

DATA

Dataset Description

This study draws on the "Medical Cost" dataset from the esteemed data science community, Kaggle.com. The dataset, curated from Brett Lantz's seminal text "Machine Learning with R," comprises the medical insurance expenses of 1338 individuals. Additionally, the dataset boasts a comprehensive range of features, including 3 categorical and 4 quantitative variables that will be elaborated in the ensuing table. The richness and complexity of this dataset ensure fertile ground for insightful analysis and interpretation. The dataset submitted from an open data source is available at the following website <https://www.kaggle.com/mirichoi0218/insurance>.

Loading Libraries

To achieve my objectives, I installed a selected set of packages and loaded several indispensable libraries. The code snippet below was instrumental in executing this process, propelling me closer to the end goal.

```
##{r}
install.packages("dplyr")
install.packages("ggplot2")
library(tidyverse)
library(corrplot)
library(class)
library(MASS)
library(ggplot2)
install.packages("leaps")
library(leaps)
```

Loading the Dataset

The following code was used to load the dataset, it was loaded from an Excel CSV file, and the dataset's correct loading was checked using a data frame.

```
##{r}
Insurance <- read.csv("C:/Users/Modupe Olayinka/OneDrive - University of Louisiana
Lafayette/Desktop/dataset/health insurance.csv", head=
TRUE)
is.data.frame(Insurance)
```

The first six rows of the dataset are displayed in the function head ()

```
##{r}
head(Insurance)
```

	age	sex	bmi	children	smoker	region	charges
	<dbl>	<chr>	<dbl>	<int>	<chr>	<chr>	<dbl>
1	19	female	27.900	0	yes	southwest	16884.924
2	18	male	33.770	1	no	southeast	1725.552
3	28	male	33.000	3	no	southeast	4449.462
4	33	male	22.705	0	no	northwest	21984.471
5	32	male	28.880	0	no	northwest	3866.855
6	31	female	25.740	0	no	southeast	3756.622

6 rows

To ascertain the dataset's dimensions, the function `dim ()` was utilized. Our analysis revealed that the dataset comprises 7 columns and 1338 rows.

```
{r}
dim(Insurance)

[1] 1338 7
```

The original structures

The output, shown below, showed that the original dataset included three modes: num, int, and Chr. I used the `str ()` function to identify the variable modes in the dataset. It is important to note that before the study begins, the dataset should only receive a light cleaning.

```
{r}
str(Insurance)

'data.frame': 1338 obs. of 7 variables:
 $ age      : int  19 18 28 33 32 31 46 37 37 60 ...
 $ sex      : chr  "female" "male" "male" "male" ...
 $ bmi      : num   27.9 33.8 33 22.7 28.9 ...
 $ children : int   0 1 3 0 0 0 1 3 2 0 ...
 $ smoker   : chr   "yes" "no" "no" "no" ...
 $ region   : chr  "southwest" "southeast" "southeast" "northwest" ...
 $ charges  : num  16885 1726 4449 21984 3867 ...
```

The `colnames ()` method was employed to exhibit the list of columns in the dataset, as exemplified below.

```
{r}
colnames(Insurance)

[1] "age" "sex" "bmi" "children" "smoker" "region" "charges"
```

Data Cleaning

First, I looked at the dataset to check if each column had the same number of values. Then, I used the following code to search for any missing data, but there wasn't any missing data found.

```
{r}
colSums(sapply(Insurance, is.na))

age sex bmi children smoker region charges
0 0 0 0 0 0 0
```

The variables "age" and "children" were recorded as whole numbers. To make it easier to analyze them, we changed the mode from "int" to "num" using the function called "as. numeric ()". The variables "sex," "smoker," and "region" were saved as text so we could easily change them during our investigation. We

will use the function called "as. factor ()" to change them from being descriptive to being a measurable value.

```
```{r}
Insurance$sex <- as.factor(Insurance$sex)
is.factor(Insurance$sex)
Insurance$smoker <- as.factor(Insurance$smoker)
is.factor(Insurance$smoker)
Insurance$region <- as.factor(Insurance$region)
is.factor(Insurance$region)
Insurance$age <- as.numeric(Insurance$age)
is.numeric(Insurance$age)
Insurance$children<- as.numeric(Insurance$children)
is.numeric(Insurance$children)
```

[1] TRUE
[1] TRUE
[1] TRUE
[1] TRUE
[1] TRUE
```

Cleaned Data

This is the updated result of the "str()" command following data cleaning.

```
Created dataset
```{r}
str(Insurance)

