Toward Approximating and Calibrating a Computer Model with Spatial Output

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This Talk

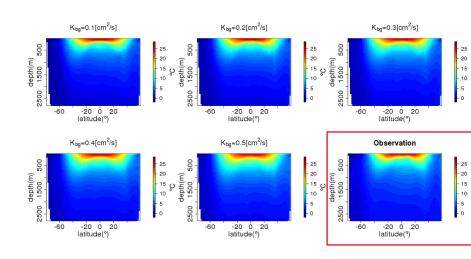
- Complex computer models are often used to make projections about future climate.
- A major source of uncertainty about these projections is due to uncertainty about climate model input parameters.
- We propose a method for learning about climate model parameters from climate model outputs and observations.
- ► Challenges: Data in the form of high-dimensional spatial fields. Complicated error structures. Non-Gaussian.
- I will describe novel computationally efficient approaches based on principal components (PC) and kernel convolution

Complex Computer Models for Climate

- Model for studying the Atlantic Meridional Overturning Circulation (AMOC): Important for maintaining equilibrium climate in Europe
- Models for studying the future of the Greenland and West Antarctic Ice Sheets

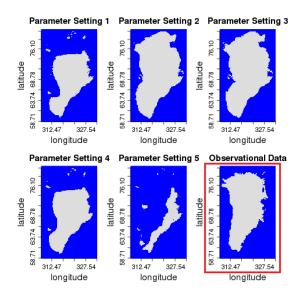
Calibration Problem

Uncertainty due to parameter K_{bg} Which parameter settings best match observations?



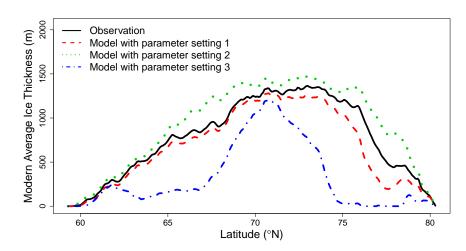
Calibration with Spatial Output

Which parameter settings best match observations?



Aggregation Approach

Which parameter settings best match *aggregated* observations?



Emulation-Calibration

- 1. Models may take a long time to run at each setting.
- 2. High-dimensional data
- 3. Potentially non-Gauassian.

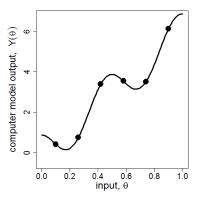
Two-stage Approach to Emulation-Calibration

- 1. Emulation step: Find fast approximation for climate model using Gaussian process (GP)
- Calibration step: Infer climate parameter using emulator and observations, while accounting for data-model discrepancy

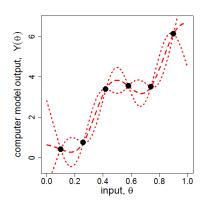
(Bhat, Haran, Olson, Keller, 2012; Liu, Bayarri and Berger, 2009)

Emulation Step

Toy example: pretend model output is a scalar



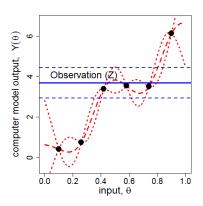
Computer model output (y-axis) vs. input (x-axis)



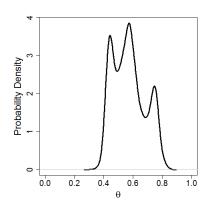
Emulation (approximation) of computer model using GP

Calibration Step

Toy example: pretend model output and observations are scalars



Combining observation and emulator



Posterior PDF of θ given model output and observation

Summary of Statistical Problem

- Goal: Learning about θ based on two sources of information:
 - ▶ **Observations***: Mean potential ocean temperature†, $\mathbf{Z} = (Z(\mathbf{s}_1), \dots, Z(\mathbf{s}_n))^T$, where $\mathbf{s}_1, \dots, \mathbf{s}_n$ are 3D locations.
 - ▶ Climate model output** for mean potential temperature $\mathbf{Y}(\theta_1), \dots, \mathbf{Y}(\theta_p)$, where each $\mathbf{Y}(\theta_i) = (Y(\mathbf{s}_1, \theta_i), \dots, Y(\mathbf{s}_n, \theta_i))^T$ is spatial field (Sriver et al., 2012).

