



筑波大学

*University of Tsukuba*

# Single-Factor Experiments: Analysis of Variance (ANOVA)


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# Before we Begin

- Go to the github repo:
  - <https://github.com/Mollinetti/Statistics-R>
- Download the script for this class! (in the 'scripts' folder, class\_4.r!)
- Run the first lines to load/install the required libraries

- 
- Introduction
  - What is ANOVA
  - Completely Randomized single-factor model
  - Single-factor Unbalanced Design
  - Single-factor Random Effects model
  - Single-factor Blocking design
  - Single-factor Repeated measures ANOVA
  - Single-factor Model Validation
  - Power test for Single-factor ANOVA



## ■ Introduction

- What is ANOVA
- Completely Randomized single-factor model
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# Introduction

- Many single-factor experiments **require more than two levels** of the factor to be considered
- Randomization of the experiments is now taken into account
- In the medical field, effects of medicines/treatments are verified for **multiple samples of populations**
- We will now call each level a **treatment**



# Introduction

- Montgomery explains the steps for a experiment:
  1. **Conjecture**: the original hypothesis that motivates the experiment.
  2. **Experiment**: the test performed to investigate the conjecture.
  3. **Analysis**: the statistical analysis of the data from the experiment.
  4. **Conclusion**: what has been learned about the original conjecture from the experiment. Often the experiment will lead to a revised conjecture, and a new experiment, and so forth.



# Introduction

- Factor levels can be chosen in two ways:
  - Fixed-effects model
  - Random effects/ Components of variance model

# Introduction

- Factor levels can be chosen in two ways:
  - Fixed-effects model
    - Specifically choose the  $a$  levels
    - Conclusions cannot be extended to treatments that were not considered
  - Random effects/ Components of variance model




# Introduction

- Factor levels can be chosen in two ways:
  - Fixed-effects model
  - Random effects/ Components of variance model
    - Random sample from a **larger population** of treatments
    - Extend the conclusion to **all treatments**
    - Knowledge about the treatments investigated is **not important**



# Introduction

- Before we start ANOVA, **let's remember:**
  - **Dependent Variable**
    - Variable being **tested and measured** in a scientific experiment
  - **Independent Variable**
    - Variable that is **changed or controlled** in a scientific experiment

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# What is ANOVA

- ANalysis Of VAriance
- ANOVA partitions the variability into two parts
- Test the hypothesis based on a comparison of two independent estimates of the population variance/mean/covariance



# What is ANOVA

- In ANOVA we want to test the significance of **the effect of** one or more **dependent variables** against three or more groups of **independent variables**
- One-way-ANOVA, repeated ANOVA, mixed effects ANOVA, ANCOVA
- two-way-ANOVA, MANOVA, MANCOVA, Single cell design

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# Completely Randomized single-factor Model

- Consider the following experiment: “We need to verify the effect of some medicine to **anemic male patients**. We have 4 groups: **control, placebo and two variations of the medicine**. Nominal hemoglobin levels for male are between **13.5 to 17.5g/dl**. **Sample size** is 20. Age is disregarded

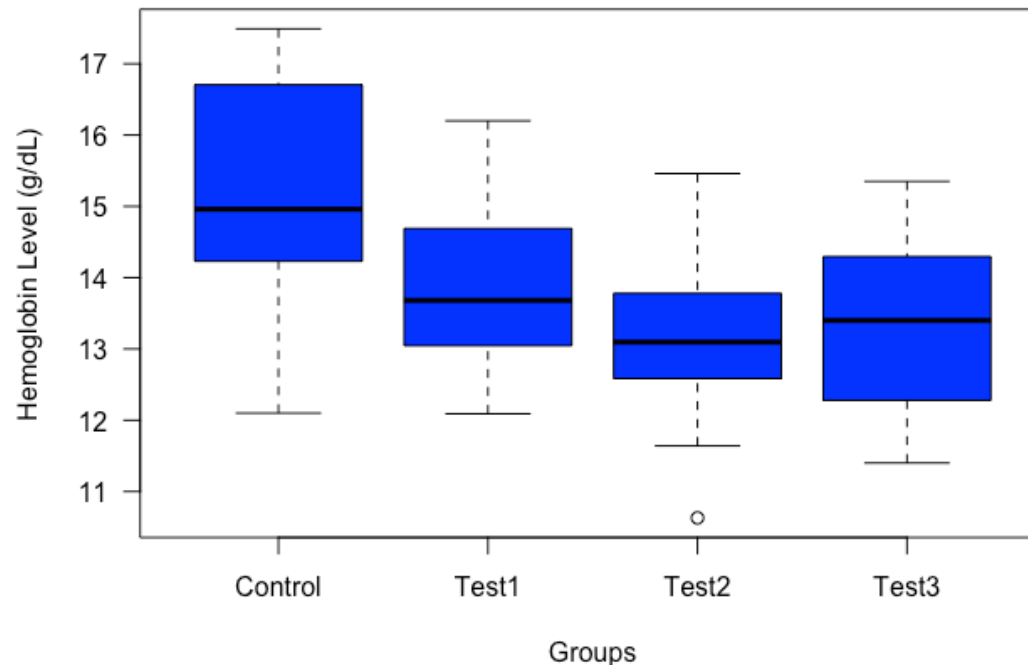
# Completely Randomized single-factor Model

- Blood samples are taken in a **completely randomized fashion**
- Need to reduce any **nuisance variable** in the experiment
- **Human error x Machine error**
- Graphical interpretation and statistic interpretation



# Completely Randomized single-factor Model

- Let's observe the boxplot of our data



# Completely Randomized single-factor Model

- Suppose we have  $a$  different levels of a single factor we wish to compare
- Observations follow the **linear model**:

$$Y_{ij} = \mu_i + \epsilon_{ij} \begin{cases} i = 1, 2, \dots, a \\ j = 1, 2, \dots, n \end{cases}$$

