Development and External Validation of a Multi-State Clinical Prediction Model for Chronic Kidney Disease Patients Progressing onto Renal Replacement Therapy and Death

# Introduction

Chronic Kidney Disease (CKD) is a progressive disease that affects the ability of the kidneys to filter toxins from the blood. Patients with End-stage Renal Disease (ESRD) are treated using Renal Replacement Therapy (RRT), which collectively describes the treatments designed to emulate the processes performed by the failing kidneys. The three most common treatment modalities are haemodialysis (HD), peritoneal dialysis (PD) and kidney transplant (Tx). The more severe stages of CKD (stages 3 - 5) affects approximately 2.6 million people over the age of 16 in England [1] with around 63 thousand adult patients registered for RRT in 2015 [2], of which 8 thousand were new patients [3].

A prognostic model is a tool which provides patients and clinicians with a measure of how likely a patient is to suffer a specific clinical event in the future [4]. They use data from previous patiets to estimate the outcomes of an individual patient. Prognostic models are used in clinical practice to influence treatment decisions such as the prescribing of statins for cardiovascular disease via the application of the QRISK models [5].

Within CKD, prognostic models have been developed to predict mortality [6]–[10], ESRD [7], the commencements of RRT [9], [11]–[13] or mortality after beginning dialysis [14]–[16]. Some previous models have used the commencement of RRT as a proxy for ESRD [17]–[19], while others have investigated the occurrence of cardiovascular events, which are common amongst CKD patients[20]–[22]. Reviews by Grams & Coresh [23], Tangri et al [24] and Ramspek et al [25], which explored the different aspects of assessing risk amongst CKD or RRT patients, found that the current landscape of CKD prediction models is lacking from both a methodological and clinical perspective [26], [27].

Methodologically, the majority of existing CKD prediction models fail to account for completing events [8], [10], [28], have high risks of bias [6], [7], [11] or are otherwise flawed compared to modern clinical prediction standards [4], [26]

In 2013, Begun et al [29] developed a multi-State model for assessing population-level progression through the severity stages of CKD (III-V), RRT and/or death, which can be used to provide a broad statement regarding a patient’s future. In 2014, Allen et al [30] applied a similar model to liver transplant recipients and their progression through the stages of CKD with a focus on the predictions of measured vs estimated glomerular filtration rate (mGFR vs eGFR). In 2017, Kulkarni et al [13] developed an MSM focusing on the categories of Calculated Panel Reactive Antibodies and kidney transplant and/or death.

Most recently, in 2018, Grams et al [31] developed a multinomial clinical prediction model for CKD patients which focused on the occurrence of RRT and/or cardiovascular events. As of the publication of this paper, this is the only currently existing CPMs of this kind for CKD patients.

However, the first three of these existing models (Begun, Allen and Kulkarni) categorise continuous variables to define their states at specific cut-offs and this has been shown to be inefficient when modelling [32]–[50]. These kinds of cut-offs can be useful when informing patients and clinicians of a patient’s diagnosis and to coincide with policy, but inherently cause a loss information when done before the data analysis stage and so these models go against the current statistical recommendations [**Cite: Something**]. These kinds of assumptions are also subject to measurement error and interval censoring [**Cite: Something**], i.e. we do not know when exactly when a patient moved from CKD Stage III to CKD Stage IV, or whether drop in estimated Glomerular Function Rate (eGFR) was temporary or inaccurate. For example, Kulkarni [13] assumes that a patient with an CPRA of (5%) is the same as a patient with an CPRA of (75%) and that a patient with an CPRA of (89.9%) is vastly different from a patient with an CPRA of (90%). Moreover, none of these models have undergone any validation process, whether internal or external [51].

It is also important to note that although these models can be used to predict patient outcomes, these models were not designed to produce individualised patient predictions as is a key aspect of a clinical prediction model; they were designed to assess the methodological advantages of MSMs in this medical field, to describe the prevalence of over time of different CKD stages and to produce population level predictions for patients with different levels of panel-reactive antibodies [52].

