



Expectation-Maximization Algorithm

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Outline

- The Concerned Problem
- EM Algorithm
- Theoretical Guarantees
- Example 1: Training Gaussian Latent-Variable Models
- Example 2: Training Gaussian Mixture Models

General Form of the Concerned Problem

- Given the joint distribution

$$p(\boldsymbol{x}, \boldsymbol{z}; \boldsymbol{\theta}),$$

where \boldsymbol{x} is the observed variable and \boldsymbol{z} is the latent variable, we need to maximize the log likelihood *w.r.t.* $\boldsymbol{\theta}$, that is,

$$\boldsymbol{\theta} = \arg \max_{\boldsymbol{\theta}} \log p(\boldsymbol{x}; \boldsymbol{\theta}),$$

where

$$p(\boldsymbol{x}; \boldsymbol{\theta}) = \sum_{\boldsymbol{z}} p(\boldsymbol{x}, \boldsymbol{z}; \boldsymbol{\theta})$$

What we have is the joint pdf, but what we need to optimize is the marginal pdf

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EM Algorithm

- Algorithm

E-step: Evaluating the expectation

M-step: Updating the parameter

$$\theta^{(\ell+1)} = \arg \max_{\theta} \mathcal{Q}(\theta; \theta^{(\ell)})$$

- Key ingredient in EM

- 1) The posteriori distribution
- 2) The expectation of joint distribution *w.r.t.* the posteriori
- 3) Maximization

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Re-representing the Log-likelihood

- The log-likelihood can be reformulated as

$$-\mathcal{L}(q, \theta) + KL(q \parallel p(z \mid x; \theta)), \text{ for } \forall \theta, q(z)$$

Remark: The KL-divergence is used to *measure the distance* between two distributions q and p , which is defined as

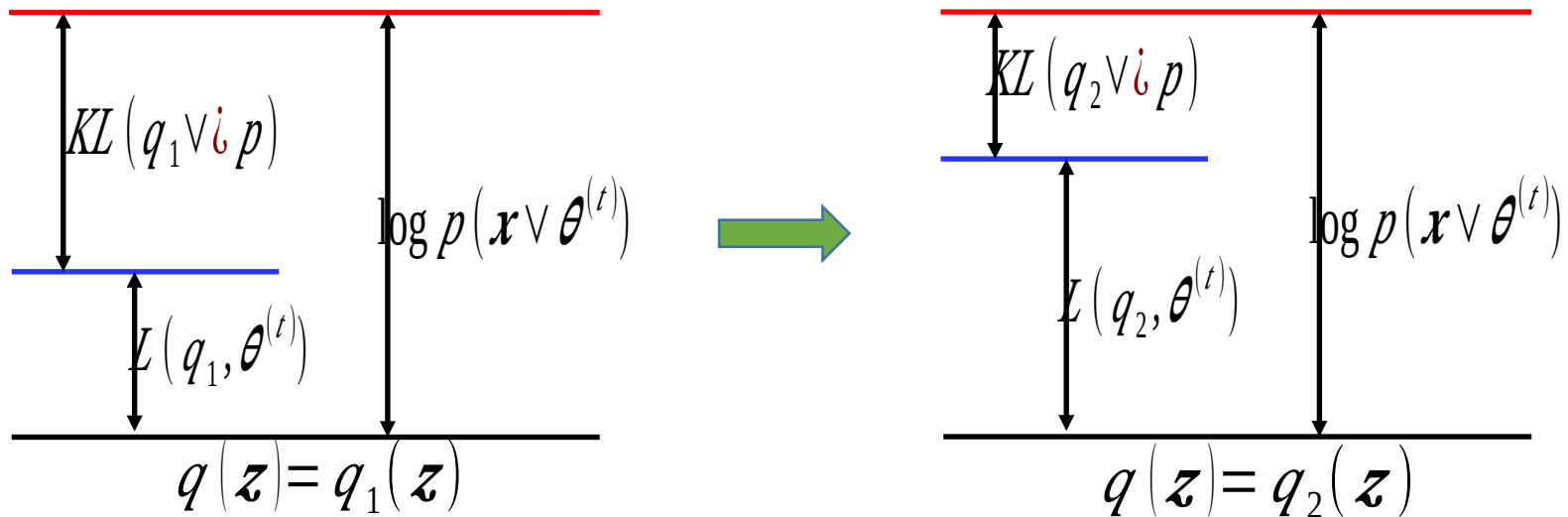
$$KL(q \parallel p)$$

- Thus, with the parameter at the -th iteration, we have

$$\log p(\mathbf{x}; \boldsymbol{\theta}^{(t)}) = \mathcal{L}(q, \boldsymbol{\theta}^{(t)}) + KL(q \vee p(\mathbf{z} \vee \mathbf{x}; \boldsymbol{\theta}^{(t)}))$$

