Gaussian processes for inference with implicit likelihoods

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Complex Scientific Models

- Scientists working in the physical and natural sciences are often interested in learning about the mechanisms or "laws" and processes underlying physical phenomena.
- ► These models may be useful for predictions/projections.
- Critical to work with the model provided by the scientists. Purely statistical approaches may not answer questions of interest or permit sound predictions.
- These scientific models may be
 - Numerical solutions of mathematical (deterministic) models or stochastic models that reflect scientific processes.
 - Translated into computer code to study simulations of the physical processes for different parameters/conditions.

Some Challenges Posed by Complex Models

- As models become more scientifically plausible, they typically become more complex. Inference for unknown parameters may be challenging for the following reasons:
 - Simulations from the model may be computationally expensive.
 - May not be possible to write closed-form expressions relating input (parameters) to output.
 - The likelihood function may be very expensive to evaluate: hard to optimize or use Monte Carlo methods.
 - There are non-ignorable discrepancies between the model and reality so even a perfectly calibrated model will not match observations well.
- ► The likelihood function is often *implicit* or has to be treated as such.

Two Examples

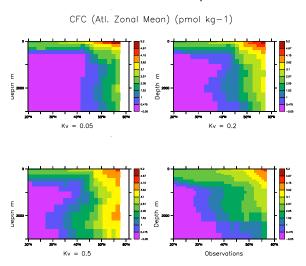
- I Climate: An Earth System Model of Intermediate Complexity (EMIC) for projecting the behavior of global ocean circulation systems.
 - Deterministic
 - Model runs are expensive
 - High-dimensional multivariate spatial process
- II Disease Dynamics: A Gravity Time Series Susceptible-Infected-Recovered (TSIR) model for the spread of infectious disease (measles).
 - Stochastic
 - Likelihood is expensive to evaluate
 - Space-time process with large number of "no incidence" observations (0s)

Climate Models: Learning About K_v

The meridional overturning circulation (MOC) has a strong influence on global climate. Collapse of this ocean "conveyor belt" may result in dramatic climate change. K_{ν} is a key climate model parameter that influences the MOC.

- ▶ K_v is a model parameter which quantifies the intensity of vertical mixing in the ocean, cannot be measured directly.
- ► Two sources of indirect information on K_v:
 - Observations of two ocean "tracers", both provide information about K_v: Carbon-14 (¹⁴C) and Trichlorofluoromethane (CFC11) collected in the 1990s (latitude, longitude, depth), zonally averaged: Z₁, Z₂
 - Climate model output of these two tracers at different values of K_v from the University of Victoria(UVic) Earth System Climate Model (Weaver et. al. 2001): Y₁(K_v), Y₂(K_v)

CFC-11 Example



- ► Bottom right: observations
- ▶ Remaining plots: climate model output at 3 settings of K_{ν} .

Deterministic Models and Emulation

Statistical interpolation



Green inputs/output = training data.

Red = the input where predictions are desired.

Input and output are typically multivariate.

Computer Model Emulation

- ► Fit an emulator ("meta model") to a training set of runs from the complex computer model.
- ► The emulator serves as a surrogate for the computer model, and is much faster/simpler. Hence, it is possible to simulate output from the emulator very quickly.
- Advantages of doing it in a probabilistic framework:
 - Uncertainties associated with interpolation (predictions), for example greater uncertainty where there is less training data information.
 - "Without any quantification of uncertainty, it is easy to dismiss computer models." (A.O'Hagan)
 - Now have a probability model.

Modeling with Gaussian Processes

- Gaussian processes (GPs) are useful models for dependent processes, e.g. time series, spatial data.
- GPs are also very useful for modeling complicated functions.

Key idea: dependence (spatial random effects) adjusts for non-linear relationships between input and output.

Gaussian Process Model Basics

- ▶ Process at location $\mathbf{s} \in D$, $D \subset \mathbb{R}^d$ is $Z(\mathbf{s}) = \mu_{\beta}(\mathbf{s}) + w(\mathbf{s})$. Location \mathbf{s} may be physical or from "input space".
- ▶ Model dependence among spatial random variables by modeling $\{w(\mathbf{s}) : \mathbf{s} \in D\}$ as a Gaussian process.
- ► For any *n* locations, $\mathbf{s}_1, \ldots, \mathbf{s}_n$, $\mathbf{w} = (w(\mathbf{s}_1), \ldots, w(\mathbf{s}_n))^T$ is multivariate normal.
- ► Convenient to specify covariance by a parametric covariance function with parameters Θ . E.g. exponential covariance: $\text{Cov}(Z(\mathbf{s}_i), Z(\mathbf{s}_j)) = \kappa \exp(-\|\mathbf{s}_i \mathbf{s}_j\|/\phi),$ $\kappa > 0, \phi > 0$. Here, $\Theta = (\kappa, \phi)$.
- ▶ Let **Z** = $(Z(\mathbf{s}_1), ..., Z(\mathbf{s}_n))^T$, so

$$\mathbf{Z}|\Theta, \boldsymbol{\beta} \sim N(\mu_{\boldsymbol{\beta}}, \Sigma(\Theta)).$$

GP Linear Model Inference

- ▶ Inference and prediction can be done via ML or Bayes.
- ▶ ML: maximize likelihood with respect to Θ , β .
- ▶ Bayes: prior on Θ , β , and MCMC to learn about $\pi(\Theta, \beta \mid \mathbf{Z})$.

GP Linear Model Prediction

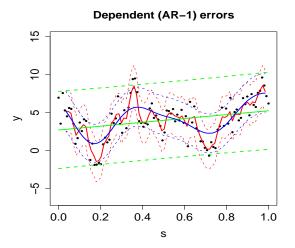
- Let the predictions at the new locations $\mathbf{s}_1^*, \dots, \mathbf{s}_m^* \in D$ be $\mathbf{Z}^* = (Z(\mathbf{s}_1^*), \dots, Z(\mathbf{s}_m^*))^T$.
- ▶ Under the GP assumption $(\mu_1, \mu_2, \Sigma$ depend on β, Θ):

$$\begin{bmatrix} \mathbf{Z} \\ \mathbf{Z}^* \end{bmatrix} \mid \Theta, \beta \sim N \begin{pmatrix} \begin{bmatrix} \boldsymbol{\mu}_1 \\ \boldsymbol{\mu}_2 \end{bmatrix}, \begin{bmatrix} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{bmatrix} \end{pmatrix}, \tag{1}$$

ML: use above with ML estimates plugged-in.

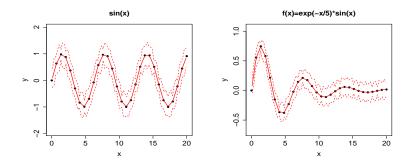
Bayes: use above, while averaging over Θ , $\beta \mid \mathbf{Z}$. This is the *posterior predictive distribution*.

GP Model for Dependence: Toy 1-D Example



Black: 1-D AR-1 process simulation. Green: independent error. (Red, blue): GP with (exponential, gaussian) covariances.

GP for Function Approximation: Toy 1-D Example



The red curves are interpolations using the same, simple GP model with constant mean μ :

 $y(x) = \mu + w(x), \{w(x), x \in (0, 20)\}$ is a zero-mean GP.

Summary of Inferential ("Calibration") Problem

Let parameter of interest be θ (here $\theta = K_v$).

Statistical problem:

- ▶ Model output is a bivariate spatial process at each θ : $\mathbf{Y} = ((\mathbf{Y}_1(\psi_1), \mathbf{Y}_2(\psi_1)), (\mathbf{Y}_1(\psi_2), \mathbf{Y}_2(\psi_2)), \dots, (\mathbf{Y}_1(\psi_K), \mathbf{Y}_2(\psi_K)),$ where $\{\psi_1, \psi_2, \dots, \psi_K\}$ is a set of plausible θ values.
- ▶ Observations: $\mathbf{Z} = (\mathbf{Z}_1, \mathbf{Z}_2)$.
- ▶ What can we learn about θ given **Z**, **Y**?

Bayesian Approach

A Bayesian framework is useful for computer model calibration:

- ▶ There is usually real prior information about θ .
- ▶ The likelihood surface for θ may often be highly multimodal and there may be identifiability issues; useful to have easy access to the full posterior distribution.
- If θ is multivariate, important to look at bivariate and marginal distributions (easier w/ sample-based approach).

Kennedy and O'Hagan (2001); Bayarri, Berger et al. (2007, 2008).

Two-stage Approach to Calibration

- 1. Find probability model for **Z** (data) using **Y** (simulations.)
 - ▶ Model relationship between $\mathbf{Z} = (\mathbf{Z}_1, \mathbf{Z}_2)$ and $\boldsymbol{\theta}$ via emulation of model output $\mathbf{Y} = (\mathbf{Y}_1, \mathbf{Y}_2)$.
 - Z is assumed to be a realization of computer model at "true"
 θ + model-data discrepancy + measurement error.
 - Discrepancy term accounts for "structural uncertainty", i.e., understanding that the the model will have systematic errors and biases. Emulation done via a Gaussian process model. Add a discrepancy term and additional source of error to this model. This provides a probability model for Z in terms of θ.
- 2. Specify a prior for θ and use observations **Z** to infer θ (parameter of interest) based on above probability model.