Where  $\mu_i = \mu + \tau_i$

# Completely Randomized single-factor Model

- Suppose we have  $a$  different levels of a single factor we wish to compare
- Observations follow the **linear model**:

Random observation  $\rightarrow$

$$Y_{ij} = \mu_i + \epsilon_{ij} \quad \begin{cases} i = 1, 2, \dots, a \\ j = 1, 2, \dots, n \end{cases}$$

Mean of the treatment  $\rightarrow \mu_i$

Error  $\rightarrow \epsilon_{ij}$

Where  $\mu_i = \mu + \tau_i$

Effect factor  $\rightarrow \tau_i$

Mean of the population (overall mean)  $\rightarrow \mu$

# Completely Randomized single-factor Model

- Each treatment can be thought of as a normal population with mean  $\mu_i$  and variance  $\sigma^2$
- Moreover, each treatment can be understood as a normally distributed variable with mean  $\mu$  plus a perturbation  $\tau_i$

# Completely Randomized single-factor Model

- For the fixed effect models, effects  $\tau_i$  are considered as deviations from overall mean  $\mu$
- So the sum of  $\tau_i$  is expected to be 0
- This is equivalent to testing the following hypotheses

$$H_0: \tau_1 = \tau_2 = \dots = \tau_a = 0$$

$$H_1: \tau_i \neq 0 \text{ for at least one } i$$

# Completely Randomized single-factor Model

■ We have the following point estimators:

□ Total sum of squares:  $SS_T = \sum_{i=1}^a \sum_{j=1}^n y_{ij}^2 - \frac{\bar{y}^2}{N}$

□ Sum square of treatments:  $SS_{Tr} = \sum_{i=1}^a \frac{y_i^2}{n} - \frac{\bar{y}^2}{N}$

□ Mean square of treatments:  $MS_{Tr} = \frac{SS_{Tr}}{(a-1)}$

□ Error sum of squares:  $SS_E = SS_T - SS_{Tr}$

□ Error mean square:  $MS_E = \frac{SS_E}{a(n-1)}$

□ F-statistic:  $F_0 = \frac{MS_{Tr}}{MS_E}$

# Completely Randomized single-factor Model

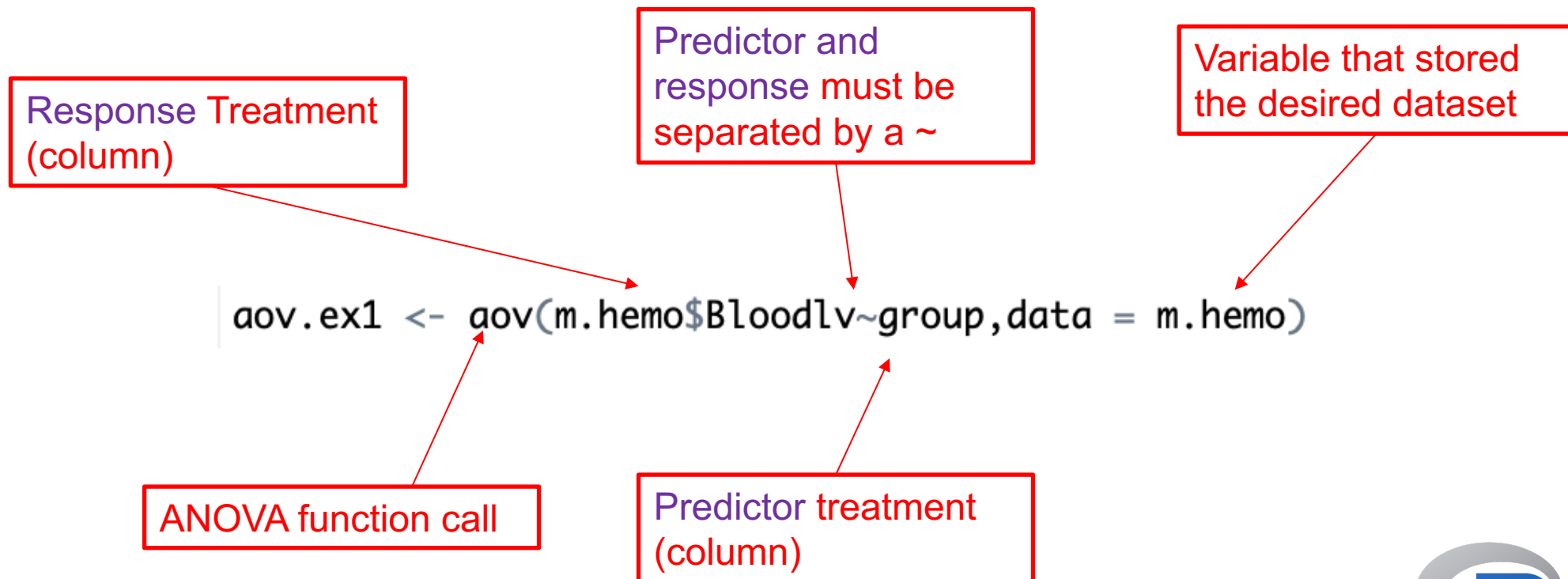
## ■ Importance of the F-statistic

■ F-statistic:  $F_0 = \frac{MS_{Tr}}{MS_E} =$  variance of the treatment means / mean of the treatment variances

- Supports the p-value in answering “Is the variance between the means of two population significantly different?”
- F statistic and Critical F-value
- If the calculated F statistic is greater than the F critical value, you can reject your null hypothesis

