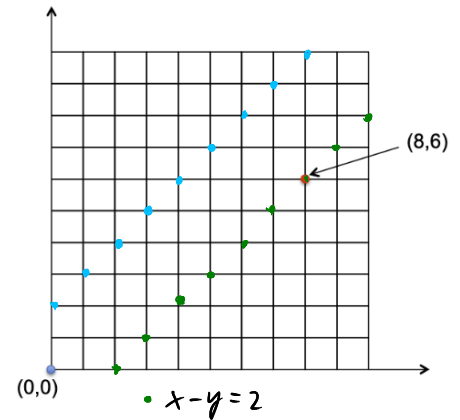


(1) Random walk of a robot

A robot is placed at the origin (the point $(0,0)$) on a two-dimension integer grid (see the figure below). Denote the position of the robot by (x,y) . The robot can either move right to $(x+1,y)$ or move up to $(x,y+1)$.



- (a) Suppose each time the robot randomly moves right or up with equal chance. What is the probability that the robot will ever reach the point $(8,6)$?
- (b) Suppose another robot has a $\frac{2}{3}$ chance to move right and a $\frac{1}{3}$ chance to move up when $x+y$ is even, otherwise it has a $\frac{1}{4}$ chance to move right and a $\frac{3}{4}$ chance to move up. It stops whenever $|x-y| \geq 2$. Find the probability that $x-y=2$ when it stops.

a) The sum $x+y$ is exactly the number of total moves made so far.

In order to reach the point $(8,6)$, it requires 14 moves in total, of which exactly 8 are to the right and 6 are up. The order in which these moves are made does not matter.

For all $i \in \mathbb{N}_0$ let X_i be the x -coordinate after i moves and Y_i the y -coordinate after i moves. Clearly, $X_i + Y_i = i$ for all $i \in \mathbb{N}_0$.

$$\mathbb{P}\left(\bigcup_{i=0}^{\infty} [(X_i, Y_i) = (8, 6)]\right) = \mathbb{P}([(X_{14}, Y_{14}) = (8, 6)]) = \mathbb{P}(X_{14} = 8) = \binom{14}{8} \left(\frac{1}{2}\right)^8 \left(\frac{1}{2}\right)^{14-8} = \binom{14}{8} \left(\frac{1}{2}\right)^{14} \approx 0.183$$

$$\begin{aligned} \text{b) } \mathbb{P}\left(\bigcup_{i=1}^{\infty} [X_i - Y_i = 2]\right) &= (pq) + (p(1-q)pq + (1-p)qpq) + \\ &\quad (p(1-q)p(1-q)pq + p(1-q)(1-p)qpq + (1-p)qp(1-q)pq + (1-p)q(1-p)qpq) + \dots \\ p &= \frac{2}{3} \\ q &= \frac{1}{4} \\ &= pq \left(1 + (p(1-q) + (1-p)q) + (p^2(1-q)^2 + 2pq(1-p)(1-q) + (1-p)^2q^2) + \dots\right) \\ &= pq \sum_{i=0}^{\infty} (p(1-q) + (1-p)q)^i = pq \frac{1}{1 - p(1-q) - (1-p)q} \\ &= \frac{2}{3} \cdot \frac{1}{4} \cdot \frac{1}{1 - \frac{2}{3} \cdot \frac{3}{4} - \frac{1}{3} \cdot \frac{1}{4}} = \frac{1}{6} \cdot \frac{1}{1 - \frac{2}{2} - \frac{1}{12}} = \frac{1}{6} \cdot \frac{12}{5} = \frac{2}{5} \end{aligned}$$

A bit more formal: R ... robot's position; $x \in \mathbb{N}_0$

$$\mathbb{P}(R = (x, x+2)) = q \mathbb{P}(R = (x, x+1)) = qp \mathbb{P}(R = (x, x)) = qp \mathbb{P}(R = (x-1, x-1)) (p(1-q) + (1-p)q)$$

$$\mathbb{P}(R = (0, 0)) = 1$$

Hence, we have a recursive formula that can easily be made explicit:

$\mathbb{P}(R = (x, x)) = (p(1-q) + (1-p)q)^x$, hence, just as above, we obtain the sum

$$\sum_{i=0}^{\infty} \mathbb{P}(R = (i, i+2)) = pq \sum_{i=0}^{\infty} (p(1-p) + (1-p)q)^i = \frac{2}{5}$$

(2) Continuous two-dimensional random variable

The joint pdf of two random variables X and Y is defined by

$$f(x, y) = \begin{cases} c(x+2y), & 0 < y < 1 \text{ and } 0 < x < 2 \\ 0, & \text{otherwise} \end{cases}$$

(a) Find the value of c and the marginal distribution of Y .

