# Chapter 7

- 7.1 Measures of predictive accuracy
- 7.2 Information criteria and cross-validation
  - Instead of 7.2, read:
     Vehtari, A., Gelman, A., Gabry, J. (2017). Practical Bayesian model evaluation using leave-one-out cross-validation and WAIC. Statistics and Computing. 27(5):1413–1432. arXiv preprint.
- 7.3 Model comparison based on predictive performance
- 7.4 Model comparison using Bayes factors
- 7.5 Continuous model expansion / sensitivity analysis
- 7.5 Example (may be skipped)

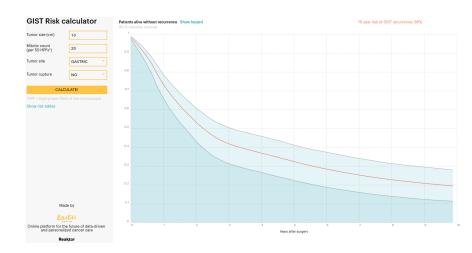
# Model assessment, selection and inference after selection

- Extra material at https://avehtari.github.io/modelselection/
   Videos, Slides, Notebooks, References
- CV-FAQ https://avehtari.github.io/modelselection/CV-FAQ.html

# Predicting concrete quality



# Predicting cancer recurrence



- ▶ True predictive performance is found out by using it to make predictions and comparing predictions to true observations
  - external validation

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  - external validation
- Expected predictive performance
  - approximates the external validation

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- Application specific utility/cost functions are important
  - eg. money, life years, quality adjusted life years, etc.

- We need to choose the utility/cost function
- ► Application specific utility/cost functions are important
  - eg. money, life years, quality adjusted life years, etc.
- ► If are interested overall in the goodness of the predictive distribution, or we don't know (yet) the application specific utility, then good information theoretically justified choice is log-score

$$\log p(y^{\text{rep}} \mid y, M),$$

- What is cross-validation
  - Leave-one-out cross-validation (elpd\_loo, p\_loo)
  - Uncertainty in LOO (SE)
- When is cross-validation applicable?
  - data generating mechanisms and prediction tasks
  - leave-many-out cross-validation
- Fast cross-validation
  - PSIS and diagnostics in loo package (Pareto k, n\_eff, Monte Carlo SE)
  - K-fold cross-validation
- Related methods (WAIC, \*IC, BF)
- Model comparison and selection (elpd\_diff, se)
- Model averaging with Bayesian stacking

# Stan and loo package

Computed from 4000 by 20 log-likelihood matrix

Monte Carlo SE of elpd\_loo is 0.1.

Pareto k diagnostic values:

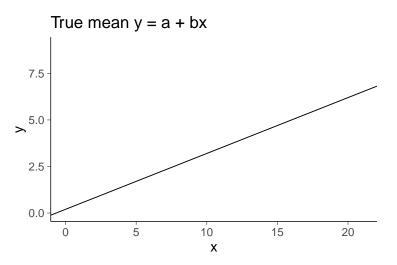
Count Pct. Min. n\_eff

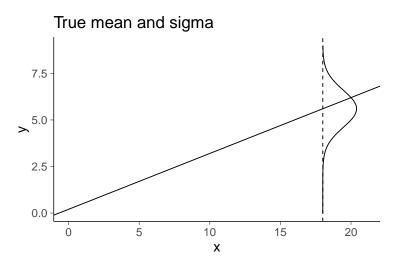
		Count	PCI.	win.	n_eff
(-Inf, 0.5]	(good)	18	90.0%	899	
(0.5, 0.7]	(ok)	2	10.0%	459	
(0.7, 1]	(bad)	0	0.0%	<NA $>$	
(1, Inf)	(very bad)	0	0.0%	<NA $>$	

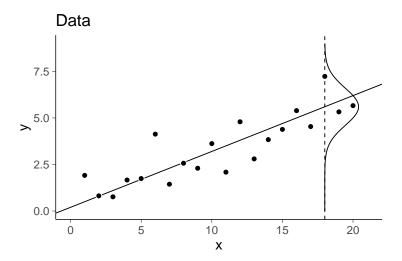
All Pareto k estimates are ok (k < 0.7). See help('pareto-k-diagnostic') for details.

Model comparison: (negative 'elpd\_diff' favors 1st model, positive favors 2nd)