# Completely Randomized single-factor Model

## ■ Anatomy of the ANOVA function in R:





# Completely Randomized single-factor Model

- Now let's run the ANOVA for our hemoglobin experiment

```
> summary(aov.ex1)
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
m.hemo\$Group	3	50.8	16.935	10.97	4.57e-06 ***
Residuals	76	117.3	1.543		

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Since  $P$  is considerably smaller than 0.1 we can safely reject  $H_0$



# Completely Randomized single-factor Model

- Since the Mean Squared error  $MS_E$  is an estimator of variance  $\sigma$ , we can build confidence intervals
- Let's verify the confidence interval for each mean  $\mu_i$



# Completely Randomized single-factor Model

- Since we **rejected  $H_0$** , we know that there is **at least one factor** that is different from the others
- How do we know what factor?
  - Fisher's least significant Difference (LSD)
  - **Tukey's test**
  - **mmc**
- We'll run a **mean-mean multiple comparison (mmc)**

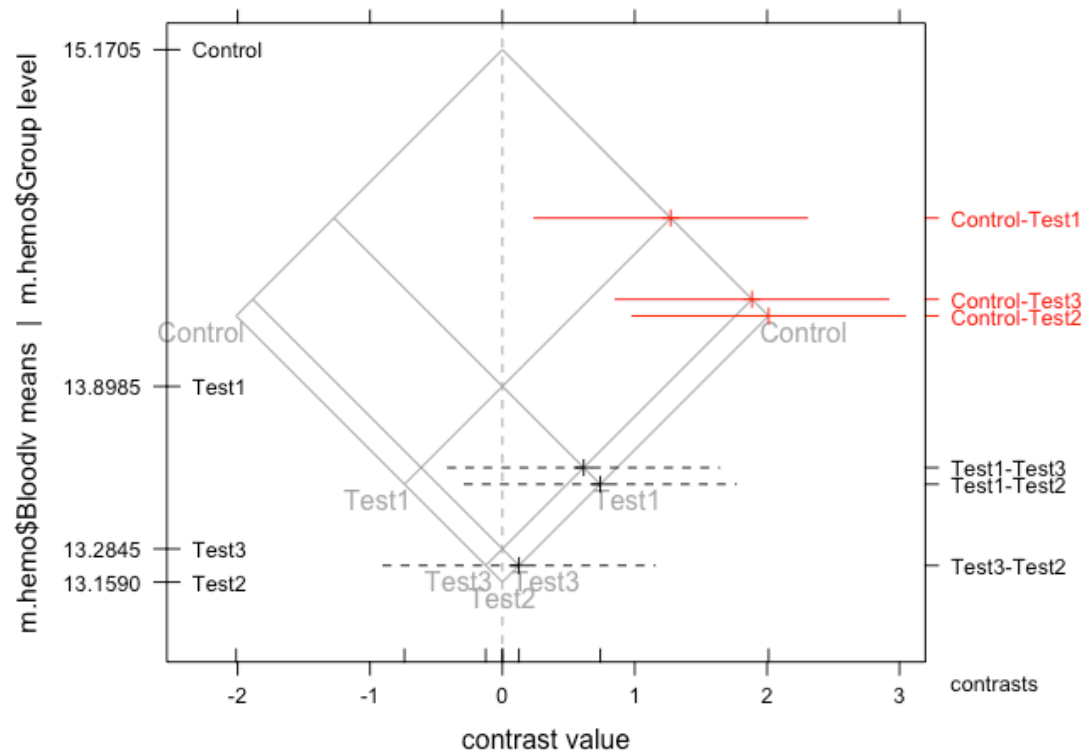


# Completely Randomized single-factor Model

- Tukey Honest Single Differences (HSD) test
  - Post-hoc analysis of the ANOVA
  - $H_0$ : no significance in the difference of means
  - $H_1$ : significance in the difference of means
- We will first plot the mmc, then plot the Tukey HSD to account for overlapping labels (tiebreaker plot)

# Completely Randomized single-factor Model

- Plotting the mmc we get a very straightforward answer:

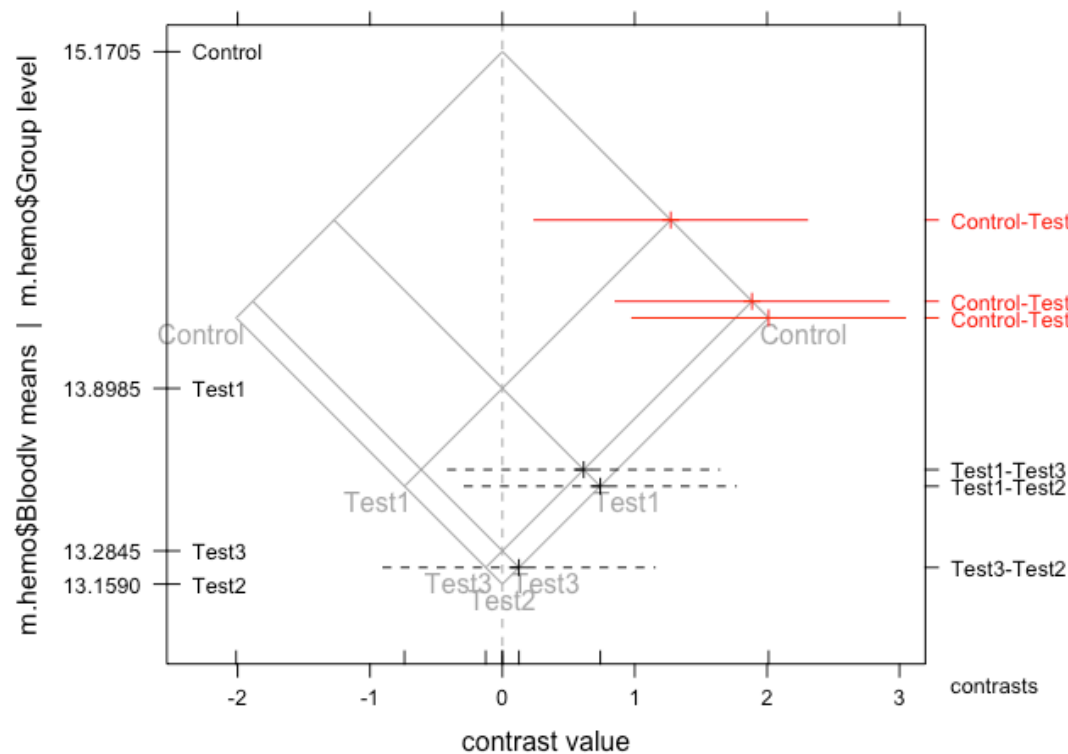


# Completely Randomized single-factor Model

## ■ How to read a mca plot:

Crossing of the means of a treatment at the levels of one factor with itself rotated 45 degrees