'data.frame': 1338 obs. of 7 variables:
 $ age : num 19 18 28 33 32 31 46 37 37 60 ...
 $ sex : Factor w/ 2 levels "female","male": 1 2 2 2 2 1 1 1 2 1 ...
 $ bmi : num 27.9 33.8 33 22.7 28.9 ...
 $ children : num 0 1 3 0 0 0 1 3 2 0 ...
 $ smoker : Factor w/ 2 levels "no","yes": 2 1 1 1 1 1 1 1 1 1 ...
 $ region : Factor w/ 4 levels "northeast","northwest",...: 4 3 3 2 2 3 3 2 1 2 ...
 $ charges : num 16885 1726 4449 21984 3867 ...
```

## Variable Description

The cleaned data used to analyze this dataset is described in the table below. The descriptions were taken from the website's author description.

<https://www.kaggle.com/mirichoi0218/insurance>.

TABLE 1 - VARIABLE DESCRIPTION

Column Name	Independent/ Dependent	Mode	Description
age	Ind	Numeric	Age of primary beneficiary
sex	Ind	Factor	Insurance contractor gender: 2 levels (female, male)
bmi	Ind	Numeric	Body mass index
children	Ind	Numeric	Number of children/dependents covered by health insurance
smoker	Ind	Factor	Smoking: 2 levels (yes, no)
region	Ind	Factor	Beneficiary's residential area in the US: 4 levels (northeast, southeast, northwest, southwest)
Charges	Dep	Numeric	Individual medical costs billed by health insurance

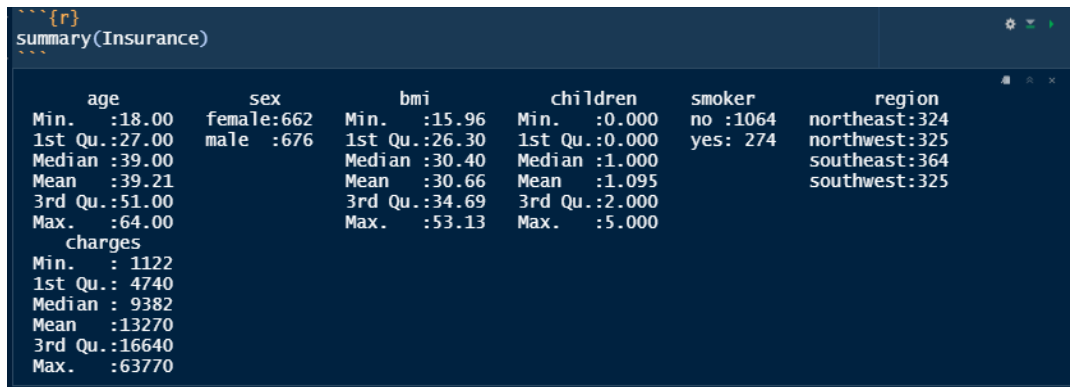
## Expectations

This project's primary goal is to forecast the medical costs that health insurance providers will charge. The cost of providing coverage to a person is estimated over a long period by insurance companies. The goal is to determine whether some people will need medical care based on an analysis of the data that is already available utilizing critical variables like BMI and smoking behaviors. Insurance providers can adjust their premiums using this information.

It is anticipated through data analysis that elements like BMI and smoking habits will have a substantial impact on insurance costs. Smokers and individuals with higher BMIs are likelier to have higher premiums than non-smokers. Several graphical methods, including bar graphs, plots, and heatmaps, will be employed to examine the dataset efficiently. Different methods, including linear regression, best subset, ridge, and lasso regressions, will be used to accurately estimate insurance costs. Techniques like K-fold cross-validation and validation set will also be employed.

## Data Analysis

I first collected a data summary before I started to analyze the dataset.



This summary gives a broad overview of how the data is split among the different features. It demonstrates that a primary beneficiary must be at least 18 years old. Additionally, the summary notes that there are equally as many males and female beneficiaries in the sample. We may also see that each primary beneficiary may have a maximum of five dependents.

Regarding smoking, the summary reveals that there are significantly more non-smokers than smokers in the dataset.

## Categorical Variables

I utilized barplots to investigate the categorical variables, Sex, Smoker, and Region, because I thought they would more accurately depict these factors.

Beneficiaries Count by Sex



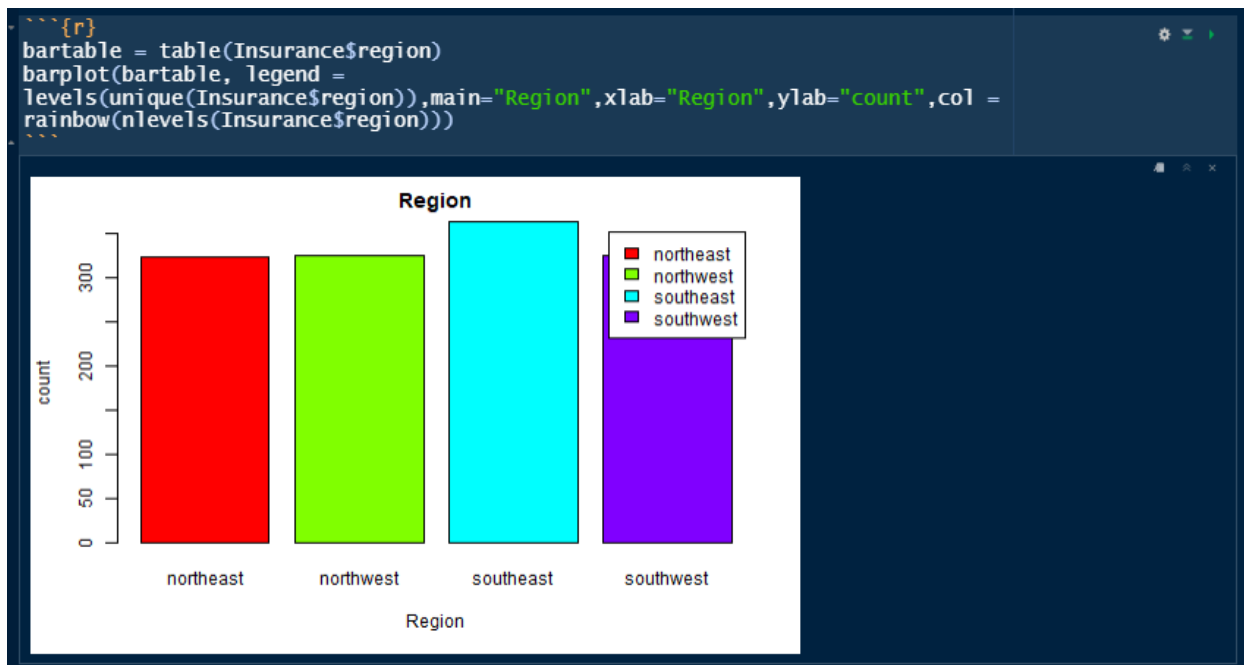
As seen in the summary command, the plot indicates that the dataset has an equal number of females and males.

### Smoking Habits



The barplot shows that in the dataset, there are noticeably more non-smokers than smokers.

### Region



We can see from the region barplot that the beneficiaries are roughly evenly divided throughout the various areas, with somewhat more people in the southeast.

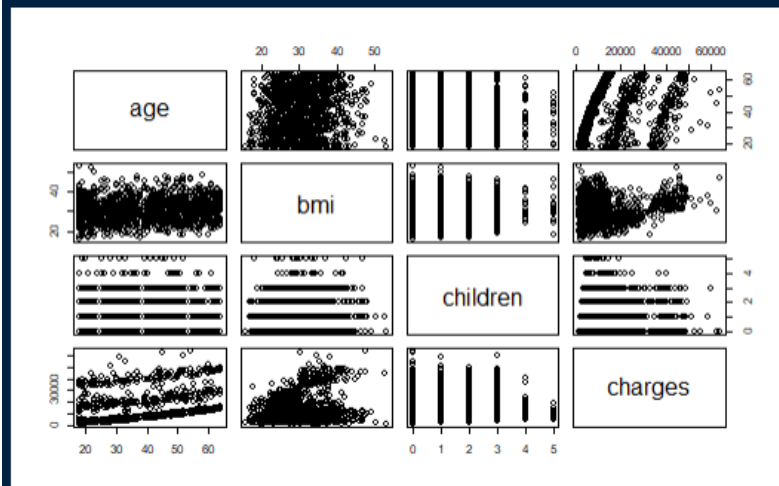
### Continuous Variables

```
{r}
cor(Insurance[, c(1,3,4,7)])
```

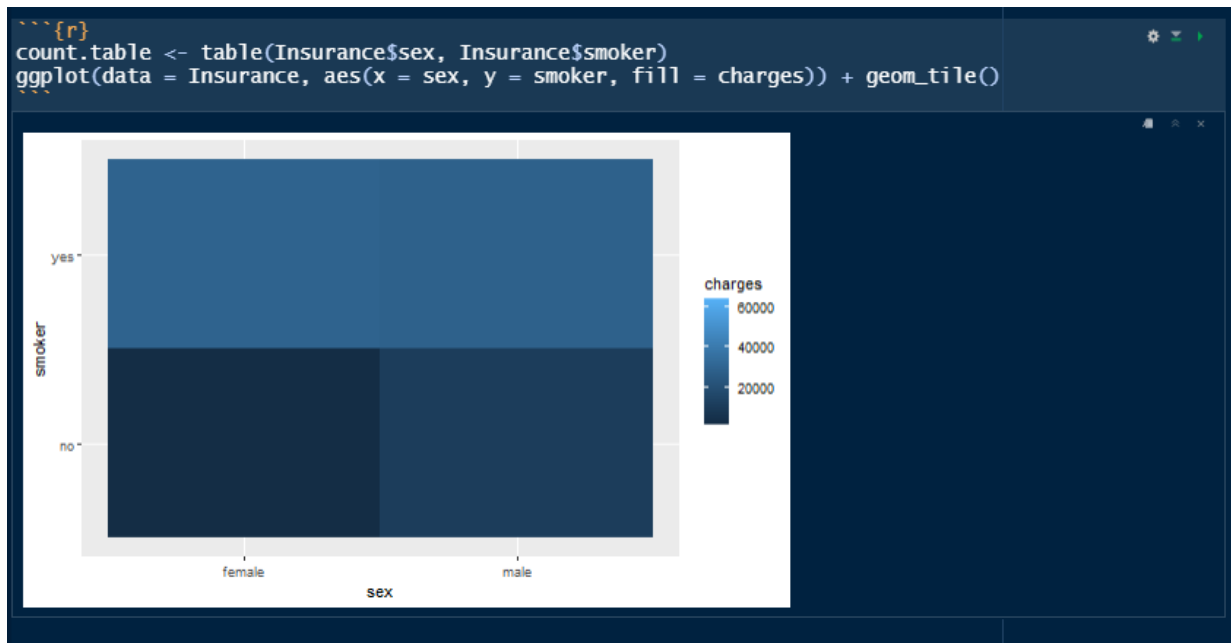
	age	bmi	children	charges
age	1.0000000	0.1092719	0.0424690	0.29900819
bmi	0.1092719	1.0000000	0.0127589	0.19834097
children	0.0424690	0.0127589	1.0000000	0.06799823
charges	0.2990082	0.1983410	0.06799823	1.00000000

To find the correlation coefficient for each of my continuous variables, I used the `cor` command. The results showed that no two variables have a meaningful association. To provide a visual depiction, I also plotted the variables using the `pairs` command. The `pairs` command verified the output from the `cor` command.