The fourth model (Grams et al [23]), is presented as a Multi-State Model and the transitions involved were studied and defined, however the underlying statistical model is a pair of multinomial logistic models analysed at 2 and 4 years. The major downside of this model is that it can only produce predictions at those predefined time points and it assumes homogeneity of transition times. For example, the first model assumes that a patient who began RRT 1 month after study entry is the same as one who began after 1 year & 11 months into the study and then the second model assumes these patients are the same as one who begins RRT at 3 years and 11 months.

Therefore, the aim of this study was to improve on previous efforts to model patient’s pathways through a Multi-State Model by choosing transition points which can be exactly identified and include states which produce a drastic difference in patient characteristics. Our modeling techniques allow for individual predictions (using a proportional hazards model) of multiple outcomes (using MSMs) at any time point (using cubic splines). The models produced by this process will then be validated, both internally and externally, to compare their results and demonstrate the transportability of the (statistically robust) clinical prediction models. We report our work in line with the TRIPOD guidelines for development and validation of clinical prediction models [26], [53].

# Methods

## Data Sources

The models were developed using data from the Salford Kidney Study (SKS) cohort of patients (previously named the CRISIS cohort), established in the Department of Renal Medicine, Salford Royal NHS Foundation Trust (SRFT). The SKS is a large longitudinal CKD cohort recruiting CKD patients since 2002. This cohort collects detailed annualised phenotypic and laboratory data, and plasma, serum and whole blood stored at -80C for biomarker and genotypic analyses. Recruitment of patients into SKS has been described in multiple previous studies [54], [55] and these have included a CKD progression prognostic factor study and to evidence the increased risk of cardiovascular events in diabetic kidney patients. In brief, any patient referred to Salford renal service (catchment population 1.5 million) who is 18 years or over and has an eGFR measurement of less than (calculated using the CKD-EPI formula [56]) was approached to be consented for the study participation.

At baseline, the data, including demographics, comorbidities, physical parameters, lab results and primary renal diagnosis are recorded in the database. Patients undergo an annual study visit and any changes to these parameters are captured. All data except blood results are collected via questionnaire by a dedicated team of research nurses. Blood results (baseline and annualised), first RRT modality and mortality outcome data are directly transferred to the database from Salford’s Integrated Record (SIR) [57]. eGFR, uPCR, comorbidity and blood results were measured longitudinally throughout patients time within the cohort.

Due to limitations in our data, we were agnostic to how long since patients were diagnosed with CKD. Therefore, we defined a patient’s start date for our model as their first date after consent at which their eGFR was recorded to be below . Some patients consented with an eGFR that was already below 60, and some entered our study later when their eGFR was measured to be below 60. This implies that our models includes both patient who have recently been diagnosed with CKD () *and* those that have been suffering with CKD for an arbitrary amount of time. This timelessness of the model means it can be applied to any patient at any time during their CKD journey.

This allows for a wider range of baseline eGFR measurements and patients who have been suffering from CKD, translating to a model which can be applied to

All patients registered in the database between October 2002 and December 2016 with available data were included in this study. As this is a retrospective convenience sample, no sample size calculations were performed prior to recruitment. All patients were followed-up within SKS until the end-points of RRT, death or loss to follow-up or were censored at their last interaction with the healthcare system prior to December 2017. Date of death for patients who commenced RRT was also available within SIR and so also included in the SKS database.

For external validation of the model, we extracted an independent cohort from the West of Scotland Electronic Renal Patient Record (SERPR). Our extract of SERPR contains all patients known to the Glasgow and Forth Valley renal service who had an eGFR measure of less than between January 2006 and January 2016. This cohort has been previously used in Chronic Kidney Disease Prognosis consortium studies investigating outcomes in patients with CKD [58] and a similar cohort has been used for the analysis of skin tumours amongst renal transplant patients. Use of anonymised data from this database has been approved by the West of Scotland Ethics Committee for use of NHS Greater Glasgow and Clyde ‘Safe Haven’ data for research.

Both the internal and external validation cohort were used as part of the multinational validation cohort used by Grams et al in their multinomial CPM discussed above [31]. In SERPR, start dates were calculated to be the firt time point where the following conditions were met:

* eGFR is measured at less than 60
* There is at least one prior eGFR measurement
* Patient is 18 or over
* Patient is not enduring an AKI [59], [60].

