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# The Effects of weather on maize yields: New evidence from Kenya

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**Abstract.** ...Applying the linear mixed effects models, we found that...

.. to be written later..

## 1. Introduction

### Paragraph 1

- Extreme weather causes disasters → early warning systems have been developed

### Paragraph 2

- What weather forecasts (measures) have been used in EWS? *ref. literature*
  - \* Mostly seasonal precip. totals and temperature averages

### Paragraph 3

- Identify difference between hazard and disaster
  - \* Not every hazard turns into disaster
  - \* For a hazard to become a disaster it needs to have **impact**
  - \* Here, we identify the key metrics which have impact on yield

### Paragraph 4

- Crop yield versus climate forecasting

### Paragraph 5

- Aim of the paper

## 2. Methods

...a case study looking at Kenya...

### 2.1. Data

In this study, we analysed the relationship between maize yields and climate. Our dataset consisted of an yearly panel of 47 counties of Kenya describing the period of 1981 – 2017. We acquired the county level yearly yield data from the Famine Early

Warning Systems Network (FEWS NET). Regarding the weather data, we exploited 0.25° resolution precipitation and temperature gridded datasets. The precipitation data were obtained from the Climate Hazards Group InfraRed Precipitation with Station data (CHIRPS) while the temperature data are available at the website of the Berkeley Earth. We calculated spatial averages of the gridded data over the counties to get a single value for each county and each point in time. The frequency of precipitation and temperature data was daily, therefore, further aggregation was needed to obtain yearly values corresponding to the yearly frequency of the yield data. Hence, each weather characteristic was further aggregated over years resulting in a county-level panel conformable with the yield data.

There are two predominant rainfall regimes in Kenya. The arid and semi-arid (ASAL) counties exhibit mostly bi-modal precipitation patterns with long rains lasting from March to May and short rains occurring between October and December. In the non-ASAL counties, the single rainy season usually starts in March and lasts until August. Following the precipitation patterns and the closely related planting and harvesting calendar, we computed yearly values of various weather characteristics as follows: For the ASAL counties, we used daily data covering October, November and December of the previous year and March, April, May of the current year and for the non-ASAL counties we used daily data covering March to August of the current year.

Following the procedure described above, we calculated a number of previously used characteristics of precipitation and temperature including indicators of floods. These measures are listed in table 2 in the appendix. The measures which we found significant were: Total seasonal precipitation and its squared values, average seasonal temperature, coefficient of variation of seasonal precipitation, coefficient of variation of average seasonal temperature (converted into Kelvin because the Celsius scale is

interval and the coefficient of variation does not have any meaning for data measured on an interval scale), maximum length of dry spell measured as number of consecutive dry days and the number of dry spells lasting for four days or more (for the purpose of this study we defined a dry day as a day when precipitation didn't exceed 1mm). We did not include squared values of the average seasonal temperature in our preferred models as it was not significant and most of the past literature has not supported the evidence of the non-linear relationship between yield and temperature.

## *2.2. Statistical approach*

Kenya consists of 47 counties with semi-autonomous county governments. As a result of the high degree of county-level autonomy, the policies and regulations often differ across the counties, hence the effects of weather on crop yield are likely to be different across the counties. Therefore, following the standard methodology, we estimated a battery of linear mixed effects models (also known as mixed models) commonly used to analyse longitudinal data (Bates, Pinheiro, Pinheiro & Bates 2000). Mixed models are suitable for analysis of panel data as they account for the panel structure of the dataset. These types of models include both fixed effects parameters and random effects. Fixed effects are analogous to parameters in a classical linear regression model and value of each effect is assumed to be fixed over all counties (Bates 2010). On the other hand, random effects are unobserved random variables. There are at least three benefits of treating a set of parameters as a random sample from some distribution. *(i)* Extrapolation of inference to a wider population *(ii)* improved accounting for system uncertainty and *(iii)* efficiency of estimation (Kery 2010b, Kery 2010a).