Mean values of each treatment

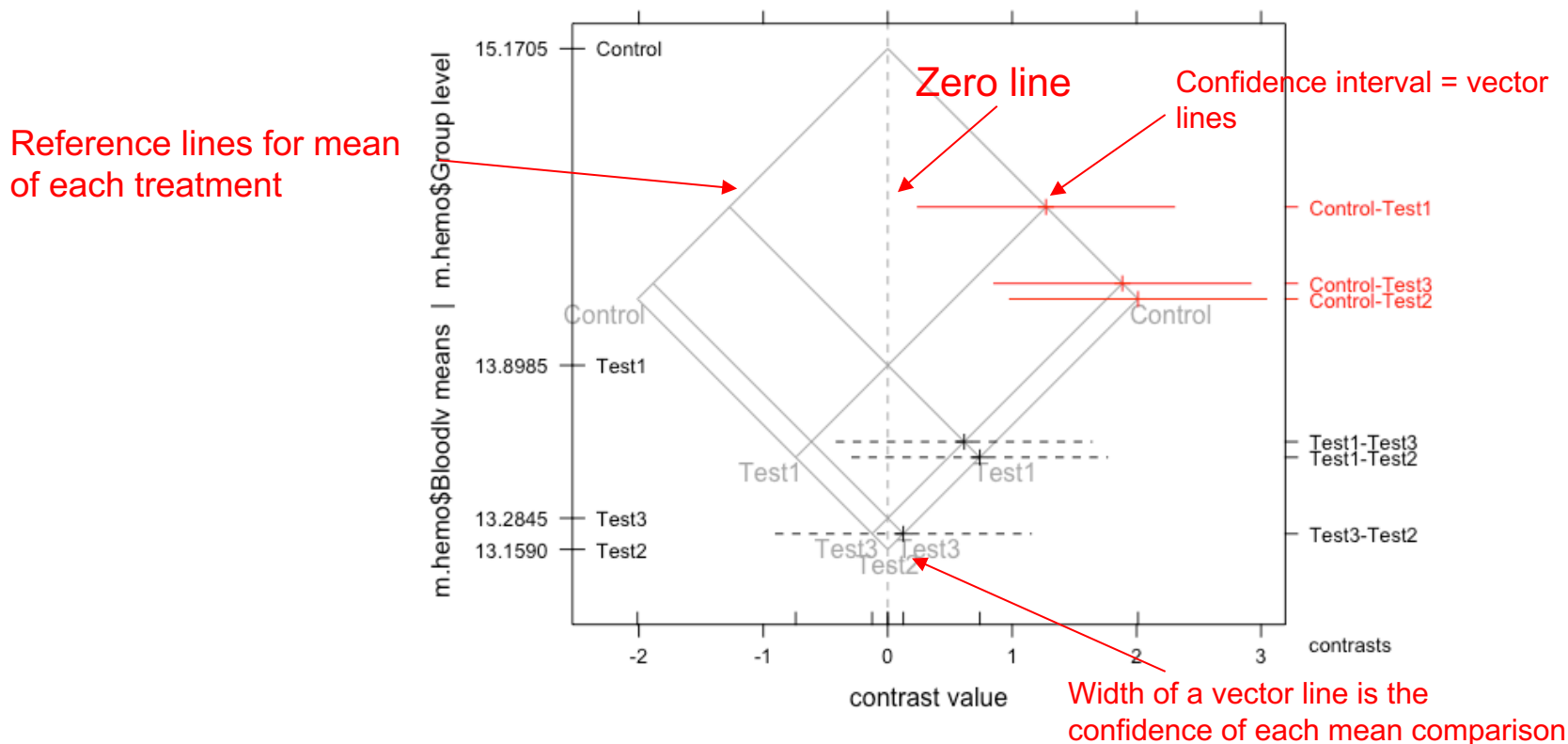


Weighted average of the means comprising each comparison

Difference in mean levels

# Completely Randomized single-factor Model

## ■ How to read a mca plot:



# Completely Randomized single-factor Model

- Contrasts are **weighted variables** in which their values **must sum to 0**
- **Example: comparing 4 means**

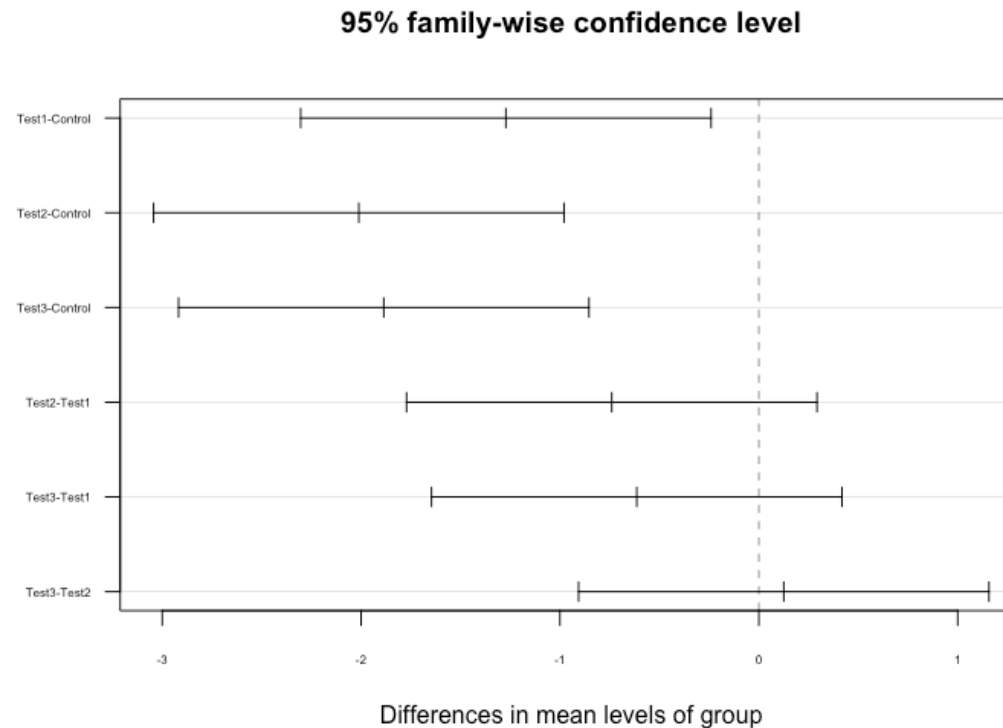
$\mu_1$	$\mu_2$	$\mu_3$	$\mu_4$
1	-1	0	0
0	0	1	-1
1	1	-1	-1

We call this an  
orthogonal contrast



# Completely Randomized single-factor Model

## ■ Plotting the Tiebreaker plot:




# Completely Randomized single-factor Model

## ■ Plotting the Tiebreaker plot:



Intervals that contain 0 indicate that the difference between the means **IS NOT SIGNIFICANT**

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# Unbalanced Design

- In some cases, the number of observations taken under each treatment may be different
- We say the the design is unbalanced
- Disadvantages over balanced design:
  - Insensitive to small departures from the assumption of equality
  - More prone to Type-II errors, less power  $\beta$

# Unbalanced Design

- Same steps as the one-way randomized ANOVA
  - Load the `assimetric_hemo_exp` dataset
  - Don't forget to drop the empty rows
  - Run the ANOVA
  - Run the mmc
  - Verify the differences



# Unbalanced Design

- We will use the `Anova()` function from the 'car' library, but we must wrap it with a linear regression function call

Linear regression  
function call


Same ~ as aov()

Type III sum-  
squares for  
unbalanced design

