# Chapter 7

- 7.1 Measures of predictive accuracy
- 7.2 Information criteria and cross-validation
  - Instead of 7.2, read:
     Vehtari, A., Gelman, A., Gabry, J. (2017). Practical
     Bayesian model evaluation using leave-one-out cross-validation and WAIC. Statistics and Computing. 27(5):1413–1432. arXiv preprint.
- 7.3 Model comparison based on predictive performance
- 7.4 Model comparison using Bayes factors
- 7.5 Continuous model expansion / sensitivity analysis
- 7.5 Example (may be skipped)

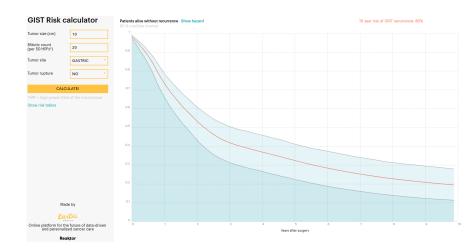
# Model assessment, selection and inference after selection

- Extra material at https://avehtari.github.io/modelselection/
  - Videos, Slides, Notebooks, References
- CV-FAQ https://avehtari.github.io/modelselection/CV-FAQ.html

# Predicting concrete quality



# Predicting cancer recurrence



- ▶ True predictive performance is found out by using it to make predictions and comparing predictions to true observations
  - external validation

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  - external validation
- Expected predictive performance
  - approximates the external validation

- We need to choose the utility/cost function
- Application specific utility/cost functions are important
  - eg. money, life years, quality adjusted life years, etc.

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- Application specific utility/cost functions are important
  - eg. money, life years, quality adjusted life years, etc.
- If are interested overall in the goodness of the predictive distribution, or we don't know (yet) the application specific utility, then good information theoretically justified choice is log-score

$$\log p(y^{\text{rep}} \mid y, M),$$

- What is cross-validation
  - Leave-one-out cross-validation (elpd\_loo, p\_loo)
  - Uncertainty in LOO (SE)
- When is cross-validation applicable?
  - data generating mechanisms and prediction tasks
  - leave-many-out cross-validation
- Fast cross-validation
  - PSIS and diagnostics in loo package (Pareto k, n\_eff, Monte Carlo SE)
  - K-fold cross-validation
- Related methods (WAIC, \*IC, BF)
- Model comparison and selection (elpd\_diff, se)
- Model averaging with Bayesian stacking

# Stan and loo package

Computed from 4000 by 20 log-likelihood matrix

```
Estimate SE
elpd loo -29.5 \ 3.3
p_loo 2.7 1.0
```

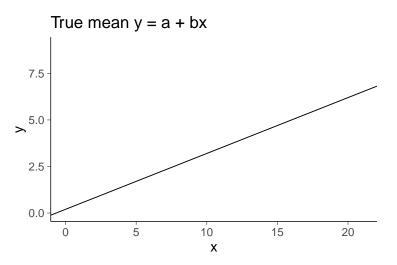
Monte Carlo SE of elpd loo is 0.1.

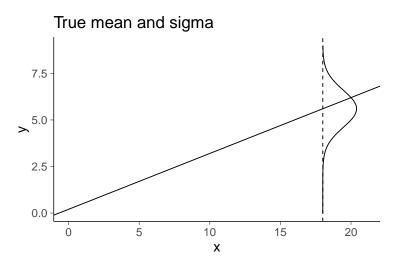
Pareto k	diagi	nostic	value	es:			
				Count	Pct.	Min.	n_eff
(-Inf, 0.	5]	(good)		18	90.0%	899	
(0.5, 0.	7]	(ok)		2	10.0%	459	
(0.7,	1]	(bad)		0	0.0%	<NA $>$	
(1, In	ıf)	(very	bad)	0	0.0%	<na></na>	

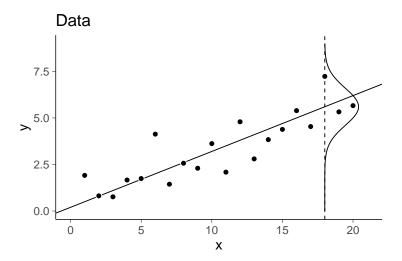
All Pareto k estimates are ok (k < 0.7). See help('pareto-k-diagnostic') for details.

Model comparison: (negative 'elpd diff' favors 1st model, positive favors 2nd)

