TOWARDS A FRAMEWORK FOR OPERATIONAL RISK IN THE BANKING SECTOR

by

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of

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in

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ABSTRACT

Towards a Framework for Operational Risk in the Banking Sector

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Researchers in health and prevention are increasingly interested in, not only if one variable affects another, but how the effect takes place. Mediation analysis is built for this very question. However, it lacks interpretable effect size estimates in situations where the mediator (intermediate variable) and/or the outcome is categorical or otherwise far from normally distributed. By integrating a powerful approach known as average marginal effects within mediation analysis—termed Marginal Mediation Analysis (MMA)—the issues regarding categorical mediators and/or outcomes is, in large part, resolved. The project discussed herein presents the development, evaluation, and application of this novel approach to understand its utility to researchers. The aims of this project were four-fold: 1) develop MMA and the software to perform it, 2) test its accuracy, robustness, and coverage via Monte Carlo simulations and develop guidelines for MMAs use, and 4) apply it to real prevention science data.

First, MMA is built on generalized linear modeling and is available in R as the Marginal Mediation package. Second, Monte Carlo simulations were used to Abstract

assess the method's accuracy, statistical power, confidence interval coverage, and ability to have the indirect plus direct effect equal the total effect. Results demonstrated accurate estimation of the effects and relatively accurate confidence interval coverage, although in some situations the confidence interval was too narrow. Further, sample sizes needed for ample statistical power ranged from 50 (large effects) to 1000 (very small effects). Finally, the indirect plus direct effect consistently equaled the total effect, although this depended in part on the sample size. Finally, the application of MMA was based on a replication of a study wherein categorical mediators and outcomes were assessed. Using MMA, the present project reassessed the research questions using more recent data. The additional information provided by MMA sheds light on the small effect sizes of the effects of interest. This information can help interventionists and lawmakers understand where efforts and resources can make the greatest impact.

(128 pages)

PUBLIC ABSTRACT

Towards a Framework for Operational Risk in the Banking Sector Mphekeleli Hoohlo

Mediation analysis is built to answer not only if one variable affects another, but how the effect takes place. However, it lacks interpretable effect size estimates in situations where the mediator (an intermediate variable) and/or the outcome is categorical or otherwise non-normally distributed. By integrating a powerful approach known as average marginal effects within mediation analysis—termed Marginal Mediation Analysis (MMA)—the issues regarding categorical mediators and/or outcomes are, in large part, resolved. This new approach allows the estimation of the indirect effects (those effects of the predictor that affect the outcome through the mediator) that are interpreted in the same way as mediation analysis with continuous, normally-distributed mediators and outcomes. This also, in turn, resolves the troubling situation wherein the indirect plus the direct effect does not equal the total effect (i.e., the total effect does not equal the total effect). By offering this information in mediation, interventionists and lawmakers can better understand where efforts and resources can make the greatest impact.

This project presents the development and the software of MMA, describes the evaluation of its performance, and reports an application of MMA to health data. The approach is successful in several aspects: 1) the software works across a wide variety of situations as the MarginalMediation R package; 2) MMA performed well and was statistically powered much like other mediation analysis approaches; and 3) the application demonstrated the increased amount of interpretable information that is provided in contrast to other approaches.

DEDICATION

This work—the dissertation and all work associated with it—is dedicated to I will always be grateful to her and for her. I also dedicate this work to my child, who patiently loved a father who was often busy working, even when at home. Finally, I dedicate this work to my parents and siblings, who kept me level-headed throughout the process, providing wise and thoughtful advice.

This work is truly evidence of the love I am surrounded by.

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CHAPTER 1

INTRODUCTION

Introduction

Introduction

The initial aim of this study is to apply a generalised linear model (GLM) that is suitable for exposure-based operational risk (EBOR) treatments within the operational risk management framework (ORMF), which effectively replaces historical loss severity curves obtained from historical loss counts, by forward-looking measures using event frequencies based on actual operational risk (OpRisk) exposures. Preliminary work on EBOR models was undertaken by (Einemann, Fritscher, and Kalkbrener, 2018). Secondly, the study provides comprehensive computational comparisons of various data-intensive techniques, versus classical statistical estimation methods for classification and regression performances between the two.

Our understanding of existing ORMF to date is limited to the assumption that financial institutions (FI's) are risk-neutral. In lieu of the afore-mentioned, this study finally seeks to invalidate the risk-neutral assumption, by means of various unsupervised learning techniques, by proposing that FI's are more risk-averse; this can be measured by analysing subtle patterns between data features and trends in the allocated risk capital estimates. In theory, a risk manager who experiences persistent/excessive losses due to particular risk events, would over-compensate cover for these particular risk types, and this would show in reduced losses in these types over time.

The health and prevention sciences are increasingly interested in not only if one variable affects another but how that effect is transmitted. For example, research suggests that chronic illness in adolescence can lead to poor health and behavioral outcomes (Pinquart and Shen, 2011). But the question remains, how does this affect take place? Does chronic illness decrease mental health which in turn causes poor outcomes? Answering this question is not only intellectually interesting but it can provide more meaningful avenues of intervention. Mediation Analysis is designed to help researchers statistically investigate these questions.

Mediation analysis is a widely used technique that allows researchers to investigate *how* one variable may affect another through an intermediate variable. Fields that seek malleable targets of intervention (e.g., the prevention sciences; Coie et al., 1993) find mediation analysis to be an indispensable tool for it allows researchers to evaluate the processes or pathways of an effect—of interventions, risk-factors, and protective-factors alike (Fairchild and MacKinnon, 2009; Hayes, 2009; Iacobucci, 2008; MacKinnon, 2008; MacKinnon, Fairchild, and Fritz, 2007; Shrout and Bolger, 2002). It uses predictors, mediators, and outcomes within a single conceptual model where "the independent [predictor] variable influences the mediator, which in

turn exerts an influence on the dependent [outcome] variable," (Serang, Jacobucci, Brimhall, and Grimm, 2017, pg. 1).

But, as it currently stands, mediation analysis is highly restricted to be used with certain types of data in certain situations. Specifically, when the hypothesized mediator and/or outcome is categorical or otherwise non-normal (e.g., binary, count, multinomial) current approaches are difficult to apply, are restricted to only a few useful cases, and are even more difficult to interpret. If the goal of the health and prevention sciences is to communicate important findings to impact policy, intervention, behavior, and future research, both *utility* (the number of situations wherein it can be used) and *interpretability* (how easily the results can be understood and applied) are key. Without these, results can be misunderstood and misconstrued, leading to false beliefs and ineffective interventions.

Therefore, this project is designed to alleviate this issue through the synthesis of two established methods, in effect creating a new framework for mediation analysis that comfortably incorporates previous work. Establishing this framework requires several key aims that this project ultimately achieves: 1) the development of the general framework and its software, 2) the evaluation of its performance in situations common to health and prevention research, and 3) the application of it to prevention data to demonstrate its use and to better understand important relationships. Before establishing this new interpretable framework, mediation analysis is discussed, highlighting its strengths and current weaknesses.

Loss collection data exercise (LCDE)

Mediation analysis, as built on linear regression (Edwards and Lambert, 2007; Hayes, 2009), combines two or more regression models to estimate the full conceptual mediation model. It is sometimes referred to as *Conditional Process Analysis* (or Moderated Mediation, Mediated Moderation) when combined with moderation (i.e., interaction effects; Hayes, 2013) and is often portrayed via path diagrams and directed acyclic graphs. Confirmatory analysis within mediation is well established for a variety of situations (e.g., Lockhart, Mackinnon, and Ohlrich, 2011) while exploratory analysis is beginning to take shape (Serang et al., 2017). Confirmatory mediation has been applied often in health behavior research—showing pathways leading to health-risk behavior such as drug use (Lockhart et al., 2017; Luk, Wang, and Simons-Morton, 2010; Shih et al., 2010; Wang, Simons-Morton, Farhart, and Luk, 2009), tobacco use (Ennett et al., 2001), and alcohol use (Catanzaro and Laurent, 2004).

Knowing the pathway of effect allows clinicians, interventionists and policy-makers to target modifiable parts of the pathway. For example, there is evidence that bully victimization in adolescence increases depression, which subsequently increases drug use (Luk et al., 2010). In this example, assuming no confounding, there are at least two immediate targets of intervention: the victimization and the depression. Interventions based on a model without the mediator will be incomplete and may fail to alleviate the risk-factor(s). Further, without the mediating effect included in the model, we are at risk of confounding, causing our estimates to

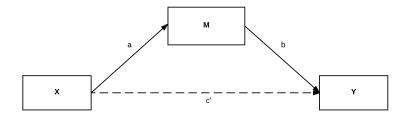


Figure 1.1: Path diagram of a simple mediation analysis model with a single predictor, a single mediator, and a single outcome.

be misleading.

In its simplest form, as shown in Figure 1.1, X is the predictor, M is the mediator (intermediate variable), and Y is the outcome. The paths labeled a and b make up the mediated effect (i.e., "indirect" effect) of X on Y whereas path c' is the direct effect of X on Y (Hayes, 2009). The total effect is equal to $a \times b + c'$, which in linear models, should equal the c estimate in the simple regression of $Y = c_0 + cX + error$. It is important to note that mediation analysis can become much more complex than that in the figure, potentially for a more causal interpretation (Hayes, 2013; Small, 2013).

Definitions

Before discussing mediation further, it is helpful to note some terminology that are often used in the field. In Figure 1.1, X is a predictor or an exogenous variable (i.e., a variable that is not predicted or influenced by something in the model; "independent variable") while M and Y are mediators and outcomes, respectively. These are also known as endogenous variables (i.e., variables that are, in part, predicted or influenced by other variables in the model). These distinctions are useful as the frameworks and assumptions of the models are discussed.

Frameworks

Two highly related frameworks exist to perform mediation analysis (Iacobucci, 2008). First, as mentioned previously, mediation analysis can be built on linear regression including ordinary least squares (OLS) and generalized linear modeling (GLM; Hayes, 2009, 2013). This requires separate models for the a paths and the b and c' paths, fit independently, to be combined into one mediation model. This approach is flexible in terms of the types of variables and model specifications as compared to the other—structural equation modeling (SEM). For example, performing moderated mediation is more straightforward in this framework (Edwards and Lambert, 2007; Hayes, 2013) than in SEM. Ultimately, the regression-based framework is what this project builds upon.

Under the SEM paradigm, all the paths are simultaneously estimated, some-

times providing more statistical power (Iacobucci, 2008). This approach notably allows more testing of the full model fit and can easily include latent variables but assumes, in general, that all variables are continuous with a multivariate normal distribution. This is a strict assumption that is difficult to assess. However, it has extensions allowing for categorical (generally ordinal) variables to be included, although this changes the estimation procedure. The issues relating to categorical mediators/outcomes are discussed in the "Analytic and Interpretation Issues with Mediation Analysis" section.

Assumptions

In his 2008 book "Introduction to Mediation Analysis," MacKinnon (2008) discusses the assumptions² of the mediation modeling procedure. Of these primary assumptions, note that there are no major differences from the assumptions of regression analysis.

- 1. Correct Functional Form. In general, mediation assumes a linear relationship between predictors and mediators/outcomes. This can be adjusted using transformations or, more pertinently, generalized linear models (e.g., logistic regression). MacKinnon (2008) also points out that it is assumed the relationships are additive; if they are not, then the correct interactions (moderators) need to be included in the model specification. This, in many ways, needs to be driven by theory and prior literature (Lockhart et al., 2011).
- 2. No Omitted Influences. A key to any mediation analysis is that variables that: 1) correlate with both the predictor and the mediator (path a), 2) correlate with both the mediator and the outcome (path b), or 3) correlate with both the predictor and the outcome (path c') are included in the model. A more general form of this assumption has been termed "sequential ignorability" (Imai, Keele, and Tingley, 2010a). This more general form includes a sensitivity analysis to assess how important deviations from this assumption are on the conclusions (Imai et al., 2010a; Imai, Keele, and Yamamoto, 2010b).
- 3. Accurate Measurement. Random measurement error produces attenuated paths (in large sample sizes) and random bias (in small sample sizes) in regression (Loken and Gelman, 2017) and therefore can affect the paths in various ways (e.g., attenuate the b path which can inflate the c' path). When possible, reliable measures and/or proper latent variable modeling should be used for this assumption to be met.
- 4. Well-Behaved Residuals. The residuals are assumed to be random, "have con-

¹The idea that SEM is "superior" to the regression paradigm was refuted by Hayes (2013) by noting that in most situations differences in the estimation is extremely minor and will not alter the conclusions. This can be seen in the small effect sizes presented in Iacobucci et al. (2007). Further, additional assumptions inherent in the SEM approach may not hold, although some are not easily tested (e.g., multivariate normality). With this said, SEM still provides a powerful framework for mediation analysis.

²The assumptions described herein are for both the regression and SEM frameworks for mediation, although, as noted above, SEM has a few additional assumptions as well.

stant variance at each value of the predictor variable" and "residual error terms are uncorrelated across equations" (pg. 55). The assumption about uncorrelated errors can stem from "No Omitted Influences" for, if there are omitted variables in both equations, the error terms will correlate. This is one of the few assumptions that can be investigated in many situations.

With the addition of temporal precedence (predictor comes before mediator) and appropriate measurement timing (the mediator is measured at the appropriate time when the effect of the predictor has occurred), the resulting estimates are asymptotically (i.e., with a large enough sample size) unbiased, allowing proper (causal) inference regarding the effects' magnitude and direction (MacKinnon, 2008). This interpretation, to aid in reproducibility, needs to be highlighted with the associated uncertainty (e.g., confidence intervals) in a meaningful metric.

Other Considerations

Causality

To discuss causality in mediation analysis, one should be familiar with the counter-factual framework (or sometimes referred to as the potential outcomes framework; Hofler, 2005). As Imai et al. (2010a) states: "the causal effect ... can be defined as the difference between two potential outcomes: one that would be realized [in the intervention] and the other that would be realized if [not in the intervention]," (pg. 3). In other words, the causal effect is the difference in potential outcomes depending on the predictor (e.g., an intervention). In reality, only one such outcome is observed—if individual "i" is assigned to the treatment group, we only observe the outcome from the treatment group and not from the control group. Imai et al. (2010a) continues: "Given this setup, the causal effect of the [intervention] can be defined as $Y_i(1) - Y_i(0)$. Because only either $Y_i(1)$ or $Y_i(0)$ is observable ... researchers often focus on the identification and estimation of the average causal effect." If the conditions are randomly assigned, this is simply $E(Y_i(1) - Y_i(0))$, or the expected value across multiple individuals and/or observations.

The counter-factual framework helps clarify causality in mediation analysis by defining the necessary conditions. Using this framework, Imai et al. (2010b) demonstrated that sequential ignorability (essentially the assumption that there are no omitted influences) is required for a causal interpretation in mediation analysis. However, this assumption is difficult to assess. Because of this difficulty, Imai et al. (2010a) developed a general mediation model that allows a researcher to assess how deviations from it, via sensitivity analysis, affect the estimates.

As will be shown, the present project incorporates the counter-factual framework intuitively. Because of its importance, this will be discussed more in the next chapter.

Modeling

Shrout and Bolger (2002) highlight a number of other important considerations in mediation analysis. First, multi-collinearity can produce problems, espe-

cially when it occurs between predictors and mediators. It can distort the statistical power of the analysis, potentially producing misleading results. The second consideration is suppression: "Suppression occurs when the indirect effect $a \times b$ has the opposite sign of the direct effect," (pg. 430). This can, if not interpreted correctly, produce confusing estimates (e.g., a positive indirect path and a negative direct path).

Shrout and Bolger (2002) also recommend using bootstrapping (also Hayes, 2009, 2013) to understand the variability in the estimates. This is due to the asymmetric distribution of indirect effects (see Figure 6 in Shrout and Bolger, 2002) that boostrapping can handle naturally. Bootstrapping uses repeated random sampling of the data with replacement and estimates the model on that sampled data. Generally, between 500 and 10,000 bootstrapped samples are used to get an accurate confidence interval. In regards to mediation analysis, bootstrapping produces as accurate (or more accurate) Type-I error rates than other methods in mediation analysis. Because of this, bootstrapping plays a major role in this project.

Two other considerations should be made regarding mediation analysis. First, only the predictors can be randomized (i.e., the mediator cannot be randomized in most situations). That is, even when the a and c' paths can portray an experimental manipulation, the b path(s) cannot. Therefore, the need for proper covariates, interpretation, reporting, and replication is even more important in mediation analysis. David MacKinnon said it well: "It is not likely that a true mechanism can be demonstrated in one statistical analysis. ... These analyses inform the next experiment that provides more information," (MacKinnon, 2008, pg. 67). Therefore, the conceptual model must be considered carefully in light of theory, prior literature, and proper covariates (Lockhart et al., 2011). Iacobucci (2008) also recommends to evaluate competing models and theories—thus presenting the effects and paths in light of alternative model specifications.

Even when done properly, replication of mediated effects is important (MacKinnon, 2008). To help make the replications most useful, the interpretation—comprising the magnitude and direction of the effect—needs to be reported with the proper uncertainty and include information on a) bivariate correlations, b) information on all relevant paths (even non-significant ones), c) must include information on the process of variable and covariate selection, d) report standardized and unstandardized results, and e) provide de-identified data and code [if possible]. It is important to note that much of this information can be included as supplemental material. In this way, results are reported that can be combined with others in order to provide "convincing evidence of [or lack of] a mediating mechanism", (MacKinnon, 2008, pg. 67).

Second, interpretation is built on combining multiple estimates, and often subsequently comparing those combinations with other estimates. For example, the indirect effect is a combination of the a and b paths and is often compared with the c' path. Therefore, if either of the a or b paths are in units that cannot be easily combined, the interpretation quickly becomes very difficult (see "Analytic and Interpretation Issues with Mediation Analysis" section). This is particularly true when the mediator(s) and/or outcome is non-normal.

This second hurdle, that of interpretation, is particularly important in this project. Below, the general interpretation guidelines are discussed, followed by when these guidelines are not straightforward.

Interpretation

In linear models, the interpretation is simple, straightforward and intuitive. The a path coefficient means: "for a one unit change in X there is an associated a units change in the mediator." Likewise, the b path coefficient means: "for a one unit change in M there is an associated b units change in the outcome, controlling for the effect of X." Finally, the c' path is: "for a one unit change in X, controlling for the effect of M, there is an associated change of c' units in the outcome." The indirect effect is $a \times b$; the total effect is $a \times b + c'$.

Each element of the mediation (i.e., the indirect, the direct, and total effects and also the individual a, b, and c' paths) needs to be considered without only trying to answer: "Is there a mediated effect?" Otherwise, researchers can lose sight of the complete story. For example, the various a paths may be important on their own (e.g., if the a path effect size is small then maybe the predictor is not a beneficial place to focus an intervention even though the effect is significant). Therefore, understanding a mediated effect is best told through several avenues: the indirect, direct, and total effects; the individual paths; these effects and paths in light of covariates; among others. This approach is also best if those effects are in meaningful and intuitive metrics.

However, once the analysis ventures into non-normal, non-linear relationships, the interpretation becomes more difficult—particularly when it comes to the indirect and total effects. For example, if the mediator is binary, often logistic regression is used to assess the a path. But that changes the a path interpretation to: "for a one unit change in X, there is an associated a log odds units change in the mediator." This interpretation is anything but intuitive. In general, the log odds are transformed into odds ratios, which improve the interpretation. But these units do not mix well with other units. This is detrimental to understanding the indirect and total effects as will be discussed further in the sections below.