This equality holds for any distribution

- Different will lead to different decomposition of



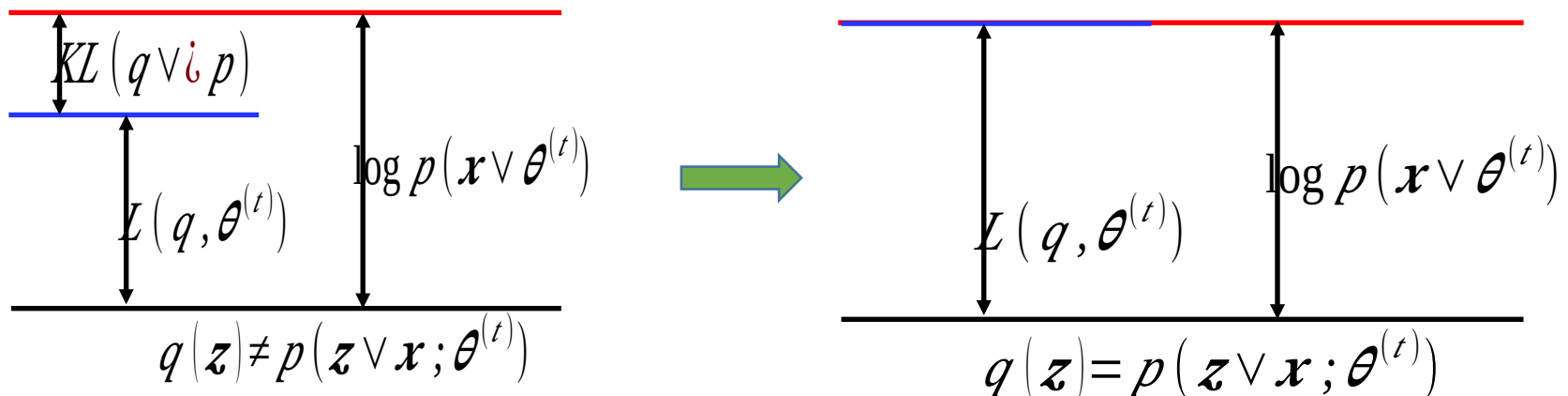
Theoretical Justification for EM

- If we set , then we have

$$KL(q \vee p(\mathbf{z} \vee \mathbf{x}; \theta^{(t)})) = 0$$

Thus, we have

$$\log p(\mathbf{x} \vee \theta^{(t)}) = \mathcal{L}(p(\mathbf{z} \vee \mathbf{x}; \theta^{(t)}), \theta^{(t)})$$



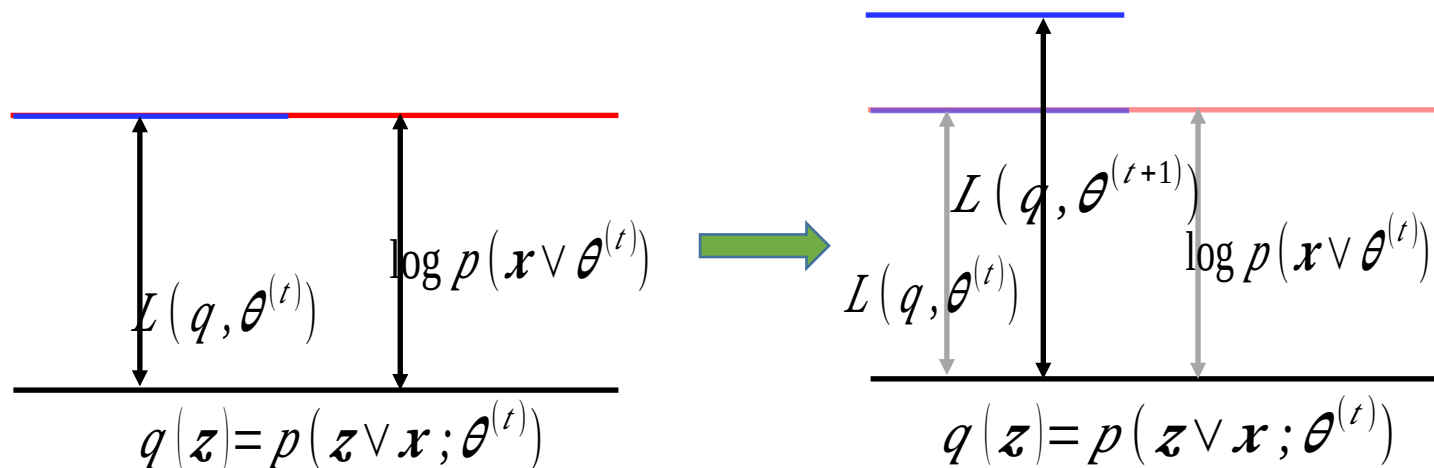
$$\log p(\mathbf{x} \vee \boldsymbol{\theta}^{(t)}) = \mathcal{L}(p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)}), \boldsymbol{\theta}^{(t)})$$

