




Study EM and MCMC algorithms for Gaussian mixture model

Dongkuan Xu

College of IST, Penn State University



Gaussian mixture model (GMM)

❑ Definition

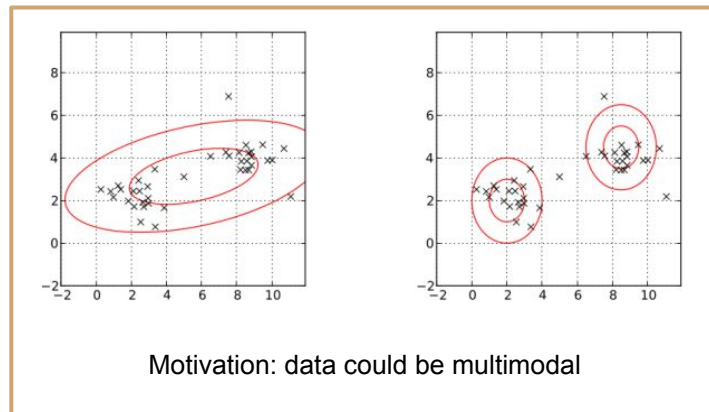
$$p(\vec{x}) = \sum_{i=1}^K \phi_i \mathcal{N}(\vec{x} \mid \vec{\mu}_i, \Sigma_i)$$
$$\mathcal{N}(\vec{x} \mid \vec{\mu}_i, \Sigma_i) = \frac{1}{\sqrt{(2\pi)^K |\Sigma_i|}} \exp\left(-\frac{1}{2} (\vec{x} - \vec{\mu}_i)^T \Sigma_i^{-1} (\vec{x} - \vec{\mu}_i)\right)$$
$$\sum_{i=1}^K \phi_i = 1$$

❑ Related work

- ❑ GMM based on EM algorithm [1-3]
- ❑ Bayesian GMM based on EM or MCMC [4-5]
- ❑ GMM with unknown mixture number [6-7]

❑ Application

- ❑ Data clustering
- ❑ Image segmentation
- ❑ Time series analysis
- ❑ Genetics



Problem 1: log-likelihood is hard to calculate when MLE

- ❑ Maximum likelihood estimation

- ❑ Likelihood function

$$\mathcal{L}(\theta; x)$$

- ❑ Maximum likelihood estimate

$$\hat{\theta} \in \{\arg \max_{\theta \in \Theta} \mathcal{L}(\theta; x)\}$$

- ❑ Log-likelihood is hard to calculate by direct derivative: summation operations in the logarithmic operations

$$\sum_{i=1}^N \log \left\{ \sum_{k=1}^K \pi_k \mathcal{N}(x_i | \mu_k, \Sigma_k) \right\}$$

Solution: Expectation-maximization (EM) for GMM

- ❑ Log-likelihood function

$$\sum_{i=1}^N \log \left\{ \sum_{k=1}^K \pi_k \mathcal{N}(x_i | \mu_k, \Sigma_k) \right\}$$

- ❑ Estimation step

$$\gamma(i, k) = \frac{\pi_k \mathcal{N}(x_i | \mu_k, \Sigma_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(x_i | \mu_j, \Sigma_j)}$$

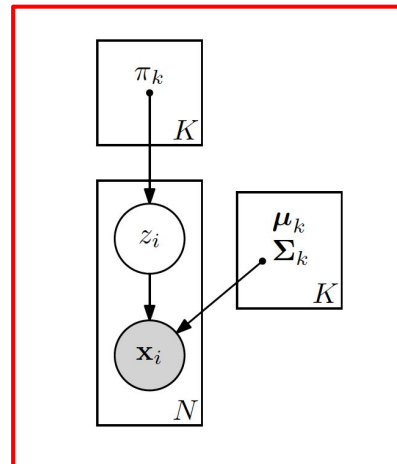
- ❑ Maximization step

$$\mu_k = \frac{1}{N_k} \sum_{i=1}^N \gamma(i, k) x_i$$

$$\Sigma_k = \frac{1}{N_k} \sum_{i=1}^N \gamma(i, k) (x_i - \mu_k)(x_i - \mu_k)^T$$