```
#using the Anova function from the car package (F-statistic is the same)  
aov.unbalanced <- Anova(lm(m.assim_hemo$Bloodlv~m.assim_hemo$Group,data = m.assim_hemo), type="III")
```

Anova function call



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# Random Effects Model

- Now, **treatments** are **random samples** from a population of treatments
- We want to draw **conclusions for the entire population of factor levels**
- *a* random factors are chosen



# Random Effects Model

- Consider variance of the treatment effects  $\tau_i$  to be  $\sigma_\tau^2$
- The variance of the response is  $\sigma_\tau^2 + \sigma^2$
- So we test Hypotheses about  $\sigma_\tau^2$

$$H_0: \sigma_\tau^2 = 0$$

$$H_1: \sigma_\tau^2 \neq 0$$

# Random Effects Model

- We first have to use the **restricted maximum likelihood (REML)** procedure and then apply ANOVA
- The computational procedure of the ANOVA table **is the same as of the fixed-effects\***
- Conclusions, however **apply for the entire population of treatments**
- **Load the 'Block\_Hemo\_exp.csv' dataset**

# Random Effects Model

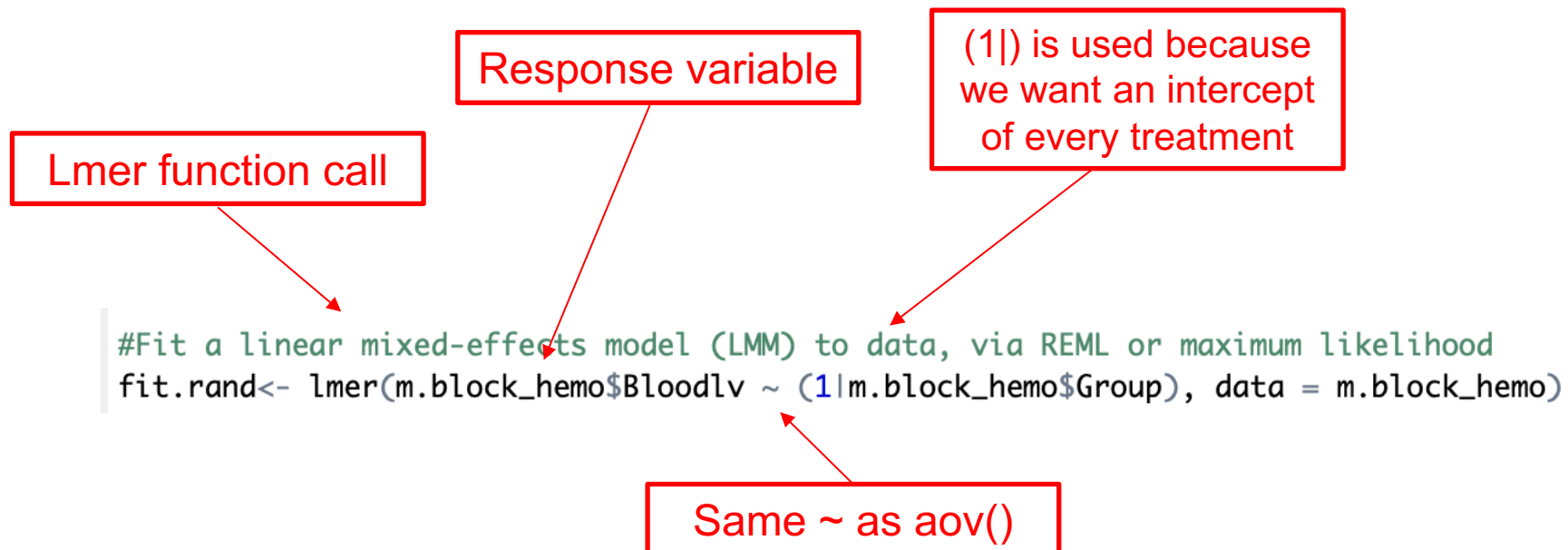
- In R, we have to use **an entirely different approach** from the regular **aov function call**.
- We call the function **lmer** from the library **lme4** (**linear mixed random effect**)

