# A Spatial Point Process Model for Viral Infections

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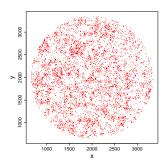
(joint work with **Josh Goldstein**, Ivan Simeonov, John Fricks, and Francesca Chiaromonte)

# Spatial point process modeling of virus infections

- Biologists are often interested in investigating the progression of viral infections
- Our goal: use data from imaging of cell cultures to study the spatial structure of an infection
- An in vitro cell culture study identifies and locates cells infected with two strains of the human respiratory syncytial virus (RSV-A and RSV-B)

### Question:

How does the presence of an infected cell impact infections in neighboring cells?



Points represent locations of cells infected with RSV.

## Outline

- Spatial point processes in the plane provide a natural framework here
- Each point represents the 2D coordinates of an infected cell
- Goal: Infer spatial interaction among cells

#### Contributions of this work:

- A new spatial point process model for the RSV data
- Inferential methods for this computationally challenging problem
- Draw scientific conclusions from fitted model

# Spatial point processes

A spatial point process is a stochastic process, a realization of which consists of a set of points  $X = (x_1, ..., x_n)$  in a bounded region  $W \subseteq \mathbb{R}^d$ .

Some SPPs can be used to model interactions:

$$f(X|\Theta) = \lambda^n \prod_{i \neq j} \phi(x_i, x_j)$$

where  $\phi(x_i, x_j)$  is the interaction function between points i and j. In the homogeneous case,  $\phi(x_i, x_j) = \phi(\|x_i - x_j\|) = \phi(r)$ 

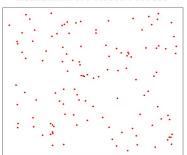
The Strauss process is a simple example,

$$\phi(r) = \begin{cases} \gamma, & 0 < r \le R \\ 1, & r > R \end{cases}$$

for  $0 \le \gamma \le 1$ . Since  $\phi(r) \le 1$  this is a repulsion point process.

# Poisson process vs. Strauss process

#### Realization of Poisson Process

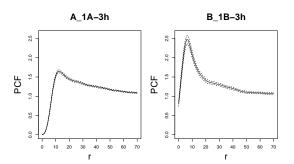


#### Realization of Strauss Process

## Exploratory analysis of RSV data: Need for a new model

The pair correlation function (PCF) g(r) is an exploratory summary statistic that tells us the attraction-repulsion behavior of points separated by distance r in a spatial point process.

• A value of g(r) > 1 indicates attraction, a tendency for points to cluster at distance r. Similarly, g(r) < 1 indicates repulsion at distance r. For our data:



Observed attraction-repulsion in RSV data varies smoothly in r.

## New attraction-repulsion model: Interaction function

Goal: Allow attraction-repulsion to vary smoothly with distance to model observed RSV behavior.

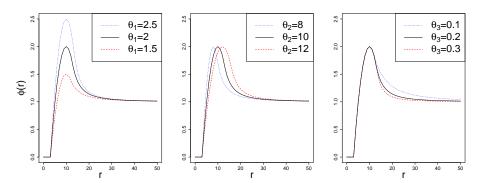
The interaction function  $\phi(r)$  is defined piecewise,

$$\phi(r) = \begin{cases} 0, & 0 \le r \le R \\ \theta_1 - \left(\frac{\sqrt{\theta_1}}{\theta_2 - R}(r - \theta_2)\right)^2, & R < r \le r_1 \\ 1 + \frac{1}{(\theta_3(r - r_2))^2}, & r > r_1 \end{cases}$$

#### where

- $\theta_1$ : value of  $\phi(\cdot)$  at the peak
- $\theta_2$ : the value of r at the peak
- $\theta_3$ : rate of descent after the peak
- R: minimum allowable distance between points
- $r_1$ ,  $r_2$ : chosen to ensure  $\phi(\cdot)$  is continuously differentiable.