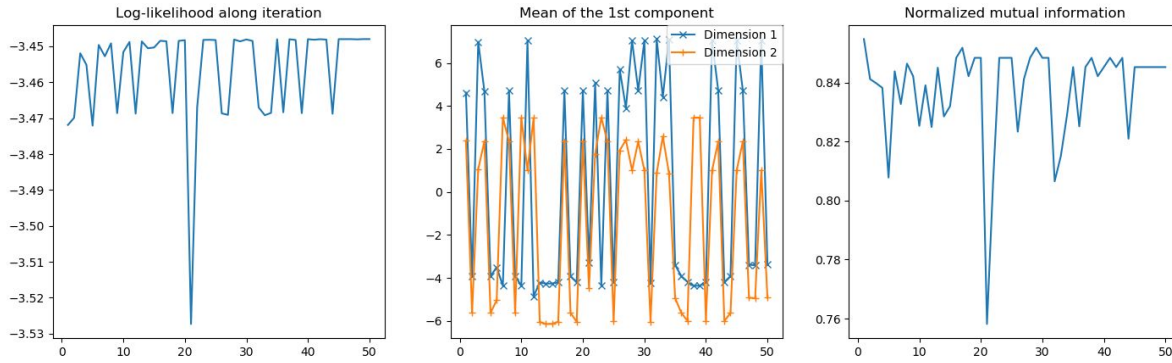
$$\pi_k = N_k / N$$

$$N_k = \sum_{i=1}^N \gamma(i, k)$$

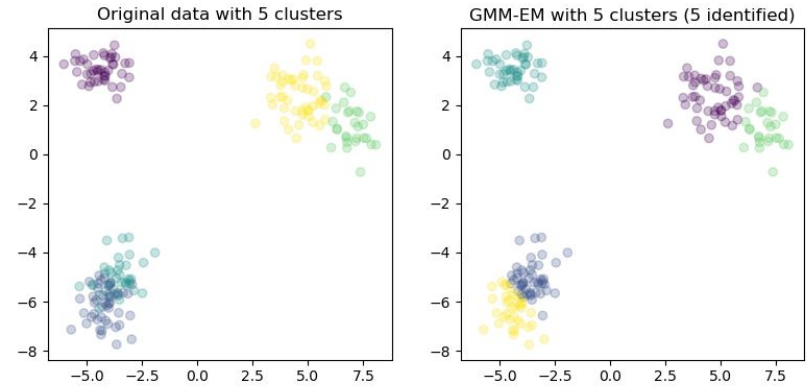


GMM with EM

- ❑ Stopping criteria
 - ❑ (1) Maximum iteration number of EM (2) A threshold about the gain on log-likelihood
- ❑ Running time per iteration: 0.00326 s
- ❑ Iteration: (1) Log-likelihood (2) Mean of the 1st component (3) Normalized mutual information (NMI)



- ❑ Clustering results:



Problem 2: Too many free parameters for EM when high-dimensional

Parameters of GMM

$$\mu_k = \frac{1}{N_k} \sum_{i=1}^N \gamma(i, k) x_i \quad \Sigma_k = \frac{1}{N_k} \sum_{i=1}^N \gamma(i, k) (x_i - \mu_k)(x_i - \mu_k)^T \quad \pi_k = N_k/N$$

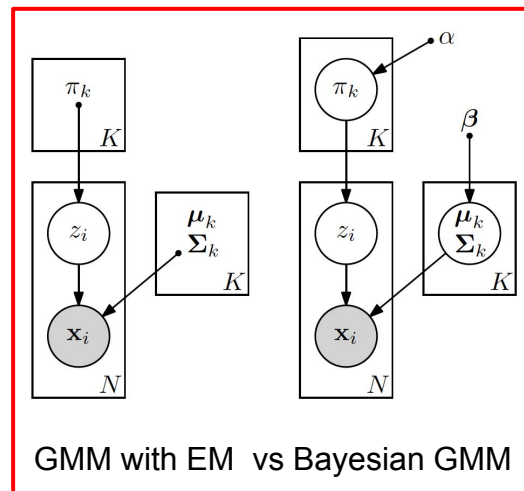
- MLE approach based on EM may fail due to singularities or degeneracies
- A Bayesian approach alleviate these by treating $\Theta = (\pi, \{\mu_k\}, \{\Sigma_k\})$ s random variables and working with distributions over Θ rather than point estimates
- Choose conjugate priors to marginalize over the parameters
 - Symmetric Dirichlet prior for mixture weights
 - Normal-inverse-Wishart (NIW) prior for component parameters

$$\pi \sim \text{Dir}(\alpha/K \mathbf{1})$$

$$z_i \sim \pi$$

$$\mu_k, \Sigma_k \sim \text{NIW}(\mathbf{m}_0, \kappa_0, \nu_0, \mathbf{S}_0)$$

$$\mathbf{x}_i \sim \mathcal{N}(\mu_{z_i}, \Sigma_{z_i})$$



Solution: Bayesian approach for GMM

❑ Algorithm of Collapsed Gibbs sampler for GMM:

Choose an initial \mathbf{z} .

for T iterations **do**

for $i = 1$ to N **do**

 Remove \mathbf{x}_i 's statistics from component z_i .

for $k = 1$ to K **do**

 Calculate $P(z_i = k | \mathbf{z}_{\setminus i}, \mathcal{X}, \boldsymbol{\alpha}, \boldsymbol{\beta}) \propto P(z_i = k | \mathbf{z}_{\setminus i}, \boldsymbol{\alpha}) p(\mathbf{x}_i | \mathcal{X}_{k \setminus i}, \boldsymbol{\beta})$.

end for

 Sample k_{new} from $P(z_i | \mathbf{z}_{\setminus i}, \mathcal{X}, \boldsymbol{\alpha}, \boldsymbol{\beta})$ after normalizing.

 Add \mathbf{x}_i 's statistics to the component $z_i = k_{\text{new}}$.

end for

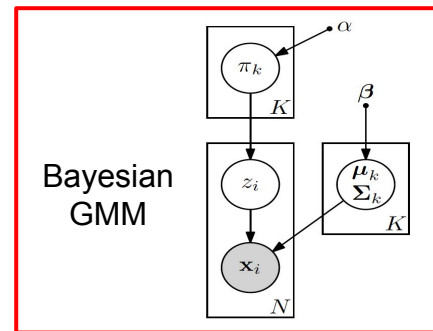
end for

▷ Gibbs sampling iterations

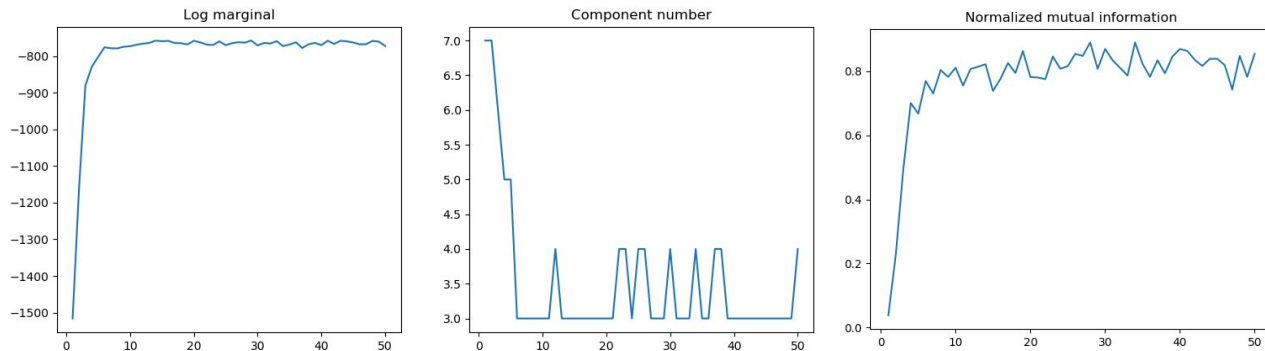
▷ Old assignment for \mathbf{x}_i

▷ Every possible component

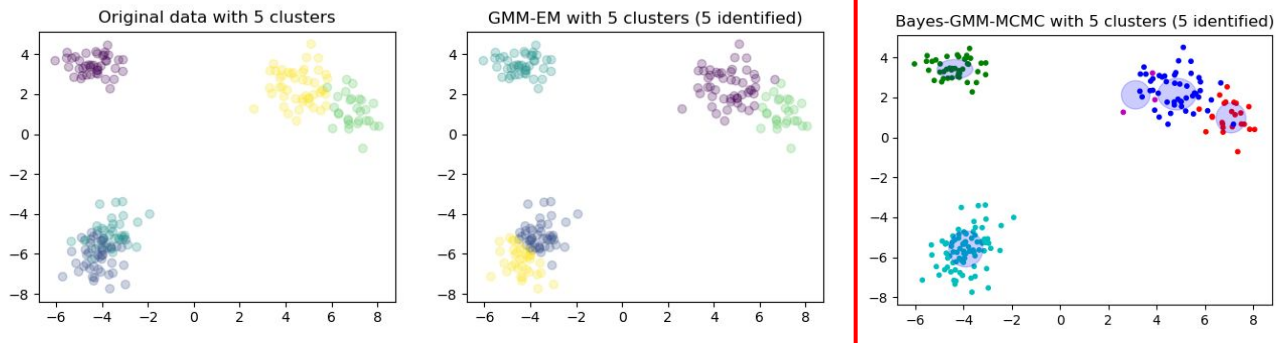
▷ New assignment for \mathbf{x}_i



- ❑ Stopping criteria:
 - ❑ (1) Maximum iteration number of EM
 - ❑ (2) A threshold about the gain of log marginal of data and component assignments: $p(X, z)$ (**possible**)
- ❑ Running time per iteration: 0.00492 (**higher than the one of GMM-EM**)
- ❑ Iteration: (1) Log marginal of $p(X, z)$ (2) Normalized mutual information (NMI) (3) Component number



- ❑ Clustering result: **a little worse than GMM-EM**

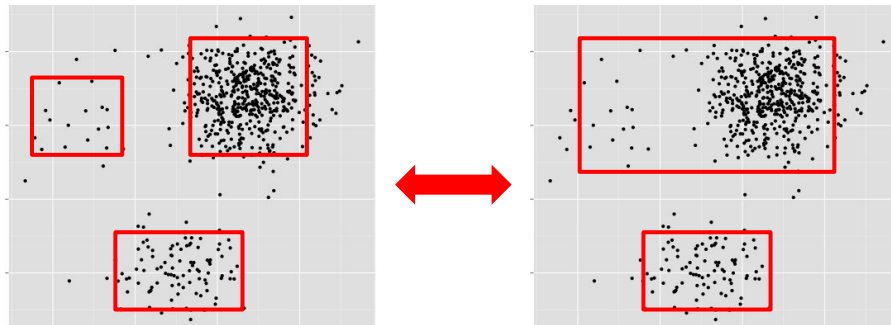


Problem 3: GMM with unknown mixture number

- Unknown mixture number