```
#Fit a linear mixed-effects model (LMM) to data, via REML or maximum likelihood  
fit.rand<- lmer(m.block_hemo$Bloodlv ~ (1|m.block_hemo$Group), data = m.block_hemo)
```



# Random Effects Model

- In R, we have to use **an entirely different approach** from the regular **aov function call**.
- We call the function **lmer** from the library **lme4** (**linear mixed random effect**)



# Random Effects Model

- Running the summary we can see the random effect associated to each treatment

```
> summary(fit.rand)
Linear mixed model fit by REML ['lmerMod']
Formula: m.block_hemo$Bloodlv ~ (1 | m.block_hemo$Group)
Data: m.block_hemo
```

REML criterion at convergence: 1465.4

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.90058	-0.71495	-0.03995	0.70689	3.04667

Random effects:

Groups	Name	Variance	Std.Dev.
m.block_hemo\$Group	(Intercept)	0.2512	0.5012
Residual		1.4504	1.2043

Number of obs: 450, groups: m.block\_hemo\$Group, 9

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	13.6314	0.1764	77.26

Random effects  
(sum them)

Estimate from the  
expected measurement  
of a donor from a  
random treatment  
(mean)




# Random Effects Model

- Now we check the random effect  $\sigma_{\tau}^2$  associated to each treatment

```
## "estimated" (better: conditional means of) random effects that we computed with the lmer  
ranef(fit.rand)
```

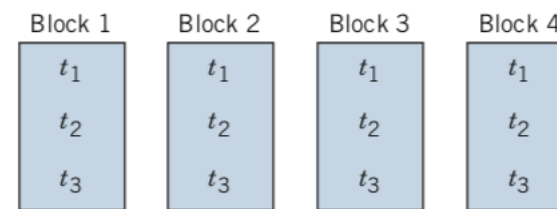
- To validate the assumptions of our model we do several plots
  - Residuals x fitted values (Tukey-Ascombe plot)
  - Qq-plot
  - Normalized qq plot



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# Blocking Design

- Reduce the variability from a nuisance factor
- Extension of the paired t-test when more than two treatments must be compared
- Selection of  $b$  blocks and running a complete replicate of the experiment in each block
- $A$  levels,  $b$  blocks



**Figure 13-8** A randomized complete block design.

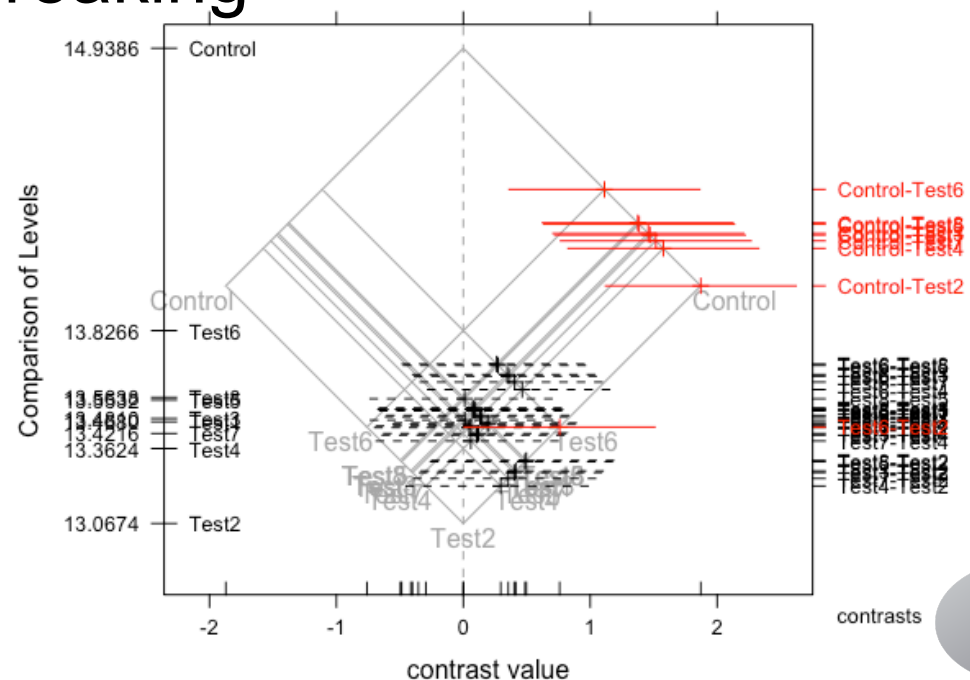



# Blocking Design

- When to use block design?
  - When you want to reduce the  $MS_E$
  - When doing a single factor experiment has much more degrees of freedom
- Generally, it is based on trial and error

# Blocking Design

- Run the ANOVA with block design at R
- Do the mmc
- Run the HSD for tie-breaking
- Check the plot




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# Repeated Measures ANOVA

- The **same participant** are observed multiple times or under all the levels of a **within-subject factor**
- **One independent variable**
  - ☐ Within-subject variable
  - ☐ Two or more categorical related group
- **One dependent variable**
  - ☐ Interval or ratio level of measurement

- 
- A within-subjects variable is an independent variable that is manipulated by testing each subject at each level of the variable
  - Example: suppose we take blood measurements on two situations:
    - After 12 hour fasting
    - Without fasting

# Repeated Measures ANOVA

- We test for the following hypotheses:

- $H_0$ : Means of measures are equal

- $H_1$ : Means of measures are not equal

$$H_0: \mu_1 = \mu_2 = \dots = \mu_a = 0$$

$$H_1: \mu_i \neq 0 \text{ for at least one } i$$

# Repeated Measures ANOVA

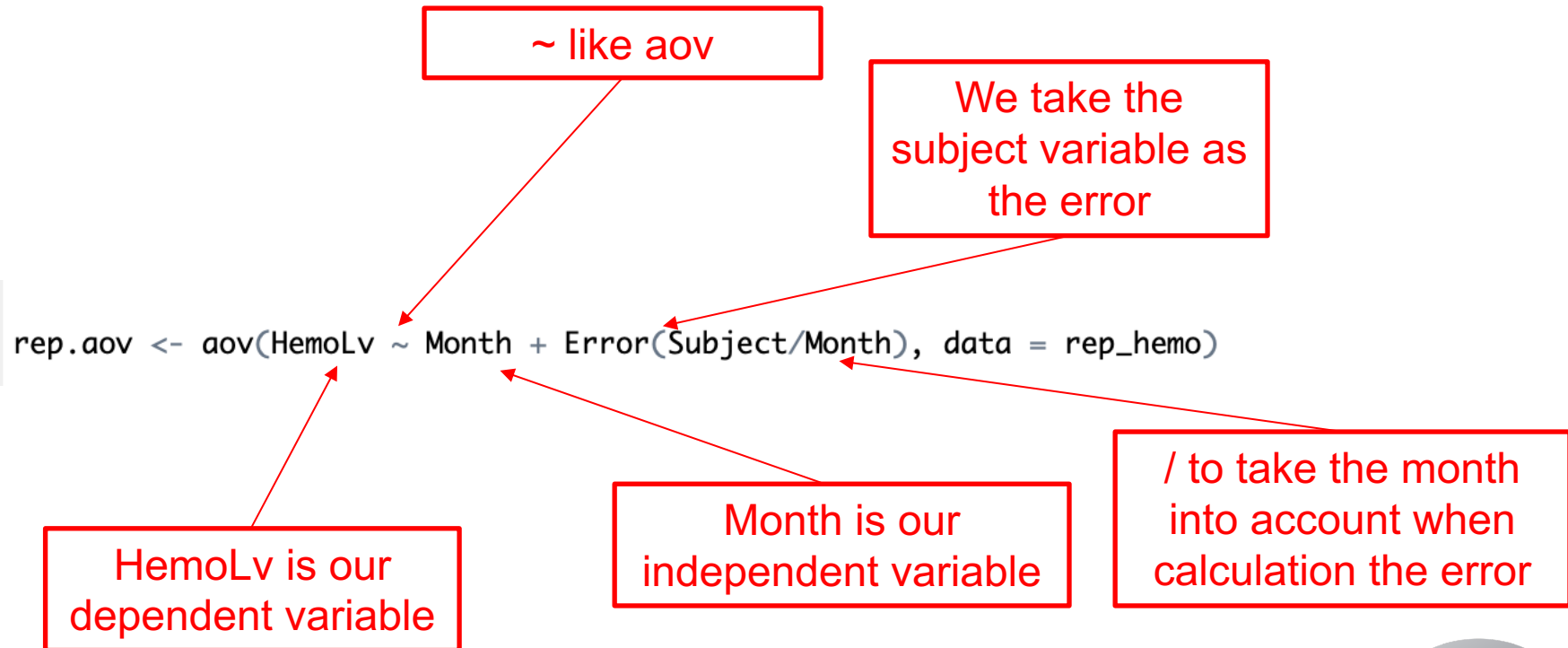
- Anova model for the repeated measures