Analytic and Interpretation Issues with Mediation Analysis

Categorical and Non-Normal Mediators/Outcomes

Mediation analysis is more difficult when the mediator and/or the outcome is not continuous, including binary variables (e.g., an individual either uses marijuana or not), ordinal variables (e.g., the self-reported confidence in social settings), other polytomous variables (e.g., sub-types of a disease), and count variables (e.g., number of hospital visits). These data situations are difficult because mediation analysis requires the mediators and/or outcomes to be continuous and approximately normal (to meet the assumption of well-behaved residuals).

There are several strategies taken in the literature to address this problem.

However, each makes its own set of assumptions and each contains limitations in interpretation. The variability in approaches and the subsequent interpretations make combining results across studies far more difficult—reducing the chance to concretely show relationships via meta-analyses and systematic reviews. The difficulty of these data situations are likely reducing reliability and interpretability for a number of reasons:

- It may be easier to ignore the assumptions that are violated when using categorical mediators and/or outcomes. Results from these analyses may not be valid.
- Different approaches produce varying assumptions and interpretations. This
 can be difficult for other researchers, clinicians, lawmakers, and laypersons
 to keep in mind, possibly leading to misunderstandings regarding results and
 their validity.
- Analyses with categorical outcomes are not typically well-emphasized in graduate training, even in more simple modeling techniques, not to mention more complex techniques like mediation analysis. With fewer individuals well-trained, more errors are likely in analyzing these data.
- The interpretation regarding analyses with categorical outcomes is often far less intuitive than with continuous outcomes. This can produce a higher cognitive load for both the researchers and those utilizing the study.

With this in mind, the following subsection discusses the current approaches to mediation analysis with categorical mediators and/or outcomes and the assumptions these approaches make.

Current Approaches

"The quest for sound methods of incorporating categorical variables is perhaps the last dilemma in mediation analysis that lacks a strong solution—it's the 'final frontier,' " (Iacobucci, 2012, pg. 583).

Although there are possibly other "frontiers" in mediation analysis, categorical mediators and/or outcomes certainly produce several challenges. Iacobucci (2008, 2012) thoroughly discusses the issues of assessing categorical variables within a mediation analysis and some of the current practices along with their associated problems³. Four approaches are of note here: 1) a series of logistic regression models, 2) polychoric correlation in SEM, 3) the method suggested by Iacobucci (2012) regarding standardization, and 4) interpreting each path separately. A fifth, but certainly least, is to ignore the distribution of the outcomes and purposefully misspecify the model. These are highlighted in Table 1.1.

Of these approaches, only the SEM approach allows any proper estimation of effect sizes. This approach was developed by Muthén (1984), which uses polychoric correlations (for ordered categorical variables) within the structural equation modeling framework. In this approach, Muthén (1984) assumes that the ordered categorical manifest variable results from a continuous latent variable, "with observed cat-

³For the present, only situations when a mediator or outcome is categorical is under consideration since a categorical predictor poses very few problems (Iacobucci, 2012).

Table 1.1: The various approaches to handling mediation with categorical mediators/outcomes.

| Approach | Pros | Cons |
|--|---|--|
| 1. Series of logistic regressions | Simple to apply in most software | Cannot obtain indirect effect size, only useful in few situations |
| 2. SEM's approach (polychoric correlation) | Powerful, well-designed, Easy to implement with proper software | Only works with ordinal variables, only standardized effect sizes |
| 3. Standardize the coefficients | Provides significance test of indirect effect | Assumptions regarding distributions, difficult to interpret beyond p-value |
| 4. Interpret each path separately | Simplest approach with proper models | Ignores some information, cannot obtain indirect effect size |
| 5. Pretend all variables are continuous | Simplest approach | Purposeful mis-specification, poor model fit |

egorical data arising through a threshold step function," (Iacobucci, 2012, pg. 583). In many cases, a continuous latent variable is a reasonable assumption. For example, in a binary variable describing whether an adolescent has ever smoked marijuana, assuming a continuous latent variable comprised of the probability of smoking may adequately represent reality. This method relies on using a probit regression "to model the relationship between the observed categorical variable and the latent normally distributed variable," (MacKinnon, 2008, pg. 320, emphasis added) and requires large sample sizes (Iacobucci, 2008). Using this latent-observed submodel, the probit is used as the threshold value to estimate the full model. Notably, both the assumption of a normally distributed latent variable and the requirement of large sample sizes are limitations of this approach.

The other approaches, including the series of logistic regression models and the interpretation of each path separately, do not allow any effect size estimation of indirect effects. Although there are some ways to discuss the proportion of mediation (Ford and Hill, 2012), this still does not provide any measures of effect size. The third approach, recommended by Iacobucci (2012), proposed using a new standardization solution using logistic regression output (for a categorical variable) that can combine linear and logistic models' estimates. It is built on the same idea as the Sobel test (MacKinnon, 2008; Sobel, 1982) but is more flexible; for "the mechanics of testing for mediation do not [need to] change whether the variables are continuous or categorical or some mix," (Iacobucci, 2012, pg. 593). However, MacKinnon and Cox (2012) criticized this approach. Ultimately, the most fatal flaw may be that a focus on the significance (relying on Null Hypothesis Significance Testing) of the effects without regard to their effect size in meaningful terms

is less interpretable and impactful.

In the end, none of these approaches can flexibly handle categorical variables of various kinds (e.g., binary, multinomial) or other non-normal distributions (e.g., costs, counts); none can produce intuitive and meaningful effect sizes and confidence intervals across these variable types; and none can consistently combine two differing types of estimates (e.g., binary mediator with continuous outcome, count mediator with binary outcome). Although the current approaches are useful in some situations—particularly the SEM approach—a more complete framework is needed.

Conclusions

Mediation analysis is a powerful framework for understanding the processes by which one variable influences another. The assumptions are not much more than that of regression analysis. The interpretation, in linear models, is straightforward and simple. However, once the analysis ventures into non-normal, non-linear relationships, the interpretation becomes more difficult—particularly when it comes to the indirect and total effects.

In the end, Iacobucci (2012) is correct in saying this problem "lacks a strong solution" (pg. 583). Although important information can be obtained from the current methods, mediation analysis with categorical mediator(s) and/or outcome(s) still misses the mark on intuitive, meaningful effect sizes.

This project aims to alleviate these issues by integrating a post-estimation approach known as *Average Marginal Effects* (AMEs) within mediation analysis. This integration can allow simple and meaningful interpretation across variable types and combinations thus far shown to be problematic. The following chapter introduces AMEs, showing their benefit in interpretation and reporting when working with non-normal variables within generalized linear models.

CHAPTER 2

AVERAGE MARGINAL EFFECTS

Any fool can know. The point is to understand. — Albert Einstein

Introduction

When the outcome variable is not continuous and/or has a distribution far from normal, researchers in prevention science often rely on generalized linear models (GLMs). The power of GLMs is clear when you consider the broad range of situations it estimates with asymptotic consistency. However, the problem with GLMs is that the estimates are not in an easily interpretable form. For example, in logistic regression (one type of GLM), the results are in "log-odds". To overcome this lack of interpretability, a simple exponentiation of the coefficient produces what is known as an odds ratio. Similarly, Poisson regression (another form of GLM), with an exponentiation, produces the risk ratio. Although some fields have adopted odds ratios (or relative risk, risk ratios, and other related metrics), these metrics have notable shortcomings.

- 1. Most of these metrics can be difficult to understand (i.e., many are not intuitive).
- 2. They cannot be combined with other metrics in a meaningful manner.

First, data have suggested that individuals, although with some variability, are able to intuitively grasp the meaning of phrases such as "highly probable" or "not likely" (Heuer Jr., 1999). Yet, this same intuition is not found in odds or odds ratios. For example, Montreuil, Bendavid, and Brophy (2005) found, of 84 articles in several epidemiology journals that used odds ratios, only 7 (8.3%) accurately

¹Asymptotic consistency refers to the ability to, as the sample size increases, produce estimates that converge to the proper value.

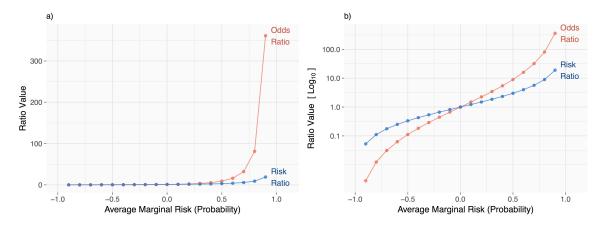


Figure 2.1: a) Comparison of odds ratio and risk ratio with the average marginal risk (probability). b) Same comparison as a) but the y-axis is rescaled (log_{10}) to better show the negative marginal risk comparisons. Both highlight the discrepancy between odds and risk ratios at various levels of marginal risk and that neither approximate the marginal risk.

interpreted the odds ratio and 22 (26%) interpreted odds as though they were probabilities ("risk"). Figure 2.1 highlights the large discrepancy between the odds ratio, the risk ratio, and the average marginal risk. Ultimately, reporting odds and interpreting them as risk is common.

The second is particularly important in the case of mediation analysis given the importance of the indirect effect (the combination of the a and b paths). If a is not in a unit that can be combined with b, then obtaining a meaningful indirect effect is not generally possible. Given its commonality and the limitation that it cannot be combined with other metrics, it is time to consider other strategies—at least in some situations.

Additive vs. Multiplicative Interpretations

The distinction between additive and multiplicative estimates is generally important but is particularly so in mediation analysis. In most quantitative research designs, the investigators are seeking information on the average effect in a population, whether this refers to an average difference across groups or an average change in the outcome for a given change in the predictor. Generally speaking, the average effect is referring to the marginal effect (i.e., the effect of a small change in the predictor in the outcome's units). In the linear regression framework, the average effect is the estimated coefficient and is interpreted additively—a one unit change in the predictor is associated with an X unit change in the outcome. Conversely, outcomes such as OR are multiplicative. Being multiplicative changes the interpretation to: a one unit change in the predictor is associated with an X times change in the outcome. Although subtle, the difference is important, especially for multi-part models (e.g., mediation analysis).

Being multiplicative indicates that the effect of the predictor changes based on the level of the predictor. For example, if the predictor is high, a small change in the predictor may have a big effect while if the predictor is low, a small change in the predictor has little effect. Figure 2.2 shows this phenomenon, where, in the outcomes original units there is an exponential function. In part a) of the figure, it is clear that a change from 2 to 3 in the predictor has a much larger effect than a change from 0 to 1. A regression would not work well here. If a log transformation is used, the relationship would be linear (and a regression can be used) but the interpretation becomes multiplicative (in this case a one unit increase in the predictor is associated with an X * 100 percent increase in the outcome).

Additive interpretations are generally the most intuitive and require less cognitive resources to understand the pattern being portrayed. In a multiplicative framework, simplicity in understanding the effect intuitively is somewhat lost (Iacobucci, 2008). Among others, this is one reason why Stata provides the margins command when dealing with two-part hurdle models.² These models break up the modeling into two parts: one for the binary part and one for the count part. In order to combine the two parts, Stata allows a transformation known as the average

²These models are often used for zero-inflated count outcomes.

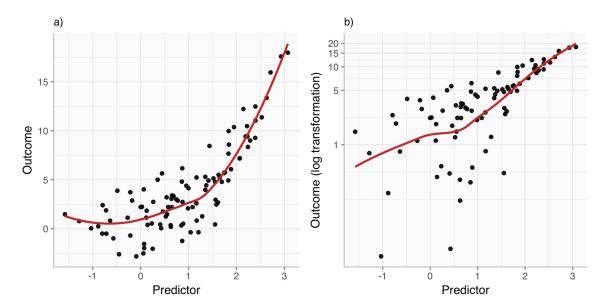


Figure 2.2: Demonstration of a non-linear relationship. a) The outcome is an exponential function of the predictor. b) When log transforming the outcome, the relationship becomes fairly linear. So the interpretation in the log transformed scale is additive, but once it is put back into the original units, it is multiplicative.

marginal effect (AME) that makes the two parts both additive, and therefore easily combined (as is also desired in mediation analysis).

Average Marginal Effect

Why Consider Average Marginal Effects?

When using GLMs, the model is fit with a link function (e.g., "logit", "probit", "log"). This change causes the marginal effect of a variable to rely on the values of the covariates in the model. This is well illustrated through an example. Say a logistic regression model was fit to the data, as shown in Equations 2.1 - 2.4, for p predictors.

$$logit(Y_i) = \beta_0 + \sum_{j=1}^p \beta_j X_{ij} + \epsilon_i$$
(2.1)

$$log(\frac{Prob(Y_i = 1)}{1 - Prob(Y_i = 1)}) = \beta_0 + \sum_{j=1}^p \beta_j X_{ij} + \epsilon_i$$
(2.2)

$$\frac{Prob(Y_i = 1)}{1 - Prob(Y_i = 1)} = e^{\beta_0 + \sum_{j=1}^p \beta_j X_{ij} + \epsilon_i}$$
 (2.3)

$$Prob(Y_i = 1) = \frac{e^{\beta_0 + \sum_{j=1}^p \beta_j X_{ij} + \epsilon_i}}{1 + e^{\beta_0 + \sum_{j=1}^p \beta_j X_i + \epsilon_i}}$$
(2.4)

This implies that the marginal effect of, say, X_{i1} is:

$$\frac{\delta Y}{\delta X_1} = \frac{e^{\beta_0 + \sum_{j=1}^p \beta_j X_{ij} + \epsilon_i}}{(1 + e^{\beta_0 + \sum_{j=1}^p \beta_j X_{ij} + \epsilon_i})^2}$$
(2.5)

That is, the marginal effect of the predictor X_1 depends on the level of each covariate for each individual (all X_{ij} 's) and each estimate (all β_j 's). This, understandably, complicates the interpretation.

Importantly, this example makes it clear that each observational unit (e.g., individual) has a unique marginal effect given her observed levels of each variable. For example, if variable X is increased by one from wherever it was observed, each individual will have different effects (i.e., different marginal effects). One individual may have a large effect; another small. Each individual has a marginal effect of a given predictor associated with her set of characteristics (covariates). To understand the average effect in the population of interest, the mean of all such marginal effects can be calculated.

The AME, then, is the averaged marginal effect across all observational units in the data. In linear models, the AME is the same as the original model estimates. This is intuitive given the AME is the marginal effect in the outcome's original units—the exact interpretation of the estimates in a linear model. However, as stated previously, in GLMs the estimates are not in the original units and therefore

must be estimated via a post-estimation calculation described below.

Definition of the Average Marginal Effect

In an instructive paper about the (now defunct) routine called "margeff" in Stata, Bartus (2005) highlights how the AME can be calculated—including the mathematical definition—and the benefits of AMEs compared to other related methods. The AME is a post-estimation calculation—it uses the model estimates and the data to provide the average effect. Bartus (2005) provided the definition of this post-estimation procedure of a continuous predictor as:

$$AME_k = \beta_k \frac{1}{n} \sum_{i=1}^n f(\beta X)$$
 (2.6)

where f refers to the probability density function of F, βX is the linear combination of the predictors (i.e., the model predicted values for each observation), and AME_k is the average marginal effect for the kth variable. This definition provides the average change in the outcome for a one unit change in the continuous variable x_k across all n observations.

Relatedly, the AME of a dummy coded variable is:

$$AME_k = \frac{1}{n} \sum_{i=1}^{n} \left[F(\beta X | x_{ki} = 1) - F(\beta X | x_{ki} = 0) \right]$$
 (2.7)

where $F(\beta X|x_{ki}=1)$ is the predicted value of the *ith* observation when the dummy variable x_k equals one and $F(\beta X|x_{ki}=0)$ is the predicted value when the dummy value of x_k equals zero holding all other variables constant. This, in effect, shows the discrete difference between the levels of the categorical variable in the outcome's original units.

Confidence Intervals

In general, two approaches are taken to estimate the confidence intervals of AMEs. The first approach is the delta method, which provides standard errors (StataCorp, 2015). Although beneficial, the second—bootstrapped confidence intervals—have proven accurate for understanding the variability in both the AME and mediation analysis. Therefore, this project uses bootstrapped confidence intervals.

The Counter-Factual Framework

As mentioned previously, this project—including the use of average marginal effects—fits within the counter-factual framework. Indeed, the definition of a dummy coded variable demonstrates this well— $[F(\beta X|x_i=1) - F(\beta X|x_i=0)]$. In essence, this answers the question: "What is the difference in the outcome when all observations of x_k are equal to one vs. when all observations of x_k are equal to zero?" The counter-factual framework strives to answer the same class of questions. When all assumptions are met, the AME is a direct, statistical answer to the causality conditions proposed by this framework.

Interpretation

Table 2.1 presents the various units that the AME will produce for the various GLM links. It is important to note that AMEs are in the outcome's original metrics whether they are probabilities, counts, or something else. The interpretation, then, is "for a one unit change in the predictor there is an associated [AME] change in the outcome."

Table 2.1: The generalized linear model link functions with their associated units of interpretation. Note: This list is not exhaustive and there are likely more GLMs that are used within prevention research.

| Link Function | Average Marginal Effect |
|-------------------|--------------------------|
| Identity | Original Continuous Unit |
| Log | Original Continuous Unit |
| Logit | Risk |
| Probit | Risk |
| Poisson | Count |
| Gamma | Count |
| Negative Binomial | Count |

Benefits and Limitations

There are several benefits to using average marginal effects with GLMs.

- Intuitive Interpretation and Few Assumptions. The first, and most obvious, benefit to using AMEs is the simplicity of the interpretation. The effect is in the units used in the modeling; it is additive (i.e., the effect is the added increase or subtracted decrease in the outcome). It provides an interpretation that imitates that of ordinary least squares regression. Relatedly, there are no difficult modeling assumptions directly tied to AMEs. Instead, the underlying models' assumptions that are used to get the AME is what is important. The only additional assumption with AME is that the effect, for each individual, is linear enough to be represented by an additive value and that the average adequately reports this.
- More Generalizable and Robust. There is evidence suggesting that AMEs are
 more robust to problems associated with GLMs (including logistic regression)
 such as unobserved heterogeneity and model mis-specification (Mood, 2010;
 Norton, 2012). This allows the estimates to be more generalizable to individuals outside of the sample.
- Low Computational Burden. Given AMEs are a post-estimation calculation,

no new models need to be fit. Instead, using the estimates of the models, the average marginal effects can be calculated. The most computationally expensive aspect of the calculation is the bootstrapped confidence intervals.

- Broadly Applicable. The AME applies to any of the generalized linear models including logistic, Poisson, gamma, beta, negative binomial, and two-part hurdle models. This provides extensive flexibility in modeling, and, once applied to mediation, will allow flexibility in modeling based on the correct functional form.
- Two-Part Models. Particularly pertinent to this project is that the calculation of AMEs have been applied to two-part models, generally of the hurdle model types, as stated earlier. In fact, this is a common routine in the Stata statistical software. This provides valuable support for the proposed approach of using AMEs in mediation analysis.

Notably, the interpretation of AMEs hold to the assumption (as found in all regressions) that it is reasonable to adjust a single covariate while holding all others constant. This may not hold in reality, although it may be necessary to gain an understanding of the individual effect of a single variable. In data that are not representative of the population (e.g., non-random sample), AMEs may be biased because an over-representation of certain covariate values may be present. This is an overall modeling problem, since GLMs also assume a random sample. In this way, this problem is not specific to AMEs.

Conclusions

The Average Marginal Effect produces intuitive, interpretable, and additive estimates of an effect. They have been applied to two-part models, similar to mediation analysis, demonstrating their utility in difficult modeling situations. The following chapter discusses the integration of AME with mediation analysis—termed

Marginal Mediation Analysis— with its interpretation and assumptions, its benefits and limitations, and the basic procedures for its use.

