

## Application 3: Analysis of an Censored Constituent using a Seasonal Model

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This application illustrates the "7-parameter model," a predefined model that has been shown to perform well for loading analyses of constituents in large ( $> 100$  square miles) watersheds (Cohn and others, 1992). In addition, the application illustrates the use of LOADEST when portions of the calibration data set are subject to censoring.

As in the previous example, a constituent with a seasonal loading pattern is considered here. In this case, constituent concentrations are assumed to vary in a continuous manner, as opposed to the abrupt changes considered in Application 2. Several of the predefined models (models 4 and 6–9; Section 3.2.2, table 7) use a first-order Fourier series (sine and cosine terms) to consider seasonality. In this application, the 7-parameter model (model 9) is developed for nutrient loading on the Potomac River. The 7-parameter model is given by:

$$\log(\text{Load}_i) = \alpha_0 + \alpha_1 \ln Q_i + \alpha_2 \ln Q_i^2 + \alpha_3 cT_i + \alpha_4 cT_i^2 + \alpha_5 \sin(2\pi dT_i) + \alpha_6 \cos(2\pi dT_i) + \epsilon_i, \quad (1)$$

where  $\ln Q_i$  is the centered log of flow,  $cT_i$  is centered decimal time, and  $dT_i$  is the decimal time for observation  $i$ . Within the model, explanatory variables one and two account for the dependence on flow, explanatory variables three and four account for the time trend, and explanatory variables five and six are a first-order Fourier series to account for seasonal variability.

The load regression model for orthophosphate data collected near USGS gaging station 01646580 on the Potomac River uses equation 1. The retrieved dataset includes 237 observations of concentrations collected from 2002 to 2010; many of the observations are below the laboratory detection limit, resulting in a censored data set. The flow data will be from USGS gaging station 01646502, located just upstream from the water-quality gage.

# 1 Retrieve and Build the Datasets

Instead of relying on a packaged dataset, this example will retrieve data from NWISweb. You must be connected to the Internet in order to replicate the results in this example.

The first step is to retrieve the water-quality and flow data. The water-quality data are retrieved using the `importNWISqw` function, which requires the station identifier and the parameter code. It also accepts beginning and ending dates. The flow data are retrieved using the `readNWIS` function, which requires only the station identifier and also accepts beginning and ending dates as well as other arguments not used. The `renCol` function simply renames the flow column so that it is more readable by humans.

```
> # Load the rloadest package, which requires the USGSwsQW and
> # other packages that contain the necessary functions
> library(dataRetrieval)
> library(rloadest)
> app3.qw <- importNWISqw("01646580", params="00660",
+   begin.date="2001-10-01", end.date="2010-09-30")
> app3.flow <- renameNWISColumns(readNWISdv("01646502", "00060",
+   startDate="2001-10-01", endDate="2010-09-30"))
```

The second step is to merge the flow data with the water-quality data to produce a calibration dataset. The function `mergeQ` extracts the flow data from the flow dataset and merges the daily flow with the sample date in the water-quality dataset. For this analysis, we assume that a sample on any given day represents a valid estimate of the mean daily concentration. It requires that the names of the dates column match between the two datasets; the column `sample_dt` in the water-quality data set is renamed to `Date` to match the date column in the flow dataset. A further requirement of `mergeQ` is that there are no replicate samples taken on the same day. In general, the concentration values with a day agree very well, for this example, simply delete the duplicated days. For other cases, it may be better to compute a mean-daily concentration.

```
> # There are duplicated samples in this dataset. Print them
> subset(app3.qw, sample_dt %in%
+   app3.qw[duplicated(app3.qw$sample_dt), "sample_dt"])
```

	site_no	sample_dt	sample_tm	tzone_cd	medium_cd
67	01646580	2004-07-08	09:45	UTC	WS
68	01646580	2004-07-08	09:50	UTC	WS
105	01646580	2006-03-09	10:30	UTC	WS
106	01646580	2006-03-09	11:30	UTC	WS
153	01646580	2008-02-05	10:15	UTC	WS
154	01646580	2008-02-05	10:20	UTC	WS
156	01646580	2008-03-04	10:15	UTC	WS
157	01646580	2008-03-04	10:20	UTC	WS
159	01646580	2008-04-02	10:15	UTC	WS
160	01646580	2008-04-02	10:20	UTC	WS
162	01646580	2008-05-06	09:15	UTC	WS
163	01646580	2008-05-06	09:20	UTC	WS
165	01646580	2008-05-15	08:45	UTC	WS
166	01646580	2008-05-15	08:50	UTC	WS
168	01646580	2008-06-04	08:45	UTC	WS
169	01646580	2008-06-04	08:50	UTC	WS

170	01646580	2008-06-10	09:45	UTC	WS
171	01646580	2008-06-10	09:50	UTC	WS
173	01646580	2008-07-01	10:15	UTC	WS
174	01646580	2008-07-01	10:20	UTC	WS
177	01646580	2008-08-04	09:15	UTC	WS
178	01646580	2008-08-04	09:20	UTC	WS
183	01646580	2008-10-01	09:15	UTC	WS
184	01646580	2008-10-01	09:20	UTC	WS
186	01646580	2008-12-03	10:15	UTC	WS
187	01646580	2008-12-03	10:20	UTC	WS
189	01646580	2009-01-12	10:45	UTC	WS
190	01646580	2009-01-12	10:50	UTC	WS
191	01646580	2009-03-03	09:45	UTC	WS
192	01646580	2009-03-03	09:50	UTC	WS
193	01646580	2009-04-01	09:15	UTC	WS
194	01646580	2009-04-01	09:20	UTC	WS
196	01646580	2009-05-05	09:45	UTC	WS
197	01646580	2009-05-05	09:50	UTC	WS
199	01646580	2009-05-14	09:45	UTC	WS
200	01646580	2009-05-14	09:50	UTC	WS
202	01646580	2009-06-03	09:45	UTC	WS
203	01646580	2009-06-03	09:50	UTC	WS
204	01646580	2009-06-15	08:45	UTC	WS
205	01646580	2009-06-15	08:50	UTC	WS
207	01646580	2009-07-01	09:15	UTC	WS
208	01646580	2009-07-01	09:20	UTC	WS
210	01646580	2009-08-11	10:45	UTC	WS
211	01646580	2009-08-11	10:50	UTC	WS
215	01646580	2009-10-06	09:45	UTC	WS
216	01646580	2009-10-06	09:50	UTC	WS
219	01646580	2010-03-04	10:15	UTC	WS
220	01646580	2010-03-04	10:25	UTC	WS
221	01646580	2010-04-06	09:15	UTC	WS
222	01646580	2010-04-06	09:25	UTC	WS
223	01646580	2010-05-04	09:30	UTC	WS
224	01646580	2010-05-04	09:40	UTC	WS
225	01646580	2010-05-11	09:15	UTC	WS
226	01646580	2010-05-11	09:25	UTC	WS