(b) Find the joint cdf of X and Y .

(c) Find the marginal distribution of X and the pdf of $Z = \frac{9}{(X+1)^2}$.

$$a) 1 \doteq \int_{\mathbb{R}^2} f d\lambda^2 = \int_0^2 \int_0^1 c(x+2y) dx dy = c \int_0^1 (2+4y) dy = c(2+2) = 4c \Leftrightarrow c = \frac{1}{4}$$

$$\forall y \in]0, 1[: \int_0^2 f(x, y) dx = \frac{1}{2} + y = \frac{1+2y}{2}$$

$$f_Y(y) = \begin{cases} y + \frac{1}{2}, & \text{if } 0 < y < 1 \\ 0, & \text{otherwise} \end{cases}$$

$$b) \int_0^y \int_0^x f(\xi, \eta) d\xi d\eta = \int_0^y \int_0^x \left(\frac{\xi}{4} + \frac{\eta}{2}\right) d\xi d\eta = \int_0^y \left(\frac{x^2}{8} + \frac{\eta x}{2}\right) d\eta = \frac{x^2 y}{8} + \frac{y^2 x}{4}$$

$$F(x, y) = \begin{cases} 0 & , \text{if } x < 0 \text{ or } y < 0 \\ \frac{x^2 y}{8} + \frac{y^2 x}{4} & , \text{if } 0 \leq x \leq 2 \text{ and } 0 \leq y \leq 1 \\ \frac{x^2}{8} + \frac{x}{4} & , \text{if } 0 \leq x \leq 2 \text{ and } y > 1 \\ \frac{y}{2} + \frac{y^2}{2} & , \text{if } x > 2 \text{ and } 0 \leq y \leq 1 \\ 1 & , \text{if } x > 2 \text{ and } y > 1 \end{cases}$$

$$c) \forall x \in]0, 2[: \int_0^1 f(x, y) dy = \int_0^1 \left(\frac{x}{4} + \frac{y}{2}\right) dy = \frac{x}{4} + \frac{1}{4}$$

$$f_X(x) = \begin{cases} \frac{1}{4}(x+1), & \text{if } 0 < x < 2 \\ 0, & \text{otherwise} \end{cases}$$

$$g:]0, 2[\rightarrow]1, 9[: x \mapsto \frac{9}{(1+x)^2}$$

$$\forall x \in]0, 2[\forall z \in]1, 9[: z = \frac{9}{(1+x)^2} \Leftrightarrow z(1+x)^2 = 9 \Leftrightarrow 1+x = \sqrt{\frac{9}{z}} \Leftrightarrow x = \sqrt{\frac{9}{z}} - 1$$

$$h:]1, 9[\rightarrow]0, 2[: z \mapsto \sqrt{\frac{9}{z}} - 1 = 3z^{-\frac{1}{2}} - 1 \Rightarrow h'(z) = -\frac{3}{2} z^{-\frac{3}{2}}$$

$$f_X(h(z)) |h'(z)| = \frac{1}{4} \left(-\frac{3}{2} z^{-\frac{3}{2}} + 1\right) \frac{3}{2} z^{-\frac{3}{2}} = \frac{3}{8} z^{-\frac{3}{2}} - \frac{9}{16} z^{-3} = \frac{3}{8} \left(z^{-\frac{3}{2}} - \frac{3}{2} z^{-3}\right)$$

$$f_Z(z) = \begin{cases} \frac{3}{8} \left(z^{-\frac{3}{2}} - \frac{3}{2} z^{-3}\right), & \text{if } 1 < z < 9 \\ 0, & \text{otherwise} \end{cases}$$

(3) Chi squared distribution

Let X and Y be independent and identically distributed (i.i.d.) $\mathcal{N}(0,1)$ random variables.

Define $Z = \min\{X, Y\}$. Show that $Z^2 \sim \chi_1^2$, i.e. show that the pdf of Z^2 is given by

$$f_{Z^2}(z) = \frac{1}{\sqrt{2\pi}} \cdot z^{-\frac{1}{2}} \cdot e^{-\frac{z}{2}} \cdot \mathbf{1}_{\{z>0\}},$$

$$\begin{aligned} P(Z < z) &= 1 - P(Z \geq z) = 1 - P(\min\{X, Y\} \geq z) = 1 - P(X \geq z \wedge Y \geq z) \\ &= 1 - P(X \geq z) P(Y \geq z) = 1 - \frac{1}{2\pi} \left(\int_z^\infty e^{-s^2/2} ds \right)^2 \end{aligned}$$

$$f_Z(z) = -\frac{1}{\pi} \int_z^\infty e^{-s^2/2} ds \left(-e^{-z^2/2} \right) = \frac{1}{\pi} e^{-z^2/2} \int_z^\infty e^{-s^2/2} ds$$

