

HW #1 Solns., EE556, Fa

$$\#1) \quad g(\underline{x}) = \ln\left(\frac{f(\underline{x}/w_1)}{f(\underline{x}/w_2)}\right) = 0 \Rightarrow$$

$$-\frac{1}{2\sigma^2} \|\underline{x} - \underline{\mu}_1\|^2 + \frac{1}{2\sigma^2} \|\underline{x} - \underline{\mu}_2\|^2 = 0$$

or

$$\|\underline{x} - \underline{\mu}_1\|^2 = \|\underline{x} - \underline{\mu}_2\|^2$$

or

$$-2\underline{x}^T \underline{\mu}_1 + \|\underline{\mu}_1\|^2 + 2\underline{x}^T \underline{\mu}_2 - \|\underline{\mu}_2\|^2 = 0$$

or

$$\underline{x}^T (\underline{\mu}_2 - \underline{\mu}_1) + \frac{\|\underline{\mu}_1\|^2 - \|\underline{\mu}_2\|^2}{2} = 0$$

$$\underbrace{-\frac{(\underline{\mu}_2 + \underline{\mu}_1)^T (\underline{\mu}_2 - \underline{\mu}_1)}{2}}_{\text{}} = 0$$

or

$$\left(\underline{x} - \left(\frac{\underline{\mu}_1 + \underline{\mu}_2}{2} \right) \right)^T (\underline{\mu}_2 - \underline{\mu}_1) = 0$$

$$\#2) \quad Y = \left(\underline{x} - \left(\frac{\underline{\mu}_1 + \underline{\mu}_2}{2} \right) \right)^T (\underline{\mu}_2 - \underline{\mu}_1)$$

$Y \geq 0 \Rightarrow$ "decide w_2 "
 else "decide w_1 "

$$Y = \underline{x}^T (\underline{\mu}_2 - \underline{\mu}_1) - \left(\frac{\|\underline{\mu}_2\|^2 - \|\underline{\mu}_1\|^2}{2} \right)$$

Q: How is Y distributed?

A: consider Y given \underline{x} is from class w_2 :

$$f_{Y/w_2}(Y/w_2) \sim N(\mu_{Y/2}, \sigma_{Y/2}^2)$$

$$\mu_{Y/2} = E[Y/w_2] = E\left[\underline{x}^T (\underline{\mu}_2 - \underline{\mu}_1) / w_2\right] - \left(\frac{\|\underline{\mu}_2\|^2 - \|\underline{\mu}_1\|^2}{2} \right)$$

$$= \underline{\mu}_2^T (\underline{\mu}_2 - \underline{\mu}_1) - \frac{(\underline{\mu}_2 + \underline{\mu}_1)^T (\underline{\mu}_2 - \underline{\mu}_1)}{2}$$

$$= \frac{1}{2} (\underline{\mu}_2 - \underline{\mu}_1)^T (\underline{\mu}_2 - \underline{\mu}_1) = \frac{1}{2} \|\underline{\mu}_2 - \underline{\mu}_1\|^2$$

Also,

$$E\left[(Y - \mu_{Y/2})^2 / w_2\right] =$$

$$E\left[\left(\underline{x}^T (\underline{\mu}_2 - \underline{\mu}_1) - (\underline{\mu}_2 - \underline{\mu}_1)^T \frac{(\underline{\mu}_1 + \underline{\mu}_2)}{2} - \underbrace{\frac{(\underline{\mu}_2 - \underline{\mu}_1)^T (\underline{\mu}_2 - \underline{\mu}_1)}{2}}_{\mu_{Y/2}}\right)^2 / w_2\right]$$

$$= E\left[\left(\underline{x}^T (\underline{\mu}_2 - \underline{\mu}_1) - \frac{(\underline{\mu}_2 - \underline{\mu}_1)^T (2\underline{\mu}_2)}{2}\right)^2 / w_2\right]$$

$$= E\left[\left((\underline{\mu}_2 - \underline{\mu}_1)^T (\underline{x} - \underline{\mu}_2)\right)^2 / w_2\right]$$

