

- 1) In a linear regression model using OLS identify the effect of correlated features on the matrix computation. (Through Derivation)

→ Multiple linear regression

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_k x_k + \epsilon$$

$\epsilon \rightarrow$ residual terms of the model.

$\beta_0, \beta_1, \beta_2, \dots, \beta_k \rightarrow$ regression coefficients.

Interaction with 2 or more variables.

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_{12} x_1 x_2 + \epsilon \quad \text{--- (1)}$$

This is called linear regression model as they can be called as linear combination of the β -parameters in the model.

$x \rightarrow$ weights and they can be non-linear

So the matrix form for linear regression is:-

$$\begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{bmatrix}_{n \times 1} = \begin{bmatrix} 1 & x_{11} & x_{21} & \dots & x_{k1} \\ 1 & x_{12} & x_{22} & \dots & x_{k2} \\ \vdots & \vdots & \vdots & \dots & \vdots \\ 1 & x_{n1} & x_{n2} & \dots & x_{kn} \end{bmatrix}_{n \times k} \begin{bmatrix} \beta_1 \\ \beta_2 \\ \vdots \\ \beta_k \end{bmatrix}_{k \times 1} + \begin{bmatrix} \epsilon_1 \\ \epsilon_2 \\ \vdots \\ \epsilon_n \end{bmatrix}_{n \times 1}$$

Criteria for Estimates

→ Select such $\hat{\beta}$ that it minimizes the sum of Squared residuals. (1)

$$e = y - X\hat{\beta}$$

Sum of squared residuals (RSS) = $e'e$.

$$[e_1 \ e_2 \ \dots \ e_n]_{1 \times n} \begin{bmatrix} e_1 \\ e_2 \\ \vdots \\ e_n \end{bmatrix}_{n \times 1} = [e_1 \times e_1 + e_2 \times e_2 + \dots + e_n \times e_n]_{1 \times 1} \rightarrow (3)$$

Sum of Squared residuals are

$$\begin{aligned} e'e &= (y - X\hat{\beta})'(y - X\hat{\beta}) \\ &= y'y - \hat{\beta}'X'y - y'X\hat{\beta} + \hat{\beta}'X'X\hat{\beta} \\ &= y'y - 2\hat{\beta}'X'y + \hat{\beta}'X'X\hat{\beta} \rightarrow (4) \end{aligned}$$

Derivate of eq (4) with respect to $\hat{\beta}$.

$$\frac{\partial e'e}{\partial \hat{\beta}} = -2X'y + 2X'X\hat{\beta} = 0 \rightarrow (5)$$

$$\frac{\partial e'e}{\partial \hat{\beta}} = +2X'X \rightarrow (6)$$

Matrix Differentiation

$$\frac{\partial a'b}{\partial b} = \frac{\partial b'a}{\partial b} = a \quad \rightarrow (6)$$

a and b are $K \times 1$ vectors.

$$\frac{\partial b'Ab}{\partial b} = 2Ab = 2b'A \quad \rightarrow (7)$$

A - Symmetric matrix.

$$\frac{\partial 2\beta'x'y}{\partial \beta} = \frac{\partial 2\beta'(x'y)}{\partial \beta} = 2x'y \quad \rightarrow (8)$$

$$\frac{\partial \beta'x'x\beta}{\partial \beta} = \frac{\partial \beta'Ax\beta}{\partial \beta} = 2Ax\beta = 2x'Ax\beta \quad \rightarrow (9)$$

$x'x = K \times K$ matrix

Normal equation from eqn 5

$$(x'x)\hat{\beta} = x'y \quad \rightarrow (10)$$

Multiplying both sides by $(x'x)^{-1}$.

$$(x'x)^{-1}(x'x)\hat{\beta} = x'y \cdot (x'x)^{-1} \quad \rightarrow (11)$$

$$I \cdot \hat{\beta} = (x'x)^{-1} \cdot x' \cdot y \quad \because (x'x)^{-1}(x'x) = I$$

$\therefore I = \text{Identity matrix}$

$$\hat{\beta} = (x'x)^{-1}x'y \rightarrow (12)$$

* Properties of the OLS Estimators

① Normal form can form. Eq(10)
 $(x'x)\hat{\beta} = x'y \rightarrow (13)$

$$(x'x)\hat{\beta} = x'(x\hat{\beta} + e)$$

∴ Substitute
 $y = x\hat{\beta} + e$

$$(x'x)\hat{\beta} = (x'x)\hat{\beta} + x'e$$

$x'e = 0$

 $\rightarrow (14)$

Matrix.

$$\begin{bmatrix} x_{11} & x_{12} & \dots & x_{1n} \\ x_{21} & x_{22} & \dots & x_{2n} \\ \vdots & \vdots & \dots & \vdots \\ x_{k1} & x_{k2} & \dots & x_{kn} \end{bmatrix} \begin{bmatrix} e_1 \\ e_2 \\ \vdots \\ e_n \end{bmatrix} = \begin{bmatrix} x_{11}e_1 + x_{12}e_2 + \dots + x_{1n}e_n \\ x_{21}e_1 + x_{22}e_2 + \dots + x_{2n}e_n \\ \vdots \\ x_{k1}e_1 + x_{k2}e_2 + \dots + x_{kn}e_n \end{bmatrix}$$

$$= \begin{bmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix} \rightarrow (15)$$

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- (1) Observed values ~~of~~ are uncorrelated with the residual.
 - (2) The sum of the residuals are zero
 - (3) Sample mean of the residual are zero
 - (4) The regression hyperplane passes through the means of the observed values.
 - (5) The predicted values ~~are~~ of y are uncorrelated with the residuals.

$$X\hat{\beta} = \hat{y} = X\hat{\beta}$$

$$\hat{y}'e = (X\hat{\beta})'e = b'X'e = 0 \rightarrow (16)$$

$$y = X\beta + e$$

↳ Assumption. States that there is a linear relationship between y and x .

$$X = n \times k$$

X - linear independent

$$E[e|X] = 0$$

$$E \begin{bmatrix} e_1|X \\ e_2|X \\ \vdots \\ e_n|X \end{bmatrix} = \begin{bmatrix} E(e_1) \\ E(e_2) \\ \vdots \\ E(e_n) \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix} \rightarrow (17)$$

$$E(\mathbf{e}\mathbf{e}'|X) = \sigma^2 \mathbf{I}$$

$$E(\mathbf{e}\mathbf{e}'|X) = E \begin{bmatrix} e_1|X \\ e_2|X \\ \vdots \\ e_n|X \end{bmatrix} [e_1|X \ e_2|X \ \dots \ e_n|X] \quad \rightarrow (19)$$

$$E(\mathbf{e}\mathbf{e}'|X) = E \begin{bmatrix} e_1^2|X & e_1 e_2|X & \dots & e_1 e_n|X \\ e_2 e_1|X & e_2^2|X & \dots & e_2 e_n|X \\ \vdots & \vdots & \ddots & \vdots \\ e_k e_1|X & e_k e_2|X & \dots & e_k e_n|X \end{bmatrix}$$

$$E(\mathbf{e}\mathbf{e}'|X) = \begin{bmatrix} \sigma^2 & 0 & \dots & 0 \\ 0 & \sigma^2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \sigma^2 \end{bmatrix} \quad \rightarrow (21)$$

$$E(\mathbf{e}\mathbf{e}'|X) = \sigma^2 \begin{bmatrix} 1 & 0 & \dots & 0 \\ 0 & 1 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & 1 \end{bmatrix} = \sigma^2 \mathbf{I} \quad \rightarrow (22)$$

$$\Omega = \sigma^2 \mathbf{I}$$

↳ Variance-covariance matrix.