

## Tutorial 5: Linear Classification

5., prior:  $P(Y=c) = \pi_c$

Class cond.:  $P(x|Y=c, \theta_c)$

→ Data: i.i.d

Likelihood function:

$$P(D|\xi\pi_c, \theta_c) = \prod_n \prod_c [P(x^{(n)}|\theta_c) \cdot \pi_c]^{\gamma_c^{(n)}}$$

$$\log P(D|\xi\pi_c, \theta_c) = \sum_n \sum_c [\underbrace{\gamma_c^{(n)} \log P(x^{(n)}|\theta_c)}_{\text{const in } \pi_c} + \gamma_c^{(n)} \cdot \log \pi_c]$$

$$\sum_c \pi_c = 1 \Rightarrow \sum_c \pi_c - 1 = 0$$

$$\mathcal{L} = \sum_n \sum_c [\gamma_c^{(n)} \cdot P(x^{(n)}|\theta_c) + \gamma_c^{(n)} \cdot \log \pi_c] - \lambda (\sum_c \pi_c - 1)$$

$$\frac{d\mathcal{L}}{d\pi_c} = \sum_n \underbrace{\gamma_c^{(n)} \cdot \frac{1}{\pi_c}}_{N_c} - \lambda$$

$$= N_c \cdot \frac{1}{\pi_c} - \lambda = 0$$

$$\pi_c = \frac{N_c}{\lambda} = \underline{\underline{\frac{N_c}{N}}}$$

$$\frac{d\mathcal{L}}{d\lambda} = - \sum_c \pi_c + 1 \stackrel{!}{=} 0$$

$$\text{use } \pi_c = \frac{N_c}{\lambda} \rightarrow \sum_c \pi_c = 1$$

$$\sum_c \frac{N_c}{\lambda} = 1 \Rightarrow \lambda = \sum_c N_c = n$$

$$6.) P(x | y=c, \theta_c) = N(x | \mu_c, \Sigma)$$

$$= \frac{1}{(2\pi)^D/2 |\Sigma|^{1/2}} \exp \left[ -\frac{1}{2} (x - \mu_c)^T \Sigma^{-1} (x - \mu_c) \right]$$

$$\log P(D | \{\pi_c, \theta_c\}) = \sum_n \sum_c [y_c^{(n)} \log p(x^{(n)} | \theta_c) + y_c^{(n)} \log \pi_c]$$

$$= \sum_n \sum_c y_c^{(n)} \left[ -\frac{D}{2} \log (2\pi) - \frac{1}{2} \log |\Sigma| \right.$$

$$\quad \quad \quad \left. - \frac{1}{2} (x^{(n)} - \mu_c)^T \Sigma^{-1} (x^{(n)} - \mu_c) + \log \pi_c \right]$$

$$= \sum_n \sum_c \left[ -\frac{1}{2} y_c^{(n)} [D \log (2\pi) + \log |\Sigma|] - 2 \log \pi \right.$$

$$\quad \quad \quad \left. + x^{(n)T} \Sigma^{-1} x^{(n)} - 2 \mu_c^T \Sigma^{-1} x^{(n)} + \mu_c^T \Sigma^{-1} \mu_c \right]$$

Derivative with respect to  $\mu_c$ :

$$\sum_n -\frac{1}{2} y_c^{(n)} \left[ -2 \Sigma^{-1} x^{(n)} - 2 \Sigma^{-1} \mu_c \right]$$

$$= \Sigma^{-1} \left[ \sum_n y_c^{(n)} x^{(n)} - N_c \mu_c \right] \stackrel{!}{=} 0$$

$$\mu_c = \frac{1}{N_c} \overline{\sum_n y_c^{(n)} x^{(n)}}$$

$$= \underline{\underline{\frac{1}{N_c} \sum_{n|x^{(n)} \in C_c} x^{(n)}}}$$

6.) Rewrite

$$\log P(D | \Sigma_{\text{IC}}, \theta_c \beta_c) = \sum_n \sum_c -\frac{1}{2} \gamma_c^{(n)} \left[ D \log(2\pi) - 2 \log \pi_c + \log |\Sigma| + \underbrace{(x^{(n)} - \mu_c)^T \Sigma^{-1} (x^{(n)} - \mu_c)}_{\text{scalar}} \right]$$