#### OrthoPhosphate.P04

67	E0.012
68	E0.009
105	<0.037
106	<0.037
153	E0.017
154	E0.018
156	<0.018
157	<0.018
159	0.021
160	E0.015
162	0.071
163	0.070

165	0.070
166	0.075
168	0.043
169	0.040
170	0.091
171	0.089
173	0.047
174	0.052
177	0.078
178	0.083
183	0.210
184	0.220
186	<0.025
187	<0.025
189	0.044
190	0.055
191	<0.025
192	<0.025
193	E0.020
194	E0.013
196	0.110
197	0.098
199	0.091
200	0.095
202	0.074
203	0.073
204	E0.016
205	E0.017
207	0.035
208	0.035
210	0.120
211	0.120
215	<0.025
216	<0.025
219	0.036
220	0.035
221	0.028
222	0.030
223	<0.025
224	<0.025
225	<0.025
226	<0.025

```

> # Remove the duplicates
> app3.qw <- subset(app3.qw, !duplicated(sample_dt))
> # Now change the date column name and merge
> names(app3.qw)[2] <- "Date"
> # Supress the plot in this merge
> app3.calib <- mergeQ(app3.qw, FLOW="Flow", DATES="Date",
+                      Qdata=app3.flow, Plot=FALSE)

```

## 2 Build the Model

The `loadReg` function is used to build the rating-curve model for constituent load estimation. The basic form of the call to `loadReg` is similar to the call to `lm` in that it requires a formula and data source. The response variable in the formula is the constituent concentration, which is converted to load per day (flux) based on the units of concentration and the units of flow. The `conc.units`, `flow.units`, and `load.units` arguments to `loadReg` define the conversion. For these data, the concentration units (`conc.units`) are "mg/L" (as orthophosphate) and are known within the column so do not need to be specified, the flow units are "cfs" (the default), and the load units for the model are "kilograms." Two additional pieces of information are required for `loadReg`—the names of the flow column and the dates column. A final option, the station identifier, can also be specified.

```
> # Create and print the load model.
> app3.lr <- loadReg(OrthoPhosphate.PO4 ~ model(9), data = app3.calib,
+   flow = "Flow", dates = "Date",
+   station="Potomac River at Chain Bridge, at Washington, DC")
> print(app3.lr)
```

\*\*\* Load Estimation \*\*\*

Station: Potomac River at Chain Bridge, at Washington, DC  
Constituent: OrthoPhosphate.PO4

Number of Observations: 210  
Number of Uncensored Observations: 179  
Center of Decimal Time: 2006.299  
Center of ln(Q): 9.3227  
Period of record: 2001-10-30 to 2010-09-02

Selected Load Model:

-----

OrthoPhosphate.PO4 ~ model(9)

Model coefficients:

	Estimate	Std. Error	z-score	p-value
(Intercept)	7.049643	0.08472	83.2130	0.0000
lnQ	1.462419	0.06069	24.0961	0.0000
lnQ2	-0.007454	0.03762	-0.1981	0.8233
DECTIME	-0.082608	0.02164	-3.8165	0.0001
DECTIME2	0.014610	0.01017	1.4364	0.1403
sin.DECTIME	-0.636680	0.09101	-6.9959	0.0000
cos.DECTIME	-0.296216	0.08009	-3.6985	0.0002

AMLE Regression Statistics

Residual variance: 0.5126

Generalized R-squared: 76.89 percent

G-squared: 307.6 on 6 degrees of freedom

P-value: <0.0001

Prob. Plot Corr. Coeff. (PPCC):

$r = 0.99$   
 $p\text{-value} = 0.0117$   
 Serial Correlation of Residuals: 0.3576

Variance Inflation Factors:

	VIF
lnQ	1.620
lnQ2	1.139
DECTIME	1.045
DECTIME2	1.160
sin.DECTIME	1.539
cos.DECTIME	1.091

Comparison of Observed and Estimated Loads

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Summary Stats: Loads in kg/d

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	Min	25%	50%	75%	90%	95%	Max
Est	137.0	382	1260	3100	10400	17500	96900
Obs	50.8	262	868	3310	9030	17300	66100

