ISTA410/INFO510 Bayesian Modelling and Inference

Lecture 4 – Independence, Distributions, Continuity

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Basic rules (so far)

Marginalization

$$P(A) = \sum_{b \in B} P(A, B)$$

Conditional probability (definition)

$$P(A|B) \equiv \frac{P(A \cap B)}{P(B)}$$

Chain (Product) Rule

$$\begin{split} P\big(A_1 \cap A_2\big) &= P(A_1)P(A_2 \big| A_1) \\ P\big(A_1 \cap A_2 \cap \dots A_N\big) &= P(A_1)P(A_2 \big| A_1)P(A_3 \big| A_1 \cap A_2) \dots P(A_N \big| A_1 \cap A_2 \cap \dots A_{N-1}) \end{split}$$

Bayes Rule

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

Normalization

$$P(X = x) = \frac{p(x)}{\sum_{x} p(x)} *$$

Structure in relations between RVs:

(Absolute, Marginal) Independence and Conditional Independence

In probability and statistics, understanding the structure in relationships between random variables (RVs) is crucial

These relationships determine how information flows between variables, how dependencies arise, and how we can simplify complex probability models

First, recall the product rule (which holds generally): P(X, Y) = P(X)P(Y|X)

$$P(A_1 \cap A_2) = P(A_1)P(A_2 | A_1)$$

$$P(A_1 \cap A_2 \cap \dots A_N) = P(A_1)P(A_2 | A_1)P(A_3 | A_1 \cap A_2) \dots P(A_N | A_1 \cap A_2 \cap \dots A_{N-1})$$

expresses the probability of a sequence of dependent events in terms of conditional probabilities

First, recall the product rule (which holds generally):

$$P(X, Y) = P(X)P(Y|X)$$

"X is independent of Y"

$$X \perp Y \Leftrightarrow P(X,Y) = P(X)P(Y)$$

$$X \perp Y \Leftrightarrow P(X|Y) = P(X) \text{ or } P(Y)=0$$

Some authors use ⊥ instead of ⊥

First, recall the product rule (which holds generally):

$$P(X, Y) = P(X)P(Y|X)$$

This can cause confusion. If P(Y) is zero, then P(X|Y) is ill-defined (divide by zero). Since if P(Y)=0, Y never happens, we have a choice to declare whether (a) X is independent from Y or (b) not. Under scrutiny, (a) make sense, and allows consistency with the first definition. The first formula works in this (weird) case because if P(Y)=0, then P(X,Y) is also 0.

"X is independent of Y"

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D.	T	117)
r_1	(I,	W)

Т	W	Р
warm	sun	0.4
warm	rain	0.1
cold	sun	0.2
cold	rain	0.3

P(T)		
T	Р	
warm	0.5	
cold	0.5	

P(W)

W	Р
sun	0.6
rain	0.4

"X is independent of Y"

$$X \perp Y \Leftrightarrow P(X,Y) = P(X)P(Y)$$
 $X \perp Y \Leftrightarrow P(X|Y) = P(X) \text{ or } P(Y)=0$

$$P_1(T, W)$$

Т	W	Р
warm	sun	0.4
warm	rain	0.1
cold	sun	0.2
cold	rain	0.3

P(T)		
T	Р	
warm	0.5	
cold	0.5	

W	P
sun	0.6
rain	0.4

 $P_2(T,W)$

Т	W	Р
warm	sun	0.3
warm	rain	0.2
cold	sun	0.3
cold	rain	0.2

Conditional Independence

$$X \perp Y \mid Z$$
 \Leftrightarrow $P(X \mid Y, Z) = P(X \mid Z)$ or $P(Y,Z)=0$ *
$$P(Y \mid X, Z) = P(Y \mid Z)$$

Conditional Independence

$$X \perp Y \mid Z$$
 \Leftrightarrow $P(X \mid Y, Z) = P(X \mid Z)$ or $P(Y, Z) = 0$ \Rightarrow $P(Y \mid X, Z) = P(Y \mid Z)$

