

~~C.V~~ ~~variable :- 2009~~

- class-lectures

A variable is a quantity whose value varies from individual to individual.

Example : The height of the third year student because it varies student to student

~~C.V~~ Stochastic variable :- A stochastic variable is variable whose values is a number determined by outcomes, time and space of the experiment

Example : Let us consider an experiment of throwing a true die. Let x_n be the ^{sim first} number of outcome of the n th throw, $n \geq 1$. For a distinct value of $n = 1, 2, \dots$ we get distinct binomial variable. $x_n : \{x_n, n \geq 1\}$. Hence x_n is a stochastic variable.

Question :- What is stochastic process? Describe the classification of stochastic process and give an example in each case

A stochastic process is a process which has some random element involved in its structure.

Mathematically : a stochastic process is a process is a set of random variables $\{x_t\}$ or $\{x(t)\}$ depending on some real parameters like time and state.

In other words, families of random variable which are the functions of say time and state are known as stochastic process.

$$\begin{aligned} P_{x_1}(s) &= \\ P_{x_2}(t) &= \end{aligned}$$

Example: Consider a simple experiment like rolling a trinomial die. Let x_n be the number of ones of the n th throw, $n \geq 1$. For a distinct value of $n = 1, 2, \dots$, we get a distinct binomial variable x_n : $\{x_n, n \geq 1\}$ which gives a family of Random.

Question ~~20/10/09~~: Discuss the classification of stochastic processes:-

Answer: One dimensional stochastic process can be classified into four types of process which are as follows:-

(i) Discrete time and discrete state space: Let x_n be the total no. of sixes in the first n throwing a die. Then the set of possible values of x_n is the finite set of non-negative integers $0, 1, 2, \dots, n$. Here the state space x_n is discrete and time n is discrete.

(ii) Discrete time and continuous state space: Let x_n be random variable which represents the weight of n th student. Here n is discrete time and the state space is continuous.

(iii) continuous time and discrete state space: Suppose that $x(t)$ is the number of auto's in $(0, t]$. Here the $x(t)$ is discrete through the $x(t)$ is defined for a continuous range of time.

IV continuous time and continuous state space

Suppose that $X(t)$ represents the random variable at time t in a particular place. In $(0, t)$ the set of possible values of $X(t)$ is continuous. Now the time.

Q. Define Markov chain and Markov process. ? and state continuous

10 Answer:-

2010/12/08 Markov chain :- The stochastic process $\{X_n; n=0, 1, 2, \dots\}$ is called a Markov chain if for $j, k, J_1, J_2, \dots, J_{n-1} \in N$ which are the outcomes or results of the experiment.

$$\begin{aligned} & \Pr \{ X_n = k | X_{n-1} = j, X_{n-2} = J_1, \dots, X_0 = J_{n-1} \} \\ &= \Pr \{ X_n = k | X_{n-1} = j \} \\ &= P_{jk} \text{ (say)} \end{aligned}$$

Whenever the first member is defined.

⇒ Markov process :- (MC)

If $\{X(t), t \in T\}$ is a stochastic process such that given the value $X(s)$, the values of $X(t), t > s$ do not depend on the values of $X(u), u < s$ then the process is said to be MC.

The uses of MC are given below:

- i) MC is widely used in physics, Engineering, Marketing and many other related areas.
- ii) In business the MC is used powerfully ^{powerful} ^{fun}. It is used to analyse the current states of product in an effort to predict its future.
- iii) It is used to forecast the behaviour of customers from random point.

Transition :-

Let $\{X_n; n=0, 1, 2, \dots\}$ be a markov chain. If $X_{n-1} = j_n$ and $X_n = K$, then the system is called transition when the state j at $(n-1)$ -th trial go to the state K at n -th trial.

$X_{n-1} = j_n$ at
 $X_n = K$

Transition probability :-

The probabilities of transition from the state j at $(n-1)$ -th trial to state K at n -th trial is called transition probability. The transition probability is basic to the study of the structure of the markov chain.

$P_{jK} = P$

Transition probability matrix (+. p.m) :-

The transition probabilities P_{jk} satisfy $P_{jk} \geq 0$ and $\sum_k P_{jk} = 1 \quad \forall j$. These probabilities may be written as,