Bias Diagnostics

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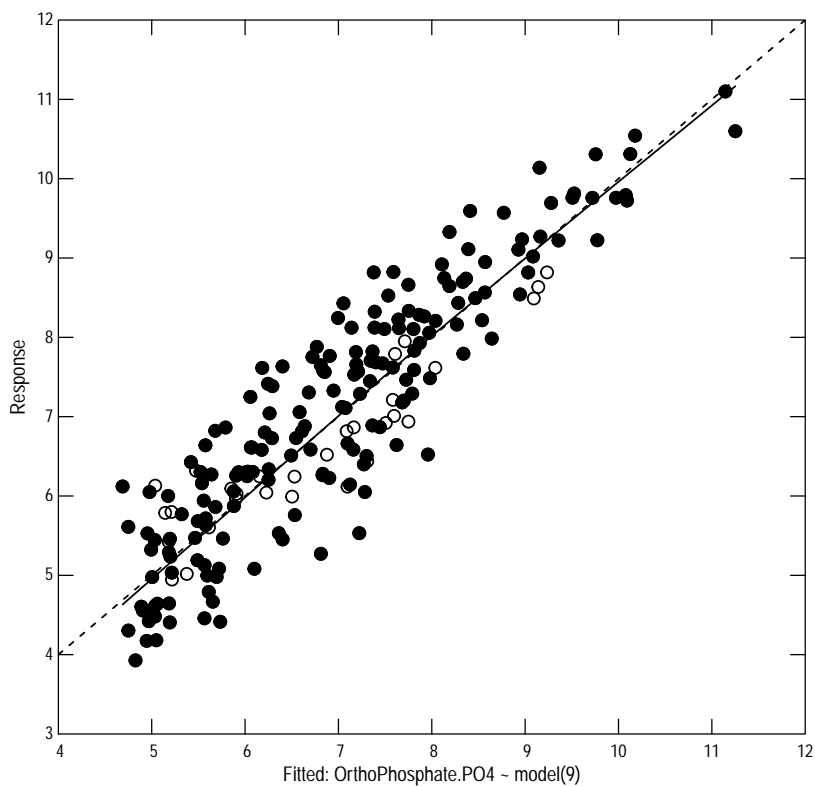
Bp: 19.08 percent  
 PLR: 1.191  
 E: 0.5439

A few details from the printed report deserve mention—the second order flow and decimal time terms have p-values that are greater than 0.05 and may not be necessary; the p-value of the PPCC test is less than 0.05, which suggests a lack of normality; the serial correlation of the residuals is 0.3576, which is quite large; and Bp is relatively large at 19.08.

### 3 Diagnostic Plots

Figure 1 shows the AMLE 1:1 line as a dashed line and the solid line is a LOWESS smooth curve. The LOWESS curve indicates a good fit.

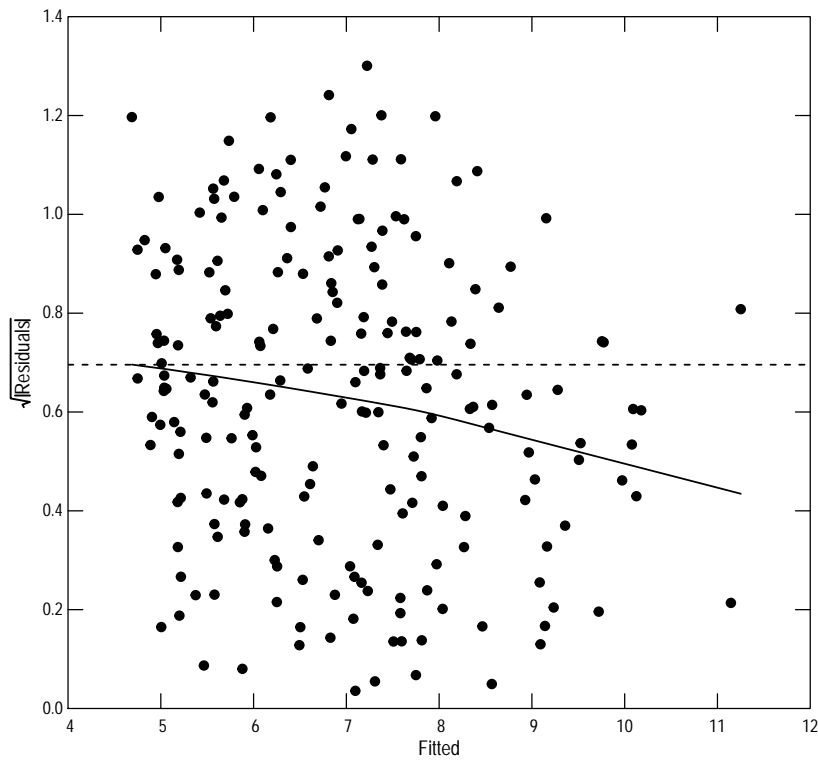
```
> # setSweave is required for the vignette.  
> setSweave("app3_01", 5, 5)  
> plot(app3.lr, which=1, set.up=FALSE)  
> graphics.off()
```



**Figure 1.** The rating-curve regression model.

Figure 2 is a scale-location (S-L) graph that is a useful graph for assessing heteroscedasticity of the residuals. The horizontal dashed line is the expected value of the square root of the absolute value of the residuals and the solid line is the LOWESS smooth. In this case, only 1 of the seven largest residuals is above the expected value line, which suggests in decreasing variance as the estimated load increases.

```
> # setSweave is required for the vignette.
> setSweave("app3_02", 5, 5)
> plot(app3.lm, which=3, set.up=FALSE)
> graphics.off()
```