```
{r}
pairs(Insurance[, c(1,3,4,7)])
```

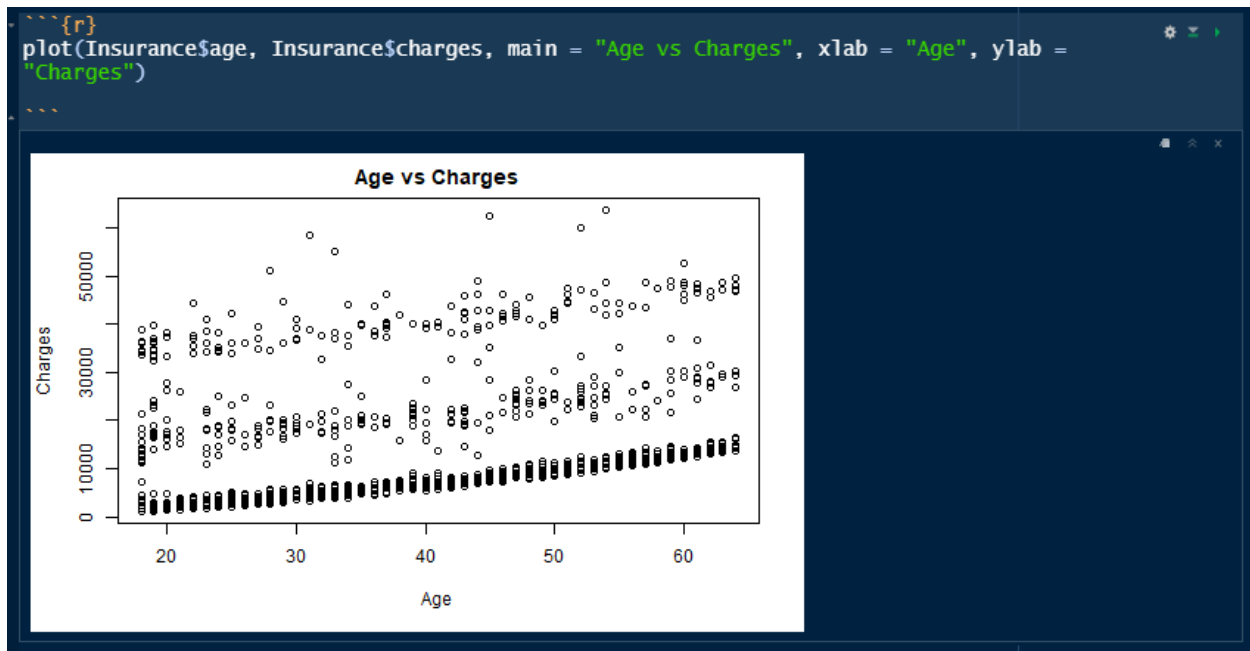


I used the `pairs()` function to create a scatter plot matrix for the numerical variables.



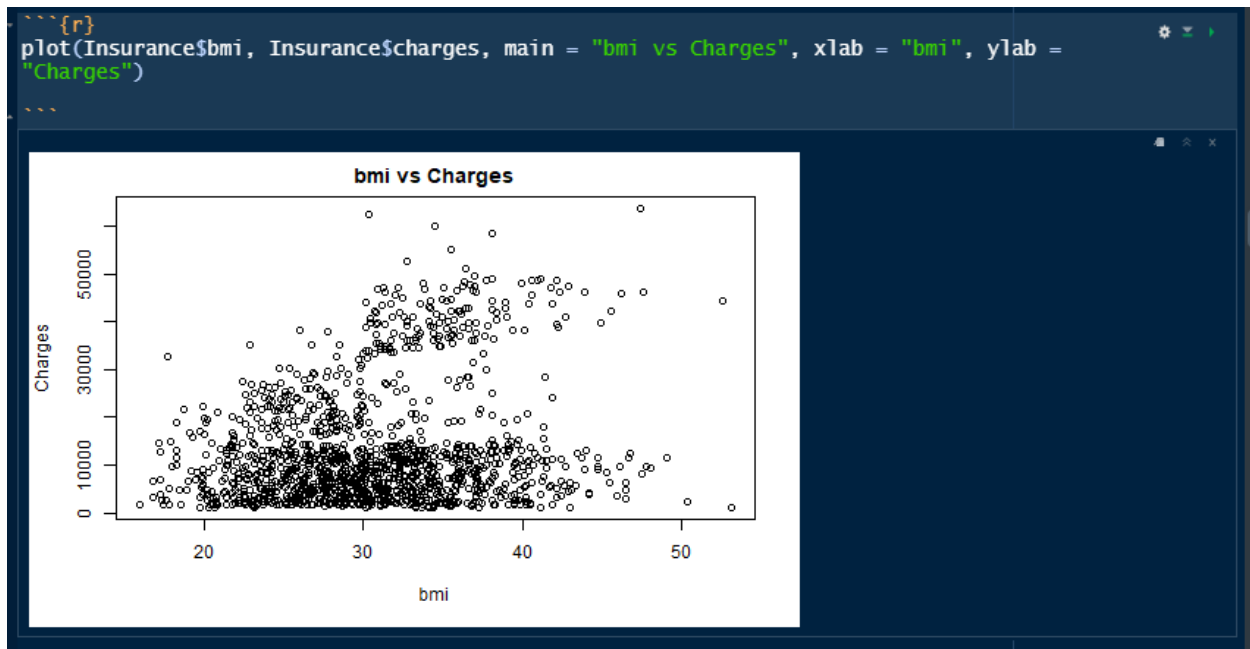
We can see from the heatmap that smokers, whether they are male or female, typically pay more for insurance than nonsmokers do. Additionally, females appear to pay less for insurance than males do for nonsmokers. This heatmap, which only applies to nonsmokers, supports the findings from the sex vs. charges barplot.

Age vs. Charges

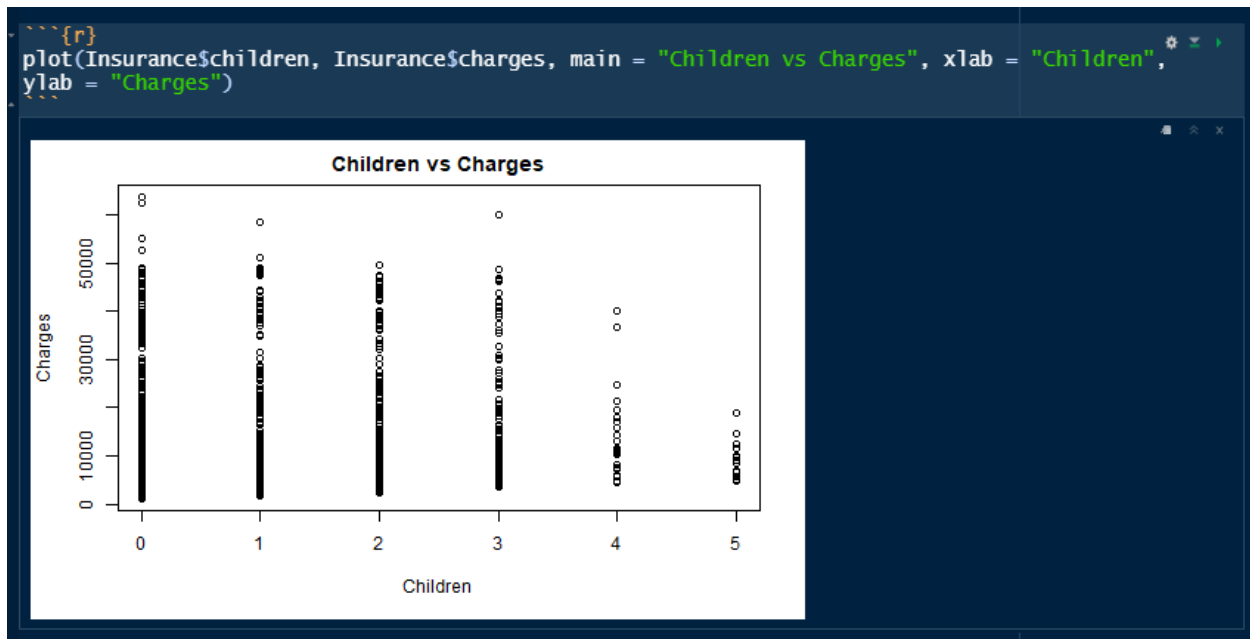


bmi vs Charges





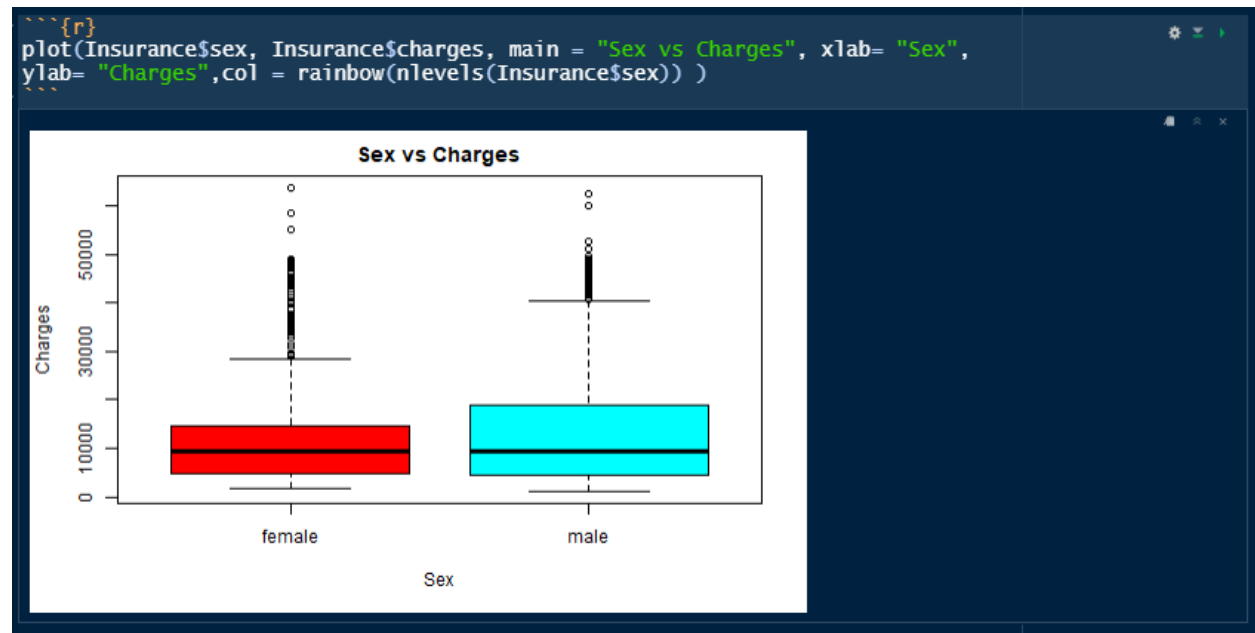
Children vs Charges



This graph demonstrates that when the number of dependents increases to four or more, charges tend to decrease.

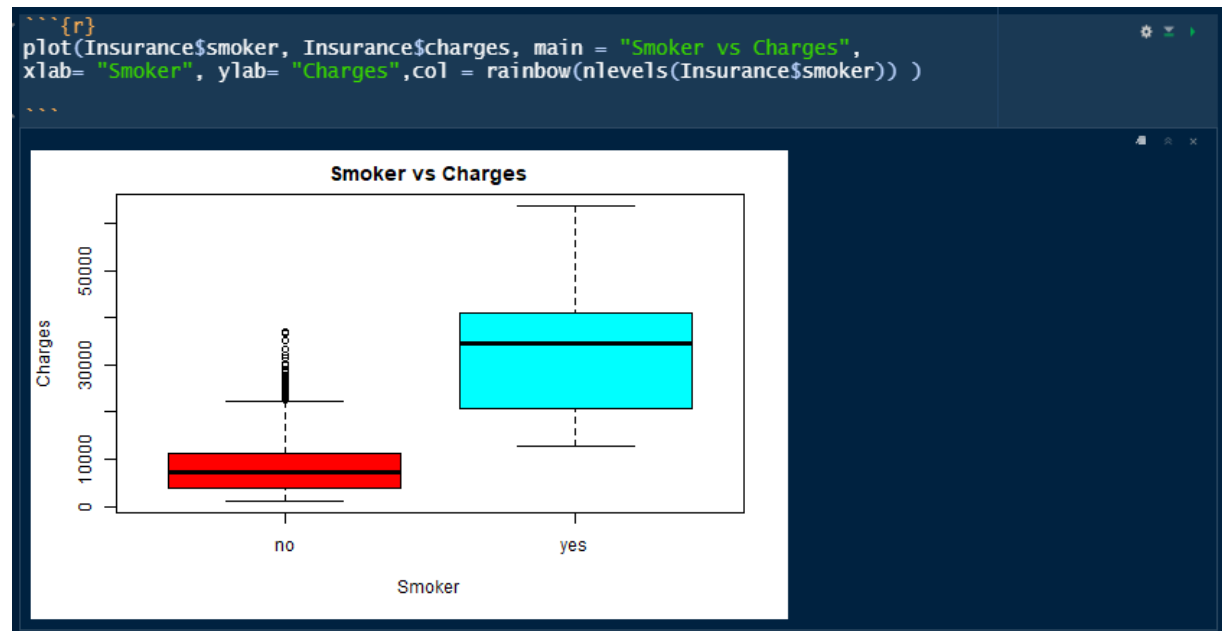
## Continuous Variable and Categorical Variables

Sex vs. Charges



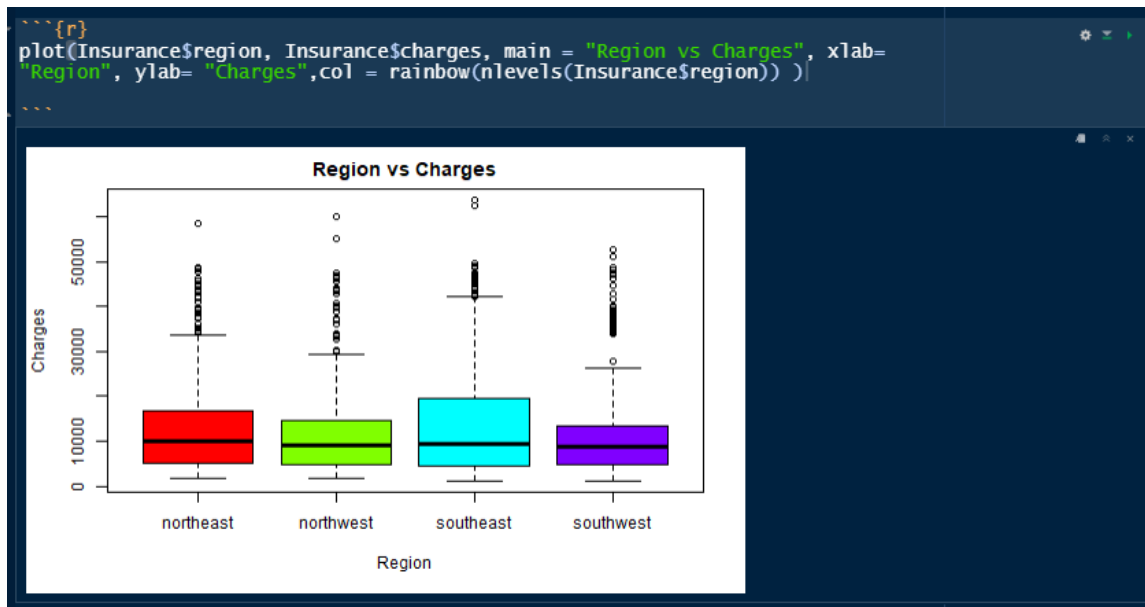
This graphic demonstrates that men typically have higher insurance than women do. To validate it, we will need to check into it more thoroughly.

Smoker vs. Charges



Smokers pay much more for insurance than non-smokers, as would be expected. This seems reasonable, given that smoking can cause various major health problems

## Region vs. Charges



The insurance payment methods are roughly equivalent. However, it appears that southeasters are charged more than persons from other parts of the country.

## Models

### 1. Multiple Linear Regression

I will begin by using all variables in multiple linear regression before determining which ones are statistically significant based on their corresponding p-values.

### Model 1

```
##{r}
insurance.lm <- lm(charges~., data = Insurance)
summary(insurance.lm)
##
```

Call:  
lm(formula = charges ~ ., data = Insurance)

Residuals:

Min	1Q	Median	3Q	Max
-11304.9	-2848.1	-982.1	1393.9	29992.8

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-11938.5	987.8	-12.086	< 2e-16 ***
age	256.9	11.9	21.587	< 2e-16 ***
sexmale	-131.3	332.9	-0.394	0.693348
bmi	339.2	28.6	11.860	< 2e-16 ***
children	475.5	137.8	3.451	0.000577 ***
smokeryes	23848.5	413.1	57.723	< 2e-16 ***
regionnorthwest	-353.0	476.3	-0.741	0.458769
regionsoutheast	-1035.0	478.7	-2.162	0.030782 *
regionsouthwest	-960.0	477.9	-2.009	0.044765 *

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6062 on 1329 degrees of freedom  
Multiple R-squared: 0.7509, Adjusted R-squared: 0.7494  
F-statistic: 500.8 on 8 and 1329 DF, p-value: < 2.2e-16

Age, BMI, children, and smokers are statistically significant factors according to this model. We will thus solely use those variables to fit another model. Despite having p-values below 0.05, the regions in the southeast and southwest will not be used at this time.

## Model 2

