Data Analysis: Machine Learning and Regression

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August Review
UChicago Financial Mathematics

Model selection

Regularized regression

Principal Components

Boosting and Bagging

In-sample

There is a tendency to over-parameterize models to make them fit the sample data too well.

- ► Are we fitting the sample-specific noise?
- Parameterizing in-sample noise leads to bad out-of-sample (OOS) performance.

R-squared will show fit always improves w/ more parameters.

- ► Adj R-squared popular, but not model specific.
- ▶ t-stat dependent on other regressors included.

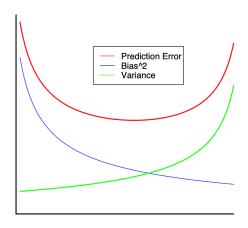
Bias-variance tradeoff

Most of our machine-learning techniques are dealing with this tradeoff.

$$\underbrace{\operatorname{var}\left[(\boldsymbol{Y}-\boldsymbol{X}\boldsymbol{b})^{2}\right]}_{\text{data variance}} = \sigma_{\epsilon}^{2} + \underbrace{\left(\boldsymbol{X}\left(\mathbb{E}\left[\boldsymbol{b}\right]-\boldsymbol{\beta}\right)\right)^{2}}_{\text{bias squared}} + \underbrace{\mathbb{E}\left[\left(\boldsymbol{X}\left(\boldsymbol{b}-\mathbb{E}\left[\boldsymbol{b}\right]\right)\right)^{2}\right]}_{\text{estimator variance}}$$

Where all moments are conditional on **X**.

Bias-Variance Tradeoff



Squared Error

Model Complexity

Out-of-sample (OOS) fit judges the model by sample points excluded from estimation.

- ► Suppose you have *N* points in your sample.
- \blacktriangleright You estimate **b** from the first *n* points.

Define,

$$\mathcal{L}^{IS}(\boldsymbol{b}) = \sum_{i=1}^{n} (y_i - \boldsymbol{x}_i' \boldsymbol{b})^2$$

$$\mathcal{L}^{OOS}(\boldsymbol{b}) = \sum_{i=n+1}^{N} (y_i - \boldsymbol{x}_i' \boldsymbol{b})^2$$

Model selection

The OOS-R-squared, \mathcal{R}_{oos}^2 is the R-squared calculation based on the IS estimation of the model, applied to OOS data.

$$\mathcal{R}_{oos}^2 = 1 - rac{\mathcal{L}^{OOS}\left(oldsymbol{b}
ight)}{\mathcal{L}^{OOS}\left(oldsymbol{b}=0
ight)}$$

where $\mathcal{L}^{OOS}\left(\boldsymbol{b}=0\right)$ is simply the null model. For instance, our usual \mathcal{R}^{2} can be written as¹,

$$\mathcal{R}_{is}^{2} = 1 - rac{\mathcal{L}^{IS}(m{b})}{\mathcal{L}^{IS}(m{b} = 0)} = 1 - rac{\sum_{i=1}^{n} e_{i}^{2}}{\sum_{i=1}^{n} (y_{i} - \bar{y})^{2}}$$

¹This assumes the regression model has an intercept.

If the training data is not representative, it may not give optimal out-of-sample performance.

$$\mathcal{R}_{is}^{2} = 1 - rac{\mathcal{L}^{IS}(\mathbf{b})}{\mathcal{L}^{IS}(\mathbf{b} = 0)} = 1 - rac{\sum_{i=1}^{n} e_{i}^{2}}{\sum_{i=1}^{n} (y_{i} - \bar{y})^{2}}$$

$$\mathcal{R}_{oos}^2 = 1 - rac{\mathcal{L}^{OOS}\left(oldsymbol{b}
ight)}{\mathcal{L}^{OOS}\left(oldsymbol{b}=0
ight)}$$

Selecting ${m b}$ to maximize ${\cal R}^2_{is}$ does not mean it will maximize ${\cal R}^2_{oos}$.