- If we update as

$$\boldsymbol{\theta}^{(t+1)} = \arg \max_{\boldsymbol{\theta}} \mathcal{L}(p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)}), \boldsymbol{\theta}),$$

then we must have the relation

$$\mathcal{L}(p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)}), \boldsymbol{\theta}^{(t+1)}) \geq \underbrace{\mathcal{L}(p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)}), \boldsymbol{\theta}^{(t)})}_{\log p(\mathbf{x} \vee \boldsymbol{\theta}^{(t)})}$$



- From the nonnegative property of KL-divergence, we know that

$$KL\left(p(\mathbf{z} \vee \mathbf{x}; \boldsymbol{\theta}^{(t)}) \vee p(\mathbf{z} \vee \mathbf{x}; \boldsymbol{\theta}^{(t+1)})\right) \geq 0$$

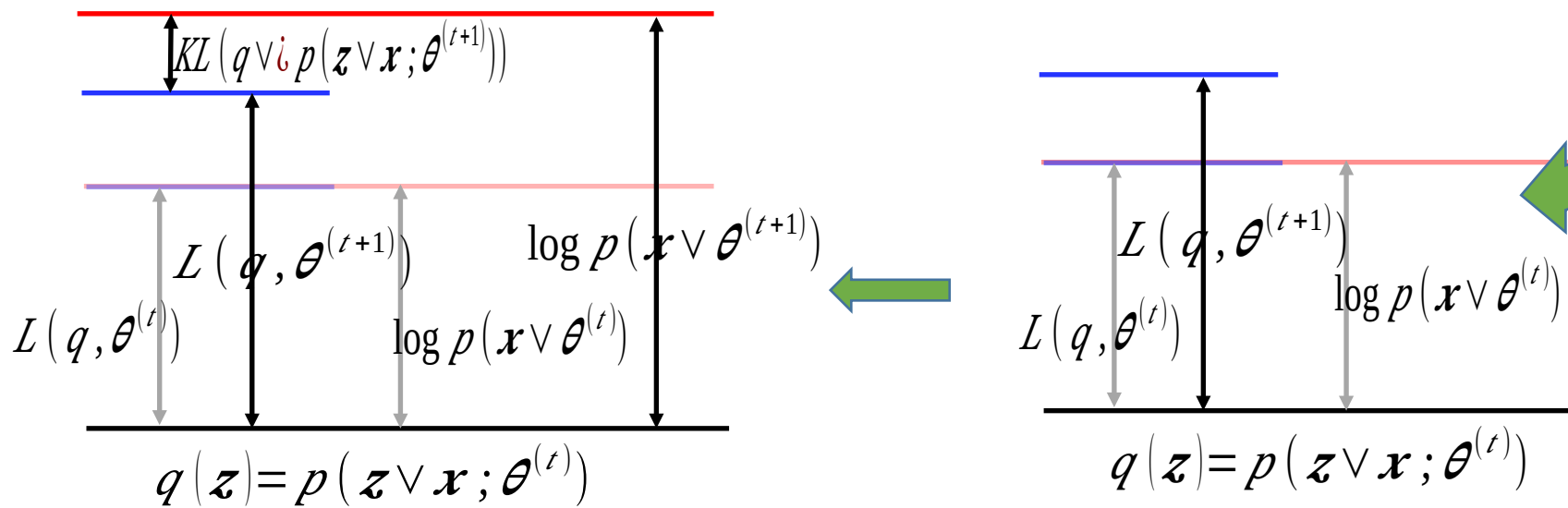
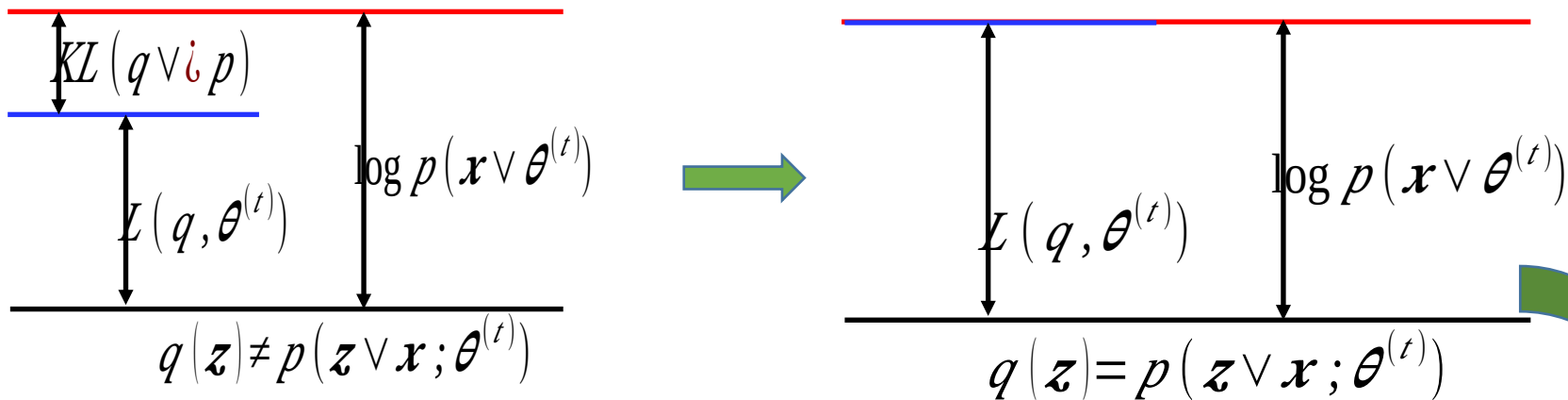
- Because holds for any , thus we have