- MLE, EM do not work

$$\begin{array}{cc} \sum_{i=1}^N \log \left\{ \sum_{k=1}^K \pi_k \mathcal{N}(x_i | \mu_k, \Sigma_k) \right\} & \gamma(i, k) = \frac{\pi_k \mathcal{N}(x_i | \mu_k, \Sigma_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(x_i | \mu_j, \Sigma_j)} \\ \text{log-likelihood} & \text{log-likelihood} \end{array}$$



- A number of possible solutions

- ❑ reversible jump MCMC (Richardson and Green 1997, Gruet et al. 1999)
- ❑ Bayes factors (Kass and Raftery 1995, Richardson and Green 1997)
- ❑ entropy distance or K-L divergence (Mengersen and Robert 1996, Sahu and Cheng 2003)
- ❑ birth-and-death processes (Stephens 2000a, Cappé et al. 2002)

Solution: Reversible jump MCMC for GMM with unknown mixture number

- ❑ Idea: birth and death moves
 - ❑ Add a new normal component in the mixture generated from the prior, or remove one component, according to the acceptance probability
- ❑ Assumption:
 - ❑ Competing models can be enumerable and represented as: $\mathcal{M} = \{\mathcal{M}_1, \mathcal{M}_2, \dots\}$
 - ❑ Current state of Markov chain is: (k, θ_k)
- ❑ GMM with RJMCMC:
 - ❑ Propose a visit from current model (k, θ_k) to next model $(\theta_{k'}, k')$
 - ❑ Accept the visit or not
 - ❑ Repeat proposing model visit until stopping criteria are met
 - ❑ Clustering based on standard GMM

Reversible jump MCMC for GMM with unknown mixture number

❑ Algorithm of GMM with RJMCMC:

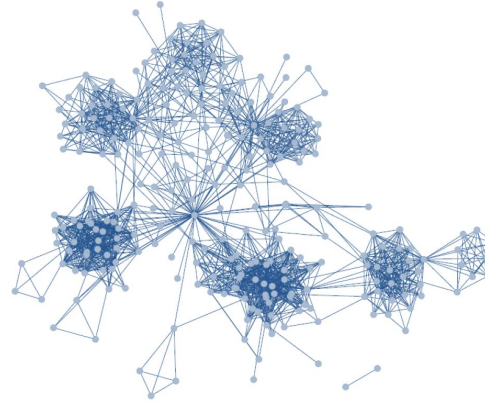
- ❑ (1) Initialize current model indicator k
- ❑ (2) Propose a visit from mode M_k to model $M_{k'}$ with probability $J(k \rightarrow k')$
- ❑ (3) Sample parameter u of GMM from a proposal density $q(u|\theta_k, k, k')$
- ❑ (4) Set $(\theta_{k'}, u') = g_{k,k'}(\theta_k, u)$, where $g_{k,k'}(\cdot)$ is a bijection, where u and u' play the role of matching the dimensions of both vectors
- ❑ (5) The acceptance probability of the new model $(\theta_{k'}, k')$ the minimum between 1 and :

$$\underbrace{\frac{p(y|\theta_{k'}, k')p(\theta_{k'})p(k')}{p(y|\theta_k, k)p(\theta_k)p(k)}}_{\text{model ratio}} \underbrace{\frac{J(k' \rightarrow k)q(u'|\theta_{k'}, k', k)}{J(k \rightarrow k')q(u|\theta_k, k, k')}}_{\text{proposal ratio}} \left| \frac{\partial g_{k,k'}(\theta_k, u)}{\partial(\theta_k, u)} \right|$$

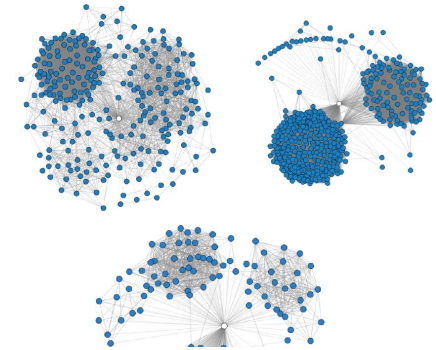
- ❑ (6) Calculate $M_{k'}$
- ❑ (7) $\text{ite} = \text{ite} + 1$
- ❑ (8) if $\text{ite} < \text{max}$ or MCMC standard error $> \varphi$, then go to (2)
- ❑ (9) output cluster assignment for all data samples

Network-structure data clustering

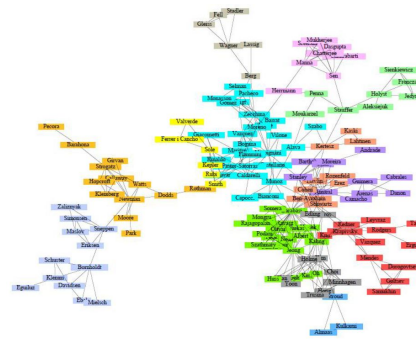
- ❑ Network data is ubiquitous
 - ❑ Web network
 - ❑ Social network
 - ❑ Biological network, etc.
- ❑ Network clustering
 - ❑ Detect sub-networks that satisfy certain properties
 - ❑ Many connections within clusters and few connections across clusters



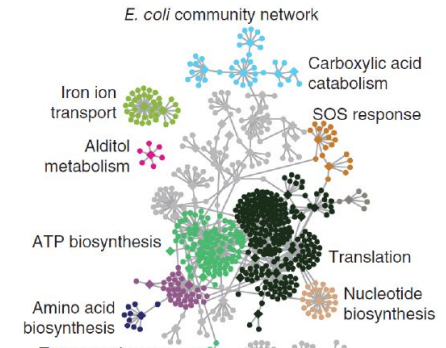
Web network



Social network



Co-author network



Gene network

Comparison between EM-GMM and Bayesian-GMM applied to network data clustering

- ❑ Methods:
 - ❑ GMM based on EM
 - ❑ Bayesian GMM based on EM
 - ❑ Bayesian GMM based on MCMC

- ❑ Data set 1: Newsgroup20 (600 instances, 6 clusters, 600-dimension, sparse)

Method	GMM-EM	Bayes-GMM-EM	Bayes-GMM-MCMC
NMI	0.1902517	0.1325872	0.0450131
Identified components	6	6	2

Comparison between EM-GMM and Bayesian-GMM applied to network data clustering

- ❑ Data set 2: Synthetic (200 instances, 10 cluster, 2-dimension)