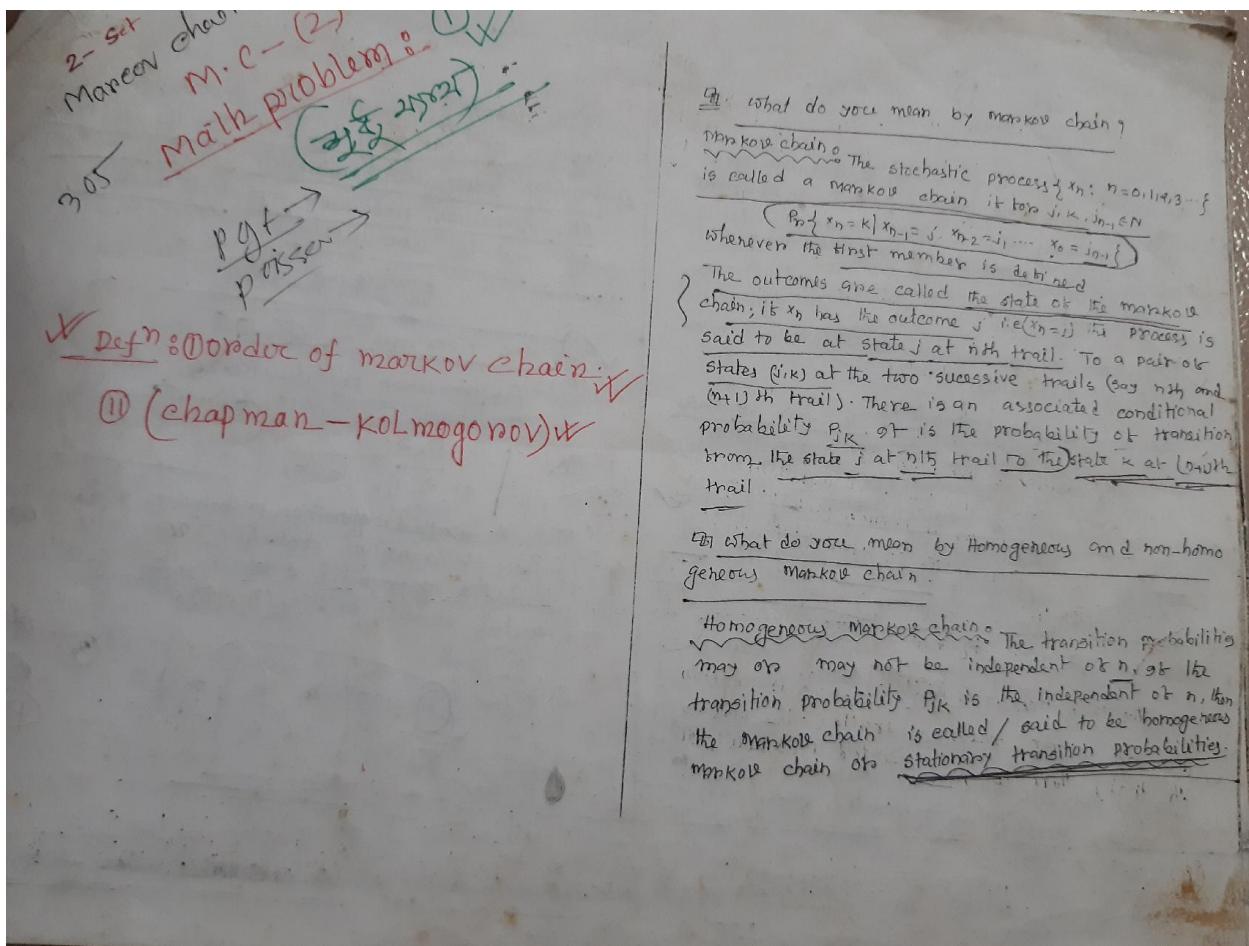
$$P = \begin{pmatrix} P_{11} & P_{12} & P_{13} & \cdots \\ P_{21} & P_{22} & P_{23} & \cdots \\ \vdots & \vdots & \vdots & \ddots \end{pmatrix}$$

which is called the t.p.m of the M.C.

Stochastic matrix :-

A square matrix with non negative elements and unit row sums is called stochastic matrix. Such that transition prob. matrix is stochastic matrix.

$$\text{i.e. } P = \begin{pmatrix} P_{11} & P_{12} & \cdots & P_{1n} \\ P_{21} & P_{22} & \cdots & P_{2n} \\ \vdots & \vdots & \ddots & \vdots \\ P_{n1} & P_{n2} & \cdots & P_{nn} \end{pmatrix}$$



(3)

Non-homogeneous markov chain: The transition prob. may or may not be independent of n . If the transition probability P_{jk} is dependent of n , then the markov chain is called non-homogeneous markov chain.

✓ Define unit-step transition probability and m-step transition probability.

Unit-step transition probability: To a pair of states (j, k) at the two successive trials (n th and $(n+1)$ th trials) there is an associated probability P_{jk} , it is the probability of transition from the state j at n th trial onto the state k at $(n+1)$ th trial. Then the corresponding transition probability $P_{jk}^{(1)}$ is called unit-step transition probability and is denoted by $P_{jk}^{(1)}$ and is defined as

$$P_{jk}^{(1)} = P\{x_{n+1} = k \mid x_n = j\}.$$

m-step transition probability: To a pair of states (j, k) at the two successive trials (n th and $(n+m)$ th trial) there is an associated probability $P_{jk}^{(m)}$

is the prob. of transition from the state j at n th trial to the state k at $(n+m)$ th trial. Then the corresponding transition probability is called m-step transition probability and it is denoted by $P_{jk}^{(m)}$ and is defined as:-

$$P_{jk}^{(m)} = P\{x_{n+m} = k \mid x_n = j\}.$$

✓ Define t.p.m with an example.

Ans: The transition probabilities P_{jk} satisfies the two conditions

$$\begin{aligned} \text{(i)} \quad & P_{jk} \geq 0 \\ \text{(ii)} \quad & \sum_{k=0}^{\infty} P_{jk} = 1 \quad \forall j \end{aligned}$$

This is called transition probability matrix or matrix of transition probability. Let P is a stochastic matrix i.e. a square matrix with non-negative elements and unit row sums.

Example:

$$P = \begin{pmatrix} 0 & 1 & 0 \\ 1/2 & 1/2 & 0 \\ 1 & 0 & 0 \end{pmatrix}$$

$$\begin{array}{l} \text{(i) } P \\ \text{(ii) } P^2 \\ \text{(iii) } P^3 \end{array}$$

To write the probabilities of transition probability matrix.

Ans: The properties of transition probability matrix are given below:-

- (i) It is square matrix.
- (ii) It has non-negative elements.
- (iii) It's row sum unit.

X Define unit step transition matrix and m step transition matrix.

so,

unit step transition matrix. let x_n be a unit step markov chain assuming values ($n=0, 1, 2, \dots, m-1$). Then its transition probability matrix P is

given by

$$P = \begin{pmatrix} P_{00} & P_{01} & P_{02} & \cdots & P_{0,m-1} \\ P_{10} & P_{11} & P_{12} & \cdots & P_{1,m-1} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ P_{m-1,0} & P_{m-1,1} & \cdots & \cdots & P_{m-1,m-1} \end{pmatrix}$$

where,

$$P_{jk}^{(1)} = \sum_{x_{n+1}=k | x_n=j} ; i, k = 0, 1, 2, \dots, m-1$$

which is called unit step transition matrix.

m step transition matrix.

m step transition matrix. let x_n be a m step markov chain assuming values ($n=0, 1, 2, \dots, m-1$) and let its transition probty matrix P is given by :-

$$P^{(m)} = \begin{pmatrix} P_{00} & P_{01} & \cdots & P_{0,m-1} \\ P_{10} & P_{11} & \cdots & P_{1,m-1} \\ \vdots & \vdots & \ddots & \vdots \\ P_{m-1,0} & P_{m-1,1} & \cdots & P_{m-1,m-1} \end{pmatrix}$$

$$P_{jk}^{(m)} = \sum_{x_{n+m}=k | x_n=j} ; i, k = 0, 1, 2, \dots, m-1$$

which is called m-step transition matrix.

q) Define double stochastic matrix.

