#### Sample solutions to exam 22-05-31

## Fråga 1

- a) We have p(2) = 1 0.1 0.3 0.2 = 0.4, so  $E[X] = 0.1 \cdot 1 + 0.4 \cdot 2 + 0.3 \cdot 3 + 0.2 \cdot 4 = 2.6$ ,  $E[X^2] = 0.1 \cdot 1^2 + 0.4 \cdot 2^2 + 0.3 \cdot 3^2 + 0.2 \cdot 4^2 = 7.6$ , so  $V[X] = E[X^2] E[X]^2 = 7.6 2.6^2 = 0.84$
- b)  $P(X \le 3 \mid X \ge 3) = \frac{P(X \le 3 \text{ and } X \ge 3)}{P(X \ge 3)} = \frac{P(X = 3)}{P(X \ge 3)} = \frac{0.3}{0.3 + 0.2} = 0.6$
- c) No, they are not independent: If they were, we would have  $P(X \le 3 \mid X \ge 3) = P(X \ge 3)$ , but  $P(X \le 3) = 0.1 + 0.4 + 0.3 = 0.8$ , and we calculated in b) that  $P(X \ge 3 \mid X \le 3) = 0.6$ .
- d)  $C(X,1/X) = E[X \cdot 1/X] E[X]E[1/X] = 1 E[X]E[1/X]$ , so we need to calculate  $E[1/X] = 0.1 \cdot \frac{1}{1} + 0.4 \cdot \frac{1}{2} + 0.3 \cdot \frac{1}{3} + 0.2 \cdot \frac{1}{4} = 0.45$ . This gives  $C(X,1/X) = 1 2.6 \cdot 0.45 = -0.17$

### Fråga 2

- a)  $F(x) = \int_{-\infty}^{x} f(y) dy = \int_{0}^{x} \frac{1}{4} y^{3} dy = \left[\frac{1}{16} y^{4}\right]_{0}^{x} = \frac{1}{16} x^{4} \text{ for } 0 \leqslant x \leqslant 2 \text{ (and } F(x) = 0 \text{ for } x < 0, F(x) = 1 \text{ for } x > 2).$   $E[X] = \int_{0}^{2} x \cdot \frac{1}{4} x^{3} dx = \int_{0}^{2} \frac{1}{4} x^{4} dx = \left[\frac{1}{20} x^{5}\right]_{0}^{2} = \frac{32}{20} \frac{0}{20} = 1.6$
- b) We use the inversion method, using the fact that if U is uniform from [0,1], then  $F_X^{-1}(U)$  has the same distribution as X. We have from part a) that  $F(x) = \frac{1}{16}x^4$ , seen as a function from [0,2] to [0,1]. To find the inverse, set y = F(x) and solve for x:

$$y = \frac{1}{16}x^4$$
, so  $x^4 = 16y$ , so  $x = 2y^{\frac{1}{4}}$ 

(As  $0 \leqslant y \leqslant 1$  and  $0 \leqslant x \leqslant 2$ ,  $x = 2y^{\frac{1}{4}}$  is the only solution.)

Applying  $F^{-1}(x) = 2x^{\frac{1}{4}}$  to the pseudorandom numbers from Re[0, 1] gives the following five simulated pseudorandom numbers.

$$t_1 = 2 \cdot 0.5755^{1/4} \approx 1.813, \quad t_2 = 2 \cdot 0.1438^{1/4} \approx 1.232, \quad t_3 = 2 \cdot 0.8224^{1/4} \approx 1.905$$

#### Fråga 3

Let  $X_1, X_2, \ldots, X_{400}$  be the amount of money (in SEK) that customers  $1, \ldots, 400$  spend at Anna's ice cream stand. By our assumptions these random variables are independent. Furthermore,

$$E[X_i] = 0.3 \cdot 10 + 0.5 \cdot 20 + 0.2 \cdot 30 = 19, \quad E[X_i^2] = 0.3 \cdot 10^2 + 0.5 \cdot 20^2 + 0.2 \cdot 30^2 = 410,$$
  
$$V[X_i] = E[X_i^2] - E[X_i]^2 = 410 - 19^2 = 49.$$

The total amount of money spent at the icecream shop is  $Y = X_1 + X_2 + \cdots + X_{400}$ , so we want to approximate  $P(Y \ge 7500)$ . Note that  $E[Y] = 400 \cdot E[X_1] = 400 \cdot 19 = 7600$  and, by independence,  $V[Y] = 400 \cdot V[X_1] = 19600$ . By the central limit theorem, we can approximate

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the distribution of Y with a normal distribution:  $Y \approx N(7600, 19600)$ . Therefore, letting  $Z \sim N(0, 1)$  be a standard normal random variable,

$$\begin{split} P(Y\geqslant 7500) &= P\left(\frac{Y-7600}{\sqrt{19600}}\geqslant \frac{7500-7600}{\sqrt{19600}}\right) \approx P\left(Z\geqslant -\frac{5}{7}\right) = 1 - P\left(Z\leqslant -\frac{5}{7}\right) \\ &= 1 - \Phi\left(-\frac{5}{7}\right) = 1 - \left(1 - \Phi\left(\frac{5}{7}\right)\right) = \Phi\left(\frac{5}{7}\right) \approx \Phi(0.71) \approx 0.7611, \end{split}$$

where we looked up the last value in the table.

### Fråga 4

- a) Let  $T_1, T_2$  be the corresponding estimators, then  $E[T_1] = 2E[X_1] E[X_2] = 2\mu \mu = \mu$ , so  $T_1$  is unbiased. Furthermore recall from the lectures that  $E[\bar{X}] = E[\bar{Y}] = \mu$ , and so  $E[T_2] = (E[\bar{X}] + E[\bar{Y}])/2 = (\mu + \mu)/2 = \mu$ , so  $T_2$  is unbiased.
- b) As all the random variables  $X_i$  and  $Y_j$  are independent, we can calculate the variances  $V[T_1] = 2^2V[X_1] + (-1)^2V[X_2] = 5$  and  $V[T_2] = \frac{1}{4}(V[\bar{X}] + V[\bar{Y}]) = \frac{1}{4}(\frac{1}{n} + \frac{100}{n}) = 25.25/n$  (recalling from the lectures that  $V[\bar{X}] = \sigma^2/n$  if  $V[X_i] = \sigma^2$ ). So  $T_1$  is more effective than  $T_2$  if 5 < 25.25/n, or equivalently n < 25.25/5 = 5.05, so this is the case for all integers  $n \le 5$  (and  $n \ge 2$ ).

# Fråga 5

a) Using the model with the estimated parameters, and letting  $x_0 = 55$ , the prediction for  $Y_0$  is

$$y_0 = \hat{m} + \hat{k}x_0 = 240 + 10.1 \cdot 55 = 795.5$$
 (milliseconds)

b) Since  $\hat{k} = 10.1 > 0$ , we are looking for a scatter plot where the points follow an 'upwards slope', which is either Figure 1 or Figure 3. As  $R^2 = 88.2\%$  is quite high, Figure 1 looks more suitable — the points are closer together and look more like a line, while they are quite spread out on Figure 3. So Figure 1 is the correct scatter plot.

(Remark: In Figure 3, we have  $R^2 = 43.1\%$ .)

c) In a linear regression model for the new server, the (new) random variables  $Y_i'$  corresponding to an input of length  $x_i'$  would follow  $Y_i' = m_{\text{new}} + k_{\text{new}}x_i' + \epsilon_i'$ , where  $\epsilon_i' \sim N(0, \sigma_{\text{new}}^2)$ . We can estimate the parameters  $m_{\text{new}}$  and  $k_{\text{new}}$  from their given relationship to m, k and the estimators for these: As  $m_{\text{new}} = 2m$  and  $k_{\text{new}} = 0.5k$ ,

$$\hat{m}_{\text{new}} = 2\hat{m} = 480, \qquad \hat{k}_{\text{new}} = 0.5\hat{k} = 5.05.$$

A request of length x should be sent to the new server if the (estimated) expected latency is smaller there than for the old server. This gives

$$480 + 5.05x < 240 + 10.1x$$

or equivalently 5.05x > 240, that is,  $x > 240/5.05 \approx 47.52$ . So all requests of length at least 48 characters should be sent to the new server, and shorter ones to the old server.