The second requirement was implemented to avoid a bias in the eGFR Rate. eGFR Rate is a measure of the change in eGFR over time and is calculated as the difference between the most recent two eGFR measurements divided by the time between them. For patients who entered the system with an eGFR < 60, their eGFR Rate would be unavailable (i.e. missing). Otherwise, patient eGFRs would *have* to drop to below 60 and thus eGFR Rate would be negative. [**I feel like this bias should have a name, but can’t think what to search to find it**]

## Development

Three separate models were developed, so we could determine a clinically viable model while maintaining model parsimony as much as possible: a Two-State, Three-State and Five-State model, each building on the previous models’ complexity (see [**Figure 1**]). The Two-State model was a traditional survival analysis where a single event (death) is considered. The Three-State model expanded on this, by splitting the Alive state into transient states of (untreated) CKD and (first) RRT; patients can therefore transition from CKD to Death or CKD to RRT, and then onto RRT to Death. The Five-State model stratifies the RRT state into HD, PD and Tx and allows similar transitions into and out of the RRT states; however, the transition from Tx to Death was not considered as it was anticipated a priori that there would be insufficient patients undergoing this transition and that the process of undergoing a transplant would be medically transformative and so it would be inappropriate to assume shared parameters before and after the transition (i.e. Tx was modelled as a second absorbing state).

Data were recorded in a time-updated manner, however all variables were measured at baseline to emulate the real-world application of the model (i.e. future prediction of states and not covariates). Variables considered as covariates were demographics (sex, age, smoking status and alcohol consumption), comorbidities (congestive cardiac failure (CCF), chronic obstructive pulmonary disease (COPD), prior cerebrovascular accident (CVA), hypertension (HT), diabetes mellitus (DM), ischemic heart disease (IHD), chronic liver disease (LD), prior myocardial infarction (MI), peripheral vascular disease (PVD) and slid tumour (ST)), physical parameters (BMI, blood pressure), blood results (haemoglobin, albumin, calcium and phosphate measures), urine protein creatinine ratio (uPCR) and primary renal diagnosis (grouped as per ERA-EDTA classifications [61]). Ethnicity was assessed in the populations, but as most patients were white, it was omitted as a potential predictor from the models.

uPCR and eGFR Rate of change were also calculated [62], [63] as the difference between the two most recent measures divided by time difference in days. Age^2, log(Age), log(eGFR Rate) and log(uPCR Rate) were considered as transformations within the model. log(-Calendar Time) were included as a covariates (cite). Calendar Time was defined as time of study entry minus 1st January 2019, ensuring CalendarTime is always negative so that patients who entered the study longer ago have a higher value for this covariate.

Intermediate states (RRT or modality) were considered to be medically transformative, and so a semi-markov (clock reset) method for analysis was considered to be well justified [64]. Each transition was modelled under a proportional hazards assumption using the Royston-Parmar technique [65] to estimate coefficients for each covariate and a restricted cubic spline (on the log-time scale) for the baseline cumulative hazard. The cumulative hazards for each transition can be combined to produce estimates for the probability of a patient being in any state at any time [66].

For variable selection, we stacked the imputed datasets together to create a larger, pseudo-population [67] and performed backwards-forwards selection based on minimising the AIC at each step ([**Cite: Something**]). This was repeated for each transition and for different numbers of evenly spaced knots in modelling the form of the baseline hazard, K={0,1,2,3,4,5}. This allowed for different transitions to use different sets of variables and numbers of knots in the final model. Some combinations of variables resulted in models that were intractable and so these models were excluded. Once a set of variables were chosen, the R-P model was applied to each imputed dataset individually and the resulting coefficients and cubic spline parameters were aggregated across imputations using Rubin’s Rules [68]. This gave a model fully defined by smooth cubic splines representing the cumulative cause-specific hazard and individualised proportional hazards for each transition.

All missing data were assumed to be missing at random and so were multiply imputed using chained equations with the Nelson-Aalen estimators for each relevant transition as predictors [69]. Some variables (smoking status and histories of COPD, LD and ST) were present in the SKS (development) dataset, but were completely missing in the SERPR extract (validation) and so these were multiply imputed from the development dataset [70].

All analysis was done in R 3.6.2 [71] using the various tidyverse packages [72], as well as the mice [73], flexsurv [74], nnet [75] and furrr [76] packages.

## Example

Once the models have been developed, we will apply them to two example patients to demonstrate their use and applicability to the general population. We will provide a direct clinical estimation of these patient outcomes based on years of nephrological experience and compare this with the results presented by our clinical prediction model.

We have chosen two (artificial) patients to use as examples of the use of our model. Their details can be seen below:

[**Details of example patients data in a two column table**]

The main difference between our two patients are [**blah blah**].

## Validation

Further to the simple individual validation performed above. A statistical validation will need to be performed to formally establish the performance of the models.

Each of the three models were internally validated in the development dataset using bootstrapping to adjust for optimism and then further externally validated in the validation dataset extracted from SERPR. The bootstrapping method was also used for both validations to produce confidence intervals around the performance metric estimates. To assess the performance in low eGFR patients, the models were also validated in subsets of the SKS and SERPR where patients had an .

For validation purposes, we consider Death and Death after RRT/HD/PD to be distinct states meaning that for the Three-State model, we have pathways a patient can take and for the Five-State model, we have . To compare across models, we combined states together to collapse down to simpler versions. We collapsed the Three-State model to a two-state structure by combining the CKD and RRT states into an Alive state. We collapsed the Five-State model to a three-state structure by combining the HD, PD and Tx into an RRT state and then further down to a two-state structure as with the Three-State model. We will report performance measures at 1-year, 2-years, 5-years and 9-years. As well as presenting the performance measures over time.

The overall accuracy of each model was assessed using the MSM adjusted Brier Score [[**Cite: Performance Metrics**]], which is a proper score function assigning 0 to a non-informative model and 1 to a perfect model, with negative numbers implying the model performs worse than assuming every patient’s state predictions are the same as the overall prevalence within the population.

The discrimination of each model was assessed using the MSM extension to the c-statistic [77] [**Cite: Performance Metrics**]. The c-statistic is a score between 0 and 1 with higher scores suggesting a better model and a c-statistic of 0.5 suggesting the model performs no better than a non-informative model.

The calibration of each model was assessed using MSM multinomial logistic regression (MLR) [78] [**Cite: Performance Metrics**] which extends the logistic regression to three or more mutually exclusive outcomes [79]. This produces an intercept vector of length and a Slope-matrix of dimension . As with the traditional calibration intercept for a well performing model, the MLR intercept values should all be as close to 0 as possible. The traditional calibration slope should be as close to 1 as possible and so the multi-state extension of the slope, the Slope-matrix should be as close to the identity matrix () as possible.

# Results

## Data Sources

As seen in table 1, The Age of both populations were centred around 64-65 with a very broad range. Due to the inclusion criteria, eGFR were capped at a maximum of 60, and was consistent across populations; however, the rate of change for eGFR was much wider in the SERPR patients than in the SKS, and it was decreasing much faster, on average ( -25 vs 0) . Blood pressure was also consistent across populations (140/75 vs 148/76 for development vs validation). The blood test results (Calcium, Albumin, Haemoglobin and Phosphate) was close together, with the further difference being Haemoglobin with an average of 123 in SKS and 109 in SERPR and a much larger standard deviation in SERPR compared to SKS (38 vs 17). Similar to the eGFR measures, the uPCR results were similar, but the rates of change were much broader in the validation dataset compared to the SKS and were generally increasing, whereas SKS remained stationary (73 vs 0). Levels of missingness were much higher in the SERPR dataset in most continuous variables.