```
rep.aov <- aov(HemoLv ~ Month + Error(Subject/Month), data = rep_hemo)
```



# Repeated Measures ANOVA

## ■ Anova model for the repeated measures





# Repeated Measures ANOVA

- For the assumption of the repeated measures for one-way ANOVA, we test:
- Normality
  - Shapiro-Wilk test
- Sphericity
  - Mauchly test (for 2 or more factors)

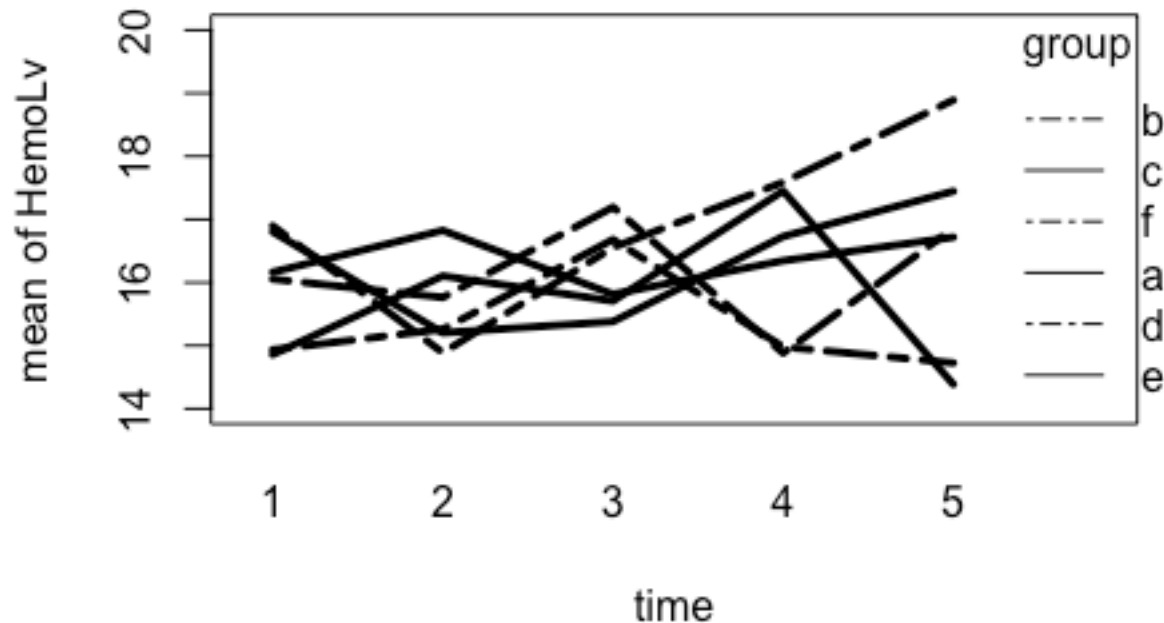
# Repeated Measures ANOVA


- Load the 'Rep\_hemo\_exp.csv' dataset
- Run the anova with a modified model
- Run the shapiro-wilk test
- Plot the means for interpretation



# Repeated Measures ANOVA

- Plot of the means



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# Model Validation

- Like the t-test, we must test our model for:
  - Normality
  - Independence
  - Heteroscedascity

# Model Validation

- Our validation varies according to which ANOVA we are dealing with:
  - One-way, unbalanced, blocking:
    - Normality: Normal Probability plot of residuals
    - Heteroscedascity: Plotting the residuals against time
    - Independence: Linear regression against the residuals
  - Random effects:
    - Normality: normal qq plots Tukey-Ascombe plot
    - Heteroscedascity: residuals against time
    - Independence: qqplot of residuals

# Model Validation

- Our validation varies according to which ANOVA we are dealing with:
  - Repeated measures:
    - Normality: Shapiro-Wilk test
    - Independence: Plot of residuals against time
    - Sphericity: Mauchly Test

# Model Validation

- R commands for each test:

- Normality: `plot()`