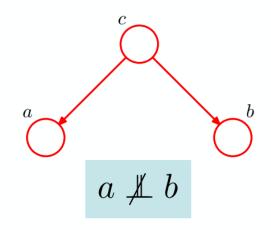
Equivalent, sometimes more convenient definition

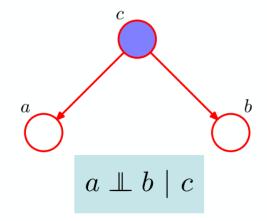
$$X \perp Y \mid Z \iff P(X,Y \mid Z) = P(X \mid Z)P(Y \mid Z)$$

Conditional Independence

Equivalent, sometimes more convenient definition

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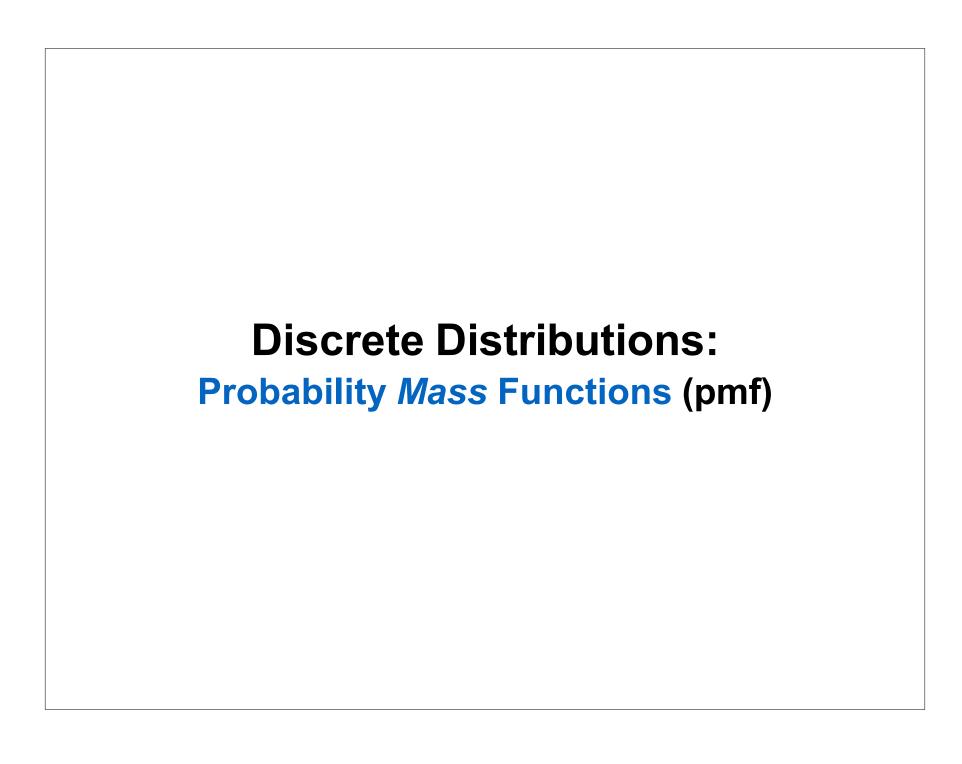


$$p(a, b, c) = p(c)p(a|c)p(b|c)$$
$$p(a, b|c) = p(a|c)p(b|c)$$

This is the definition of $a \perp \!\!\! \perp b \mid c$

(what the graph represents in general)

(right side: with c observed, p(c) = 1.0; left side: we're *conditioning* on c)

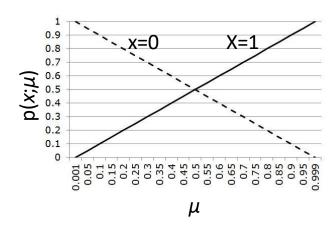


Bernoulli Distribution

$$x \in \{0,1\}$$
 (e.g., 1 is "heads" and 0 is "tails")

$$p(x = 1 | \mu) = \mu$$
 and $p(x = 0 | \mu) = 1 - \mu$

$$p(x|u) = \begin{cases} u & \text{if } x=1\\ 1-u & \text{if } x=0 \end{cases}$$

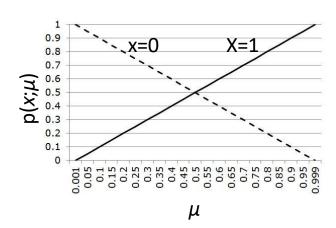


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$$Bern(x|\mu) = \mu^{x}(1 - \mu)^{(1-x)}$$



Binomial Distribution

Probability distribution for getting *m* "heads" out of *N* tosses.