$$= \sum_n \sum_c -\frac{1}{2} \gamma_c^{(n)} \left[ D \log(2\pi) - 2 \log \pi_c + \log |\Sigma| + \text{tr}(\Sigma^{-1} (x^{(n)} - \mu_c)(x^{(n)} - \mu_c)^T) \right]$$

Derivative w.r.t.  $\Sigma^{-1}$

$$\begin{aligned} \frac{\partial}{\partial \Sigma} \text{tr}(AB) &= B^T \\ \frac{d}{dA} \log |A| &= (A^{-1})^T \\ \log |A^{-1}| &= -\log |A| \end{aligned} \quad \left. \begin{aligned} \log |\Sigma| &= -\log |\Sigma^{-1}| \\ \log |\Sigma^{-1}| &= -\log |\Sigma| \end{aligned} \right\}$$

$$\rightarrow \sum_n \sum_c -\frac{1}{2} \gamma_c^{(n)} \left[ -\Sigma + \sum_n \sum_c \gamma_c^{(n)} (x^{(n)} - \mu_c)(x^{(n)} - \mu_c)^T \right] = 0$$

$$\begin{aligned} \rightarrow \Sigma &= \frac{1}{N} \sum_n \sum_c \gamma_c^{(n)} (x^{(n)} - \mu_c)(x^{(n)} - \mu_c)^T // S_c = \frac{1}{N_c} \sum_n \gamma_c^{(n)} (x^{(n)} - \mu_c)(x^{(n)} - \mu_c)^T \\ &= \sum_c \frac{N_c}{N} S_c \end{aligned}$$

$$1., \text{ prior } \hat{p}(Y=0) = \frac{1}{2}$$

$$\hat{p}(Y=1) = \frac{1}{2}$$

$$P(x|Y=0) = \lambda_0 e^{-\lambda_0 x}$$

$$P(x|Y=1) = \lambda_1 e^{-\lambda_1 x}$$

$$\rightarrow P(Y|x)$$

$$Y \in \{0, 1\}$$

$$P(Y|x) = \begin{cases} .. & \text{if } Y=0, 1 \\ 0 & \text{else} \end{cases}$$

$$\rightarrow \text{Bernoulli: } q^Y(1-q)^{1-Y}$$

$$Y_{\text{pred}} = \arg \max_k p(Y=k|x)$$

$$P(Y=1|x) > P(Y=0|x)$$

$$\Leftrightarrow \frac{P(Y=1|x)}{P(Y=0|x)} > 1$$

$$\Leftrightarrow \ln \frac{P(Y=1|x)}{P(Y=0|x)} > 0$$

$$\ln \frac{P(Y=1|x)}{P(Y=0|x)}$$

$$= \ln \frac{P(x|Y=1) \cdot P(Y=1) - \bar{P}(x)}{P(x) \cdot P(x|Y=0) P(Y=0)}$$

$$= \ln \frac{\lambda_1 \exp[-\lambda_1 x] \cdot \frac{1}{2}}{\lambda_0 \exp[-\lambda_0 x] \cdot \frac{1}{2}}$$

$$= \ln \frac{\lambda_1}{\lambda_0} + (\lambda_0 - \lambda_1)x > 0$$

$$(\lambda_0 - \lambda_1)x > \ln \lambda_0 - \ln \lambda_1$$

Case 1:  $(\lambda_0 - \lambda_1) > 0 \Leftrightarrow \lambda_0 > \lambda_1$

$$x > \frac{\ln \lambda_0 - \ln \lambda_1}{\lambda_0 - \lambda_1}$$

Case 2:  $(\lambda_0 - \lambda_1) < 0 \Leftrightarrow \lambda_0 < \lambda_1$

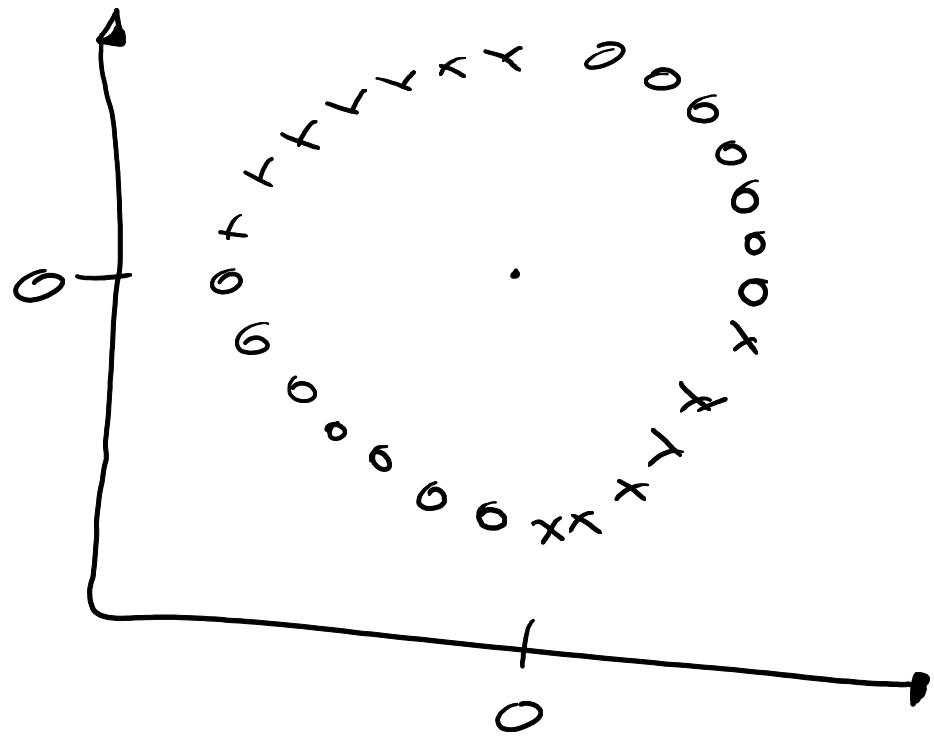
$$x < \frac{\ln \lambda_0 - \ln \lambda_1}{\lambda_0 - \lambda_1}$$

$x \in \left\{ \begin{array}{l} \left( \frac{\ln \lambda_0 - \ln \lambda_1}{\lambda_0 - \lambda_1}, \infty \right) \text{ if } \lambda_0 > \lambda_1 \\ [0, \frac{\ln \lambda_0 - \ln \lambda_1}{\lambda_0 - \lambda_1}) \text{ if } \lambda_0 < \lambda_1 \end{array} \right.$

- 2.)
- separating hyperplane  
→ all points on correct side
  - magnitude of  $w \rightarrow \infty$
  - prevent: regularization

$$\begin{aligned}
 3.) P(y=k|x) &= \frac{e^{w_k^T x}}{\sum_i e^{w_i^T x}} \quad (\text{softmax}) \\
 &= \frac{e^{xP(w_1^T x)}}{e^{xP(w_1^T x)} + e^{xP(w_0^T x)}} \\
 &= \frac{1}{1 + \frac{e^{xP(w_0^T x)}}{e^{xP(w_1^T x)}}} \\
 &= \frac{1}{1 + e^{xP[-(w_1 - w_0)^T x]}} \quad || \widehat{w} = w_1 - w_0 \\
 &= \frac{1}{1 + e^{xP(-\widehat{w}^T x)}} \quad \text{Sigmoid} \\
 &= \sigma(\widehat{w}^T x)
 \end{aligned}$$

4.)



$$\phi(x_1, x_2) = x_1 x_2$$