| **Variable** | **SKS (Development)** | **SERPR (Validation)** |
| --- | --- | --- |
| Age | 64.378 (14.573) [ 20.000, 94.000] < 0 ( 0.00%)> | 65.880 ( 13.734) [ 18.000, 98.000] < 0 ( 0.00%)> |
| eGFRa | 30.368 (14.122) [ 3.577, 59.965] < 0 ( 0.00%)> | 36.132 ( 13.668) [ 1.651, 59.998] < 0 ( 0.00%)> |
| eGFR.Ratea | -0.015 ( 1.528) [-19.107, 33.781] <1,278 (42.87%)> | -25.476 (623.048) [-8,755.272, 9,260.375] < 0 ( 0.00%)> |
| SBPb | 140.193 (21.839) [ 77.000, 220.000] < 50 ( 1.67%)> | 147.746 ( 25.984) [ 82.000, 258.000] <6,880 (88.61%)> |
| DBPb | 74.555 (11.754) [ 36.000, 159.000] < 52 ( 1.74%)> | 76.263 ( 13.853) [ 35.000, 128.000] <6,879 (88.60%)> |
| BMIc | 28.848 ( 6.114) [ 13.182, 61.466] < 572 (19.18%)> | 29.331 ( 6.722) [ 15.343, 48.301] <7,681 (98.93%)> |
| Albumind | 42.152 ( 4.186) [ 12.000, 52.000] < 60 ( 2.01%)> | 36.490 ( 5.647) [ 7.000, 53.000] <3,455 (44.50%)> |
| Calciume | 2.302 ( 0.147) [ 1.209, 3.660] < 68 ( 2.28%)> | 2.408 ( 0.158) [ 1.419, 3.610] <5,113 (65.85%)> |
| Haemoglobind | 122.977 (17.068) [ 61.000, 195.000] < 72 ( 2.41%)> | 108.588 ( 38.446) [ 6.250, 208.000] <3,968 (51.10%)> |
| Phosphatee | 1.162 ( 0.284) [ 0.430, 3.710] < 87 ( 2.91%)> | 1.203 ( 0.311) [ 0.370, 4.370] <5,127 (66.03%)> |
| uPCRf | 0.112 ( 0.195) [ 0.000, 2.025] < 245 ( 8.21%)> | 0.184 ( 0.477) [ 0.000, 6.390] <7,513 (96.76%)> |
| uPCR.Ratef | -0.096 ( 3.035) [-70.727, 28.198] <1,777 (59.61%)> | 73.177 (465.186) [ -2.255, 3,051.403] <7,721 (99.44%)> |
| a(ml/min/1.73m^2) or per year; b(mmHG); c(kg/m^2); d(g/l); e(mmol/l); f(gmol/l) or per year | | |
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Table 1: Population demographics for the continuous variables presented as: mean (sd) [min,max] <number missing (percent missing)>

Table 2 shows a breakdown of the categorical variables across the populations. In the development population, there are far more males than females, which is actually against what is believed to occur in the general public [**Cite: Something**], whereas in the validation population the proportions are much more matched. Most patients were white in the SKS dataset, and ethnicity has extremely high missingness in SERPR, which also contributed to its omission from the model. The majority of the SKS patients were former smokers, however this information was unavailable in the SERPR dataset. Primary Renal Diagnosis suffered from very high levels of missingness in the validation dataset, but was much better recorded in the development dataset (although still far from perfect).

| **Variable** | **Category** | **SKS (Development)** | **SERPR (Validation)** |
| --- | --- | --- | --- |
| Gender | |  |  |
| Male | | 1,865 (62.56 %) | 3,915 (50.42 %) |
| Female | | 1,116 (37.43 %) | 3,849 (49.57 %) |
| Ethnicity | |  |  |
| White | | 2,875 (96.44 %) | 683 ( 8.79 %) |
| Asian | | 75 ( 2.51 %) | 12 ( 0.15 %) |
| Black | | 21 ( 0.70 %) | 7 ( 0.09 %) |
| Other | | 10 ( 0.33 %) | 2 ( 0.02 %) |
| <Missing> | | 0 ( 0.00 %) | 7,060 (90.93 %) |
| Smoking Status | |  |  |
| Former | | 1,535 (51.49 %) |  |
| Non-Smoker | | 979 (32.84 %) |  |
| Smoker | | 379 (12.71 %) |  |
| Former 3Y | | 46 ( 1.54 %) |  |
| <Missing> | | 42 ( 1.40 %) |  |
| Renal Diagnosis | |  |  |
| Systemic diseases affecting the kidney | | 1,304 (43.74 %) | 299 ( 3.85 %) |
| Glomerular disease | | 442 (14.82 %) | 225 ( 2.89 %) |
| Tubulointerstitial disease | | 268 ( 8.99 %) | 164 ( 2.11 %) |
| Miscellaneous renal disorders | | 227 ( 7.61 %) | 188 ( 2.42 %) |
| Familial / hereditary nephropathies | | 173 ( 5.80 %) | 102 ( 1.31 %) |
| <Missing> | | 567 (19.02 %) | 6,786 (87.40 %) |