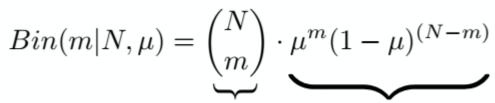
$$Bin(m|N,\mu) = \begin{pmatrix} N \\ m \end{pmatrix} \cdot \mu^m (1-\mu)^{(N-m)}$$
 Number of ways to get m heads in N tosses N tosses

Example event N=3, m=2
HHT
HTH
THH

where
$$\binom{N}{m} = \frac{N!}{(N-m)!m!}$$

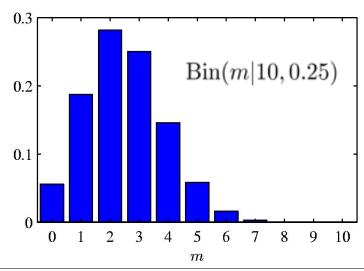
Binomial Distribution

Probability distribution for getting *m* "heads" out of *N* tosses.



Number of ways to get *m* heads in N tosses

Probability of **each** way to get *m* heads in *N* tosses



where
$$\binom{N}{m} = \frac{N!}{(N-m)!m!}$$

Categorical Distribution

- Simple extensions to Bernoulli to multiple outcomes (e.g., a six sided die)
- Let *K* be the number of possible outcomes (e.g., for six sided die, *K*=6)
- Now we use vectors for μ and x, i.e., μ and x.
- **x** is a vector of 0's and exactly one 1 for observed outcomes (e.g., rolling 3 with a six sided die is (0,0,1,0,0,0)).
- μ is a vector of real values representing the distribution over which element of **x** will be value 1.

$$p(\mathbf{x}|\boldsymbol{\mu}) = \prod_{k=1}^{K} \boldsymbol{\mu}_k^{x_k}$$
 (note that $\sum_{k=1}^{K} \boldsymbol{\mu}_k = 1$)

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Multinomial Distribution

Extension of binomial to multiple outcomes.

Let K be the number of outcomes, with probabilities μ .

 $Mult(m_1, m_2, ..., m_K \mu, N)$ denotes the probability of seeing m_1 of outcome 1, m_2 of outcome 2, and so on.

Multinomial Distribution

Extension of binomial to multiple outcomes.

Let K be the number of outcomes, with probabilities μ .

$$Mult(m_1, m_2, ..., m_K | \boldsymbol{\mu}, N) = \begin{pmatrix} N \\ m_1 & m_2 & ... & m_K \end{pmatrix} \prod_{k=1}^K \mu_k^{m_k}$$

where
$$\begin{pmatrix} N \\ m_1 & m_2 & \dots & m_K \end{pmatrix} = \begin{pmatrix} N! \\ \hline m_1! & m_2! & \dots & m_K! \end{pmatrix}$$

and
$$\sum_{k=1}^{K} m_k = N$$



Continuous Spaces

- Outcome space is observation of a real value
 - E.g., person's height, temperature outside, etc...
- Example: a random variable X that can take on any value in [0,1] with equal probability
 - We say that X is uniformly distributed.
- Challenge: P(X = x) = 0
- Instead of assigning probabilities to points, we instead assign probabilities to regions (within some range or interval):

$$P(x_1 < X < x_2)$$
 but not $P(X = x)$

Continuous Spaces

For a Uniform Distribution on [0,1]

The **PDF** is f(x)=1 for $0 \le x \le 1$

The probability of picking a number in some subrange [x1,x2] is given by

$$P(x_1 \leq X \leq x_2) = \int_{x_1}^{x_2} p(x) \, dx \qquad \begin{array}{l} \text{Here,} \\ \text{p(x)=1} \\ \text{dx = x}_2 - \text{x}_1 \end{array}$$

$$P(0.2 \le x \le 0.5) = 0.5 - 0.2 = 0.3$$

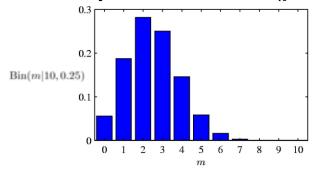
To compute the probability that X lies in some range, we compute the definite integral of the density function:

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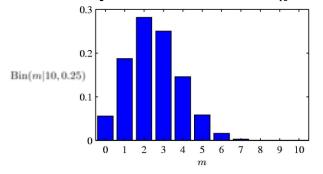
With discrete outcome spaces... Probability **Mass** Function (**pmf**)



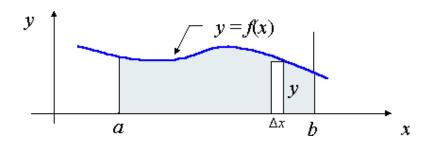
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Probability **Density** Function (**pdf**)

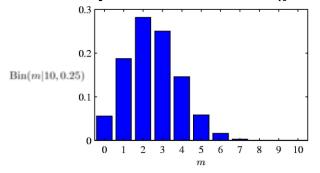


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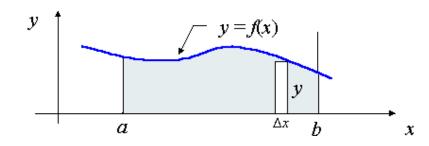
$$P(x_1 \le X \le x_2) = \int_{x_1}^{x_2} p(x) \, dx$$

$$\rho = \frac{m}{V}_{\text{density}}$$

With discrete outcome spaces... Probability **Mass** Function (**pmf**)



Probability **Density** Function (pdf)



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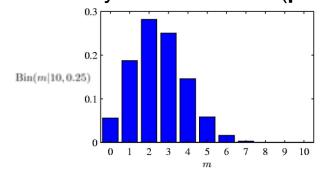
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mass

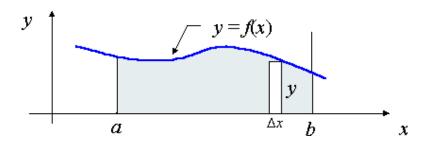
density volume

$$\rho = \frac{m}{V}_{\text{volume}}$$

With discrete outcome spaces... Probability **Mass** Function (**pmf**)



Probability **Density** Function (pdf)



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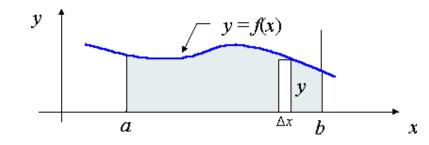
mass

density volume

$$p: \mathbb{R} \mapsto \mathbb{R}$$
$$p(x) \ge 0$$
$$\int_{-\infty}^{\infty} p(x) \, dx = 1$$

Note: p(x) can be > 1!

Probability **Density** Function (**pdf**)



The actual probability of an event is obtained by integrating the PDF over a specific interval

To compute the probability that X lies in some range, we compute the definite integral of the density function:

$$P(x_1 \le X \le x_2) = \int_{x_1}^{x_2} p(x) \, dx$$

mass

density volume

$$p(x) = p(x_0, x_1, ..., x_n)$$
 Probability vector is just a joint probability!

Joint
$$P(x_1 \le X \le x_2, y_1 \le Y \le y_x) = \int_{x=x_1}^{x_2} \int_{y=y_1}^{y_2} p(x, y) dx dy$$

$$\text{Conditional} \qquad P(x_1 \leq X \leq x_2, Y = y) = \int_{x=x_1}^{x_2} p(x|Y = y) \, dx$$

Marginalization
$$P(y) = \int_{x=x_1}^{x_2} p(x,y) dx$$
 (where $x_1 \le X \le x_2$ describes the sample space of X)