Formally, a linear mixed model can be described by the distribution of two vectors of random variables: the response  $\mathcal{Y}$  and the vector of random effects  $\mathcal{B}$ . The distribution of  $\mathcal{B}$  is multivariate normal and the conditional distribution of  $\mathcal{Y}$  given  $\mathcal{B} = \mathbf{b}$  is

multivariate normal of a form

$$(\mathcal{Y}|\mathcal{B} = \mathbf{b}) \sim N(\mathbf{X}\beta + \mathbf{Zb}, \sigma^2\mathbf{I}), \quad (1)$$

where  $\mathbf{X}$  is an  $n \times p$  model matrix of fixed effects,  $\beta$  is a  $p$ -dimensional fixed-effects parameter,  $\mathbf{Z}$  is an  $n \times q$  model matrix for the  $q$ -dimensional vector of random-effects variable  $\mathcal{B}$  evaluated at  $\mathbf{b}$  and  $\sigma$  is a scale factor. The distribution of  $\mathcal{B}$  can be written as:

$$\mathcal{B} \sim N(0, \Sigma), \quad (2)$$

where  $\Sigma$  is a  $q \times q$  positive semi-definite variance-covariance matrix.

It was apparent from the histogram of the maize yield data that the distribution of maize yield was closer to a log-normal than to a normal distribution. Therefore, we opted for a log-linear functional form of our models as is a common practice in the agricultural economics. Furthermore, we scaled all predictors by subtracting mean and dividing them by standard deviation to avoid convergence problems.

To find the most suitable set of fixed effects, we adapted a stepwise selection procedure which minimized the Akaike's Information Criterion (AIC). The first step of the procedure was the estimation of the model with the complete set of our weather measures in the fixed effects and with the random intercepts. After this, we performed a search through the subsets of the fixed effects minimizing the AIC and allowing for steps in both 'backward' and 'forward' directions (that is removing and adding the predictors). We decided not to include number of heatwave days and average maximum daily temperature although including them as fixed effects would slightly improve the AIC. The reasons for not including them were their strong correlation with average seasonal temperature (the correlation coefficient was 0.981 in the case of maximum daily temperature and 0.517 for the number of heatwave days with the p-values smaller

than  $1 \times 10^{-6}$  in both cases), high variance inflation factor (VIF) of the maximum temperature (5.397) and the lack of significance of their individual t-tests (the p-values were 0.121 for the maximum daily temperature and 0.075 for the number of heatwave days). Furthermore, including heatwave days and maximum daily temperature would only lead to a small change in the AIC from 2080.877 to 2076.439 (these values differ from the AIC presented in table 3 in the appendix because the final estimates reported in table 3 were estimated using the restricted maximum likelihood (REML) whereas the models were re-estimated using the ML method for comparison<sup>‡</sup>). In addition, the coefficients of both heatwave days and maximum daily temperature were positive, which is opposite to what we expected. Hence, we concluded that the decrease in the AIC and the positive signs are likely to be results of the high correlation with average temperature and we did not include them in our preferred specification.

We further applied a series of likelihood ratio tests to test for the significance of random slopes. No subset of the random slopes improved the fit; therefore, besides the fixed effects, our final model only included the random intercepts.<sup>§</sup>

To verify our results, we adapted an alternative step-down model building approach using the Satterthwaite's method to determine the p-values of the individual t-tests (Kuznetsova, Brockhoff & Christensen 2017). We started with a model which included the complete set of our weather measures in both fixed effects and random effects. In the first part of the procedure, the insignificant random effects were removed one by one. In each iteration, the insignificant variable with the highest p-value was removed.

<sup>‡</sup> REML is generally the preferred estimation method for the mixed effects models as the ML estimates of the variance component are biased, however, the REML criterion is not meaningful for comparison of the models with different fixed effects. Therefore, we re-estimated the models by ML for the purpose of comparison (Zuur, Ieno, Walker, Saveiliev & Smith 2009).

<sup>§</sup> Although adding the random slope of the coefficient of variation improved the AIC from 2122.226 to 2122.110, we decided to not to include the coefficient of variation of temperature in the random effects. The reason for not including the random slope is that the difference in AIC is negligible and the F-test of the comparison of the two models is insignificant favouring the simpler model (p-value= 0.128).