CHAPTER 3

MARGINAL MEDIATION ANALYSIS

Without an interpretable scale, it is difficult to use effect size to communicate results in a meaningful and useful way. — Preacher and Hayes, 2011

Introduction

The proposed integration of average marginal effects and mediation analysis is designed to resolve two major obstacles currently found in mediation analysis:

- 1. The difficulty of performing mediation analysis with categorical mediators and/or outcomes, and
- 2. The lack of reliable and flexible effect size estimates in mediation analysis—particularly with categorical mediators and/or outcomes.

These issues are relatively common in prevention work (e.g., Ford and Hill, 2012; B. Hoeppner, Hoeppner, and Abroms, 2017; Wong and Brower, 2013) and the current approaches are not adequate—as was discussed at length in the previous chapters. In this chapter, the integration of Average Marginal Effects (AMEs) and mediation analysis—Marginal Mediation Analysis (MMA)—is discussed, including its interpretation and assumptions as well as its benefits and limitations. It is expected that this adjustment to both the modeling and the interpretation will help researchers in the health and prevention sciences to be able to model their data in the most properly-specified way and be able to communicate their findings clearly.

Definition of Marginal Mediation Analysis

The form of the general marginal mediation model, including the postestimation step, are demonstrated in the following equations, where Equations 3.1 and 3.2 demonstrate the mediation estimation while Equations 3.3 and 3.7 show the post-estimation procedures.

$$M_{ij} = a_0 + \sum_{k=1}^{p} a_k x_{ki} + \epsilon_i$$
 for $j = 1, ..., m$ mediators (3.1)

$$Y_i = \beta_0 + \sum_{j=1}^m b_j M_{ij} + \sum_{k=1}^p c'_k x_{ki} + \epsilon_i$$
 (3.2)

for the *ith* individual, for k = 1, ..., p predictors, and j = 1, ..., m mediators. The paths are all labeled with their common term (e.g., path a is labeled a). Combining these two equations provides the full mediation model. Using these models, we apply the post-estimation of the average marginal effects as presented by Bartus (2005). For a continuous x_k variable, the average marginal effect of path a is:

$$AME_k^a = a_k \frac{1}{n} \sum_{i=1}^n f(aX)$$
 (3.3)

where f refers to the probability density function, aX is the linear combination of the predictors (i.e., the model predicted values for each observation), and AME_k^a is the average marginal effect of the a path for the kth variable. Ultimately, Equation 3.3 is identical to that of the following:

$$AME_k^a = \frac{1}{n} \sum_{i=1}^n \frac{f(aX_1) - f(aX_2)}{2h}$$
 (3.4)

where

$$aX_{1} = \begin{bmatrix} ax_{11} & ax_{12} & \dots & ax_{1k} + h & \dots & ax_{1p} \\ ax_{21} & ax_{22} & \dots & ax_{2k} + h & \dots & ax_{2p} \\ \vdots & \vdots & \ddots & \vdots & \ddots & \vdots \\ ax_{n1} & ax_{n2} & \dots & ax_{nk} + h & \dots & ax_{np} \end{bmatrix}$$
(3.5)

and

$$aX_{2} = \begin{bmatrix} ax_{11} & ax_{12} & \dots & ax_{1k} - h & \dots & ax_{1p} \\ ax_{21} & ax_{22} & \dots & ax_{2k} - h & \dots & ax_{2p} \\ \vdots & \vdots & \ddots & \vdots & \ddots & \vdots \\ ax_{n1} & ax_{n2} & \dots & ax_{nk} - h & \dots & ax_{np} \end{bmatrix}$$
(3.6)

Both $f(aX_1)$ and $f(aX_2)$ are the model predicted value for the outcome given the small change due to h. Equations 3.4, 3.5, and 3.6 use a very small h value (default is 1×10^{-7}). This provides the change in the average predicted value for a very small increase and a very small decrease in in the x_k variable. This is also described in depth by Leeper (2017) since it is the strategy employed by margins. This approach is flexible—especially in the R statistical environment with the improvement of derivative computation that Leeper provided.

Similarly, the AME of a dummy coded variable in the a path is:

$$AME_k^a = \frac{1}{n} \sum_{i=1}^n \left[F(aX|x_{ki} = 1) - F(aX|x_{ki} = 0) \right]$$
 (3.7)

where $F(aX|x_{ki}=1)$ is the predicted value of the *ith* observation when the dummy variable equals one and $F(aX|x_{ki}=0)$ is the predicted value when the dummy value equals zero. This is the same approach used by margins.

Notably, these same post-estimation equations (3.3 and 3.7) can be used for the b and c' paths as well.

Interpretation

The interpretation of Marginal Mediation is based on the original units of the mediator(s) and outcome(s). Because there are so many possible combinations of GLM types within mediation analysis, instead of outlining every combination, the basic principles are presented that apply to all situations.

Principle 1: The individual paths are interpreted based on the corresponding endogenous variable's original metric.

The individual paths have interpretations identical to those of AMEs, as discussed in Chapter 2. Therefore, the a path depends on the type of mediator and modeling approach chosen. For example, for a binary mediator, the a path is a probability (risk); a mediator representing a count has an a path that is in the same count units.

Principle 2: The indirect effect, as a combination of the a and b paths, are interpreted based on the outcome's original metric.

The indirect effect is joining two paths, possibly of different metrics. However, the interpretation is still straightforward: the entire effect will be in the outcome's original metric. An example may be beneficial to highlight this principle.

Suppose there are data with a hypothesized binary mediator (depression or no depression) and a continuous outcome (quality of life). A logistic regression is used to model path a and a linear regression is used for the b and c' paths. After calculating the AME of the paths, path a is in units of risk (of depression) while path b is the difference in the quality of life between depressed and not depressed individuals. By combining paths a and b through multiplication, we get the effect of the predictor on depression risk then what that depression risk does to quality of life; that is, the effect of the predictor on quality of life through depression.

Principle 3: Both the direct and total effects are interpreted based on the outcome's original metric.

Similar to Principle 2, the direct and total effects are in the outcome's original units. This is intuitive given that the AME of the direct effect is in the outcome's units, and it is not combined with any other path. For the total effect, the indi-

rect and direct effect (which are both in the same units) are added together to get the complete effect. It was expected that, as in linear regression (but that is lacking in other situations), $a \times b + c' = c$. That is, the indirect and direct effects together equals the total effect. This expectation was tested, as is described in the next chapter.

Effect Sizes

A major advantage of this framework is the way effect sizes can be used intuitively.

"First, virtually all effect size indices should be scaled appropriately, given the measurement and the question of interest. Without an *interpretable scale*, it is difficult to use effect size to communicate results in a meaningful and useful way.... Second, it should be emphasized that effect size estimates are themselves sample statistics and thus will almost certainly differ from their corresponding population values. Therefore, it is important to report confidence intervals for effect sizes..." (Preacher and Kelley, 2011, pg. 95, emphasis added).

The interpretation of MMA makes it clear that both of these aspects of proper effect size estimation and reporting are adequately represented. Of particular note, unlike many effect sizes that are only useful in certain research questions, the effect sizes produced by AMEs—and thus found in Marginal Mediation Analysis—are flexibly oriented to "be scaled appropriately" to best "communicate results in a meaningful and useful way" for a wide variety of situations.

Additionally, Preacher and Kelley (2011) continue: "it is important to develop a way to gauge the effect size of the product term ab itself," (pg. 95). That is, not only does the effect size of the individual paths need to be meaningful but the product of $a \times b$ must be as well. Marginal Mediation Analysis provides this compara-

bility as the indirect effect can be compared without problems with the direct and total effects (they are all in the same units).

Assumptions

The assumptions inherent in MMA are the same as those presented in Chapter 1 regarding mediation analysis. The only additional assumption regards the ability of the effect to be represented additively (i.e., can the effect be represented linearly after accounting for the marginal effect for each observation?). In linear models, this is already included as an implicit assumption. For other models, although the relationship may not be linear in the outcome, taking the average of the effects across the observations is assumed to be representative of the relationship across the sample.

Reproducibility and Interpretability

As mentioned previously, categorical mediators/outcomes are generally not difficult to model using GLMs—only the interpretation is difficult. Generalized linear models, in conjunction with AMEs, allow researchers to use more correct functional forms, thereby reducing the justification to fit poorly specified models that have an easier interpretation.

With this framework, MMA can be applied across the GLM spectrum and essentially any combination of GLM types. For example, a marginal mediation model is defined when the mediator is binary and the outcome is continuous; when the mediator is a count and the outcome is ordinal; when the mediator is continuous and the outcome is binary. Each has a straightforward, yet informative, interpretation as outlined by the principles above. This attribute alone can increase the reproducibility of research using mediation analysis.

Finally, the interpretation across the paths and effects is straightforward and

flexible. Other researchers, laypersons, lawmakers, and clinicians can assess the direction, magnitude, meaning, and utility of findings much easier—thus, increasing the reach and impact of research. In mediation analysis, this can prove largely beneficial given the already complex nature of the modeling scheme. By simplifying the interpretation, less cognitive resources are required to gain a basic understanding of the findings; instead, more resources can be used to understand how to apply it and assess future research questions based on the findings.

CHAPTER 4

METHODS

For precept must be upon precept, precept upon precept; line upon line, line upon line; here a little, and there a little. — Isaiah 28:10, King James Bible

Introduction

As presented in the previous chapter, Marginal Mediation Analysis (MMA) has the ability to simplify the interpretation of mediated effects in a wide variety of situations, particularly in situations where an effect size otherwise does not exist (e.g., indirect effects when the mediator or outcome is categorical). In this chapter, methods fashioned to develop MMA and evaluate its performance are discussed via three phases:

- 1. Development of MMA
- 2. Monte Carlo Simulation Study of MMA
- 3. Application of MMA

These phases were designed to provide the theory and the software to perform MMA, assess the method's ability to accurately estimate the underlying effects, develop the guidelines of its use in finite samples, and apply it to real-world prevention data by replicating a recent study (Ford and Hill, 2012) that used a categorical mediator and categorical outcomes. Below, each phase is described in depth.

Phase I: Development of Marginal Mediation Analysis

To be useful to public health, psychological, and prevention researchers, the incorporation of average marginal effects within mediation analysis must happen in two ways: in theory and in software. This phase is focused on understanding the

properties of MMA and on developing the software necessary to perform it.

Properties of MMA

Building on the mediation framework discussed by Hayes (2009) and by Edwards and Lambert (2007), MMA was established on linear regression—either ordinary least squares (OLS) for continuous outcomes/mediators or maximum likelihood (via GLMs) for categorical outcomes/mediators. In this framework, two or more regression equations are combined to provide the overall mediation model as discussed in Chapter 1. This method then adds a post-estimation step (Chapter 2) into this mediation framework.

The form of the general marginal mediation model, including the postestimation step, were discussed in Chapter 3. Using this general framework, various considerations were made in the development of the method. First, an appropriate manner in which to integrate moderation (interaction effects) into the framework is important. Because of the work by Edwards and Lambert (2007), this included assessing the reduced form¹ of the models in addition to visualizations of the predicted values across levels of the moderator. Second, it has been noted by MacKinnon (2008) that in non-linear models the $a \times b + c'$ generally does not equal the c path as it does in linear models (Chapter 11 in his book). To assess whether these are equal within MMA, both a basic analysis and a Monte Carlo simulation (phase II) was used.

Software Development

A major aspect of this first phase is the development of the software for researchers to apply MMA. This software is provided via the R statistical environment given R is free, widely used by researchers in health and prevention, and ex-

¹Reduced form refers to only having exogenous variables on the right-hand side of a regression equation (i.e., substituting the predictors of the mediators into the equation).

tensions to the software via "packages" are efficiently disseminated through the Comprehensive R Archive Network (CRAN). It consists of a number of functions to fit the model and assess the model's fit while efficiently producing the paths and effects in proper units.

Because of the flexibility of numerical derivation methods and the speed improvements by Thomas Leeper (2017), numerical derivatives were used to obtain the marginal effects for each observational unit. From here, means of the marginal effects were calculated for each variable in the model. To assess the model uncertainty, bootstrapping via the boot R package was applied. This approach relies on repeated resampling from the original sample with replacement. In general, this method does not rely on any distributional assumptions and works well with asymmetric distributions (as is found in indirect effects).

The package applies the best practices for both computational speed and user readability (Wickham, 2015), allowing other researchers to extend the package more easily. Additionally, several built-in tests will inform the functionality of the package before beginning Phase 2. These tests were performed on Linux, Mac, and Windows platforms. Finally, the package uses Git as the version control system. The necessary functions were developed first so that the package tests and simulations could begin. The usability of the package that is important in the disseminated version were developed afterwards.

Phase II: Monte Carlo Simulation Study of Marginal Mediation Analysis

The evaluation of MMA is an essential step in understanding its properties and robustness and further assess the performance of the software. The consistency of MMA, the statistical power at various sample sizes, and the accuracy of the bootstrapped confidence intervals were all tested via a Monte Carlo simulation study (Carsey and Harden, 2013; Paxton, Curran, Bollen, Kirby, and Chen, 2001).

In the simulation, data were simulated to come from a population of known parameters. A literature review of mediation analysis in prevention work high-lighted the appropriateness of the population parameters chosen. The results of the simulation helped in the development of the guidelines for using MMA in practice.

Literature Review

Before performing the Monte Carlo simulations, a review of the literature is recommended (Paxton et al., 2001). This review focused on the use of mediation analysis in prevention research where the analyses contained categorical mediators and/or outcomes. This review included all recent articles (2008 - 2017) found that clearly identified a mediator or outcome that was categorical in nature. This search relied on terms such as "generalized linear models", "logistic", "dichotomous", "polytomous", and "count" in conjunction with "mediation analysis" across the Scopus database.

Simulations

Monte Carlo simulations, via the R statistical environment version 3.4.2, assessed the finite properties of MMA. Monte Carlo simulation was selected due to its simplicity in generating informative results and its high success in the literature (e.g., Graham, Olchowski, and Gilreath, 2007; Nylund, Asparouhov, and Muthen, 2007). Here, 500 data sets were simulated for each combination of experimental conditions (Carsey and Harden, 2013; Paxton et al., 2001). The data were simulated from a known population with a researcher specified causal model (i.e., the "population model"). The model consisted of either a binary mediator (0 = "No", 1 = "Yes") or a count variable (Poisson distribution), a continuous outcome, and a continuous predictor while also varying the sample size and the effect sizes for a total of 90 unique combinations of the conditions (see Table 4.1).

The a path population model is defined below for when the mediator is binary, where the $Prob(M_u = 1)$ is a latent continuous variable with a logistic relationship with the predictors and the ϵ_i is normally distributed with a mean of 0 and a standard deviation of 1.

$$log(\frac{Prob(M_u = 1)_i}{1 - Prob(M_u = 1)_i}) = a_0 + a_1 x + \epsilon_i$$
(4.1)

The observed variable, M_o , is defined as follows:

$$M_o = \begin{cases} 0, & \text{if } Prob(M_u = 1) < .5, \\ 1, & \text{otherwise.} \end{cases}$$

$$(4.2)$$

The a path population model for when the mediator is a count is shown below where M_u is a latent continuous variable with an exponential relationship with the predictors and the ϵ_i is normally distributed with a mean of 0 and a standard deviation of 1.

$$log(M_u) = a_0 + a_1 x + \epsilon_i \tag{4.3}$$

The observed variable, M_o , is defined as follows:

$$M_o = Po(\lambda = M_u) \tag{4.4}$$

This creates an observed, count variable that has λ values based on the latent mediator.

The b and c' paths population model is identical to Equation 3.2 with only one predictor and a single binary, or count, mediator, as shown below.

$$M_i = a_0 + a_1 x_1 + \epsilon_{mi} \tag{4.5}$$

| Table 4.1: The various experiment | al conditions of | f the Monte | Carlo simulation s | study. |
|-----------------------------------|------------------|-------------|--------------------|--------|
|-----------------------------------|------------------|-------------|--------------------|--------|

| Independent Variables | Conditions |
|--|---|
| Sample size Effect size of a path Effect size of b path Effect size of c' path | 50, 100, 200, 500, 1000 small, moderate, large small, moderate, large moderate |
| Type of mediator Total conditions | binary, count |
| Total conditions | 90 |

Table 4.2: The effect sizes of the a and b paths in the Monte Carlo Simulation.

| Size | Odds Ratio (a path) | Risk Ratio (a path) | r (b path) |
|----------|---------------------|---------------------|------------|
| Small | 1.58 | 1.34 | 0.10 |
| Moderate | 3.44 | 1.82 | 0.30 |
| Large | 6.73 | 3.01 | 0.50 |

$$Y_{i} = \beta_{0} + bM_{i} + c'x_{i} + \epsilon_{yi} \tag{4.6}$$

Table 4.1 highlights the conditions that were varied for each simulation. A distinct MMA model was applied to each of the 500 data sets for each possible combination of experimental conditions. This means 45,000 MMA models were fit. Using eight cores of powerful core i7 computers, these computations were finished over a span of several days.

Notably, the effect sizes for both the binary mediator and count mediator (a path) were the odds ratios and risk ratios corresponding to small, moderate, and large effect sizes. These are found in Table 4.2.

The focus of the simulations was to gauge the accuracy, power, and coverage of MMA at estimating the population effects while undergoing the experimental conditions. The dependent variables were:

- 1. bias (i.e., is the mean of the estimates at the population mean?),
- 2. power (i.e., how often does the null properly get rejected?),

3. confidence interval coverage (i.e., does the confidence interval cover the proper interval?), and

4. consistency regarding how closely $a \times b + c'$ is to c (i.e., does the indirect plus the direct effect equal the total effect?).

The effects of the conditions on these outcomes were assessed via visualizations and descriptive tables.

Guideline Development

Recommendations from the simulation study were documented, including necessary sample sizes, bias in various conditions, and the accuracy of bootstrapped confidence intervals in each condition. The documentation will be available in manual form online on the R website and GitHub.

Phase III: Application of Marginal Mediation Analysis

During the third phase, all important aspects of MMA discovered throughout the first two phases were used to replicate previous work regarding the relationship of adolescent religiosity with substance use (Figure 4.1). This study was selected given it used:

- 1. a large sample with a mix of binary and continuous mediators and outcomes,
- 2. one of the better and common statistical approaches, and
- 3. data that were publicly available and that had a more recent release to investigate.

Data

To replicate this study, data from the 2014 (most recent) release of National Survey on Drug Use and Health (NSDUH; Ford and Hill, 2012) were used. As described by Ford and Hill (2012), the NSDUH is "an ongoing study sponsored by the

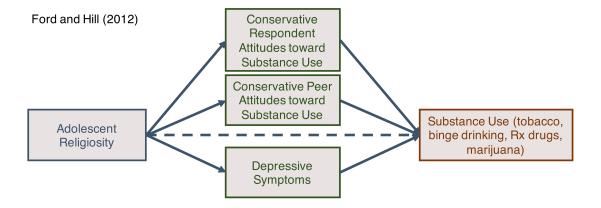


Figure 4.1: Path diagram of the replicated models from Ford and Hill (2012). The depression mediator and all of the outcomes are categorical.

