

## Poisson Process

**Definition:** A Counting Process  $N(t)$  is a Poisson Process  $N(t)$  if

- # of arrivals in  $(t_0, t_1]$ ,  $N(t_1) - N(t_0)$ , is a Poisson RV with expected value  $\lambda(t_1 - t_0)$
- # of arrivals in each non-overlapping interval are independent random variable

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## Poisson Process

- Process rate  $(\lambda) = E[N(t)] / t$
- $M = N(t_1) - N(t_0) = \text{Poisson RV}$

$$P_M(m) = \begin{cases} \frac{[\lambda(t_1 - t_0)]^m}{m!} e^{-\lambda(t_1 - t_0)} & m = 0, 1, 2, \dots \\ 0 & \text{Otherwise} \end{cases}$$

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## Joint PMF

**Theorem:** Poisson Process  $N(t)$  of rate  $\lambda$ ,  
Joint PMF of  $N(t_1), \dots, N(t_k)$ ,  $t_1 < \dots < t_k$

$$P_{N(t_1)N(t_2)N(t_3)\dots N(t_k)}(n_1, n_2, n_3, \dots, n_k) \\ = \begin{cases} \frac{\alpha_1^{n_1} e^{-\alpha_1}}{n_1!} \frac{\alpha_2^{(n_2 - n_1)} e^{-\alpha_2}}{(n_2 - n_1)!} \dots \frac{\alpha_k^{(n_k - n_{k-1})} e^{-\alpha_k}}{(n_k - n_{k-1})!}, & 0 \leq n_1 \leq n_2 \leq \dots \leq n_k \\ 0, & \text{otherwise} \end{cases} \\ \alpha_i = \lambda(t_i - t_{i-1})$$

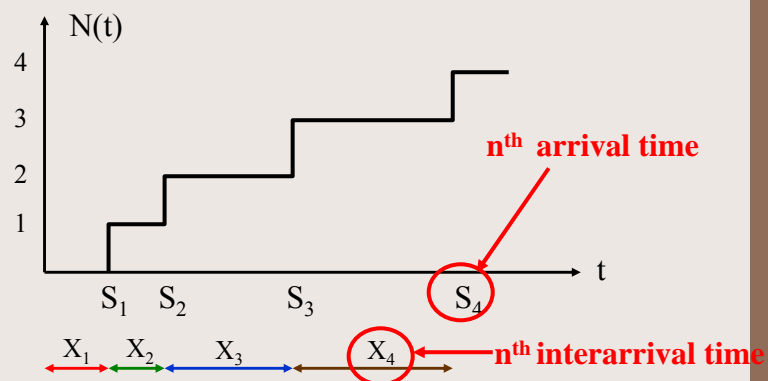
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## Joint PMF

$$P_{N(t_1)N(t_2)N(t_3)\dots N(t_k)}(n_1, n_2, n_3, \dots, n_k) \\ = P_{N(t_1)}(n_1) P_{N(t_2)|N(t_1)}(n_2 | n_1) P_{N(t_3)|N(t_2)N(t_1)}(n_3 | n_1, n_2) \dots \\ P_{N(t_k)|N(t_1)N(t_2)\dots N(t_{k-1})}(n_k | n_1, n_2, \dots, n_{k-1}) \\ = P_{N(t_1)}(n_1) P_{N(t_2)|N(t_1)}(n_2 | n_1) P_{N(t_3)|N(t_2)}(n_3 | n_2) \dots \\ P_{N(t_k)|N(t_{k-1})}(n_k | n_{k-1}) \\ = \begin{cases} \frac{\alpha_1^{n_1} e^{-\alpha_1}}{n_1!} \frac{\alpha_2^{(n_2 - n_1)} e^{-\alpha_2}}{(n_2 - n_1)!} \dots \frac{\alpha_k^{(n_k - n_{k-1})} e^{-\alpha_k}}{(n_k - n_{k-1})!}, & 0 \leq n_1 \leq n_2 \leq \dots \leq n_k \\ 0, & \text{otherwise} \end{cases} \\ \alpha_i = \lambda(t_i - t_{i-1})$$

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## Counting Process



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## Example

- A mobile station transmits data packet as Poisson process with rate 12 packets/sec
- Find PMF # of packets transmitted in the  $k^{\text{th}}$  hour
  - Find Joint PMF of # of packets transmitted in the  $k^{\text{th}}$  hour and  $z^{\text{th}}$  hour

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## Example

- Let  $N_k$  = # of packets transmitted in  $k^{\text{th}}$  hour
- # packets in each hour is IID

$$P_{N_k}(n) = \begin{cases} \frac{[12(3600-0)]^n}{n!} e^{-12(3600-0)} & n = 0, 1, 2, \dots \\ 0 & \text{Otherwise} \end{cases}$$

$$= \begin{cases} \frac{[43200]^n}{n!} e^{-43200} & n = 0, 1, 2, \dots \\ 0 & \text{Otherwise} \end{cases}$$

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## Example

- Joint PMF of # of packets transmitted in the  $k^{\text{th}}$  hour and  $z^{\text{th}}$  hour

$$P_{N_k, N_z}(n_k, n_z) = \begin{cases} \frac{\alpha_k^{n_k} e^{-\alpha_k}}{n_k!} \frac{\alpha_z^{n_z} e^{-\alpha_z}}{n_z!} & n_k = 0, 1, \dots \\ & n_z = 0, 1, \dots \\ 0 & \text{Otherwise} \end{cases}$$

$$= \begin{cases} \frac{\alpha^{(n_k+n_z)}}{n_k! n_z!} e^{-2\alpha} & n_k = 0, 1, \dots \\ & n_z = 0, 1, \dots \\ 0 & \text{Otherwise} \end{cases}$$

$\alpha = \alpha_k = \alpha_z = \lambda T = 12(3600-0) = 43200$

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## Interarrival Time

**Theorem:** Poisson Process of rate  $\lambda$ , the interarrival times  $X_1, X_2, \dots$  are an iid random sequence with **Exponential PDF**

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x} & x \geq 0 \\ 0 & \text{Otherwise} \end{cases}$$

## Interarrival Time

**Theorem:** A Counting Process with *independent exponential interarrival time*  $X_1, X_2, \dots$  with  $E[X_i] = 1/\lambda$  is a **Poisson Process** of rate  $\lambda$

## Brownian Motion Process

- A continuous time, continuous value random process

**Definition:** A Brownian Motion Process  $X(t)$  has the property that

- $X(0) = 0$
- For  $\tau > 0$ ,  $X(t + \tau) - X(t) = \text{Gaussian RV}$  with  $E[X(t + \tau) - X(t)] = 0$  and  $\text{Var}[X(t + \tau) - X(t)] = \alpha\tau$  that is independent of  $X(t')$  for  $t' \leq t$

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## Brownian Motion Process

Gaussian RV

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{\frac{-(x - \mu)^2}{2\sigma^2}}$$

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## Joint PDF

- **Theorem:** For the Brownian motion process  $X(t)$ , joint PDF of  $X(t_1), \dots, X(t_k)$

$$f_{X(t_1), \dots, X(t_k)}(x_1, \dots, x_k) = \prod_{n=1}^k \frac{1}{\sqrt{2\pi\alpha(t_n - t_{n-1})}} e^{\frac{-(x_n - x_{n-1})^2}{2\alpha(t_n - t_{n-1})}}$$

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## Expected Value

$$\mathbf{X(t): } X(t_1) \rightarrow f_{X(t_1)}(x) \rightarrow E[X(t_1)]$$

**Definition:** The expected value of a stochastic process  $X(t)$  is the deterministic function

$$\mu_x(t) = E[X(t)]$$

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## Covariance of X and Y

**Definition:**  $\text{Cov}[X, Y] = E[(X - \mu_x)(Y - \mu_y)]$

$$\begin{aligned}\text{Cov}[X, Y] &= E[XY - \mu_x Y - \mu_y X + \mu_x \mu_y] \\ &= E[XY] - \mu_x E[Y] - \mu_y E[X] + \mu_x \mu_y \\ &= E[XY] - \mu_x \mu_y - \mu_x \mu_y + \mu_x \mu_y\end{aligned}$$

**Theorem:**  $\text{Cov}[X, Y] = E[XY] - \mu_x \mu_y$

- High covariance = an observation of X provides an accurate indication of Y

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## Autocovariance

**Definition:** The autocovariance function of a stochastic process  $X(t)$  is

$$C_X(t, \tau) = \text{Cov}[X(t), X(t + \tau)]$$

Note

- For  $\tau = 0 \rightarrow C_X(t, 0) = \text{Var}[X(t)]$

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## Autocovariance

- Auto = self = same process
- For a same process  $X(t)$ , in 2 different times  $t_1 = t$  and  $t_2 = t + \tau$
- For high covariance = a sample function, is unlikely to change in  $\tau$  interval
- For near zero covariance = rapid change

How much the sample function is likely to change in the  $\tau$  interval after  $t$

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## Correlation

**Definition:** The correlation of  $X$  and  $Y$  is  $r_{X,Y}$   
$$r_{X,Y} = E[XY]$$

**Theorem:** 
$$\text{Cov}[X,Y] = r_{X,Y} - \mu_x \mu_y$$

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## Autocorrelation

- **Definition:** The autocorrelation function of a stochastic process  $X(t)$  is

$$R_X(t, \tau) = E [ X(t) X(t+\tau) ]$$

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## Autocovariance & Autocorrelation

**Theorem:** The autocovariance and autocorrelation functions of a process  $X(t)$  satisfy

$$C_X(t, \tau) = R_X(t, \tau) - \mu_X(t) \mu_X(t + \tau)$$

Note:

Autocovariance → use  $X(t)$  to predict a future  $X(t+\tau)$

Autocorrelation → describe the power of a random signal

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## Random Sequence

For a discrete time process, the sample function is described by the ordered sequence of random variable  $X_n = X(nT)$

- **Definition:** A random sequence  $X_n$  is an ordered sequence of random variable  $X_0, X_1, \dots$

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## Autocovariance & Autocorrelation of a Random Sequence

**Definition:** The autocovariance of a random sequence  $X_n$  is

$$C_X[m, k] = \text{Cov}[X_m, X_{m+k}]$$

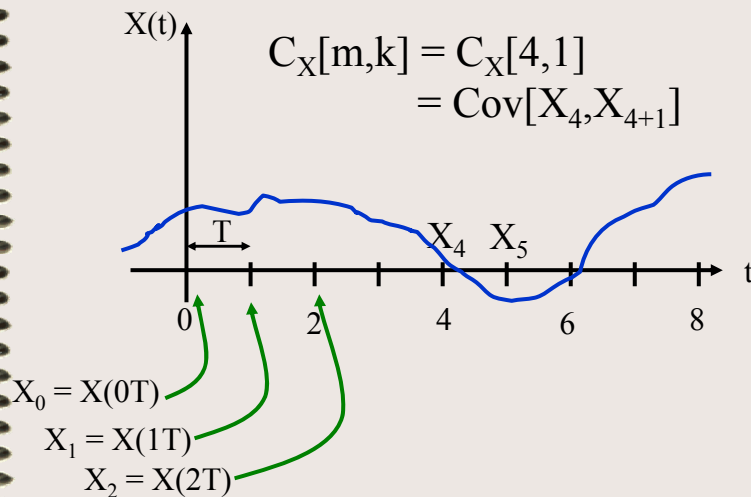
**Definition:** The autocorrelation of a random sequence  $X_n$  is

$$R_X[m, k] = E[X_m X_{m+k}]$$

$m$  and  $k$  are integers

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## Autocovariance & Autocorrelation of a Random Sequence



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## Stationary Process

- For a general random process  $X(t)$ , normally,  
at  $t_1$ :  $X(t_1)$  has pdf =  $f_{X(t_1)}(x)$  [depends on  $t_1$ ]
- For a stationary random process  $X(t)$ ,  
at  $t_1$ :  $X(t_1)$  has pdf =  $f_{X(t_1)}(x)$  [not depend on  $t_1$ ]