```
{r}
insurance.lm <- lm(charges~.-sex-region, data = Insurance)
summary(insurance.lm)
```

Call:  
lm(formula = charges ~ . - sex - region, data = Insurance)

Residuals:

Min	1Q	Median	3Q	Max
-11897.9	-2920.8	-986.6	1392.2	29509.6

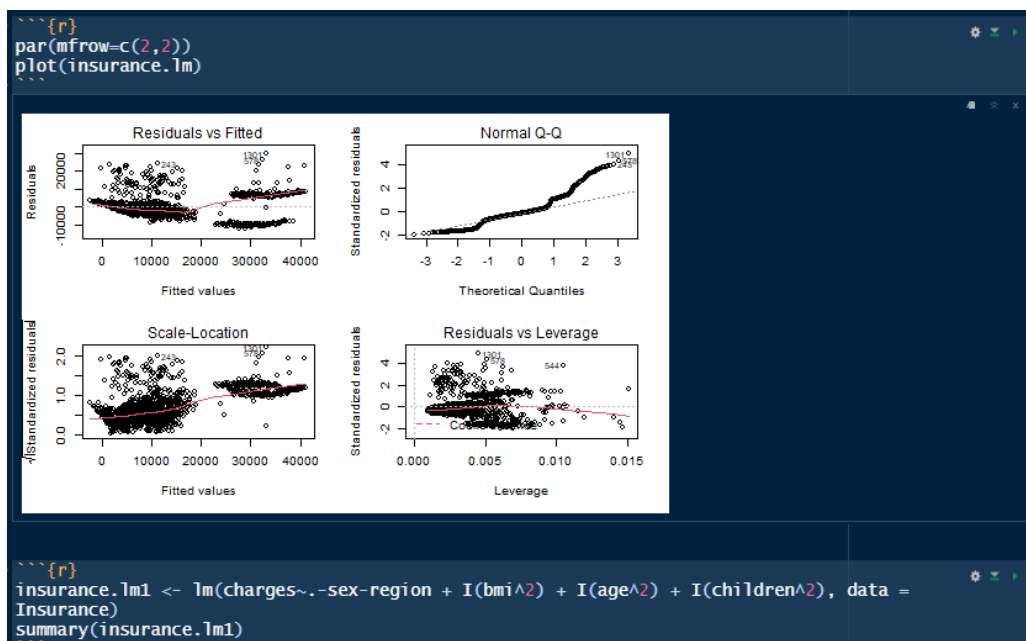
Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-12102.77	941.98	-12.848	< 2e-16 ***
age	257.85	11.90	21.675	< 2e-16 ***
bmi	321.85	27.38	11.756	< 2e-16 ***
children	473.50	137.79	3.436	0.000608 ***
smokeryes	23811.40	411.22	57.904	< 2e-16 ***

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6068 on 1333 degrees of freedom  
Multiple R-squared: 0.7497, Adjusted R-squared: 0.7489  
F-statistic: 998.1 on 4 and 1333 DF, p-value: < 2.2e-16

Despite having a slightly lower R-squared than the previous model, this model reveals that all the variables included are statistically significant.



The residual vs. fitted graphic demonstrates a pattern that supports the data's nonlinearity. The residual vs. leverage plot and the scale-location plot show evidence of high high-leverages and outliers, respectively.

To see if the model can be enhanced, we will then alter our variables.

### Model 3