Z and $\mathbf{Y}(\theta_i)$'s are *n*-dimensional vectors

For MOC example: output at each θ_i is a high-dimensional spatial field. n = 61,051 locations, p = 250 runs.

*World Ocean Atlas 2009

**University of Victoria (UVic) Earth System Climate Model †Averaged over 1955-2006

GP for Computer Model Emulation

- ► Fit GP to *np*-dimensional data $\mathbf{Y} = (\mathbf{Y}(\theta_1)^T, \dots, \mathbf{Y}(\theta_p)^T)^T$ for interpolation.
- Covariance used for
 - non-linear relationship between parameter and model output (model output as a function of parameter)
 - non-linear spatial surface (model output as a function of location)
- Covariance function example:

$$\begin{aligned} \mathsf{Cov}\left(\mathsf{Y}(\mathbf{s}, \boldsymbol{\theta}), \mathsf{Y}(\mathbf{s}', \boldsymbol{\theta}'); \boldsymbol{\xi}\right) = & \kappa \exp\left(-\frac{g\left(\mathbf{s}, \mathbf{s}'\right)}{\phi_{\mathbf{s}}}\right) \exp\left(-\frac{\left\|\boldsymbol{\theta} - \boldsymbol{\theta}'\right\|}{\phi_{\boldsymbol{\theta}}}\right) \\ & + \zeta \mathit{I}(\boldsymbol{\theta} = \boldsymbol{\theta}') \mathit{I}(\mathbf{s} = \mathbf{s}') \end{aligned}$$

where g is geodesic distance, and $\xi = (\kappa, \phi_s, \phi_\theta, \zeta)$ is covariance parameter.

Step 1: Emulation (Approximating Computer Model)

- Find MLE for covariance parameter ξ , denoted by $\hat{\xi}$
- ▶ Get $\eta(\theta_{NEW}, \mathbf{Y})$ for prediction at any $\theta_{NEW} \in \Theta$:
 - GP gives

$$\left(\begin{array}{c} \mathbf{Y} \\ \mathbf{Y}(\boldsymbol{\theta}_{NEW}) \end{array}\right) \sim N \left(\left(\begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array}\right)_{n(p+1)\times 1}, \left(\begin{array}{cc} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{array}\right)_{n(p+1)\times n(p+1)}\right)$$

Emulator:

$$\boldsymbol{\eta}\left(\boldsymbol{\theta}_{\textit{NEW}}, \boldsymbol{Y}\right) = \boldsymbol{Y}(\boldsymbol{\theta}_{\textit{NEW}}) | \boldsymbol{Y} \sim \mathcal{N}\left(\boldsymbol{\Sigma}_{21} \boldsymbol{\Sigma}_{11}^{-1} \boldsymbol{Y}, \boldsymbol{\Sigma}_{22} - \boldsymbol{\Sigma}_{21} \boldsymbol{\Sigma}_{11}^{-1} \boldsymbol{\Sigma}_{12}\right)$$

Step 2: Calibration (Inferring Input Parameter)

Probability model for Z based on

$$\mathbf{Z} = \boldsymbol{\eta}(\boldsymbol{\theta}, \mathbf{Y}) + \boldsymbol{\delta},$$

where *n*-dimensional spatial field δ is model-observation discrepancy with covariance parameter ξ_{δ} .

▶ Inference for θ based on posterior distribution

$$\pi(\boldsymbol{\theta}, \boldsymbol{\xi}_{\delta} | \mathbf{Z}, \mathbf{Y}, \hat{\boldsymbol{\xi}}) \propto \underbrace{L(\mathbf{Z} | \mathbf{Y}, \boldsymbol{\theta}, \boldsymbol{\xi}_{\delta}, \hat{\boldsymbol{\xi}})}_{\text{likelihood given by above}} \times \underbrace{\rho(\boldsymbol{\theta}) \times \rho(\boldsymbol{\xi}_{\delta})}_{\text{priors for } \boldsymbol{\theta} \text{ and } \boldsymbol{\xi}_{\delta}}$$

with emulator parameter $\hat{\xi}$ fixed at value estimated in emulation step.