$$\log p(\mathbf{x}; \boldsymbol{\theta}^{(t+1)}) = \underbrace{\mathcal{L}\left(p(\mathbf{z} \vee \mathbf{x}; \boldsymbol{\theta}^{(t)}), \boldsymbol{\theta}^{(t+1)}\right)}_{\geq \log p(\mathbf{x}; \boldsymbol{\theta}^{(t)})} + \underbrace{KL\left(p(\mathbf{z} \vee \mathbf{x}; \boldsymbol{\theta}^{(t)}) \vee p(\mathbf{z} \vee \mathbf{x}; \boldsymbol{\theta}^{(t+1)})\right)}_{\geq 0}$$

- Thus, we can see that

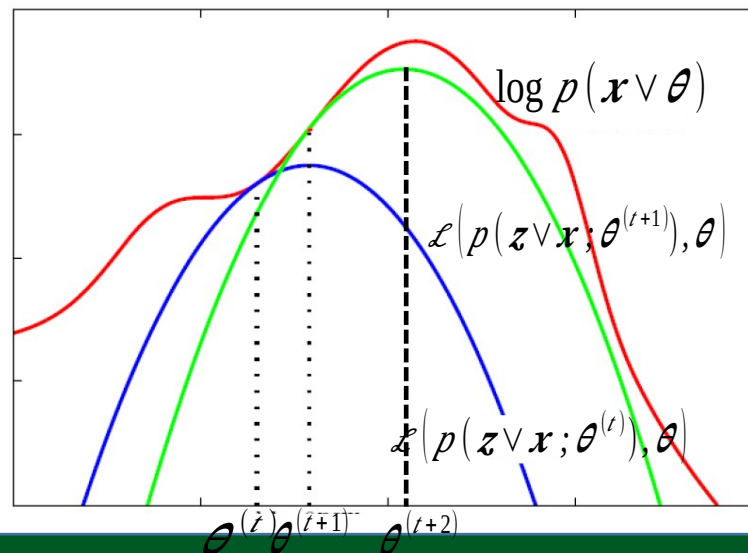
$$\boxed{\log p(\mathbf{x}; \boldsymbol{\theta}^{(t+1)}) \geq \log p(\mathbf{x}; \boldsymbol{\theta}^{(t)})}$$

EM algorithm can guarantee the increase of likelihood at each step



A View in the Parameter Space

- 1) E-step (t): deriving the expression $q(z|x; \theta^{(t)})$ given the model parameter $\theta^{(t)}$
- 2) M-step (t): computing the optimal value $\theta^{(t+1)}$
- 3) E-step ($t+1$): deriving the expression for $q(z|x; \theta^{(t+1)})$ given the model parameter $\theta^{(t+1)}$
- 4) Repeating the above process until convergence