Table 2: Population demographics for the categorical variables presented as number (percent)

Overall, there were high levels of comorbidities within the SKS population as shown in table 3, but these levels were much lower in the SERPR population, possibly due to the data extraction processed (where data is un-recorded, no history is assumed). In SKS, most comorbidities were at over 80% prevalence, apart from diabetes mellitus, which had a lower prevalence of 33% and over 97% (2,891) patients had a history of liver disease. In SERPR, hypertension was the highest prevalence in SERPR at 40% (3,122), followed by diabetes mellitus at 20% (1,546) and cerebrovascular accident was the lowest prevalence at 2.36% (184). Liver disease, chronic obstructive pulmonary disease and solid tumour data were unavailable in the SERPR data.

| **Variable** | **SKS (Development)** | **SERPR (Validation)** |
| --- | --- | --- |
| DM | 992 (33.32%) < 4 (0.13%)> | 1,546 (19.91%) < 0 ( 0.00%)> |
| CCF | 2,414 (81.08%) < 4 (0.13%)> | 406 ( 5.22%) < 0 ( 0.00%)> |
| MI | 2,492 (83.70%) < 4 (0.13%)> | 556 ( 7.16%) < 0 ( 0.00%)> |
| IHD | 2,393 (80.38%) < 4 (0.13%)> | 867 (11.16%) < 0 ( 0.00%)> |
| PVD | 2,485 (83.47%) < 4 (0.13%)> | 376 ( 4.84%) < 0 ( 0.00%)> |
| CVA | 2,727 (91.60%) < 4 (0.13%)> | 184 ( 2.36%) < 0 ( 0.00%)> |
| COPD | 2,411 (80.98%) < 4 (0.13%)> |  |
| LD | 2,891 (97.11%) < 4 (0.13%)> |  |
| ST | 2,570 (86.32%) < 4 (0.13%)> |  |
| HT | 2,546 (91.48%) <198 (6.64%)> | 3,122 (40.21%) < 0 ( 0.00%)> |

Table 3: Population comorbidity prevalence for the two populations presented as number (percent) <number missing (percent missing)>

[**Data about Time Points (i.e. median event times)**]

## Development

Table @(ref:PH-Betas) shows the full results from the Three-State Models, the results for the Two-State and Five-State Models can be seen in [**Supplementary Material**]. Older patients are predicted to be likely to transition to RRT. Increased rates of decline of eGFR were associated with the transiton from CKD to RRT.