$$\begin{array}{ccc} \text{elpd\_diff} & \text{se} \\ -0.2 & 0.1 \end{array}$$

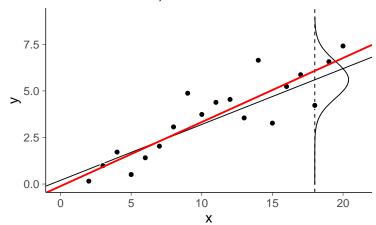




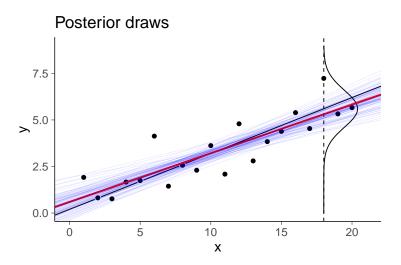


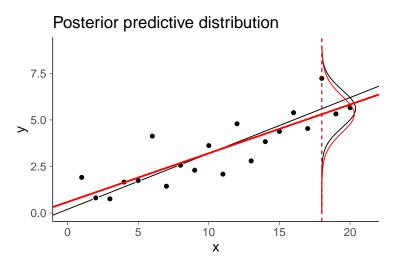


## Posterior mean, alternative data realisation

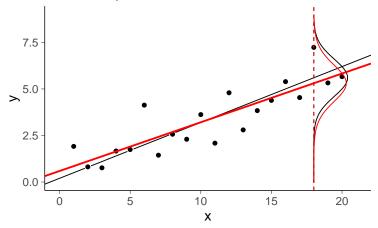




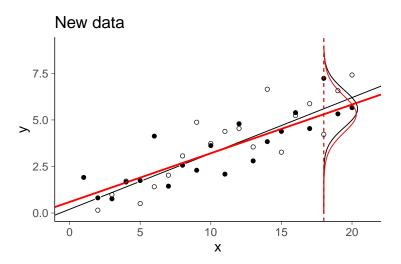


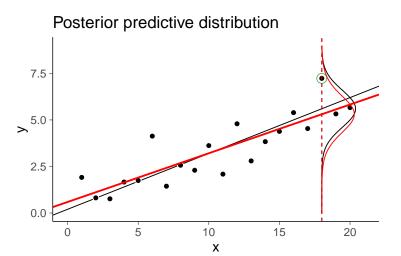


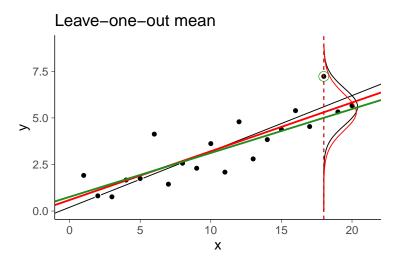
## Posterior predictive distribution

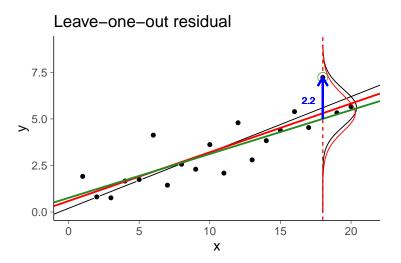


$$p(\tilde{y} \mid \tilde{x} = 18, x, y) = \int p(\tilde{y} \mid \tilde{x} = 18, \theta) p(\theta \mid x, y) d\theta$$

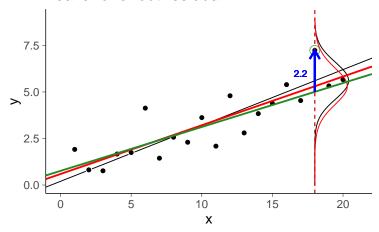






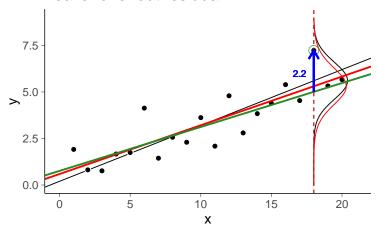


#### Leave-one-out residual



$$y_{18} - E[p(\tilde{y} \mid \tilde{x} = 18, x_{-18}, y_{-18})]$$

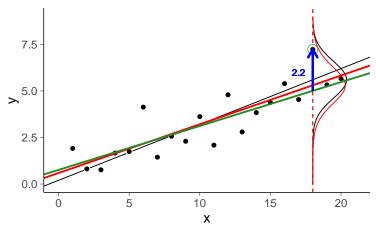
#### Leave-one-out residual



$$y_{18} - E[p(\tilde{y} \mid \tilde{x} = 18, x_{-18}, y_{-18})]$$

Can be use to compute, e.g., RMSE, R<sup>2</sup>, 90% error

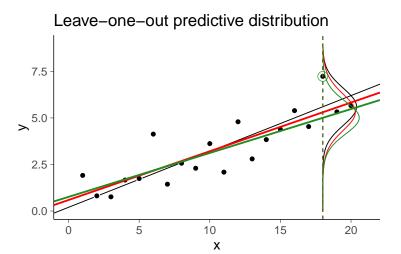
#### Leave-one-out residual



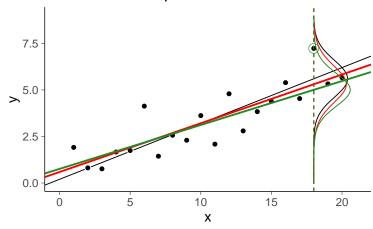
$$y_{18} - E[p(\tilde{y} \mid \tilde{x} = 18, x_{-18}, y_{-18})]$$

Can be use to compute, e.g., RMSE, R<sup>2</sup>, 90% error

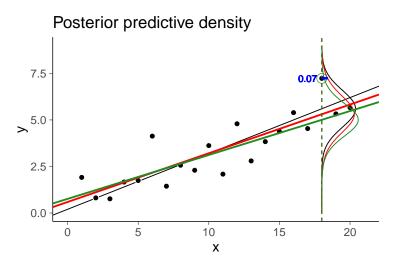
See LOO-R<sup>2</sup> at avehtari.github.io/bayes\_R2/bayes\_R2.html



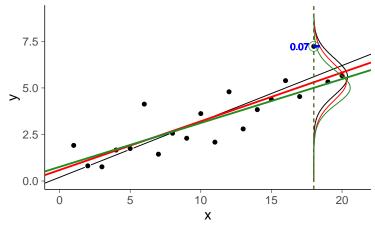
## Leave-one-out predictive distribution



$$p(\tilde{y} \mid \tilde{x} = 18, x_{-18}, y_{-18}) = \int p(\tilde{y} \mid \tilde{x} = 18, \theta) p(\theta \mid x_{-18}, y_{-18}) d\theta$$

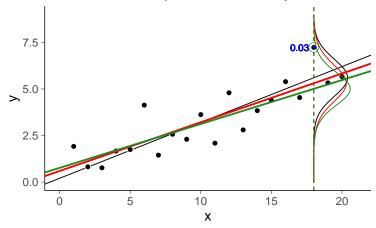


# Posterior predictive density

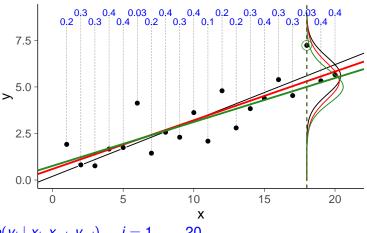


$$p(\tilde{y}=y_{18}\mid \tilde{x}=18,x,y)\approx 0.07$$

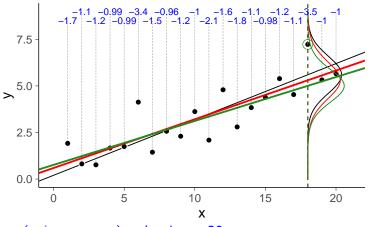
## Leave-one-out predictive density



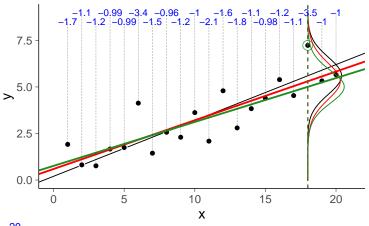
$$p(\tilde{y} = y_{18} \mid \tilde{x} = 18, x, y) \approx 0.07$$
  
 $p(\tilde{y} = y_{18} \mid \tilde{x} = 18, x_{-18}, y_{-18}) \approx 0.03$ 



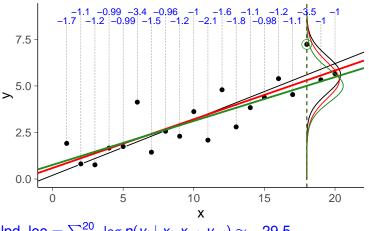
$$p(y_i \mid x_i, x_{-i}, y_{-i}), \quad i = 1, \dots, 20$$



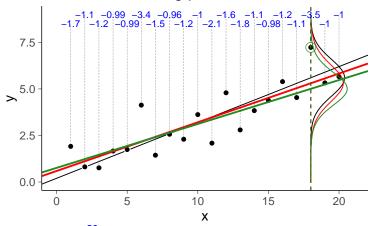
 $\log p(y_i \mid x_i, x_{-i}, y_{-i}), \quad i = 1, \dots, 20$ 



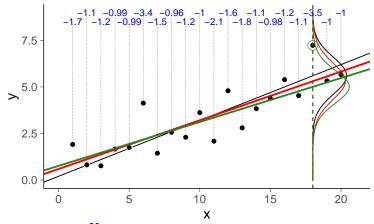
 $\sum_{i=1}^{20} \log p(y_i \mid x_i, x_{-i}, y_{-i}) \approx -29.5$ 



elpd\_loo =  $\sum_{i=1}^{20} \log p(y_i \mid x_i, x_{-i}, y_{-i}) \approx -29.5$ 

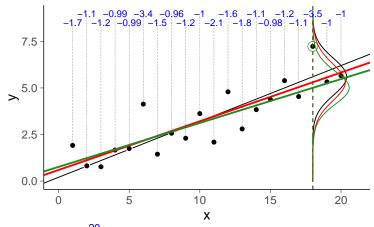


elpd\_loo =  $\sum_{i=1}^{20} \log p(y_i \mid x_i, x_{-i}, y_{-i}) \approx -29.5$ an estimate of log posterior pred. density for new data

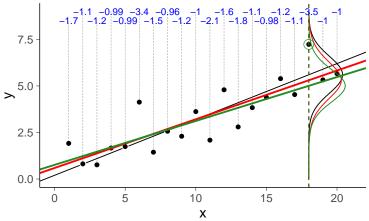


elpd\_loo = 
$$\sum_{i=1}^{20} \log p(y_i \mid x_i, x_{-i}, y_{-i}) \approx -29.5$$

$$lpd = \sum_{i=1}^{20} log p(y_i | x_i, x, y) \approx -26.8$$

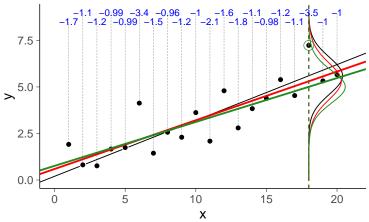


elpd\_loo = 
$$\sum_{i=1}^{20} \log p(y_i \mid x_i, x_{-i}, y_{-i}) \approx -29.5$$
  
lpd =  $\sum_{i=1}^{20} \log p(y_i \mid x_i, x, y) \approx -26.8$   
p loo = lpd - elpd loo  $\approx 2.7$ 

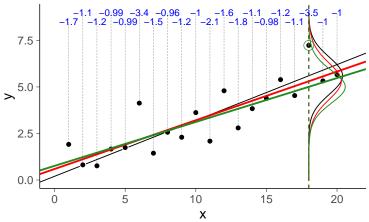


elpd\_loo = 
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asymptotically approaches p in case of regular faithful model

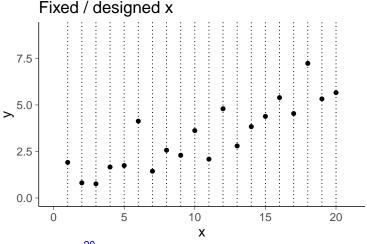


elpd\_loo = 
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p\_loo = lpd - elpd\_loo  $\approx 2.7$   
asymptotically approaches  $p$  in case of regular faithful model



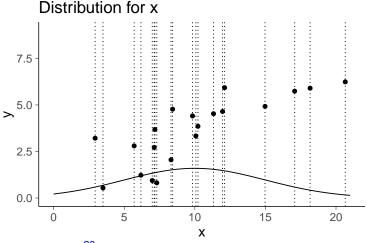
elpd\_loo = 
$$\sum_{i=1}^{20} \log p(y_i \mid x_i, x_{-i}, y_{-i}) \approx -29.5$$
  
SE = sd(log  $p(y_i \mid x_i, x_{-i}, y_{-i})) \cdot \sqrt{20} \approx 3.3$ 

see Vehtari, Gelman & Gabry (2017a) and Vehtari & Ojanen (2012) for more



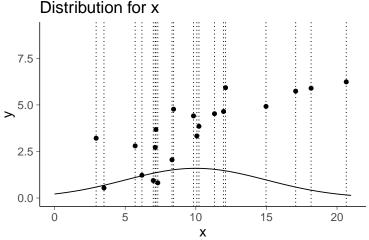
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LOO is ok for fixed / designed x. SE is uncertainty about  $y \mid x$ .



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LOO is ok for random x. SE is uncertainty about  $y \mid x$  and x.



elpd\_loo = 
$$\sum_{i=1}^{20} \log p(y_i \mid x_i, x_{-i}, y_{-i}) \approx -29.5$$
  
SE = sd(log  $p(y_i \mid x_i, x_{-i}, y_{-i})) \cdot \sqrt{20} \approx 3.3$ 

LOO is ok for random x. SE is uncertainty about  $y \mid x$  and x. Covariate shift can be handled with importance weighting or modelling

#### 100 package

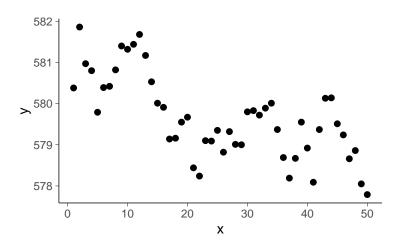
Computed from 4000 by 20 log-likelihood matrix

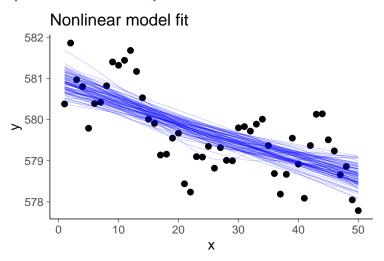
Monte Carlo SE of elpd\_loo is 0.1.

Pareto k diagnostic values:

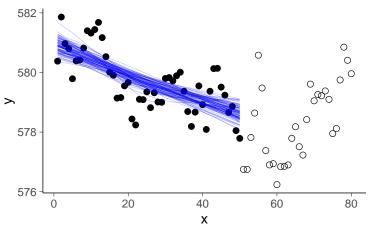
				Count	Pct.	Min.	n_eff
(-Inf, C	).5]	(good)		18	90.0%	899	
(0.5, 0)	1.7]	(ok)		2	10.0%	459	
(0.7,	1]	(bad)		0	0.0%	<na></na>	
(1, 1	nf)	(very	bad)	0	0.0%	<na></na>	

All Pareto k estimates are ok (k < 0.7). See help('pareto-k-diagnostic') for details.

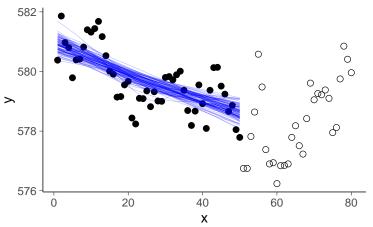




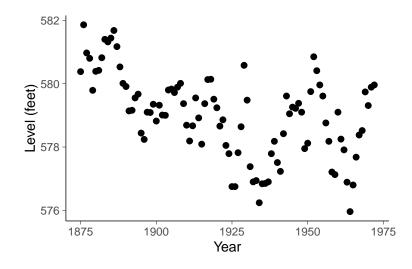




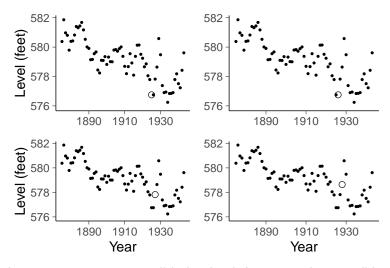




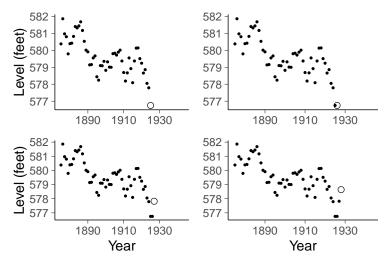
Extrapolation is more difficult



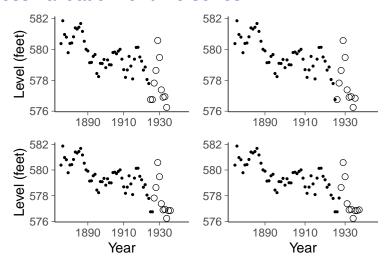
Can LOO or other cross-validation be used with time series?



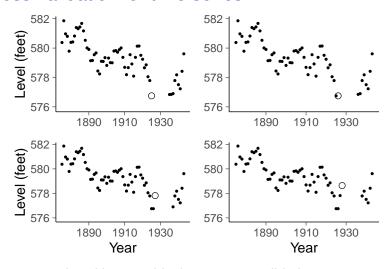
Leave-one-out cross-validation is ok for assessing conditional model



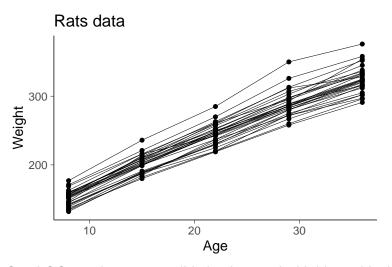
Leave-future-out cross-validation is better for predicting future



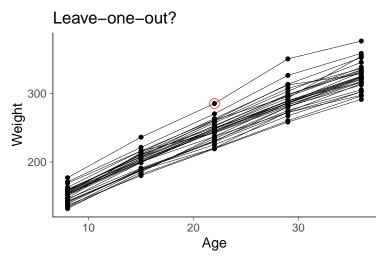
*m*-step-ahead cross-validation is better for predicting further future

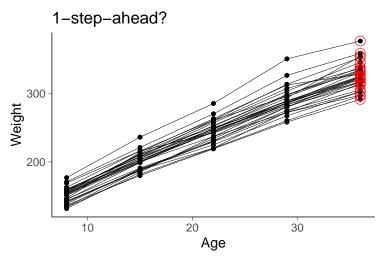


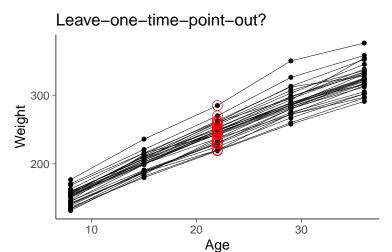
m-step-ahead leave-a-block-out cross-validation

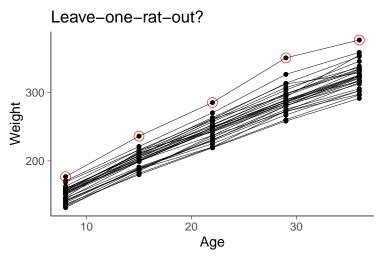


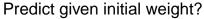
Can LOO or other cross-validation be used with hierarchical data?

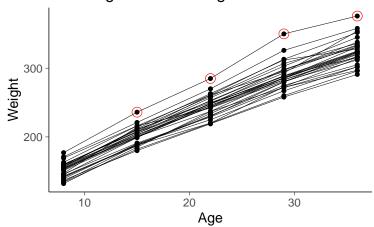












# Summary of data generating mechanisms and prediction tasks

- You have to make some assumptions on data generating mechanism
- Use the knowledge of the prediction task if available
- Cross-validation can be used to analyse different parts, even if there is no clear prediction task

#### Fast cross-validation

- Pareto smoothed importance sampling LOO (PSIS-LOO)
- K-fold cross-validation

• We want to compute

$$p(y_i \mid x_{-i}, y_{-i}) = \int p(y_i \mid x_i, \theta) p(\theta \mid x_{-i}, y_{-i}) d\theta$$

- We want to compute  $p(y_i \mid x_{-i}, y_{-i}) = \int p(y_i \mid x_i, \theta) p(\theta \mid x_{-i}, y_{-i}) d\theta$
- Proposal distribution is full posterior  $\theta^{(s)} \sim p(\theta \mid x, y)$
- Target distribution is LOO-posterior  $p(\theta \mid x_{-i}, y_{-i})$

• We want to compute  $p(y_i \mid x_{-i}, y_{-i}) = \int p(y_i \mid x_i, \theta) p(\theta \mid x_{-i}, y_{-i}) d\theta$ 

• Proposal distribution is full posterior 
$$\theta^{(s)} \sim p(\theta \mid x, y)$$

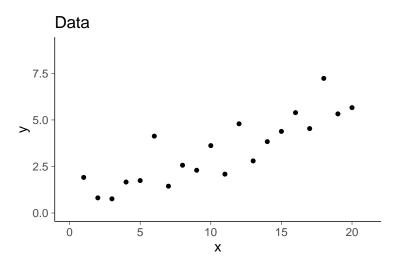
- Target distribution is LOO-posterior  $p(\theta \mid x_{-i}, y_{-i})$
- Importance ratio

$$w_i^{(s)} = \frac{p(\theta^{(s)} \mid x_{-i}, y_{-i})}{p(\theta^{(s)} \mid x, y)} \propto \frac{1}{p(y_i \mid \theta^{(s)})}$$

• We want to compute  $p(y_i \mid x_{-i}, y_{-i}) = \int p(y_i \mid x_i, \theta) p(\theta \mid x_{-i}, y_{-i}) d\theta$ 

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- Importance ratio

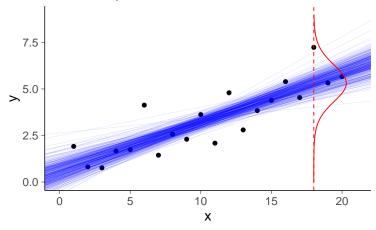
$$w_{i}^{(s)} = \frac{p(\theta^{(s)} \mid x_{-i}, y_{-i})}{p(\theta^{(s)} \mid x, y)} \propto \frac{1}{p(y_{i} \mid \theta^{(s)})}$$
$$\tilde{w}_{i}^{(s)} = \frac{w_{i}^{(s)}}{\sum_{s'=1}^{S} w_{i}^{(s')}}$$



## Posterior draws 7.5 > 5.0 2.5 0.0 10 15 20 5 Χ

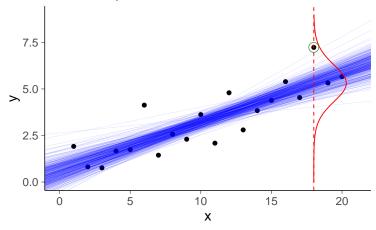
$$\theta^{(s)} \sim p(\theta \mid x, y)$$

#### Posterior predictive distribution



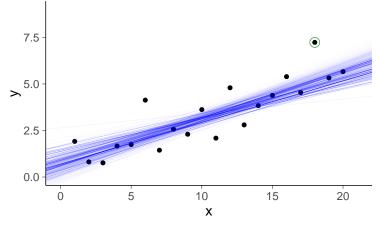
$$\theta^{(s)} \sim p(\theta \mid x, y), \quad p(\tilde{y} \mid \tilde{x}, x, y) \approx \frac{1}{S} \sum_{s=1}^{S} p(\tilde{y} \mid \tilde{x}, \theta^{(s)})$$

#### Posterior predictive distribution



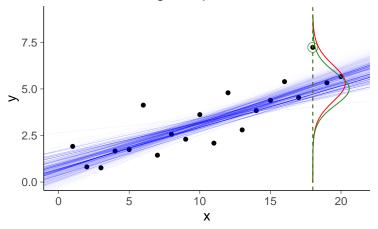
$$\theta^{(s)} \sim p(\theta \mid x, y), \quad p(\tilde{y} \mid \tilde{x}, x, y) \approx \frac{1}{S} \sum_{s=1}^{S} p(\tilde{y} \mid \tilde{x}, \theta^{(s)})$$

#### PSIS-LOO weighted draws



$$\theta^{(s)} \sim p(\theta \mid x, y), \quad w_i^{(s)} = p(\theta^{(s)} \mid x_{-i}, y_{-i}) / p(\theta^{(s)} \mid x, y)$$

#### PSIS-LOO weighted predictive distribution



$$\theta^{(s)} \sim p(\theta \mid x, y), \quad w_i^{(s)} = p(\theta^{(s)} \mid x_{-i}, y_{-i}) / p(\theta^{(s)} \mid x, y)$$
$$p(y_i \mid x_i, x_{-i}, y_{-i}) \approx \sum_{s=1}^{S} [\tilde{w}_i^{(s)} p(y_i \mid x_i, \theta^{(s)})]$$

### Pareto smoothed importance sampling LOO

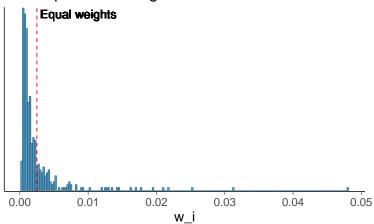
We want to compute

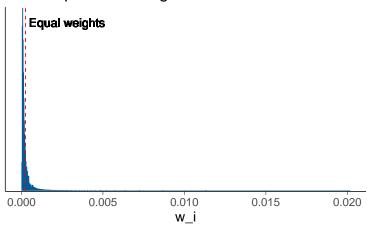
$$p(y_i \mid x_{-i}, y_{-i}) = \int p(y_i \mid x_i, \theta) p(\theta \mid x_{-i}, y_{-i}) d\theta$$

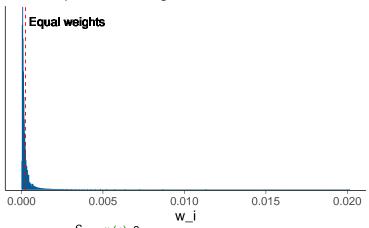
- Proposal distribution is full posterior  $\theta^{(s)} \sim p(\theta \mid x, y)$
- Target distribution is LOO-posterior  $p(\theta \mid x_{-i}, y_{-i})$
- Importance ratio

$$w_i^{(s)} = \frac{p(\theta^{(s)} \mid x_{-i}, y_{-i})}{p(\theta^{(s)} \mid x, y)} \propto \frac{1}{p(y_i \mid \theta^{(s)})}$$
$$\tilde{w}_i^{(s)} = \frac{w_i^{(s)}}{\sum_{s'=1}^S w_i^{(s')}}$$

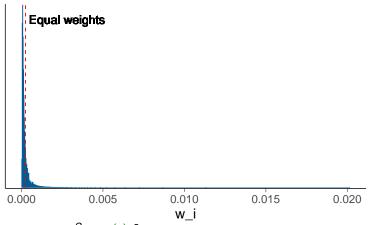
- The variability of importance weights matter
  - Pareto-k diagnostic
  - Pareto smoothed importance sampling LOO (PSIS-LOO)







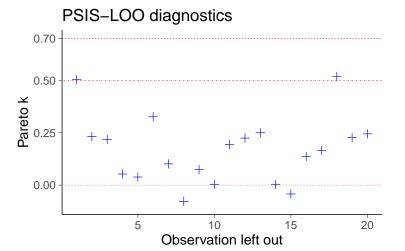
ESS 
$$\approx 1/\sum_{s=1}^{S} (\tilde{w}^{(s)})^2 \approx 459$$



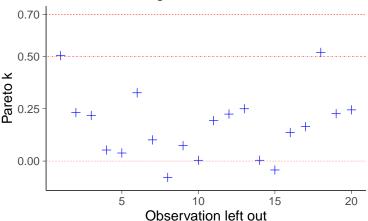
ESS 
$$\approx 1/\sum_{s=1}^{S} (\tilde{\textit{w}}^{(s)})^2 \approx 459$$

Pareto  $\hat{k} \approx 0.52$ 

- Pareto  $\hat{k}$  estimates the tail shape which determines the convergence rate of PSIS. Less than 0.7 is ok.



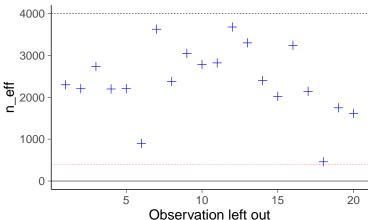
### PSIS-LOO diagnostics



#### Pareto k diagnostic values:

		Count	Pct.	Min .	n_eff
(-Inf, 0.5]	(good)	18	90.0%	899	
(0.5, 0.7]	(ok)	2	10.0%	459	
(0.7, 1]	(bad)	0	0.0%	<NA $>$	
(1, Inf)	(very bad)	0	0.0%	<NA $>$	

### PSIS-LOO diagnostics



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### 100 package

Computed from 4000 by 20 log-likelihood matrix

Monte Carlo SE of elpd\_loo is 0.1.

#### Pareto k diagnostic values:

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All Pareto k estimates are ok (k < 0.7). See help('pareto-k-diagnostic') for details.

see more in Vehtari, Gelman & Gabry (2017b)

# Pareto smoothed importance sampling (PSIS)

- Replace the largest weights with ordered statistics of the fitted Pareto distribution
  - equivalent to using model to filter the noise out of the weights

See more in Vehtari, Simpson, Gelman, Yao & Gabry (2021)

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- Replace the largest weights with ordered statistics of the fitted Pareto distribution
  - equivalent to using model to filter the noise out of the weights
- Reduced variability compared to the plain IS
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- Asymptotically consistent under some mild conditions

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### Stan code

$$\log(w_i^{(s)}) = \log(1/p(y_i \mid x_i, \theta^{(s)})) = -\log_{\text{lik}[i]}$$

#### Stan code

```
\log(w_i^{(s)}) = \log(1/p(y_i \mid x_i, \theta^{(s)})) = -\log_{\text{lik}[i]}
model {
  alpha ~ normal(pmualpha, psalpha);
  beta ~ normal(pmubeta, psbeta);
  y ~ normal(mu, sigma);
generated quantities {
  vector[N] log lik;
  for (i in 1:N)
    log lik[i] = normal lpdf(v[i] | mu[i], sigma);
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RStanARM and brms compute log\_lik by default

### loo()

 RStan (log\_lik in gen. quantities), RStanARM, brms loo(fit)

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- RStan (log\_lik in gen. quantities), RStanARM, brms loo(fit)
- CmdStanR (eventually will work as others)

```
loocmd <- function(fit, ...) {
  loo(fit$draws("log_lik"),
  r_eff=relative_eff(fit$draws("log_lik"))),
  ...
}
loocmd(fit)</pre>
```

# What if many high Pareto- $\hat{k}$ 's

rstan::loo(..., moment\_match = TRUE)
 brms::loo(..., moment\_match = TRUE)
 support moment matching algorithm by Paananen et al.
 (2021, Implicitly adaptive importance sampling), see also
 http://mc-stan.org/loo/articles/loo2-moment-matching.html

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- Use K-fold-CV (more about this soon)

# Pareto smoothed importance sampling CV variants

- PSIS-LOO for hierarchical models
  - leave-one-group out is challenging for PSIS-LOO
  - see Merkel, Furr and Rabe-Hesketh (2018) for an approach using quadrature integration
  - Stan demo of the challenges and integrated LOO at https://avehtari.github.io/modelselection/roaches.html

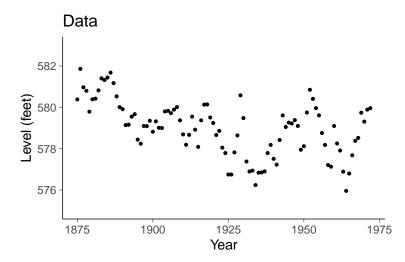
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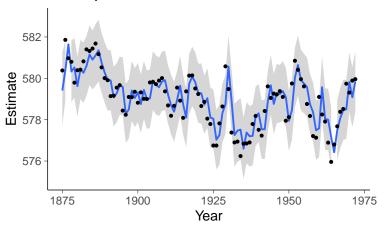
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- PSIS-LOO for non-factorizable models
  - mc-stan.org/loo/articles/loo2-non-factorizable.html
- PSIS-LOO for time series
  - Approximate leave-future-out cross-validation (LFO-CV)
    - mc-stan.org/loo/articles/loo2-lfo.html

# LFO-CV



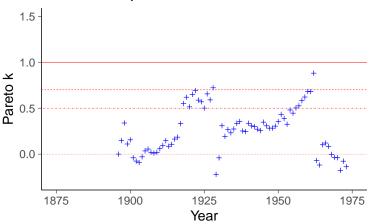
### LFO-CV





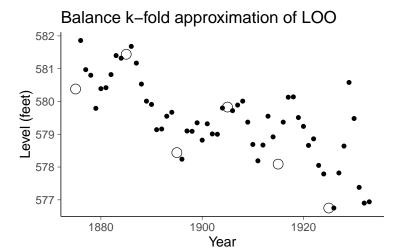
### LFO-CV

### PSIS-1-step-ahead with refits

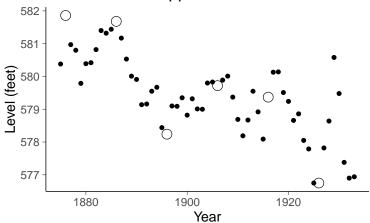


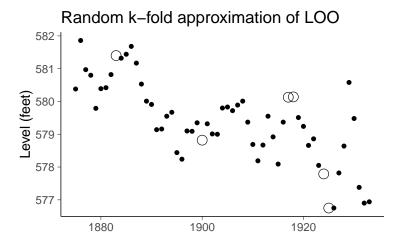
#### K-fold cross-validation

- K-fold cross-validation can approximate LOO
  - the same use cases as with LOO
- K-fold cross-validation can be used for hierarchical models
  - good for leave-one-group-out
- K-fold cross-validation can be used for time series
  - with leave-block-out

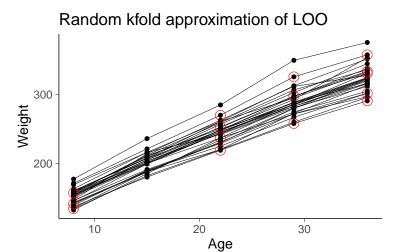


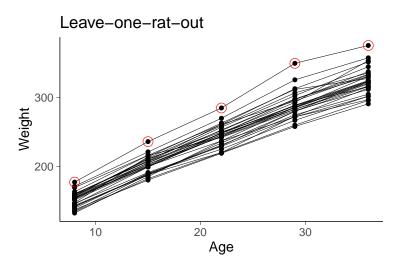
# Balance k-fold approximation of LOO





Year





#### K-fold-CV code

- RStan, CmdStanR
   See vignette http://mc-stan.org/loo/articles/loo2-elpd.html
- RStanARM, brms kfold(fit)
- Alternative data divisions

```
kfold_split_random()
kfold_split_balanced()
kfold_split_stratified()
```

#### WAIC vs PSIS-LOO

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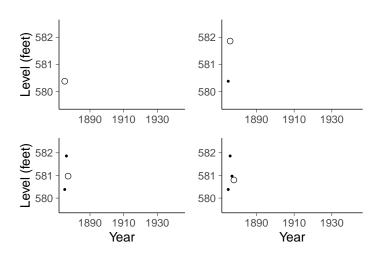
- WAIC has same assumptions as LOO
- PSIS-LOO is more accurate
- PSIS-LOO has much better diagnostics
- LOO makes the prediction assumption more clear, which helps if K-fold-CV is needed instead
- Multiplying by -2 doesn't give any benefit (Watanabe didn't multiply by -2)

#### \*IC

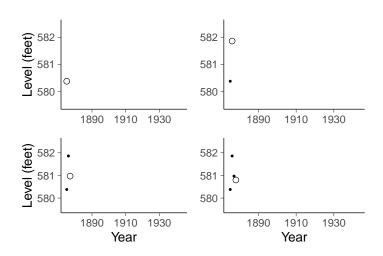
- AIC uses maximum likelihood estimate for prediction
- DIC uses posterior mean for prediction
- BIC is an approximation for marginal likelihood
- TIC, NIC, RIC, PIC, BPIC, QIC, AICc, ...

 Like leave-future-out 1-step-ahead cross-validation but starting with 0 observations

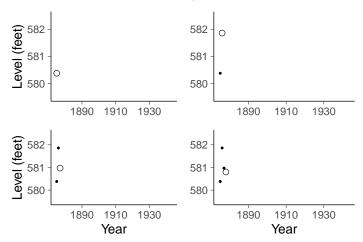
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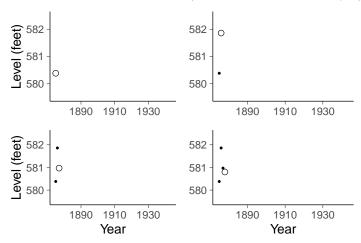
- Like leave-future-out 1-step-ahead cross-validation but starting with 0 observations
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  - unstable in case of misspecified models also asymptotically



#### Cross-validation for model assessment

- CV is good for model assessment when application specific utility/cost functions are used
  - e.g. 90% absolute error

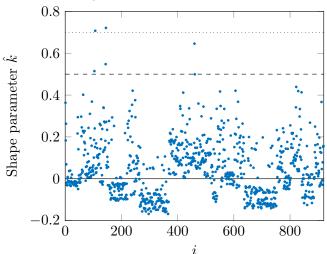
#### Cross-validation for model assessment

- CV is good for model assessment when application specific utility/cost functions are used
  - e.g. 90% absolute error
- Also useful in model checking in similar way as posterior predictive checking (PPC)
  - model misspecification diagnostics (e.g. Pareto-k and p\_loo)
  - checking calibration of leave-one-out predictive posteriors (ppc\_loo\_pit in bayesplot)

see demos avehtari.github.io/modelselection/

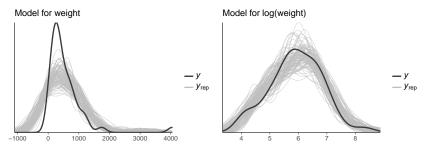
### Radon example

### PSIS-LOO diagnostics



see Vehtari, Gelman & Gabry (2017a)

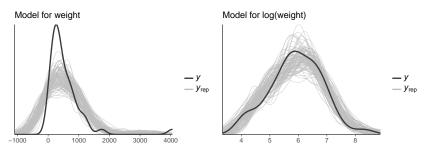
Posterior predictive checking is often sufficient



Predicting the yields of mesquite bushes.

Gelman, Hill & Vehtari (2020): Regression and Other Stories, Chapter 11.

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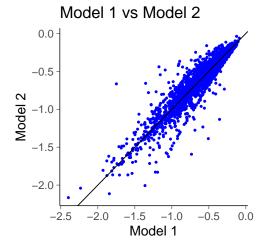
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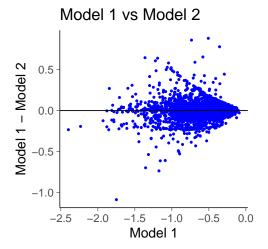
- BDA3, Chapter 6
- Gabry, Simpson, Vehtari, Betancourt, Gelman (2019). Visualization in Bayesian workflow. JRSS A, https://doi.org/10.1111/rssa.12378
- mc-stan.org/bayesplot/articles/graphical-ppcs.html

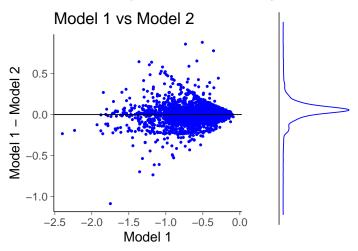
 With good priors that keep the prior on predictive space consistent, there is no need to do model selection to avoid overfitting

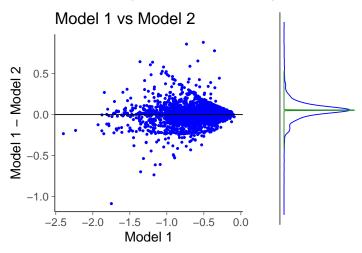
- Logistic regression for predicting probability of switching well with high arsenic level in rural Bangladesh
  - Model 1: log(arsenic) + distance
  - Model 2: log(arsenic) + distance + education level

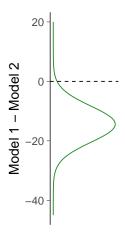


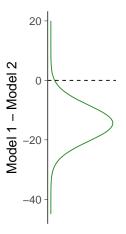
Model 1:  $\widehat{\text{elpd}}_{\text{Loo}}(\mathbf{M}_a \mid y^{\text{obs}}) \approx -1952$ , SE=16 Model 2:  $\widehat{\text{elpd}}_{\text{Loo}}(\mathbf{M}_b \mid y^{\text{obs}}) \approx -1938$ , SE=17











### 8 schools - Model comparison

No difference between pooled and hierarchical for predicting the future observations for a new school (exchangeble with the schools in the data).

### Poisson vs Hurdle-Poisson example

