

CPSC 538R PROJECT REPORT

# Lifted Parallel Tempering for Training Restricted Boltzmann Machines

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# 1 Introduction

Restricted Boltzmann Machines (RBMs) are Markov random fields with applications including density estimation, dimensionality reduction and collaborative filtering.

## 2 Restricted Boltzmann Machines

An RBM is a Markov random field where the underlying graph is a complete bipartite graph separating visible units  $\mathbf{v} = (v_1, \dots, v_n) \in \{0, 1\}^n$  from the latent hidden units  $\mathbf{h} = (h_1, \dots, h_m) \in \{0, 1\}^m$  as shown in Figure 2.

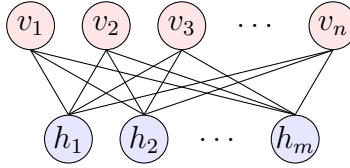


Figure 1: Graphical representation of an RBM.

Furthermore an RBM is assumed to have the Gibbs distribution given by

$$P(\mathbf{v}, \mathbf{h}) = \frac{1}{Z} \exp(-E(\mathbf{v}, \mathbf{h})), \quad (1)$$

where the energy  $E(\mathbf{v}, \mathbf{h})$  has the form,

$$\begin{aligned} E(\mathbf{v}, \mathbf{h}) &= - \sum_{i=1}^n \sum_{j=1}^m v_i w_{ij} h_j - \sum_{i=1}^n b_i v_i - \sum_{j=1}^m c_j h_j \\ &\equiv -\mathbf{v}^T \mathbf{W} \mathbf{h} - \mathbf{b}^T \mathbf{v} - \mathbf{c}^T \mathbf{h}, \end{aligned}$$

for some  $\mathbf{W} \in \mathbb{R}^{n \times m}$ ,  $\mathbf{b} \in \mathbb{R}^n$ , and  $\mathbf{c} \in \mathbb{R}^m$ . We have  $w_{i,j}$  encodes the interaction potential between  $v_i$ , and  $h_j$  and the  $b_i$ ,  $c_j$  encodes the local potential for  $v_i$  and  $h_j$  respectively.

### 2.1 RBM Likelihood

Suppose we have data  $\mathcal{D} = \{\mathbf{v}^1, \dots, \mathbf{v}^\ell\}$ , we want to find  $\mathbf{W}, \mathbf{b}, \mathbf{c}$ , that maximize the log-likelihood function given by

$$\mathcal{L}(\theta|\mathcal{D}) = \log P(\mathcal{D}|\theta)$$

As shown in §4.2 in [FI14]), the gradient of  $\mathcal{L}$  can be written as,

$$\frac{\partial \mathcal{L}}{\partial \theta} = \mathbb{E}_{P_{data}} \left( \frac{\partial E}{\partial \theta} \right) - \mathbb{E}_P \left( \frac{\partial E}{\partial \theta} \right) \quad (2)$$

where  $P_{data}$  is the empirical distribution of the data. So by taking the partial derivatives with respect to  $w_{ij}, b_i, c_j$ , we get the gradient updates given by,

$$\begin{aligned}\frac{\partial \mathcal{L}}{\partial w_{ij}} &= \mathbb{E}_{P_{data}}(v_i h_j) - \mathbb{E}_P(v_i h_j) \\ \frac{\partial \mathcal{L}}{\partial b_i} &= \mathbb{E}_{P_{data}}(v_i) - \mathbb{E}_P(v_i) \\ \frac{\partial \mathcal{L}}{\partial c_j} &= \mathbb{E}_{P_{data}}(h_j) - \mathbb{E}_P(h_j)\end{aligned}$$

Hinton [Hin02] showed that this update also emerges by minimizing the KL divergence between the data distribution and the equilibrium distribution over the visible variables.

Therefore in order to optimize for our model parameters, we need to compute the above expectations, which in general are intractable. We thus resort to approximating them using MCMC. The natural MCMC method for us to use is Gibbs sampling as the bipartite structure of our graphical model tells us that  $v_i \perp\!\!\!\perp v_k \mid \mathbf{h}$  and  $h_j \perp\!\!\!\perp h_l \mid \mathbf{v}$ . Thus we get the following conditional distributions

$$P(v_i = 1 \mid \mathbf{h}) = \sigma\left(b_i + \sum_j w_{ij} h_j\right), \quad (3)$$

$$P(h_j = 1 \mid \mathbf{v}) = \sigma\left(c_j + \sum_i w_{ij} v_i\right), \quad (4)$$

where  $\sigma(x) = (1 + e^{-x})^{-1}$  is the sigmoid function.

### 3 Contrastive Divergence

Contrastive Divergence (*CD*) is the standard way for training RBMs [Hin02]. *CD* approximates  $\mathbb{E}_P\left(\frac{\partial E}{\partial \theta}\right)$  with  $\langle \frac{\partial E}{\partial \theta} \rangle^k$ , the Monte-Carlo average after sampling the hidden and visible units using  $k$  Gibbs Sampling updates.

CD begins by initializing the visible units to a given data vector. Then, in what is known as a "negative phase", the hidden units are updated via Gibbs sampling, with the visible units clamped. This is followed by a "positive" phase, where the visible units are updated with Gibbs sampling, while the hidden units are clamped using

$$P(h_j = 1 \mid \mathbf{v}) = \sigma\left(c_j + \sum_i w_{ij} v_i\right).$$

The resulting state of the visible and hidden units gives us the sample  $\langle \frac{\partial E}{\partial \theta} \rangle^0$ . One can repeat this process, updating the visible and hidden units one layer at a time to obtain the estimate  $\langle \frac{\partial E}{\partial \theta} \rangle^k$  for any arbitrary  $k$  [Hin12]. When  $k > 1$ , we will call this algorithm CD- $k$ .

To obtain an unbiased estimator of  $\mathbb{E}_P\left(\frac{\partial E}{\partial \theta}\right)$ , this process must run to equilibrium, requiring many iterations. However, this is computationally expensive. Hinton [Hin02] notes that after 1 iteration, it can be seen in which direction the model is wandering. Hinton proposes the simplified learning rule

$$\Delta \theta = \left\langle \frac{\partial E}{\partial \theta} \right\rangle^0 - \left\langle \frac{\partial E}{\partial \theta} \right\rangle^1$$

CD guarantees that weights are updated in the correct direction, albeit with wrong magnitudes [Tie08].

By using a small  $k$ , one gets biased estimates of the probability distribution underlying the model. Thus, one does not necessarily reach a minimum of the negative log likelihood during training. More problematically, Fischer and Igel [?] showed that the log-likelihood can even decrease as more iterations of gradient ascent are taken.

### 3.1 PCD

Although CD works for simple models, it can lead to bad models if the mixing rate of the Markov Chain is low. Tieleman [Tie08] notes that each time a Markov Chain is reinitialized with a data vector to approximate  $\mathbb{E}_P\left(\frac{\partial E}{\partial \theta}\right)$ , valuable information about the model is lost. Tieleman proposes the use of persistent chains. Rather than being initialized to a data vector, chains "persist", beginning with the final state of a previous Markov Chain. This variant of contrastive divergence is called Persistent contrastive divergence (PCD).

However, Tieleman, and Hinton [TH09] note that PCD still relies on approximating the gradient of the KL divergence. PCD minimizes a difference of KL divergences

$$KL(P_{data} || P) - KL(P^t || P),$$

where  $P_t$  is the distribution of persistent Markov chain after time  $t$ . The second term is an error incurred for not running the chain to equilibrium.

Desjardins [DCB<sup>+</sup>10] notes that although the error term is small relative to the desired term in the beginning of training, as  $t$  increases, the error term dominates due to the desired term vanishing, and the mixing of the chain decreasing. PCD encourages the model to settle to recently visited modes in successive iterations, leading to few deep minima in the energy.

## 4 Tempering Methods

As we have discussed previously, the problem of learning the model for an RBM can be boiled down to effectively we can approximate  $\mathbb{E}_P\left(\frac{\partial E}{\partial \theta}\right)$ . In order to do this, we had to resort to Gibbs Sampling. Gibbs sampling performs local updates to generate a Markov Chain with stationary distribution  $P(\mathbf{v}, \mathbf{h})$  which is complex and multi-modal. Despite the theoretical guaranteed convergence of Gibbs's sampling, the local nature of the updates means that chain can have a very difficult time traversing the modes of  $P$  and has a tendency to get trapped exploring the local areas of our sample space of high probability. This means that certain modes are over-represented, while other are under-represented and leading to a poor approximation of our expectation.

Parallel tempering (PT) is a class of Monte Carlo methods first introduced by Geyer in [?] that improve the mixing of our MCMC chain, by simultaneously running  $N$  MCMC chains in parallel but at different temperatures. The higher temperature chains focus on exploration of the state space, while the room temperature chain focus on the accuracy of samples. In order for there to be communication between the hot and room temperature chains, we periodically propose swaps between the chains in such a way that does not propose

## 4.1 Parallel Tempering

Let  $P$  be a Gibbs distribution over some sample space  $\Omega$  of the form,

$$P(x) = \frac{1}{Z} \exp(-E(x)).$$

Note that  $E$  can be that of an RBM but does not have to be. We define our inverse-temperature space as  $\mathcal{B} = [\beta_{\min}, 1]$ , for  $0 \leq \beta_{\min} < 1$ . Given  $\beta \in \mathcal{B}$ , let  $P^{(\beta)}$  the probability distribution identified with the un-normalized density  $P(x)^\beta$ , which we can write

$$P^{(\beta)}(x) = \frac{1}{Z(\beta)} \exp(-\beta E(x)),$$

where  $Z(\beta)$  is the partition function.

Given a sequence of inverse temperatures  $0 \leq \beta_{\min} = \beta_0 < \beta_1 < \dots < \beta_N = 1$ , we define the measure  $\tilde{P}$  on  $\Omega^{N+1}$  by,

$$\tilde{P} = P^{(\beta_0)} \times \dots \times P^{(\beta_N)}.$$

In parallel tempering we construct a Markov chain  $\tilde{X}_n = (X_n^0, \dots, X_n^N)$  on  $\Omega^N$  with stationary distribution  $\tilde{\pi}$ . We update  $X_n$  by alternate between two different type of dynamics: exploration and communication.

During the exploration phase, we allow for each  $X_n^i$  to independently explore  $\Omega$  according to  $P^{(\beta_i)}$  via our favourite MCMC algorithm such as Gibbs Sampling. Since MCMC updates of each component  $i$  leaves  $P^{(\beta_i)}$  invariant, we have the exploration phase leaves  $\tilde{P}$  invariant.

During the communication phase, we propose a swap between the  $\beta_i$  and  $\beta_j$  components of  $\tilde{X}_n$ , in other words we swap  $X_n^i$  and  $X_n^j$  in  $\tilde{X}_n$ . To simplify notation, given  $x = (x^0, \dots, x^i, \dots, x^j, \dots, x^N)$ , we will define  $x_{(i,j)} = (x^0, \dots, x^j, \dots, x^i, \dots, x^N)$ . So the proposed state during the communication phase is  $(\tilde{X}_n)_{(i,j)}$ . In order to keep  $\tilde{X}_n$  stationary with respect to  $\tilde{P}$ , we can make this update reversible by accepting this proposed swap according to the Metropolis acceptance,

$$\begin{aligned} \alpha &= 1 \wedge \frac{\tilde{P}((\tilde{X}_n)_{(i,j)})}{\tilde{P}(\tilde{X}_n)} \\ &= 1 \wedge \frac{P^{(\beta_i)}(X_n^j) P^{(\beta_j)}(X_n^i)}{P^{(\beta_i)}(X_n^i) P^{(\beta_j)}(X_n^j)} \\ &= 1 \wedge \exp [(\beta_j - \beta_i)(E(X_n^j) - E(X_n^i))] \end{aligned}$$

We refer to the sequence  $(X_n^i)_{n \geq 0}$  as the  $i^{th}$  chain, the sequence of states at a particular temperature  $\beta_i$ . We refer to the sequence  $R^i = (X_0^i, \dots, X_n^i, X_{n+1}^j, X_{n+2}^j, \dots)$  as the  $i^{th}$  replica. The  $i^{th}$  replica follows the trajectory of the  $i^{th}$  state as it's temperature changes. In this case, there was a successful swap at time  $n$  between temperatures at states  $\beta_i$  and  $\beta_j$ .

Since the likelihood of acceptance is proportional to the change in temperature, we usually restrict to our swaps to nearest neighbours.

## 4.2 Lifted Parallel Tempering

In parallel tempering because of detailed balance condition on the temperature the replica must be proposed a swap that can potentially raise or lower the temperature. This forces the replica to traverse the  $\mathcal{B}$  in a random walk fashion and leads to diffusive behaviour. This is not favourable as the goal is to have the information from the high temperature states reach room temperature as quickly as possible, without getting distracted along the way. That is where Lifted Parallel Tempering (LPT) comes to the rescue. LPT is a non-reversible version of PT aims to remedy this diffusive behaviour. Before we introduce it, we will need to introduce the Generalized Metropolis Hastings framework as outlined in [Wu17].

### 4.2.1 Generalized Metropolis-Hastings Framework

Suppose we probability measure  $P$  over measurable space  $(\Omega, \mathcal{F})$  and  $S : (\Omega, \mathcal{F}) \rightarrow (\Omega, \mathcal{F})$  is a measure preserving involution for  $P$ . Suppose we have Metropolis-proposal kernel  $Q(y|x)$ , we will use this to construct a new proposal  $\tilde{Q}$  given by  $\tilde{Q}(y|x) = Q(S(dy)|x)$ . We will then accept this proposal with probability

$$\alpha(x, y) = \frac{\pi(y)\tilde{Q}(y|x)}{\pi(x)\tilde{Q}(x|y)}.$$

This induces the reversible Metropolis transition kernel,

$$\tilde{K}(y|x) = \alpha(x, y)\tilde{Q}(y|x) + \left(1 - \int \alpha(x, y)\tilde{Q}(y|x)dy\right) \delta_x(y).$$

which has stationary distribution  $P$ . Since  $\tilde{K}$  and  $S$  preserve  $P$ , then so does  $K \equiv S \circ \tilde{K}$ , which is given by,

$$K(y|x) = \alpha(x, y)Q(y|x) + \left(1 - \int \alpha(x, y)Q(y|x)dy\right) \delta_{S(x)}(y).$$

Effectively, we make a proposal according to  $Q$ , which is accepted with probability  $\alpha$  and apply  $S$  if the proposal is rejected. This in general produces an non-reversible scheme as the composition of two reversible kernels is not necessarily reversible.

### 4.2.2 Lifted Parallel Tempering Algorithm

Suppose we have a partition  $P = \{\beta_0, \dots, \beta_N\}$  for our inverse-temperature space  $\mathcal{B} = [\beta_{\min}, 1]$ , where

$$0 \leq \beta_{\min} = \beta_0 < \dots \beta_N = 1.$$

Similar to PT, we have  $N$  additional chains running simultaneously at inverse temperatures according to  $P$ , with joint distribution  $\tilde{\pi}^P = \pi^{(\beta_0)} \times \dots \times \pi^{(\beta_N)}$ . We denote the state at time  $n$  by  $X_n = (X_n^0, \dots, X_n^N)$ , where  $X^i$  is at temperature  $\beta_i$ . In addition we introduce the lifted parameters  $(\eta_n, \varepsilon_n) \in \{0, 1, \dots, N\} \times \{\pm 1\}$ . Here  $\eta_n$  represents temperature index of the  $\eta_0$ -th replica and  $\varepsilon_n$  represents the direction the replica will attempt to swap, i.e, the next proposed swap will be between  $X_n^{\eta_n}$  and  $X_n^{\eta_n + \varepsilon_n}$ .

To deal with the boundary cases, note that the replica is at temperature  $\beta_0$  or  $\beta_N$  then it must be proposed a swap with the state at temperature  $\beta_1$  and  $\beta_{N-1}$  respectively. So we have

$$(\eta_n, \varepsilon_n) \in \{0, 1, \dots, N\} \times \{\pm 1\} \setminus \{(0, -1), (N, 1)\} \equiv \mathcal{X}.$$

Let  $Z_n = (X_n, \eta_n, \varepsilon_n)$  be a Markov chain in enlarged space will be  $\Omega^{N+1} \times \mathcal{X}$ . Suppose we are working with the deterministic proposal

$$Q(y|z) = \delta_{\phi(z)}(y)$$

given by,

$$\phi(X, \eta, \varepsilon) = \begin{cases} (X_{(\eta, \eta+\varepsilon)}, \eta + \varepsilon, \varepsilon) & 0 < \eta + \varepsilon < N, \\ (X_{(\eta, \eta+\varepsilon)}, \eta + \varepsilon, -\varepsilon) & \text{Otherwise.} \end{cases}$$

We also define the involution  $S : \Omega^{N+1} \times \mathcal{X}$  given by,

$$S(X, \eta, \varepsilon) = \begin{cases} (X, \eta, -\varepsilon) & 0 < \eta + \varepsilon < N, \\ (X, \eta, \varepsilon) & \text{Otherwise,} \end{cases}$$

which preserves the measure  $\tilde{\pi}^P \times \text{Unif}(\mathcal{X})$ . Together  $Q$  and  $S$  give us the recipe for the lifted parallel tempering algorithm. Details are outlined in [Wu17].

### 4.3 Deterministic Even/Odd Algorithm

### 4.4 PT for RBM

## 5 Experiments

## 6 References

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