Interpreting \mathcal{R}_{oos}^2

OOS R-squared is much different than the usual in-sample stat.

- $m{\mathcal{R}}^2_{\mathit{is}} \in [0,1]$ for models with an intercept (or de-meaned data.)
- $ightharpoonup \mathcal{R}^2_{oos}$ can be (and often is) negative, if the model does worse than no model.
- ▶ Models with higher \mathcal{R}_{is}^2 often have lower \mathcal{R}_{oos}^2 .

IS estimation by construction improves IS R-squared, but if it is fitting IS noise, then OOS R-squared gets worse.

How do we get the OOS data?

- One can hold out the final portion of the sample, but then results are impacted by specifics of this one period.
- ► A more general approach is to use K-folds, which just means multiple subsamples.

K-Folds

- ▶ Randomly split the data into *K* subsamples.
- ► For each subsample, estimate the model excluding it and treating it as the OOS data.
- Using completely randomized subsamples can be a problem if the data has serial correlation. In that case, one could use sequential subsamples.

Model selection

Regularized regression

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OLS

OLS solves

$$oldsymbol{b} = \arg\min_{oldsymbol{b}_o} \sum_{i=1}^n (y_i - oldsymbol{x}_i' oldsymbol{b}_o)^2$$

- ► This is an unconstrained optimization.
- **b** minimizes the sum of squared errors.

How would you solve a regression model where the vector, \boldsymbol{b} , must satisfy an equality constraint?

- We could add constraints explicitly.
- ▶ Or we could add a penalty to the objective function.

Regularized regression

Suppose we have k regressors,

$$\boldsymbol{b} = \arg\min_{\boldsymbol{b}_o} \sum_{i=1}^{n} (y_i - \boldsymbol{x}_i' \boldsymbol{b}_o)^2 + \lambda \sum_{i=1}^{k} c(b_o^i)$$

where

- $lackbox{b}_o^i$ is the *i*-th element of the candidate solution vector, $oldsymbol{b}_o$.
- $ightharpoonup c(\cdot)$ is the regularizing function for non-zero elements of $m{b}$.
- ▶ We assume c is a non-negative function with c(0) = 0.
- \blacktriangleright λ is a data-specific parameter.

Question

▶ If we believe the data are explained by a linear (in β) model,

$$y = \alpha + \mathbf{x}'\mathbf{\beta} + \epsilon$$

then how do we justify estimating it with the regularized, (constrained,) model?

- Suppose the classic OLS assumptions hold. Is there any justification for using regularized regressions?
- ▶ Is the regularized/constrained estimator biased/consistent? Does the regularized/constrained estimator reduce estimator variance?

Regularizing functions

There are several popular regularizing functions.

Stepwise	L ₀	$c(\boldsymbol{b}) = \mathbb{1}_{\boldsymbol{b} \neq 0}$
LASSO	L_1	$c(\boldsymbol{b}) = \boldsymbol{b} $
Ridge	L_2	$c(\mathbf{b}) = \mathbf{b} ^2$
Elastic Net		$c(\boldsymbol{b}) = \alpha \boldsymbol{b} + \frac{1-\alpha}{2} \boldsymbol{b} ^2$
Log		$c(m{b}) = \log(1 + m{b})$

Figure: Source: Taddy, 2019

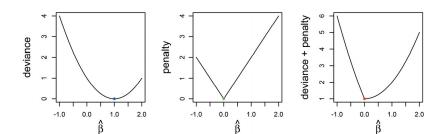


Figure: Source: Taddy, 2019

- ▶ Both Ridge and LASSO are in this family of estimators. What is the difference in their specification?
- ▶ Which estimator leads to sparse models?

Choosing the regularizing function

Choose based on what you want to penalize.

- LASSO produces many 0 values.
- Ridge allows many non-zero values but severely punishes large estimated effects.
- Elastic net is a mix.
- ► Log penalizes strongly for non-zero elements, but penalizes weakly for large values.