```
##{r}
insurance.lm1 <- lm(charges~.-sex-region + I(bmi^2) + I(age^2) + I(children^2), data = Insurance)
summary(insurance.lm1)
```

Call:  
lm(formula = charges ~ . - sex - region + I(bmi^2) + I(age^2) + I(children^2), data = Insurance)

Residuals:

Min	1Q	Median	3Q	Max
-10551	-3114	-1196	1702	30359

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-13518.329	3498.607	-3.864	0.000117 ***
age	-87.357	82.479	-1.059	0.289726
bmi	792.804	206.940	3.831	0.000134 ***
children	1272.677	371.985	3.421	0.000642 ***
smokeryes	23813.533	408.529	58.291	< 2e-16 ***
I(bmi^2)	-7.542	3.251	-2.320	0.020496 *
I(age^2)	4.322	1.028	4.204	2.8e-05 ***
I(children^2)	-185.366	100.799	-1.839	0.066142 .

---  
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6021 on 1330 degrees of freedom  
Multiple R-squared: 0.7541, Adjusted R-squared: 0.7528  
F-statistic: 582.7 on 7 and 1330 DF, p-value: < 2.2e-16

This transformation produced the highest R-squared and adjusted R-squared value after experimenting with other transformations. However,  $I(\text{children}^2)$ 's p-value is higher than 0.05, indicating that this variable is not statistically significant. Therefore, we shall fit a model without it.

### Model 4

```
##{r}
insurance.lm1 <- lm(charges~.-sex-region + I(bmi^2) + I(age^2) , data = Insurance)
summary(insurance.lm1)
```

Call:  
lm(formula = charges ~ . - sex - region + I(bmi^2) + I(age^2), data = Insurance)

Residuals:

Min	1Q	Median	3Q	Max
-10532	-3085	-1211	1671	30071

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-13808.067	3498.184	-3.947	8.32e-05 ***
age	-57.539	80.942	-0.711	0.477289
bmi	788.095	207.109	3.805	0.000148 ***
children	641.361	143.373	4.473	8.36e-06 ***
smokeryes	23845.198	408.531	58.368	< 2e-16 ***
I(bmi^2)	-7.449	3.253	-2.289	0.022210 *
I(age^2)	3.957	1.010	3.920	9.32e-05 ***

---  
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6026 on 1331 degrees of freedom  
Multiple R-squared: 0.7535, Adjusted R-squared: 0.7524  
F-statistic: 678 on 6 and 1331 DF, p-value: < 2.2e-16

Although this model has a lower R-squared, we will keep it for now and investigate other models to improve how well they fit our dataset.

## 2. Best Subset Regression

```
##{r}
regfit.full <- regsubsets(charges~., Insurance)
summary(regfit.full)
```

Subset selection object  
Call: regsubsets.formula(charges ~ ., Insurance)  
8 Variables (and intercept)

	Forced in	Forced out
age	FALSE	FALSE
sexmale	FALSE	FALSE
bmi	FALSE	FALSE
children	FALSE	FALSE
smokeryes	FALSE	FALSE
regionnorthwest	FALSE	FALSE
regionsoutheast	FALSE	FALSE
regionsouthwest	FALSE	FALSE

1 subsets of each size up to 8  
Selection Algorithm: exhaustive

	age	sexmale	bmi	children	smokeryes	regionnorthwest	regionsoutheast	regionsouthwest
1 ( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
2 ( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
3 ( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
4 ( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
5 ( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
6 ( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
7 ( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
8 ( 1 )	" "	" "	" "	" "	" "	" "	" "	" "

```
##{r}
reg.summary <- summary(regfit.full)
reg.summary$rsq
```

[1] 0.6197648 0.7214008 0.7474772 0.7496945 0.7501113 0.7507814 0.7508839 0.7509130

```
##{r}
par(mfrow = c(2,2))
plot(reg.summary$rss ,xlab="Number of Variables ",ylab="RSS", type="l")
plot(reg.summary$adjr2 ,xlab="Number of Variables ", ylab="Adjusted RSq",type="l")
which.max(reg.summary$adjr2)
```

[1] 6

```

{r}
par(mfrow = c(2,2))
plot.new()
points(6, reg.summary$adjr2[6], col = "red", cex = 2, pch = 20)
plot(reg.summary$cp, xlab = "Number of Variables", ylab = "Cp", type = "l")
which.min(reg.summary$cp)

```

[1] 6

```

{r}
par(mfrow = c(2,2))
plot.new()
points(6, reg.summary$cp[6], col = "red", cex = 2, pch = 20)
plot(reg.summary$bic, xlab = "Number of Variables", ylab = "BIC", type = "l")
which.min(reg.summary$bic)

```

[1] 4

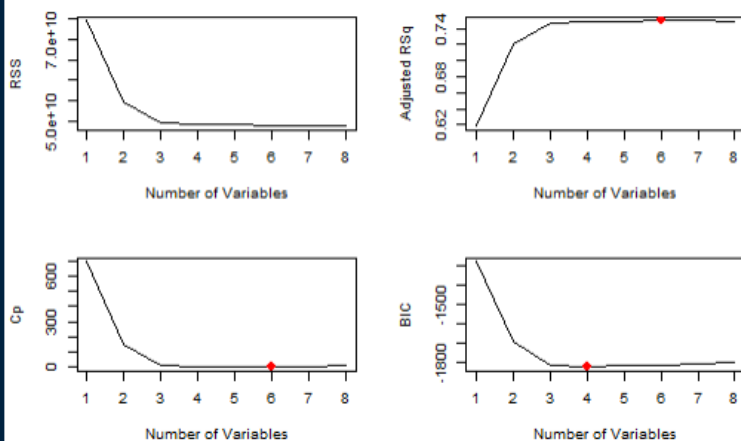
```

{r}
par(mfrow = c(2,2))

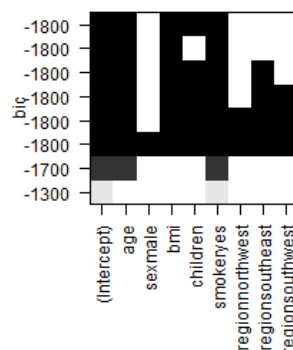
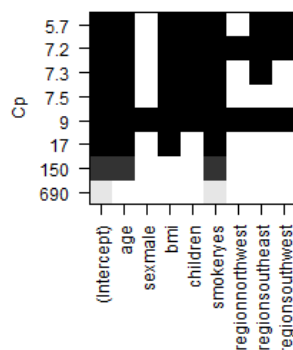
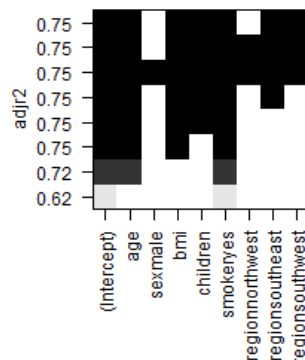
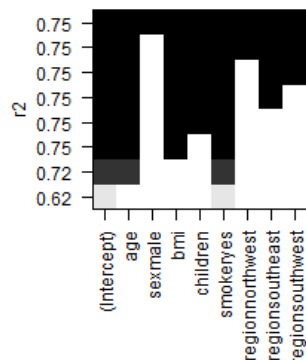
Plot RSS versus Number of Variables
plot(reg.summary$rss, xlab="Number of Variables", ylab="RSS", type="l")
Plot Adjusted R-squared versus Number of Variables
plot(reg.summary$adjr2, xlab="Number of Variables", ylab="Adjusted RSq", type="l")
points(6, reg.summary$adjr2[6], col="red", cex=2, pch=20)
Plot Cp versus Number of Variables
plot(reg.summary$cp, xlab="Number of Variables", ylab="Cp", type="l")
Add a red point at position 6
points(6, reg.summary$cp[6], col="red", cex=2, pch=20)

