

Lecture Notes VII – Double descent

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“Double descent” on a simple example

Reading HTF Ch.: , Murphy Ch.: , Bach Ch.: Ch.10.2.3

Linear regression when $n > d$

- ▶ We describe a very simple linear regression situation (following Bach section 10.2.3)
- ▶ For it, we are able to explicitly obtain the expected estimation error $E[\|\theta_{\text{true}} - \hat{\theta}\|^2]$
- ▶ Surprisingly, the variance of this error decreases with d , and the error itself has a limit proportional to $\|\theta_{\text{true}}\|^2$.
- ▶ Input distribution $x^{1:n} \sim N(0, I_d)$, noise $\epsilon^{1:n} \sim N(0, \sigma^2)$
- ▶ Model $y^i = (x^i)^T \theta_{\text{true}} + \epsilon^i$.
- ▶ Denote $X \in \mathbb{R}^{n \times d}$, $y, \epsilon \in \mathbb{R}^n$ the usual input matrix, output, and noise vectors respectively
- ▶ Denote $K = XX^T \in \mathbb{R}^{n \times n}$ the Gram matrix (or kernel matrix). We assume K is non-singular
- ▶ From Lecture IV, [The Implicit Bias of Gradient Descent](#) we know that
 - ▶ When X is full rank n , the equation $y = X\theta$ has multiple solutions θ
 - ▶ Gradient Descent converges to the min norm solution $\hat{\theta} = X^T K^{-1} y$

The expected estimation error $MSE(\theta_{true}) = E[\|\theta_{true} - \hat{\theta}\|^2]$

- Decompose $\hat{\theta}$

$$\hat{\theta} = X^T K^{-1} y = X^T K^{-1} (X\theta_{true} + \epsilon) = X^T K^{-1} X \theta_{true} + X^T K^{-1} \epsilon \quad (1)$$

- Then,

$$MSE(\theta_{true}) = E_{X,\epsilon}[\|\theta_{true} - \hat{\theta}\|^2] \quad (2)$$

$$= \underbrace{E_X[\theta_{true}^T (I_d - X^T K^{-1} X) \theta_{true}]}_{\text{bias}^2} + \underbrace{E_{X,\epsilon}[\epsilon^T K^{-1} X X^T K^{-1} \epsilon]}_{\text{variance}} \quad (3)$$

- The Variance term becomes

$$Var = E_{X,\epsilon}[\epsilon^T K^{-1} \epsilon] \quad (4)$$

$$= E_X[\text{trace } K^{-1}] \sigma^2 \quad \text{Wishart!} \quad (5)$$

$$= \sigma^2 \frac{n}{d - n - 1} \quad (6)$$

The expected estimation error $MSE(\theta_{true}) = E[\|\theta_{true} - \hat{\theta}\|^2]$

- ▶ The Bias² term:
- ▶ Note that $\theta_P = X^T K^{-1} X \theta_{true}$ is the orthogonal projection of θ_{true} on the row space of X , and $\theta_{true}^T X^T K^{-1} X \theta_{true} = \|\theta_P\|^2$.
- ▶ The subspace is a random subspace of dimension n in \mathbb{R}^d . By spherical symmetry, the length of the projection of a fixed vector on a random subspace is the same with that of a projecting a random vector of length (squared) $\|\theta_{true}\|^2$ on a fixed subspace, e.g. the first d unit vectors in \mathbb{R}^d . The latter expected value is easy to compute

$$E[\|\theta_P\|^2] = \frac{n}{d} \|\theta_{true}\|^2 \quad (7)$$

Exercise Proving this is a moderately easy exercise

- ▶ Hence,

$$\text{bias}^2 = E_X[\theta_{true}^T (I_d - X^T K^{-1} X) \theta_{true}] = \frac{d-n}{d} \|\theta_{true}\|^2 \quad (8)$$

The expected estimation error $MSE(\theta_{true}) = E[\|\theta_{true} - \hat{\theta}\|^2]$

- ▶ Finally

$$MSE(\theta_{true}) = E[\|\theta_{true} - \hat{\theta}\|^2] = \frac{d-n}{d} \|\theta_{true}\|^2 + \sigma^2 \frac{n}{d-n-1} \quad (9)$$

for $d > n+1$

- ▶ When $d \rightarrow \infty$, the variance $\rightarrow 0$ and the bias² $\rightarrow \|\theta_{true}\|^2$