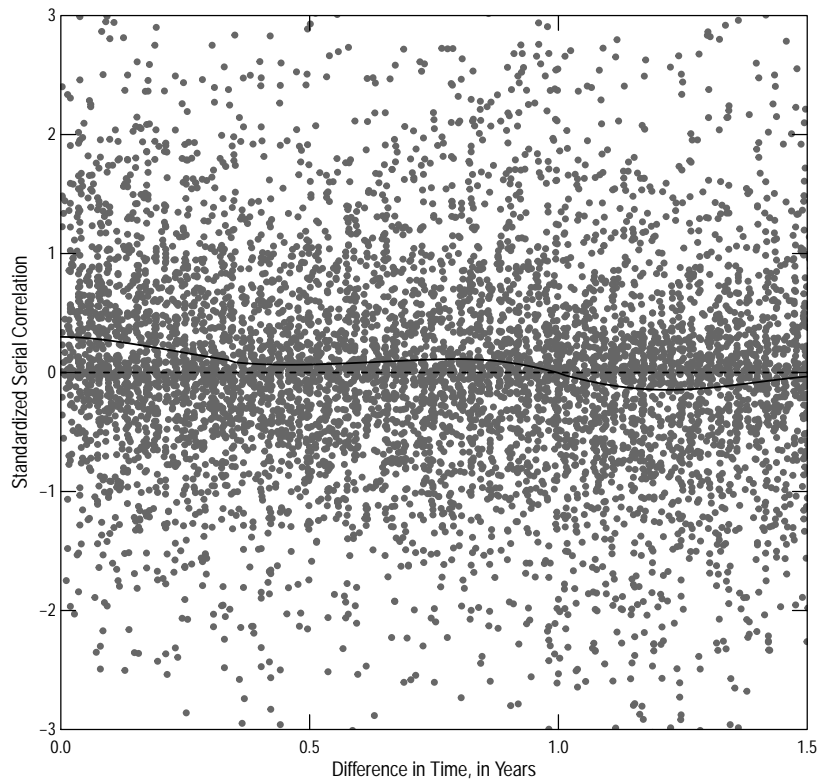


**Figure 2.** The scale-location graph for the regression model.



The correlogram in figure 3 is a adaptation of the correlogram from time-series analysis, which deals with regular samples. The horizontal dashed line is the zero value and the solid line is a kernel smooth rather than a LOWESS line. The kernel smooth gives a better fit in this case. The solid line should be very close to the horizontal line. In this case, there is a suggestion of a long-term lack of fit because the solid line is above the horizontal line for a 1-year lag.

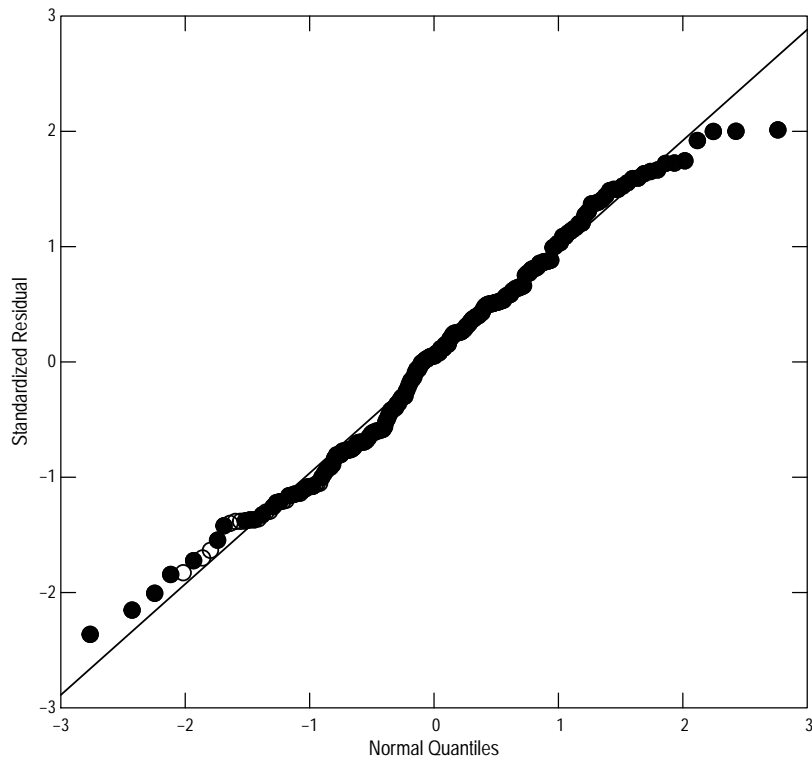
```
> # setSweave is required for the vignette.
> setSweave("app3_03", 5, 5)
> plot(app3.lr, which=4, set.up=FALSE)
> graphics.off()
```



**Figure 3.** The correlogram from the regression model.

Figure 4 shows the q-normal plot of the residuals. The visual appearance of figure 4 confirms the results of the PPCC test in the printed output—the largest residuals trail off the line.

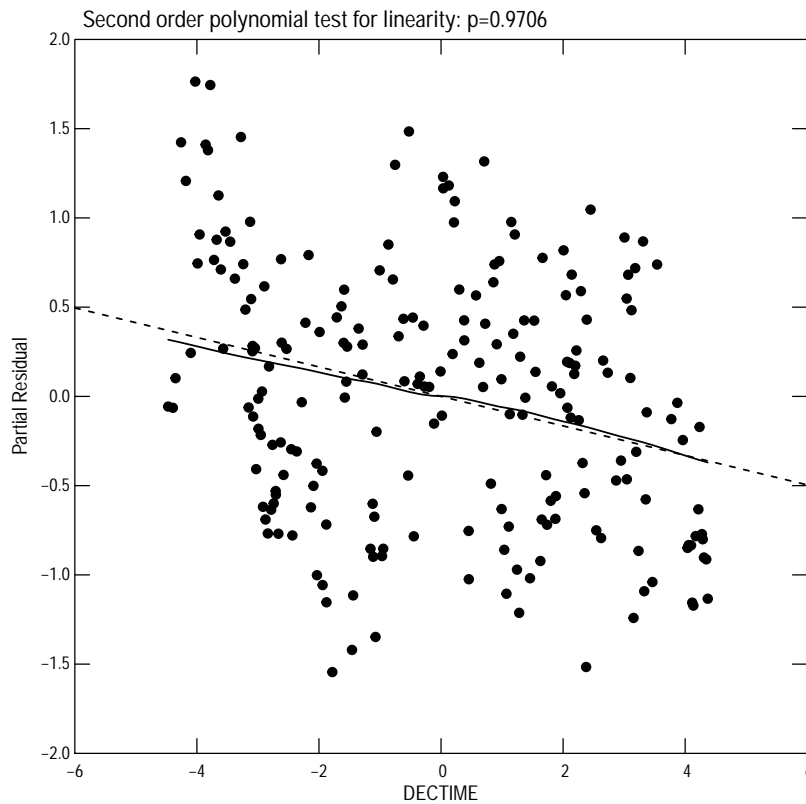
```
> # setSweave is required for the vignette.  
> setSweave("app3_04", 5, 5)  
> plot(app3.lr, which=5, set.up=FALSE)  
> graphics.off()
```



**Figure 4.** The Q-normal plot of the residuals.

Figure 5 shows the partial residual plot for decimal time (DECTIME). This one was selected because of the long-term lack of fit over time suggested by figure 3. The dashed line is the linear fit and the solid line is the LOWESS smooth. In this case, the LOWESS smooth does follow the fitted line, but there is a distinct pattern in the left part of the graph—most of the residuals are above the line up to a DECTIME value of about -3 and then most residuals are below the line to about -1, after which the residuals are fairly well behaved.

```
> # setSweave is required for the vignette.  
> setSweave("app3_05", 5, 5)  
> plot(app3.lm, which="DECTIME", set.up=FALSE)  
> graphics.off()
```

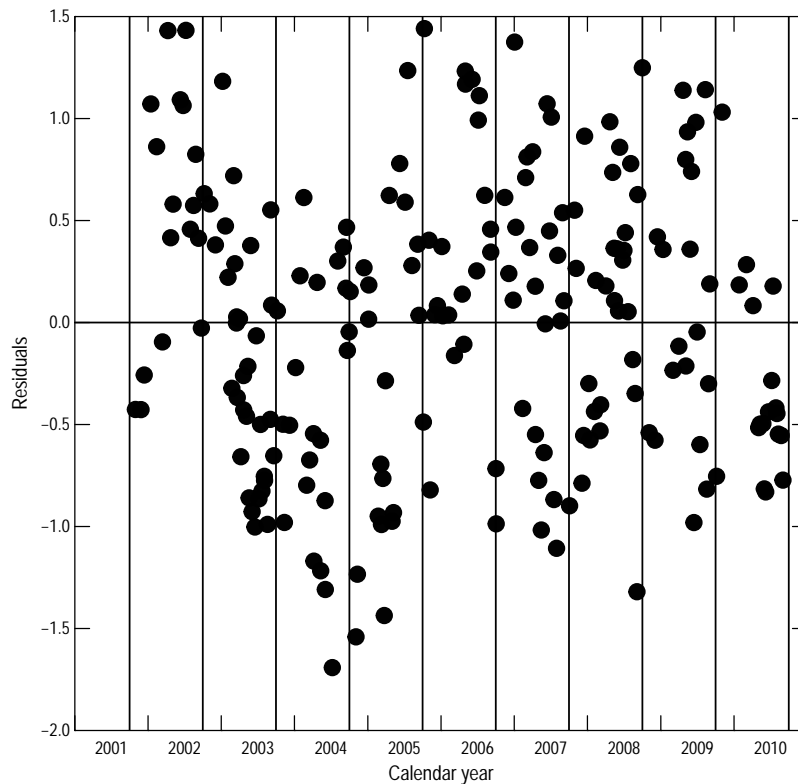


**Figure 5.** The partial residual plot for decimal time.

## 4 Further Diagnostics

Figure 5 suggested a distinct pattern in the residuals in the early part of the record. Figure 6 replots the residuals on a date axis, so that it will be easier to relate to the date. The second call to `refLine` adds vertical lines at the water years.

```
> # setSweave is required for the vignette.  
> setSweave("app3_06", 5, 5)  
> timePlot(app3.calib$Date, residuals(app3.lm),  
+         Plot=list(what="points"), ytitle="Residuals")  
> refLine(horizontal=0)  
> refLine(vertical=as.Date("2001-10-01") + years(0:9))  
> graphics.off()
```



**Figure 6.** Model residuals by date.

The residuals for water-year 2002 are mostly greater than 0; those for water-year 2003 are trending down and those for water-year 2004 are mostly less than 0. This raises the question about whether more persistent flow patterns affect the relation between flow and concentration. The code immediately below computes the average (first line) and water-year average (second line) flow. The pattern of water-year average flows closely matches the pattern of the residuals.

```
> mean(app3.flow$Flow)
```

```
[1] 12656.48
```

```
> with(app3.flow, tapply(Flow, waterYear(Date), mean))
```

2002	2003	2004	2005	2006	2007
4661.712	23446.740	18394.617	12782.082	9337.562	11086.959
2008	2009	2010			
11035.164	9488.548	13663.671			

## 5 Modeling Flow Anomalies

Vecchia and others (2008) describe an approach for breaking down stream flow into what they call anomalies—long- to intermediate-term deviations from average flow and the residual high-frequency variation or daily residuals. That approach can be very useful in cases such as this where there is a strong relation between flow and concentration, but relatively persistent patterns of flow are not captured.

The first step in modeling flow anomalies required retrieving data for a longer period of time. We'll retrieve data from two years prior to the start of the sampling record that we are working with. The additional two years of record were selected because the first one year would be all missing values and one additional year to establish a pattern going into 2001. Then we'll construct a single 1-year anomaly, which seems to make sense from figure 6. This 6-parameter anomaly model is given by:

$$\log(\text{Load}_i) = \alpha_0 + \alpha_1 \text{A1yr}_i + \alpha_2 \text{HFV}_i + \alpha_3 dT_i + \alpha_4 \sin(2\pi dT_i) + \alpha_5 \cos(2\pi dT_i) + \epsilon_i, \quad (2)$$

where  $\text{A1yr}_i$  is the 1-year anomaly of the log of flow,  $\text{HFV}_i$  is remaining high-frequency variation in the log of flow, and  $dT_i$  is the decimal time for observation  $i$ .

```
> app3.anom <- renameNWISColumns(readNWISdv("01646502", "00060",
+   startDate="1999-10-01", endDate="2010-09-30"))
> app3.anom <- cbind(app3.anom, anomalies(log(app3.anom$Flow),
+   a1yr=365))
> # The head would show missing values for a1yr and HFV
> tail(app3.anom)
```

	agency_cd	site_no	Date	Flow	Flow_cd	a1yr
4013	USGS	01646502	2010-09-25	1550	A	0.011527670
4014	USGS	01646502	2010-09-26	1430	A	0.010749096
4015	USGS	01646502	2010-09-27	2020	A	0.009923879
4016	USGS	01646502	2010-09-28	1770	A	0.009010897
4017	USGS	01646502	2010-09-29	1900	A	0.008259015
4018	USGS	01646502	2010-09-30	7080	A	0.010839955
						HFV
4013						-1.58411007
4014						-1.66391198
4015						-1.31766370
4016						-1.44886868
4017						-1.37724246
4018						-0.06440338