$$= (\underline{\mu}_2 - \underline{\mu}_1)^T E\left[(\underline{x} - \underline{\mu}_2)(\underline{x} - \underline{\mu}_2)^T / w_2\right] (\underline{\mu}_2 - \underline{\mu}_1)$$

$$= \|\underline{\mu}_2 - \underline{\mu}_1\|^2 \cdot \sigma^2$$

S₀:

$$\begin{aligned}\text{Prob}[\text{error}] &= \frac{1}{2} \cdot \frac{1}{\sqrt{2\pi} \sigma_{Y/2}} \int_{-\infty}^0 e^{-\frac{(y - \mu_{Y/2})^2}{2\sigma_{Y/2}^2}} dy \\ &\quad + \frac{1}{2} \cdot \frac{1}{\sqrt{2\pi} \sigma_{Y/2}} \int_0^{\infty} e^{-\frac{(y - \mu_{Y/2})^2}{2\sigma_{Y/2}^2}} dy \\ &\stackrel{\text{by symmetry}}{=} \frac{1}{\sqrt{2\pi} \sigma_{Y/2}} \int_{-\infty}^0 e^{-\frac{(y - \mu_{Y/2})^2}{2\sigma_{Y/2}^2}} dy\end{aligned}$$

$$\text{Let } r = \frac{y - \mu_{Y/2}}{\sigma_{Y/2}} \Rightarrow \sigma_{Y/2} dr = dy$$

Can then rewrite:

$$\begin{aligned}\text{Prob}[\text{error}] &= \frac{1}{\sqrt{2\pi} \sigma_{Y/2}} \int_{-\infty}^{-\mu_{Y/2}/\sigma_{Y/2}} e^{-\frac{r^2}{2}} \sigma_{Y/2} dr \\ &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\|\underline{\mu}_2 - \underline{\mu}_1\|/2\sigma} e^{-r^2/2} dr \\ &= \frac{1}{\sqrt{2\pi}} \int_{\|\underline{\mu}_2 - \underline{\mu}_1\|/2\sigma}^{\infty} e^{-r^2/2} dr\end{aligned}$$

can calculate via
standard tables for
normal c.d.f.

$$\#3) \quad g_i(\underline{x}) = \ln(f(\underline{x} | \underline{w}_i) P[\underline{w}_i])$$

$$= -\frac{1}{2} (\underline{x} - \underline{\mu}_i)^T \Sigma^{-1} (\underline{x} - \underline{\mu}_i) + \left\{ \begin{array}{l} \text{nonessential} \\ \text{terms (in} \\ \text{this case)} \end{array} \right\}$$

$$g_1(\underline{x}) - g_2(\underline{x}) =$$

$$\begin{aligned} & \left(-\frac{1}{2} \underline{x}^T \Sigma^{-1} \underline{x} + \frac{1}{2} \underline{\mu}_1^T \Sigma^{-1} \underline{x} + \frac{1}{2} \underline{x}^T \Sigma^{-1} \underline{\mu}_1 - \frac{1}{2} \underline{\mu}_1^T \Sigma^{-1} \underline{\mu}_1 \right) \\ & - \left(-\frac{1}{2} \underline{x}^T \Sigma^{-1} \underline{x} + \frac{1}{2} \underline{\mu}_2^T \Sigma^{-1} \underline{x} + \frac{1}{2} \underline{x}^T \Sigma^{-1} \underline{\mu}_2 - \frac{1}{2} \underline{\mu}_2^T \Sigma^{-1} \underline{\mu}_2 \right) \end{aligned}$$

Cancel \nearrow same \nearrow

$$\underline{\mu}_1^T \Sigma^{-1} \underline{x} = \underline{x}^T \Sigma^{-1} \underline{\mu}_1 \quad \text{-- why?}$$

$$\begin{aligned} A: \quad (\underline{\mu}_1^T \Sigma^{-1} \underline{x})^T &= (\Sigma^{-1} \underline{x})^T (\underline{\mu}_1^T)^T \\ &= \underline{x}^T (\underbrace{\Sigma^{-1}}_{\text{symmetric}})^T \underline{\mu}_1 = \underline{x}^T \Sigma^{-1} \underline{\mu}_1 \end{aligned}$$