Plot BIC versus Number of Variables
plot(reg.summary$bic, xlab="Number of Variables", ylab="BIC", type="l")
Add a red point at position 6
points(4, reg.summary$bic[4], col="red", cex=2, pch=20)

```



```
{r}
par(mfrow = c(1,2))
plot(regfit.full,scale="r2")
plot(regfit.full,scale="adjr2")
plot(regfit.full,scale="Cp")
plot(regfit.full,scale="bic")
```



After considering all the available information, I think that the most appropriate model comprises of four variables, which include age, BMI, number of children, and smoking status.



### 3. Forward and Backward Stepwise Selection

```
##{r}
regfit.fwd<- regsubsets(charges~., data = Insurance, method = "forward")
summary(regfit.fwd)
```

Subset selection object  
Call: regsubsets.formula(charges ~ ., data = Insurance, method = "forward")  
8 Variables (and intercept)

		Forced in	Forced out
age		FALSE	FALSE
sexmale		FALSE	FALSE
bmi		FALSE	FALSE
children		FALSE	FALSE
smokeryes		FALSE	FALSE
regionnorthwest		FALSE	FALSE
regionsoutheast		FALSE	FALSE
regionsouthwest		FALSE	FALSE

1 subsets of each size up to 8  
Selection Algorithm: forward

		age	sexmale	bmi	children	smokeryes	regionnorthwest	regionsoutheast	regionsouthwest
1	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
2	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
3	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
4	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
5	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
6	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
7	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
8	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "

```
##{r}
regfit.bwd <- regsubsets(charges~., data = Insurance, method = "backward")
summary(regfit.bwd)
```

Subset selection object  
Call: regsubsets.formula(charges ~ ., data = Insurance, method = "backward")  
8 Variables (and intercept)

		Forced in	Forced out
age		FALSE	FALSE
sexmale		FALSE	FALSE
bmi		FALSE	FALSE
children		FALSE	FALSE
smokeryes		FALSE	FALSE
regionnorthwest		FALSE	FALSE
regionsoutheast		FALSE	FALSE
regionsouthwest		FALSE	FALSE

1 subsets of each size up to 8  
Selection Algorithm: backward

		age	sexmale	bmi	children	smokeryes	regionnorthwest	regionsoutheast	regionsouthwest
1	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
2	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
3	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
4	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
5	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
6	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
7	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "
8	( 1 )	" "	" "	" "	" "	" "	" "	" "	" "

Both forward and backward stepwise selection methods yielded identical variable selection for each model.

## Validation Set Approach

```
set.seed(1)
train <- sample(c(TRUE, FALSE), nrow(Insurance), rep = TRUE)
test <- (!train)
regfit.best <- regsubsets(charges~., data = Insurance[train,],)
test.mat <- model.matrix(charges~., data = Insurance[test,])
val.errors = rep(NA, 8)
for(i in 1:8){
 coefi <- coef(regfit.best, id=i)
 pred <- test.mat[, names(coefi)] %*% coefi
 val.errors[i] = mean((Insurance$charges[test] - pred)^2)
}
val.errors
```

```
[1] 60092451 43450714 38869494 38860514 38853079 38426858 38471229 38489704
```

```
which.min(val.errors)
```

```
[1] 6
```

```
coef(regfit.best, 6)
```

(Intercept)	age	bmi	children	smokeryes	regionsoutheast
regionsouthwest					
-10490.6818	248.8687	305.4671	275.0695	23132.9314	-1183.8361
-980.2071					

The model with six variables is the best one, according to the validation set approach. Since these six variables might not be the same as the ones picked for the training batch, I will use the entire model to identify them.

```
regfit.best <- regsubsets(charges~., data = Insurance)
coef(regfit.best, 6)
```

(Intercept)	age	bmi	children	smokeryes	regionsoutheast
regionsouthwest					
-12165.3824	257.0064	338.6413	471.5441	23843.8749	-858.4696
-782.7452					

In this instance, the training set and the entire dataset both chose the six previously indicated variables.

## Cross Validation

```
##{r}
set.seed(1)
folds <- sample(1:k, nrow(Insurance), replace = TRUE)
cv.errors <- matrix(NA, k, 8, dimnames = list(NULL, paste(1:8)))

for(j in 1:k) {
 # Convert categorical variables to numeric format
 x_train <- model.matrix(charges ~ ., data = Insurance[folds != j,])

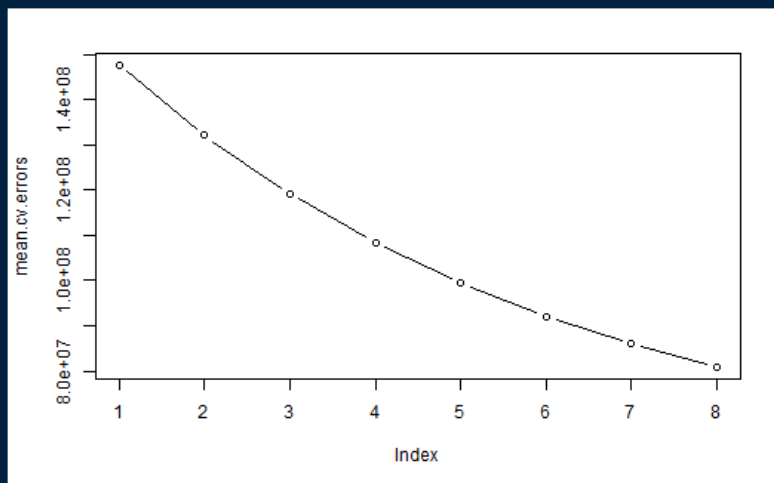
 # Fit cross-validated elastic net model
 cv.fit <- cv.glmnet(x = x_train, y = Insurance$charges[folds != j], alpha = 1)

 for(i in 1:8) {
 # Predict on test fold and calculate CV error
 x_test <- model.matrix(charges ~ ., data = Insurance[folds == j,])
 pred <- predict(cv.fit, newx = x_test, s = cv.fit$lambda[i], type = "response")
 cv.errors[j, i] = mean((Insurance$charges[folds == j] - pred)^2)
 }
}

mean.cv.errors <- apply(cv.errors, 2, mean)
mean.cv.errors
```

1	2	3	4	5	6	7	8
147632577	132079572	119164511	108439756	99533649	92137614	85995446	80894426

```
##{r}
par(mfrow=c(1,1))
plot(mean.cv.errors ,type="b")
```



```
##{r}
reg.best <- regsubsets(charges~., data = Insurance)
coef(reg.best,4)
```

(Intercept)	age	bmi	children	smokeryes
-12102.7694	257.8495	321.8514	473.5023	23811.3998

The plot indicates that the cross-validation method resulted in a four-variable model. To obtain these four variables, we can apply the best subset selection technique to the full model.

#### 4. Ridge Regression

```
##{r}
set.seed(1)
x <- model.matrix(charges~., Insurance)[-1]
y <- Insurance$charges
grid <- 10^seq(10, -2, length=100)
train<- sample(1:nrow(x), nrow(x)/1.3)
test<- (-train)
y.test <- y[test]
cv.out <- cv.glmnet(x[train,], y[train], alpha = 0)
best.lam <- cv.out$lambda.min
glm.mod <- glmnet(x[train,], y[train], alpha = 0, lambda = grid, thresh = 1e-12)
glm.pred <- predict(glm.mod, s=best.lam, newx = x[test,])
mean((glm.pred - y.test)^2)
```