This was repeated until only the significant predictors remained. After this, we applied an analogous procedure to the fixed effects until we were left only with the significant variables. As recommended in literature, we considered the level of significance  $\alpha = 0.1$  for random effects and  $\alpha = 0.05$  for fixed effects (Kuznetsova et al. 2017). This method led to the same model as the one obtained with our primary AIC-based method described above.

According to the conditional Lagrange multiplier (LM) test developed by Baltagi & Li (1991) and Baltagi & Li (1995), the errors exhibited a within group autocorrelation structure in our models (the p-value of the LM test was  $1.1 \times 10^{-13}$ ). To further investigate the autocorrelation structure of the errors, we estimated a number of models (with the subset of predictors chosen as described above) with an ARMA( $p,q$ ) error structure. In particular, we estimated all variants of our model with ARMA( $p,q$ ) errors such that  $p \leq 2$  and  $q \leq 2$ . Comparing the AIC criteria and using the corresponding likelihood ratio statistics, we found that the most appropriate error correlation structure was ARMA(1,1). The value of the AIC criterion of the model with the ARMA(1,1) error structure was 2122.2 while the AIC was 2129.2 for the model with ARMA(1,0) errors. The p-value of the likelihood ratio test of the comparison of the model with ARMA(1,1) errors and the model with ARMA(1,0) errors was smaller than 0.003, hence the ARMA(1,1) error structure turned out to be a better fit. The AIC criteria and the likelihood ratio tests of all the models with ARMA( $p,q$ ) errors are summarised in table 3 in the appendix.

### **3. Results**

Besides our main model, we estimated the same specification for two additional subsamples of the Kenyan counties. In particular, we estimated the model for the

subsample of the ASAL counties and for the subsample of the non-ASAL counties separately to compare the effects of various weather characteristics across the two areas with the significantly different climate. The estimates of all three models are summarised in table 1. The first two columns represent the model for all counties, the third and fourth columns include estimates for the subsample of the ASAL counties and the last two columns describe the estimates based on the subsample of the non-ASAL counties. All predictors were standardised, therefore the estimates and the corresponding p-values can be interpreted as measures of relative importance. As an additional measure of relative importance we included F-values of the type *III* analysis of variance (ANOVA) in table 1. The F-tests were conducted using the marginal rather than the sequential sum of squares. That is, for each variable, the F-test is a test of significance of that part of explained sum of squares which can be contributed to the particular explanatory variable. In other words, the F-value is a test statistic of comparison of the preferred model (with all the variables in table 1) and the model without the explanatory variable in question but including all other predictors. The main advantage of the type *III* ANOVA is that the test statistics do not depend on the order in which the effects are entered in the model.

The preferred specification has a log-linear functional form, therefore the estimates in table 1 are not directly interpretable as marginal effects and their effects on yield are multiplicative. The predicted marginal effects of all coefficients are depicted in figures 1 in this section and in figures 1 and 2 in the appendix. The units of all predictors in all three figures are multiples of their standard deviations since the predictors were standardised. The exponents of the estimates (which can be interpreted as the additive marginal effects on yield) can be found in table 4 in the appendix.

In the following, we will discuss the estimates based on the sample of all counties



and after that we will focus on the estimates based on the two subsamples.