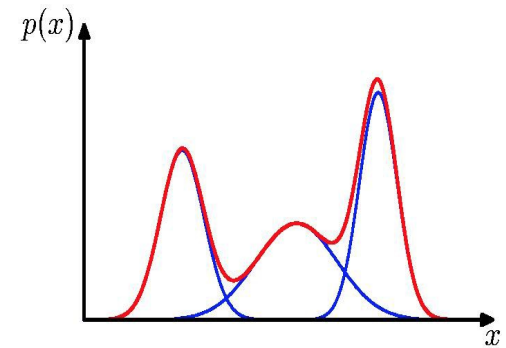


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- **Example 1: Gaussian Mixture Models**
- Example 2: Training Probabilistic PCA Models

Gaussian Mixture Model Review

- For a Gaussian mixture distribution, *i.e.*,



it can be represented as the marginal distribution of the joint distribution

$$p(\mathbf{x}, \mathbf{z}) = p(\mathbf{x} | \mathbf{z}) p(\mathbf{z})$$

- follows the categorical distribution with parameter

EM: E-step

- The posteriori distribution

$$\mathbf{p}^{(i)}$$

- denotes the one-hot vector with the i -th element being 1

- The log of the joint distribution

$$\log \pi(\mathbf{y}, \mathbf{z})$$

Note that $\mathbf{p}^{(i)}$ can only be a one-hot vector

- The expectation

➤ Due to , we have

- Therefore, we have

- Taking into gives

$$\mathcal{Q}(\theta; \theta^{(t)}) = \sum_{k=1}^K \gamma_k^{(t)} \left[-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_k)^T \boldsymbol{\Sigma}_k^{-1} (\mathbf{x} - \boldsymbol{\mu}_k) - \frac{1}{2} |\boldsymbol{\Sigma}_k| + \log \pi_k \right] + C$$

- is the constant

- So far, only one data example is considered
- If data for are considered, the becomes

$$\mathcal{Q}(\theta; \theta^{(t)}) = \frac{1}{N} \sum_{n=1}^N \sum_{k=1}^K \gamma_{nk}^{(t)} \left[-\frac{1}{2} (\mathbf{x}^{(n)} - \boldsymbol{\mu}_k)^T \boldsymbol{\Sigma}_k^{-1} (\mathbf{x}^{(n)} - \boldsymbol{\mu}_k) - \frac{1}{2} |\boldsymbol{\Sigma}_k| + \log \pi_k \right] + C$$

EM: M-step

- By taking derivatives *w.r.t.* μ_k and Σ_k and setting them to zero, we obtain the optimal μ_k as

$$\mu_k^{(t+1)} = \frac{1}{N_k} \sum_{n=1}^N \gamma_{nk} \mathbf{x}_n$$

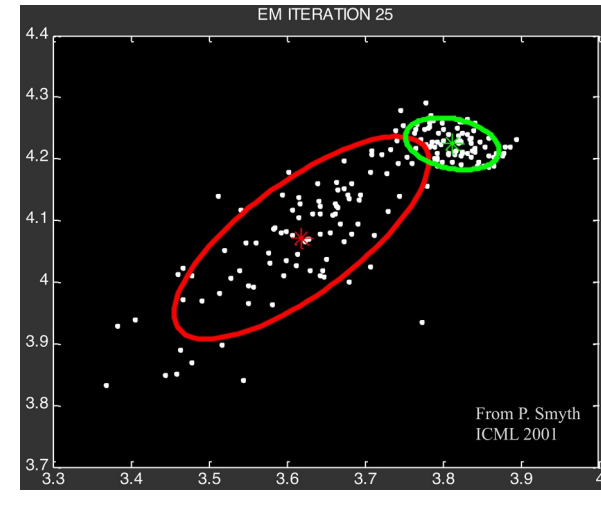
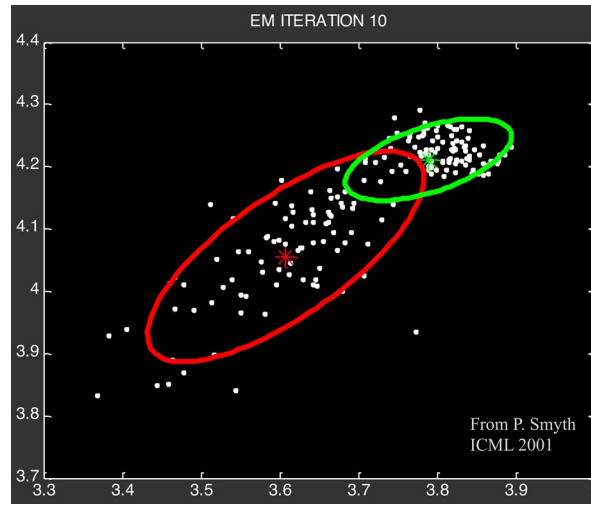
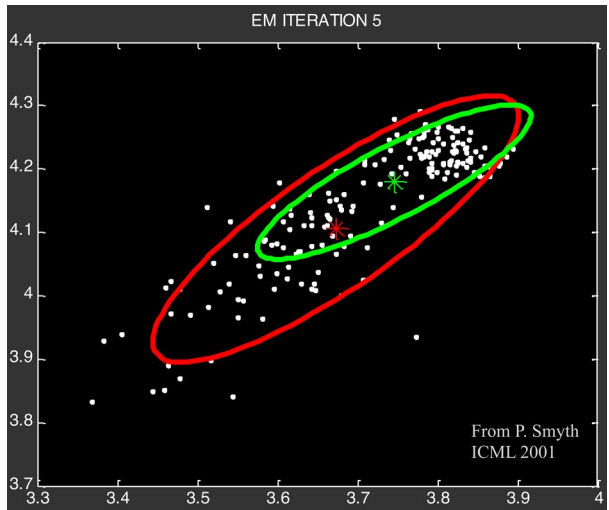
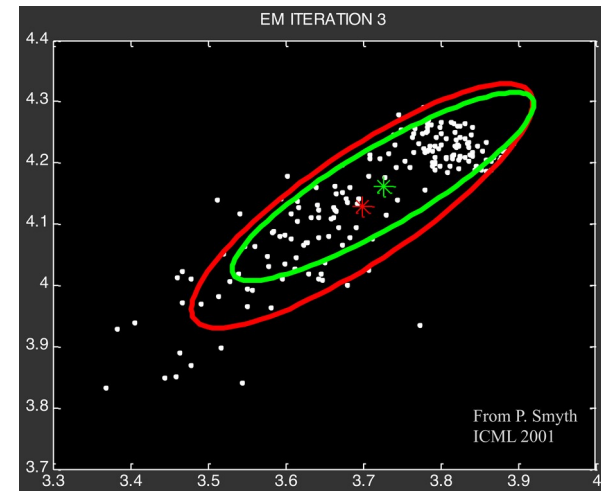
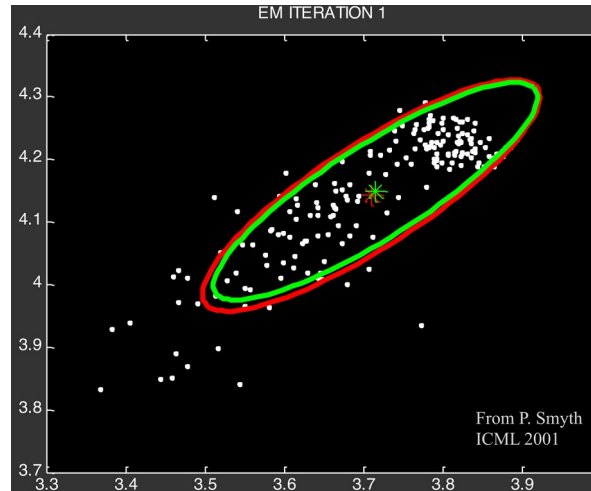
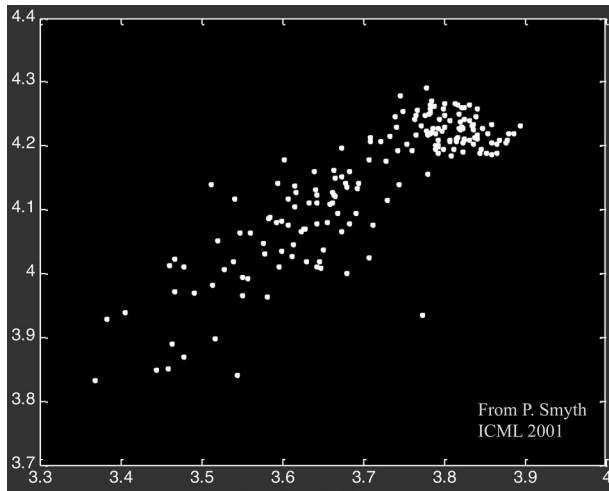
$$\Sigma_k^{(t+1)} = \frac{1}{N_k} \sum_{n=1}^N \gamma_{nk} (\mathbf{x}_n - \mu_k^{(t+1)}) (\mathbf{x}_n - \mu_k^{(t+1)})^T$$

$$\pi_k^{(t+1)} = \frac{N_k}{N}$$

where N_k is the effective number of examples assigned to the k -th class

Summary of EM Algorithm

- Given the current estimate $\theta^{(t)}$, update η as
- Given the η , update θ and $\hat{\theta}$ as