### Stationary Process

= same random variable at all time

= no statistical properties change with time

$$f_{X(t_1)}(x) = f_{X(t_1 + \tau)}(x) = f_X(x)$$

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## Stationary Process

**Definition:** A stochastic process  $X(t)$  is stationary iff for all sets of time  $t_1, \dots, t_m$  and any time different  $\tau$ ,

$$f_{X(t_1), \dots, X(t_m)}(x_1, \dots, x_m) = f_{X(t_1 + \tau), \dots, X(t_m + \tau)}(x_1, \dots, x_m)$$

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## Stationary Random Sequence

**Definition:** A random sequence  $X_n$  is stationary iff for any finite sets of time instants  $n_1, \dots, n_m$  and any time different  $k$ ,

$$f_{X_{n_1}, \dots, X_{n_m}}(x_1, \dots, x_m) = f_{X_{n_1 + k}, \dots, X_{n_m + k}}(x_1, \dots, x_m)$$

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## Stationary Process

**Theorem:** A stationary process  $X(t)$ ,

$$\begin{aligned}\mu_X(t) &= \mu_X \\ R_X(t, \tau) &= R_X(0, \tau) = R_X(\tau) \\ C_X(t, \tau) &= R_X(0, \tau) - \mu_X^2 = C_X(\tau)\end{aligned}$$

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## Stationary Random Sequence

**Theorem:** A stationary random sequence  $X_n$ , for all  $m$

$$\begin{aligned}E[X_m] &= \mu_X \\ R_X[m, k] &= R_X[0, k] = R_X[k] \\ C_X[m, k] &= R_X[0, k] - \mu_X^2 = C_X[k]\end{aligned}$$

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## Wide Sense Stationary

**Definition:**  $X(t)$  is a wide sense stationary random process iff for all  $t$ ,

$$E[X(t)] = \mu_X$$

$$R_X(t, \tau) = R_X(0, \tau) = R_X(\tau)$$

**Definition:**  $X_n$  is a wide sense stationary random sequence iff for all  $n$ ,

$$E[X_n] = \mu_X$$

$$R_X[n, k] = R_X[0, k] = R_X[k]$$

## Wide Sense Stationary

- For every **stationary** process or sequence, it is also **wide sense stationary**.
- However, if it is a **wide sense stationary** it may or may not be **stationary**.

## Example

- For  $n = \text{even}$   
 $X_n = \pm 1$  with prob =  $\frac{1}{2}$  ( $n = \text{even}$ )
- For  $n = \text{odd}$   
 $X_n = -1/3$  with prob =  $9/10$   
 $X_n = 3$  with prob =  $1/10$
- Stationary ?
  - No
- Wide sense stationary ?
  - Mean = 0 for all  $n$
  - $C_X(t, \tau) = 0$  for  $\tau > 0$
  - $C_X(t, \tau) = 1$  for  $\tau = 0$
  - Yes , it's wide sense stationary

## Wide Sense Stationary

**Theorem:** For a wide sense stationary process  $X(t)$ ,

$$\begin{aligned}R_X(0) &\geq 0 \\R_X(\tau) &= R_X(-\tau) \\|R_X(\tau)| &\leq R_X(0)\end{aligned}$$



## Wide Sense Stationary

**Theorem:** For a wide sense stationary sequence  $X_n$ ,

$$R_X[0] \geq 0$$

$$R_X[k] = R_X[-\tau]$$

$$|R_X[k]| \leq R_X[0]$$

## Average Power

- From Ohm's Law :  $V = IR$
- For  $v(t)$ ,  $i(t)$ ,  $R \Omega$ , the instantaneous power dissipated,  $P(t)$ ,

$$P(t) = v^2(t)/R = i^2(t)R$$

- For  $R = 1 \Omega$ ,  $P(t) = v^2(t) = i^2(t)$
- For a voltage or current is a sample function of random process,  $x(t,s)$

$$\rightarrow P \text{ across } 1 \Omega \text{ resistor} = x^2(t,s)$$

## Average Power

- Define  $x^2(t,s)$  as the instantaneous power of  $x(t,s)$
- For a  $X(t)$ ,  $X^2(t)$  is the instantaneous of power  $X(t)$

### Definition:

For a wide sense stationary process  $X(t)$ ,

$$R_X(0) = E[X^2(t)]$$

## Homework

- 6.2.2
- 6.3.3
- 6.4.1
- 6.5.3
- 6.7.2
- 6.8.2