$$\begin{array}{ccc} \text{elpd\_diff} & \text{se} \\ -0.2 & 0.1 \end{array}$$

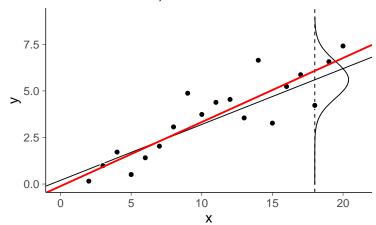




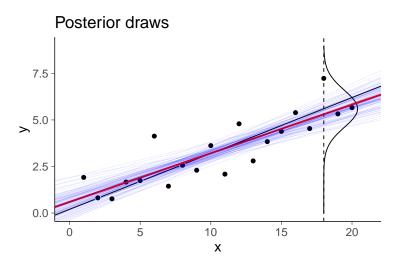


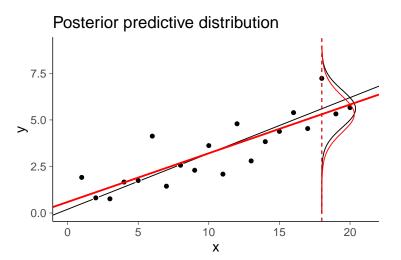


## Posterior mean, alternative data realisation

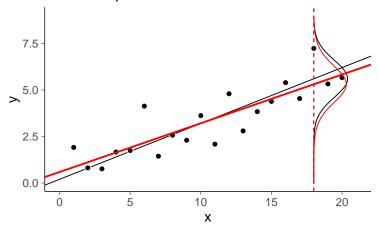




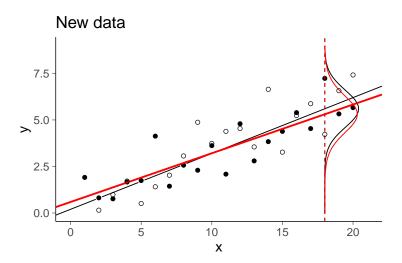


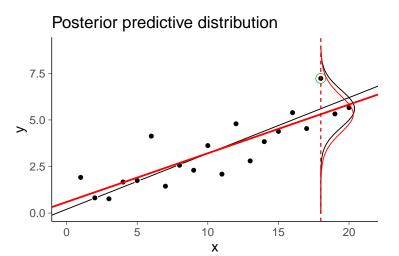


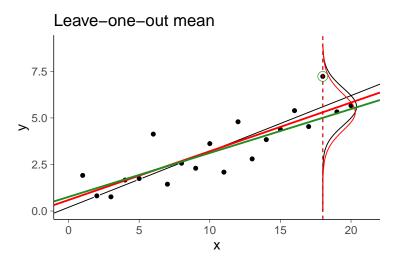
## Posterior predictive distribution

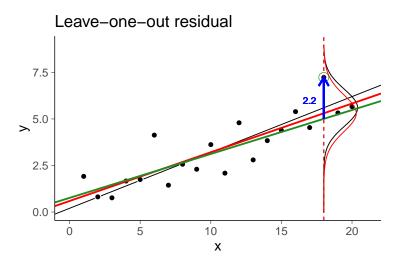


$$p(\tilde{y} \mid \tilde{x} = 18, x, y) = \int p(\tilde{y} \mid \tilde{x} = 18, \theta) p(\theta \mid x, y) d\theta$$

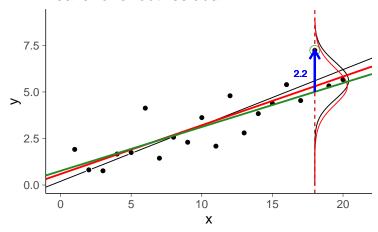






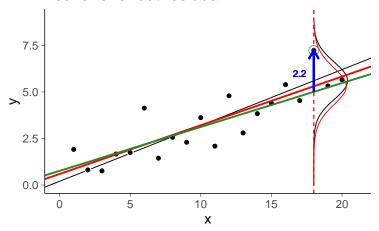


#### Leave-one-out residual



$$y_{18} - E[p(\tilde{y} \mid \tilde{x} = 18, x_{-18}, y_{-18})]$$

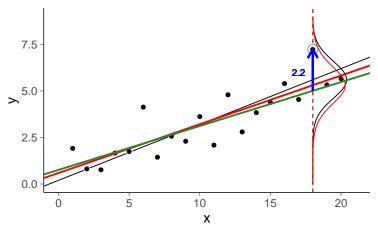
#### Leave-one-out residual



$$y_{18} - E[p(\tilde{y} \mid \tilde{x} = 18, x_{-18}, y_{-18})]$$

Can be use to compute, e.g., RMSE, R<sup>2</sup>, 90% error

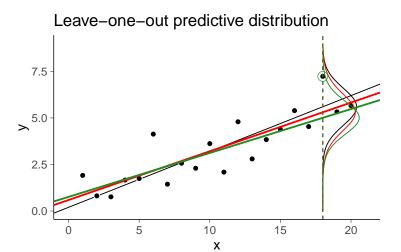
#### Leave-one-out residual



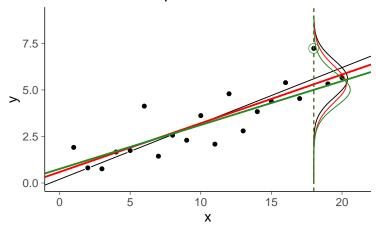
$$y_{18} - E[p(\tilde{y} \mid \tilde{x} = 18, x_{-18}, y_{-18})]$$

Can be use to compute, e.g., RMSE, R<sup>2</sup>, 90% error

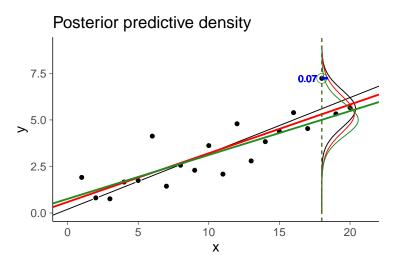
See LOO-R<sup>2</sup> at avehtari.github.io/bayes\_R2/bayes\_R2.html



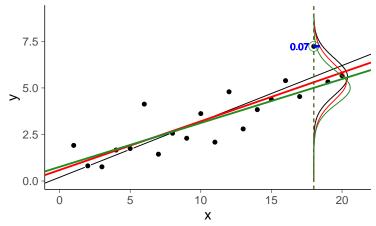
## Leave-one-out predictive distribution



$$p(\tilde{y} \mid \tilde{x} = 18, x_{-18}, y_{-18}) = \int p(\tilde{y} \mid \tilde{x} = 18, \theta) p(\theta \mid x_{-18}, y_{-18}) d\theta$$

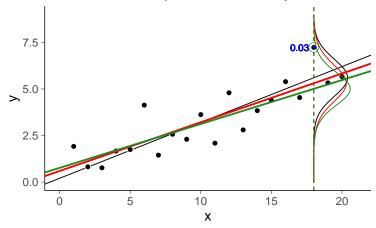


## Posterior predictive density



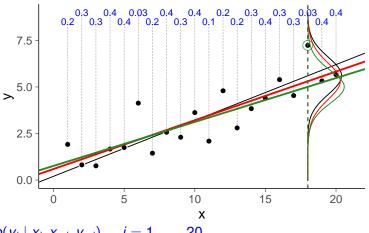
$$p(\tilde{y} = y_{18} \mid \tilde{x} = 18, x, y) \approx 0.07$$

## Leave-one-out predictive density



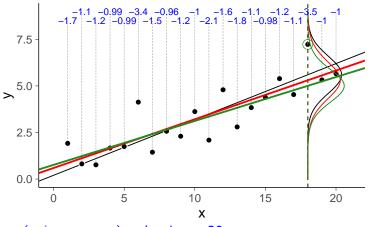
$$p(\tilde{y} = y_{18} \mid \tilde{x} = 18, x, y) \approx 0.07$$
  