Ans: A non-negative square matrix is said to be double stochastic matrix if all the rows and columns sum are unity.

Example: $P = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix}$

$$\left\{ \begin{array}{l} 0+1=1 \\ 1+0=1 \end{array} \right. \quad \left\{ \begin{array}{l} 0+0=0 \\ 1+1=2 \end{array} \right. \quad \left\{ \begin{array}{l} 0+0=0 \\ 1+1=2 \end{array} \right. =$$

The above square is said to be double stochastic matrix since its row and column sums are unity.

Tutorial

XQ How do you calculate proby dist' with the help of transition probability.

Ans: The proby dist' of x_0, x_1, \dots, x_n can be computed in terms of the transition probability matrix P_{jk} and the initial dist' of x_0 . Suppose for simplicity $x_0 = 0$, then we get

$$\Pr\{x_0=0, x_1=1, \dots, x_{n-2}=1, x_{n-1}=j, x_n=k\}$$

$$= \Pr\{x_n=k | x_{n-1}=j, x_{n-2}=1, \dots, x_0=0\}$$

$$\Pr\{x_{n-1}=j, \dots, x_0=0\}$$

$$= \Pr\{x_n=k | x_{n-1}=j\} \cdot \Pr\{x_{n-1}=j | x_{n-2}=1, \dots, x_0=0\}$$

$$\Pr\{x_{n-2}=1, x_{n-3}=1, \dots, x_0=0\}$$

$$= \Pr\{x_n=k | x_{n-1}=j\} \cdot \Pr\{x_{n-1}=j | x_{n-2}=1\} \cdot \Pr\{x_{n-2}=1 | x_{n-3}=1, \dots, x_0=0\}$$

$$= \Pr\{x_n=k | x_{n-1}=j\} \cdot \Pr\{x_{n-1}=j | x_{n-2}=1\} \cdot \Pr\{x_{n-2}=1 | x_{n-3}=1, \dots, x_0=0\}$$

$$= \Pr\{x_n=k | x_{n-1}=j, x_{n-2}=1, \dots, x_0=0\}$$

$$= \Pr\{x_0=0\} \cdot b_{0j} \cdot P_{jk}$$

Thus

$$\Pr\{x_0=0, x_1=1, \dots, x_{n-2}=1, x_{n-1}=j, x_n=k\}$$

$$= \Pr\{x_0=0\} \cdot b_{0j} \cdot P_{jk} \quad (\text{show})$$

Ex: Let $\{x_n, n \geq 0\}$ be a Markov chain with three states 0, 1 and 2 with transition matrix

$$P = \begin{pmatrix} 0 & 1 & 2 \\ 3/4 & 1/4 & 0 \\ 1/4 & 1/2 & 1/4 \\ 0 & 3/4 & 1/4 \end{pmatrix}. \quad \text{ans}$$

and initial dist' $\Pr\{x_0=i\} = r_i : i=0, 1, 2$. Find

$$(1) P_{12}^{(2)}, (2) P_{13}^{(2)}, (3) P_{32}^{(2)}, (4) P_{23}, (5) P_{21}^{(1)}$$

$$(6) P_{12}^{(1)}, (7) P_{13}^{(1)}, (8) \Pr\{x_2=2, x_1=1, x_0=2\}$$

$$(9) \Pr\{x_3=1, x_2=2, x_1=1, x_0=2\}$$

Solution: we have

$$P = \begin{pmatrix} 0 & 1 & 2 \\ 3/4 & 1/4 & 0 \\ 1/4 & 3/4 & 0 \\ 0 & 0 & 1/4 \end{pmatrix}$$

$$\begin{aligned} P^2 &= \begin{pmatrix} 3/4 & 1/4 & 0 \\ 1/4 & 3/4 & 0 \\ 0 & 0 & 1/4 \end{pmatrix} \begin{pmatrix} 3/4 & 1/4 & 0 \\ 1/4 & 3/4 & 0 \\ 0 & 0 & 1/4 \end{pmatrix} \\ &= \begin{pmatrix} 0 & 1 & 2 \\ 15/16 & 5/16 & 1/16 \\ -5/16 & 8/16 & 2/16 \\ 3/16 & 9/16 & 4/16 \end{pmatrix} \end{aligned}$$

$$(I) \quad \underline{P_{13}^{(2)}} = \Pr \{ \underline{x_2=1} | \underline{x_1=1}, \underline{x_0=2} \} = \frac{8}{16}$$

$$(II) \quad \underline{P_{12}^{(2)}} = \Pr \{ \underline{x_3=2} | \underline{x_1=1}, \underline{x_0=1} \} = \frac{3}{16}$$

$$(III) \quad \underline{P_{140.}^{(2)}} = \Pr \{ \underline{x_4=0} | \underline{x_2=1} \} = \frac{3}{16}$$

$$(IV) \quad \underline{P_{21}^{(1)}} = \Pr \{ \underline{x_1=0} | \underline{x_0=2} \} = \frac{3}{4}$$

$$(V) \quad \underline{P_{12}^{(1)}} = \Pr \{ \underline{x_2=2} | \underline{x_1=1} \} = \frac{1}{4}$$

$$(VI) \quad \Pr \{ \underline{x_2=2}, \underline{x_1=1}, \underline{x_0=2} \}$$

$$= \Pr \{ \underline{x_2=2} | \underline{x_1=1} \} \Pr \{ \underline{x_1=1} | \underline{x_0=2} \} \Pr \{ \underline{x_0=2} \}$$