U.S. Substance Abuse and Mental Health Services Administration that dates back to the 1970s" (page 4) and collects data on drug use of individuals 12 years and older across the United States. Ford's study used the 2007 data release while the replication uses the 2014. Several measures² were used to replicate the findings of Ford and Hill (2012):

- 1. Four substance use outcomes (tobacco use, prescription drug use, marijuana use, and illicit drug use).
- 2. Religiosity was based on the average response across four items relating to church attendance, the importance of religious beliefs to the individual, and participation in faith-based activities. Higher scores indicated more religiosity.
- 3. Respondent attitudes toward substance use was also the average response based on four items gauging the individual's response to someone their age using substances. Higher scores indicate more conservative attitudes.
- 4. Peer attitudes toward substance use is similar to the respondent attitudes except that it was asked how the individual's friends would feel about someone their age using substances. Again, the average response was used where

²Notably, this replication study omits the "heavy drinking" outcome because positive responses to it were very rare in the 2014 data.

higher scores indicate more conservative attitudes.

5. "Psychological well-being was indicated by major depression," (page 5) which was measured as at least five of the nine possible depression symptoms listed in the survey.

Substance Use

As stated previously, the substance use outcomes are tobacco use, prescription drug use, marijuana use, and illicit drug use.

- Tobacco use was defined as one of the following three items: 1) cigarette use within the last year, 2) smokeless tobacco use within the last year, and 3) cigar use within the last year.
- Prescription drug use consisted of four groups of drugs that are being used either without a prescription or for the sole purpose of obtaining a high within the last year: pain relievers, tranquilizers, stimulants, and sedatives.
- Marijuana use was a single item: marijuana use within the last year.
- Illicit drug use was defined as using any of the following drugs within the last year: cocaine, crack, heroin, hallucinogens, LSD, PCP, ecstasy, inhalants, or meth.

Each outcome was coded as dichotomous: use or no use within the last year.

Religiosity

Adolescent religiosity was the mean response across four items: 1) the number of times attended religious activities in past year, 2) religious beliefs are important, 3) religious beliefs influence decisions, and 4) the amount of participation in religious activities. The higher the average score the more the adolescent is considered to be religious.

Respondent and Peer Attitudes

The respondent's conservative views on drug use is the average of four items, answering the question "How do you feel about someone your age using [cigarettes daily, marijuana, marijuana monthly, drinking daily]?" Similarly, the peer's conservative views on drug use is the average of four items, answering the question "How do you think your close friends would feel about you using [cigarettes daily, marijuana, marijuana monthly, drinking daily]?"

Psychological Well-being

Finally, psychological well-being was defined as having had a major depressive episode in the past year. This was a binary (yes or no) variable based on "if they reported experiencing at least five of the following: felt sad, empty, or depressed most of the day or discouraged; lost interest or pleasure in most things; experienced changes in appetite or weight; sleep problems; other noticed you were restless or lethargic; felt tired or low energy nearly every day; felt worthless nearly every day; inability to concentrate or make decisions; any thoughts or plans of suicide," (Ford and Hill, 2012, pg. 5).

Analyses

The mediation analyses were replicated from Ford and Hill (2012). Although the mediation analysis is performed differently herein, the model specifications were identical to that employed there.

Importantly, Ford and Hill (2012) say: "we use the categorical data method outlined by MacKinnon (2008) to formally test the indirect effects," (pg. 5). This approach uses a significance test based on the estimates of both a and b and their standard errors. However, as stated throughout this project, the significance alone is insufficient information to provide for a mediation analysis; effect sizes are also

necessary. Because of this, Ford and Hill (2012) continue by discussing the amount of the association between the predictor and outcome, in percentage units, that the mediator accounted for. This approach is useful but has some notable shortcomings. First, depending on the level of multi-collinearity in the models, the standard errors of the estimates can be inefficient which reduces the statistical power of this test. Second, it does not provide the effect size measures that would be most useful (e.g., the effect a one unit increase in the predictor has on the outcome through the mediator). Third, the measure is consistently too conservative with binary outcomes (Jiang and Vanderweele, 2015).

For the replication, then, each of the mediation models reported were run using MMA in place of the techniques employed by Ford and Hill (2012). Four distinct MMA models, one for each of the substance use outcomes, were assessed. These were all controlling for (adjusting for) parental attitudes towards substance use, age, race, sex, and income. The models included 500 bootstrapped samples to obtain 95% confidence intervals.

Further, using a variant of the "difference method," the amount of the total effect that was mediated was calculated using the following:

Proportion mediated =
$$\frac{indirect}{indirect + direct}$$

Finally, the information provided through the use of MMA was also compared to that produced in the original paper.³

Conclusions

Ultimately, the goal of this project is to develop, evaluate and apply a method that can provide meaningful interpretation in mediation when the mediator and/or

 $^{^3}$ All analysis code is available on the Open Science Framework (osf.io/753kc) and in the Appendices of this document.

outcome is categorical. Each phase builds on this goal, as is discussed in the following chapters starting with the presentation of the results of Phase I and Phase II regarding the theory, software, and evaluation of MMA.

CHAPTER 5

PHASES I & II: DEVELOPMENT AND TESTING OF MMA

Exploring the unknown requires tolerating uncertainty. — Brian Greene

Introduction

The results of both Phase I (the development of the method and its software) and Phase II (the Monte Carlo simulations) regarding Marginal Mediation Analysis (MMA) are presented in this chapter.

Developmental Considerations

The general MMA framework was discussed in Chapter 3. This framework was extended with some important additional considerations, including integrating moderation and analytically assessing the relation between the decomposed total effect and the total effect.

Moderation

Moderation (interaction) is sometimes hypothesized to occur in conjunction with mediation. Moderation is any situation where the effect of a variable on another *depends* on the value of a third variable. This phenomenon, in conjunction with mediation, is often referred to as conditional process analysis, moderated mediation, or mediated moderation—depending on the source and situation.

An example of one of the many possible moderated mediation models is found in Figure 5.1 (for more examples, see Hayes, 2013). In this example, the moderator (denoted W in the figure), moderates that relationship between X and M. In other words, the effect of X on M depends on the value of W. This further suggests that

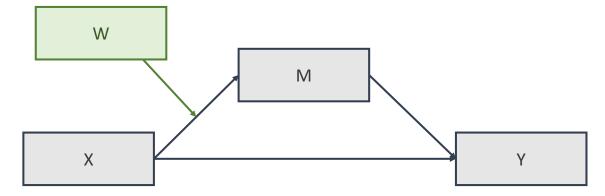


Figure 5.1: An example of moderated mediation, where the moderator (denoted W), moderates the effect of X on M (the a path).

the effect of X on Y, through M, also depends on the value of W.

In general, interactions make interpretation more difficult. In linear models, the interpretation of the interaction estimate becomes: "a one unit increase in X is associated with a $\beta_x + \beta_{int} \times W$ effect on the outcome." That is, to understand the size, and direction, of the effect of X, the level of W must be considered. For example, if W is a categorical variable with values of 0 and 1, and the following regression was estimated:

$$\hat{M}_i = 1.0 + 5.0X + 3.0W + .5X * W \tag{5.1}$$

then:

- 1. X is associated with a 5.0 increase in M when W is 0, and
- 2. X is associated with a 5.5 increase in M when W is 1.

The same logic holds for continuous moderators, although representative values of W must be chosen instead of using all possible values.

Yet, in non-linear situations, this becomes more strenuous. However, using average marginal effects, the interpretation can be like that of linear models. In general, this has been done by selecting various, representative values of W at which the average marginal effect is assessed (StataCorp, 2015). If W is categorical then

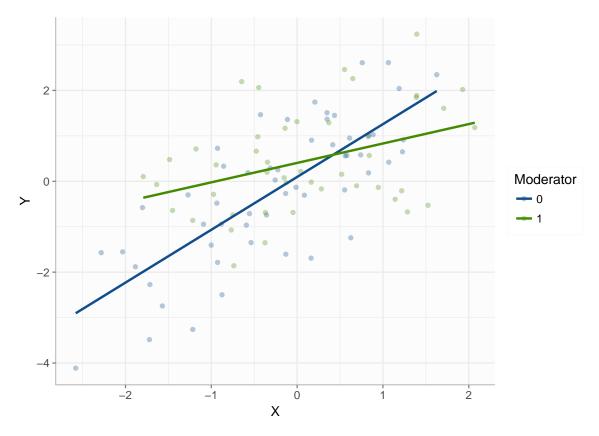


Figure 5.2: An example of a common approach to visualizing a moderation effect where the effect of X on Y is shown for each level of W.

all observed values can be used. Notably, in linear and generalized linear models, moderation is probably best understood using visualizations—showing the effect of X at various levels of W (see Figure 5.2).

In relation to MMA, moderation can be understood at both 1) an individual path level and 2) a complete model level. At an individual path level, moderation is understood as it is in non-mediating regression situations. For example, if path a is moderated, the effect of X on M can be understood via visualizations or representative values of W can be inserted into the regression equation, as in the example above.

To understand it in relation to the complete model, the framework discussed by Edwards and Lambert (2007) suggests using the *reduced form* of the mediation

model to understand the moderation in the context of the indirect and direct effects. The reduced form refers to having only exogenous variables on the right-hand side of the equation (i.e., substituting the estimates of the a path into the b/c path model) as shown below.

Starting with the non-reduced form, we have M_i , an endogenous variable, on the right hand side.

$$Y_i = \beta_0 + b_1 M_i + c_1 \dot{x}_i + \epsilon_{yi} \tag{5.2}$$

Using the a path model, and assuming the same model specification as in Figure 5.1, we can substitute in the predictors of M_i .

$$Y_{i} = \beta_{0} + b_{1}(a_{0} + a_{1}x_{i} + a_{2}w_{i} + a_{3}x_{i} * w_{i}) + c_{1}^{'}x_{i} + \epsilon_{yi}$$

$$(5.3)$$

This form is now reduced so that only exogenous variables are on the right-hand side. Using these estimates, it is now possible to assess the effect of x on Y when it depends on the level of w using the same approach with the individual paths. Importantly, using these estimates, the moderated effect of x can be visualized as well.

The Decomposed Total Effect Equals The Total Effect

When using the average marginal effect, the decomposed total effect $(a \times b + c')$ equals that of the original total effect (c). Winship and Mare (1983), demonstrated that, using calculus, an outcome variable Y can be decomposed by its total differential

$$dY = \frac{\delta Y}{\delta X}dX + \frac{\delta Y}{\delta M}dM \tag{5.4}$$

which implies the general formula

$$\frac{dY}{dX} = \frac{\delta Y}{\delta X} + \frac{\delta Y}{\delta M} \frac{dM}{dX} \tag{5.5}$$

(pg. 83, the symbols were altered to match that of the present project). That is, the total effect is equal to the direct plus indirect effects. If the average marginal effect is a good estimate of the derivative (or partial derivative), then:

$$\frac{dY}{dX} = \frac{\delta Y}{\delta X} + \frac{\delta Y}{\delta M} \frac{dM}{dX} = c' + b \times a = c \tag{5.6}$$

Therefore, it is expected that regardless of the distributions of the mediators or outcomes $a \times b + c' = c$.

This is further demonstrated with finite sampling properties in the Monte Carlo simulation in Phase II using both binary and count mediators.

Standardized Effects

It is often of considerable worth to understand standardized effects. These can be defined in numerous ways, depending on the situation and types of variables that are being used. In situations where the outcome is continuous, MMA can use a partial standardization approach discussed by Preacher and Hayes (2011) where the outcome is standardized using its own standard deviation. This produces interpretations that are based on the change in the outcome in standard deviation units. If using a dichotomous predictor, this essentially becomes a standardized mean difference (e.g., Cohen's D). It is also possible to standardize both the continuous outcomes and continuous predictors to obtain a partial correlation metric from these models as well.

As discussed below, the software to perform MMA includes the outcome standardization for continuous outcomes but does not provide standardization tech-

niques for the predictors. Future iterations of the software will include this as well.¹

Software Development

The software package developed for MMA is called MarginalMediation and is freely available via the R statistical environment. The software allows straightforward use of MMA across continuous, binary, and count mediators and/or outcomes (other distributions also work but have not been extensively tested). The computation is done in several steps:

- 1. Function and model checks
- 2. a path average marginal effect estimates
- 3. b and c' path average marginal effect estimates
- 4. Bootstrapped confidence intervals
- 5. Formatting and printing of the output

This strategy was undertaken to help in error-checking and allows the function to print informative output to the user during the modeling, which is especially useful for situations with large samples and many bootstrapped samples.

Functions

Using the package is based on a single function—mma()—that provides the main functionality (Figure 5.3). mma() is built on several other functions that perform specific duties that allow the simple syntax. The main functions of the package are shown in Table 5.1.

¹The researcher can standardize the continuous predictors and outcomes before performing MMA, in which case they can obtain the partial correlation estimates with the current version of the software.

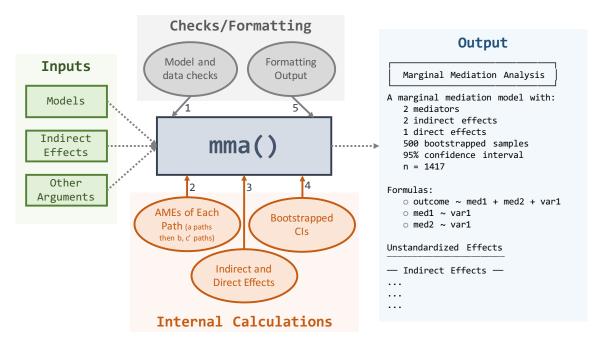


Figure 5.3: The structure of the mma() function. From left to right: (1) inputs inform the function of the model specifications, the indirect effects to be reported, and other arguments; (2) internal processes including checks/formatting and internal calculations, (3) the output with some truncated example output. Notably, the numbers along the solid arrows pointing at the mma() function show the order of the operations, namely checks, AMEs, indirect and direct effects, bootstrapped confidence intervals, and formatting of the output.

Computation of the Marginal Effect

MarginalMediation uses built-in R functionality that allows for relatively fast computation of the marginal effects. The approach taken here is identical to that of the margins R package (Leeper, 2017), as described in Chapter 3. This is repeated here. Specifically, for continuous predictors, the numerical derivative is used as shown below where a is the general symbol for the model estimates.

$$AME_k = \frac{1}{n} \sum_{i=1}^n \frac{f(aX_1) - f(aX_2)}{2h}$$
 (5.7)

| Table 5.1: Functions used in the Marginal Mediation R pa |
|--|
|--|

| Type | Function | Behavior |
|--------------------|----------------------------------|---------------------------------|
| Main Function | mma() | Performs the full MMA model |
| Marginal Function | amed() | Computes the average |
| | | marginal effects of a given |
| | | GLM model |
| Moderated Media- | <pre>mod_med()</pre> | Computes the marginal ef- |
| tion | | fects at various levels of a |
| | | moderator (still in testing) |
| Checks and format- | These functions perform | Check model specification and |
| ting | behind the scenes | function requirements |
| Other Functions | <pre>mma_std_ind_effects()</pre> | Obtain the standardized in- |
| | and | direct and direct effects from |
| | <pre>mma_std_dir_effects()</pre> | the model |
| Other Functions | ${\tt mma_ind_effects()}$ and | Obtain the unstandardized |
| | <pre>mma_dir_effects()</pre> | indirect and direct effects |
| | | from the model |
| Other Functions | <pre>perc_med()</pre> | Obtain the percent of media- |
| | | tion for each specified path in |
| | | the model |

where

$$aX_{1} = \begin{bmatrix} ax_{11} & ax_{12} & \dots & ax_{1k} + h & \dots & ax_{1p} \\ ax_{21} & ax_{22} & \dots & ax_{2k} + h & \dots & ax_{2p} \\ \vdots & \vdots & \ddots & \vdots & \ddots & \vdots \\ ax_{n1} & ax_{n2} & \dots & ax_{nk} + h & \dots & ax_{np} \end{bmatrix}$$

$$(5.8)$$

and

$$aX_{2} = \begin{bmatrix} ax_{11} & ax_{12} & \dots & ax_{1k} - h & \dots & ax_{1p} \\ ax_{21} & ax_{22} & \dots & ax_{2k} - h & \dots & ax_{2p} \\ \vdots & \vdots & \ddots & \vdots & \ddots & \vdots \\ ax_{n1} & ax_{n2} & \dots & ax_{nk} - h & \dots & ax_{np} \end{bmatrix}$$

$$(5.9)$$

With a small h (default is 1×10^{-7}), this produces the average marginal effect across

all the observations (e.g., the average change in the predicted value for a very small increase and a very small decrease in in the x_k variable).

For discrete predictors, the discrete difference is used as shown below,

$$AME_k = \frac{1}{n} \sum_{i=1}^{n} \left[F(\beta X | x_{ki} = 1) - F(\beta X | x_{ki} = 0) \right]$$
 (5.10)

where $F(\beta X|x_{ki}=1)$ is the predicted value of the *ith* observation when the dummy variable x_k equals one and $F(\beta X|x_{ki}=0)$ is the predicted value when the dummy value of x_k equals zero holding all other variables constant. This, in effect, shows the discrete difference between the levels of the categorical variable in the outcome's original units.

These approaches are employed in MarginalMediation due to their flexibility across GLM types and model specifications. For example, it can handle many types of models (e.g., linear, GLM, multilevel) and can produce more interpretable estimates of the marginal effects of predictors that have quadratic terms (e.g., age and age^2).

Standardization

As was briefly noted earlier, partial standardization wherein the outcome is standardized is possible when the outcome is continuous. In these situations, the output of MarginalMediation will include both unstandardized and standardized effects (see Figure 5.4).

Examples of Software Use

To briefly demonstrate the use of mma(), fictitious data were first generated, where X, M, and Y are continuous. Using these data (called df1), the following R code demonstrates the use of mma() in the simplest case.

First, the individual sub-models are fit thereby creating pathbc and patha which are both glm objects. Then, the b and c paths (pathbc) model object is the first argument to mma(), followed by the a paths (in this case only a single a path but multiple—separated by commas—can be included). The necessary argument is the ind_effects. This argument expects a vector or list of quoted paths, where the paths are the form "predictor-mediator". In this case, the predictor is called X and the mediator is called M.

The fit object as created by mma() contains a number of elements, including the indirect effects, the direct effects, the confidence interval, and the original data. Figure 5.4 provides an example of how the output could look if the fit object is printed. This output provides both unstandardized effects (both indirect and direct) that are in the units of the outcome and standardized effects—using the standard deviation of the outcome as recommended by MacKinnon (2008)—which are in the standard deviation units of the outcome.

Further, it can be assessed whether, in this case, the indirect plus the direct effects equal the total effect. Here, the total effect is 1.921 which is equal to the indirect effect (0.885) plus the direct effect (1.036). This suggests that comparisons between the effects can be confidently made.

If a covariate, X2 is added to the data, this can be easily added to the model as shown below.

```
Marginal Mediation Analysis
A marginal mediation model with:
  1 mediators
  1 indirect effects
  1 direct effects
  100 bootstrapped samples
  95% confidence interval
  n = 100
Formulas:
  ○ Y ~ X + M
  0 M ~ X
Unstandardized Effects

    Indirect Effects —

    A-path B-path Indirect Lower Upper
X-M 0.94825 0.93282 0.88454 0.64805 1.16212
— Direct Effects —
 Direct Lower Upper
X 1.0358 0.8144 1.24912
Standardized Effects
— Indirect Effects —
   Indirect Lower Upper
X-M 0.36097 0.26446 0.47425
— Direct Effects —
 Direct Lower Upper
X 0.4227 0.33235 0.50976
```

Figure 5.4: An example of the output from the mma() function.

It is also possible to access various aspects of these MMA model fit objects.

```
perc_med(fit2, "X-M")
```

This informs the researcher that the indirect effect accounts for approximately 62%

of the total effect from X to Y in fit2.

Monte Carlo Simulation Study

With the software package MarginalMediation, the simulations were able to assess the package's functionality and the overall framework's ability to estimate the underlying effects accurately. First, to assess the appropriateness of the experimental conditions, a literature review was conducted.

Literature Review

Studies were sought that saliently reported results wherein both mediation analysis and generalized linear models were used. Since 2012, this produced 57 articles (via Scopus).² Among these, three general categories of articles were found:

- 1. Articles that were methodologically building on mediation analysis.
- Articles that applied mediation where a mediator and/or outcome was categorical and the authors used the "difference method" (MacKinnon, 2008) to assess the amount of mediation.
- 3. Articles that applied mediation where a mediator and/or outcome was categorical and the authors either used the structural equation modeling approach or did not include the categorical mediator and/or outcome in the mediation.

Of these, number two was most prevalent. The literature suggested that the parameters selected for the simulations were relevant, particularly the small effect sizes and large sample sizes. Most studies used extant, large questionnaire data sets and the majority were cross-sectional.

Importantly, this search demonstrated the commonality of the "difference method" as discussed by MacKinnon (2008). This method relies on the following:

²The search terms included: "mediation analysis" and ["logistic" or "generalized linear models" or "GLM" or "poisson"].

$$a \times b + c' = c \tag{5.11}$$

$$a \times b = c - c' \tag{5.12}$$

In essence, this says that it is possible to estimate the indirect effect, that is $a \times b$ by assessing the difference c-c'. However, in situations where the decomposed total effect does not equal the total effect, this method may not be valid although many studies still used this approach in categorical data situations.

Simulations

The Monte Carlo simulation produced 45,000 marginal mediation models (although including the bootstrapped intervals there were 22.5 million models run). These simulated models were run on powerful Core i7 computers over the span of several days. The following subsections discuss the results of the simulations in regard to each outcome of interest.

Decomposed Total Effect Equals The Total Effect

One of the major questions about the performance of MMA regards whether the decomposed total effect equals the total effect $(a \times b + c' = c)$. Table 5.2 high-lights the average discrepancy between the decomposed total effect and the total effect divided by the total effect (thereby adjusting the discrepancy for the size of the total effect). Clearly, on an average level, deviations are extremely small, generally < .5% discrepancy, with the majority < .1% discrepancy. The discrepancies also decrease in size as the sample size increases.

Figure 5.5 presents the individual simulated differences between the decomposed total effect and the total effect. Once assessing the individual discrepancies, two patterns are of note:

1. There are larger discrepancies for smaller sample sizes and larger effect sizes.

Table 5.2: Average discrepancies between the decomposed total effect and the total effect for various sample sizes.

| | Mediator | | |
|-------------|----------|---------|--|
| Sample Size | Binary | Count | |
| 50 | 0.0017 | -0.0080 | |
| 100 | 0.0024 | -0.0028 | |
| 200 | 0.0007 | 0.0027 | |
| 500 | 0.0001 | -0.0036 | |
| 1000 | -0.0000 | -0.0027 | |

2. Besides a single outlier in the count condition (Panel b), most discrepancies are small.

First, the largest discrepancies are where the sample sizes are small (n = 50) and the effect sizes are larger. This is intuitive in that as the effect size is larger, the amount of discrepancy that is still considered small also increases (i.e., variability simply due to the estimates being on a larger scale). For both binary and count mediators, the discrepancies, even in the large effect sizes, are very small as the sample increases to n = 1000. Given the literature review, this is a sample size that is often possible in the health and prevention sciences. Further, most effect sizes in the literature were moderate or smaller. These conditions had low variability in the discrepancy.

Second, in the count condition there is a clear outlying value (>2 for the n = 50 and large/large effect size condition). Other than this value—across the binary and count mediators—all other values are relatively close to zero. For the binary mediator condition, the scale was in risk (probability) units. The discrepancies, here, in the n = 50 condition are notable in their size while the other conditions had discrepancies that are essentially within rounding error. For the count mediator condition, the scale of the total effect was in count units. The outlier is notable in its large discrepancy in these units; most other values were essentially within rounding error of the effect size.

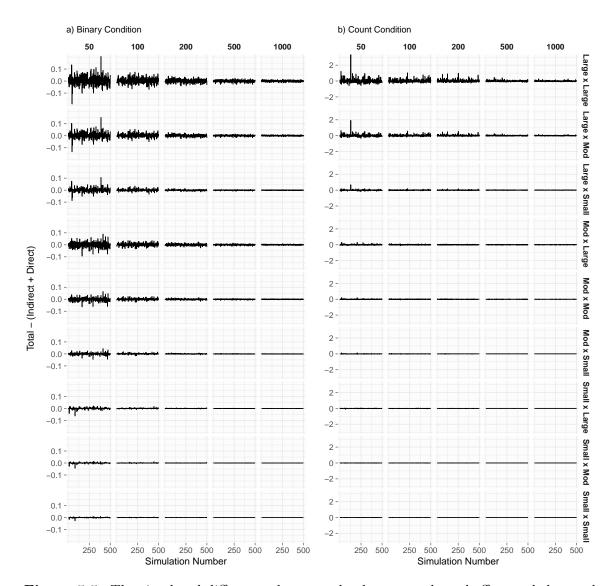


Figure 5.5: The simulated differences between the decomposed total effect and the total effect. The discrepancies are higher for smaller sample sizes and larger effect sizes.

Ultimately, this provides evidence of MMAs ability to estimate values that let the $a \times b + c' = c$ condition to hold, even in individual applications. This, however, is somewhat dependent on the sample size. As for differences across the effect sizes of the a and b paths, as an effect size increases so does the level of "rounding error." That is, in large effects, larger discrepancies are still a minor deviation than if the effect was small. Therefore, the main aspect of this finding is that sample size is important in the accuracy of the indirect plus direct equaling the total effect.

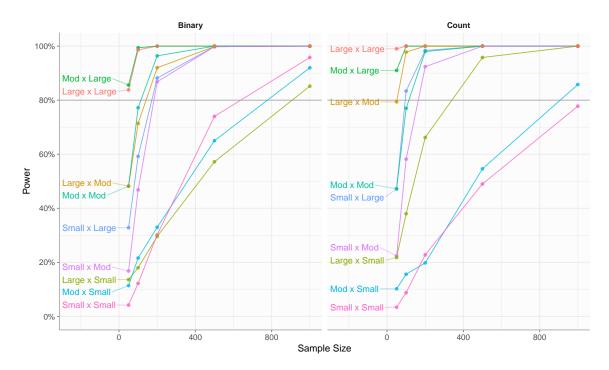


Figure 5.6: The simulated levels of power per tested sample size (x-axis) and effect size of the indirect path (color; a combination of the a path by the b path) stratified by the distribution of the mediator (binary or count).

Statistical Power

Figure 5.6 shows the statistical power of MMA across the various conditions. The figure shows the statistical power at each tested sample size for each combination of effect sizes (e.g., "Mod x Large" is a moderate a path effect size and a large b path effect size). Overall, most effect size combinations are adequately powered at a sample size of 200 across both binary and count mediator conditions. Interestingly, the "Small x Small" condition had more power at higher sample sizes than "Large x Small", which is contrary to intuition. However, the issue here was the issue of *complete separability* wherein the estimates and the standard errors are either biased or not estimable in logistic regression. With a large effect and a large sample size, this became common, thus reducing the statistical power in these conditions. The count mediator condition did not have this issue.

Overall, the method has the statistical power for even very small indirect effects with a sample size of 1000. As mentioned before, sample sizes greater than 1000 are common in the literature suggesting the method can be used even to detect small effect sizes.

Estimation Accuracy

It is also important for MMA to estimate the expected parameters. Figure 5.7 highlights that MMA is consistent in estimating the underlying effects for each combination of effect sizes across the various sample sizes. In the figure, which is stratified by the combination of effect sizes, shows the population parameter (vertical lines) and the estimated values (the density distributions). Overall, the distributions are centered at the true population parameter in each situation across the conditions.

As also seen in Figure 5.5, there is more variability in the estimation for larger effect sizes than for smaller. Again, this variability is likely due to the estimates being on a larger scale.

Confidence Interval Coverage

Finally, the confidence interval coverage is shown in Figure 5.8. Panel a) of the figure shows the overview—that the confidence interval coverage is around the 95% line for both the binary and count mediator conditions. However, looking at it much more closely in Panel b) it is clear that there is some deviation from the 95% line, particularly in the binary mediator condition. This is not a major deviation but an important one, nonetheless. Given the use of the percentile bootstrapping method herein, it may be important to apply other bootstrapping approaches such as the Bias-Corrected Bootstrap.

This finding of the indirect effect having confidence intervals that were too

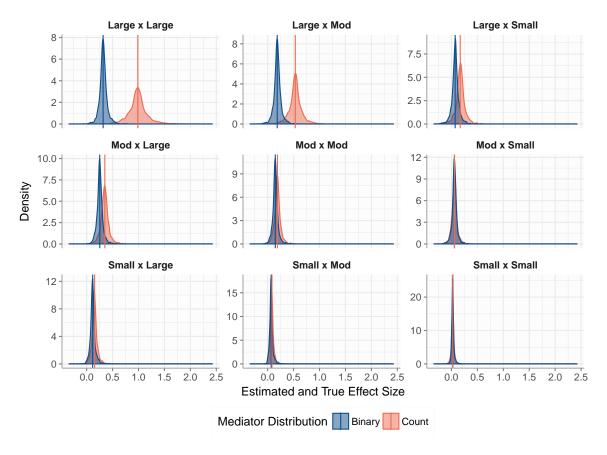


Figure 5.7: The simulated accuracy per tested sample size (x-axis) and effect size of the indirect path (color; a combination of the a path by the b path).

narrow has been found previously for the percentile bootstrapped (as applied in MMA; MacKinnon, Lockwood, and Williams, 2004). However, MacKinnon et al. (2004) also found the bootstrap methods, including the percentile approach, is among the best of the tested approaches. Other approaches, including the Monte Carlo confidence interval can be tested in future studies.

Conclusion

Marginal Mediation Analysis shows promise in its ability to accurately estimate models wherein the mediator is a binary or a count variable. Results regarding the decomposed total effect equaling the total effect are positive, although the estimation accuracy of this relationship depends on the sample and effect sizes. The

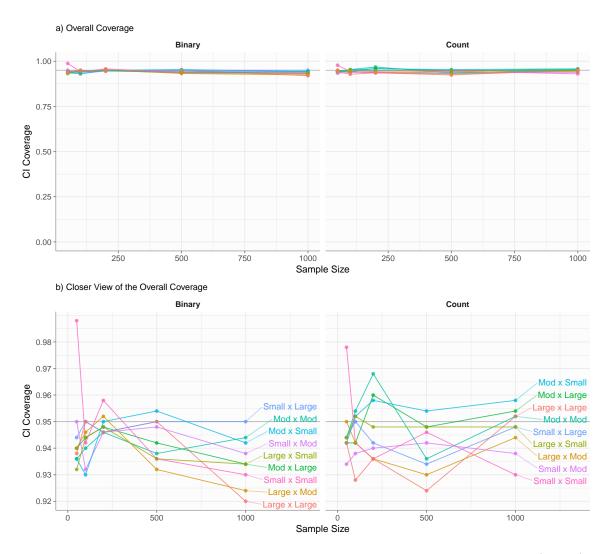


Figure 5.8: The simulated confidence interval coverage per tested sample size (x-axis) and effect size of the indirect effect. The "ideal" level is set at 0.95. Panel a) shows that the confidence intervals from a broad perspective. Panel b) provides a closer look at the individual patterns of the confidence intervals around the ideal level of 0.95.

statistical power is comparable to other modern mediation techniques wherein even small effect sizes can be estimated with a sample size of 1000. The estimation is consistent, ultimately averaging at the true population value. The confidence interval coverage was often too narrow for the binary mediator condition—sometimes having coverage of just above 92%—but it was approximately correct for the count mediator condition with some variability around 95%. Finally, the software for MMA is free to use in the R statistical environment in the MarginalMediation

package. This allows researchers to begin to use the approach with little overhead. All in all, MMA appears to be a practical approach to difficult mediation situations.

CHAPTER 6

PHASE III: APPLICATION OF MMA

The power of intuitive understanding will protect you from harm until the end of your days. — Lao Tzu

Introduction

In 2012, Ford and Hill published an article that used some of the most common approaches to mediation when a mediator and/or outcome is categorical. Specifically, they used:

- 1. the difference method (MacKinnon, 2008),
- 2. the "categorical data method outlined by MacKinnon (2008)" (pg. 5) to assess the significance of the difference method, and
- 3. the percent of the total effect that was mediated.

These three approaches are not only common but likely some of the best approaches in this situation. However, as stated in Chapter 4, these have some notable shortcomings. First, the standard errors can be inefficient and biased if there is a high degree of multi-collinearity (or the degree to which there is perfect separability) in any of the models. The significance of the difference method depends on these standard error estimates. Second, it does not provide the effect size measures that would be most useful [e.g., the effect of increasing the predictor on the outcome through the mediator(s)]. Third, the difference method is consistently too conservative with binary outcomes (Jiang and Vanderweele, 2015).

To build on the important findings from Ford and Hill (2012), this study replicates their work using more recent data from 2014 while using MMA to obtain ef-

¹Perfect separability is where a predictor can perfectly predict the outcome in logistic regression.

fect sizes and confidence intervals for the indirect and direct effects.

Results

```
## Error in eval(expr, envir, enclos): object 'da36361.0001' not found
## Error in tolower(names(d)): object 'd' not found
## Error in eval(lhs, parent, parent): object 'd' not found
## Error in map(.x, .f, ...): object 'd1' not found
## Error in map lgl(.x, .p, ...): object 'd1' not found
## Error in eval(lhs, parent, parent): object 'd1' not found
## Error in library(here): there is no package called 'here'
## Error in here("Data/NSDUH_2014_Results.rda"): could not find function "here"
## Error in library(survey): there is no package called 'survey'
## Error in svydesign(ids = ~1, strata = ~vestr, weights = ~analwt_c, data = d1):
## Error in svyglm(self ~ religious + age2 + irsex + newrace2 + irfamin3 + : could
## Error in svyglm(peer ~ religious + age2 + irsex + newrace2 + irfamin3 + : could
## Error in svyglm(dep ~ religious + age2 + irsex + newrace2 + irfamin3 + : could
## Error in coef(obj): object 'svy a1' not found
## Error in rownames(est1) = c("Respondent", "Peer", "Depression"): object 'est1'
## Error in data.frame(est1): object 'est1' not found
## Error in svyglm(model, design = design, family = "quasibinomial"): could not fi
## Error in UseMethod("vcov"): no applicable method for 'vcov' applied to an object
## Error in rownames(est2) = c("Tobacco", "Rx", "Marijuana", "Illicit"): object 'e
## Error in data.frame(est2): object 'est2' not found
```

The descriptive statistics are found in Table 6.1 for the 13,600 adolescents in the sample. Overall, these sample statistics were very similar to the 2007 sample used by Ford and Hill. However, the prevalence of drug use across each category dropped since 2007, although marijuana use did not drop substantially (13.85% in

2007 to 12.7% in 2014). Heavy drinking (10.4% in 2007) had only a single positive response in the entire sample of adolescents in 2014. Unfortunately, the number of major depressive episodes increased from 8.4% in 2007 to 11.3% in 2014. Attitudes regarding drug use were essentially identical as that in 2007 for the respondent, peer, and parent (and each had high reliabilities—all $\alpha \geq .80$ —comparable to 2007). Finally, the attitudinal measures and the measure of religiosity had high reliabilities.

Table 6.1: Descriptive statistics of the sample.

| | $\begin{array}{c} \text{Mean/ Percent (SD)} \\ \text{n} = 13,600 \end{array}$ |
|--------------------------------------|---|
| Drug use | |
| Prescription drug misuse (past year) | 5.7% |
| Tobacco use (past year) | 11.8% |
| Heavy drinking (past 30 days) | 0% |
| Marijuana use (past year) | 12.7% |
| Other illicit drug use (past year) | 3.5% |
| Demographics | |
| Female | 49.0% |
| Race (Non-White) | 45.8% |
| Income (2x poverty level) | 55.4% |
| Major Depression Episode | 11.3% |
| Attitudinal measures | |
| Respondent (range 1-3) | $2.6 \ (\alpha = .86)$ |
| Peer (range 1-3) | $2.5 \; (\alpha = .88)$ |
| Parent (range 1-3) | $2.9 \; (\alpha = .84)$ |
| Religiosity | $\alpha = .80$ |

Four MMA models were used to assess the pathways from adolescent religiosity to substance use, one for each outcome (any tobacco use, prescription drug misuse, marijuana use, and other illicit drug use). Each model controlled for parental conservative attitudes toward substance use, the adolescents' family income, and the adolescents' age, race, and sex. Figure 6.1 presents the individual paths in the

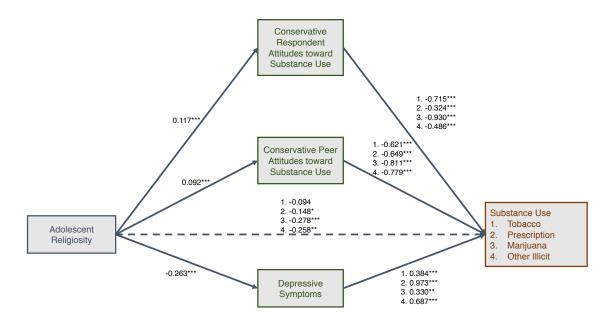


Figure 6.1: Results of the mediation models' individual paths regarding religiosity and substance use. Note: *** p < .001, ** p < .01, * p < .05

model units. Therefore, the paths leading to the conservative attitudes (both respondent and peer attitudes) are in the attitude metric with a range from 1 - 3. The paths leading to depression and the substance outcomes are all in log-odds. As the figure highlights, most paths were statistically significant at p < .05.

Because MMA provides information about each of the indirect effects naturally in the same units, it is possible to assess the amount mediated by each mediator while also controlling for the other mediators in a straightforward manner—without having to fit several other models and assess each c - c'. Table 6.2 presents the amount of the total effect of religiosity on substance use that is mediated through respondent conservative attitudes, peer conservative attitudes, and depression. Overall, the effect of religiosity on substance use is heavily mediated by the hypothesized mediators, more so for tobacco use than the others.²

In addition to this information, MMA provides information regarding the in-

²Using the approach used in Ford and Hill, the total mediated effects are slightly different than those estimated via the indirect and direct effects. This may be due to the weighting of the sample; an important area for further research into MMAs performance.

Table 6.2: The percent of mediation (the percent of the total effect) by path for each outcome.

| | Outcome | | | |
|------------------|---------|--------------|-----------|---------|
| Mediator | Tobacco | Prescription | Marijuana | Illicit |
| Respondent Views | 34.2 | 14.0 | 23.2 | 14.1 |
| Peer Views | 23.2 | 22.0 | 15.8 | 17.7 |
| Depression | 4.0 | 9.2 | 1.8 | 4.4 |
| Total Mediation | 61.4 | 45.2 | 40.8 | 36.1 |

direct and direct effects in the same units. Figure ?? highlights the indirect and direct effects with their associated 95% confidence intervals in the average marginal effects. All of the effects here are in risk (probability) units—i.e., risk of tobacco use, prescription misuse, marijuana use, or illicit drug use. Although all indirect effects and most direct effects are significant, the effect size estimates are particularly important here as the meaningfulness of these significant effects can be overlooked.

These resulting effects are all small, with most effects less than 0.01 (i.e., less than a single risk unit). That is, most effects show changes in the risk of the outcome by less than a single unit. For example, if adolescent religiosity is increased by one unit, its effect on the risk of tobacco use, through respondent attitudes, is a decrease of 0.007; through peer attitudes a decrease of 0.005; through depression a decrease of 0.001; and directly a decrease of 0.008. The total effect, then, is approximately -0.021. Therefore, if an individual has a risk of using tobacco at 50%, by increasing religiosity by one unit (holding the covariates constant), on average that individual's risk would decrease to 47.9%. About 0.013 of the effect of religiosity on tobacco use is mediated while 0.008 is direct from religiosity. Ultimately, these findings highlight the fact that the effect sizes, especially the indirect and direct effect sizes, are valuable companions to the p-values.

Error in loadNamespace(name): there is no package called 'anteo'

Conclusions

The replication highlighted several important facets of the important work by Ford and Hill (2012). First, it simplifies the interpretation of the model by using average marginal effects. Second, it highlighted the effect sizes in terms of risk of substance use. This allowed the relatively small effects to be understood, not only in their significance, but in their meaning. Ultimately, MMA provided a more straightforward approach and substantially more information for each model than other mediation approaches.

CHAPTER 7

DISCUSSION

A model is a simplification or approximation of reality and hence will not reflect all of reality. ... While a model can never be "truth," a model might be ranked from very useful, to useful, to somewhat useful, to, finally, essentially useless. — Burnham and Anderson, 2002

General Discussion

Models are simply representations of reality. This is also true of mediation models even though they are used to model more complex relations. The value in using such models is generally seen in their ability to provide opportunities for intervention or prevention.¹

As has been discussed throughout this project, mediation models are most useful when the model communicates both the significance (e.g., p-values) and meaningfulness (e.g., effect sizes). One without the other can be misleading, potentially resulting in faulty interventions and policies. Together, significance and meaning tell a more complete story of the data The significance helps researchers understand uncertainty; effect sizes communicate the potential for intervention to actually make meaningful changes in important outcomes.

However, some situations wherein mediation analysis is applied can provide a lack of interpretable information, particularly in terms of the effect sizes. This limits the usefulness of the model, whether or not it is an accurate representation

¹Although not a major aspect of this project, it is important to note the predictive power a specific model has. For example, a model may show a significant effect in a logistic regression but it may not predict the outcome above chance. In this case, the question turns from "does X have an effect on Y?" to "does it predict Y?" This distinction has important implications when it comes to intervention and prevention but involves a number of important concepts that cannot be covered herein (e.g., model specification including penalty parameters, modeling approach such as tree based approaches). This general idea—that of the importance of predictive accuracy—is discussed in Hastie, Tibshirani, and Friedman (2009).

of reality. It is for this purpose that Marginal Mediation Analysis (MMA) was developed. It provides mediation analysis with the tools to communicate both significance and meaning. This is coming forth at an opportune time, as the American Psychological Association, among others, have called for more focus on effect sizes and less attention on p-values (Cumming, 2014).

Findings from the Three Phases

This project has presented the development of the approach and its software, the evaluation of its performance in possible real-world scenarios, and the application of it to health data regarding adolescent substance use. In its first phase, this project produced MMA with its accompanying software—the MarginalMediation R package. The software is freely available and allows for researchers to quickly apply it. The main function, mma(), is relatively quick even with the bootstrapping, and produces thorough output.

The next phase used Monte Carlo simulations to evaluate MMA and ways that it can possibly be improved. For example, given the results regarding the confidence interval coverage, it may be of benefit to try alternative approaches, either adaptations of bootstrapping (e.g., Bias-Corrected Bootstrap) or others. MacKinnon et al. (2004) found that Monte Carlo confidence intervals performed well and, therefore, may also be a valuable addition to MMA.

These simulations further demonstrated a trade-off regarding sample sizes and effect sizes: large effect sizes can be found in small samples but those same conditions provide much more variability in estimating the total effect accurately. Overall, these findings demonstrate the ability for the sample size, as it increases, to reduce bias and solidify relationships that should hold in mediation models (e.g., $a \times b + c' = c$).

The Monte Carlo simulations also allowed for the testing of the software.

Some situations, once simulated, demonstrated a need for change, often regarding the speed of the software, its accuracy, and necessary checks to avoid more serious problems. Ultimately, there was a natural feedback loop between the simulations and the software that were developed interactively. Once a stable version of the software was achieved, the reported simulations were all run based on that version (v0.5.0).

In the final phase, the application study highlighted important information regarding the MMA approach and adolescent health. The application study replicated work by Ford and Hill (2012), which was chosen to replicate for three major reasons:

- 1. the application study used a large sample with a mix of binary and continuous mediators and outcomes (common in the literature),
- 2. the statistical approach is one of the better approaches (also common in the literature), and
- 3. the data were open and a more recent release was available to investigate.

Although MMA can benefit the researcher in many situations, the benefits of using MMA are particularly clear within the context of this application study. Most importantly, MMA provided more information, in the form of effect size estimates, that help instruct on the meaningfulness of the results (Cumming, 2014). As Preacher and Kelley (2011) state: "it is important to develop a way to gauge the effect size of the product term ab itself," (pg. 95). That is, not only does the effect size of the individual paths need to be meaningful but the product of $a \times b$ must be as well. Although nearly all effects were significant, with such a large sample size significance tests alone can be misleading. For this study, the addition of the effect sizes are helpful to understand that each estimated effect was small. This provides a more complete view of the relationships tested herein.

Second, in terms of the substantive findings, there is strong evidence across

many studies that adolescent religiosity is related to substance use. This is shown here as well. Consistently, religiosity was negatively related to the four substance use outcomes. About half of the total effect of religiosity on substance use was mediated by personal and peer attitudes about substance use. Depression also mediated the relationship, but to a much lesser degree.

Although not definitive, this study in conjunction with Ford and Hill (2012) presents evidence that religiosity may impact substance use outcomes through attitudes towards substance use. More research, particularly research with longitudinal data, are needed to further test and understand these relationships and their ability to inform intervention or policy.

Limitations

The MMA approach has two notable limitations. First, mediation analysis assumes no measurement error in the mediators. Although latent variable methods can help with this (Iacobucci, 2008; Lockhart et al., 2011), the data necessary are not always available and the integration of average marginal effects within SEM is not clearly defined as of yet. Ultimately, the estimates are only as good as the measurements. Second, it may also be difficult for researchers to accept given the novelty of average marginal effects in the field. This is being alleviated through the use of various introductions to average marginal effects and its use in other fields (Barrett & Lockhart, in preparation).

Of course, the Monte Carlo simulation did not test for all conditions present in real-world modeling. Although it accounted for the main influences, there are other possible important influences that may impact the performance of the method, including missing values and model mis-specification. These are important influences to assess in future projects. Finally, the application study used cross-sectional data. This makes it difficult to demonstrate causality and puts additional

pressure on the ability to control for confounding.

Future Research

Several foreseeable areas of investigation can prove useful for understanding and extending MMA. First, the application highlighted an important area for future inquiry—MMA with survey weighted data. The application study used data that were collected via a complex survey design and were therefore weighted. Further research is needed to understand MMAs behavior in these situations.

Second, this project specifically assessed binary and count mediators. Another important type of variable that could play an important role in mediation is "time-to-event" or survival data. Future research is needed to understand how this type of data with its accompanying statistical approaches can fit into MMA.

Third, MMA relies on the *sequential ignorability* assumption as described by Imai et al. (2010a). A sensitivity analysis is available to assess how important deviations from this assumption are on the estimates and conclusions (Imai et al., 2010a, 2010b). Integrating this sensitivity analysis would be a valuable addition to the approach. It likely would be a natural integration but this integration would need to be tested.

Relatedly, it could also be useful to look at using instrumental variables to help appease sequential ignorability. Although generally not applied in conjunction with mediation analysis, the approach could prove useful for MMA specifically and mediation analysis as a whole.

Lastly, the integration of latent variable approaches, including latent class analysis, is an important step in making this approach more broadly applicable. Work regarding average marginal effects, categorical data, and structural equation models would be an important contribution as well.

Conclusions

The results of the development, simulations, and application all show that MMA holds much promise in extending mediation analysis more fully to situations where the mediators and/or outcomes are categorical or non-normally distributed. Although further work is necessary to understand MMAs performance across more situations, the results of this project demonstrates its utility for common health and prevention research.

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APPENDICES

Appendix A: R Code for Chapter 5

Required: R Packages from CRAN

```
if (!require(tidyverse)){
   install.packages("tidyverse")
   library(tidyverse)
}
if (!require(furniture)){
   install.packages("furniture")
   library(furniture)
}
if (!require(here)){
   install.packages("here")
   library(here)
}
if (!require(devtools)){
   install.packages("devtools")
   library(devtools)
}
```

Required: R Packages from GitHub

```
if (!require(MarginalMediation)){
  devtools::install_github("tysonstanley/MarginalMediation")
  library(MarginalMediation)
}
```

Examples from Chapter 5

Figure 5.2 on page 42

```
set.seed(843)
library(tidyverse)
tibble::data_frame(
   X = rnorm(100),
   W = rbinom(100, 1, .5),
   Y = X + .5*W + -.5*X*W + rnorm(100)
) %>%
   ggplot(aes(X, Y, group = factor(W), color = factor(W))) +
        geom_point(alpha = .3) +
        geom_smooth(method = "lm", se=FALSE) +
        scale_color_manual(values = c("dodgerblue4", "chartreuse4")) +
        labs(color = "Moderator") +
        anteo::theme_anteo_wh()
ggsave("figures/fig_interaction_effect.pdf",
        width = 7, height = 5, units = "in")
```

Monte Carlo Simulation

Notably, the code for both the binary mediator condition and the count mediator condition we run via the Terminal as, once the directory was where the R file was located:

```
Rscript Analyses_MMMC_scriptBinary.R 'c(1:45)'
and
Rscript Analyses_MMMC_scriptCount.R 'c(1:45)'
```

Binary Mediator

```
## Marginal Mediation: Monte Carlo Simulation Study
     BINARY Mediator
## Tyson S. Barrett
##
## devtools::install_github("tysonstanley/MarginalMediation")
args <- commandArgs(TRUE)</pre>
args <- eval(parse(text = args))</pre>
library(MarginalMediation)
library(tidyverse)
## Create all combinations of independent variables
cond_binary = expand.grid(
  samplesize = c(50, 100, 200, 500, 1000),
           = c(.55, 1.45, 2.22),
            = c(.24, .62, 1.068),
  effectb
            = c(.3)
  effectc
)
## Population Models
## Binary Mediator
data_genB = function(ps, reps, samplesize, effecta, effectb, effectc){
  set.seed(84322)
  Xc = rnorm(ps)
  z = effecta*Xc + rnorm(ps, 0, 1)
  pr = 1/(1+exp(-z))
  M = rbinom(ps, 1, pr)
  Y = effectb*M + effectc*Xc + rnorm(ps, 0, 1)
  M = factor(M)
  df = data.frame(Y, M, Xc)
  bin = vector("list", reps)
  print(cbind(samplesize, effecta, effectb))
  print(lm(Y ~ M + Xc)$coefficients)
  print(lm(scale(Y) ~ M + Xc)$coefficients)
  med = amed(glm(M ~ Xc, df, family = "binomial"))
```

```
for (i in 1:reps){
    d = df[sample(ps, samplesize), ]
    pathbc = glm(Y \sim M + Xc, data = d)
    patha = glm(M ~ Xc, data = d, family = "binomial")
    bin[[i]] = mma(pathbc, patha,
          ind_effects = c("Xc-M"),
          boot = 500)
    bin[[i]] = list("IndEffects" = bin[[i]]$ind_effects,
                    "DirEffects" = bin[[i]]$dir_effects,
                    "Boot"
                                = bin[[i]]$boot,
                                 = lm(Y ~ Xc, d)$coefficients,
                    "Total"
                    "MedSize" = med)
    cat("\r", i)
  print(exp(glm(M ~ Xc, family = "binomial")$coefficients))
  return(bin)
}
i = 0
for (j in args){
  set.seed(84322)
  i = i + 1
  cat("\nNumber:", j, "\n\n")
  out = data_genB(1e6, 500,
                  cond_binary[args[[i]],1],
                  cond_binary[args[[i]],2],
                  cond_binary[args[[i]],3],
                  cond_binary[args[[i]],4])
  save(out, file = paste0("Sims_Data/Binary2_",
                          cond_binary[args[[i]],1], "_",
                          cond_binary[args[[i]],2], "_",
                          cond_binary[args[[i]],3], "_",
                          cond_binary[args[[i]],4], ".rda"))
  cat("\nNumber:", j, "\n\n")
  cat("\nConditions Complete:\n",
      " Sample size =", cond_binary[args[[i]],1],
                   =", cond_binary[args[[i]],2],
      "\n A path
      "\n B path
                     =", cond_binary[args[[i]],3],
      "\n C path =", cond_binary[args[[i]],4], "\n")
}
```

Count Mediator

```
## Marginal Mediation: Monte Carlo Simulation Study
## COUNT Mediator
## Tyson S. Barrett
##
## devtools::install_github("tysonstanley/MarginalMediation")
args <- commandArgs(TRUE)</pre>
args <- eval(parse(text = args))</pre>
library(MarginalMediation)
library(tidyverse)
## Create all combinations of independent variables
cond_count = expand.grid(
  samplesize = c(50, 100, 200, 500, 1000),
  effecta = c(.3, .6, 1.1),
  effectb = c(.084, .265, .49),
  effectc = c(0, .3)
)
## Population Models
## Count Mediator
data_genC = function(ps, reps, samplesize, effecta, effectb, effectc){
  set.seed(84322)
  Xc = rnorm(ps)
 m1 = exp(effecta * Xc)
  M = rpois(ps, lambda=m1)
  Y = effectb*M + effectc*Xc + rnorm(ps, 0, 1)
  df = data.frame(Y, M, Xc)
  poi = vector("list", reps)
  print(cbind(samplesize, effecta, effectb))
  print(lm(Y ~ M + Xc)$coefficients)
  print(lm(scale(Y) ~ M + Xc)$coefficients)
  med = amed(glm(M ~ Xc, df, family = "poisson"))
  for (i in 1:reps){
    d = df[sample(ps, samplesize), ]
    pathbc = glm(Y \sim M + Xc, data = d)
    patha = glm(M ~ Xc, data = d, family = "poisson")
    poi[[i]] = mma(pathbc, patha,
                   ind_effects = c("Xc-M"),
                   boot = 500)
    poi[[i]] = list("IndEffects" = poi[[i]]$ind_effects,
                    "DirEffects" = poi[[i]]$dir_effects,
                    "Boot"
                               = poi[[i]]$boot,
                    "Total"
                               = lm(Y \sim Xc, d)$coefficients,
                    "MedSize"
                                = med)
```

```
cat("\r", i)
 }
 print(exp(glm(M ~ Xc, family = "poisson")$coefficients))
 return(poi)
}
i = 0
for (j in args){
 set.seed(84322)
  i = i + 1
  cat("\nNumber:", j, "\n\n")
  out = data_genC(1e6, 500,
                 cond_count[args[[i]],1],
                  cond_count[args[[i]],2],
                  cond_count[args[[i]],3],
                  cond_count[args[[i]],4])
  save(out, file = paste0("Sims_Data/Count2_",
                          cond_count[args[[i]],1], "_",
                          cond_count[args[[i]],2], "_",
                          cond_count[args[[i]],3], "_",
                          cond_count[args[[i]],4], ".rda"))
  cat("\nNumber:", j, "\n\n")
  cat("\nConditions Complete:\n",
      " Sample size =", cond_count[args[[i]],1],
      "\n A path =", cond_count[args[[i]],2],
                    =", cond_count[args[[i]],3],
      "\n B path
      "\n C path =", cond_count[args[[i]],4], "\n")
}
```

Monte Carlo Simulation Data Analyses

Data Preparations for tables and figures around page 53

```
options(na.rm=TRUE)
library(tidyverse)
library(furniture)
library(here)
es = read.csv("effect_sizes.csv") %>%
  data.frame(row.names = .$size) %>%
  select(-size)
filenames = list.files(here("Sims_Data/"),
                       pattern = ".rda")
tot = indc = indd =
  dirc = dird = vector("list", length(filenames))
for (i in filenames){
  cat("File:", i, "\n")
  load(paste0(here("Sims_Data/", i)))
  tot[[i]] = lapply(out, function(x) x$Total) %>%
    do.call("rbind", .) %>%
    data.frame %>%
    mutate(type = strsplit(i, " ")) %>%
    mutate(ss = map_chr(type, ~.x[2])) %>%
    mutate(dist = map_chr(type, ~.x[1])) %>%
    mutate(boot = map_chr(dist, ~ifelse(grep1("2$", .x), 500, 100))) %>%
    mutate(dist = map_chr(dist, ~gsub("2", "", .x))) %>%
    mutate(ap = map_chr(type, ~.x[3])) %>%
    mutate(bp = map_chr(type, ~.x[4])) %>%
    mutate(cp = map_chr(type, ~gsub("\.rda", "", .x[5]))) \%
    select(Xc, ss, dist, boot, ap, bp, cp)
  indc[[i]] = lapply(out, function(x) x$IndEffects[1, ]) %>%
    do.call("rbind", .) %>%
    data.frame %>%
    mutate(type = gsub(".rda", "", i)) %>%
    mutate(type = strsplit(type, "_")) %>%
    mutate(ss = map_chr(type, ~.x[2])) \%
    mutate(dist = map_chr(type, ~.x[1])) %>%
    mutate(boot = map_chr(dist, ~ifelse(grep1("2$", .x), 500, 100))) %>%
    mutate(dist = map_chr(dist, ~gsub("2", "", .x))) %>%
    mutate(ap = ifelse(dist == "Count",
                         ifelse(map_chr(type, ~.x[3]) == "0.3", "Small",
                         ifelse(map\_chr(type, ~.x[3]) == "0.6", "Mod",
                                "Large")),
```

```
ifelse(map_chr(type, ~.x[3]) == "0.55", "Small",
                         ifelse(map_chr(type, \sim .x[3]) == "1.45", "Mod",
                                "Large")))) %>%
               = ifelse(map_chr(type, ~.x[4]) == "0.24", "Small",
   mutate(bp
                  ifelse(map_chr(type, ~.x[4]) == "0.62", "Mod",
                  ifelse(map_chr(type, ~.x[4]) == "1.068", "Large",
                  ifelse(map_chr(type, ~.x[4]) == "0.084", "Small",
                  ifelse(map_chr(type, ~.x[4]) == "0.265", "Mod",
                         "Large")))))) %>%
    mutate(cp = map_chr(type, ~.x[5])) %>%
    mutate(power = ifelse(Lower > 0 & Upper > 0, 1, 0)) %>%
   mutate(ind_cat = paste(ap, "x", bp)) %>%
   mutate(true = ifelse(dist == "Count",
                          es[ind_cat, "ind_count"],
                          es[ind_cat, "ind_binary"])) %>%
    mutate(ap_size = ifelse(dist == "Count",
                            es[ind_cat, "a_count"],
                            es[ind_cat, "a_binary"])) %>%
   mutate(bp_size = ifelse(dist == "Count",
                            es[ind_cat, "b_count"],
                            es[ind_cat, "b_binary"])) %>%
    mutate(ci
                 = ifelse(true < Upper & true > Lower, 1, 0)) %>%
    select(-type)
 dirc[[i]] = lapply(out, function(x) x$DirEffects[1, ]) %>%
    do.call("rbind", .) %>%
    data.frame %>%
    mutate(type = gsub(".rda", "", i)) %>%
   mutate(type = strsplit(type, "_")) %>%
   mutate(ss = map_chr(type, ~.x[2])) %>%
   mutate(dist = map_chr(type, ~.x[1])) %>%
   mutate(boot = map_chr(dist, ~ifelse(grepl("2$", .x), 500, 100))) %>%
    mutate(dist = map_chr(dist, ~gsub("2", "", .x))) %>%
   mutate(ap = map_chr(type, ~.x[3])) %>%
   mutate(bp = map_chr(type, ~.x[4])) %>%
   mutate(cp = map_chr(type, ~.x[5])) %>%
   mutate(ci
               = ifelse(Lower > 0 & Upper > 0 & cp > 0, 1,
                  ifelse(Lower < 0 & Upper > 0 & cp == 0, 1, 0))) %>%
   mutate(power = ifelse(Lower > 0 & Upper > 0, 1, 0)) %>%
   mutate(true = cp) %>%
   select(-type)
}
ind1 = do.call('rbind', indc) %>%
 mutate(var = "continuous") %>%
 select(Indirect, Lower, Upper, ss, dist,
         boot, ap, bp, cp, ci, true, power,
```

Table 5.2 on page 53

Figure 5.5 on page 54

```
## Total = Total
total_total = cbind(tot1[,1], ind1[1:45000,1], dir1[1:45000,1]) %>%
 data.frame %>%
 set_names(c("total", "indirect", "direct")) %>%
 mutate(sample = ind1[1:45000, "ss"]) %>%
 mutate(sample = factor(sample,
                         levels = c("50", "100", "200", "500", "1000"))) %>%
 mutate(ap = ind1[1:45000,"ap"],
         bp = ind1[1:45000,"bp"],
         dist = ind1[1:45000, "dist"],
         true = ind1[1:45000,"true"]) %>%
 mutate(true = as.numeric(as.character(true))) %>%
 mutate(diff = (total - (indirect + direct))) %>%
 mutate(est = (total - (indirect + direct))/true) %>%
 mutate(eff = paste(ap, "x", bp)) %>%
 group_by(sample, eff, dist) %>%
 mutate(index = 1:n())
p1tot = total_total %>%
 filter(dist == "Binary") %>%
 ggplot(aes(index, diff)) +
   geom_path() +
    scale_x_continuous(breaks = c(250, 500),
                       labels = c("250", "500")) +
    scale_y_continuous(breaks = c(-.1,0,.1)) +
   facet_grid(eff~sample) +
   anteo::theme_anteo_wh() +
```

```
theme(panel.spacing = unit(.1, "cm"),
          strip.text.y = element_blank()) +
    labs(x = "Simulation Number",
         y = "Total - (Indirect + Direct) \n",
         subtitle = "a) Binary Condition")
p2tot = total_total %>%
  filter(dist == "Count") %>%
  ggplot(aes(index, diff)) +
  geom_path() +
  scale_x_continuous(breaks = c(250, 500),
                     labels = c("250", "500")) +
  scale_y_continuous(breaks = c(-2,0,2)) +
  coord_cartesian(ylim = c(-2.9, 3.1)) +
  facet_grid(eff~sample) +
  anteo::theme_anteo_wh() +
  theme(panel.spacing = unit(.1, "cm")) +
  labs(x = "Simulation Number",
       y = "",
       subtitle = "b) Count Condition")
plot_total = gridExtra::grid.arrange(p1tot, p2tot, ncol = 2)
ggsave("figures/fig_total_total.pdf",
      plot = plot_total,
       width = 10, height = 10, units = "in")
```

Figures 5.6, 5.7, and 5.8 on pages 55, 57, and 58, respectively.

```
## Accuracy, Power, Coverage
inds = ind1 \%
 mutate(accuracy = (as.numeric(Indirect) -
                       (as.numeric(ap_size) *
                          as.numeric(bp_size)))) %>%
 group_by(ss, dist, boot = as.numeric(boot), ap_size,
           bp_size, cp, var, ap, bp) %>%
 summarize(Ind = mean(Indirect),
           Low = mean(Lower),
           Hi = mean(Upper),
            ci = mean(ci),
            power = mean(power),
            acc = mean(accuracy)) %>%
 ungroup
dirs = dir1 %>%
 group_by(ss, dist, boot = as.numeric(boot), cp, var) %>%
 summarize(Dir = mean(Direct),
           Low = mean(Lower),
            Hi = mean(Upper),
            ci = mean(ci),
            power = mean(power))
```

```
ggplot(inds, aes(x = as.numeric(ss), y = power,
                 color = paste(ap, "x", bp),
                 group = interaction(ap, bp, boot, dist, cp))) +
  geom_hline(yintercept = .8, color = "darkgrey") +
  geom_line(alpha = .8) +
  geom_point(alpha = .8) +
  facet_grid(~ dist, space = "free", scales = "free") +
  anteo::theme_anteo_wh() +
  theme(panel.spacing = unit(.25, "cm"),
        axis.line = element_line(color = "darkgrey"),
        legend.position = "none") +
  scale_y\_continuous(breaks = c(0, .2, .4, .6, .8, 1),
                     labels = scales::percent) +
  labs(y = "Power",
       x = "Sample Size") +
  ggrepel::geom_text_repel(data = inds %>%
              filter(ss == 50),
            aes(label = paste(ap, "x", bp)),
            nudge_x = -150) +
  coord_cartesian(xlim = c(-250, 1000),
                  ylim = c(0,1))
ggsave("figures/sim_fig_power.pdf",
       width = 10, height = 6, units = "in")
ggplot(ind1, aes(x = Indirect, fill = dist,
                 color = dist,
                 group = interaction(ap, bp, boot, dist, cp))) +
  geom_density(alpha = .5) +
  geom_vline(aes(xintercept = true, color = dist)) +
  facet_wrap(~ paste(ap, "x", bp), scales = "free") +
  anteo::theme_anteo_wh() +
  theme(panel.spacing = unit(.25, "cm"),
        axis.line = element line(color = "darkgrey"),
        legend.position = "bottom") +
  labs(y = "Density",
       x = "Estimated and True Effect Size",
       fill = "Mediator Distribution",
       color = "Mediator Distribution") +
  scale_fill_manual(values = c("dodgerblue4", "coral2")) +
  scale_color_manual(values = c("dodgerblue4", "coral2"))
ggsave("figures/sim_fig_acc.pdf",
       width = 8, height = 6, units = "in")
p1 = ggplot(inds, aes(x = as.numeric(ss), y = ci,
                      color = paste(ap, "x", bp),
                      group = interaction(ap, bp, boot, dist, cp))) +
```

```
geom_hline(alpha = .8, yintercept = .95, color = "darkgrey") +
  geom_line(alpha = .8) +
  geom_point(alpha = .8) +
  facet_grid(~ dist, space = "free", scales = "free") +
  anteo::theme_anteo_wh() +
  theme(panel.spacing = unit(.25, "cm"),
        axis.line = element_line(color = "darkgrey"),
        legend.position = "none") +
  labs(y = "CI Coverage",
       x = "Sample Size",
       subtitle = "a) Overall Coverage") +
  coord_cartesian(ylim = c(0,1))
p2 = ggplot(inds, aes(x = as.numeric(ss), y = ci,
                      color = paste(ap, "x", bp),
                      group = interaction(ap, bp, boot, dist, cp))) +
  geom_hline(alpha = .8, yintercept = .95, color = "darkgrey") +
  geom_line(alpha = .8) +
  geom_point(alpha = .8) +
  facet_grid( ~ dist, space = "free", scales = "free") +
  anteo::theme_anteo_wh() +
  theme(panel.spacing = unit(.25, "cm"),
        axis.line = element_line(color = "darkgrey"),
        legend.position = "none") +
  labs(y = "CI Coverage",
       x = "Sample Size",
       subtitle = "b) Closer View of the Overall Coverage") +
  ggrepel::geom_text_repel(data = inds %>%
                             filter(ss == 1000),
                           aes(label = paste(ap, "x", bp)),
                           nudge_x = 250,
                           segment.alpha = .4) +
  coord_cartesian(xlim = c(0, 1350)) +
  scale_y = continuous(breaks = c(.92, .93, .94, .95, .96, .97, .98))
p3 = gridExtra::grid.arrange(p1,p2, ncol = 1)
ggsave("figures/sim_fig_ci.pdf",
      plot = p3,
       width = 10, height = 9, units = "in")
```

Appendix B: R Code for Chapter 6

Required: R Packages from CRAN

```
if (!require(tidyverse)){
  install.packages("tidyverse")
  library(tidyverse)
}
if (!require(furniture)){
  install.packages("furniture")
  library(furniture)
}
if (!require(here)){
  install.packages("here")
  library(here)
if (!require(devtools)){
  install.packages("devtools")
  library(devtools)
}
if (!require(survey)){
  install.packages("survey")
  library(survey)
}
```

Required: R Packages from GitHub

```
if (!require(MarginalMediation)){
  devtools::install_github("tysonstanley/MarginalMediation")
  library(MarginalMediation)
}
```

Data Preparation

Data preparation using the 2014 National Survey on Drug Use and Health, as described in Chapter 4.

```
library(tidyverse)
library(furniture)
## Load data
load("Data/NSDUH_2014_Results.rda")
d = da36361.0001
names(d) = tolower(names(d))
rm(da36361.0001)
## Variables
d1 = d \%
 select(
   ## ----- ##
   ## Outcomes
         (1,2,8,11,12 = within last year) ##
   ## ----- ##
   ## Tobacco Outcome
   cigrec, ## cig
   chewrec, ## chew
   cigarrec, ## cigar
   #pipe30dy, ## pipe (30 days here instead)
   ## Heavy Drinking Outcome
   dr5day, ## 1+ is within last 30 days
   ## Rx Outcome
   analrec, ## pain relievers
   tranrec, ## tranquilizers
   stimrec, ## stimulants
   sedrec, ## sedatives
   ## Marijuana Outcome
   mjrec, ## marijuana
   ## Other Illicit Outcome
   cocrec, ## cocaine
   crakrec, ## crack
   herrec, ## heroin
   hallrec, ## hallucinogens
   lsdrec, ## LSD
   pcprec, ## PCP
   ecsrec, ## ecstacy
   inhrec, ## inhalants
```

```
methrec, ## meth
## Mediators
## Mean response (higher = more cons) ##
## ----- ##
## Self Views Mediator
yegpkcig, ## someone your age cig
yegmjevr, ## someone your age mj
yegmjmo, ## someone your age mj monthly
yegaldly, ## someone your age drinking daily
## Peer Views Mediator
yefpkcig, ## you cig
yefmjevr, ## you mj
yefmjmo, ## you mj monthly
yefaldly, ## you drinking daily
## Psychological Well-being (Major Depression)
ymdeyr, ## past year major depressive epidosde (MDE)
## Predictor
                                   ##
## Cronbach's Alpha
## Standardized mean level
## ----- ##
## Religiosity
yerlgsvc, ## past 12, times at church (1-6
yerlgimp, ## religious beliefs are important (1-4 strong dis to strong agree)
yerldcsn, ## religious belief influence decisions (1-4)
yefaiact, ## religious activities
## Control Variables
                                  ##
## ----- ##
## Parental Attitudes
yeppkcig, ## parents feel about cig
yepmjevr, ## parents feel about mj
yepmjmo, ## parents feel about mj monthly
yepaldly, ## parents feel about drinking daily
## Demographics
age2, ## age
catage, ## age category (1 = 12-17 year old)
irsex, ## gender (1 = male)
newrace2, ## race (1 = White, 2-7 non-white)
irfamin3, ## total family income (6 = 50,000 - 74,999)
```

```
poverty2, ## not used in the study but could be for ours
             ## (1 = poverty, 2 = low middle, 3 = middle class or more)
    ## ----- ##
    ## Sampling Variables
   ## ----- ##
   analwt_c, ## sample weight
   vestr,
            ## analysis stratum
   verep
            ## analysis replicate
    ) %>%
  filter(catage == "(1) 12-17 Years Old")
## Data Cleaning
dich = function(x){
 x = ifelse(grepl("(01)|(02)|(08)|(11)", x), 1, 0)
map_to = function(x){
 lbls = sort(levels(x))
 lbls = (sub("^{([0-9]+\)} + (.+$)", "\1", lbls))
 x = as.numeric(gsub("^\(0*([0-9]+)\\).+$", "\\1", x))
}
d1[, c(1:18)] = map_df(d1[, c(1:18)], ~dich(.x))
d1[, c(19:36)] = map_if(d1[, c(19:36)], is.factor, ~map_to(.x))
## Creating final modeling variables
d1 = d1 \%
  mutate(tobacco = ifelse(rowSums(cbind(cigrec, chewrec,
                                      cigarrec)) > 0, 1, 0),
        drink = dr5day,
              = ifelse(rowSums(cbind(analrec, tranrec,
                                      stimrec, sedrec)) > 0, 1, 0),
        mari = ifelse(mjrec == 1, 1, 0),
        illicit = ifelse(rowSums(cbind(cocrec, crakrec,
                                      herrec, hallrec,
                                      lsdrec, pcprec,
                                      ecsrec, inhrec,
                                      methrec)) > 0, 1, 0)) %>%
  mutate(self = rowMeans(cbind(yegpkcig, yegmjevr, yegmjmo, yegaldly)),
        peer = rowMeans(cbind(yefpkcig, yefmjevr, yefmjmo, yefaldly))) %>%
  mutate(dep = washer(ymdeyr, 2, value = 0)) %>%
  mutate(religious = rowMeans(cbind(scale(yerlgsvc),
                                  scale(yerlgimp),
                                  scale(yerldcsn),
                                  scale(yefaiact)))) %>%
  mutate(parent = rowMeans(cbind(yeppkcig, yepmjevr,
```

yepmjmo, yepaldly)))

Models

```
## Sampling Design
library(survey)
design = svydesign(ids = ~1,
                   strata = ~vestr,
                   weights = ~analwt_c,
                   data = d1)
## All a Path Models
## Unadjusted
svy_a1 = svyglm(self ~ religious, design = design)
svy_a2 = svyglm(peer ~ religious, design = design)
svy_a3 = svyglm(dep ~ religious, design = design,
                family = 'quasibinomial')
## Adjusted
svy_a12 = svyglm(self ~ religious + age2 +
                   irsex + newrace2 + irfamin3 + parent,
                 design = design)
svy_a22 = svyglm(peer ~ religious + age2 +
                   irsex + newrace2 + irfamin3 + parent,
                 design = design)
svy_a32 = svyglm(dep ~ religious + age2 +
                   irsex + newrace2 + irfamin3 + parent,
                 design = design,
                 family = 'quasibinomial')
## All b and c' Path Models (drink such low prevalence that it was not included)
svy_bc = svy_bc2 = list()
for (i in c("tobacco", "rx", "mari", "illicit")){
  ## Unadjusted Model
  model = as.formula(pasteO(i, "~ self + peer + dep + religious"))
  svy_bc[[i]] = svyglm(model, design = design, family = "binomial")
  ## Adjusted Model
  model2 = as.formula(paste0(i, "~ self + peer + dep + religious + age2 +
                             irsex + newrace2 + irfamin3 + parent"))
  svy_bc2[[i]] = svyglm(model2, design = design, family = "binomial")
}
library(MarginalMediation)
## Tobacco
fit_tob = mma(svy_bc[["tobacco"]],
              svy_a1,
```

```
svy_a2,
              svy_a3,
              ind_effects = c("religious-self",
                               "religious-peer",
                               "religious-dep"),
              boot = 500)
fit_tob2 = mma(svy_bc2[["tobacco"]],
               svy_a12,
               svy_a22,
               svy_a32,
               ind_effects = c("religious-self",
                                "religious-peer",
                                "religious-dep"),
               boot = 500)
## Rx
fit_rx = mma(svy_bc[["rx"]],
              svy_a1,
              svy_a2,
              svy_a3,
              ind_effects = c("religious-self",
                               "religious-peer",
                               "religious-dep"),
              boot = 500)
fit_rx2 = mma(svy_bc2[["rx"]],
               svy_a12,
               svy_a22,
               svy_a32,
               ind_effects = c("religious-self",
                                "religious-peer",
                                "religious-dep"),
               boot = 500)
## Marijuana
fit_mar = mma(svy_bc[["mari"]],
             svy_a1,
             svy_a2,
             svy_a3,
             ind_effects = c("religious-self",
                              "religious-peer",
                              "religious-dep"),
             boot = 500)
fit_mar2 = mma(svy_bc2[["mari"]],
              svy_a12,
              svy_a22,
              svy_a32,
              ind_effects = c("religious-self",
```

```
"religious-peer",
                               "religious-dep"),
              boot = 500)
## Illicit
fit_ill = mma(svy_bc[["illicit"]],
              svy_a1,
              svy_a2,
              svy_a3,
              ind_effects = c("religious-self",
                              "religious-peer",
                              "religious-dep"),
              boot = 500)
fit_ill2 = mma(svy_bc2[["illicit"]],
               svy_a12,
               svy_a22,
               svy_a32,
               ind_effects = c("religious-self",
                               "religious-peer",
                                "religious-dep"),
               boot = 500)
save(fit_tob, fit_tob2,
     fit_rx, fit_rx2,
     fit_mar, fit_mar2,
     fit_ill, fit_ill2,
     file = here("Data/NSDUH_2014_Results.rda"))
library(MarginalMediation)
library(tidyverse)
library(here)
load(file = here("Data/NSDUH_2014_Results.rda"))
## Extract direct effects
directs_fx = function(..., type){
  list(...) %>%
    map(~.x$dir_effects) %>%
    do.call("rbind", .) %>%
    data.frame(.) %>%
    select(Direct, Lower, Upper) %>%
    data.frame(., row.names =
                 gsub("religious", "Religiousity (Direct)", row.names(.))) %>%
    rownames_to_column() %>%
    mutate(Outcome = c(rep("Tobacco", 1), rep("Prescription", 1),
                       rep("Marijuana", 1), rep("Illicit", 1))) %>%
    select(Outcome, rowname, Direct, Lower, Upper) %>%
```

```
set_names(c("Outcome", "Path", "Estimate", "Lower", "Upper")) %>%
    mutate(CI = paste0("(", round(Lower,4), ", ", round(Upper,4), ")")) %>%
    select(-CI) %>%
    mutate(type = type)
}
directs_un = directs_fx(fit_tob,
                        fit_rx,
                        fit_mar,
                        fit_ill,
                        type = "Unadjusted")
directs_adj = directs_fx(fit_tob2,
                         fit_rx2,
                         fit_mar2,
                         fit_ill2,
                         type = "Adjusted")
## Extract indirect effects and bind to directs
inds_fx = function(..., type){
  list(...) %>%
    map(~.x$ind_effects) %>%
    do.call("rbind", .) %>%
    data.frame(.) %>%
    select(Indirect, Lower, Upper) %>%
    data.frame(., row.names = gsub("religious-", "Religiousity Through",
                                   row.names(.))) %>%
    data.frame(., row.names = gsub("dep", "\nDepression",
                                   row.names(.))) %>%
    data.frame(., row.names = gsub("self", "\nRespondent Views",
                                   row.names(.))) %>%
    data.frame(., row.names = gsub("peer", "\nPeer Views",
                                   row.names(.))) %>%
    rownames_to_column() %>%
    mutate(Outcome = c(rep("Tobacco", 3),
                       rep("Prescription", 3),
                       rep("Marijuana", 3),
                       rep("Illicit", 3))) %>%
    select(Outcome, rowname, Indirect, Lower, Upper) %>%
    set_names(c("Outcome", "Path", "Estimate",
                "Lower", "Upper")) %>%
    mutate(CI = paste0("(", round(Lower,4), ", ",
                       round(Upper,4), ")")) %>%
    select(-CI) %>%
    mutate(type = type)
unadjusted = inds_fx(fit_tob,
                     fit_rx,
                     fit_mar,
```

```
fit ill,
                     type = "Unadjusted") %>%
  rbind(directs_un)
adjusted = inds_fx(fit_tob2,
                   fit_rx2,
                   fit_mar2,
                   fit_ill2,
                   type = "Adjusted") %>%
  rbind(directs_adj)
inds = rbind(unadjusted, adjusted) %>%
  data.frame %>%
  mutate(type = factor(type,
                       levels = c("Unadjusted", "Adjusted"))) %>%
  mutate(Path = gsub("[0-9]","", Path)) %>%
  mutate(Outcome = factor(Outcome,
                          levels = c("Tobacco", "Prescription",
                                      "Marijuana", "Illicit")))
## Odds ratios and linear effects as done in Ford and Hill
## Sampling Design
library(survey)
design = svydesign(ids = ~1,
                   strata = ~vestr,
                   weights = ~analwt_c,
                   data = d1)
## All a Path Models
svy_a1 = svyglm(self ~ religious + age2 +
                  irsex + newrace2 + irfamin3 + parent,
                design = design)
svy_a2 = svyglm(peer ~ religious + age2 +
                  irsex + newrace2 + irfamin3 + parent,
                design = design)
svy_a3 = svyglm(dep ~ religious + age2 +
                  irsex + newrace2 + irfamin3 + parent,
                design = design,
                family = 'quasibinomial')
## path a
patha_fx = function(obj, row){
  cbind(coef(obj)[row],
            confint(obj)[row,1],
            confint(obj)[row,2])
}
est1 =
  rbind(
    patha_fx(svy_a1, "religious"),
```

```
patha_fx(svy_a2, "religious"),
    patha_fx(svy_a3, "religious")
)
rownames(est1) = c("Respondent", "Peer", "Depression")
est1 = data.frame(est1) %>%
  set_names(c("Estimate", "Lower", "Upper"))
## All c and c' Path Models
svy_c = svy_c1 = list()
for (i in c("tobacco", "rx", "mari", "illicit")){
  model = as.formula(paste0(i, "~ religious + age2 +
                            irsex + newrace2 + irfamin3 + parent"))
  svy_c[[i]] = svyglm(model, design = design, family = "quasibinomial")
  model2 = as.formula(paste0(i, "~ self + peer + dep + religious + age2 +
                             irsex + newrace2 + irfamin3 + parent"))
  svy_c1[[i]] = svyglm(model2, design = design, family = "quasibinomial")
}
## Odds ratios of c and c' path models
pathc_fx = function(obj, row, drug){
  cbind(coef(obj[[drug]])[row],
            confint(obj[[drug]])[row,1],
            confint(obj[[drug]])[row,2])
}
est2 =
  rbind(
    cbind(pathc_fx(svy_c, "religious", "tobacco"),
          pathc_fx(svy_c1, "religious", "tobacco")),
    cbind(pathc_fx(svy_c, "religious", "rx"),
          pathc_fx(svy_c1, "religious", "rx")),
    cbind(pathc_fx(svy_c, "religious", "mari"),
          pathc_fx(svy_c1, "religious", "mari")),
    cbind(pathc_fx(svy_c, "religious", "illicit"),
          pathc_fx(svy_c1, "religious", "illicit"))
rownames(est2) = c("Tobacco", "Rx",
                   "Marijuana", "Illicit")
est2 = data.frame(est2) %>%
  set_names(c("c", "c_Lower", "c_Upper", "c1", "c1_Lower", "c1_Upper"))
```

Tables and Figures

```
Table 6.1 on page 62
```

```
## overall table1 (not adjusted for survey weights)
d1 %>%
  table1(factor(rx), factor(tobacco),
         factor(drink), factor(mari), factor(illicit),
         irsex, factor(ifelse(newrace2 == "(1) NonHisp White", 0, 1)),
         factor(
           ifelse(poverty2 == "(3) Income > 2X Fed Pov Thresh (See comment above)",
                  1, 0)),
         factor(dep), self, peer, parent, religious,
         type = c("simple", "condense"),
         var_names = c("Prescription", "Tobacco",
                       "Heavy Drinking", "Marijuana",
                       "Other Illicit", "Sex",
                       "Race (Non-White)",
                       "Income (2x poverty)",
                       "Major Depression Episode",
                       "Respondent", "Peer",
                       "Parent", "Religiosity"),
         output = "latex2")
## Survey weighted
library(survey)
design = svydesign(ids = ~1,
                   strata = ~vestr,
                   weights = ~analwt_c,
                   data = d1)
svymean(~rx, design)
svymean(~tobacco, design)
svymean(~drink, design)
svymean(~mari, design)
svymean(~illicit, design)
svymean(~irsex, design)
svymean(~newrace2, design)
svymean(~poverty2, design)
svymean(~dep, design, na.rm=TRUE)
svymean(~self, design, na.rm=TRUE)
svymean(~peer, design, na.rm=TRUE)
svymean(~parent, design, na.rm=TRUE)
svymean(~religious, design, na.rm=TRUE)
## alpha of religiosity
with(d1,
  psych::alpha(cbind(scale(yerlgsvc),
                     scale(yerlgimp),
```

```
scale(yerldcsn),
                     scale(yefaiact))))
## alpha of respondent
with(d1,
  psych::alpha(cbind(yegpkcig, yegmjevr,
                     yegmjmo, yegaldly)))
## alpha of peer
with(d1,
  psych::alpha(cbind(yefpkcig, yefmjevr,
                     yefmjmo, yefaldly)))
## alpha of parent
with(d1,
 psych::alpha(cbind(yeppkcig, yepmjevr,
                     yepmjmo, yepaldly)))
## number of heavy drinking responses
sum(d1$drink)
```

Table 6.2 on page 64

```
perc_fx = function(obj){
  obj$ind_effects[,3]/(obj$dir_effects[,1] +
                         sum(obj$ind_effects[,3]))
}
percent_ind = cbind(perc_fx(fit_tob2),
                    perc_fx(fit_rx2),
                    perc_fx(fit_mar2),
                    perc_fx(fit_ill2)) %>%
  data.frame %>%
  set_names(c("Tobacco", "Prescription", "Marijuana", "Illicit")) %>%
  map_df(~.x*100) %>%
  mutate(Mediator = c("Respondent Views", "Peer Views", "Depression")) %>%
  select(Mediator, Tobacco, Prescription, Marijuana, Illicit)
percent_ind = percent_ind %>%
  rbind(data.frame(
    Mediator = "Total",
    Tobacco = sum(percent_ind$Tobacco),
    Prescription = sum(percent_ind$Prescription),
    Marijuana = sum(percent_ind$Marijuana),
    Illicit = sum(percent_ind$Illicit)
  ))
library(xtable)
xtable(percent_ind, digits = 1)
```

Table ?? on page ??

```
est21 = est2 %>%
  rownames_to_column() %>%
  group_by(rowname) %>%
  summarize(perc = ((c - c1)/c)*100)
library(xtable)
xtable(est21, digits = 1) %>%
  print.xtable(include.rownames = FALSE)
```

Figure ?? on page ??

```
p = position_dodge(width = .2)
inds %>%
  filter(type == "Adjusted") %>%
  ggplot(aes(Path, Estimate, group = type, color = type)) +
    geom_hline(yintercept = 0, color = "darkgrey") +
    geom_point(position = p, alpha = .8) +
    geom_errorbar(aes(ymin = Lower, ymax = Upper),
                  position = p, alpha = .8,
                  width = .3) +
    facet_wrap(~Outcome) +
    coord_flip() +
    anteo::theme_anteo_wh() +
    theme(legend.position = "bottom",
          axis.line = element_line(color = "darkgrey"),
          panel.spacing = unit(.3, "in")) +
    scale_color_manual(values = c("chartreuse4", "coral2"),
                       guide = FALSE) +
    labs(x = "", y = "",
         color = "")
```

CURRICULUM VITA

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EDUCATION

Ph.D. Quantitative Psychology | Expected: Feb 2018

Utah State University

Marginal Mediation Analysis: A New Framework for Interpretable Mediated Effects

B.S. Economics May 2014

Utah State University

 $Cum\ Laude$

B.S. Psychology | May 2014

Utah State University

Cum Laude

RESEARCH AREAS

Primary: Model Interpretability, Mediation Analysis, R Programming

SECONDARY: Health Behavior and Chronic Conditions, Machine Learning

RESEARCH EXPERIENCE

Research Assistant, Prevention Science Lab, Utah State University 2016-Present

Advisor: Ginger Lockhart, PhD

Statistical and Data Science Consultant, Utah State University 2016-Present

Advisors: Sarah Schwartz, PhD and Jamison Fargo, PhD

WOC Research Assistant, Veteran Affairs SLC

2016-Present

Advisors: Vanessa Stevens, PhD and Richard Nelson, PhD

Research Assistant, NCHAM, Utah State University

2016-Present

Advisor: Karl White, PhD

PUBLISHED WORKS IN REFEREED JOURNALS

- * Denotes that I was the project methodologist or data scientist
 - **Barrett, T.S.** & Brignone, E. (2017). Furniture for Quantitative Scientists. R Journal, 9(2), 142-146.
 - *Borrie, S.A., Lansford, K.L., & **Barrett, T.S.** (2017). Generalized Adaptation to Dysarthric Speech. *Journal of Speech, Language, and Hearing Research*, 60, 3110-3117. doi: 10.1044/2017_JSLHR-S-17-0127.
 - Barrett, T.S. & White, K.R. (2017). Trends in Hearing Loss Among Adolescents. Pediatrics, 140(6), e20170619. doi: 10.1542/peds.2017-0619.
 - Fargo, J.D., Brignone, E., Metraux, S., Peterson, R., Carter, M.E., Barrett, T., Palmer, M., Redd, A., Samore, M.H., & Gundlapalli, A.V. (2017). Homelessness Following Disability-Related Discharges from Active Duty Military Service in Afghanistan and Iraq. *Disability and Health Journal*, 10, 592-599. doi: 10.1016/j.dhjo.2017.03.003.
 - *Brown, L.T., Mohr, K.A.J., Wilcox, B.R., & **Barrett, T.S.** (2017). The effects of dyad reading and text difficulty on third-graders' reading achievement. *Journal of Educational Research*. doi: 10.1080/00220671.2017.1310711
 - Brignone, E., Gundlapalli, A.V., Blais, R.K., Kimerling, R., **Barrett, T.S.**, Nelson, R.E., Carter, M.E., Samore, M.A., & Fargo, J.D. (2017). Increased Health Care Utilization and Costs among Veterans with a Positive Screen for Military Sexual Trauma. *Medical Care*, 55, S70-S77. doi: 10.1097/MLR.00000000000000767
 - Long, E., **Barrett, T.S.**, & Lockhart, G. (2017). Network-behavior dynamics of adolescent friendships, alcohol use, and physical activity. *Health Psychology*, 36(6), 577-586. doi: 10.1037/hea0000483
 - *Borrie, S.A., Lansford, K.L., & **Barrett, T.S.** (2017). Rhythm perception and its role in perception and learning of dysrhythmic speech. *Journal of Speech, Language, and Hearing Research, 60*(3), 1-10. doi: 10.1044/2016_JSLHR-S-16-0094
 - Behl, D., White, K., **Barrett, T.**, Blaiser, K., Callow-Heusser, C., & Croshaw, T. (2017). A Multisite Study Evaluating the Benefits of Early Intervention via Telepractice. *Infant and Young Children*, 30(2), 147-161.
 - Barrett, T.S. & White, K. R. (2016). Prevalence and Trends of Childhood Hearing Loss Based on Federally-funded National Surveys: 1994–2013. Journal of Early Hearing Detection and Intervention, 1(2), 8-16.
 - Doutré, S. M., **Barrett, T.S.**, Greenlee, J. & White, K. R. (2016). Losing Ground: Awareness of Congenital Cytomegalovirus in the United States. *Journal of Early Hearing Detection and Intervention*, 1(2), 39-48.

Munoz, K., Rusk, S.E.P., Nelson, L., Preston, E., White, K.R., **Barrett, T.S.** & Twohig, M.P. (2016). Pediatric Hearing Aid Management: Parent Reported Needs for Learning Support. *Ear and Hearing Journal.* 37(6), 703-709.

Borgogna, N., Lockhart, G., Grenard, J, **Barrett, T.**, Shiffman, S. & Reynolds, K. (2015). Ecological Momentary Assessment of Urban Adolescents' Technology Use and Cravings for Unhealthy Snacks and Drinks: Differences by Ethnicity and Sex. *Journal of the Academy of Nutrition and Dietetics*, 115(5), 759-766.

Under Review/In Revision

- Barrett, T.S. & Lockhart, G. Efficient exploration of many variables and interactions using regularized regression. *Prevention Science*. In Revision.
- *Leopold, S., Healy, E. W., Youngdahl, C., **Barrett, T.S.**, & Apoux, F. Speech-material and talker effects in speech band importance. *Journal of the Acoustical Society of America*. In Revision.

OTHER PUBLICATIONS

R for the Health, Behavioral, and Social Sciences — A primer on using R for researchers in health, behavioral, and social sciences including importing data, data manipulation, data modeling, and data visualization. Currently published online at tysonstanley.github.io/Rstats/.

The furniture R package (Published on CRAN and GitHub)

— Contains functions to help with several aspects of research in the health, behavioral, and social sciences. The main functions—table1() and tableC()—produce descriptive statistics and correlations in well-formatted tables (as commonly seen as the "Table 1" of academic journals). Over 8,000 downloads (see tysonbarrett.com/furniture).

The MarginalMediation R package (Published on GitHub)

- Contains functions to perform Marginal Mediation. Papers discussing the method and software are forthcoming (see tysonbarrett.com/MarginalMediation).
- The anteo R package (Published on GitHub) Contains functions to help in interpreting machine learning models. Still under active development.

SELECTED PRESENTATIONS

Barrett, T.S., & Lockhart, G. (2017). Enhancing the Exploration and

- Communication of Big Data in Prevention Science. Poster presented at the Annual Meeting of the Society of Prevention Research, Washington, DC. Received "Distinguished Poster Award" and "Abstract of Distinction."
- Barrett, T.S., & Lockhart, G. (2017). Exploring the Predictors of Marijuana Use Among Adolescents with Asthma. Oral presentation at the Utah State University Research Symposium, Logan, UT.
- Sanghavi, K., White, K., **Barrett, T.S.**, Wylie, A., Raspa, M., Cashman, D., Vogel, B. Caggana, M. & Bodurtha, J. (2017). Poster presented at the Early Hearing Detection and Intervention Conference, Atlanta, GA. Received "Outstanding Poster Award."
- Brignone, E., Gundlapalli, A.V., **Barrett, T.S.**, Blais, R.K., Nelson, R.E., Carter, M.E., Kimerling, R., Samore, M.H., Fargo, J.D. (2016). Cost of Care among Male and Female Veterans with a Positive Screen for Military Sexual Trauma. Poster presented at the 2016 Annual Meeting of the International Conference of Psychology, Yokohama, Japan.
- Barrett, T.S., Munoz, K. & White, K. (2016). How well do parent report hearing loss in their children? Poster presented at the Early Hearing Detection and Intervention Conference, San Diego, CA.
- Barrett, T.S., Munoz, K. & White, K. (2016). Accounting for Temporary Loss in National Studies on Hearing Loss. Poster presented at the Early Hearing Detection and Intervention Conference, San Diego, CA.
- Barrett, T.S., Munoz, K. & White, K. (2016). An Evaluation of Early Intervention delivered via Video Conferencing. Poster presented at the Early Hearing Detection and Intervention Conference, San Diego, CA.
- Stevens, V., **Barrett, T.S.** & Nelson, R. (2016). Distribution and Daily Cost of Care in a Pediatric Hospital. Oral presentation to the Pediatric Guidance Council of Intermountain Healthcare, Salt Lake City, UT.
- Barrett, T.S., Munoz, K. & White, K. (2015). Refinements to estimating prevalence of hearing loss in children. Poster presented at the Utah State University Research Symposium, Logan, UT.
- Barrett, T.S., Prante, M., Peterson, R., Fargo, J.D., Pyle, N. (2014). Predictors of employability among homeless youth. Poster presented at the Psi-Chi Undergraduate Research Conference at Idaho State University, Pocatello, ID. Best Undergraduate Poster Presentation Award.
- Barrett, T.S., Holland, D. (2014). Nascent Entrepreneurship, Impulsivity,

and Self- Efficacy. Poster presented at the Research on Capitol Hill, Salt Lake City, UT.

Holland, D., **Barrett, T.S.** (2013). Impulsivity in young entrepreneurs. Round table discussion at the Babson Business Conference, Paris, France.

AWARDS

- Abstract of Distinction (Annual Meeting of the Society of Prevention Research)
- Distinguished Poster Award (Annual Meeting of the Society of Prevention Research)
- Outstanding Poster Award (Annual Meeting of EHDI)
- Dean's Scholarship (full tuition for two years)
- Best Poster Presentation (Psi-Chi Undergraduate Conference)

TEACHING INTERESTS

PRIMARY INTEREST: QUANTITATIVE METHODS

- Undergraduate and Graduate Statistics (ANOVA and Regression [OLS, GLM])
- Multilevel Modeling (Hierarchical Linear Modeling, GEE, Mixed Effects)
- Mediation Analysis (Marginal Mediation, Moderated Mediation)
- Reproducible Research (Research Methods, Open Science Framework)
- R for the Social Sciences (Undergraduate and Graduate Level)
- Exploratory Data Analysis
- Structural Equation Modeling (Psychometrics and Measurement Models, Mixture Modeling)
- Research Methods (Undergraduate and Graduate Level)

SECONDARY INTEREST: PUBLIC HEALTH

- Research in Public Health
- Disabilities (Hearing Loss, Developmental Disabilities)

TEACHING EXPERIENCE

Instructor

- R for the Health, Behavioral, Educational, and Social Sciences I and II
 - 2016 Present
 - Created, Developed, and Taught
 - Graduate Level
 - Most Recent Student-Responded Ratings:
 - * Overall: 4.5 out of 5.0
 - * Teaching: 4.9 out of 5.0
 - * Course: 4.8 out of 5.0

Teaching Assistant

- Econometrics I (Graduate Level)
 - Fall 2016
- Psychological Statistics (Undergraduate Level)
 - -2014-2015

METHODOLOGICAL TRAINING

- Regression, Generalized Linear Models
- Mixed Effects, Generalized Linear Models
- Machine Learning
 - CART
 - Random Forest
 - Regularized Regression
 - Boosting/Bagging
 - Cross-Validation
- Social Network Analysis
- Data Visualization
 - Static and Dynamic Visuals

SOFTWARE AND PROGRAMMING EXPERIENCE

- 1. R and RMarkdown: Expert (4 years of daily use)
 - Database Queries, Creation, and Management
 - Data Analytics
 - Website Creation
 - Program Development and Deployment
- 2. REDCap: Moderate Experience (1 year of weekly use)
 - Database Creation

- Data Entry and Survey Creation and Deployment
- Data Management via REDCap API
- 3. SQL: Some Experience (4 years of occasional use)
 - Database Queries and Database Management Python: Minor Experience (1 year of occasional use)
 - Data Analytics