

# Chapter 12

- 12.1 Efficient Gibbs samplers (not part of the course)
- 12.2 Efficient Metropolis jump rules (not part of the course)
- 12.3 Further extensions to Gibbs and Metropolis (not part of the course)
- 12.4 Hamiltonian Monte Carlo (important)
- 12.5 Hamiltonian dynamics for a simple hierarchical model (useful example)
- 12.6 Stan: developing a computing environment (useful intro)

## Extra material for HMC / NUTS

- An introduction for applied users with good visualizations: Monnahan, Thorson, and Branch (2016) Faster estimation of Bayesian models in ecology using Hamiltonian Monte Carlo. <https://dx.doi.org/10.1111/2041-210X.12681>
- A review of why HMC works: Neal (2012). MCMC using Hamiltonian dynamics. <https://arxiv.org/abs/1206.1901>
- The No-U-Turn Sampler: Hoffman and Gelman (2014). The No-U-Turn Sampler: Adaptively Setting Path Lengths in Hamiltonian Monte Carlo. <https://jmlr.csail.mit.edu/papers/v15/hoffman14a.html>
- Multinomial variant of NUTS: Betancourt (2018). A Conceptual Introduction to Hamiltonian Monte Carlo. <https://arxiv.org/abs/1701.02434>

## Extra material for Stan

- Gelman, Lee, and Guo (2015) Stan: A probabilistic programming language for Bayesian inference and optimization. [http://www.stat.columbia.edu/~gelman/research/published/stan\\_jebbs\\_2.pdf](http://www.stat.columbia.edu/~gelman/research/published/stan_jebbs_2.pdf)
- Carpenter et al (2017). Stan: A probabilistic programming language. Journal of Statistical Software 76(1). <https://doi.org/10.18637/jss.v076.i01>
- Stan User's Guide, Language Reference Manual, and Language Function Reference (in html and pdf) <https://mc-stan.org/users/documentation/>
  - easiest to start from Example Models in User's guide
- Basics of Bayesian inference and Stan, part 1 Jonah Gabry & Lauren Kennedy (StanCon 2019 Helsinki tutorial)
  - <https://www.youtube.com/watch?v=ZRpo41I02KQ&index=6&list=PLuwyh42iHquU4hUBQs20hkBsKSMrp6H0J>
  - <https://www.youtube.com/watch?v=6cc4N1vT8pk&index=7&list=PLuwyh42iHquU4hUBQs20hkBsKSMrp6H0J>

# Chapter 12 demos

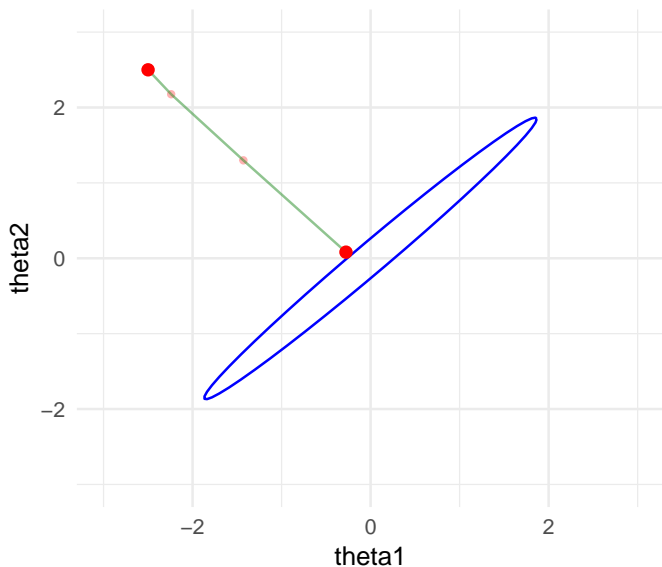
- demo12\_1: HMC
- <https://chi-feng.github.io/mcmc-demo/>
- <http://elevanth.org/blog/2017/11/28/build-a-better-markov-chain/>
- cmdstanr\_demo, rstan\_demo
- <http://sumsar.net/blog/2017/01/bayesian-computation-with-stan-and-farmer-jons/>
- <http://mc-stan.org/documentation/case-studies.html>
- <https://mc-stan.org/cmdstanr/>
- <https://mc-stan.org/rstan/>

# Hamiltonian Monte Carlo

- Uses log density (negative log density is called energy)
- Uses gradient of log density for more efficient sampling

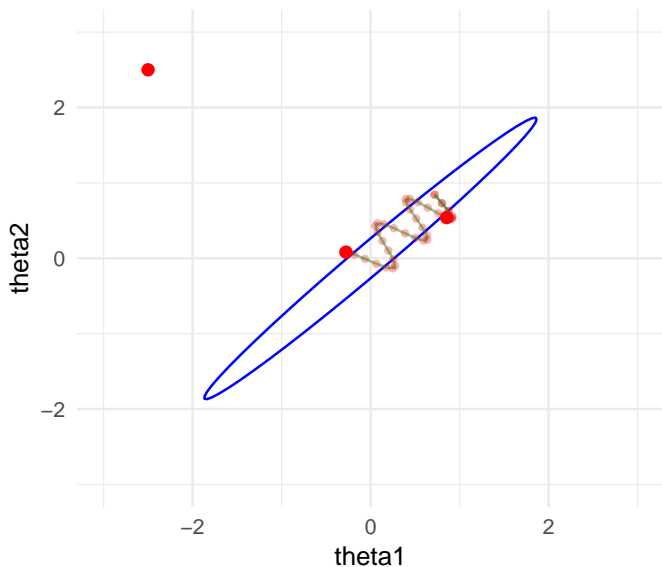
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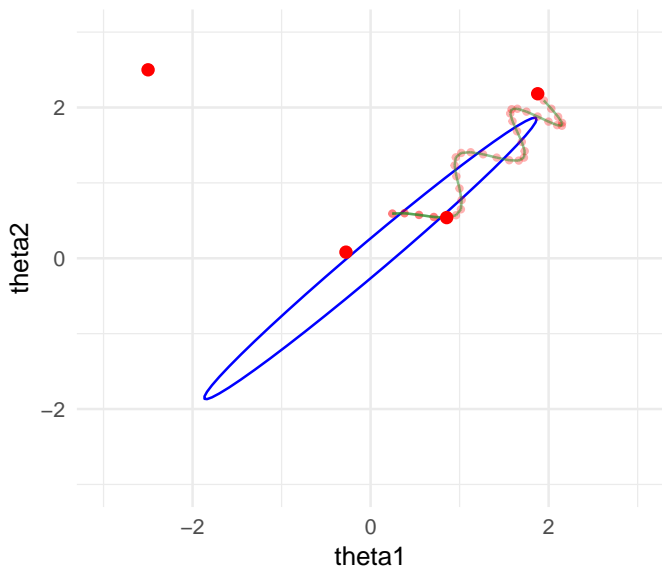
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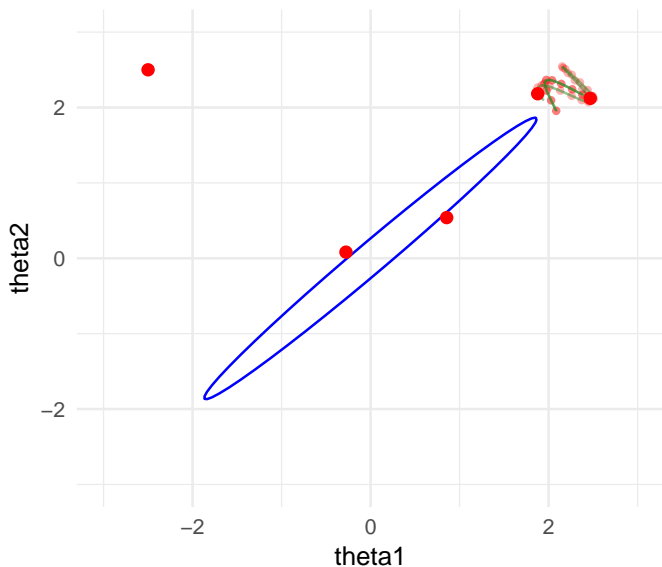
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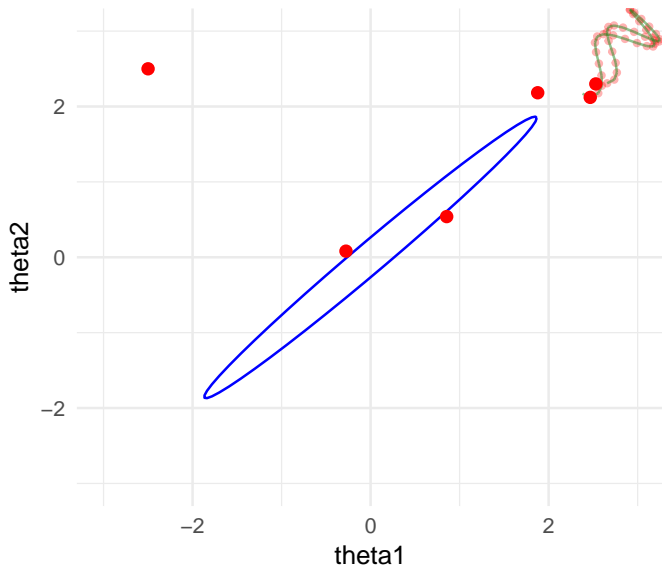
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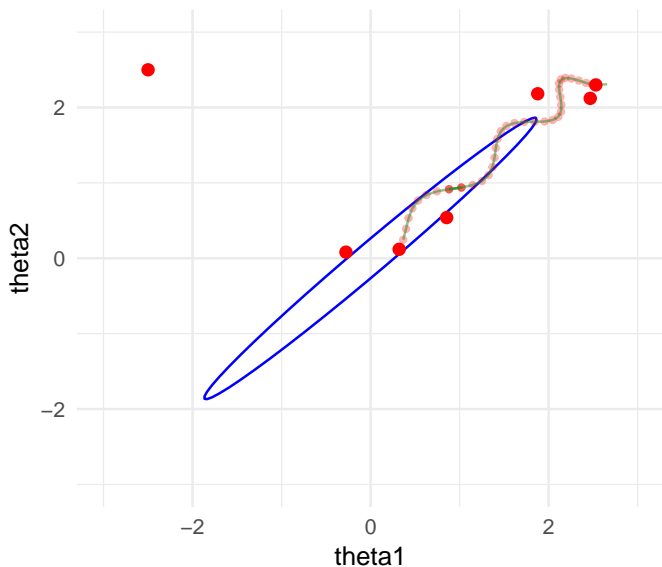
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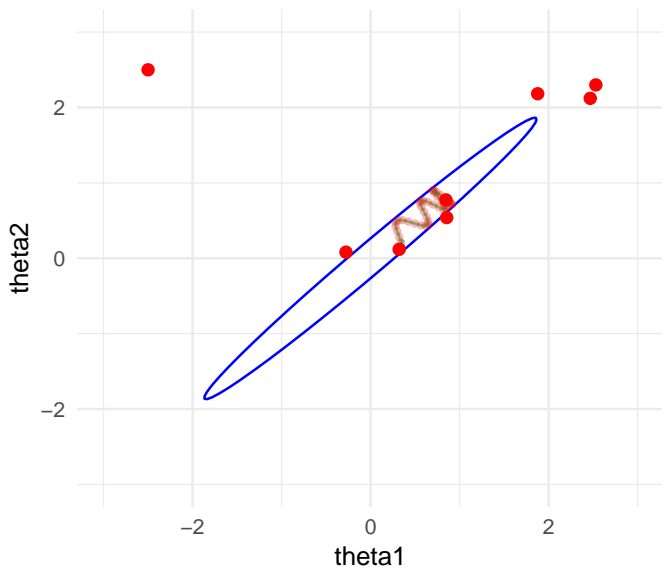
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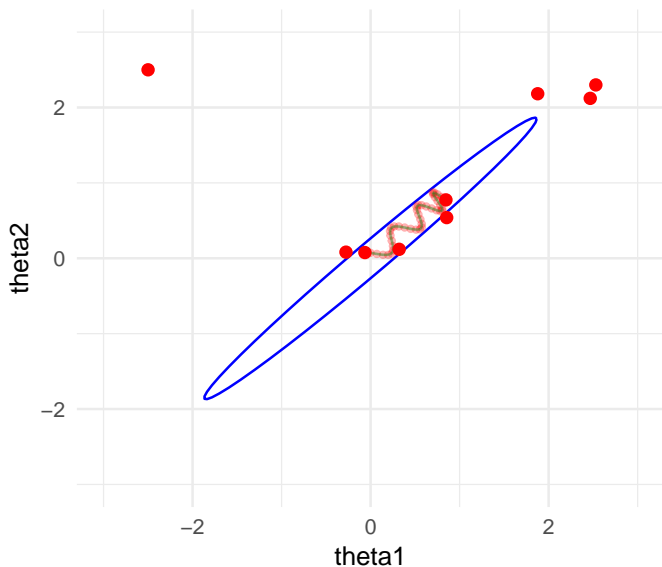
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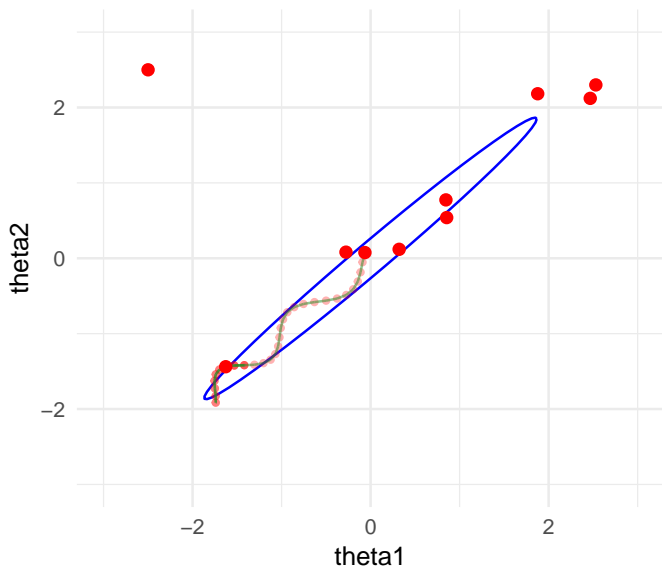
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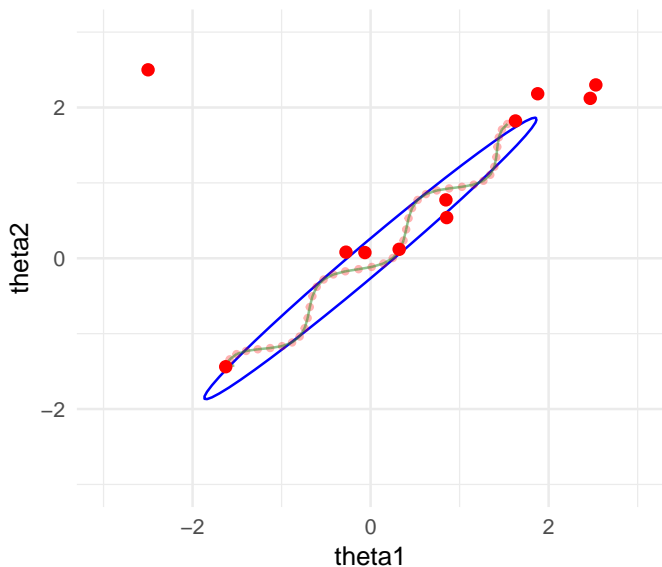
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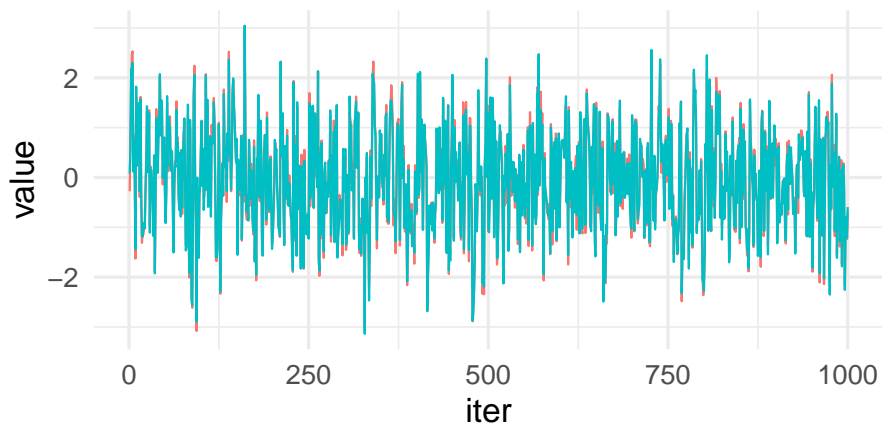
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## Trends

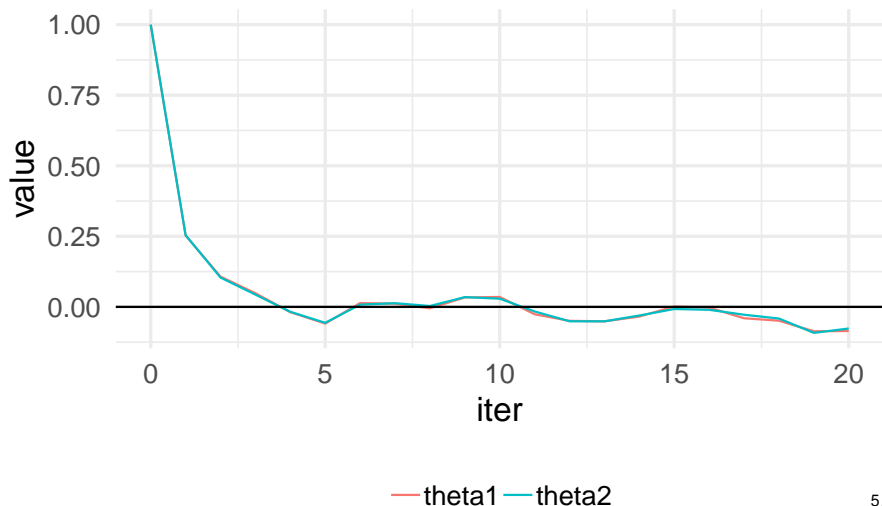




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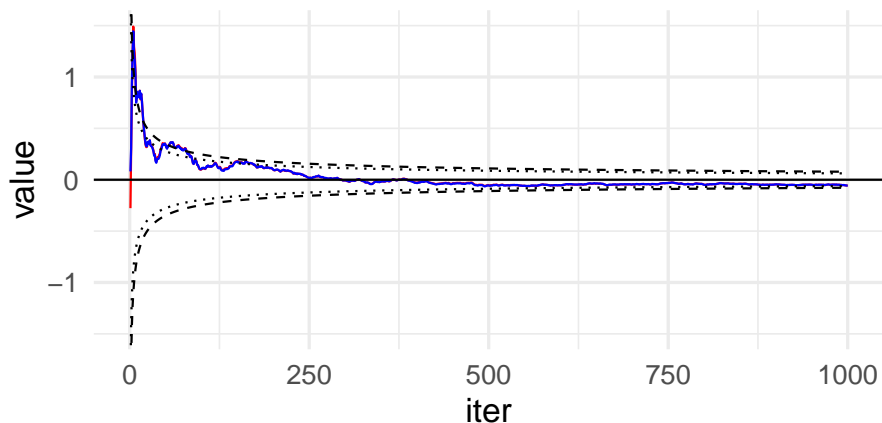
## Autocorrelation function



# Hamiltonian Monte Carlo

- Uses log density (negative log density is called energy)
- Uses gradient of log density for more efficient sampling

## Cumulative averages



- theta1 — theta2 - - 95% interval for MCMC error ..... 95% interval for independent samples

# Hamiltonian Monte Carlo / No-U-Turn sampling

1. HMC basics (static HMC)
2. HMC + leapfrog discretization + Metropolis (static HMC)
3. NUTS + slice sampling + Metropolis (dynamic HMC)
4. NUTS + multinomial (dynamic HMC)

# Hamiltonian Monte Carlo

- Previously
  - Factorizing: sample from
    - 1)  $p(\sigma^2)$ ,
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jump distribution is a combination of proposal distribution and point mass at the previous value
- HMC
  - Augment with  $\phi$  (the same dimensionality as  $\theta$ )
  - - 1) sample directly from  $p(\phi)$ ,
    - 2) make a special joint Metropolis step for  $p(\theta, \phi) = p(\theta)p(\phi)$

# Hamiltonian Monte Carlo

- 1) Sample from  $p(\phi)$ 
  - define  $p(\phi) = \text{normal}(0, 1)$
- 2) Metropolis update for  $p(\theta, \phi) = p(\theta)p(\phi)$ 
  - proposal from Hamiltonian dynamic simulation



# Hamiltonian dynamic simulation

- Statistical mechanics and canonical distribution

$$\begin{aligned} p(\theta, \phi) &= p(\theta)p(\phi) \\ &= \frac{1}{Z} \exp(-(U(\theta) + K(\phi))) \\ &= \frac{1}{Z} \exp(-H(\theta, \phi)) \end{aligned}$$

where

- $U$  is potential energy function
- $K$  is kinetic energy function
- $H$  is Hamiltonian energy function
- $\phi$  is called a momentum variable

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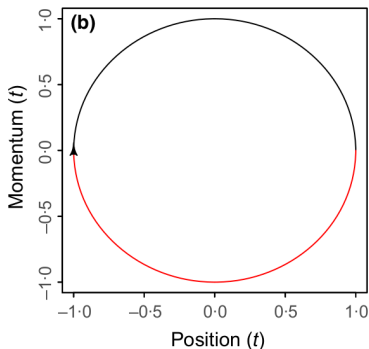
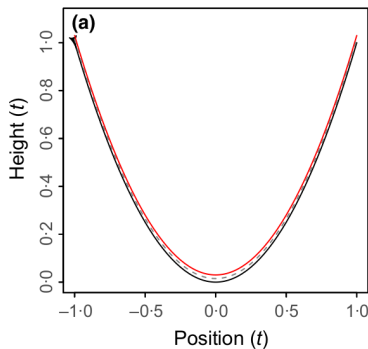
where

- $U$  is potential energy function
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  - $\phi$  is called a momentum variable
- The potential energy is the negative log density  
 $U(\theta) = -\log(p(\theta)) + C$

# Hamiltonian dynamic simulation

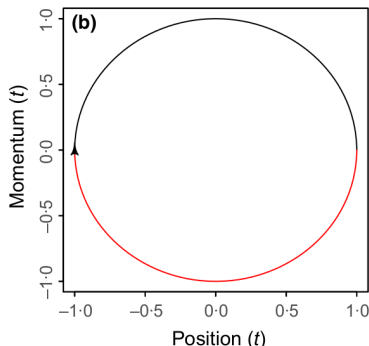
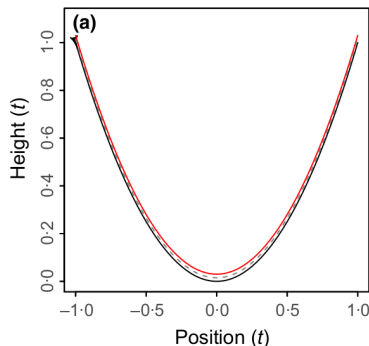
Equations of motion, use also the gradient

$$\frac{d\theta_i}{dt} = \frac{\partial H}{\partial \phi_i}$$
$$\frac{d\phi_i}{dt} = -\frac{\partial H}{\partial \theta_i}$$



# Hamiltonian Monte Carlo

- 1) Sample from  $p(\phi)$ 
  - define  $p(\phi) = \text{normal}(0, 1)$
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  - proposal from Hamiltonian dynamic simulation  
 $p(\theta, \phi) \propto \exp(-H(\theta, \phi))$

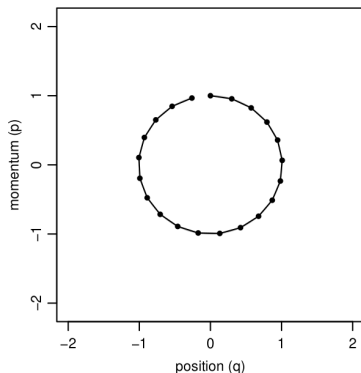


From Monnahan et al (2017)

# Leapfrog discretization

- Leapfrog discretization
  - preserves volume
  - reversible
  - discretization error does not usually grow in time

(c) Leapfrog Method, stepsize 0.3

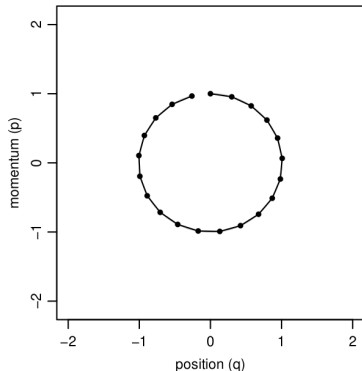


From Neal (2012)

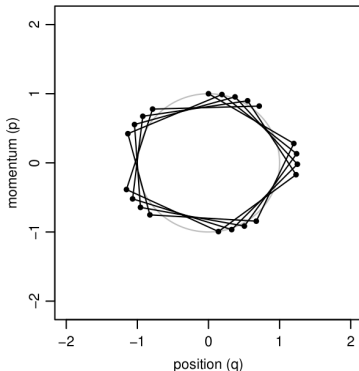
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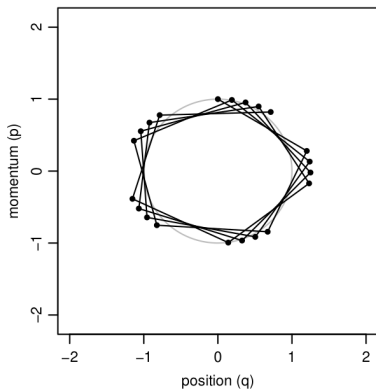
(d) Leapfrog Method, stepsize 1.2



From Neal (2012)

# Leapfrog discretization + Metropolis

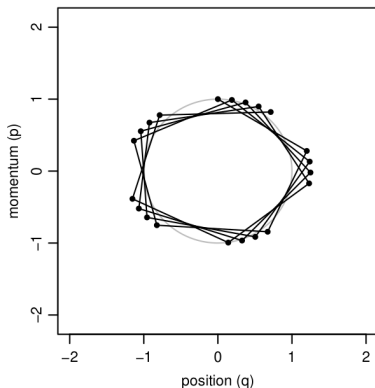
- Leapfrog discretization
  - due to the discretization error the simulation steps away from the constant contour



From Neal (2012)

# Leapfrog discretization + Metropolis

- Leapfrog discretization
  - due to the discretization error the simulation steps away from the constant contour
- Metropolis step with  $r = \exp(H(\theta^*, \phi^*) - H(\theta^{(t-1)}, \phi^{(t-1)}))$ 
  - accept if the Hamiltonian energy in the end is higher
  - accept with some probability if the Hamiltonian energy in the end is lower

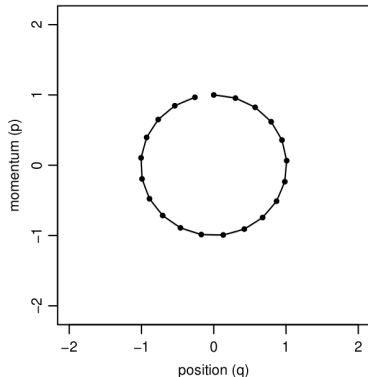




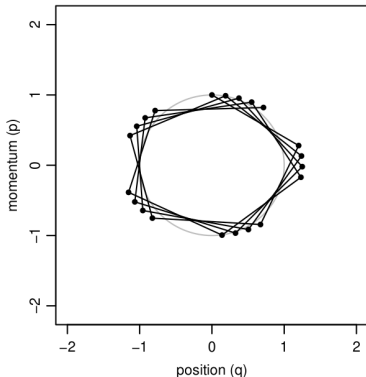
# Two steps of Hamiltonian Monte Carlo

- Perfect simulation keeps  $p(\theta, \phi)$  constant

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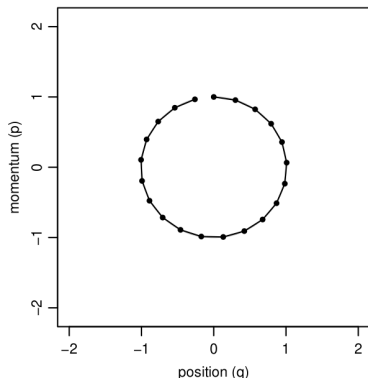


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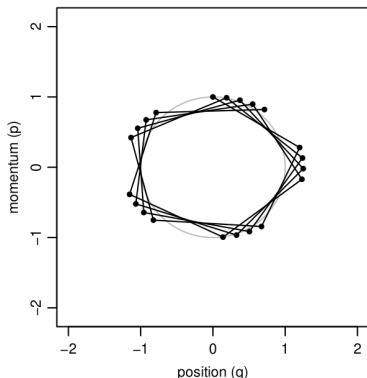
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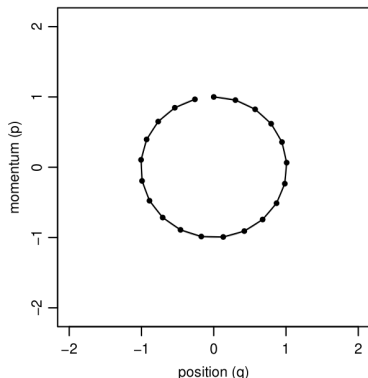


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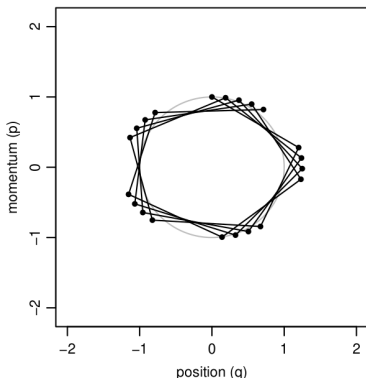
# Two steps of Hamiltonian Monte Carlo

- Perfect simulation keeps  $p(\theta, \phi)$  constant
- Discretized simulation keeps changes in  $p(\theta, \phi)$  small
- Alternating sampling from  $p(\phi)$  is crucial for moving to  $(\theta, \phi)$  points with different joint density

(c) Leapfrog Method, stepsize 0.3



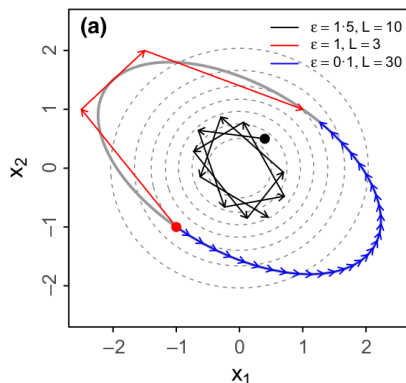
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# Leapfrog discretization, step size

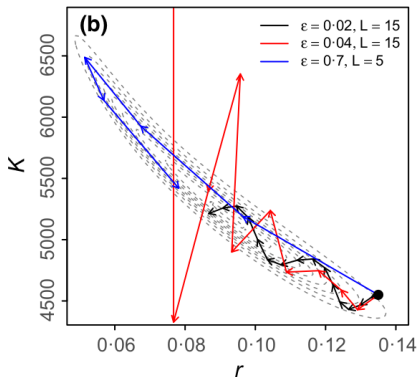
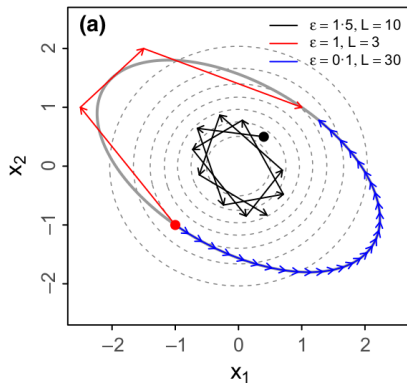
- Small step size  $\rightarrow$  high acceptance rate, but many log density and gradient evaluations
- Big step size  $\rightarrow$  less log density and gradient evaluations, but lower acceptance rate



From Monnahan et al (2017)

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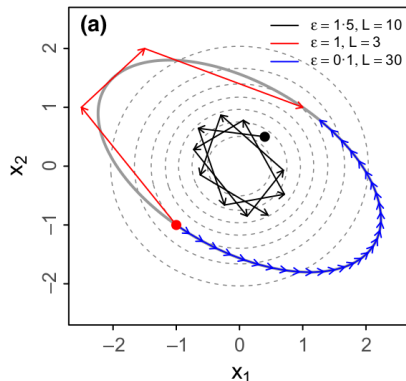
- Small step size  $\rightarrow$  high acceptance rate, but many log density and gradient evaluations
- Big step size  $\rightarrow$  less log density and gradient evaluations, but lower acceptance rate and the simulation may diverge



From Monnahan et al (2017)

# Leapfrog discretization, the number of steps

- Many steps can reduce random walk
- Many steps require many log density and gradient evaluations



From Monnahan et al (2017)

# Static Hamiltonian Monte Carlo

- Fixed number of steps
- Demo <https://chi-feng.github.io/mcmc-demo/>

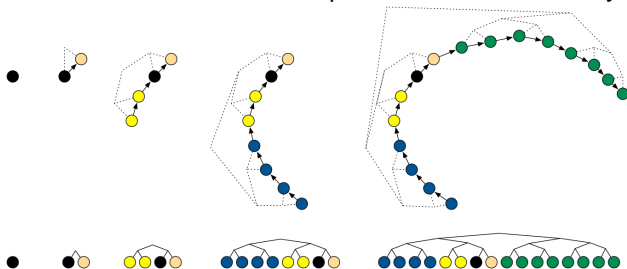
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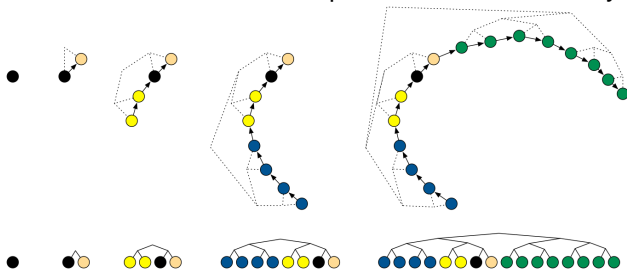
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# Mass matrix and the step size adaptation