$g: \mathbb{R} \setminus \{0\} \rightarrow \mathbb{R}^+ : z \mapsto z^2$ has got two right inverses, namely

$$h_-: (0, \infty) \rightarrow (-\infty, 0) : u \mapsto -\sqrt{u} \quad \text{and} \quad h_+: (0, \infty) \rightarrow (0, \infty) : u \mapsto \sqrt{u}$$

$h_-^{-1}(u) = -\frac{1}{2} u^{-\frac{1}{2}}$ and $h_+^{-1}(u) = \frac{1}{2} u^{-\frac{1}{2}}$, hence for all $u \in \mathbb{R}^+$, we obtain

$$\begin{aligned} f_{Z^2}(u) &= f_Z(\sqrt{u}) \frac{1}{2\sqrt{u}} + f_Z(-\sqrt{u}) \frac{1}{2\sqrt{u}} \\ &= \frac{1}{\pi} e^{-\frac{u}{2}} \int_{\sqrt{u}}^\infty e^{-s^2/2} ds \frac{1}{2\sqrt{u}} + \frac{1}{\pi} e^{-\frac{u}{2}} \int_{-\sqrt{u}}^\infty e^{-s^2/2} ds \\ &= \frac{1}{\sqrt{2\pi}} u^{\frac{1}{2}} e^{-\frac{u}{2}} \int_{-\infty}^\infty \frac{1}{\sqrt{2\pi}} e^{-s^2/2} ds = \frac{1}{\sqrt{2\pi}} u^{\frac{1}{2}} e^{-\frac{u}{2}} \end{aligned}$$

$$\text{Thus, } f_{Z^2}(u) = \frac{1}{\sqrt{2\pi}} u^{\frac{1}{2}} e^{-\frac{u}{2}} \mathbb{1}_{\mathbb{R}^+}(u)$$

(4) Random variables on the unit disk

Let (X, Y) be uniformly distributed on the unit disk $\{f(x; y) : x^2 + y^2 \leq 1\}$. Let

$$R = \sqrt{X^2 + Y^2}.$$

Find the cdf, pdf, and the expectation the random variable R .

$$K := \{(x, y) \in \mathbb{R}^2 \mid x^2 + y^2 \leq 1\}$$

$$1 \stackrel{!}{=} \int_K c \, d\lambda^2 = c \int_0^1 \int_0^{2\pi} r \, d\varphi \, dr = c\pi \Leftrightarrow c = \frac{1}{\pi}$$

$$f_{X,Y}(x, y) = \begin{cases} \frac{1}{\pi}, & \text{if } (x, y) \in K \\ 0, & \text{otherwise} \end{cases}$$

$$h: \mathbb{R}^+ \times [0, 2\pi[\rightarrow \mathbb{R}^2: (r, \varphi) \mapsto (r \cos(\varphi), r \sin(\varphi))$$

$$f_{R,\varphi}(r, \varphi) = f_{X,Y}(h(r, \varphi)) \cdot r = \begin{cases} \frac{r}{\pi}, & \text{if } 0 \leq r \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$f_R(r) = \int_0^{2\pi} f_{R,\varphi}(r, \varphi) \, d\varphi \cdot \mathbb{1}_{[0,1)}(r) = 2r \cdot \mathbb{1}_{[0,1)}(r)$$

$$F_R(r) = \int_0^r f_R(s) \, ds = \begin{cases} 0, & \text{if } r \leq 0 \\ r^2, & \text{if } 0 < r < 1 \\ 1, & \text{otherwise} \end{cases}$$

$$\mathbb{E}(R) = \int_{-\infty}^{\infty} f_R(r) \cdot r \, dr = \int_0^1 2r^2 \, dr = \frac{2}{3}$$

(5) Transformations

Suppose X and Y are independent gamma distributed random variables with $X \sim \text{Gamma}(\alpha_1, \beta)$ and $Y \sim \text{Gamma}(\alpha_2, \beta)$. Consider the following two random variables

$$U = X + Y \quad \text{and} \quad V = \frac{X}{X + Y}.$$

(a) Show that $U \sim \text{Gamma}(\alpha_1 + \alpha_2, \beta)$.