\Rightarrow

$$g_1(\underline{x}) - g_2(\underline{x}) = \underline{x}^T \Sigma^{-1} \underline{\mu}_1 - \frac{1}{2} \underline{\mu}_1^T \Sigma^{-1} \underline{\mu}_1 - \underline{x}^T \Sigma^{-1} \underline{\mu}_2 + \frac{1}{2} \underline{\mu}_2^T \Sigma^{-1} \underline{\mu}_2$$

$$= \underline{x}^T (\Sigma^{-1} \underline{\mu}_1 - \Sigma^{-1} \underline{\mu}_2) - \frac{1}{2} \underline{\mu}_1^T \Sigma^{-1} \underline{\mu}_1 + \frac{1}{2} \underline{\mu}_2^T \Sigma^{-1} \underline{\mu}_2$$

$$\underline{x}^T \Sigma^{-1} (\underline{\mu}_1 - \underline{\mu}_2)$$

$$\frac{1}{2} (\underline{\mu}_2 - \underline{\mu}_1)^T \Sigma^{-1} (\underline{\mu}_1 + \underline{\mu}_2)$$

$$= (\underline{\mu}_1 - \underline{\mu}_2)^T \Sigma^{-1} \left(\underline{x} - \frac{(\underline{\mu}_1 + \underline{\mu}_2)}{2} \right) = 0$$

$$\#4) \quad P(x/w_1) \sim N\left(\begin{pmatrix} 2 \\ 0 \end{pmatrix}, \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}\right)$$

$$P(x/w_3) \sim N\left(\begin{pmatrix} 0 \\ 2 \end{pmatrix}, \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}\right)$$

$$P(x/w_2) \sim N\left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 1 & 0 \\ 0 & 2 \end{pmatrix}\right)$$

$$P(w_1) = P(w_2) = P(w_3) = 1/3$$

Let's find boundaries between all pairs of classes:

Between classes 1 + 3, we know (from class) that the decision boundary is the line:

$$(\mu_1, -\mu_2)^T (\underline{x} - \underline{x}_0) = 0.$$

We can choose $\underline{x}_0 = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$ in this case \Rightarrow

$$B_{13} = \left\{ \underline{x} = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix} \text{ s.t. } (1 \ -1) \begin{pmatrix} x_1 \\ x_2 \end{pmatrix} = 0 \right\}$$

The boundary between classes 1 + 2 is given by:

$$-\frac{1}{2} (\underline{x} - \underline{\mu}_1)^T \Sigma^{-1} (\underline{x} - \underline{\mu}_1) - \frac{1}{2} \log |\Sigma_1| =$$

$$-\frac{1}{2} (\underline{x} - \underline{\mu}_2)^T \Sigma^{-1} (\underline{x} - \underline{\mu}_2) - \frac{1}{2} \log |\Sigma_2|$$

or

$$(x_1 - 2)^2 + x_2^2 + \frac{1}{2} \log 1 = x_1^2 + \frac{1}{2} x_2^2 + \log 2$$

or

$$\frac{1}{2} x_2^2 - 4x_1 + 4 - \log 2 = 0$$

$$B_{12} = \left\{ \underline{x} = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix} \text{ s.t. } \frac{1}{8} x_2^2 = (x_1 - 1 + \frac{\log 2}{4}) \right\}$$

The boundary between classes 2 and 3 is given by:

$$x_1^2 + (x_2 - 2)^2 = x_1^2 + \frac{1}{2}x_2^2 + \log 2$$

$$\text{or} \quad \frac{1}{8}x_2^2 = x_2 - 1 + \frac{\log 2}{4}$$

$$B_{23} = \{x : x_2 = 0.96 \text{ or } x_2 = 7.04\}$$

Rough sketch :

$$A = (.94, .94)$$

$$B = (7.06, 7.06)$$