The next step is to merge the flow and anomaly data with the water-quality data.

```
> # Suppress the plot in this merge and overwrite app3.calib
> app3.calib <- mergeQ(app3.qw, FLOW=c("Flow", "a1yr", "HFV"),
+   DATES="Date",
+   Qdata=app3.anom, Plot=FALSE)
```

The final step is to construct the model. Note that flow is not a necessary part of the model because it is represented by the anomalies, linear time is represented by `dectime(Date)`, and the seasonal components by `fourier(Date)`. Note also that decimal time is not centered, but could be by using the `center` function.

```
> app3.lra <- loadReg(OrthoPhosphate.P04 ~ a1yr + HFV + dectime(Date)
+                     + fourier(Date),
+                     data = app3.calib,
+                     flow = "Flow", dates = "Date",
+                     station="Potomac River at Chain Bridge, at Washington, DC")
> print(app3.lra)
```

\*\*\* Load Estimation \*\*\*

Station: Potomac River at Chain Bridge, at Washington, DC  
 Constituent: OrthoPhosphate.P04

Number of Observations: 210  
 Number of Uncensored Observations: 179  
 Center of Decimal Time: 2006.299  
 Center of ln(Q): 9.3227  
 Period of record: 2001-10-30 to 2010-09-02

Selected Load Model:

-----

OrthoPhosphate.P04 ~ a1yr + HFV + dectime(Date) + fourier(Date)

Model coefficients:

	Estimate	Std. Error	z-score	p-value
(Intercept)	201.10035	38.05116	5.285	0
a1yr	0.69175	0.11572	5.978	0
HFV	1.54017	0.05256	29.303	0
dectime(Date)	-0.09696	0.01897	-5.112	0
fourier(Date)sin(k=1)	-0.69938	0.07869	-8.888	0
fourier(Date)cos(k=1)	-0.33704	0.07024	-4.798	0

AMLE Regression Statistics

Residual variance: 0.4021

Generalized R-squared: 81.6 percent

G-squared: 355.5 on 5 degrees of freedom

P-value: <0.0001

Prob. Plot Corr. Coeff. (PPCC):

r = 0.9924

p-value = 0.0402

Serial Correlation of Residuals: 0.1914

Variance Inflation Factors:

	VIF
a1yr	1.043
HFV	1.546
dectime(Date)	1.048
fourier(Date)sin(k=1)	1.453
fourier(Date)cos(k=1)	1.082

Comparison of Observed and Estimated Loads

```

-----
      Summary Stats: Loads in kg/d
-----
      Min 25%  50%  75%   90%   95%   Max
Est 104.0 380 1070 2770 10600 25100 75200
Obs  50.8 262  868 3310  9030 17300 66100

Bias Diagnostics
-----
      Bp: 22.63 percent
      PLR: 1.226
      E: 0.6723

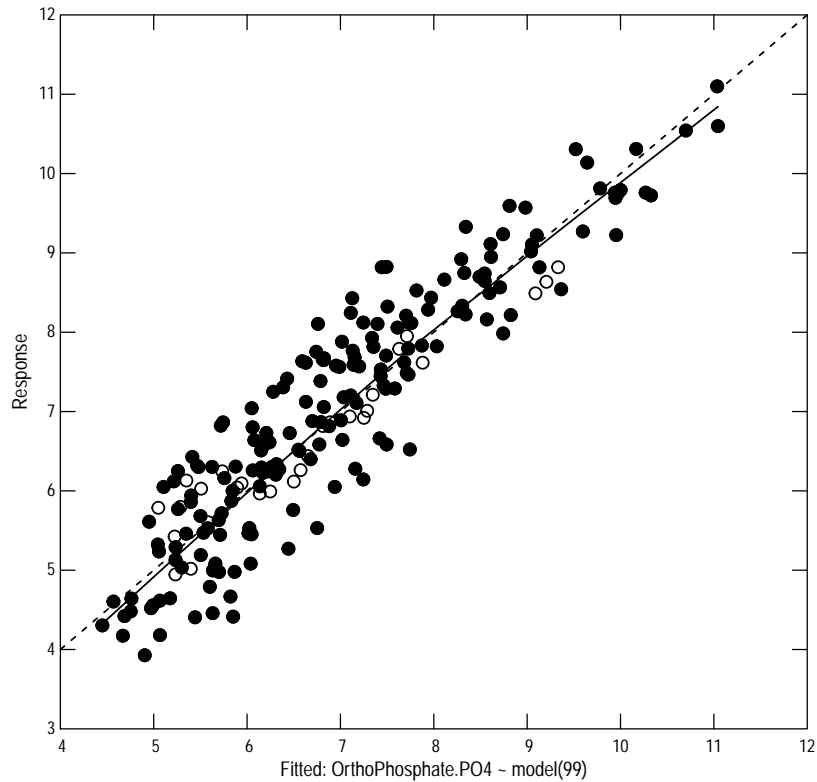
```

The residual variance is much smaller than the original model, 0.4021 rather than 0.5126. The PPCC p-value is still less than 0.05, but much closer to 0.05. The serial correlation of the residuals is much smaller than the original 0.1914 rather than 0.3576. But the Bp statistic is a bit larger 22.63 percent rather than 19.08. In spite of the larger Bp statistic, the Nash-Sutcliffe statistic (E) is larger 0.6723 rather than 0.5439. All of this suggests a better model. Review some of the diagnostic plots.



Figure 7 shows the AMLE 1:1 line as a dashed line and the solid line is a LOWESS smooth curve. The LOWESS curve indicates a good fit.