| **Var** | **CKD -> Dead** | **CKD -> RRT** | **RRT -> Dead** |
| --- | --- | --- | --- |
| (Age-60) | 0.161 ( -0.051, 0.374) | -0.041 ( -0.051, -0.031) | 0.063 ( 0.050, 0.076) |
| (Age-60)^2 | -0.000 ( -0.002, 0.000) | -0.000 ( -0.000, -0.000) |  |
| log(Age) | -5.725 ( -17.969, 6.518) |  |  |
| eGFR | -0.013 ( -0.019, -0.006) | -0.095 ( -0.108, -0.082) | 0.011 ( -0.001, 0.025) |
| eGFR Rate |  | 0.055 ( -0.021, 0.131) | -0.056 ( -0.363, 0.250) |
| log(eGFR Rate) | 0.042 ( -0.125, 0.210) |  | 0.227 ( -0.770, 1.225) |
| Gender : Female | -0.235 ( -0.371, -0.099) | -0.277 ( -0.455, -0.099) |  |
| SmokingStatus : Former\_3Y | -0.212 ( -0.879, 0.453) | -0.133 ( -0.757, 0.490) | -0.282 ( -1.082, 0.518) |
| SmokingStatus : Non\_Smoker | -0.198 ( -0.345, -0.051) | -0.162 ( -0.364, 0.039) | -0.294 ( -0.598, 0.009) |
| SmokingStatus : Smoker | 0.356 ( 0.160, 0.551) | 0.175 ( -0.076, 0.428) | 0.387 ( 0.068, 0.706) |
| SBP | -0.001 ( -0.004, 0.002) | 0.005 ( -0.000, 0.011) |  |
| DBP | 0.006 ( 0.000, 0.013) | 0.006 ( -0.001, 0.015) |  |
| BMI |  |  |  |
| Primary Renal : Familial / hereditary nephropathies | -0.424 ( -0.854, 0.006) | 1.029 ( 0.720, 1.338) | -0.562 ( -1.084, -0.040) |
| Primary Renal : Glomerular disease | -0.394 ( -0.635, -0.154) | -0.165 ( -0.465, 0.134) | -0.488 ( -0.883, -0.094) |
| Primary Renal : Miscellaneous renal disorders | -0.263 ( -0.505, -0.021) | -0.649 ( -1.143, -0.155) | 0.033 ( -0.553, 0.620) |
| Primary Renal : Tubulointerstitial disease | -0.463 ( -0.741, -0.184) | -0.265 ( -0.577, 0.046) | -0.310 ( -0.803, 0.181) |
| Albumin | -0.044 ( -0.064, -0.024) | -0.032 ( -0.059, -0.004) | -0.044 ( -0.079, -0.009) |
| Calcium | 0.280 ( -0.192, 0.752) | -0.515 ( -1.207, 0.177) |  |
| Haemoglobin | -0.013 ( -0.017, -0.008) | -0.005 ( -0.012, 0.001) | -0.005 ( -0.014, 0.003) |
| Phosphate | 0.511 ( 0.132, 0.890) | 0.869 ( -0.059, 1.799) |  |
| uPCR | 0.125 ( -0.318, 0.569) | 0.700 ( 0.112, 1.288) | -0.108 ( -0.736, 0.519) |
| uPCR Rate |  | -0.019 ( -0.045, 0.005) | 0.036 ( -0.062, 0.136) |
| log(uPCR Rate) |  | 0.218 ( -0.310, 0.747) | -0.198 ( -0.534, 0.137) |
| DM | 0.122 ( -0.011, 0.255) | 0.141 ( -0.074, 0.358) | 0.200 ( -0.096, 0.496) |
| CCF | -0.394 ( -0.535, -0.253) |  | -0.299 ( -0.597, -0.002) |
| MI | -0.246 ( -0.397, -0.094) | 0.234 ( -0.061, 0.530) | 0.186 ( -0.199, 0.572) |
| IHD | 0.102 ( -0.041, 0.245) | -0.077 ( -0.334, 0.179) | -0.097 ( -0.424, 0.228) |
| PVD | -0.248 ( -0.394, -0.103) | -0.168 ( -0.405, 0.068) | -0.183 ( -0.492, 0.126) |
| CVA | -0.070 ( -0.252, 0.111) |  | -0.168 ( -0.577, 0.240) |
| COPD | -0.289 ( -0.433, -0.145) |  |  |
| LD | -0.169 ( -0.578, 0.239) | -0.316 ( -0.731, 0.097) | -0.270 ( -0.858, 0.318) |
| ST | -0.274 ( -0.431, -0.117) | -0.181 ( -0.516, 0.153) | -0.278 ( -0.611, 0.055) |
| HT |  | 0.274 ( -0.176, 0.726) | -0.416 ( -1.104, 0.271) |

Table 4: Hazard Ratios for the Three-State Model

Female patients were more likely to remain in the CKD state than Males, or to remain in the RRT state once there. Smokers were more likely than Non-/Former Smokers to undergo any transition, apart from CKD to Tx. Blood results had effects on all transitions in some way, and Primary Renal Diagnosis were strongly related to predictions. Patients with Liver Disease were much more likely to undergo a Transplant than those without Liver Disease ().

[[1]] function(x) exp(case\_when(0 <= t & t < 3 ~ 1.19795*log(t)+17.68798, 3 <= t & t < 443 ~ -9e-05*log(t)3+3e-04*log(t)2+1.19761log(t)+17.68811, 443 <= t & t < 873 ~ 0.15869log(t)3-2.9019*log(t)2+18.88018*log(t)-18.22403, 873 <= t & t < 1295 ~ -0.30096*log(t)3+6.43659*log(t)2-44.36299log(t)+124.54338, 1295 <= t & t < 1876 ~ -0.04158log(t)3+0.86028*log(t)2-4.40166*log(t)+29.08554, 1876 <= t & t < 2738 ~ 0.51263*log(t)3-11.67048*log(t)2+90.03919log(t)-208.17245, 2738 <= t & t < 5497 ~ -0.23992log(t)3+6.19863*log(t)2-51.3924*log(t)+164.96467, 5497 <= t ~ 1.98997*log(t)+11.72244)) <environment: 0x0000021a11dfc7b8>