Computational challenge: $np \times np$ covariance matrix of **Y** (reminder: n = 61,051 p = 250)

Main Idea: reduced-dimensional PCAs

► Consider model outputs at $\theta_1, \dots, \theta_p$ as if they were replicates of a multivariate process, thereby obtaining their PCs

$$\begin{pmatrix} Y(\mathbf{s}_{1}, \theta_{1}) & \dots & Y(\mathbf{s}_{n}, \theta_{1}) \\ \vdots & \ddots & \vdots \\ Y(\mathbf{s}_{1}, \theta_{p}) & \dots & Y(\mathbf{s}_{n}, \theta_{p}) \end{pmatrix}_{p \times n} \Rightarrow \begin{pmatrix} Y_{1}^{R}(\theta_{1}) & \dots & Y_{J_{y}}^{R}(\theta_{1}) \\ \vdots & \ddots & \vdots \\ Y_{1}^{R}(\theta_{p}) & \dots & Y_{J_{y}}^{R}(\theta_{p}) \end{pmatrix}_{p \times J_{y}}$$

▶ PCs pick up characteristics of model output that vary most across input parameters $\theta_1, \dots, \theta_p$.

Emulation Using PCs

- Fit 1-dimensional GP for each series $Y_j^R(\theta_1), \dots, Y_j^R(\theta_p)$
- ▶ $\eta(\theta, \mathbf{Y}^R)$: J_y -dimensional emulation process for PCs, \mathbf{Y}^R is collection of PCs
- ► Computation reduces from $\mathcal{O}(n^3p^3)$ to $\mathcal{O}(J_yp^3)$ (1.2 × 10²¹ to 1.0 × 10⁸ flops).
- ► Emulation for original output: compute $\mathbf{K}_y \eta(\theta, \mathbf{Y}^R)$ where \mathbf{K}_y is matrix of scaled eignvectors

Calibration in Reduced Dimensions

Probability model for dimension-reduced observation Z^R:

$$\begin{split} \mathbf{Z} &= \underbrace{\mathbf{K}_{\mathcal{Y}} \boldsymbol{\eta}(\boldsymbol{\theta}, \mathbf{Y}^R)}_{\text{emulator}} + \underbrace{\mathbf{K}_{\mathcal{d}} \boldsymbol{\nu}}_{\text{discrepancy}} + \underbrace{\boldsymbol{\epsilon}}_{\text{observation error}}, \\ \Rightarrow & \mathbf{Z}^R = (\mathbf{K}^T \mathbf{K})^{-1} \mathbf{K}^T \mathbf{Z} = \left(\begin{array}{c} \boldsymbol{\eta}(\boldsymbol{\theta}, \mathbf{Y}^R) \\ \boldsymbol{\nu} \end{array} \right) + (\mathbf{K}^T \mathbf{K})^{-1} \mathbf{K}^T \boldsymbol{\epsilon}, \end{split}$$

with combined basis $[\mathbf{K}_y \ \mathbf{K}_d]$, knot process $\nu \sim N(\mathbf{0}, \kappa_d \mathbf{I})$, and observational error $\epsilon \sim N(\mathbf{0}, \sigma^2 \mathbf{I})$.

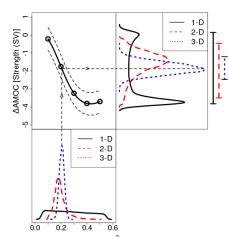
▶ Infer θ through posterior distribution

$$\pi(\boldsymbol{\theta}, \kappa_{\boldsymbol{d}}, \sigma^2 | \mathbf{Z}^R, \mathbf{Y}^R) \propto \underbrace{L(\mathbf{Z}^R | \mathbf{Y}^R, \boldsymbol{\theta}, \kappa_{\boldsymbol{d}}, \sigma^2)}_{\text{likelihood given by above}} \underbrace{p(\boldsymbol{\theta}) p(\kappa_{\boldsymbol{d}}) p(\sigma^2)}_{\text{priors}}$$

Results

Computational efficiency allows us to calibrate using unaggregated data. E.g. Atlantic Meriodonoal Overturning Circulation (AMOC).