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  - mass matrix is estimated during the adaptation phase of the warm-up
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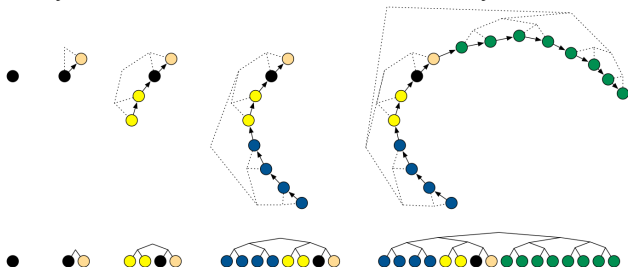
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- Step size
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- After adaptation the algorithm parameters are fixed and some more iterations run to finish the warmup

# Max tree depth diagnostic

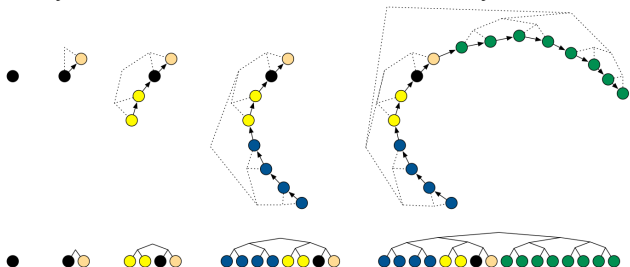
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  - the dynamic simulation is build as a binary tree



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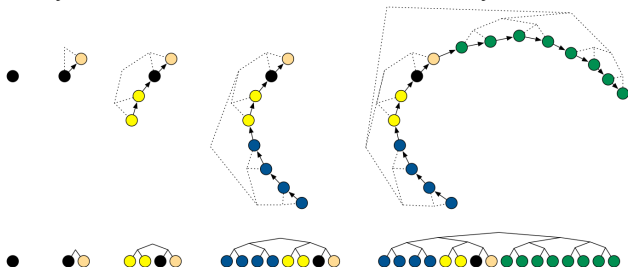


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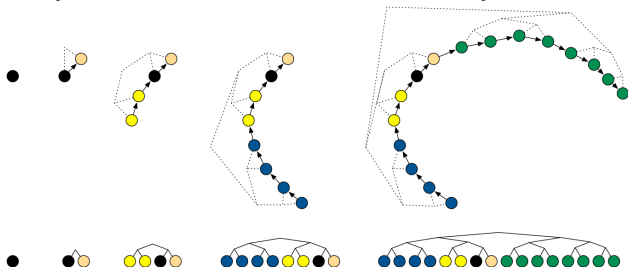


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  - moderate inefficiency doesn't invalidate the result

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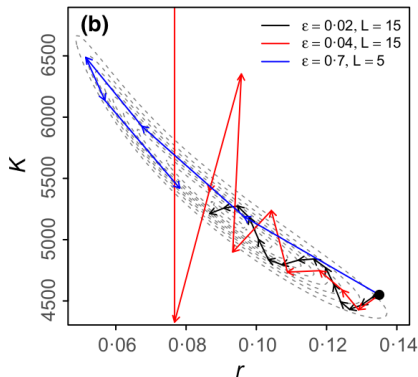


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- Indicates inefficiency in sampling leading to higher autocorrelations and lower ESS ( $S_{\text{eff}}$ )
  - very low inefficiency can indicate problems that need to be inverse-distance
  - moderate inefficiency doesn't invalidate the result
- Different parameterizations matter

# Divergences

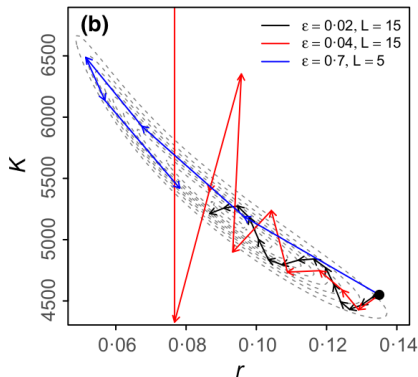
- HMC specific: indicates that Hamiltonian dynamic simulation has problems with unexpected fast changes in log-density (compared to the used step size)
  - indicates possibility of biased estimates



From Monnahan et al (2017)

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From Monnahan et al (2017)

- Demo <https://chi-feng.github.io/mcmc-demo/>

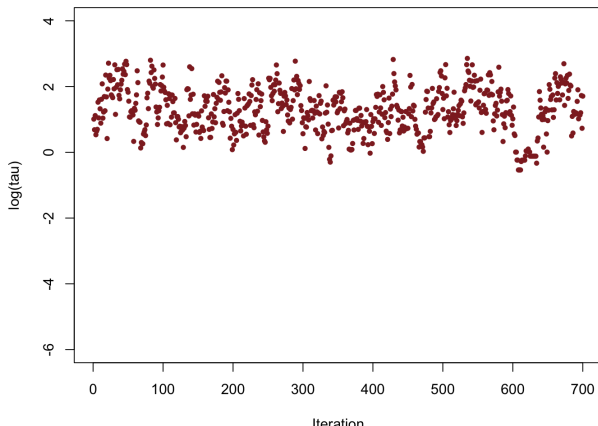


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- [http://mc-stan.org/users/documentation/case-studies/divergences\\_and\\_bias.html](http://mc-stan.org/users/documentation/case-studies/divergences_and_bias.html)

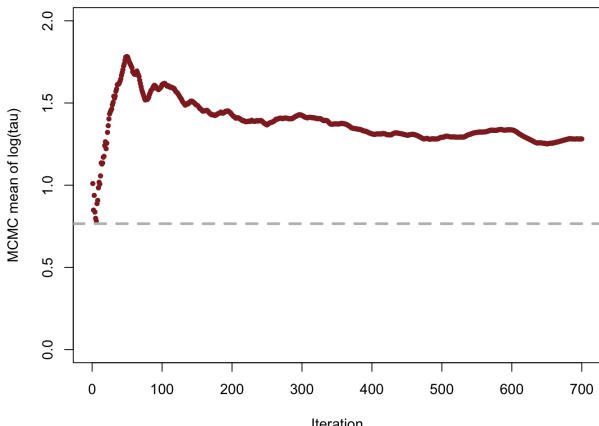
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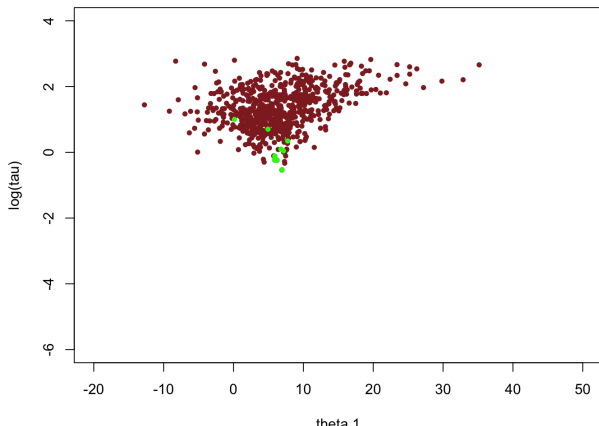
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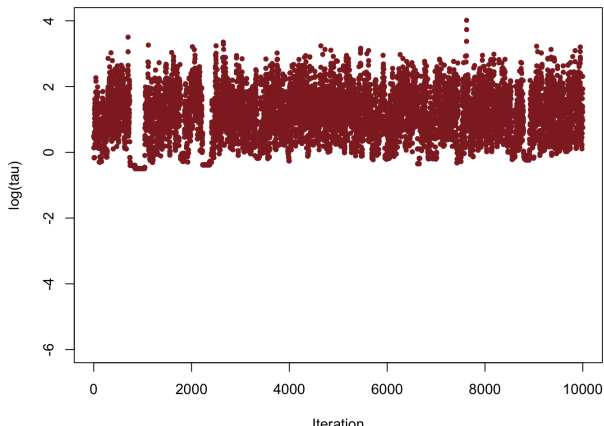
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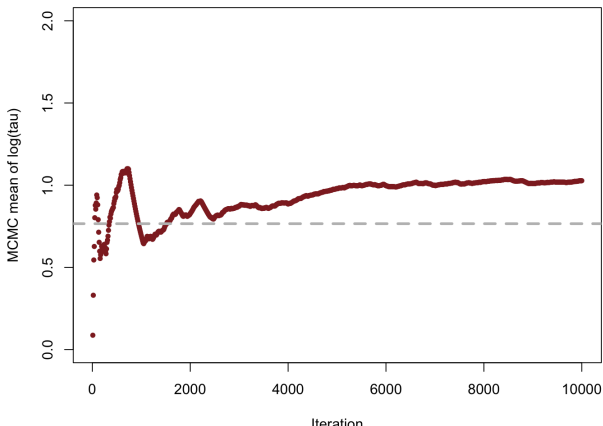
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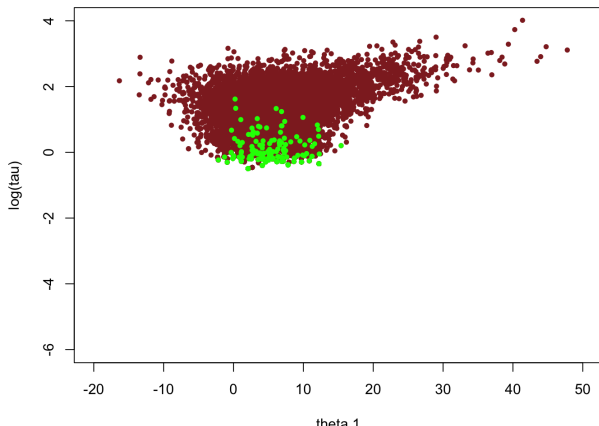
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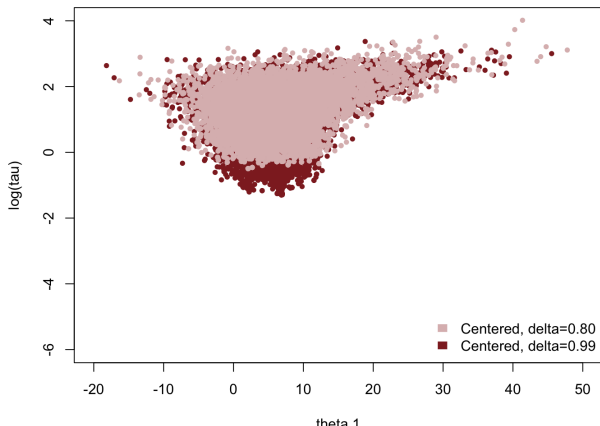
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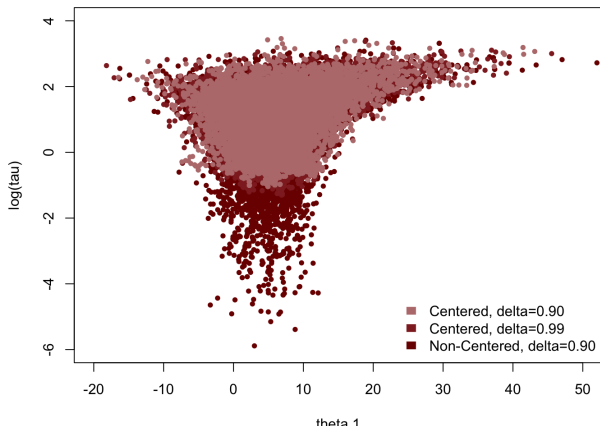
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- Funnels
  - optimal step size depends on location
- Multimodal
  - difficult to move from one mode to another
- Long-tailed with non-finite variance and mean
  - efficiency of exploration is reduced
  - central limit theorem doesn't hold for mean and variance

# Probabilistic programming language

- Wikipedia “A probabilistic programming language (PPL) is a programming language designed to describe probabilistic models and then perform inference in those models”

# Probabilistic programming language

- Wikipedia “A probabilistic programming language (PPL) is a programming language designed to describe probabilistic models and then perform inference in those models”
- To make probabilistic programming useful
  - inference has to be as automatic as possible
  - diagnostics for telling if the automatic inference doesn't work
  - easy workflow (to reduce manual work)
  - fast enough (manual work replaced with automation)

# Probabilistic programming

- Enables agile workflow for developing probabilistic models
  - language
  - automated inference
  - diagnostics
- Many frameworks Stan, PyMC, Pyro (Uber), TFP (Google), Turing.jl, JAGS, ELFI, ...



# Stan - probabilistic programming framework

- Language, inference engine, user interfaces, documentation, case studies, diagnostics, packages, ...
  - autodiff to compute gradients of the log density



mc-stan.org

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mc-stan.org

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mc-stan.org

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  - autodiff to compute gradients of the log density
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- Several full time developers, 40+ developers, more than 100 contributors
- R, Python, Julia, Scala, Stata, Matlab, command line interfaces
- More than 140 R packages using Stan



mc-stan.org

# Stan

- Stanislaw Ulam (1909-1984)
  - Monte Carlo method
  - H-Bomb

## Binomial model - Stan code

```
data {  
  int<lower=0> N;      // number of experiments  
  int<lower=0,upper=N> y; // number of successes  
}  
  
parameters {  
  real<lower=0,upper=1> theta; // parameter of the binomial  
}  
  
model {  
  theta ~ beta(1,1);      //prior  
  y ~ binomial(N,theta); // observation model  
}
```

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- Data type and size are declared
- Stan checks that given data matches type and constraints

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```

- Data type and size are declared
- Stan checks that given data matches type and constraints
  - If you are not used to strong typing, this may feel annoying, but it will reduce the probability of coding errors, which will reduce probability of data analysis errors

# Binomial model - Stan code

```
parameters {  
  real <lower=0,upper=1> theta ;  
}
```

- Parameters may have constraints
- Stan makes transformation to unconstrained space and samples in unconstrained space
  - e.g. log transformation for <lower=a>
  - e.g. logit transformation for <lower=a,upper=b>

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- Parameters may have constraints
- Stan makes transformation to unconstrained space and samples in unconstrained space
  - e.g. log transformation for <lower=a>
  - e.g. logit transformation for <lower=a,upper=b>
- For these declared transformation Stan automatically takes into account the Jacobian of the transformation (see BDA3 p. 21)

## Binomial model - Stan code

```
model {  
  theta ~ beta(1,1);      // prior  
  y ~ binomial(N,theta); // likelihood  
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## Binomial model - Stan code

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model {  
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  y ~ binomial(N,theta); // likelihood  
}
```

~ is syntactic sugar and this is equivalent to

```
model {  
  target += beta_lpdf(theta | 1, 1);  
  target += binomial_lpmf(y | N, theta);  
}
```

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- `target` is the log posterior density

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model {  
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$\sim$  is syntactic sugar and this is equivalent to

```
model {  
  target += beta_lpdf(theta | 1, 1);  
  target += binomial_lpmf(y | N, theta);  
}
```

- `target` is the log posterior density
- `_lpdf` for continuous, `_lpmf` for discrete distributions  
(discrete for the left hand side of `|`)



## Binomial model - Stan code

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model {  
  theta ~ beta(1,1);      // prior  
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- `target` is the log posterior density
- `_lpdf` for continuous, `_lpmf` for discrete distributions (discrete for the left hand side of `|`)
- for Stan sampler there is no difference between prior and likelihood, all that matters is the final `target`
- you can write in Stan language any program to compute the log density (Stan language is Turing complete)

# Stan

- Stan compiles (transpiles) the model written in Stan language to C++
  - this makes the sampling for complex models and bigger data faster
  - also makes Stan models easily portable, you can use your own favorite interface

# CmdStanR

RStan

```
library(cmdstanr)
```

```
options(mc.cores = 1)
```

```
d_bin <- list(N = 10, y = 7)
```

```
mod_bin <- cmdstan_model(stan_file = 'binom.stan')
```

```
fit_bin <- mod_bin$sample(data = d_bin)
```

# CmdStanR

RStan

```
library(cmdstanr)
```

```
options(mc.cores = 1)
```

```
d_bin <- list(N = 10, y = 7)
```

```
mod_bin <- cmdstan_model(stan_file = 'binom.stan')
```

```
fit_bin <- mod_bin$sample(data = d_bin)
```

# PyStan

PyStan

```
import pystan
import stan_utility

data = dict(N=10, y=8)
model = stan_utility.compile_model('binom.stan')
fit = model.sampling(data=data)
```

# PyStan

PyStan

```
import pystan
import stan_utility

data = dict(N=10, y=8)
model = stan_utility.compile_model('binom.stan')
fit = model.sampling(data=data)
```

# Stan

- Compilation (unless previously compiled model available)
- Warm-up including adaptation
- Sampling
- Generated quantities
- Save posterior draws
- Report divergences,  $n_{\text{eff}}$ ,  $\hat{R}$



# Difference between proportions

- An experiment was performed to estimate the effect of beta-blockers on mortality of cardiac patients
- A group of patients were randomly assigned to treatment and control groups:
  - out of 674 patients receiving the control, 39 died
  - out of 680 receiving the treatment, 22 died

# Difference between proportions

```
data {  
  int<lower=0> N1;  
  int<lower=0> y1;  
  int<lower=0> N2;  
  int<lower=0> y2;  
}  
parameters {  
  real<lower=0,upper=1> theta1;  
  real<lower=0,upper=1> theta2;  
}  
model {  
  theta1 ~ beta(1,1);  
  theta2 ~ beta(1,1);  
  y1 ~ binomial(N1,theta1);  
  y2 ~ binomial(N2,theta2);  
}  
  
generated quantities {  
  real oddsratio;  
  oddsratio = (theta2/(1-theta2))/(theta1/(1-theta1));  
}
```

# Difference between proportions

```
data {  
  int<lower=0> N1;  
  int<lower=0> y1;  
  int<lower=0> N2;  
  int<lower=0> y2;  
}  
parameters {  
  real<lower=0,upper=1> theta1;  
  real<lower=0,upper=1> theta2;  
}  
model {  
  theta1 ~ beta(1,1);  
  theta2 ~ beta(1,1);  
  y1 ~ binomial(N1,theta1);  
  y2 ~ binomial(N2,theta2);  
}  
  
generated quantities {  
  real oddsratio;  
  oddsratio = (theta2/(1-theta2))/(theta1/(1-theta1));  
}
```

## Difference between proportions

```
generated quantities {  
  real oddsratio;  
  oddsratio = (theta2/(1 - theta2))/(theta1/(1 - theta1));  
}
```

- generated quantities is run after the sampling

# Difference between proportions

```
d_bin2 <- list(N1 = 674, y1 = 39, N2 = 680, y2 = 22)
mod_bin2 <- cmdstan_model(stan_file = 'binom2.stan')
fit_bin2 <- mod_bin2$sample(data = d_bin2, refresh=1000)
```

```
> Running MCMC with 4 parallel chains...
```

```
Chain 1 Iteration:    1 / 2000 [  0%] (Warmup)
Chain 1 Iteration: 1000 / 2000 [ 50%] (Warmup)
Chain 1 Iteration: 1001 / 2000 [ 50%] (Sampling)
Chain 1 Iteration: 2000 / 2000 [100%] (Sampling)
```

```
...
```

```
All 4 chains finished successfully.
Mean chain execution time: 0.0 seconds.
Total execution time: 0.2 seconds.
```

# Difference between proportions

```
fit_bin2$summary()
```

	variable	mean	median	sd	mad	q5	q95	rhat	ess_bulk	ess_tail
1	lp__	-253.	-253.	1.0	0.74	-255.	-253.	1.0	1751.	2231.
2	theta1	0.059	0.059	0.0093	0.0093	0.045	0.075	1.0	3189.	2657.
3	theta2	0.034	0.033	0.0069	0.0067	0.023	0.046	1.0	3229.	2163.
4	oddsratio	0.57	0.55	0.16	0.15	0.35	0.87	1.0	2998.	2685.

# Difference between proportions

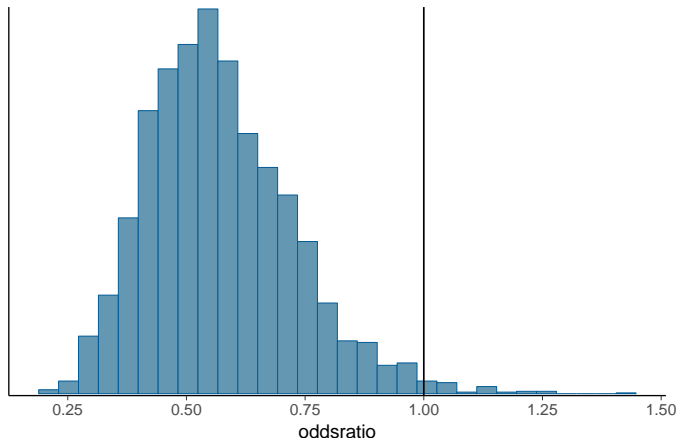
```
fit_bin2$summary()
```

	variable	mean	median	sd	mad	q5	q95	rhat	ess_bulk	ess_tail
1	lp__	-253.	-253.	1.0	0.74	-255.	-253.	1.0	1751.	2231.
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4	oddsratio	0.57	0.55	0.16	0.15	0.35	0.87	1.0	2998.	2685.

- `lp__` is the log density, ie, same as `target`

## Difference between proportions

```
draws <- fit_bin2$draws()  
mcmc_hist(draws, pars = 'oddsratio') +  
  geom_vline(xintercept = 1) +  
  scale_x_continuous(breaks = c(seq(0.25, 1.5, by=0.25)))
```



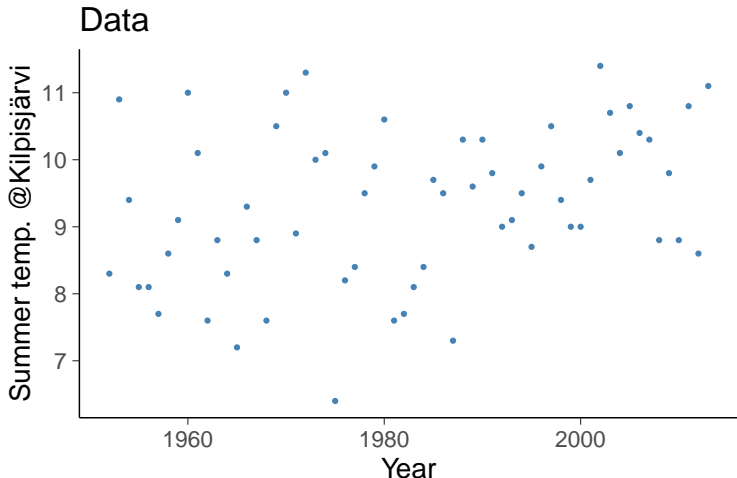


# Shinystan

- Graphical user interface for analysing MCMC results

# Kilpisjärvi summer temperature

- Temperature at Kilpisjärvi in June, July and August from 1952 to 2013
- Is there change in the temperature?



# Gaussian linear model

```
data {  
  int<lower=0> N; // number of data points  
  vector[N] x; //  
  vector[N] y; //  
}  
parameters {  
  real alpha;  
  real beta;  
  real<lower=0> sigma;  
}  
transformed parameters {  
  vector[N] mu;  
  mu <- alpha + beta*x;  
}  
model {  
  y ~ normal(mu, sigma);  
}
```

# Gaussian linear model

```
data {  
    int <lower=0> N; // number of data points  
    vector[N] x; //  
    vector[N] y; //  
}
```

- difference between `vector[N] x` and `array[N] real x`

# Gaussian linear model

```
data {  
    int<lower=0> N; // number of data points  
    vector[N] x; //  
    vector[N] y; //  
}
```

- difference between `vector[N] x` and `array[N] real x`
- no integer vectors: `array[N] int x`

# Gaussian linear model

```
parameters {  
  real alpha;  
  real beta;  
  real<lower=0> sigma;  
}  
transformed parameters {  
  vector[N] mu;  
  mu <- alpha + beta*x;  
}
```

- transformed parameters are deterministic transformations of parameters and data

## Priors for Gaussian linear model

```
data {  
  int<lower=0> N; // number of data points  
  vector[N] x; //  
  vector[N] y; //  
  real pmualpha; // prior mean for alpha  
  real psalpha; // prior std for alpha  
  real pmubeta; // prior mean for beta  
  real psbeta; // prior std for beta  
}  
...  
transformed parameters {  
  vector[N] mu;  
  mu <- alpha + beta*x;  
}  
model {  
  alpha ~ normal(pmualpha, psalpha);  
  beta ~ normal(pmubeta, psbeta);  
  y ~ normal(mu, sigma);  
}
```

# Student-t linear model

```
...
parameters {
  real alpha;
  real beta;
  real<lower=0> sigma;
  real<lower=1,upper=80> nu;
}
transformed parameters {
  vector[N] mu;
  mu <- alpha + beta*x;
}
model {
  nu ~ gamma(2,0.1);
  y ~ student_t(nu, mu, sigma);
}
```

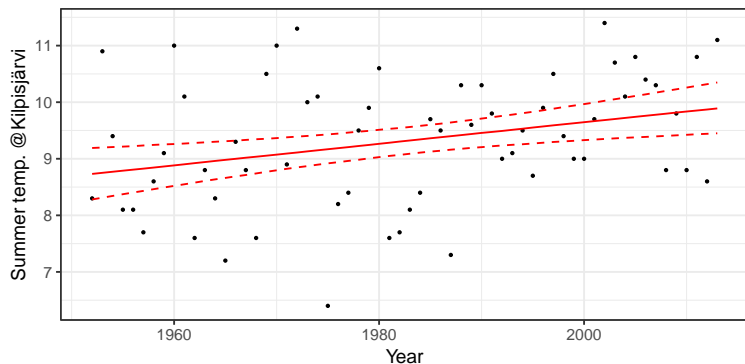


# Priors

- Prior for temperature increase?