## New attraction-repulsion model: Interaction function



- $\phi(r) > 1$ : attraction, points tend to cluster at distance r
- $\phi(r) < 1$ : repulsion

## Attraction-repulsion model

The likelihood can be written as

$$\mathcal{L}(X|\Theta) = \frac{f(X|\Theta)}{c(\Theta)}, f(X|\Theta) = \lambda^n \Big[ \prod_{i=1}^n e^{\min \Big[ \sum_{i \neq j} log(\phi(x_i, x_j)), k \Big]} \Big]$$

### Model parameters:

- ullet  $\lambda$  is the intensity of the process
- $\theta_1, \theta_2, \theta_3$  control the shape of  $\phi(r)$ .
- R is the minimum distance allowed between points
- k is a truncation constant necessary to prevent "clumping" behavior

**Important:**  $c(\Theta)$  is intractable. This makes computing very challenging.

### Inference

Let  $\Theta = (\lambda, k, \theta_1, \theta_2, \theta_3)$ . Likelihood  $\mathcal{L}(X|\Theta)$ . Is maximum likelihood an option?

- Unknown normalizing function poses a major challenge.
- Existing methods: MCMC-maximum likelihood (e.g. Geyer and Thompson, 1992), essentially importance sampling combined with MCMC, should work in principle.
- However, this algorithm is infeasible in practice:
  - Problem 1: Initial values are important and difficult to obtain.
  - Problem 2: Standard errors are difficult because we need an estimate of inverse Hessian, in turn need estimate of gradients of unnormalized likelihood. Difficult/intractable for our model.
- Bayesian inference (contrary to what some people might think?) is more tractable and convenient in this normalizing function problem.

# Bayesian Inference

- Bayesian inference to the rescue: turns out that, although it is challenging, computing for Bayesian inference is feasible even though MCMCMI F is not.
- ullet Bayesian inference for  $\Theta$  is based on the posterior distribution

$$\pi(\Theta|X) \propto \mathcal{L}(X|\Theta)p(\Theta) = \frac{f(X|\Theta)p(\Theta)}{c(\Theta)}$$

- Markov chain Monte Carlo (MCMC) is a convenient approach to learning about  $\pi(\Theta|X)$ .
- Choose a gamma prior on k, prior for remaining parameters are uniform over a scientifically plausible range.

## **MCMC**

- Construct a Markov chain with stationary distribution  $\pi(\Theta|X)$ .
- In MCMC, propose  $\Theta'$  from  $q(\Theta, \Theta')$  and calculate the following acceptance probablity:

$$\alpha = \min\left(1, \frac{p(\Theta')q(\Theta', \Theta)f(X|\Theta')}{p(\Theta)q(\Theta, \Theta')f(X|\Theta)} \frac{c(\Theta)}{c(\Theta')}\right)$$

- The intractable normalizing constant  $c(\Theta)$  does not cancel.
- Traditional MCMC methods cannot be applied.

# Auxiliary Variable MCMC algorithm

Idea: Introduce an auxiliary variable. That is, augment the target distribution.

- Perfect sampling for  $f(X \mid \Theta)$  unavailable so cannot use Møller et al. (2006) and Murray et al. (2007).
- Double Metropolis-Hastings algorithm of Liang (2010).
- Two nested MCMC samplers; the "inner sampler" generates an auxiliary point pattern at each step of the "outer" sampler.

# Computational challenges

The largest datasets consist of 13,000-14,000 spatial locations. For data this large, the nested samplers are expensive; the inner sampler must be run for thousands of iterations at each step of the outer sampler.

- Inner sampler updates fast in practice since we only propose to add or remove a single point (birth-death sampler)
- R too slow, code in C and optimize.
- Greatly reduce computing by truncating the interaction function for large values of r (evaluate  $\phi(r)$  to 1 when  $r > R_{max}$ ).
- Inference for three replicates of the largest dataset takes a few days on the Lion-X cluster.

### Conclusions

Can make meaningful scientific conclusions as a result of inference on model parameters across multiple RSV experiments (e.g. RSV-B infected cells have a higher propensity to lump together than RSV-A; suggests RSV-B induces stronger increase in susceptibility to infection).

### Advantages of our method:

- Model captures the complex scale-varying attraction and repulsion behavior observed in the RSV dataset; this flexibility is not available in other models
- Parametric specification of the interaction function lets us draw meaningful conclusions based on parameter inference
- We use a simulation-based approach with PCFs to check the goodness of fit for our model to the data.
- Inference works well for simulated examples we recover truth.

Goldstein, Haran, Simeonov, Fricks, Chiaromonte Biometrics, 2015.