 $p(\tilde{y} = y_{18} \mid \tilde{x} = 18, x_{-18}, y_{-18}) \approx 0.03$ 

## Leave-one-out predictive densities



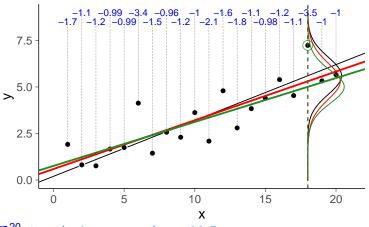
 $p(y_i \mid x_i, x_{-i}, y_{-i}), \quad i = 1, \dots, 20$ 

## Leave-one-out log predictive densities



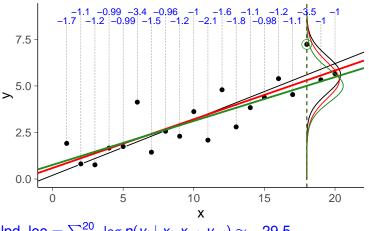
 $\log p(y_i\mid x_i,x_{-i},y_{-i}),\quad i=1,\ldots,20$ 

## Leave–one–out log predictive densities



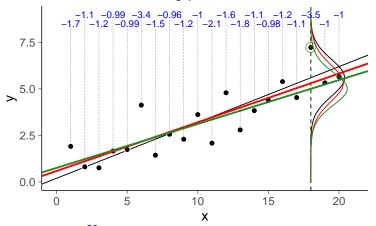
 $\sum_{i=1}^{20} \log p(y_i \mid x_i, x_{-i}, y_{-i}) \approx -29.5$ 

## Leave-one-out log predictive densities



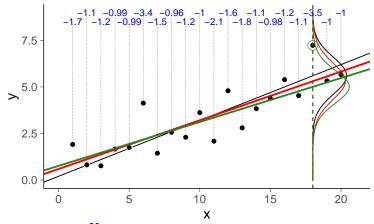
elpd\_loo =  $\sum_{i=1}^{20} \log p(y_i \mid x_i, x_{-i}, y_{-i}) \approx -29.5$ 

## Leave-one-out log predictive densities



elpd\_loo =  $\sum_{i=1}^{20} \log p(y_i \mid x_i, x_{-i}, y_{-i}) \approx -29.5$  an estimate of log posterior pred. density for new data

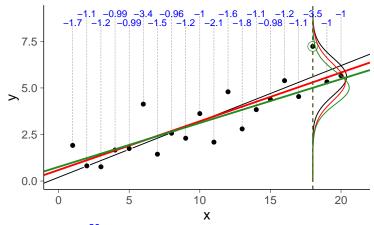
#### Leave–one–out log predictive densities



elpd\_loo = 
$$\sum_{i=1}^{20} \log p(y_i \mid x_i, x_{-i}, y_{-i}) \approx -29.5$$

$$lpd = \sum_{i=1}^{20} log p(y_i | x_i, x, y) \approx -26.8$$

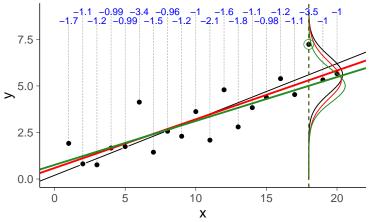
#### Leave–one–out log predictive densities



elpd\_loo = 
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$$p\_loo = lpd - elpd\_loo \approx 2.7$$

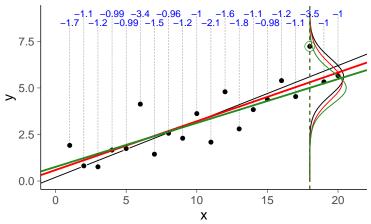
#### Leave-one-out log predictive densities



elpd\_loo = 
$$\sum_{i=1}^{20} \log p(y_i \mid x_i, x_{-i}, y_{-i}) \approx -29.5$$
  
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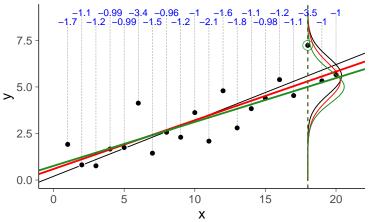
asymptotically approaches p in case of regular faithful model

#### Leave-one-out log predictive densities



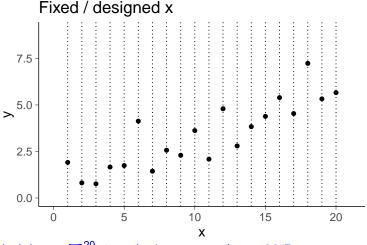
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#### Leave-one-out log predictive densities



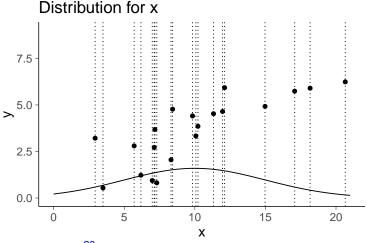
elpd\_loo = 
$$\sum_{i=1}^{20} \log p(y_i \mid x_i, x_{-i}, y_{-i}) \approx -29.5$$
  
SE = sd(log  $p(y_i \mid x_i, x_{-i}, y_{-i})) \cdot \sqrt{20} \approx 3.3$ 

see Vehtari, Gelman & Gabry (2017a) and Vehtari & Ojanen (2012) for more



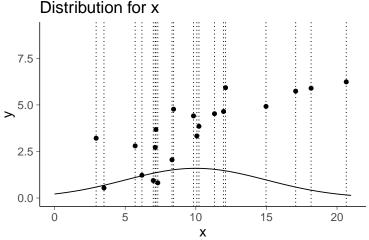
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LOO is ok for fixed / designed x. SE is uncertainty about  $y \mid x$ .



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LOO is ok for random x. SE is uncertainty about  $y \mid x$  and x.



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LOO is ok for random x. SE is uncertainty about  $y \mid x$  and x. Covariate shift can be handled with importance weighting or modelling

#### 100 package

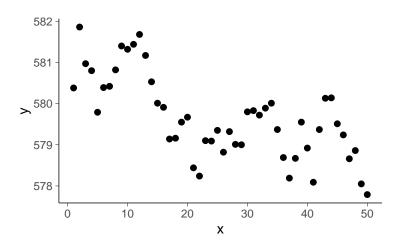
Computed from 4000 by 20 log-likelihood matrix

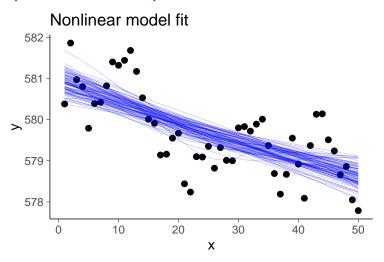
Monte Carlo SE of elpd\_loo is 0.1.

Pareto k diagnostic values:

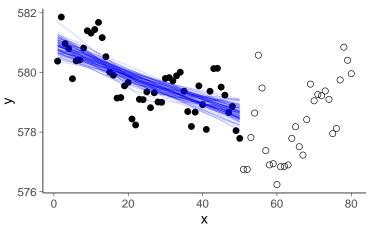
		Count	Pct.	Min.	n_eff
(-Inf, 0.5]	] (good)	18	90.0%	899	
(0.5, 0.7	] (ok)	2	10.0%	459	
(0.7, 1	] (bad)	0	0.0%	<NA $>$	
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All Pareto k estimates are ok (k < 0.7). See help('pareto-k-diagnostic') for details.

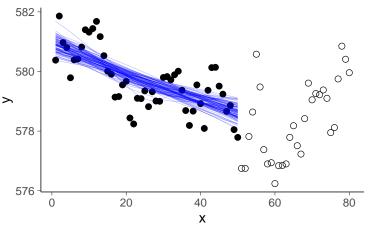




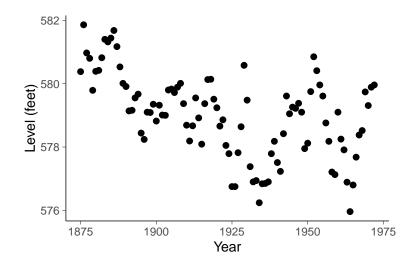




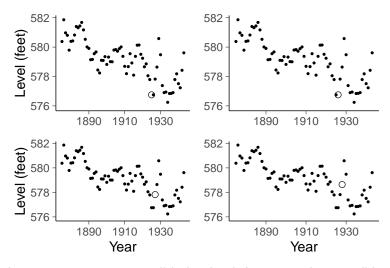




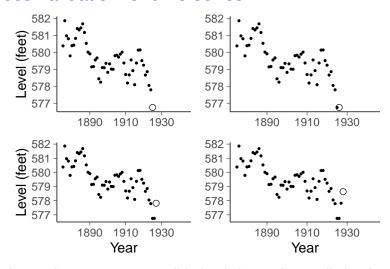
Extrapolation is more difficult



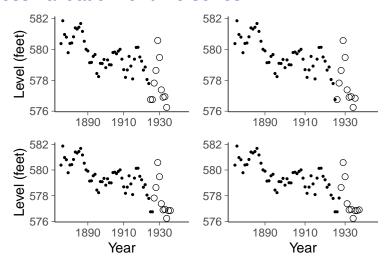
Can LOO or other cross-validation be used with time series?



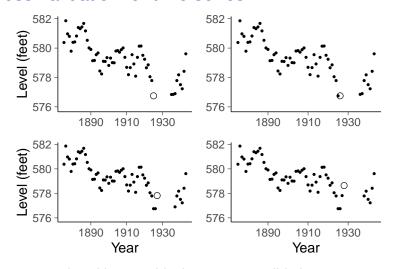
Leave-one-out cross-validation is ok for assessing conditional model



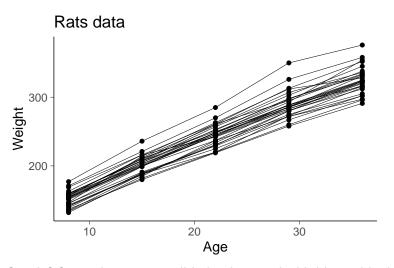
Leave-future-out cross-validation is better for predicting future



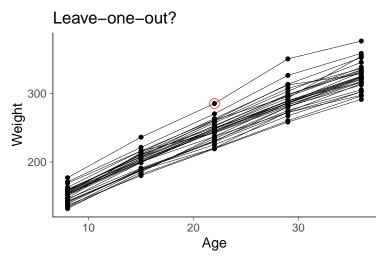
*m*-step-ahead cross-validation is better for predicting further future

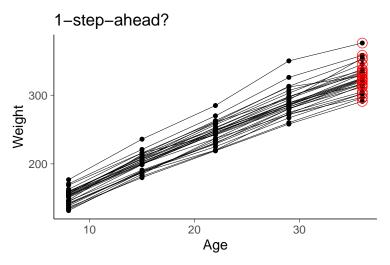


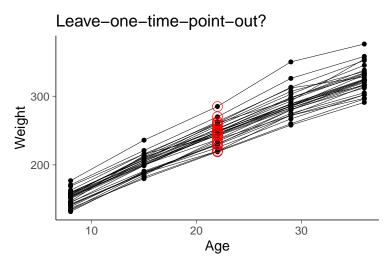
m-step-ahead leave-a-block-out cross-validation

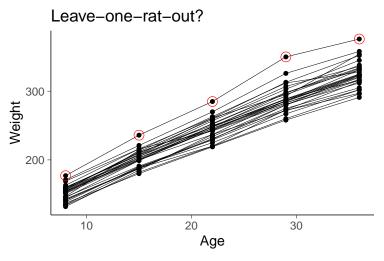


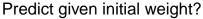
Can LOO or other cross-validation be used with hierarchical data?