(b) Show that U and V are also independent random variables.

$$f_X(x) = \frac{1}{\Gamma(\alpha_1) \beta^{\alpha_1}} x^{\alpha_1-1} e^{-\frac{x}{\beta}} \mathbb{1}_{\mathbb{R}^+}(x) \quad \alpha_1, \alpha_2, \beta > 0$$

$$f_Y(y) = \frac{1}{\Gamma(\alpha_2) \beta^{\alpha_2}} y^{\alpha_2-1} e^{-\frac{y}{\beta}} \mathbb{1}_{\mathbb{R}^+}(y)$$

a) As X and Y are independent, we have $f_{X,Y}(x,y) = f_X(x) f_Y(y)$

$$g: \mathbb{R}^2 \rightarrow \mathbb{R}^+ \times (0,1): (x,y) \mapsto (x+y, \frac{x}{x+y})$$

$$\forall x,y,u,v \in \mathbb{R}^+: (u,v) = (x+y, \frac{x}{x+y}) \Leftrightarrow u = x+y \wedge v = \frac{x}{x+y} \Leftrightarrow u = x+y \wedge v = \frac{x}{u} \Leftrightarrow y = u-x \wedge x = vu \Leftrightarrow x = vu \wedge y = u(1-v)$$

Hence, $h: \mathbb{R}^+ \times (0,1) \rightarrow \mathbb{R}^2: (u,v) \mapsto (vu, u(1-v))$ is the inverse of g

$$dh(u,v) = \begin{pmatrix} v & u \\ 1-v & -u \end{pmatrix}, \text{ and } \det(dh(u,v)) = -uv - u(1-v) = -u$$

$$f_{u,v}(u,v) = f_{X,Y}(h(u,v)) |\det(dh(u,v))| = f_X(vu) f_Y(u(1-v)) u$$

$$C(\alpha) := \frac{1}{\Gamma(\alpha) \beta^\alpha}$$

$$f_u(u) = \int_{\mathbb{R}} f(u,v) dv = C(\alpha_1) C(\alpha_2) u^{(\alpha_1+\alpha_2)-1} \int_{\mathbb{R}} \mathbb{1}_{\mathbb{R}^+}(vu) \mathbb{1}_{\mathbb{R}^+}(u(1-v)) v^{\alpha_1-1} (1-v)^{\alpha_2-1} e^{-\frac{uv}{\beta}} e^{-\frac{u(1-v)}{\beta}} dv$$

$$e^{-\frac{uv}{\beta}} e^{-\frac{u(1-v)}{\beta}} = e^{-\frac{uv}{\beta}} e^{\left(\frac{uv}{\beta} - \frac{u}{\beta}\right)} = e^{-\frac{u}{\beta}}, \text{ and } uv > 0 \wedge u(1-v) > 0 \Leftrightarrow u > 0 \wedge 0 < v < 1$$

$$\int_0^1 v^{\alpha_1-1} (1-v)^{\alpha_2-1} dv = B(\alpha_1, \alpha_2) = \frac{\Gamma(\alpha_1) \Gamma(\alpha_2)}{\Gamma(\alpha_1 + \alpha_2)}, \text{ hence} \quad \alpha := \alpha_1 + \alpha_2$$

$$\text{Finally, we obtain } f_u(u) = \frac{1}{\Gamma(\alpha_1 + \alpha_2) \beta^{(\alpha_1 + \alpha_2)}} u^{(\alpha_1 + \alpha_2)-1} e^{-\frac{u}{\beta}} \mathbb{1}_{\mathbb{R}^+}(u)$$

$$b) f_V(v) = \int_{\mathbb{R}} f(u,v) du = C(\alpha_1) C(\alpha_2) v^{\alpha_1-1} (1-v)^{\alpha_2-1} \mathbb{1}_{(0,1)}(v) \int_{\mathbb{R}^+} u^{\alpha-1} e^{-\frac{u}{\beta}} du$$

$$\frac{u}{\beta} = w \Leftrightarrow \frac{dw}{du} = \frac{1}{\beta} \Leftrightarrow du = \beta dw$$

$$\int_{\mathbb{R}^+} u^{\alpha-1} e^{-\frac{u}{\beta}} du = \int_{\mathbb{R}^+} (\beta w)^{\alpha-1} e^{-w} \beta dw = \beta^\alpha \Gamma(\alpha)$$

$$\text{Thus, } f_V(v) = \frac{1}{\Gamma(\alpha_1) \Gamma(\alpha_2)} v^{\alpha_1-1} (1-v)^{\alpha_2-1} \Gamma(\alpha) \mathbb{1}_{(0,1)}(v)$$

$$f_u(u) f_V(v) = \frac{1}{\Gamma(\alpha_1) \Gamma(\alpha_2)} v^{\alpha_1-1} (1-v)^{\alpha_2-1} u^{\alpha_1-1} u^{\alpha_2-1} u \mathbb{1}_{(0,1)}(v) \mathbb{1}_{\mathbb{R}^+}(u) = f_{u,v}(u,v)$$

Hence, U and V are independent!