```
[1] 47024809
```

```
##{r}
glm.coef <- predict(glm.mod, type="coefficients", s=best.lam)[1:9,]
glm.coef
```

(Intercept)	age	sexmale	bmi	children	smokeryes
regionnorthwest	regionsoutheast	regionsouthwest			
-10235.7545	250.1924	139.2664	291.0564	432.4727	21633.0029
-307.2834	-654.8357	-656.0909			

Ridge regression produced the optimal model that includes all variables, which is unsurprising.

#### 5. Lasso Regression

```
##{r}
set.seed(1)
cv.out <- cv.glmnet(x[train,], y[train], alpha = 1)
best.lam <- cv.out$lambda.min
lasso.mod <- glmnet(x[train,], y[train], alpha = 1, lambda = grid)
lasso.pred<- predict(lasso.mod, s=best.lam,newx = x[test,])
mean((lasso.pred - y.test)^2)
```

```
[1] 45671404
```

```
##{r}
lasso.coef <- predict(lasso.mod, type="coefficients", s=best.lam)[1:9,]
lasso.coef
```

(Intercept)	age	sexmale	bmi	children	smokeryes
regionnorthwest	regionsoutheast	regionsouthwest			
-11651.568001	266.630591	0.000000	303.235459	394.635541	23245.603786
-1.513378	-510.147078	-371.040819			

Lasso regression resulted in a lower test MSE than ridge regression. Additionally, the lasso model has one less variable (specifically, 'sex male') than the ridge model.

## 6. Least Square Regression

```
{r}
train.df <- data.frame(Insurance[train,])
test.df <- data.frame(Insurance[test,])
lm.fit <- lm(charges~., data=train.df)
lm.pred <- predict(lm.fit, test.df, type = c("response"))
mean((lm.pred - test.df$charges)^2)
```

```
[1] 45478363
```

```
{r}
err.lm <- mean((lm.pred - test.df$charges)^2)
err.ridge <- mean((glm.pred - y.test)^2)
err.lasso <- mean((lasso.pred - y.test)^2)
err.all <- c(err.lm, err.ridge, err.lasso)
barplot(err.all, xlab = "Models", ylab = "Test MSE", names= c("lm", "ridge", "lasso"))
```



When I examine all three data sets, I can see that they all produce a similar test MSE ("mean squared error"). However, the lasso and least square sets produce a slightly lower test MSE, and the least square set produces the smallest value.



R-squared values for all three models are greater than 0.7, however, the ridge model has the lowest value. In general, I am confident in the projections' accuracy.

## 7. Qualitative Analysis Using BMI as a Categorical Variable

In this section, I'll test my ability to predict outcomes using the body mass index (BMI) as a categorical variable. I'll classify someone as "obese" if their BMI is 30 or higher, and I'll give them a BMI score of 1 for being obese and 0 for not being. The data set will then be subjected to four different analyses (LDA, QDA, Logistic Regression, and KNN) to determine how well each method predicts outcomes and to determine the test error.

```
{r}
Insurance$bmi1 <- ifelse(Insurance$bmi > 30, 1, 0)
set.seed(1)
subset <- sample(nrow(Insurance), nrow(Insurance)*0.7)
datatrain <- Insurance[subset,]
datatest <- Insurance[-subset,]
dim(datatest)
```

```
[1] 402 8
```

```
{r}
dim(datatrain)
```

```
[1] 936 8
```

a. Linear Discriminant Analysis

```
##{r}
attach(Insurance)
lda.fit <- lda(bmi1~., data = datatrain)
lda.predict <- predict(lda.fit, datatest)
predictions <- lda.predict$class
actual <- datatest$bmi1
table(predictions, actual)
```

	actual	
predictions	0	1
0	181	13
1	3	205

$(13+3)/(402) = 0.0398$  is the test error.

b. Quadratic Discriminant Analysis

```
##{r}
qda.fit <- qda(bmi1~., data = datatrain)
qda.predict <- predict(qda.fit, datatest)
predictions <- qda.predict$class
table(predictions, actual)
```

	actual	
predictions	0	1
0	176	14
1	8	204

$(14+8)/402 = 0.05547$  is the test error.

c. Logistic Regression

```
##{r}
logistic.fit <- glm(bmi1~., data = datatrain, family = binomial)
logistic.probs <- predict(logistic.fit, datatest, type = "response")
logistic.pred <- rep(0, length(datatest$bmi1))
logistic.pred[logistic.probs>0.5]=1
table(logistic.pred, actual)
```

Warning: glm.fit: algorithm did not converge  
Warning: glm.fit: fitted probabilities numerically 0 or 1 occurred

	actual	
logistic.pred	0	1
0	183	2
1	1	216

$(2+1)/402 = 0.00746$  is the test error.

#### d. K Nearest Neighbor

```
library(class)
train.x <- data.matrix(datatrain)
test.x <- data.matrix(datatest)
train.y <- data.matrix(datatrain$bmi1)
test.y <- data.matrix(datatest$bmi1)
knn.predict <- knn(train.x, test.x, train.y, k=1)
table(knn.predict, test.y)
```

	test.y	
knn.predict	0	1
0	122	79
1	62	139

$(79+62)/402 = 0.35$  is the test error for  $K=1$

```
knn.predict2 <- knn(train.x, test.x, train.y, k=5)
table(knn.predict2, test.y)
```

	test.y	
knn.predict2	0	1
0	110	91
1	74	127

$(91+74)/402 = 0.41$  is the test error for  $K=5$

```
knn.predict3 <- knn(train.x, test.x, train.y, k=10)
table(knn.predict3, test.y)
```

	test.y	
knn.predict3	0	1
0	108	95
1	76	123

$(95+76)/402 = 0.425$  is the test error for  $K=10$

Given that KNN's test error is significantly larger than that of QDA, LDA, and logistic regression, it doesn't appear to be an appropriate method for this dataset.

The minimal test error (0.00746) is provided by logistic regression compared to the other three classification techniques.

## Summary

Implementing various graphical techniques and regression methods facilitated in-depth analysis of the Insurance dataset aimed at uncovering the key factors that impact insurance charges. The primary objective was to assist insurance companies in establishing an appropriate premium price. The analysis revealed that smokers were subject to significantly higher charges than non-smokers. However, an unexpected finding emerged, with non-smoking males incurring higher charges than their female



counterparts. Furthermore, the graphical analysis highlighted that those individuals with four or five children had lower charges than those with fewer children.

Multiple linear regression was employed to develop the most accurate predictive model, resulting in an impressive R-squared value of 0.7535 and an adjusted R-squared of 0.7525. To further optimize the model, I utilized the best subset selection, which identified four critical variables, namely age, BMI, children, and smoking status. The validation procedures, including validation set and cross-validation, revealed divergent models with six and four variables, respectively. Moreover, the forward and backward stepwise selection methods yielded the same models, adding further robustness to the findings.

Additionally, I evaluated the dataset using least square, lasso, and ridge regressions. The analysis demonstrated that least square and lasso regression outperformed ridge regression, exhibiting smaller test mean squared error (MSE) and higher R-squared. Finally, four classification algorithms were applied, namely LDA, QDA, logistic regression, and KNN, utilizing BMI as a categorical variable. My findings indicated that logistic regression had the smallest test error (0.00746), with LDA and QDA closely following suit. However, KNN exhibited relatively high test errors, indicating its unsuitability for my predictive model.