- We compare 1D (depth profile) and 2D (zonal average) with 3D (unaggregated) data.
- Inference with 3D data leads to sharper inference for θ.
- Inference using 3D data is more robust to changes in prior specifications for discrepancy parameters.



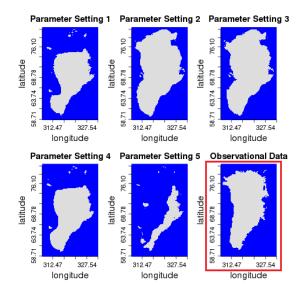
Discussion and Ongoing Work

Dimension reduction-based approach:

- ▶ Very fast, scales well with *n*, number of spatial locations
- Very easy to use: Automatic emulation step
- Works for a number of other multivariate settings, e.g. time series, multiple time series, multiple spatial output
- We can deal with aggregated spatial binary data using our methods.
- Calibration with non-Gaussian spatial data, e.g. 0-1 (ice-no ice) involves developing a reduced-dimensional approach, building upon a logistic PCA (Lee et al., 2010)

New Challenge: Calibration with Spatial Binary Output

Again, which output best matches the observations?

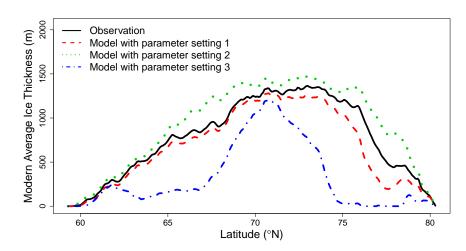


Calibration with Binary Output

- Standard Gaussian process approach does not apply
- Our reduced-dimensional approach also does not apply
- Some options:
 - Aggregation/averaging to obtain "more Gaussian" output, then apply our methods
 - New approach that applies to binary output. Challenging: naive application of spatial generalized mixed model to such data is infeasible

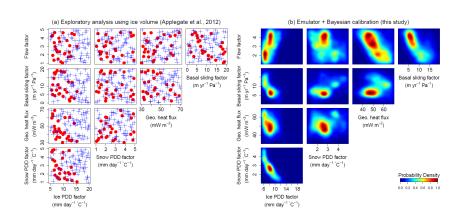
Aggregation Approach

Which parameter settings best match *aggregated* observations?

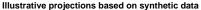


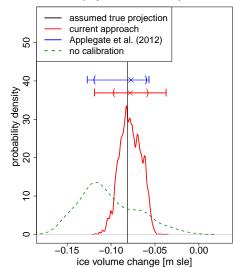
How Does Statistical Rigour Help?

Left: sensible but non-rigourous vs Right: sound statistics "Underneath the hood": (i) accounting for (epistemic) uncertainties in emulation, (ii) real probability distributions.



Ice Volume Change Projection





Ongoing Work

- Would like to use the original binary data. Hence, reduced-dimensional calibration for binary spatial data.
- Computational issues are even more delicate because a naive latent variable approach would result in severe computational issues
 - 1. Use binary analogue to regular PCAs.
 - 2. Discrepancy modeling is tricky...

Acknowledgments

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- Won Chang, University of Chicago
- David Pollard, Earth and Environmental Systems Institute (EESI), Penn State U.
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- Klaus Keller, Geosciences, Penn State U.
- Roman Olson, The University of New South Wales

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- ▶ NSF CDSE/DMS-1418090

Relevant Manuscripts

- Chang, W., M. Haran, R. Olson, and K. Keller (2014): Fast dimension-reduced climate model calibration, *Annals of Applied Statistics*
- Chang, W., Applegate, P., Haran, M. and Keller, K. (2014) Probabilistic calibration of a Greenland Ice Sheet model using spatially-resolved synthetic observations: toward projections of ice mass loss with uncertainties, Geoscientific Model Development

Appendix: Cross-Validation for Emulator

Example of leave-10%-out cross validation result:

