

INFO 510 Bayesian Modelling and Inference

Lecture 5 – Conjugate priors and uninformative priors

Dr. Kunal Arekar
College of Information
University of Arizona, Tucson

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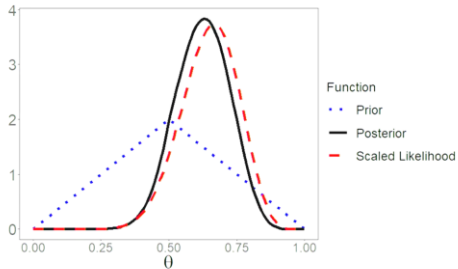
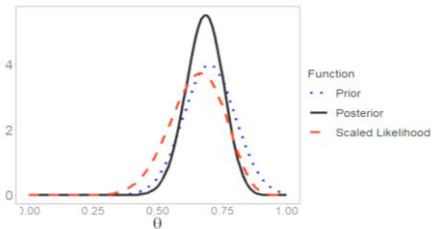
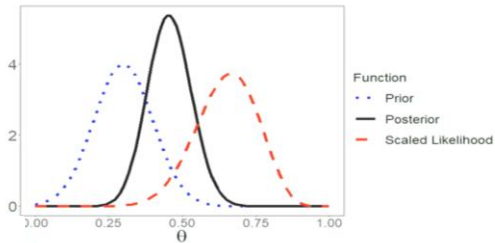
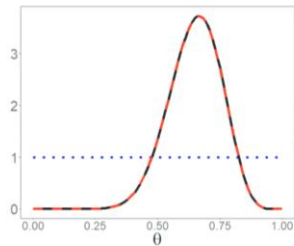
■ Conjugacy

- Conjugate priors
- Mixture of conjugate prior

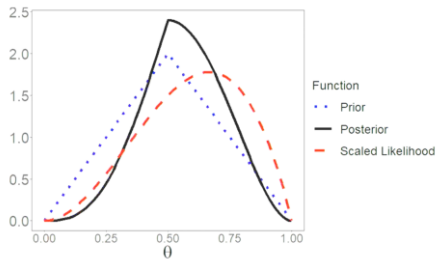
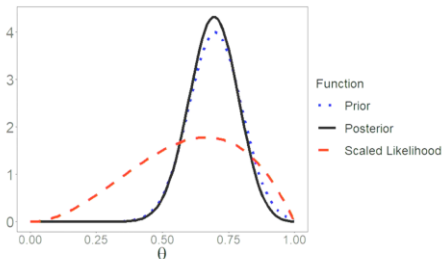
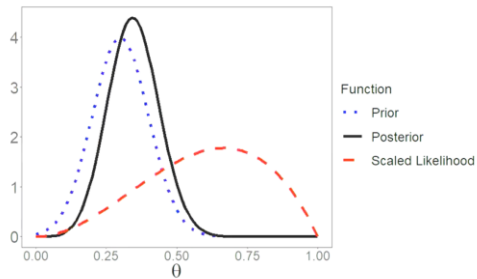
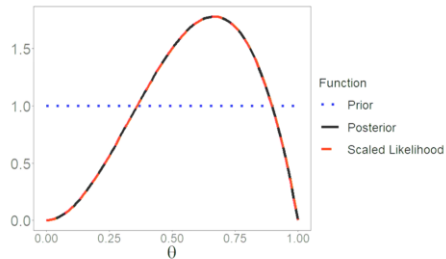
■ Uninformative priors

- Jeffreys prior

Different priors good data



Different priors bad data



Conjugacy

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- If the posterior distribution $p(\theta|X)$ are in the same family as the prior probability distribution $p(\theta)$, the prior and posterior are then called conjugate distributions, and the prior is called a conjugate prior for the likelihood function $p(X|\theta)$.
- All members of the exponential family have conjugate priors.

Brief List of Conjugate Models

Likelihood	Prior	Posterior
Binomial	Beta	Beta
Negative Binomial	Beta	Beta
Poisson	Gamma	Gamma
Geometric	Beta	Beta
Exponential	Gamma	Gamma
Normal (mean unknown)	Normal	Normal
Normal (variance unknown)	Inverse Gamma	Inverse Gamma
Normal (mean and variance unknown)	Normal/Gamma	Normal/Gamma
Multinomial	Dirichlet	Dirichlet

The Conjugate Beta Prior

- The Beta distribution is conjugate to the Binomial distribution.

$$\begin{aligned} p(\theta|x) &= p(x|\theta)p(\theta) = \text{Binomial}(n, \theta) * \text{Beta}(a, b) = \\ &\binom{n}{x} \theta^x (1 - \theta)^{n-x} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{(a-1)} (1 - \theta)^{b-1} \\ &\propto \theta^x (1 - \theta)^{n-x} \theta^{(a-1)} (1 - \theta)^{b-1} \end{aligned} \tag{9}$$

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- The posterior distribution is simply a $\text{Beta}(x + a, n - x + b)$ distribution.
- Effectively, our prior is just adding $a - 1$ successes and $b - 1$ failures to the dataset.

Coin Flipping Example: Model

- Use a **Bernoulli likelihood** for coin X landing ‘heads’,

$$p(X = 'H'|\theta) = \theta, \quad p(X = 'T'|\theta) = 1 - \theta,$$

$$p(X|\theta) = \theta^{I(X='H')} (1 - \theta)^{I(X='T')}.$$

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$$p(\theta|a, b) = \frac{\theta^{a-1}(1-\theta)^{b-1}}{B(a, b)} \propto \theta^{a-1}(1-\theta)^{b-1}$$

$$\text{with } B(a, b) = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}.$$

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- Remember that probabilities sum to one so we have

$$1 = \int_0^1 p(\theta|a, b) d\theta = \int_0^1 \frac{\theta^{a-1}(1-\theta)^{b-1}}{B(a, b)} d\theta = \frac{1}{B(a, b)} \int_0^1 \theta^{a-1}(1-\theta)^{b-1} d\theta$$

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which helps us compute integrals since we have

$$\int_0^1 \theta^{a-1}(1-\theta)^{b-1} d\theta = B(a, b).$$

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- If we observe 'HHH' then our **posterior** distribution is

$$\begin{aligned} p(\theta|HHH) &= \frac{p(HHH|\theta)p(\theta)}{p(HHH)} && \text{(Bayes' rule)} \\ &\propto p(HHH|\theta)p(\theta) && (p(HHH) \text{ is constant}) \\ &= \theta^3(1-\theta)^0p(\theta) && \text{(likelihood def'n)} \\ &= \theta^3(1-\theta)^0\theta^{a-1}(1-\theta)^{b-1} && \text{(prior def'n)} \\ &= \theta^{3+a-1}(1-\theta)^{b-1}. \end{aligned}$$

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- Which we've written in the form of a **Beta** distribution,

$$\theta|HHH \sim \text{Beta}(3+a, b),$$

which let's us skip computing the integral $p(HHH)$.

Coin Flipping Example: Estimates

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- Posterior predictive,

$$\begin{aligned} p(H|HHH) &= \int_0^1 p(H|\theta)p(\theta|HHH)d\theta \\ &= \int_0^1 \text{Ber}(H|\theta)\text{Beta}(\theta|3 + a, b)d\theta \\ &= \int_0^1 \theta\text{Beta}(\theta|3 + a, b)d\theta = \mathbb{E}[\theta] \\ &= \frac{(3 + a)}{(3 + a) + b} = \frac{4}{5} = 0.8. \end{aligned}$$

Coin Flipping Example: Effect of Prior

- We assume all θ equally likely and saw HHH,
 - ▣ ML predict it will always land heads.
 - ▣ Bayes predict probability of landing heads is only 80%.
 - ▣ Takes into account other ways that HHH could happen.

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 - ❑ Posterior predictive would be $\frac{3+1}{3+1+1} = 0.80$.
- Beta(3, 3) prior is like seeing 3 heads and 3 tails (stronger uniform prior),
 - ❑ Posterior predictive would be $\frac{3+3}{3+3+3} = 0.667$.
- Beta(100, 1) prior is like seeing 100 heads and 1 tail (biased),
 - ❑ Posterior predictive would be $\frac{3+100}{3+100+1} = 0.990$.
- Beta(0.01, 0.01) biases towards having unfair coin (head or tail),
 - ❑ Posterior predictive would be $\frac{3+0.01}{3+0.01+0.01} = 0.997$.

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 - ❑ Posterior predictive would be $\frac{3+0.01}{3+0.01+0.01} = 0.997$.
- Dependence on (a, b) is where people get uncomfortable:
 - ❑ But basically the same as choosing regularization parameter λ .
 - ❑ If your prior knowledge isn't misleading, you **will not overfit**.

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- For example, suppose we are modelling coin tosses, and we think the coin is either fair, or is biased towards heads. This cannot be represented by a Beta distribution. However, we can model it using a mixture of two Beta distributions. For example, we might use:

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- If θ comes from the first distribution, the coin is fair, but if it comes from the second it is biased towards heads.

Mixtures of Conjugate Priors (Cont.)

- The prior has the form

$$p(\theta) = \sum_k p(z = k)p(\theta|z = k) \quad (12)$$

where $z = k$ means that θ comes from mixture component k ,
 $p(z = k)$ are called the **prior mixing weights**, and each $p(\theta|z = k)$
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Posterior can also be written as a mixture of conjugate

- distributions as follows:

$$p(\theta|X) = \sum_k p(z = k|X)p(\theta|X, z = k) \quad (13)$$

where $p(z = k|X)$ are the **posterior mixing weights** given by

$$p(z = k|X) = \frac{p(z = k)p(X|z = k)}{\sum_{k'} p(z = k'|X)p(\theta|X, z = k')} \quad (14)$$

Uninformative Priors

- If we don't have strong beliefs about what θ should be, it is common to use an **uninformative** or **non-informative** prior, and to **let the data speak for itself**.
- Designing uninformative priors is tricky.

Uninformative Prior for the Bernoulli

- Consider the Bernoulli parameter

$$p(x|\theta) = \theta^x (1 - \theta)^{n-x} \quad (15)$$

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- What is the most uninformative prior for this distribution?

An **uninformative prior** should:

Influence the outcome as little as possible

Reflect minimal knowledge about θ

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- By decreasing the magnitude of the pseudo counts, we can lessen the impact of the prior. By this argument, the most uninformative prior is $Beta(0, 0)$, which is called **Haldane prior**.

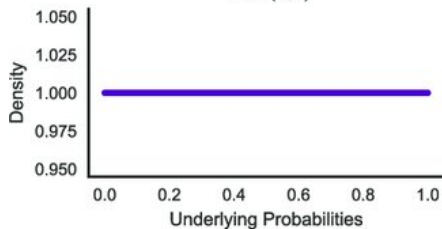
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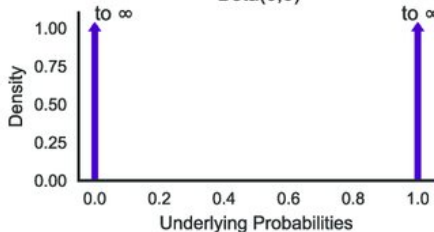
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 - **Haldane prior** is an **improper** prior; it does not integrate to 1.
 - **Haldane prior** results in the posterior $Beta(x, n - x)$ which will be proper as long as $n - x \neq 0$ and $x \neq 0$.

Bayes' Prior
Beta(1,1)



Haldane's Prior
Beta(0,0)



Haldane Prior: Beta(0,0)

$$\theta|D \sim \text{Beta}(N_1, N_0)$$

$$E[\theta|D] = N_1 / (N_1 + N_0)$$

which is the MLE estimate

if $N_1 = 0$ or $N_0 = 0$, then the posterior is:
Beta(0, N_0) or Beta(N_1 , 0) \Rightarrow not normalizable!

Posterior mean is undefined.

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 - **Haldane prior** results in the posterior $Beta(x, n - x)$ which will be proper as long as $n - x \neq 0$ and $x \neq 0$.
- We will see that the "right" uninformative prior is **Beta(1, 1)**.

Jeffreys Prior

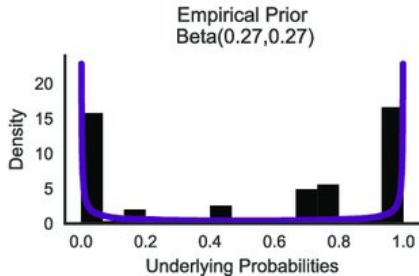
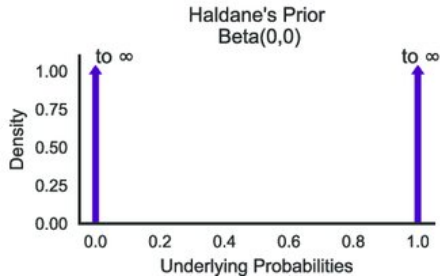
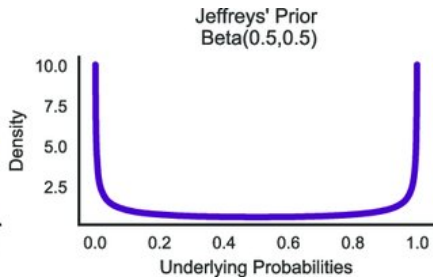
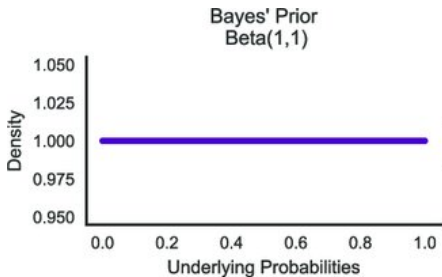
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Imagine you're describing a shape, like a circle. Whether you describe it using its radius (r) or its diameter ($d=2r$), your lack of knowledge about the circle's size should remain the same

Similarly, if we change the parameter representation of θ to $\phi=h(\theta)$ using a transformation h , our prior $p(\theta)$ should adjust so that it remains "uninformative" in both representations

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- **Jeffreys prior** is the prior that satisfies $p(\theta) \propto \sqrt{\det I(\theta)}$, where $I(\theta)$ is the **Fisher information** for θ , and is **invariant under reparametrization** of the parameter vector θ .



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$$I(\theta) = E\left[\left(\frac{\partial}{\partial \theta} \log f(X; \theta)\right)^2 \middle| \theta\right] \quad (16)$$

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$$I(\theta) = E\left[\left(\frac{\partial}{\partial \theta} \log f(X; \theta)\right)^2 \middle| \theta\right] \quad (16)$$

- If $\log f(X; \theta)$ is twice differentiable with respect to θ , then

$$I(\theta) = -E\left[\frac{\partial^2}{\partial \theta^2} \log f(X; \theta) \middle| \theta\right] \quad (17)$$

Reparametrization for Jeffreys Prior: One parameter case

- For an alternative parametrization ϕ we can derive $p(\phi) = \sqrt{I(\phi)}$ from $p(\theta) = \sqrt{I(\theta)}$, using the **change of variables theorem** and the definition of **Fisher information**:

$$\begin{aligned} p(\phi) &= p(\theta) \left| \frac{d\theta}{d\phi} \right| \propto \sqrt{I(\theta) \left(\frac{d\theta}{d\phi} \right)^2} = \sqrt{E \left[\left(\frac{d \ln L}{d\theta} \right)^2 \right] \left(\frac{d\theta}{d\phi} \right)^2} \\ &= \sqrt{E \left[\left(\frac{d \ln L}{d\theta} \frac{d\theta}{d\phi} \right)^2 \right]} = \sqrt{E \left[\left(\frac{d \ln L}{d\phi} \right)^2 \right]} = \sqrt{I(\phi)} \end{aligned} \tag{18}$$

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- The **Fisher information** is the expected information

$$I(\theta) = E[J(\theta|X)|X \sim \theta] = \frac{\theta}{\theta^2} + \frac{1-\theta}{(1-\theta)^2} = \frac{1}{\theta(1-\theta)} \quad (22)$$

- Hence Jeffreys prior is

$$p(\theta) \propto \theta^{-\frac{1}{2}} (1-\theta)^{-\frac{1}{2}} \propto \text{Beta}\left(\frac{1}{2}, \frac{1}{2}\right). \quad (23)$$

What NOT to do when considering priors

Do NOT choose a prior that assigns 0 probability/density to possible values of the parameter

Do NOT base the prior on the observed data

Do NOT feel like you have to find that one, perfect prior

Do NOT worry too much about the prior!