[**Baseline Hazards**]

## Example

## Validation

# Discussion

We have used data provided by SKS to develop a Multi-State Clinical Prediction Model and then validated this model within the SKS and SERPR datasets. Within our Models, the cause of a patient’s renal disease had the widest effect on patient outcomes meaning that outcomes are highly dependent on ERA-EDTA classification of the diagnosis. Most groupings resuted in a lowered hazard of death and an increased hazard of RRT compared to the baseline of Systemic diseases.

[**Baseline Hazards go up or down?**]r xx(“Validation results”)

Paragraph 2

In all transitions relating to the

The application of a Multi-state clinical prediction model to this field is novel and gives a powerful tool for providing individualised predictions of multiple outcomes at a wide range of time points. The general inclusion criteria for the development dataset, and the wide range of patient ages and measurements allows for the model to be applied to a broad spectrum of patients.

Although the inclusion criteria for SKS were broad, the demographics of the local area resulted in homogeneity of ethnicity, which may create a limitation to the applicability of our model. The Renal Department at SRFT is a tertiary care facility for CKD sufferers and is well renowned for its capabilities of care meaning that it is likely to attract less-healthy patients from a wider catchment area, making the cohort of patients in the development population in worse condition than the general population of CKD patients.

There were also high levels of missingness in the eGFR and uPCR rates of changes would also produce a bias, due to these measures likely being missing not at random. The derivation of the validation dataset ensured that all patients had an eGFR Rate measurement; this was done to avoid data missing not at random (only negative or missing data would be available as patient’s eGFR dropped to less than 60), however deriving data in this way could itself induce a survivor bias in the start date used for patients.

In the Five-State Model, We omitted the analysis of the Tx to Dead state due to the anticipated low number of events within the SKS dataset. The lowest number of events for a transition was therefore PD to Dead, which had only [**Number of PD Events**]. Altogether, we considered [**Number of covariates**] and so this equates to an EPV [**EPV**], which is below the usual rule of thumb of 10 EPV. We also compared this to the work done by [**Cite: Events-per for CPMs, Riley Group**] and, even with generous prior information, [**Which models satisfied the criteria?**].

We have assumed a proportional hazards relationship between the predictors and probability of survival, which is considered by some to be a strong assumption to make, however we acknowledge this limitation, and the authors believe that it is mitigated by the flexibility that the assumption permits. In addition to the general PH assumption, the R-P model requires the assumption that the log cumulative hazard function follows a cubic spline, (however this is a fairly weak assumption [**Cite: Something**]), which is modelled as part of the regression. We did not assess the viability of these models as it was believed this assumption to make our results more understandable.

[**More limitations**]

Despite this, we believe that the advantages of such a model far outweigh these shortcomings.

Compared to the raw internal validation, the model performance during the external validation was worse for all metrics. However, once adjusted for optimism, the results were much more cohesive which implies that the model is highly transportable to a new population without much alterations being required. Due to the differences in the healthcare systems of England and Scotland, it can be appreciated that despite the populations being similar, their care would be different enough to emphasise a larger difference between our populations than that shown in our (relatively homogeneous) populations [**Cite: Kidney England vs Scotland**].

Although not directly assessing causality in regards to state-transitions, our Three-State model can be used by clinicians to either expediate or delay transitions of patient onto RRT, if it is believed that this would be beneficial. Alternatively, the Five-State Model can be interpreted to provide information regarind *which* treatment a patient should undergo.

Further research would be needed to establish the efficacy of the use of such a model in clinical practice [80] by comparing it to standard care and establishing whether the use of our model improves patient outcomes, however we have clearly demonstrated the *effectiveness* of our model.

All three models produced for this work performed well in terms of accuracy, calibration and discrimination when applied internally and externally. This shows directly that the models are suitable for use in populations similar to both our development and our validation datasets. It can also be concluded that the models can be transported to any population with a similar healthcare system to the UK.

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