```
plot(aov.ex1)
```

- Heteroscedascity: `plot(residuals())`

```
plot(aov.ex1$res)
```

- Independence: `lm()`

```
lm.aov1<- lm(formula = abs(aov.ex1$res) ~ group)
```





# Model Validation

## ■ R commands for each test:

- Normality: **plot()**

`plot(aov.ex1)`

- Heteroscedascity: **plot(residuals())**

`plot(aov.ex1$res)`

Residuals column of  
the anova

- Independence: **lm()**


Don't forget the ~

Desired variable

Function call

`lm.aov1<- lm(formula = abs(aov.ex1$res) ~ group)`



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# Power test for ANOVA

- We will use the `power.anova.test` in R, for that we need:
  - Number of groups
  - Between group variance
  - Within group variance
  - Confidence level  $\alpha$
  - Sample size or desired power  $\beta$



# Power test for ANOVA

## ■ Using our randomized experiment example:

```
#calculate the within group variance
hemo_wvar <- anova(aov.ex1)["Residuals", "Mean Sq"]
#do the power test
p <- power.anova.test(groups = length(colMeans(hemo)), between.var = var(hemo_means),
  within.var = hemo_wvar, power=0.85, sig.level=0.05, n=NULL)
```

Between group  
variance

Desired power

Within group  
variance

```
n<- ceiling(p$n)
```

```
#do the power test with sample size
```

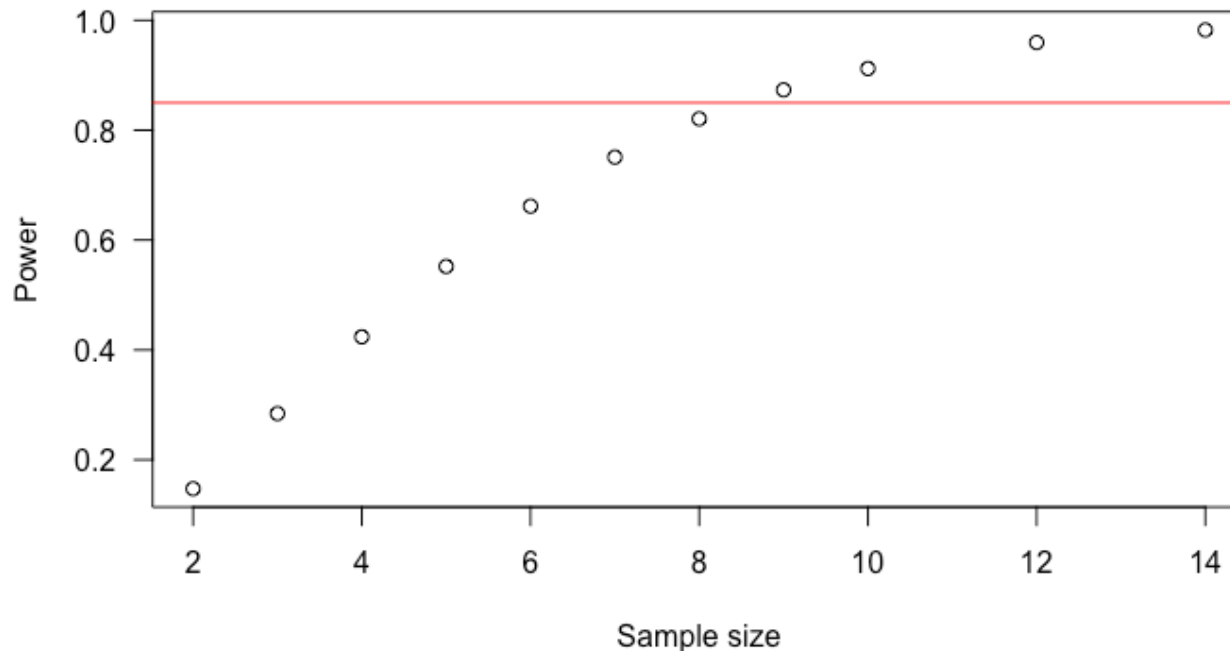
```
p2 <- power.anova.test(groups = length(colMeans(hemo)), between.var = var(hemo_means), within.var = hemo_wvar, sig.level=0.05, n=n)
```

Calculated sample  
size



# Power test for ANOVA

- Plotting different sample sizes to measure the increase in power



# Next Episode

- ANOVA is not over yet! We will talk about the general case of ANOVA (MANOVA, two-way)
- However, we must understand how to fit linear models (linear regression)
- Next stop, Linear Regression!
- And a brief explanation about single-factor ANCOVA



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# Single-Factor Experiments: Analysis of Variance (ANOVA)

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