comparison is controlled at ϵ .

false discovery rate: Since the units 10,000 is a little large, it may cause that we might have many incorrect rejections. With this error rate, we allow more incorrect rejections of null hypothesis as the number of true rejections increases, but the ratio is limited. And the value of ϵ can be flexibly selected according to needs.

strong familywise error rate: It controls that probability of making any false rejections is at ϵ or under it, regardless of the number of correct rejections made.

cons:

per comparison error rate: It makes no correction for multiple comparisons, therefore there may be many Type I errors.

false discovery rate: It may increase more Type I errors.

strong familywise error rate: The value of ϵ is traditionally more fixed.

So in this experiment, at each location, we test the null hypothesis that the expected gamma ray emission rate is equal to the known background rate and clean up the locations that are significantly different. So, I would like to use per comparison error rate because it controls each individual test under the error rate.

P5.9

In this experiment, one student has a p-value less than 5% for telepathy, and one student has a p-value less than 5% for precognition. If we choose 0.05 as the significance level, in this case, p-value < 5%. Then, we conclude that the null hypothesis that any single person could exhibit any kind of ESP is false and should be rejected.