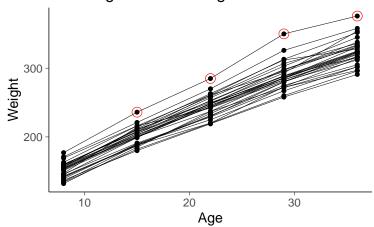












# Summary of data generating mechanisms and prediction tasks

- You have to make some assumptions on data generating mechanism
- Use the knowledge of the prediction task if available
- Cross-validation can be used to analyse different parts, even if there is no clear prediction task

#### Fast cross-validation

- Pareto smoothed importance sampling LOO (PSIS-LOO)
- K-fold cross-validation

• We want to compute

$$p(y_i \mid x_{-i}, y_{-i}) = \int p(y_i \mid x_i, \theta) p(\theta \mid x_{-i}, y_{-i}) d\theta$$

- We want to compute  $p(y_i \mid x_{-i}, y_{-i}) = \int p(y_i \mid x_i, \theta) p(\theta \mid x_{-i}, y_{-i}) d\theta$
- Proposal distribution is full posterior  $\theta^{(s)} \sim p(\theta \mid x, y)$
- Target distribution is LOO-posterior  $p(\theta \mid x_{-i}, y_{-i})$

• We want to compute  $p(y_i \mid x_{-i}, y_{-i}) = \int p(y_i \mid x_i, \theta) p(\theta \mid x_{-i}, y_{-i}) d\theta$ 

• Proposal distribution is full posterior 
$$\theta^{(s)} \sim p(\theta \mid x, y)$$

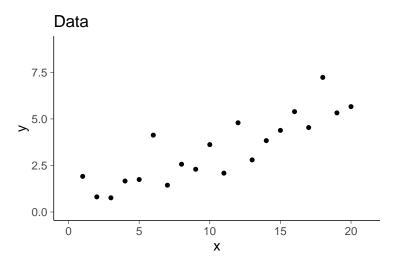
- Target distribution is LOO-posterior  $p(\theta \mid x_{-i}, y_{-i})$
- Importance ratio

$$w_i^{(s)} = \frac{p(\theta^{(s)} \mid x_{-i}, y_{-i})}{p(\theta^{(s)} \mid x, y)} \propto \frac{1}{p(y_i \mid \theta^{(s)})}$$

• We want to compute  $p(y_i \mid x_{-i}, y_{-i}) = \int p(y_i \mid x_i, \theta) p(\theta \mid x_{-i}, y_{-i}) d\theta$ 

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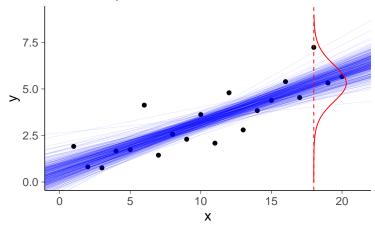
$$w_i^{(s)} = \frac{p(\theta^{(s)} \mid x_{-i}, y_{-i})}{p(\theta^{(s)} \mid x, y)} \propto \frac{1}{p(y_i \mid \theta^{(s)})}$$
$$\tilde{w}_i^{(s)} = \frac{w_i^{(s)}}{\sum_{s'=1}^{S} w_i^{(s')}}$$



## Posterior draws 7.5 > 5.0 2.5 0.0 10 15 20 5 Χ

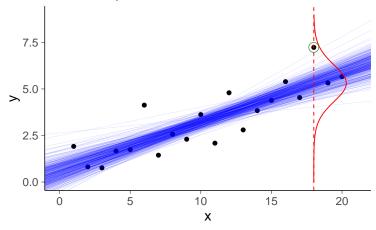
$$\theta^{(s)} \sim p(\theta \mid x, y)$$

#### Posterior predictive distribution



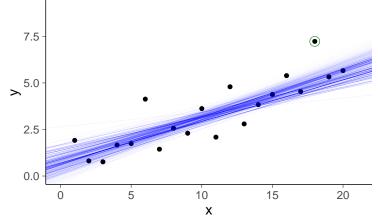
$$\theta^{(s)} \sim p(\theta \mid x,y), \quad p(\tilde{y} \mid \tilde{x},x,y) \approx \frac{1}{S} \sum_{s=1}^{S} p(\tilde{y} \mid \tilde{x},\theta^{(s)})$$

#### Posterior predictive distribution



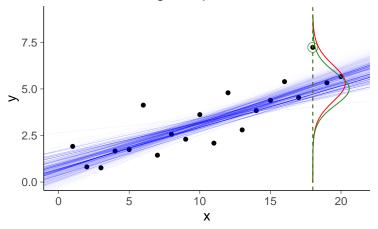
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#### PSIS-LOO weighted draws



$$\theta^{(s)} \sim p(\theta \mid x, y), \quad w_i^{(s)} = p(\theta^{(s)} \mid x_{-i}, y_{-i}) / p(\theta^{(s)} \mid x, y)$$

#### PSIS-LOO weighted predictive distribution



$$\begin{split} \theta^{(s)} &\sim p(\theta \mid x, y), \quad w_i^{(s)} = p(\theta^{(s)} \mid x_{-i}, y_{-i}) / p(\theta^{(s)} \mid x, y) \\ p(y_i \mid x_i, x_{-i}, y_{-i}) &\approx \sum_{s=1}^{S} [\tilde{w}_i^{(s)} p(y_i \mid x_i, \theta^{(s)})] \end{split}$$

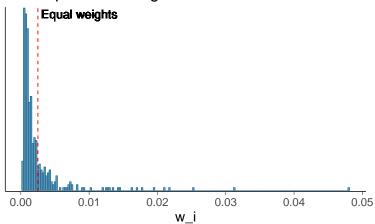
We want to compute

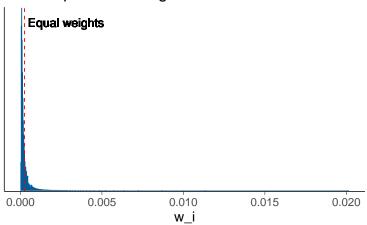
$$p(y_i \mid x_{-i}, y_{-i}) = \int p(y_i \mid x_i, \theta) p(\theta \mid x_{-i}, y_{-i}) d\theta$$