# Fråga 6

- a) The numbers on the arrows coming out of each state have to sum up to 1, so  $* = 1 \frac{1}{4} \frac{1}{4} = \frac{1}{2}$ .
- b) The transition matrix is

$$\begin{pmatrix} 1/4 & 0 & 3/4 \\ 1 & 0 & 0 \\ 1/4 & 1/2 & 1/4 \end{pmatrix}$$

c) The initial distribution of  $X_0$  is represented by the vector  $p^{(0)} = \begin{pmatrix} 0.5 & 0 & 0.5 \end{pmatrix}$ . Then

$$p^{(1)} = p^{(0)}\mathbf{P} = \begin{pmatrix} 0.5 & 0 & 0.5 \end{pmatrix} \cdot \begin{pmatrix} 1/4 & 0 & 3/4 \\ 1 & 0 & 0 \\ 1/4 & 1/2 & 1/4 \end{pmatrix} = \begin{pmatrix} 0.25 & 0.25 & 0.5 \end{pmatrix}$$

so  $P(X_1 = 2) = 0.5$  (the third element of the vector).

d) 
$$P(X_0 = 0 \mid X_1 = 2) = \frac{P(X_0 = 0, X_1 = 2)}{P(X_1 = 2)} = \frac{0.5 \cdot 0.75}{0.5} = 0.75.$$

e) We need to find  $a, b, c \in [0, 1]$  with a + b + c = 1 so that

$$(a \quad b \quad c) \cdot \begin{pmatrix} 1/4 & 0 & 3/4 \\ 1 & 0 & 0 \\ 1/4 & 1/2 & 1/4 \end{pmatrix} = (a \quad b \quad c)$$

This gives  $\frac{1}{4}a+b+\frac{1}{4}c=a$ ,  $\frac{1}{2}c=b$ ,  $\frac{3}{4}a+\frac{1}{4}c=c$ . From the third equation we get  $\frac{3}{4}a=\frac{3}{4}c$ , so a=c. Plugging this and the second equation into a+b+c=1 gives  $c+\frac{1}{2}c+c=1$ , which solves to  $c=\frac{2}{5}$ ,  $a=c=\frac{2}{5}$ ,  $b=\frac{1}{2}c=\frac{1}{5}$ . So the stationary distribution is represented by the vector

$$\begin{pmatrix} \frac{2}{5} & \frac{1}{5} & \frac{2}{5} \end{pmatrix}$$

#### Fråga 7

- a) If some emails are spam with probability p=0.2 independently, we can view this as randomly deleting the Poisson events with probability p=0.2, independently. By the thinning property, the result is still a Poisson process, with intensity  $(1-p)\lambda=0.8\cdot 2=1.6$ . The number of non-spam emails in an interval of length 8 (hours) is Poisson distributed with parameter  $8\cdot 1.6=12.8$ . So the probability that you get 12 emails is  $e^{-12.8}\cdot 12.8^{12}/12!\approx 11.1\%$ .
- b) The numbers of emails in each interval are distributed according to Poisson distributions with with parameters  $\lambda_1 = 3\lambda = 6$ ,  $\lambda_2 = 1 \cdot \lambda = 2$  and  $\lambda_3 = 2\lambda = 4$ , respectively. Furthermore, as the intervals are disjoint, the numbers are independent. Thus, by the addition theorem for the Poisson distribution, the total number of emails during the outages has the distribution  $Po(\lambda_1 + \lambda_2 + \lambda_3) = Po(12)$ .

c) The number of spam emails over time is given by a Poisson process with intensity  $\lambda \cdot p = 2 \cdot 0.2 = 0.4$  (to see this, note that we can view this as the number of all emails with the non-spam emails deleted, which happens with probability 0.8, so by the thinning property this is a Poisson process with intensity  $(1 - (1 - 0.8))\lambda = 0.2\lambda$ ). Thus the waiting time for the first event has the distribution Exp(1/0.4) = Exp(2.5), which has expectation 2.5 (hours).