```
\begin{array}{ccc} & & \text{elpd\_diff} & \text{se\_diff} \\ \text{Hurdle-Poisson} & 0.0 & 0.0 \\ \text{Poisson} & -215.9 & 22.1 \end{array}
```

Clear difference (which was also obvious in posterior predictive checks)

1. The models make very similar predictions

2. The models are misspecified with outliers in the data

3. The number of observations is small

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  - if  $|\widehat{\text{elpd}}_{\text{Loo}}(M_a, M_b \mid y^{\text{obs}})| < 4$ , SE is not reliable, but the difference is small anyway
  - selecting a "wrong" model has small cost
  - in nested case the skewness favors the simpler model
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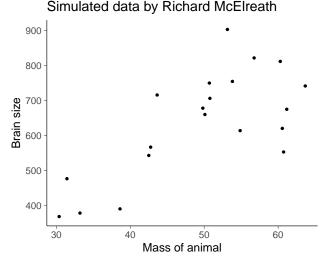
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- 3. The number of observations is small
  - in nested case the skewness favors the simpler model
  - any inference with small *n* is difficult
  - if  $|\widehat{\text{elpd}}_{\text{Loo}}(M_a, M_b \mid y^{\text{obs}})| > 4$ , model is well specified, and n > 100 then the normal approximation is good

- In nested case, often easier and more accurate to analyse posterior distribution of more complex model directly avehtari.github.io/modelselection/betablockers.html
  - instead of comparing Model 1:  $y \sim \text{normal}(\alpha, \sigma)$  vs Model 2:  $y \sim \text{normal}(\alpha + \beta x, \sigma)$  look at the posterior of  $\beta$  directly

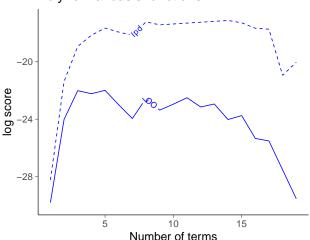
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Polynomial basis functions



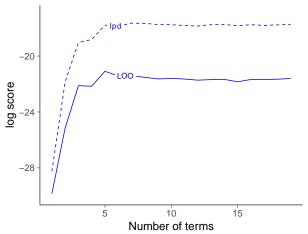
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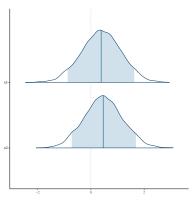
#### Gaussian process basis functions



### Model is not needed to avoid overfitting

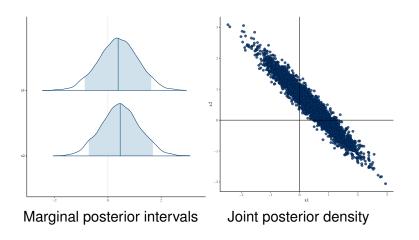
- No overfiiting when using good priors that keep the prior on the predictive space approximately constant when more components are added, e.g.
  - Gaussian procesees
  - (regularized) Horseshoe for sparsity
  - R2-D2 and R2-D2-M2 for prior on R2

# Sometimes predictive model comparison can be useful



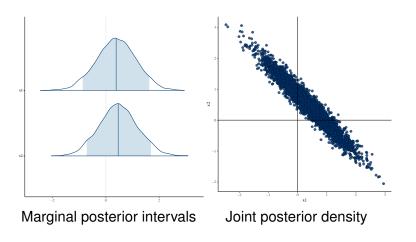
Marginal posterior intervals

# Sometimes predictive model comparison can be useful



rstanarm + bayesplot

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rstanarm + bayesplot

see also Collinear demo

- Continuous expansion including all models?
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# Model averaging

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### Model averaging

- Prefer continuous model expansion
- If needed integrate over the model space = model averaging
- Bayesian stacking may work better than BMA in case of misspecified models or small data
  - Yao, Vehtari, Simpson, & Gelman (2018)

#### Cross-validation and model selection

- · Cross-validation can be used for model selection if
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- Overfitting in selection process is not unique for cross-validation

### Selection induced bias and overfitting

- Selection induced bias in cross-validation
  - same data is used to assess the performance and make the selection
  - the selected model fits more to the data
  - the CV estimate for the selected model is biased
  - recognized already, e.g., by Stone (1974)

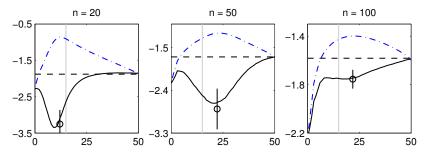
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- Performance of the selection process itself can be assessed using two level cross-validation, but it does not help choosing better models

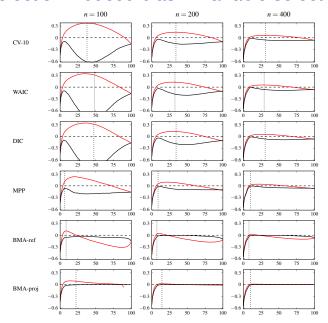
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  - same data is used to assess the performance and make the selection
  - the selected model fits more to the data
  - the CV estimate for the selected model is biased
  - recognized already, e.g., by Stone (1974)
- Performance of the selection process itself can be assessed using two level cross-validation, but it does not help choosing better models
- Bigger problem if there is a large number of models as in covariate selection

### Selection induced bias in variable selection



#### Selection induced bias in variable selection



- It's good to think predictions of observables, because observables are the only ones we can observe
- Cross-validation can simulate predicting and observing new data
- Cross-validation is good if you don't trust your model
- Different variants of cross-validation are useful in different scenarios
- Cross-validation has high variance, and if you trust your model you can beat cross-validation in accuracy

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