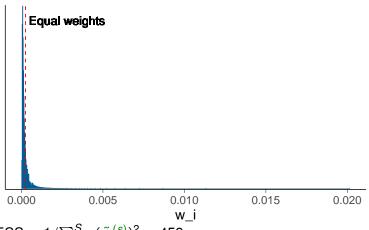
- Proposal distribution is full posterior  $\theta^{(s)} \sim p(\theta \mid x, y)$
- Target distribution is LOO-posterior  $p(\theta \mid x_{-i}, y_{-i})$
- Importance ratio

$$w_i^{(s)} = \frac{p(\theta^{(s)} \mid x_{-i}, y_{-i})}{p(\theta^{(s)} \mid x, y)} \propto \frac{1}{p(y_i \mid \theta^{(s)})}$$
$$\tilde{w}_i^{(s)} = \frac{w_i^{(s)}}{\sum_{s'=1}^{S} w_i^{(s')}}$$

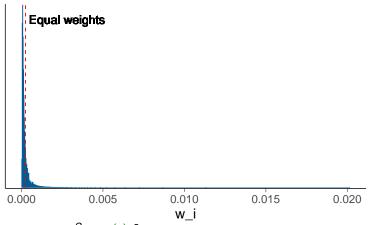
- The variability of importance weights matter
  - Pareto-k diagnostic
  - Pareto smoothed importance sampling LOO (PSIS-LOO)







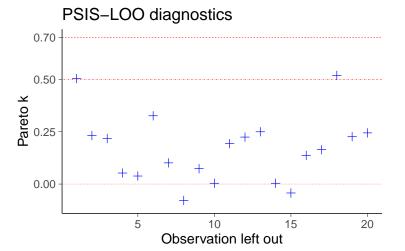
ESS 
$$\approx 1/\sum_{s=1}^{S} (\tilde{w}^{(s)})^2 \approx 459$$



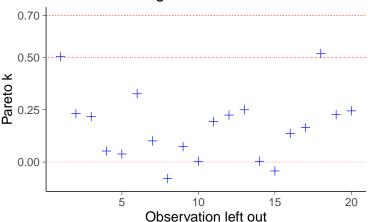
ESS 
$$\approx 1/\sum_{s=1}^{S} (\tilde{\textit{w}}^{(s)})^2 \approx 459$$

Pareto  $\hat{k} \approx 0.52$ 

- Pareto  $\hat{k}$  estimates the tail shape which determines the convergence rate of PSIS. Less than 0.7 is ok.



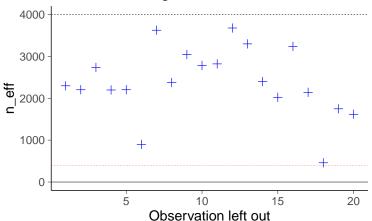
### PSIS-LOO diagnostics



#### Pareto k diagnostic values:

		Count	Pct.	Min .	n_eft
(-Inf, 0.5]	(good)	18	90.0%	899	
(0.5, 0.7]	(ok)	2	10.0%	459	
(0.7, 1]	(bad)	0	0.0%	<NA $>$	
(1, Inf)	(very bad)	0	0.0%	<NA $>$	

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### 100 package

Computed from 4000 by 20 log-likelihood matrix

Monte Carlo SE of elpd\_loo is 0.1.

# Pareto k diagnostic values:

		Count	Pct.	Mın.	n_eff
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All Pareto k estimates are ok (k < 0.7). See help('pareto-k-diagnostic') for details.

see more in Vehtari, Gelman & Gabry (2017b)

## Pareto smoothed importance sampling (PSIS)

- Replace the largest weights with ordered statistics of the fitted Pareto distribution
  - equivalent to using model to filter the noise out of the weights

See more in Vehtari, Simpson, Gelman, Yao & Gabry (2021)

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# Pareto smoothed importance sampling (PSIS)

- Replace the largest weights with ordered statistics of the fitted Pareto distribution
  - equivalent to using model to filter the noise out of the weights
- Reduced variability compared to the plain IS
- Reduced bias compared to the truncated IS
- Asymptotically consistent under some mild conditions

See more in Vehtari, Simpson, Gelman, Yao & Gabry (2021)

### Stan code

$$\log(w_i^{(s)}) = \log(1/p(y_i \mid x_i, \theta^{(s)})) = -\log_{\text{lik}[i]}$$