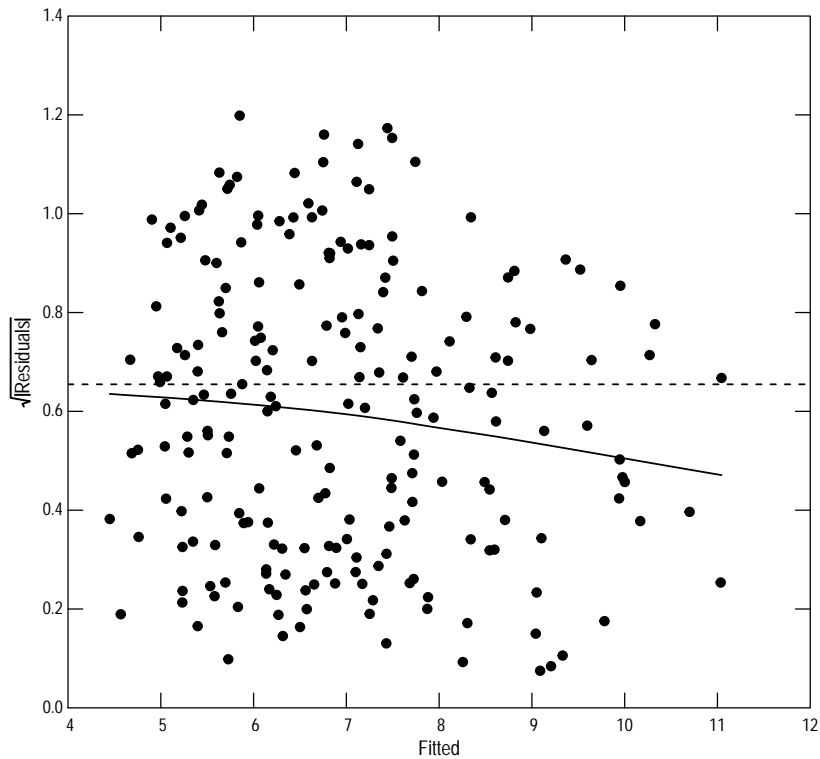
```
> # setSweave is required for the vignette.  
> setSweave("app3_07", 5, 5)  
> plot(app3.lra, which=1, set.up=FALSE)  
> graphics.off()
```



**Figure 7.** The revised rating-curve regression model.

Figure 8 shows the S-L graph, which indicates some decrease in variance for larger fitted values than for smaller.

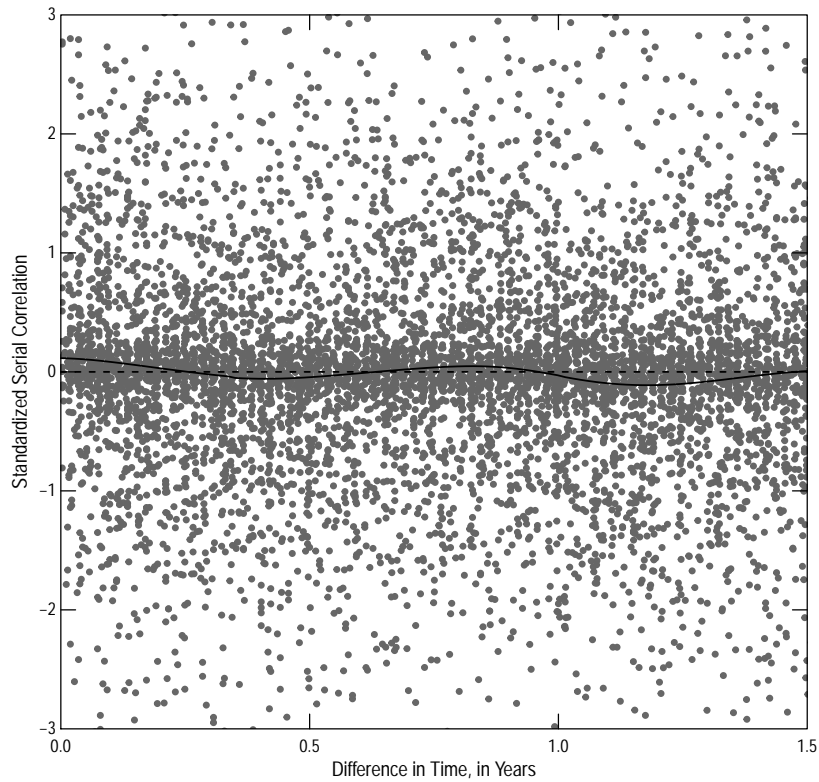
```
> # setSweave is required for the vignette.  
> setSweave("app3_08", 5, 5)  
> plot(app3.lra, which=3, set.up=FALSE)  
> graphics.off()
```



**Figure 8.** The scale-location graph for the revised regression model.

The correlogram in figure 9 shows more variability than one would like, but no distinct long-term or seasonal patterns.

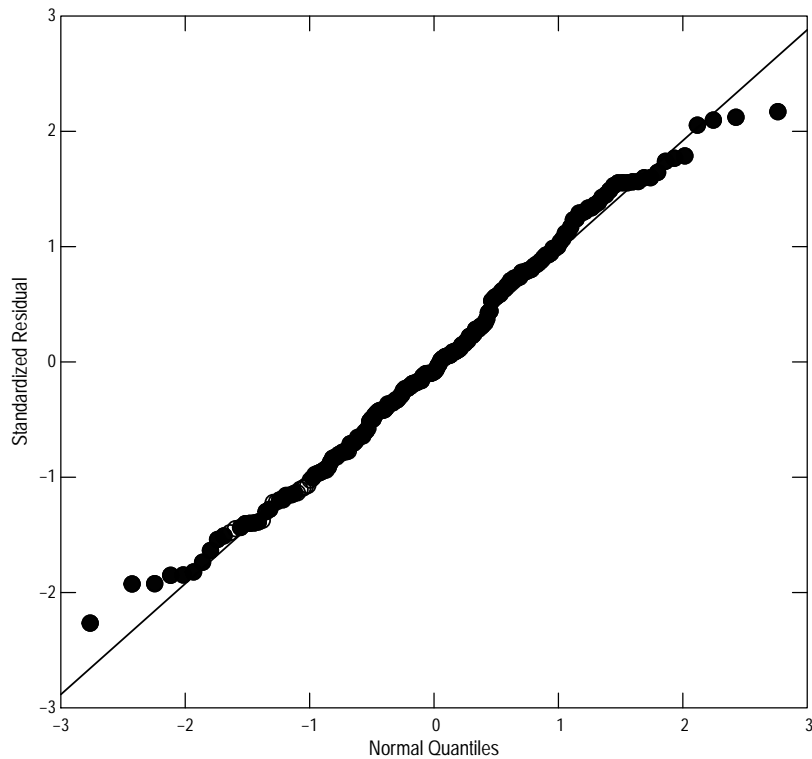
```
> # setSweave is required for the vignette.  
> setSweave("app3_09", 5, 5)  
> plot(app3.lra, which=4, set.up=FALSE)  
> graphics.off()
```