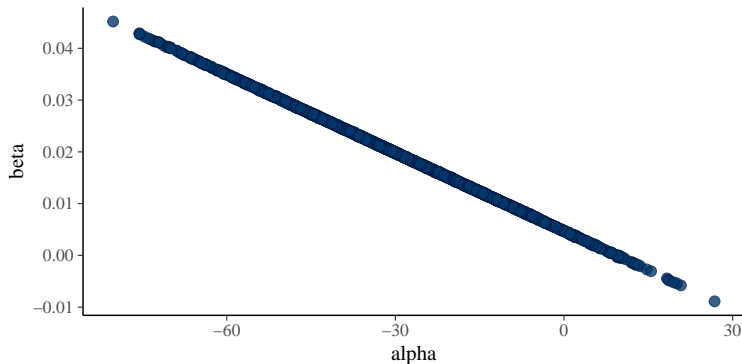
# Kilpisjärvi summer temperature

Posterior fit



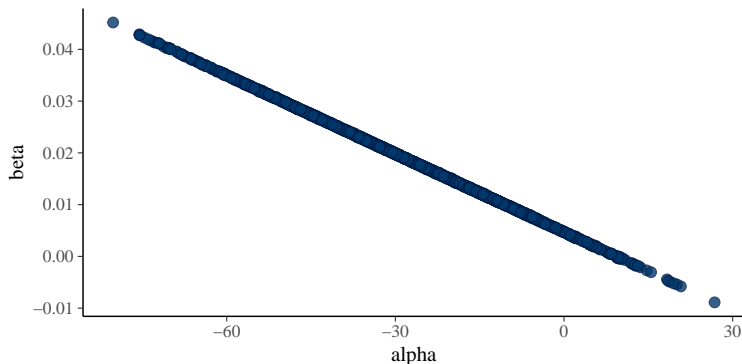
# Kilpisjärvi summer temperature

Posterior draws of alpha and beta



# Kilpisjärvi summer temperature

Posterior draws of alpha and beta



Warning: 1 of 4000 (0.0%) transitions hit the maximum treedepth limit  
See <https://mc-stan.org/misc/warnings> for details.

# Linear regression model in Stan

```
data {  
  int<lower=0> N; // number of data points  
  vector[N] x; //  
  vector[N] y; //  
  real xpred; // input location for prediction  
}  
transformed data {  
  vector[N] x_std;  
  vector[N] y_std;  
  real xpred_std;  
  x_std = (x - mean(x)) / sd(x);  
  y_std = (y - mean(y)) / sd(y);  
  xpred_std = (xpred - mean(x)) / sd(x);  
}
```

# RStanARM

- RStanARM provides simplified model description with pre-compiled models
  - no need to wait for compilation
  - a restricted set of models

Two group Binomial model:

```
d_bin2 <- data.frame(N = c(674, 680), y = c(39,22), grp2 = c(0,1))  
fit_bin2 <- stan_glm(y/N ~ grp2, family = binomial(), data = d_bin2,  
                    weights = N)
```

# RStanARM

- RStanARM provides simplified model description with pre-compiled models
  - no need to wait for compilation
  - a restricted set of models

## Two group Binomial model:

```
d_bin2 <- data.frame(N = c(674, 680), y = c(39,22), grp2 = c(0,1))  
fit_bin2 <- stan_glm(y/N ~ grp2, family = binomial(), data = d_bin2,  
                    weights = N)
```

## Gaussian linear model

```
fit_lin <- stan_glm(temp ~ year, data = d_lin)
```

# BRMS

- BRMS provides simplified model description
  - a larger set of models than RStanARM, but still restricted
  - need to wait for the compilation

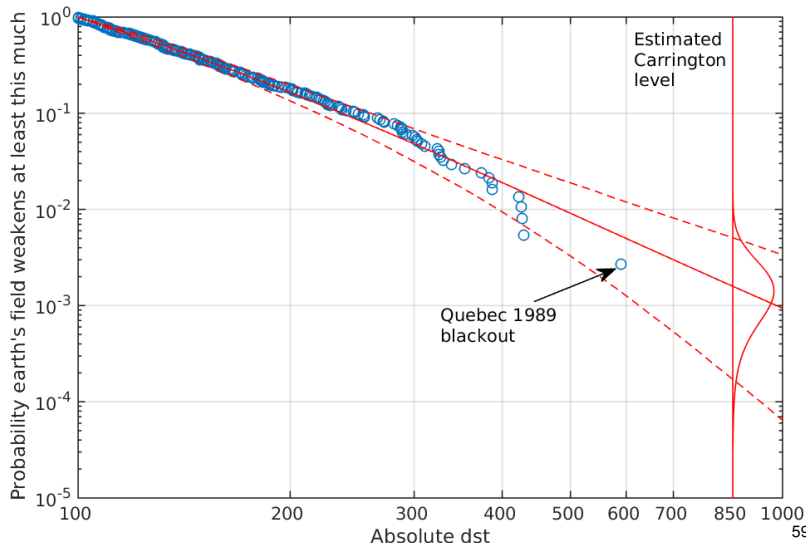
```
fit_bin2 <- brm(y/N ~ grp2, family = binomial(), data = d_bin2,  
               weights = N)
```

```
fit_lin_t <- brm(temp ~ year, data = d_lin, family = student())
```



# Extreme value analysis

## Geomagnetic storms



# Extreme value analysis

```
data {  
  int<lower=0> N;  
  vector<lower=0>[N] y;  
  int<lower=0> Nt;  
  vector<lower=0>[Nt] yt;  
}  
transformed data {  
  real ymax;  
  ymax <- max(y);  
}  
parameters {  
  real<lower=0> sigma;  
  real<lower=-sigma/ymax> k;  
}  
model {  
  y ~ gpareto(k, sigma);  
}  
generated quantities {  
  vector[Nt] predccdf;  
  predccdf<-gpareto_ccdf(yt,k,sigma);  
}
```

# Functions

```
functions {  
  real gpareto_lpdf(vector y, real k, real sigma) {  
    // generalised Pareto log pdf with mu=0  
    // should check and give error if k<0  
    // and max(y)/sigma > -1/k  
    int N;  
    N <- dims(y)[1];  
    if (fabs(k) > 1e-15)  
      return -(1+1/k)*sum(log1pv(y*k/sigma)) -N*log(sigma);  
    else  
      return -sum(y/sigma) -N*log(sigma); // limit k->0  
  }  
  vector gpareto_ccdf(vector y, real k, real sigma) {  
    // generalised Pareto log ccdf with mu=0  
    // should check and give error if k<0  
    // and max(y)/sigma < -1/k  
    if (fabs(k) > 1e-15)  
      return exp((-1/k)*log1pv(y/sigma*k));  
    else  
      return exp(-y/sigma); // limit k->0  
  }  
}
```

# Other packages

- R
  - posterior — posterior handling and diagnostics
  - shinystan — interactive diagnostics
  - bayesplot — visualization and model checking (see model checking in Ch 6)
  - tideybayes and ggdist — more visualization
  - loo — cross-validation model assessment and comparison (see Ch 7)
  - projpred — projection predictive variable selection
  - priorsense — prior and likelihood sensitivity diagnostics
- Python
  - ArviZ — visualization, and model checking and assessment (see Ch 6 and 7)

# Different interfaces

- CmdStanR / CmdStanPy
  - Lightweight interface on top of commandline program CmdStan
  - Lacks some features that are not needed in this course, but is usually easier to install
- RStan / PyStan
  - C++ functions of Stan are called directly from R / Python
  - Higher integration between R/Python and Stan, but maybe more difficult to install due to more requirements of compatible C++ compilers and libraries
- More recent useful R packages
  - posterior: for handling posterior draws, convergence diagnostics, and summaries
  - tidybayes + ggdist: pretty plots