Relation to Soft K -Means

- When restricting , the updating of GMM becomes

$$\pi_k \leftarrow \frac{\sum_{n=1}^N \gamma_{nk}}{N}$$

$$\mu_k \leftarrow \frac{\sum_{n=1}^N \gamma_{nk} \mathbf{x}_n}{\sum_{n=1}^N \gamma_{nk}}$$

where

- Updates in soft K -means

$$r_{nk} = \frac{e^{-\beta \|\mathbf{x}^{(n)} - \mu_k\|^2}}{\sum_{i=1}^K e^{-\beta \|\mathbf{x}^{(n)} - \mu_i\|^2}}$$

$$\mu_k \leftarrow \frac{\sum_{n=1}^N r_{nk} \mathbf{x}_n}{\sum_{n=1}^N r_{nk}}$$

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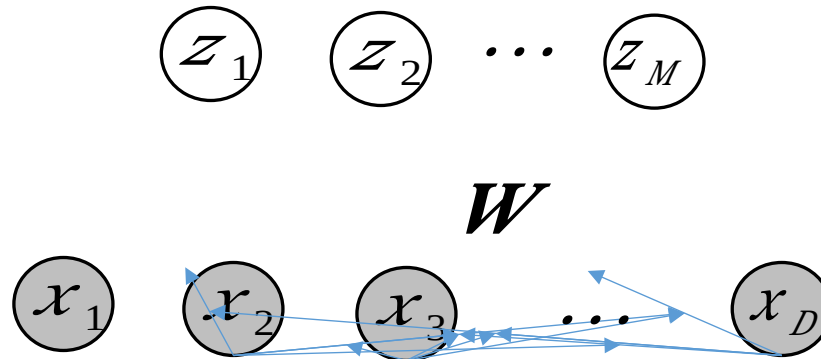
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Probabilistic PCA Review

- Probabilistic PCA model

Prior distribution: $p(\mathbf{z}) = \mathcal{N}(\mathbf{z}; \mathbf{0}, \mathbf{I})$

Likelihood function: $p(\mathbf{x} | \mathbf{z}) = \mathcal{N}(\mathbf{x}; \mathbf{W}\mathbf{z} + \boldsymbol{\mu}, \sigma^2 \mathbf{I})$



- The objective is to maximize the \log likelihood $w.r.t.$ all training data points

Two EM Steps

- It is a latent-variable model, thus we can use **EM** to optimize it

Remark: maximizing is equivalent to

- *Reminder:* Key integrant in EM
 - **E-step:** Expectation *w.r.t.* the posteriori

$$\mathcal{Q}(\theta; \theta^{(t)}) = \sum_{n=1}^N \mathbb{E}_{p(\mathbf{z}_n \vee \mathbf{x}_n; \theta^{(t)})} [\log p(\mathbf{x}_n, \mathbf{z}_n; \theta)]$$

- **M-step:** Maximization

$$\theta^{(t+1)} = \arg \max_{\theta} \mathcal{Q}(\theta; \theta^{(t)})$$

E-Step: Evaluating

- From

$$p(\mathbf{x}, \mathbf{z}; \theta) = \frac{1}{(2\pi\sigma^2)^{D/2}} e^{-\frac{\|\mathbf{x} - \mathbf{W}\mathbf{z} - \boldsymbol{\mu}\|^2}{2\sigma^2}} \cdot \frac{1}{(2\pi)^{M/2}} e^{-\frac{\|\mathbf{z}\|^2}{2}}$$

we obtain

$$\log p(\mathbf{x}, \mathbf{z}; \theta) = -\frac{D}{2} \log 2\pi\sigma^2 - \frac{M}{2} \log 2\pi - \frac{\|\mathbf{x} - \mathbf{W}\mathbf{z} - \boldsymbol{\mu}\|^2}{2\sigma^2} - \frac{\|\mathbf{z}\|^2}{2}$$

- Thus, we have

$$\mathcal{Q}(\theta; \theta^{(t)}) = \sum_{n=1}^N \left(-\frac{1}{2\sigma^2} \|\boldsymbol{\mu}\|^2 + \frac{1}{\sigma^2} (\mathbf{x} - \boldsymbol{\mu})^T \mathbf{W} \mathbb{E}_{\mathbf{z}_n} [\mathbf{z}_n] - \frac{1}{2\sigma^2} \text{Tr} \left(\mathbf{W}^T \mathbf{W} \mathbb{E}_{\mathbf{z}_n} [\mathbf{z}_n \mathbf{z}_n^T] \right) + C \right)$$