**Figure 9.** The correlogram from the revised regression model.

Figure 10 shows the q-normal plot of the residuals. The largest residuals trail off the line for this analysis but not quite as much as in the original model.

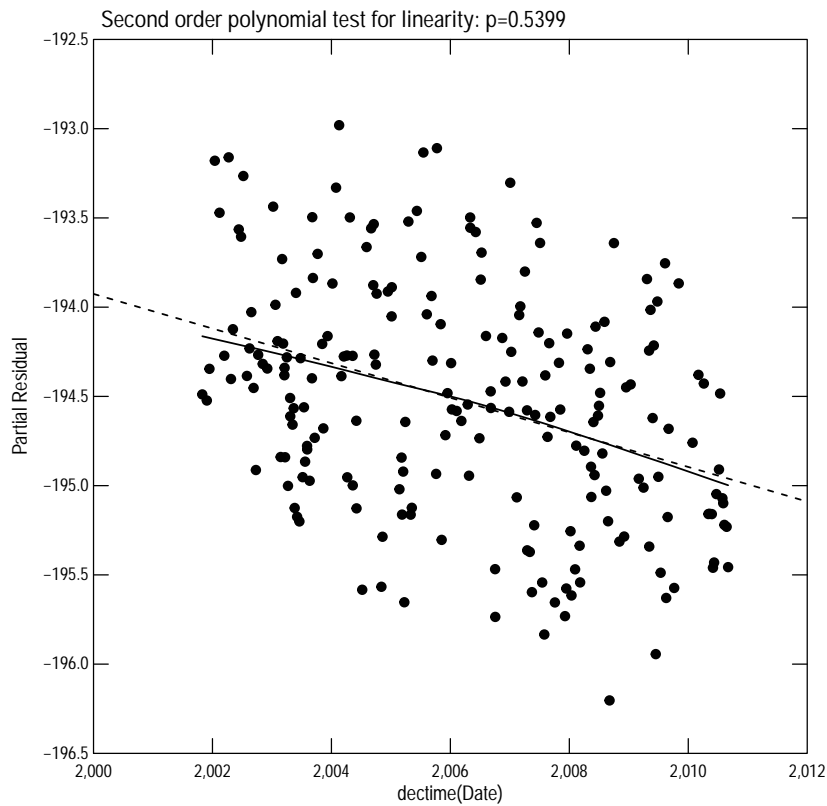
```
> # setSweave is required for the vignette.  
> setSweave("app3_10", 5, 5)  
> plot(app3.lra, which=5, set.up=FALSE)  
> graphics.off()
```



**Figure 10.** The Q-normal plot of the residuals from the revised model.

Figure 11 shows the partial residual plot for decimal time. This one was selected because of the long-term lack of fit over time suggested by figure 3. The dashed line is the linear fit and the solid line is the LOWESS smooth. In this case, the LOWESS smooth does follow the fitted line, but there is a distinct pattern in the left part of the graph—most of the residuals are above the line up to a DECTIME value of about -3 and then most residuals are below the line to about -1, after which the residuals are fairly well behaved.

```
> # setSweave is required for the vignette.
> setSweave("app3_11", 5, 5)
> plot(app3.lra, which="dectime(Date)", set.up=FALSE)
> graphics.off()
```



**Figure 11.** The partial residual plot for decimal time.

## 6 Load Estimates

Because we used anomalies in the regression model, we must be very careful to use the same anomalies in the estimation data. The data that were retrieved to compute the anomalies include dates outside of the calibration period, so must be subsetted to the calibration period. We'll compute load estimates for the water years 2002 through 2010

```
> app3.est <- subset(app3.anom, Date > as.Date("2001-09-30"))
> predLoad(app3.lra, newdata = app3.est, by="water year",
+          print=TRUE)
```

```
-----
Constituent Output File Part IIa: Estimation (test for extrapolation)
Load Estimates for 2001-10-01 to 2010-09-30
-----
```

```
Streamflow Summary Statistics
-----
```

```
WARNING: The maximum estimation data set steamflow exceeds the maximum
calibration data set steamflow. Load estimates require extrapolation.
```

```
-----
Constituent Output File Part IIb: Estimation (Load Estimates)
Load Estimates for 2001-10-01 to 2010-09-30
-----
```

```
Flux Estimates, in kg/d, using AMLE
-----
```

	Period	Ndays	Flux	Std.Err	SEP	L95	U95
1	WY 2002	365	1220.648	164.5140	193.7133	884.9994	1642.238
2	WY 2003	365	8037.721	837.9254	999.7867	6256.4054	10168.877
3	WY 2004	366	2671.495	315.5778	403.1849	1968.4296	3544.931
4	WY 2005	365	1591.845	124.5184	175.5853	1275.4593	1962.831
5	WY 2006	365	1570.207	101.3389	179.1237	1248.4155	1949.573
6	WY 2007	365	1537.921	111.4383	165.1147	1239.7052	1886.133
7	WY 2008	366	1786.773	163.4341	238.2895	1365.2750	2297.535
8	WY 2009	365	1214.415	108.0001	156.0877	937.2396	1547.987
9	WY 2010	365	1423.184	169.7617	220.0813	1040.5768	1901.012