**Table 1. *Mixed effects model:***  
*Log of maize yield and weather, ARMA(1,1) errors*

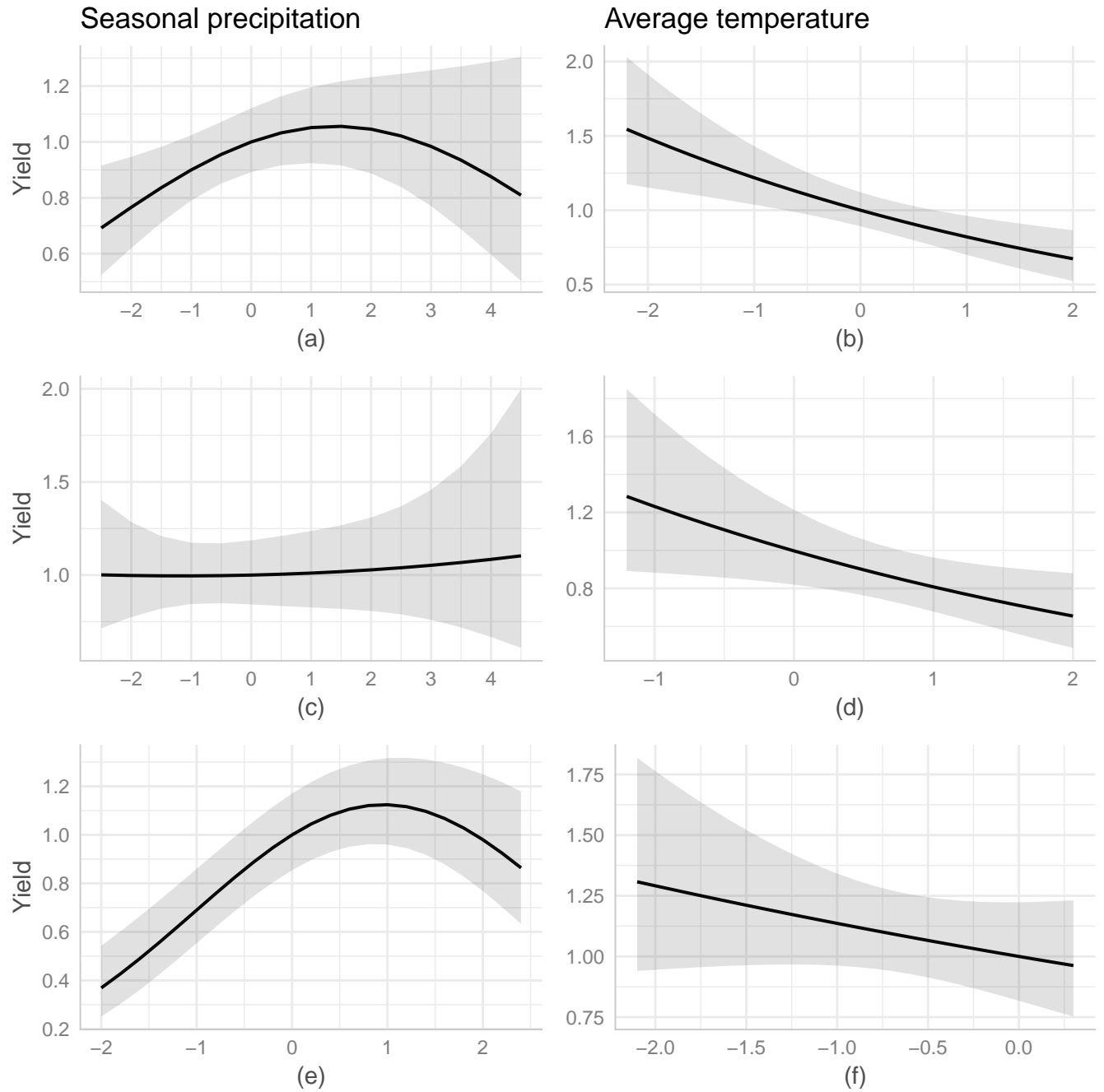
Fixed effects:	<i>All counties</i>		<i>ASAL</i>		<i>non-ASAL</i>	
	Estimate	F-value <sup>a</sup>	Estimate	F-value <sup>a</sup>	Estimate	F-value <sup>a</sup>
Intercept	0.259***	19.908	0.241*	5.084	0.342**	9.891
Prec. total	0.078*	5.358	0.007	0.028	0.244***	19.140
Prec. total sq.	-0.028*	4.263	0.003	0.037	-0.127***	23.384
Prec. c. of var.	-0.079•	3.278	-0.031	0.237	-0.097	2.326
Dry spell -length	-0.067*	4.800	-0.181**	6.841	-0.012	0.164
Dry spells $\geq$ 4 d.	-0.063*	4.848	-0.156**	8.032	-0.011	0.087
Temp. - average	-0.199***	12.010	-0.214*	5.409	-0.128	1.525
Temp. c. of var.	0.042•	2.961	0.032	0.355	0.058*	5.372
<b>Random effects:</b>						
Intercept						
<i>Number of observations:</i>	1300		698		602	