### Stan code

```
\log(w_i^{(s)}) = \log(1/p(y_i \mid x_i, \theta^{(s)})) = -\log_{\text{lik}[i]}
model {
  alpha ~ normal(pmualpha, psalpha);
  beta ~ normal(pmubeta, psbeta);
  y ~ normal(mu, sigma);
generated quantities {
  vector[N] log lik;
  for (i in 1:N)
    log lik[i] = normal lpdf(v[i] | mu[i], sigma);
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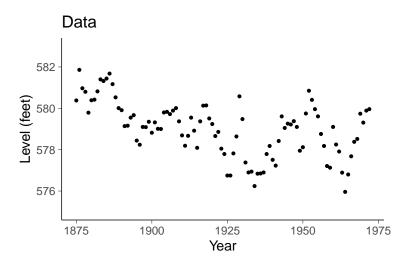
RStanARM and BRMS compute log\_lik by default

- PSIS-LOO for hierarchical models
  - leave-one-group out is challenging for PSIS-LOO
  - see Merkel, Furr and Rabe-Hesketh (2018) for an approach using quadrature integration
  - Stan demo of the challenges and integrated LOO at https://avehtari.github.io/modelselection/roaches.html

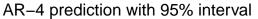
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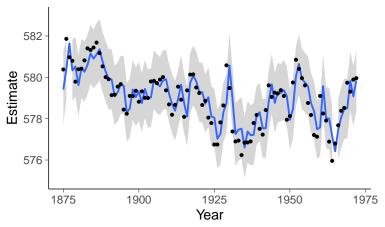
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  - mc-stan.org/loo/articles/loo2-non-factorizable.html
- PSIS-LOO for time series
  - Approximate leave-future-out cross-validation (LFO-CV)
    - mc-stan.org/loo/articles/loo2-lfo.html

# LFO-CV



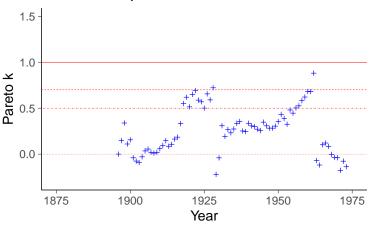
### LFO-CV





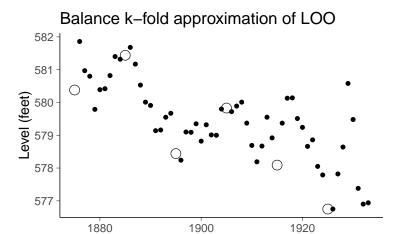
### LFO-CV

### PSIS-1-step-ahead with refits



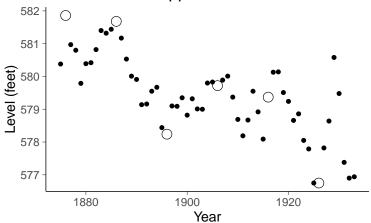
### K-fold cross-validation

- K-fold cross-validation can approximate LOO
  - the same use cases as with LOO
- K-fold cross-validation can be used for hierarchical models
  - good for leave-one-group-out
- K-fold cross-validation can be used for time series
  - with leave-block-out

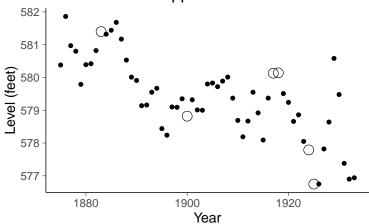


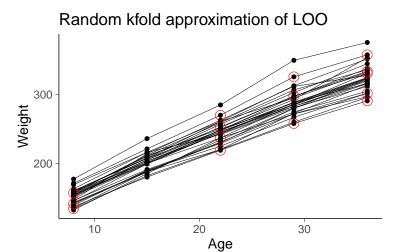
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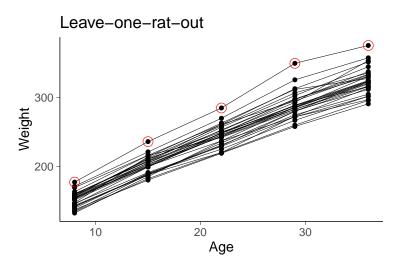
### Balance k-fold approximation of LOO

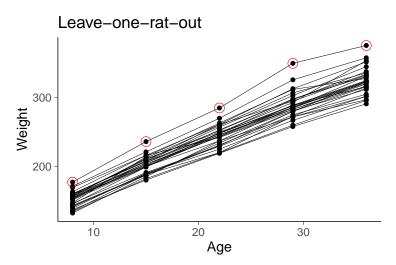


# Random k-fold approximation of LOO









kfold\_split\_random()
kfold\_split\_balanced()
kfold\_split\_stratified()

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- Multiplying by -2 doesn't give any benefit (Watanabe didn't multiply by -2)

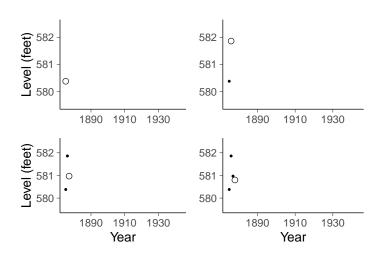
### \*IC

- AIC uses maximum likelihood estimate for prediction
- DIC uses posterior mean for prediction
- BIC is an approximation for marginal likelihood
- TIC, NIC, RIC, PIC, BPIC, QIC, AICc, ...

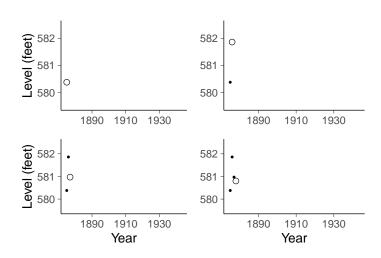
### Marginal likelihood / Bayes factor

 Like leave-future-out 1-step-ahead cross-validation but starting with 0 observations

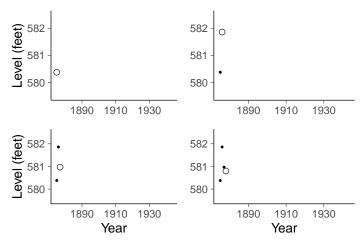
 Like leave-future-out 1-step-ahead cross-validation but starting with 0 observations



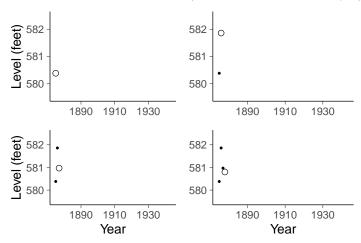
- Like leave-future-out 1-step-ahead cross-validation but starting with 0 observations
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#### Cross-validation for model assessment

- CV is good for model assessment when application specific utility/cost functions are used
  - e.g. 90% absolute error

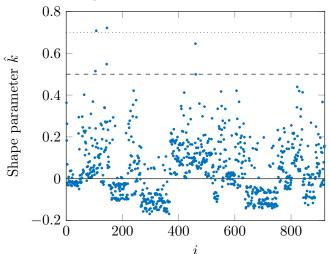
#### Cross-validation for model assessment

- CV is good for model assessment when application specific utility/cost functions are used
  - e.g. 90% absolute error
- Also useful in model checking in similar way as posterior predictive checking (PPC)
  - model misspecification diagnostics (e.g. Pareto-k and p\_loo)
  - checking calibration of leave-one-out predictive posteriors (ppc\_loo\_pit in bayesplot)

see demos avehtari.github.io/modelselection/

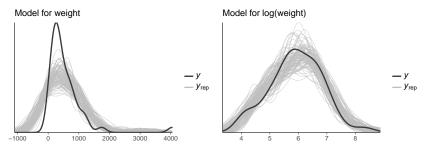
#### Radon example

#### PSIS-LOO diagnostics



see Vehtari, Gelman & Gabry (2017a)

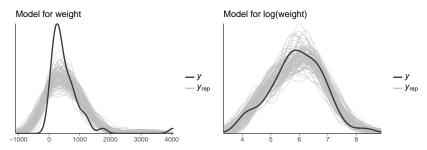
Posterior predictive checking is often sufficient



Predicting the yields of mesquite bushes.

Gelman, Hill & Vehtari (2020): Regression and Other Stories, Chapter 11.