Inference with Multiple Spatial Fields: Step 1

Need (i) flexible model for relationship between \mathbf{Y}_1 and \mathbf{Y}_2 , (ii) computational tractability.

Model (Y₁, Y₂) as a hierarchical model: Y₁|Y₂ and Y₂ as Gaussian processes (cf. Royle and Berliner, 1999.)

$$\begin{split} \mathbf{Y}_1 \mid \mathbf{Y}_2, \boldsymbol{\beta}_1, \boldsymbol{\xi}_1, \boldsymbol{\gamma} &\sim \textit{N}(\boldsymbol{\mu}_{\boldsymbol{\beta}_1}(\boldsymbol{\theta}) + \mathbf{B}(\boldsymbol{\gamma})\mathbf{Y}_2, \boldsymbol{\Sigma}_{1.2}(\boldsymbol{\xi}_1)) \\ \mathbf{Y}_2 \mid \boldsymbol{\beta}_2, \boldsymbol{\xi}_2 &\sim \textit{N}(\boldsymbol{\mu}_{\boldsymbol{\beta}_2}(\boldsymbol{\theta}), \boldsymbol{\Sigma}_2(\boldsymbol{\xi}_2)) \end{split}$$

- ▶ $\mathbf{B}(\gamma)$ is a matrix relating \mathbf{Y}_1 and \mathbf{Y}_2 , with parameters γ .
- ▶ The covariances of the Gaussian processes depend on both **s** (spatial distance) and θ (distance in parameter space).
- \triangleright β s, ξ s are regression, covariance parameters.

Inference with Multiple Spatial Fields: Step 2

- ► Emulation: Fit GP via maximum likelihood, then obtain predictive distribution at locations of observations.
- ► Add model discrepancy and measurement error:

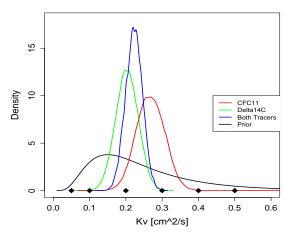
$$\mathbf{Z} = \boldsymbol{\eta}(\mathbf{Y}, \boldsymbol{\theta}) + \boldsymbol{\delta}(\mathbf{Y}) + \boldsymbol{\epsilon}$$

- where $\delta(\mathbf{Y}) = (\delta_1 \ \delta_2)^T$ is the model discrepancy, also modeled as a GP. $\epsilon = (\epsilon_1 \ \epsilon_2)^T$ is the observation error.
- Model discrepancy term can make crucial adjustment to θ estimates (Bayarri, Berger et al. 2007; Bhat et al., 2010).
- ▶ Use Markov chain Monte Carlo (MCMC) to estimate $\pi(\theta \mid \mathbf{Z}, \mathbf{Y})$, integrating out remaining parameters.
- Separating stages: 'modularization' (e.g. Liu, Bayarri, Berger, 2009). Computational advantages + reduce identifiability issues.

Computational Issues

- ► Matrix computations are O(N³), where N is the number of observations. Here: N is in tens of thousands. Computationally intractable without some dimension reduction.
- Need long MCMC runs since there may be multimodality issues, and the chain mixes slowly.
- We use reduced rank approach based on kernel mixing (Higdon, 1998): continuous process created by convolving a discrete white noise process with a kernel function.
- ► Special structure + Sherman-Woodbury-Morrison identity + Sylvester's Theorem used to reduce matrix computations.
- In MLE step: take advantage of structure of hierarchical model to reduce computations.

Results for K_{ν} inference



posteriors: only CFC-11, only $\Delta^{14}C$, both CFC-11 & $\Delta^{14}C$. Result: $\mathbf{K_v}$ pdf suggests weakening of MOC in the future.

Summary

- 1. Our approach is to perform inference in two stages:
 - Obtain a probability model connecting CFC-11, Δ¹⁴C tracer observations to K_v by fitting a Gaussian process model to climate model runs.
 - Using this probability model, infer a posterior density for K_v from the observations.
- We model multivariate spatial data via a flexible hierarchical structure.
- We use kernel mixing to obtain patterned covariances, making computations tractable for large data sets.

We can use inferred K_v in the climate model to project the MOC. We find that the MOC weakens over the next 50 years.

II. Infectious Disease Models

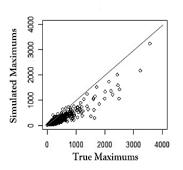
- Gravity-TSIR model: Space-time model for the spread of measles. Unknown parameters of this model control the dynamics of the spread of this disease e.g. how the disease spreads as a function of distance between locations.
- Thousands of latent variables e.g. number of immigrants moving from one location to another.
- ► Rich space-time data set from England and Wales. Time points × locations = 546 × 952 = 519,792.
 Potential for learning about parameters, but also poses computational challenges.

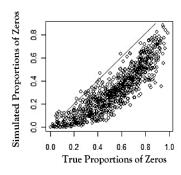
Inference for Gravity TSIR Model Parameters

- ► An approximate grid-based Markov chain Monte Carlo provides a way out of the computational challenges.
- However, traditional likelihood-based/Bayesian inference does not result in a fitted model that reproduces scientifically relevant features of the data.
- Instead, fit GP to summary statistics of model runs where summaries are based on scientifically relevant features.
- Calibration based on using this GP with the data results in improved inference.

Traditional Likelihood-Based Approach

Simulations from fitted model do not match up well with the data for important characteristics of the process.



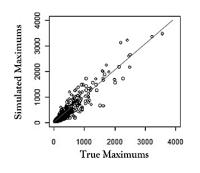


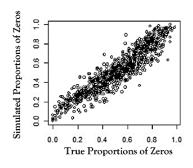
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- Instead, fit GP to summary statistics of model runs where summaries are based on scientifically relevant features.
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GP-based Calibrations Using Key Summaries

Simulations from fitted model are a much better match.





Summary

- Gaussian processes are a powerful tool for problems where the likelihood is implicit and simulating from the model is expensive.
- GPs are useful for deterministic and stochastic models.
- Important to take computational complexity into account and devise strategies to expedite computing.
- When traditional likelihood-based approaches are unsatisfactory because they ignore scientifically important features of the data, GP-based approaches can provide an alternative that directly uses the important features of the data.
- Limitation: computationally intractable when the number of parameters of interest (dimensionality of θ) is large.

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- Grenfell, B.T., Bjørnstad, O. N. and Kappey, J. (2001), "Traveling waves and spatial hierarchies in measles epidemics." *Nature*.
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II. Infectious Disease Models

- Infectious disease models are useful for investigating key questions in biology. They are of practical use in the management and control of infectious diseases, including immunization and epidemic control strategies.
- Here: focus on statistical inference for the Gravity-TSIR model, which models spatiotemporal dynamics. This model presents several inferential and computational challenges.

Simple SIR models

Basic SIR models classify individuals as one of **susceptible** (S), **infected** (I) or **recovered** (R).

- Individuals are born into the susceptible class.
- Susceptible individuals have never come into contact with the disease and are able to catch the disease, after which they move into the infected class.
- Infected individuals spread the disease to susceptibles, and remain in the infected class (the infected period) before moving into the recovered class.
- Individuals in the recovered class are assumed to be immune for life.

Gravity T-SIR model

Extension of the discrete time-series SIR (T-SIR) model (Bjornstad et al.2002; Grenfell et al. 2002) with explicit formulation of the spatial transmission between different host communities.

Notation:

- ▶ $I_{k,t}$ number of **infected** individuals in city k at time t.
- ▶ $S_{k,t}$ number of **susceptible** individuals in city k at time t.
- $d_{k,i}$ **distance** between cities k and j.
- ▶ $N_{k,t}$ **population** of city k at time t.
- ▶ $B_{k,t}$ local number of new hosts (**births**) in city k at time t.
- L_{k,t} number of infected people moved (**immigrants**) to city k at time t.
- ▶ *T* cities, *K* time points.

Modeling incidences

Following Xia, Bjornstad and Grenfell (2004):

▶ Number of incidences of a disease at time t + 1 for city k,

$$I_{k,t+1} = \mathsf{Poisson}(\lambda_{k,t+1}), \text{ where } \lambda_{k,t+1} = \beta_t \mathcal{S}_{k,t} (I_{k,t} + L_{k,t})^{\alpha}.$$

• α , $\{\beta_t\}$ are local transmission parameters.