Notes: Standard errors in brackets; •  $p < 0.1$ ; \*  $p < 0.05$ ; \*\*  $p < 0.01$ ; \*\*\*  $p < 0.001$

<sup>a</sup> Marginal (type III) sum of squares. The F-statistics correspond to the sum of squares attributable to each fixed effect.

The estimates of the model based on all the Kenyan counties are summarised in the first two columns of table 1. Consistently with our assumptions, seasonal precipitation and average temperature have significant effect on yields. As we expected, the effect of temperature is negative. The linear term of precipitation is positive while the squared term is negative and as we can see in figure 1(a), the marginal effect of precipitation is hill-shaped

Besides seasonal precipitation and average temperature, the length and number of dry spells



**Figure 1.** Predicted multiplicative marginal effects of seasonal precipitation (left column) and average seasonal temperature (right column) on maize yields assuming that all other variables are fixed at their means (which is zero, as the predictors were standardised). The first row (a, b) represents the model for all counties, the second row (c, d) is based on the subsample of arid and semi-arid counties (ASAL) and the third row (e, f) represents the model for the non-ASAL counties. Precipitation and temperature (x-axis) are in multiples of their standard deviations as the models were standardised. The effects are multiplicative as the models are in the log-linear form.

- Verbal description and interpretation of the results. Discussing goodness of fit using various criteria such as AIC or alternatives to  $R^2$ .
- relative importance of the individual variables and its measures
- show VIF
- If we get the yield data for the period from 2015 onwards: Out of sample predictions and comparison with the real data

## Appendix

*should be at the end of the main text but before list of references*

**Table 1. List of arid and semi-arid (ASAL) and non-ASAL counties**

<b>ASAL:</b>	Baringo, Embu, Garissa, Isiolo, Kajiado, Kilifi Kitui, Kwale, Laikipia, Lamu, Makueni, Mandera, Marsabit, Meru, Mombasa, Narok, Nyeri, Samburu, Taita-Taveta, Tana River, Tharaka Nithi, Turkana, Wajir, West Pokot
<b>non-ASAL:</b>	Bomet, Bungoma, Busia, Elgeyo Marakwet, Homa Bay, Kakamega, Kericho Kiambu, Kirinyaga, Kisii, Kisumu, Machakos, Migori, Murang'a, Nakuru, Nyamira, Nyandarua, Siaya, Trans Nzoia, Uasin Gishu, Vihiga

**Table 2.** Precipitation and temperature measures considered

- 
- Total rainfall over the rainy season
  - Coefficient of variation of the rainfall during the rainy season
  - Maximum length of dry spell during the rainy season (in number of days)
  - Number of drought spells during the rainy season: a drought spell defined as 4 consecutive days without rain or more<sup>a</sup>
  - Number of drought spells during the rainy season: a drought spell defined as 10 consecutive days without rain or more<sup>a</sup>
  - Number of drought spells during the rainy season: a drought spell defined as 20 consecutive days without rain or more<sup>a</sup>
  - Average temperature during the rainy season
  - Standard deviation of temperature during the rainy season
  - Cumulative degree days during the rainy season (excluding the days when maximum temperature above 30°C or below 10°C)
  - Number of heatwave days during the rainy season when max. temperature > 35°C
  - Maximum daily rainfall - to control for possible floods
  - Sum of precipitation amount on days where precipitation is above 90<sup>th</sup> percentile of precip. of the whole period<sup>b</sup>
  - Sum of precipitation amount on days where precipitation is above 95<sup>th</sup> percentile of precip. of the whole period<sup>c</sup>
  - Sum of precipitation amount on days where precipitation is above 99<sup>th</sup> percentile of precip. of the whole period<sup>d</sup>
  - Number of days where precipitation is above 90<sup>th</sup> percentile of precip. of the whole period<sup>b</sup>
  - Number of days where precipitation is above 95<sup>th</sup> percentile of precip. of the whole period<sup>c</sup>
  - Number of days where precipitation is above 99<sup>th</sup> percentile of precip. of the whole period<sup>d</sup>
- 

<sup>a</sup> Threshold for a dry day was considered 1mm.

<sup>b</sup> 90<sup>th</sup> percentile was calculated for the subsample of wet days, that is the days where precipitation is above or equal to 1mm.

<sup>c</sup> 95<sup>th</sup> percentile was calculated for the subsample of wet days, that is the days where precipitation is above or equal to 1mm.

<sup>d</sup> 99<sup>th</sup> percentile was calculated for the subsample of wet days, that is the days where precipitation is above or equal to 1mm.

**Table 3.** Comparison of models with different error autocorrelation structure

Error autocorrelation structure	AIC	Likelihood ratio vs. ARMA(1,1) <sup>a</sup>	
		Statistic	p-value
<i>None</i>	2205.4	87.17	$< 1 \times 10^{-4}$
ARMA(1,0)	2129.2	8.99	0.0027
ARMA(0,1)	2145.0	24.73	$< 1 \times 10^{-4}$
ARMA(1,1)	2122.2	— — —	— — —
ARMA(2,1)	2124.1	0.15	0.6990
ARMA(1,2)	2124.2	0.01	0.9181
ARMA(2,2)	2125.9	0.31	0.8549

<sup>a</sup> Likelihood ratio test of a comparison of the model in each row against the ARMA(1,1) error structure model. ARMA(1,1) error structure seems to be the most suitable as all lower-order correlation structure models are rejected against ARMA(1,1) while ARMA(1,1) is not rejected against any of the higher order structures.

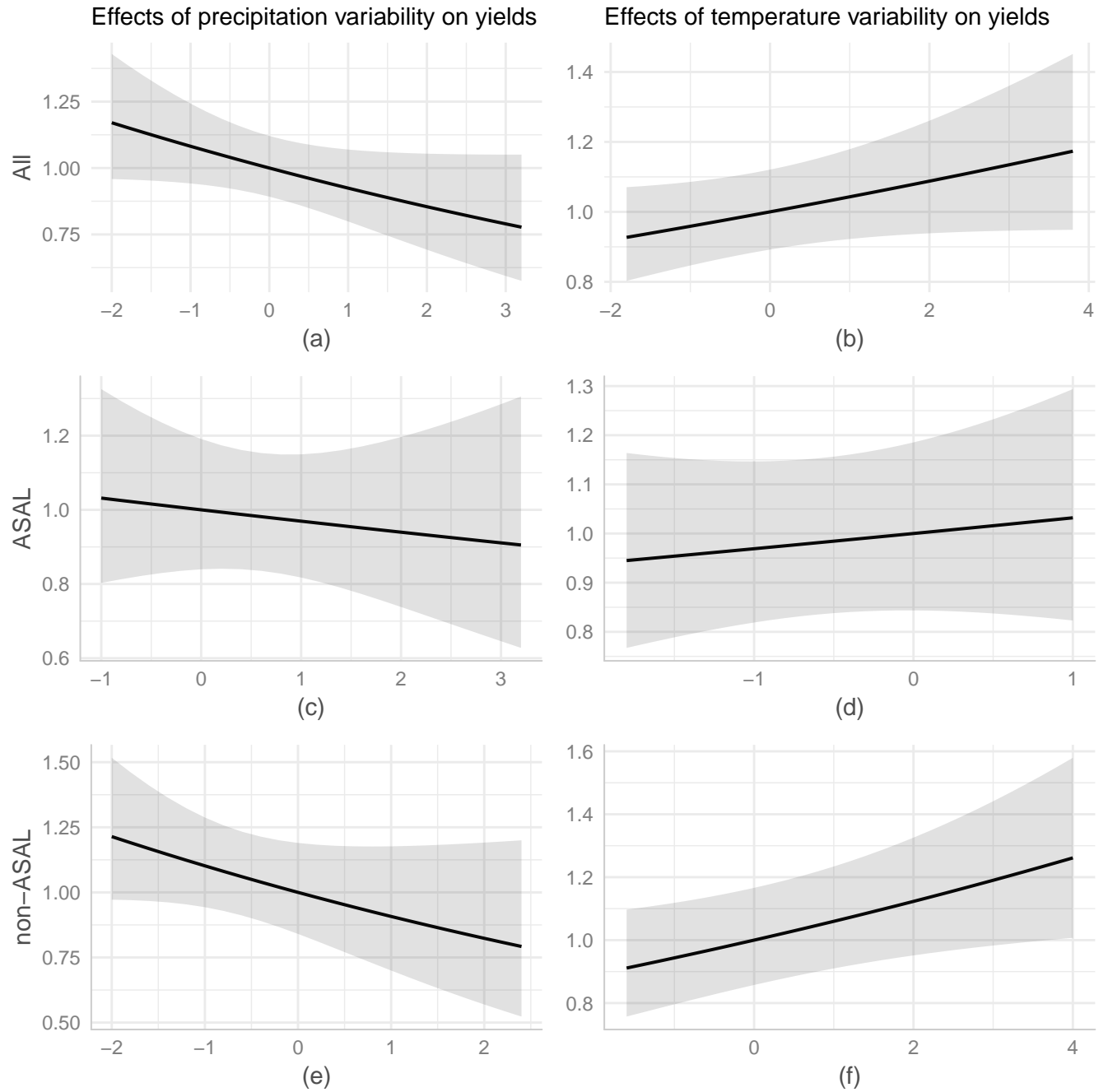
Possibly include a table of all values which I get from the lme or lme4 models summary in R, that is. correlation of the fixed effects etcetera

Maybe also a table with standard errors here

**Table 4. Mixed effects model:** exponents of the coefficient estimates Log of maize yield and weather, ARMA(1,1) errors

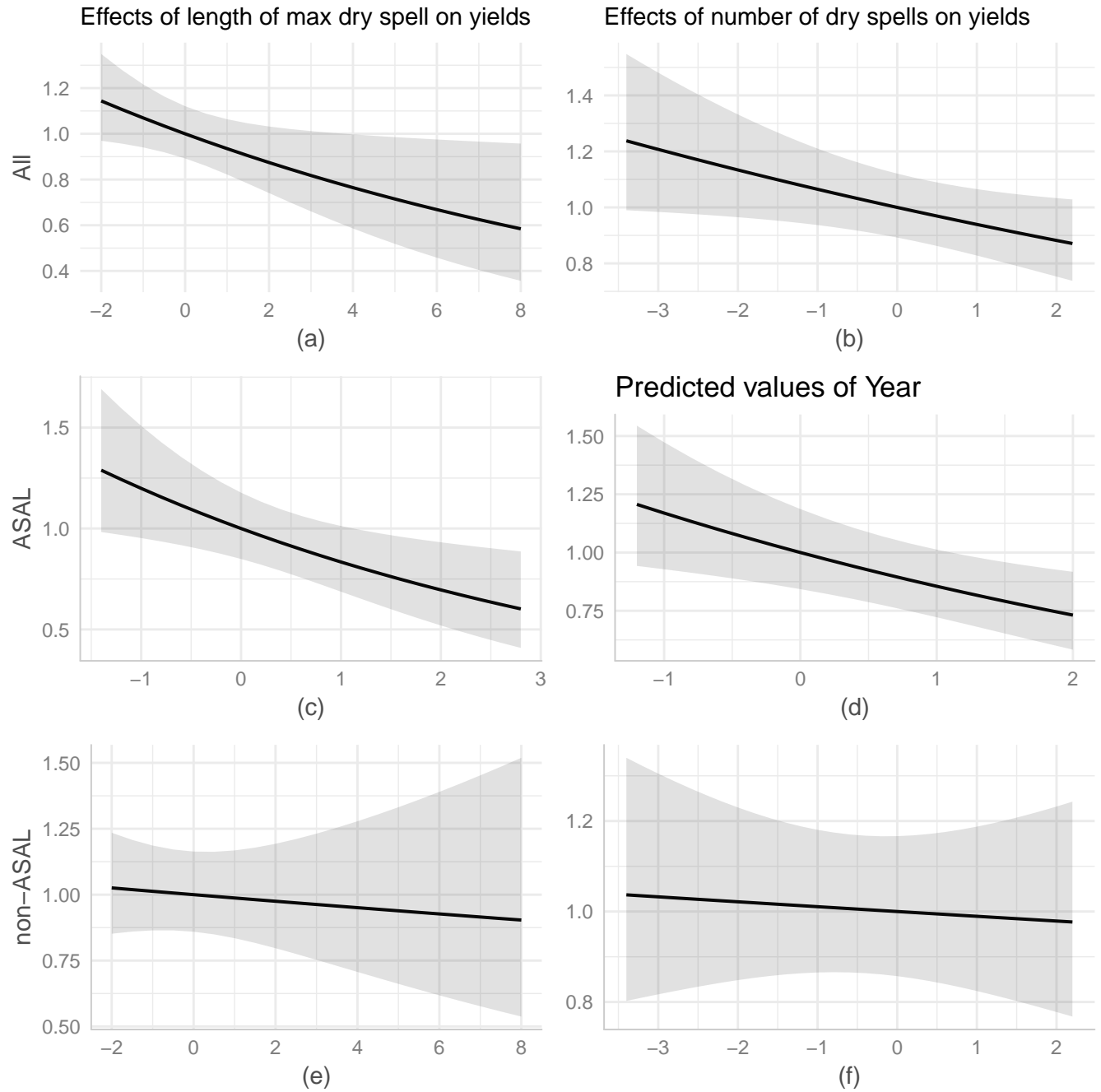
<b>Fixed effects:</b>	<i>All counties</i>	<i>ASAL</i>	<i>non-ASAL</i>
Intercept	1.296***	1.272*	1.408**
Prec. total	1.081*	1.007	1.277***
Prec. total sq.	0.973*	1.003	0.881***
Prec. c. of var.	0.924•	0.970	0.907
Dry spell -length	0.935*	0.834**	0.988
Dry spells $\geq$ 4 d.	0.939*	0.855**	0.989
Temp. - average	0.820***	0.808*	0.880
Temp. c. of var.	1.043•	1.032	1.060 *
<b>Random effects:</b>			
Intercept			
<i>Number of observations:</i>	1300	698	602

Notes: Standard errors in brackets; •  $p < 0.1$ ; \*  $p < 0.05$ ; \*\*  $p < 0.01$ ; \*\*\*  $p < 0.001$



**Figure 1.** Predicted multiplicative marginal effects of coefficient of variation (CV) of precipitation (left column) and CV of temperature (right column) on maize yields assuming that all other variables are fixed at their means (which is zero, as the predictors were standardised). The first row (a, b) represents the model for all counties, the second row (c, d) is based on the subsample of arid and semi-arid counties (ASAL) and the third row (e, f) represents the model for the non-ASAL counties. CV of precipitation and temperature (x-axis) are in multiples of their standard deviations as the models were standardised. The effects are multiplicative because the models are in the log-linear form.





**Figure 2.** Predicted multiplicative marginal effects of length of maximum dry spell in days (left column) and number of dry spells lasting for four days or more (right column) on maize yields assuming that all other variables are fixed at their means (which is zero, as the predictors were standardised). The first row (a, b) represents the model for all counties, the second row (c, d) is based on the subsample of arid and semi-arid counties (ASAL) and the third row (e, f) represents the model for the non-ASAL counties. Maximum length of dry spell and number of dry spells (x-axis) are in multiples of their standard deviations as the models were standardised. The effects are multiplicative because the models are in the log-linear form.

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