$$= P_{12}^{(1)} \times P_{21}^{(1)} \times \Pr \{ \underline{x_0=2} \}$$

$$= \frac{3}{4} \times \frac{1}{4} \times \frac{1}{2}$$

$$= \frac{3}{32}$$

$$(VII) \quad \Pr \{ \underline{x_3=1}, \underline{x_2=2}, \underline{x_1=1}, \underline{x_0=2} \}$$

$$= \Pr \{ \underline{x_3=1} | \underline{x_2=2} \} \Pr \{ \underline{x_2=2} | \underline{x_1=1} \} \Pr \{ \underline{x_1=1} | \underline{x_0=2} \}$$

$$= P_{21}^{(1)} \times P_{12}^{(1)} \times P_{11}^{(1)} \times \Pr \{ \underline{x_0=2} \}$$

$$= \frac{3}{4} \times \frac{1}{4} \times \frac{3}{4} \times \frac{1}{2}$$

$$= \frac{3}{64}$$

~~Prob~~ Problem: Given the transition matrix

$$P = \begin{pmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} & 0 \\ 0 & 0 & 1 \end{pmatrix} \text{ with states } 1, 2, 3$$

and initial distn $P_{\text{ini}} \{x_0 = i\} : i = 1, 2, 3$

- Find (I) $P_{13}^{(0)}$ (II) $P_{13}^{(2)}$ (III) $P_{13}^{(3)}$ (IV) $P_{23}^{(0)}$ (V) $P_{23}^{(2)}$ (VI) $P_{23}^{(3)}$ (VII) $P_{33}^{(0)}$ (VIII) $P_{33}^{(2)}$ (IX) $P_{33}^{(3)}$.

Solution: We have,

$$P = \begin{pmatrix} 1 & 2 & 3 \\ 0 & \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} & 0 \\ 0 & 0 & 1 \end{pmatrix}$$

$$P^T = \begin{pmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} & 0 \\ 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} & 0 \\ 0 & 0 & 1 \end{pmatrix}$$

$$= \begin{pmatrix} \frac{1}{4} & \frac{3}{4} & 0 \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{2} & \frac{1}{2} & 0 \end{pmatrix}$$

$$\begin{aligned} P^3 &= P \cdot P = \begin{pmatrix} \frac{1}{4} & \frac{3}{4} & 0 \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \\ \frac{1}{2} & \frac{1}{2} & 0 \end{pmatrix} \begin{pmatrix} 0 & \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} & 0 \\ 0 & 0 & 1 \end{pmatrix} \\ &= \begin{pmatrix} \frac{1}{8} & \frac{1}{8} & \frac{1}{8} \\ \frac{1}{8} & \frac{1}{8} & \frac{1}{8} \\ \frac{1}{4} & \frac{1}{4} & \frac{1}{4} \end{pmatrix} \end{aligned}$$

$$(I) P_{13}^{(0)} = P_{\text{ini}} \{x_2 = 3 | x_1 = 1\} = \frac{1}{2}$$

$$(II) P_{13}^{(2)} = P_{\text{ini}} \{x_2 = 3 | x_1 = 1\} = 0$$

$$(III) P_{13}^{(3)} = P_{\text{ini}} \{x_2 = 3 | x_1 = 1\} = \frac{1}{8}$$

$$(IV) P_{23}^{(0)} = P_{\text{ini}} \{x_2 = 3 | x_1 = 2\} = 0$$

$$(V) P_{23}^{(2)} = P_{\text{ini}} \{x_2 = 3 | x_1 = 2\} = \frac{1}{8}$$

$$(VI) P_{23}^{(3)} = P_{\text{ini}} \{x_2 = 3 | x_1 = 2\} = \frac{1}{8}$$

$$(VII) P_{33}^{(0)} = P_{\text{ini}} \{x_2 = 3 | x_1 = 3\} = 0$$

$$(VIII) P_{33}^{(2)} = P_{\text{ini}} \{x_2 = 3 | x_1 = 3\} = \frac{1}{8}$$

$$(IX) P_{33}^{(3)} = P_{\text{ini}} \{x_2 = 3 | x_1 = 3\} = \frac{1}{8}$$

Question: How to construct the t.p.m?

2011: Question (Hortoz)

Soln.: ~~2009~~ problem. Let x_n be the position of a particle performing a random walk with absorbing barriers at states 0 and 4. The unit-step transition probabilities are given below:-

$$\Pr\{x_n = n+1 \mid x_{n-1} = n\} = p$$

$$\Pr\{x_n = n-1 \mid x_{n-1} = n\} = q$$

$$\Pr\{x_n = 0 \mid x_{n-1} = 0\} = 1 = p_{00}^{(1)} = 1$$

$$\Pr\{x_n = 4 \mid x_{n-1} = 4\} = 1 = p_{44}^{(1)} = 1$$

Solution: The transition matrix is given by

		States of x_n				
		0	1	2	3	4
States of x_{n-1}	0	p_{00}	p_{01}	p_{02}	p_{03}	p_{04}
	1	p_{10}	p_{11}	p_{12}	p_{13}	p_{14}
2	p_{20}	p_{21}	p_{22}	p_{23}	p_{24}	
	p_{30}	p_{31}	p_{32}	p_{33}	p_{34}	
3	p_{40}	p_{41}	p_{42}	p_{43}	p_{44}	

$$= \begin{pmatrix} 0 & 0 & 1 & 2 & 3 & 4 \\ 1 & 2 & 0 & 0 & 0 & 0 \\ 2 & 0 & 0 & p & 0 & 0 \\ 3 & 0 & 0 & 0 & p & 0 \\ 4 & 0 & 0 & 0 & 0 & p \\ 0 & 0 & 0 & 0 & 0 & 1 \end{pmatrix}$$

(a) Define order of a markov chain.

Ans: Let A Markov chain $\{x_n\}$ is said to be of order $s(s=1, 2, 3)$ if for all n .

$$\Pr\{x_n = k \mid x_{n-1} = j_1, x_{n-2} = j_2, \dots, x_{n-s} = j_{s-1}\}$$

$$= \Pr\{x_n = k \mid x_{n-1} = j, \dots, x_{n-s} = j\}$$

Whenever the l.h.s is defined.

A Markov chain $\{x_n\}$ is said to be of order one (simply a Markov chain) if

$$\Pr\{x_n = k \mid x_{n-1} = j, x_{n-2} = j_1, \dots\}$$

$$= \Pr\{x_n = k \mid x_{n-1} = j\}$$

$$= p_{jk}$$

~~2010~~ Problem Draw the transition graph of the following Markov chain

$$\begin{pmatrix} \frac{5}{8} & \frac{3}{18} & 0 \\ 0 & \frac{1}{4} & \frac{3}{18} \\ 0 & \frac{3}{8} & \frac{1}{8} \end{pmatrix}$$

Solution: we have,

$$P = \begin{pmatrix} 0 & \frac{1}{8} & \frac{2}{8} \\ 1 & \frac{5}{8} & \frac{3}{8} \\ 2 & \frac{3}{8} & \frac{1}{8} \end{pmatrix}$$

The graph of the above chain is

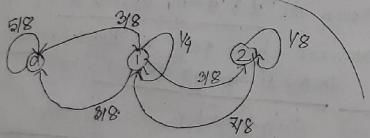


Fig:-1: Transition graph of Markov chain

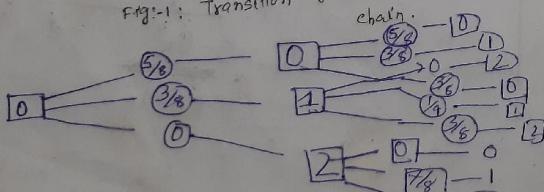


Fig: Transition trees;

Problem: Draw the transition graph of the following Markov chain

$$\begin{pmatrix} \frac{3}{4} & \frac{1}{4} & 0 \\ \frac{1}{4} & \frac{1}{2} & \frac{1}{4} \\ 0 & \frac{3}{4} & \frac{1}{4} \end{pmatrix}$$

Solution: let us define states

$$\text{states } \begin{pmatrix} 0 & 1 & 2 \\ 1 & \frac{3}{4} & \frac{1}{4} \\ 2 & \frac{1}{4} & \frac{3}{4} \end{pmatrix}$$

The graph of the above chain is

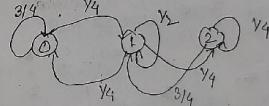
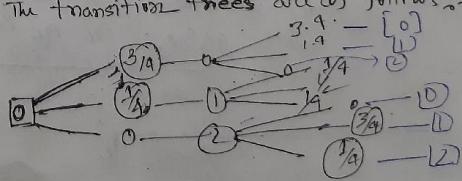


Fig: Transition graph of Markov chain.

The transitional trees occur as follows:-



~~Problem~~ Draw the transition graph of the following markov chain

$$\begin{pmatrix} 0 & y_2 & y_2 \\ y_2 & 0 & y_2 \\ y_2 & y_2 & 0 \end{pmatrix}$$

Solution: Let us define states

$$\begin{matrix} & \text{states} \\ \text{states} & \begin{pmatrix} 0 & y_2 & y_2 \\ y_2 & 0 & y_2 \\ y_2 & y_2 & 0 \end{pmatrix} \end{matrix}$$

The transition graph of the above chain is

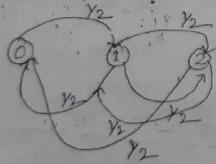
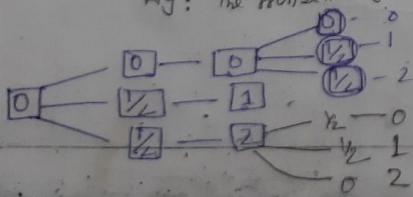


Fig: The transition graph of Markov chain.



2014, 2011

For state j and prove that Chapman-Kolmogorov equation

Show how these equations are used to generate n -step transition probability

Statement: For a unit step transition prob "s" gives the probability from the state j at a trial state K at the next following trial.

Note,

The $(n+m)$ step-transition probabilities is

$$P_{jk}^{(n+m)} = \sum_n P_{jk}^{(n)} \cdot P_{jk}^{(m)} = \sum_n P_{jn}^{(n)} P_{nk}^{(m)}$$

Proof: If state K can be reached from the state j in two steps through some intermediate state n then we can write, for fixed n

$$P_n [x_{n+2} = k | x_{n+1} = r | x_n = j]$$

$$= P_n [x_{n+2} = k | x_{n+1} = r] P_n [x_{n+1} = r | x_n = j]$$

$$= P_n [x_{n+2} = k | x_{n+1} = r] P_r [x_{n+1} = r | x_n = j]$$

$$P_{jk}^{(2)} = P_r^{(1)} P_k^{(1)}$$

since the intermediate state n can assume value

$n=1, 2, \dots$ then we set

$$P_{jk}^{(2)} = P_r [x_{n+2} = k | x_n = j]$$

$$= \sum_n P_n [x_{n+2} = k, x_{n+1} = m | x_n = j]$$

$$= \sum_n P_n [x_{n+2} = k | x_{n+1} = m] P_m [x_{n+1} = m | x_n = j]$$

$$= \sum_n P_n^{(1)} P_m^{(0)}$$

which is the Chapman-Kolmogorov equation for two step.

Again we can write,

$$P_{jk}^{(4)} = \sum_n P_n [x_{n+m} = k | x_n = j]$$

$$= \sum_n P_n [x_{n+m} = k | x_{n+m} = r]$$

$$\cdot P_r [x_{n+m} = r | x_n = j]$$

$$= \sum_n P_n^{(1)} P_r^{(m)}$$

(Similarly we get) $P_{jk}^{(m+1)} = \sum_n P_n^{(m)} P_r^{(1)}$

which is the Chapman-Kolmogorov equation for $(m+1)$ step.