- denotes the expectation *w.r.t.* the distribution
- means the trace operation, and is irrelevant to *and*

M-Step: Maximization

- The global optimal μ is already known to be $\bar{\mathbf{x}}$, so we fix

$$\mu = \bar{\mathbf{x}}$$

- By deriving

$$\frac{\partial \mathcal{Q}(\boldsymbol{\theta}; \boldsymbol{\theta}^{(t)})}{\partial \mathbf{W}} = -\frac{1}{\sigma^2} \sum_{n=1}^N \left(\mathbf{W} \mathbb{E}_{\mathbf{z}_n} [\mathbf{z}_n \mathbf{z}_n^T] - (\mathbf{x} - \bar{\mathbf{x}}) \mathbb{E}_{\mathbf{z}_n} [\mathbf{z}_n^T] \right)$$

and setting $\frac{\partial \mathcal{Q}}{\partial \mathbf{W}} = 0$, we obtain

$$\mathbf{W}^{(t+1)} \leftarrow \left(\sum_{n=1}^N (\mathbf{x}_n - \bar{\mathbf{x}}) \mathbb{E}_{\mathbf{z}_n} [\mathbf{z}_n^T] \right) \left(\sum_{n=1}^N \mathbb{E}_{\mathbf{z}_n} [\mathbf{z}_n \mathbf{z}_n^T] \right)^{-1}$$

How to get the expectations and

- Given the data \mathbf{X} , and fixing β , it can be derived that the posterior is

$$p(\mathbf{z}_n | \mathbf{x}_n) = \mathcal{N}(\mathbf{z}_n | \mathbf{M}^{-1} \mathbf{W}^T (\mathbf{x}_n - \bar{\mathbf{x}}), \sigma^2 \mathbf{M}^{-1})$$

where

- From the distribution, we can easily obtain

$$\mathbb{E}_{\mathbf{z}_n}[\mathbf{z}_n] = \mathbf{M}^{-1} \mathbf{W}^T (\mathbf{x}_n - \bar{\mathbf{x}})$$

$$\mathbb{E}_{\mathbf{z}_n}[\mathbf{z}_n \mathbf{z}_n^T] = \sigma^2 \mathbf{M}^{-1} + \mathbb{E}_{\mathbf{z}_n}[\mathbf{z}_n] \mathbb{E}_{\mathbf{z}_n}[\mathbf{z}_n^T]$$

Using 'completing the square' trick to derive the posteriori

$$\begin{aligned}
 \log p(\mathbf{x}, \mathbf{z}; \theta) &= \underbrace{-\frac{D}{2} \log 2\pi \sigma^2 - \frac{M}{2} \log 2\pi}_{C_1} - \frac{\|\mathbf{x} - \mathbf{W}\mathbf{z} - \boldsymbol{\mu}\|^2}{2\sigma^2} - \frac{\|\mathbf{z}\|^2}{2} \\
 &\stackrel{!}{=} \underbrace{C_1 - \frac{1}{2\sigma^2} (\|\mathbf{x}\|^2 - 2\boldsymbol{\mu}^T \mathbf{x} + \|\boldsymbol{\mu}\|^2)}_{\phi(\mathbf{x})} - \frac{1}{2\sigma^2} (-2\mathbf{x}^T \mathbf{W}\mathbf{z} + 2\boldsymbol{\mu}^T \mathbf{W}\mathbf{z} + \|\mathbf{W}\mathbf{z}\|^2) - \frac{1}{2} \|\mathbf{z}\|^2 \\
 &\stackrel{!}{=} \phi(\mathbf{x}) + \frac{1}{\sigma^2} (\mathbf{x} - \boldsymbol{\mu})^T \mathbf{W}\mathbf{z} - \frac{1}{2\sigma^2} \mathbf{z}^T \mathbf{M}\mathbf{z} \\
 &\stackrel{!}{=} -\frac{1}{2\sigma^2} \left(\mathbf{z} - \mathbf{M}^{-1} \mathbf{W}^T (\mathbf{x} - \boldsymbol{\mu}) \right)^T \mathbf{M} \left(\mathbf{z} - \mathbf{M}^{-1} \mathbf{W}^T (\mathbf{x} - \boldsymbol{\mu}) \right) + \eta(\mathbf{x}) \\
 &\quad \Rightarrow \\
 &\quad \Rightarrow
 \end{aligned}$$

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- Thank You!