

# APPROXIMATE BAYESIAN COMPUTATION FOR ARCHIMEDEAN COPULA MODELS

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$$\pi(\theta|y) \propto p(y|\theta)\pi(\theta)$$

**Problem:** How to perform Bayesian inference when the likelihood function  $p(y|\theta)$  is computationally intractable?

**Solution:** If we can easily simulate from the likelihood, ABC methods provide a possible way.

$$\pi_{\text{ABC}}(\theta|y) \propto \int \mathcal{I}(\|y^* - y\| < \epsilon) \pi(\theta|y^*) dy^* \approx \pi(\theta|y)$$

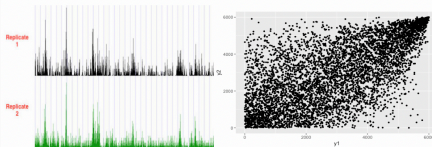
# MOTIVATION: ARCHIMEDEAN COPULA MODELS

**J-dimensional copula:**  $C(u_1, \dots, u_J) = P(U_1 \leq u_1, \dots, U_J \leq u_J)$  where  $U_1, \dots, U_J$  are uniform random variables and  $C : [0, 1]^J \rightarrow [0, 1]$ .

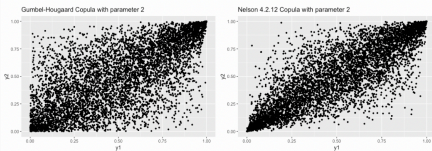
**J-dimensional Archimedean copula:**

$\psi(C(u_1, \dots, u_J); \theta) = \psi(u_1; \theta) + \dots + \psi(u_J; \theta)$  where  $\theta$  is an **association parameter** describing the strength of the dependence between  $U_j$ 's, and  $\psi$  is a generator function specific to each Archimedean copula such that  $\psi : [0, 1] \rightarrow [0, \infty)$ .

ChIP-Seq data and its rank scatter plot



2 dimensional Archimedean copulas with  $\theta = 2$



Gumbel Hougaard copula generator function:

$$\psi(u) = (-\log(u))^\theta, \theta \in [1, \infty)$$

So its distribution function is,

$$C(u_1, \dots, u_n) = \psi^{-1}(\psi(u_1) + \dots + \psi(u_n)) = \exp\{-((-\log(u_1))^\theta + \dots + (-\log(u_n))^\theta)^{1/\theta}\}$$

Using ABC to estimate  $\theta$  is possible because,

- likelihood is unavailable
- it easy to simulate from the model

Requirements:

- a proposal,  $q(\cdot)$
  - the observed data,  $y$
  - a distance function,  $||\cdot||$
  - a tolerance level,  $\epsilon$
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## Rejection Sampling (Sisson et al. 2018)

1. Simulate  $\theta^* \sim q(\cdot)$
2. Simulate  $y \sim p(\cdot|\theta^*)$
3. if  $||y - y_{\text{obs}}|| < \epsilon$  then accept  $\theta^*$   
with probability  $\frac{\pi(\theta^*)}{Kq(\theta^*)}$ .
4. Repeat above steps N times.

## MCMC (Marjoram et al. 2003)

1. Simulate  $\theta^* \sim q(\cdot)$
2. Simulate  $y \sim p(\cdot|\theta^*)$
3. if  $||y - y_{\text{obs}}|| < \epsilon$  then accept  $\theta^*$   
with probability  
 $\min\{1, \frac{\pi(\theta^*)q(\theta|\theta^*)}{\pi(\theta)q(\theta^*|\theta)}\}$  else stay at  $\theta$ .
4. Repeat above steps N times.

- Should we use the entire dataset  $y$  or an appropriate summary statistic,  $s$ ?
- Choice of distance function and summary statistic.
- Computationally expensive?
- Choice of tolerance level and other tuning parameters.

## ABC MCMC: SIMULATION STUDY

- Prior on  $\theta$  is **Gamma(shape=1, scale=2)** truncated at 1.
- Proposal is a random walk i.e.  $N(\theta^{(i-1)}, \sigma^2)$  truncated at 1.
- Tolerance,  $\epsilon$  and  $\sigma^2$  are chosen such that the **acceptance rate was 20-35%**.
- **Absolute difference** is used as the distance function for two dimensions and for more than 2, **Frobenius norm** is used.
- For each simulation,  $y_{\text{obs}}$  consisted of 100 data points.
- Number of simulations = 100 and  $N=10000$ .

Dimensions:	2		3		5	
Summary statistic	Estimate	MSE	Estimate	MSE	Estimate	MSE
Spearman's rank $\rho$	3.06	1.32	2.94	1.12	2.19	0.07
Kendall's $\tau$	2.60	0.47	2.55	0.39	2.16	0.05

**Table:** Results from ABC MCMC algorithm for Gumbel-Hougaard Copula with  $\theta = 2$ .

# ABC MCMC WITH KERNEL FUNCTION

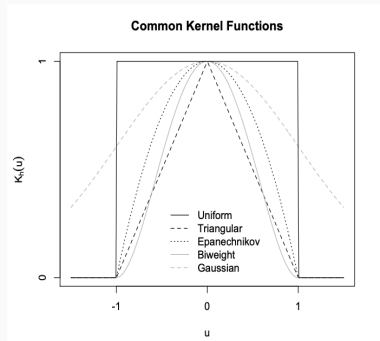
**Kernel function:**  $K_\epsilon(u) = \frac{1}{\epsilon} K(\frac{u}{\epsilon})$  where  $K(u) = \frac{1}{\sqrt{2\pi}} e^{-u^2/2}$

**Previously:** Accept  $\theta^*$  with probability  $\min\{1, \frac{\pi(\theta^*)q(\theta|\theta^*)}{\pi(\theta)q(\theta^*|\theta)}\}$ .

**Now:** Accept  $\theta^*$  with probability  $\min\{1, \frac{K_\epsilon(||s^* - s_{obs}||) \pi(\theta^*)q(\theta|\theta^*)}{K_\epsilon(||s^{(l-1)} - s_{obs}||) \pi(\theta)q(\theta^*|\theta)}\}$ .

Dimension	Estimate	MSE
2	2.24	0.12
3	2.14	0.09
5	2.18	0.10

**Table:** Results from ABC MCMC algorithm (using Kendall's  $\tau$ ) with kernel function for Gumbel Hougaard Copula with  $\theta = 2$ .



Note: Smoothing functions are discussed in Sisson et al. 2018 for ABC Rejection algorithms and something similar in Fernhead and Prangle 2011.



## CONCLUSION

- For large data sets, using  $y_{\text{obs}}$  instead of summary statistic is not advisable due to computational costs.
- ABC MCMC requires a lot of tuning to run well. If possible, a different ABC algorithm can be preferred.
- Kendall's  $\tau$  estimates better than Spearman's rank correlation.
- Estimation for higher dimensional copula works better because there is more information for a single parameter.
- ABC MCMC with kernel function seems like a better approach in this case, at least for lower dimensional copula.