Question

▶ Under what circumstances might Ridge be useful?

$$oldsymbol{b}^{ ext{ridge}} = ig(oldsymbol{\mathcal{X}}'oldsymbol{\mathcal{X}} + \lambda\,\mathcal{I}_kig)^{-1}ig(oldsymbol{\mathcal{X}}'oldsymbol{Y}ig)$$

- Deals with multicolinearity and ill-conditioning.
- Often is derived from a Bayesian model, related to Black-Litterman.
- ▶ Will not reduce dimensionality—non-zero components will still be non-zero.

▶ Under what circumstances might LASSO be useful?

- ► The constraint does not vanish for small components of **b**, so it forces components to zero.
- ► Essentially LASSO docks each component of **b** by a fixed amount, so it still allows very large values.
- ► LASSO is the most widespread regularized regression.

Nonlinearity

With regularized regression, we lose linearity.

- Scaling of X now matters, because it can't just be offset by rescaling β!
- ightharpoonup Try using standard-deviation deflated X.
- ▶ Or change regularization function to $\tilde{c}(b^i) = \sigma^i c(b^i)^2$

This re-scaling ensures that the regularization applies less to X data that varies little, and thus requires larger β .

²Many packages have a setting to do this automatically.

Tuning parameters

The parameter λ is estimated from the data.

- Find the minimum λ such that the estimate is regularized to 0: $\mathbf{b} = 0$.
- Estimate the model for a sequence of diminishing λ until getting to $\lambda=0$, which produces the OLS estimates.

To choose among λ , use

- Information criteria: Akaike (AIC), Bayesian (BIC), etc.
- Cross validation on K-folds.

K-Fold Cross-validation

Use Cross-Validation (CV) on K-folds.

- Split the data into K random subsets.
- ▶ For k = 1...K.
- Use all data except subsample k to estimate the model.
- Calculate the OOS errors using subsample k.

Regularized regression w/ CV

Use K-fold CV to optimize λ .

- \triangleright Select a sequence of λ_i .
- Divide into K-folds.
- \blacktriangleright For each λ_i , estimate the model using K-fold CV to obtain K sets of OOS loss, (such as sum of squared errors.)
- \triangleright Select the λ_i that minimizes the simple average of the (K sets of) OOS loss values.

Outline

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Dimension reduction

Suppose we have a $k \times 1$ vector of data \mathbf{x}_t , with covariance matrix, Σ . Suppose k is large.

- We may seek a smaller set of variables p << k that explains most of the variation explained by x.
- ► Consider these *p* factors, $z^1, z^2, ..., z^p$.

Find the vector, \mathbf{w}^1 that maximizes

$$s.t. \mathbf{w}' \mathbf{w} = 1$$

This says \mathbf{w}^1 is a linear combination of \mathbf{x}_t that maximizes the variation explained, rescaled.

Iterating

The full set of factors, \mathbf{w}^i could then be developed iteratively...

$$arg \max \mathbf{w}' \mathbf{\Sigma} \mathbf{w}$$
 $s.t. (\mathbf{w}^i)' \mathbf{w}^j = 0, i \neq j$ $(\mathbf{w}^i)' \mathbf{w}^j = 1, i = j$

We can iterate to obtain K vectors which will exactly span the space spanned by \mathbf{x}_{t} .

- ightharpoonup Or we can stop at p, obtaining a reduced dimensionality.
- This dimensionality will maximally explain the space spanned by x_t.
- ► The new space has an orthogonal basis.

The principal components are the factors constructed,

$$\boldsymbol{z}_{t}^{i}=\left(\boldsymbol{w}^{i}\right)\boldsymbol{x}_{t}$$

We do not need to solve the optimization above!

- ightharpoonup One can show that \mathbf{w}^i is the i-th eigenvector of Σ .
- ► The i-th eigenvalue gives the total variation explained by the i-th vector.

Thus, PCA is obtained via eigenvector decomposition!

- ▶ We have one time-series to explain, **Y**.
- ▶ There are 10 explanatory time-series, collected in **X**.
- We calculate the first three Principal Components, Z, using standard methods.

Do these principal components maximize the R-squared relative to any other 3-factors extracted from \boldsymbol{X} ?

► Suppose a bond trader wishes to trade only the 10 most important bonds out of a set of 100 bonds. Will PCA be useful for this selection?

- ▶ PCA reduces the dimensionality—but in a rotated vector space.
- ► In terms of the original space, any one PC is a combination of every original variable.
- Use LASSO or other regularized regressions to get sparsity.
- ► PCA is useful for state-space reduction for statistical and mathematical reasons—not for interpretability.

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When deciding on the number of regressors, it is common to see the following:

- ► Include all regressors under consideration.
- ► Check the t-stats (p-values) of each regressor.
- ► Re-run the regression on only the statistically significant regressors.

This is known as backward stepwise regression. (BSR).

Though common, BSR should be avoided.

- ► If regressors are strongly correlated, neither may show significance even though one is.
- ► The p-values are being generated on an overfit model and may have substantial small-sample bias.

Forward Step Regression

Forward Step Regression (FSR) builds up the model instead of cutting it down.

- ► Suppose you have *J* regressors under consideration.
- ightharpoonup Fit a univariate regression on each of the J regressors.
- ▶ Choose the regressor, (call it j,) with the highest in-sample fit (R-squared?).
- ► Estimate all bivariate regressions that include regressor *j*.
- ► Continue until hitting a limit of regressors, a threshold R-squared, or using a selection rule such as AIC/BIC.

Marginal regression

Marginal regression returns a principal factor based on covariation in x, but also impact on y.

For each component of x, (column of X,) run a univariate regression, (assume we have de-meaned the data,)

$$y = \phi^j x^j + u$$

- ▶ Stack up these univariate regression coefficients, ϕ^j into a vector, ϕ .
- Calculate the factor,

$$z^1 = x' \phi$$

▶ Use z^1 to analyze/predict y via OLS:

$$y = z^1 \beta^1 + \epsilon$$

Partial Least Squares

Instead of stopping and analyzing y given z^1 , repeat the marginal regression.

- Namely, take the sample residuals e from the regression of y on z^1 . Do a marginal regression on them to obtain z^2 .
- Use z^1 and z^2 to analyze/predict y with OLS.
- ► Keep repeating for more factors.

- ► What would it mean to apply "boosting" to marginal regression?
- ▶ In the bias-variance tradeoff, what is boosting meant to help?

Boosting

Boosting refers to iteratively using a (typically simple model.)

- ► For instance, we can "boost" the marginal regression model by re-running it on the residuals from the final predictions.
- ► Thus, boosting builds new predictions in a sequence from simpler predictions.
- PLS is simply "boosted" MLS.
- Boosting is useful if we have a model with low variance but substantial bias.

Classification and Regression Trees (CART) are a supervised learning tool.

- ► For **x** and y, the CART groups y according to thresholds of the associated x data.
- ▶ There are an impossible number of potential trees / splits.
- ► CART starts at the top and seeks the component of *x*, and its threshold value, that minimizes,

$$\sum_{i \in left} \left(y^i - \bar{y}^{left} \right)^2 + \sum_{i \in right} \left(y^i - \bar{y}^{right} \right)^2$$

► CART can model complex relationships but is prone to overfitting.

Question

- ▶ How do we get random forests from regression trees?
- ► Why are random forests useful?
- Does a forest improve the bias or the variance of an individual tree?

Bagging

Bootstrap aggregating, or "Bagging", is useful when an estimator is unbiased but highly variable.

- Run the model over many bootstrapped samples.
- ▶ This yields many observations from the variable, unbiased estimator distribution.
- Take the average, and you have a less variable, unbiased estimator.

Supervised vs unsupervised

- Regression models are supervised. They model x based on feedback from associated y data.
- ▶ PCA, k-means clustering, etc. are unsupervised. They organize x without any feedback from v.
- ▶ Partial Least Squares (PLS) takes a PCA-type factorization of x based on information in y.