In general we have

$$P_{jk}^{(m+n)} = \sum_r P_{jr}^{(n)} P_{rk}^{(m)}$$

$$= \sum_r P_{jr}^{(n)} P_{rk}^{(m)} \dots (1)$$

which is the Chapman-Kolmogorov equation for $m+n$ step.

Note From (1) we have

$$P_{jk}^{(m+n)} \geq P_{jr}^{(m)} P_{rk}^{(n)}$$

Let $P = P_{jk}$ be the transition matrix of the unit-step transition and $P^{(m)} = P_{jk}^{(m)}$ be the transition matrix for m unit-step transition. Put $m=2$ then we get $P^{(2)}$.

$$\text{i.e. } P^{(2)} = P \cdot P = P^2$$

similarly

$$P^{(m+1)} = P^m \cdot P = P \cdot P^m$$

$$\text{and } P^{(m+n)} = P^m \cdot P^n = P^n \cdot P^m$$

✓

~~W.P.~~ 2010.1.1

Problem 9: Let $\{x_n\}_{n \geq 0}$ be a homogeneous two-state Markov chain with multi-step transition matrix

$$P = \begin{pmatrix} q & p \\ p & q \end{pmatrix}$$

Show that

$$P^n = \begin{pmatrix} q + \frac{(q-p)^n}{2} & \frac{1}{2} - \frac{(q-p)^n}{2} \\ \frac{1}{2} - \frac{(q-p)^n}{2} & \frac{1}{2} + \frac{(q-p)^n}{2} \end{pmatrix}$$

Solution we have

$$P = \begin{pmatrix} q & p \\ p & q \end{pmatrix}$$

Now,

$$\begin{aligned} P^2 &= P \cdot P = \begin{pmatrix} q & p \\ p & q \end{pmatrix} \begin{pmatrix} q & p \\ p & q \end{pmatrix} \\ &= \begin{pmatrix} q^2 + p^2 & pq + pq \\ pq + pq & p^2 + q^2 \end{pmatrix} \\ &= \begin{pmatrix} p^2 + q^2 & 2pq \\ 2pq & p^2 + q^2 \end{pmatrix} \end{aligned}$$

$$\begin{aligned} P^3 &= \begin{pmatrix} \frac{(q+p)^2}{2} + \frac{(q-p)^2}{2} & \frac{(q+p)^2}{2} - \frac{(q-p)^2}{2} \\ \frac{(q+p)^2}{2} - \frac{(q-p)^2}{2} & \frac{(q+p)^2}{2} + \frac{(q-p)^2}{2} \end{pmatrix} \\ &= \begin{pmatrix} \frac{1}{2} + \frac{(q-p)^2}{2} & \frac{1}{2} - \frac{(q-p)^2}{2} \\ \frac{1}{2} - \frac{(q-p)^2}{2} & \frac{1}{2} + \frac{(q-p)^2}{2} \end{pmatrix} \end{aligned}$$

$$P^3 = P \cdot P \cdot P$$

$$= \begin{pmatrix} p^2 + q^2 & 2pq \\ 2pq & p^2 + q^2 \end{pmatrix} \begin{pmatrix} q & p \\ p & q \end{pmatrix}$$

$$\begin{aligned} &= \begin{pmatrix} q^3 + p^2q + 2pq^2 & p^3 + p^2q + 2pq^2 \\ p^3 + p^2q + 2pq^2 & q^3 + p^2q + 2pq^2 \end{pmatrix} \\ &\quad \xrightarrow{\text{Simplification}} \begin{pmatrix} q^3 + 3pq^2 & p^3 + 3pq^2 \\ p^3 + 3pq^2 & q^3 + 3pq^2 \end{pmatrix} \end{aligned}$$

$$\begin{aligned}
 &= \begin{pmatrix} \frac{(q-p)^3}{2} + \frac{(p+q)^3}{2} & \frac{(p+q)^3}{2} - \frac{(q-p)^3}{2} \\ \frac{(p+q)^3}{2} - \frac{(q-p)^3}{2} & \frac{(q-p)^3}{2} + \frac{(p+q)^3}{2} \end{pmatrix} \\
 &= \begin{pmatrix} \frac{1}{2} + \frac{(q-p)^3}{2} & \frac{1}{2} - \frac{(q-p)^3}{2} \\ \frac{1}{2} - \frac{(q-p)^3}{2} & \frac{1}{2} + \frac{(q-p)^3}{2} \end{pmatrix}
 \end{aligned}$$

similarly

$$P^4 = \begin{pmatrix} \frac{1}{2} + \frac{(q-p)^4}{2} & \frac{1}{2} - \frac{(q-p)^4}{2} \\ \frac{1}{2} - \frac{(q-p)^4}{2} & \frac{1}{2} + \frac{(q-p)^4}{2} \end{pmatrix}$$

In general we have

$$P^n = \begin{pmatrix} \frac{1}{2} + \frac{(q-p)^n}{2} & \frac{1}{2} - \frac{(q-p)^n}{2} \\ \frac{1}{2} - \frac{(q-p)^n}{2} & \frac{1}{2} + \frac{(q-p)^n}{2} \end{pmatrix}$$

(showed)

✓ 2010

Suppose that the initial distribution is

$$\{x_0=0\} \text{ and } P_{00} \{x_0=0\} = a$$

$$\text{and } P_{10} \{x_0=1\} = b = 1-a.$$

Applying Bayes' rule, show that

$$P_{01} \{x_0=0 | x_1=0\} = \frac{a \{1 + (q-p)^n\}}{1 + (a-b)(q-p)^n}$$

Solution we have:

$$P^n = \begin{pmatrix} 0 & \frac{1}{2} + \frac{(q-p)^n}{2} & \frac{1}{2} - \frac{1}{2} (q-p)^n \\ \frac{1}{2} - \frac{1}{2} (q-p)^n & \frac{1}{2} + \frac{1}{2} (q-p)^n \end{pmatrix}$$

Now

$$P_{00} \{x_0=0\} = a$$

$$P_{00} \{x_0=1\} = b = 1-a$$

$$\therefore P_{01} \{x_n=0, x_0=0\} = P_{00} \{x_n=0 | x_0=0\} P_{00} \{x_0=0\}$$

$$= P_{00}^{(n)} a$$

$$\text{Again, } = a P_{00}^{(n)}.$$

$$P_{10} \{x_n=0, x_0=1\} = P_{10} \{x_n=0 | x_0=1\} P_{10} \{x_0=1\}$$

$$= P_{10}^{(n)} b$$

$$= b P_{10}^{(n)}.$$

Applying Bayes rule we have

$$\begin{aligned}
 \Pr\{x_0=0 | x_n=0\} &= \frac{\Pr(x_0=0 | x_n=0) \Pr(x_0=0)}{\Pr(x_0=0 | x_n=0) \Pr(x_0=0)} \\
 &\quad + \Pr(x_0=0 | x_1=0) \Pr(x_1=0) \\
 &= \frac{a P_{00}^{(b)}}{a P_{00}^{(b)} + b P_{10}^{(b)}} \\
 &= \frac{a \left\{ \frac{1}{2} + \frac{1}{2} (2-p)^n \right\}}{a \left\{ \frac{1}{2} + \frac{1}{2} (2-p)^n \right\} + b \left\{ \frac{1}{2} - \frac{1}{2} (2-p)^n \right\}} \\
 &= \frac{\frac{a}{2} \left\{ 1 + (2-p)^n \right\}}{a \left\{ \frac{1}{2} + \frac{1}{2} (2-p)^n \right\} + \frac{1}{2} - \frac{1}{2} (2-p)^n} \\
 &= \frac{\frac{a}{2} \left\{ 1 + (2-p)^n \right\}}{\frac{a}{2} + \frac{a}{2} (2-p)^n + \frac{b}{2} - \frac{b}{2} (2-p)^n} \\
 &= \frac{\frac{a}{2} \left\{ 1 + (2-p)^n \right\}}{\frac{1}{2}(a+1-a) + \frac{b-a}{2} (2-p)^n}
 \end{aligned}$$

$$\Pr\{x_0=0 | x_n=0\} = \frac{a \left\{ 1 + (2-p)^n \right\}}{1 + (a-b) (2-p)^n} \quad (\text{shaded})$$

problem 95

$$P = \begin{pmatrix} 0 & 1 \\ 1-a & a \\ b & 1-b \end{pmatrix}$$

$$\Pr\{x_0=1\} = P_1 = \frac{1}{2} + \Pr\{x_0=0\} = \frac{1}{2} - \frac{1}{2} p$$

$$\text{Show that } \Pr_1 = \frac{a}{a+b} + \left(1 - \frac{a}{a+b}\right) \left(1-a-b\right)^n \quad \checkmark$$

solutions we have

$$P = \begin{pmatrix} 1-a & a \\ b & 1-b \end{pmatrix}$$

$$\begin{aligned}
 P_1 &= \left(\frac{b}{a+b} + \left(1-\frac{a}{a+b}\right) \frac{a}{a+b} \right) \left(\frac{a}{a+b} - \left(1-a-b\right) \frac{a}{a+b} \right) \\
 &= \left(\frac{b}{a+b} - \left(1-a-b\right) \frac{b}{a+b} \right) \left(\frac{a}{a+b} + \left(1-a-b\right) \frac{b}{a+b} \right)
 \end{aligned}$$

$$(1-a-b) \left(\frac{b}{a+b} + \left(1-\frac{a}{a+b}\right) \frac{a}{a+b} \right) \left(\frac{a}{a+b} - \left(1-a-b\right) \frac{a}{a+b} \right)$$

$$\begin{aligned}
 &= \begin{pmatrix} (1-a)^2 ab + a(1-a)^2 + b(1-b) \\ a(1-a)(1-b) + ? \\ b(1-a) + b(1-b) \\ ab + (1-b)^2 \end{pmatrix} \\
 &= \begin{pmatrix} 1-2a+a^2 ab & a - a^2 - ab \\ ab - b^2 - ab & 1-2b+b^2 + ab \end{pmatrix} \\
 &= \begin{pmatrix} \frac{b}{a+b} + (1-a-b)^n \frac{a}{a+b} & \frac{a}{a+b} - (1-a-b)^n \frac{a}{a+b} \\ \frac{b}{a+b} - (1-a-b)^n \frac{b}{a+b} + \frac{a}{a+b} + (1-a-b)^n \frac{b}{a+b} \end{pmatrix}
 \end{aligned}$$

Similarly

$$p_{11} = \begin{pmatrix} \frac{b}{a+b} + (1-a-b)^n \frac{a}{a+b} & \frac{a}{a+b} - (1-a-b)^n \frac{a}{a+b} \\ \frac{b}{a+b} - (1-a-b)^n \frac{b}{a+b} & \frac{a}{a+b} + (1-a-b)^n \frac{b}{a+b} \end{pmatrix}$$

We know that

$$\begin{aligned}
 p_{11} &= \Pr\{X_1 = 1\} \\
 &= \Pr\{X_1 = 1 | X_0 = 0\} + \Pr\{X_1 = 1 | X_0 = 1\}
 \end{aligned}$$

$$\begin{aligned}
 &= p_{11}^{(0)} \Pr\{X_0 = 1\} + p_{11}^{(1)} \Pr\{X_0 = 1\} \\
 \text{Now given that} \\
 \Pr\{X_0 = 1\} &= p_1 = 1 - q \\
 \text{Now,} \\
 p_{11} &= p_{11}^{(0)} p_1 + p_{11}^{(1)} q \\
 &= p_{11}^{(0)} p_1 + p_{11}^{(1)} (1 - p_1) \\
 &= p_{11}^{(0)} p_1 + p_{11}^{(1)} - p_{11}^{(1)} p_1 \\
 &= p_{11}^{(0)} + p_{11}^{(1)} - p_{11}^{(1)} p_1 \\
 &= \frac{a}{a+b} - (1-a-b)^n \frac{a}{a+b} + p_1 \left\{ \frac{a}{a+b} + (1-a-b)^n \frac{b}{a+b} \right. \\
 &\quad \left. - \frac{a}{a+b} + (1-a-b)^n \frac{a}{a+b} \right\} \\
 &= \frac{a}{a+b} - (1-a-b)^n \frac{a}{a+b} + p_1 (a-b)^n \cdot \frac{(a+b)}{(a+b)} \\
 &= \frac{a}{a+b} - (1-a-b)^n \frac{a}{a+b} + p_1 (a-b)^n \\
 &= \frac{a}{a+b} - (1-a-b)^n \frac{a}{a+b} + p_1 (1-a-b)^n
 \end{aligned}$$

$$\therefore p_{11} = \frac{a}{a+b} - (1-a-b)^n \frac{a}{a+b} + p_1 (1-a-b)^n \quad \text{shown}$$

Problem 9

$$P = \begin{pmatrix} 1-(1-c)P & (1-c)P \\ (1-c)(1-P) & (1-c)P+c \end{pmatrix}$$

Show that $p_n = c p_{n-1} + (1-c)P$.

Solution: we have

$$P = \begin{pmatrix} 0 & 1 \\ 1-(1-c)P & (1-c)P \\ (1-c)(1-P) & (1-c)P+c \end{pmatrix}$$

We know that

$$\begin{aligned} p_n &= \Pr\{X_n = 1\} \\ &= \Pr\{X_{n-1} = 1 | X_n = 1\} \Pr\{X_n = 1 | X_{n-1} = 1\} + \Pr\{X_{n-1} = 0 | X_n = 1\} \Pr\{X_n = 1 | X_{n-1} = 0\} \\ &= p_{n-1}^{(1)} p_{n-1} + p_{n-1}^{(0)} q_{n-1} \\ &= 2(1-c)P + (1-c)P(1-p_{n-1}) \\ &= (1-c)P p_{n-1} + c p_{n-1} + (1-c)P - (1-c)P p_{n-1} \\ &= c p_{n-1} + (1-c)P \quad (\text{shown}) \end{aligned}$$

C.V.

(b) Define persistent, transient, null persistent and non-null persistent, ergodic.

Ans:

Persistent: A state j is said to be persistent or recurrent if and only if $F_{jj} = \infty$ i.e. return to state j is certain.

Transient: A state j is said to be transient if and only if $F_{jj} < \infty$ i.e. return to state j is uncertain.

Null persistent: A persistent state j is said to be null persistent if and only if $\mu_j = \infty$ i.e. the mean recurrence time is infinite.

Non-null persistent: A persistent state j is said to be non-null persistent if and only if $\mu_j < \infty$ i.e. the mean recurrence time is less than infinite.

Ergodic: A non-null persistent and aperiodic state of a markov chain is said to be ergodic.

Q Define recurrent, positive recurrent and null recurrent markov chain.

Ans

recurrent: A state j is said to be recurrent if and only if $F_{jj} = 1$ return to the state j is certain.

positive recurrent: A recurrent state j is said to be positive recurrent if and only if $\mu_{jj} < \infty$ i.e. the mean recurrence time is infinite.

null recurrent: A recurrent state j is said to be null recurrent if and only if $\mu_{jj} = \infty$ i.e. the mean recurrence time is less than infinite.

Q Distinguish between a persistent and a transient state.

Ans The difference between a persistent state and transient state are given below:-

(i) If $F_{jj} = 1$ then the state j is said to be persistent but if $F_{jj} \leq 1$ then the state j is said to be transient state.

(ii) For state j , the persistent state return to the state j is certain but transient state is uncertain.

(iii) Therefore, are two types of persistent state null persistent and non-null persistent but transient has no type.

Q when is a state null persistent? when it is periodic.

Ans A persistent state j is said to be null persistent if and only if $\mu_{jj} < \infty$ i.e. the mean recurrence time is infinite.

Let us consider a transition probability matrix P .

In general $p_{2n} = p^n$ and $p = p^{2n+1}$ so that $p_{2n} > 0$ and $p_{2n+1} = 0$ for each i . Then the state is said to periodic.

which but transient state is uncorrected

~~Q~~ Question : construct the transition matrix w.r.t
the following statement :-

012 0010 22 11 000 12 102 10

Answers :-

0 1 2

$$P = \begin{matrix} 0 & \left(\begin{matrix} III = \frac{3}{8} & II = \frac{3}{8} & I = \frac{2}{8} \\ III = \frac{9}{12} & II = \frac{1}{12} & I = \frac{2}{12} \\ I = \frac{1}{5} & III = \frac{3}{5} & II = \frac{1}{5} \end{matrix} \right) \\ 1 & \\ 2 & \end{matrix}$$

$$\therefore P = \begin{pmatrix} \frac{3}{8} & \frac{3}{8} & \frac{2}{8} \\ \frac{9}{12} & \frac{1}{12} & \frac{2}{12} \\ \frac{1}{5} & \frac{3}{5} & \frac{1}{5} \end{pmatrix}$$

$$P = \begin{matrix} 0 & \left(\begin{matrix} 0 & 1 & 2 \\ III & II & I \\ III & II & I \end{matrix} \right) \\ 1 & \\ 2 & \end{matrix}$$

$$P = \begin{matrix} 0 & \left(\begin{matrix} III = 3 & II = 3 & I = 2 \\ III = 1 & II = 1 & I = 1 \\ III = 1 & II = 1 & I = 1 \end{matrix} \right) \\ 1 & \\ 2 & \end{matrix} \rightarrow + n P \quad \begin{pmatrix} \frac{3}{8} & \frac{3}{8} & \frac{2}{8} \\ \frac{9}{12} & \frac{1}{12} & \frac{2}{12} \\ \frac{1}{5} & \frac{3}{5} & \frac{1}{5} \end{pmatrix}$$