Modeling susceptibles

Number of susceptible individuals at time t + 1 for city k is then modeled via balance equation (Bartlett, 1957):

$$S_{k,t+1} = S_{k,t} + B_{k,t} - I_{k,t+1}$$

► Finally, unobserved number of infected immigrants moved to city *k* at time *t* is modeled as:

$$L_{k,t} = \text{Gamma}(m_{k,t}, 1),$$

where

$$m_{k,t} = \theta N_{k,t}^{\tau_1} \sum_{j=1, j \neq k}^{K} \frac{(I_{jt})^{\tau_2}}{d_{k,j}^{\rho}}, \quad \theta, \tau_1, \tau_2, \rho > 0.$$

Statistical inference for measles

Measles data

- The UK Registrar General's data for 952 cities in England and Wales for years 1944-1966 of biweekly incidences of measles. Very rich spatio-temporal data.
- Data for number of susceptibles from standard susceptible reconstruction algorithms (cf. Fine and Clarkson, 1982)

Parameters of the model:

- ▶ Reliable estimates of local transmission parameters α and $\{\beta_t\}$ are assumed known from previous work (Bjornstad et al. 2001).
- ▶ **Goal**: Infer unknown gravity parameters: θ , τ_1 , τ_2 , ρ .

Challenges with likelihood-based inference

- ▶ Dimensions of the data (*TK*): 546*952 = 519,792.
- ▶ Number of infected immigrants $\{L_{k,t}\}$ are unobserved.
- ► The likelihood function is complicated:
 - ▶ Involves integrating over 519,792 latent variables.
 - Very expensive calculations per iteration.
- Approximate Bayesian computation (ABC) approaches are infeasible since simulating draws from this model is computationally expensive.

A simplified model and gridded MCMC

Simplify the model by fixing the number of immigrants (latent variables) at their means.

- Likelihood evaluations are still very expensive.
- Studying likelihood surface, learning about variability of estimates is computationally infeasible.

Gridded Metropolis-Hastings:

- We evaluate expensive parts of the likelihood on a grid of parameter values (can use parallel processors for this) and store these in a look-up table.
- M-H algorithm on discretized parameter space (on grid).
 M-H ratio evaluation is now much faster.

Results

- ► The gridded MCMC algorithm produces posterior distributions similar to a non-gridded MCMC algorithm, but much faster.
- Conclusions based on a simulation study:
 - Serious identifiability issues. Can only infer 2 of the 4 parameters.
 - In simulation studies: posterior (and likelihood) surface is peaked away from the true parameter values. There's a significant shift (bias) in parameter estimates.

Alternative approach

- Instead of likelihood-based approach, focus on important biological 'signatures' of the process. E.g. proportion of zeros (# of times no disease incidences in a city).
- ► Borrow ideas from computer model emulation, calibration (cf. Sacks et al., 1989.)
 - Simulate realizations from the gravity model at different parameter values.
 - 2. Use the signatures to define summary statistics.
 - 3. Find distance between summary statistics for the simulated process and the observations.
 - 4. Fit a Gaussian process to this distance, as a function of the parameters.
 - Can obtain a likelihood and perform Bayesian inference for the gravity model parameters using the observations.

Inferential approach outline

- Gravity parameters, $\Theta = (\theta, \tau_1, \tau_2, \rho)$.
- ► Summary statistics (distance to observations) based on simulations at Θ_i , i = 1, ..., n parameter settings, $\mathbf{Y} = (\mathbf{Y}(\Theta_1), ..., \mathbf{Y}(\Theta_n))$.
- Model stochastic model output **Y** using a Gaussian process: **Y** | β , $\xi \sim N(\mu_{\beta}(\Theta), \Sigma(\xi, \Theta))$. Infer β , ξ : regression, covariance parameters.
- ► Model summary statistic for real data set **Z**:
- ▶ **Z** = η (**Y**, θ) + δ _Ψ(**Y**, Θ) + ϵ _{σ ²}(**Y**) where η is a random variable with predictive distribution derived above. δ is a discrepancy function, modeled as Gaussian process, and ϵ is a vector of i.i.d. errors.
- ▶ Infer posterior $\pi(\Theta, \Psi, \sigma^2 \mid \mathbf{Z}, \mathbf{Y})$ using MCMC.

Conclusions

- Our GP-based emulation approach appears to produce unbiased estimates of the parameters.
- With estimated parameters, the model is able to reproduce well the signatures of the disease process.
- This is the first statistically rigorous approach to this problem: estimates of uncertainty, joint distributions of parameters, predictions/variability from fitted model.

Caveats and future work:

- Our statistical approach unearths serious identifiability issues: can still only learn about 2 parameters at most.
- Computational concerns only allow for a limited number of model forward runs.

Key references

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