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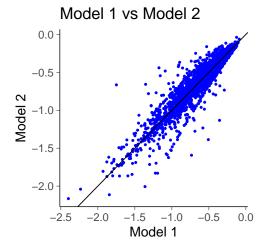
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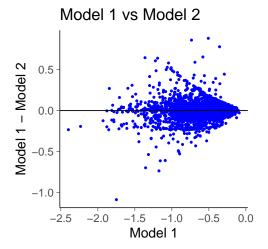
- BDA3, Chapter 6
- Gabry, Simpson, Vehtari, Betancourt, Gelman (2019). Visualization in Bayesian workflow. JRSS A, https://doi.org/10.1111/rssa.12378
- mc-stan.org/bayesplot/articles/graphical-ppcs.html

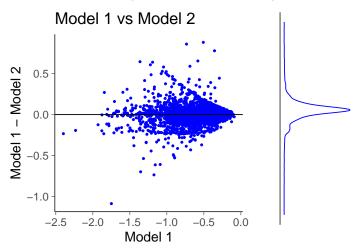
 With good priors that keep the prior on predictive space consistent, there is no need to do model selection to avoid overfitting

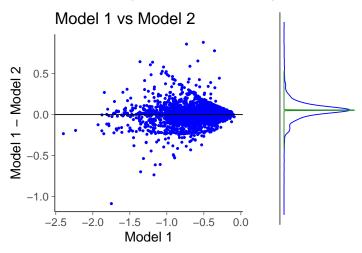
- Logistic regression for predicting probability of switching well with high arsenic level in rural Bangladesh
  - Model 1: log(arsenic) + distance
  - Model 2: log(arsenic) + distance + education level

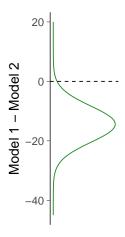


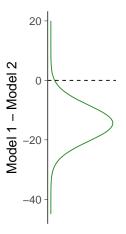
Model 1:  $\widehat{\text{elpd}}_{\text{Loo}}(\mathbf{M}_a \mid y^{\text{obs}}) \approx \text{-1952}$ , SE=16 Model 2:  $\widehat{\text{elpd}}_{\text{Loo}}(\mathbf{M}_b \mid y^{\text{obs}}) \approx \text{-1938}$ , SE=17











1. The models make very similar predictions

2. The models are misspecified with outliers in the data

3. The number of observations is small

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  - if  $|\widehat{\text{elpd}}_{\text{Loo}}(M_a, M_b \mid y^{\text{obs}})| < 4$ , SE is not reliable, but the difference is small anyway
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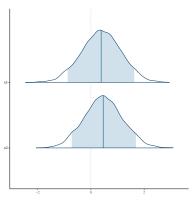
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  - in nested case the skewness favors the simpler model
  - any inference with small *n* is difficult
  - if  $|\widehat{\text{elpd}}_{\text{Loo}}(M_a, M_b \mid y^{\text{obs}})| > 4$ , model is well specified, and n > 100 then the normal approximation is good

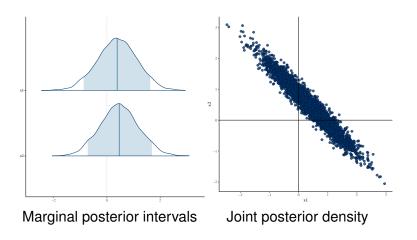
 In nested case, often easier and more accurate to analyse posterior distribution of more complex model directly avehtari.github.io/modelselection/betablockers.html

## Sometimes predictive model comparison can be useful



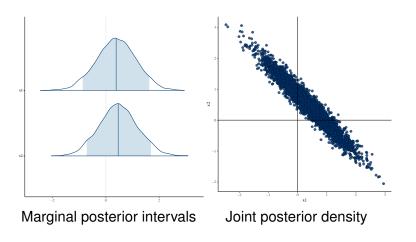
Marginal posterior intervals

## Sometimes predictive model comparison can be useful



rstanarm + bayesplot

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see also Collinear demo

- Continuous expansion including all models?
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# Model averaging

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### Model averaging

- Prefer continuous model expansion
- If needed integrate over the model space = model averaging
- Bayesian stacking may work better than BMA in case of misspecified models or small data
  - Yao, Vehtari, Simpson, & Gelman (2018)

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- Overfitting in selection process is not unique for cross-validation

### Selection induced bias and overfitting

- Selection induced bias in cross-validation
  - same data is used to assess the performance and make the selection
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  - the CV estimate for the selected model is biased
  - recognized already, e.g., by Stone (1974)

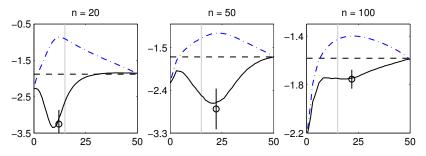
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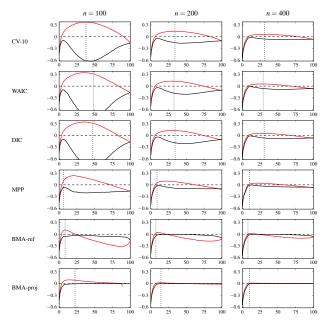
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- Bigger problem if there is a large number of models as in covariate selection

#### Selection induced bias in variable selection



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- It's good to think predictions of observables, because observables are the only ones we can observe
- Cross-validation can simulate predicting and observing new data
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