Stochastic Processes

Homework 5

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03/29/2017

Problem 1

- X_n is our Markov chain with state space $S = \{(S,S), (S,C), (C,S), (C,C)\}$
- X_0 corresponds to day 0 and 1 where it is given that $X_0 = (S,S)$
- X₁ corresponds to day 2 and 3, and X₂ corresponds to day 4 and 5
- We are interested in day 5 being sunny which means day 4 can be sunny or cloudy, therefore by using the law of total probability, this state of day 5 being sunny can be computed as:

$$Pr(X_2 = (S,S) | X_0 = (S,S)) + Pr(X_2 = (C,S) | X_0 = (S,S))$$

• Using n-step transition probability for a Markov chain, we can rewrite the above probability as:

$$P_{(S,S),(S,S)}^{(2)} + P_{(S,S),(C,S)}^{(2)}$$
; where P is the given transition probability matrix

• P⁽²⁾ is computed by P * P:

$$P^{(2)} = \begin{matrix} (S,S) & (S,C) & (C,S) & (C,C) \\ (S,S) & \begin{bmatrix} 0.49 & 0.21 & 0.12 & 0.18 \\ 0.20 & 0.20 & 0.12 & 0.48 \\ 0.35 & 0.15 & 0.20 & 0.30 \\ 0.10 & 0.10 & 0.16 & 0.64 \end{matrix}$$

• Finally, the probability that it is sunny on day 5 is:

$$P_{(S,S),(S,S)}^{(2)} + P_{(S,S),(C,S)}^{(2)} = 0.49 + 0.12 = 0.61$$

- The long run fraction of sunny days is computed by the limiting distribution for sunny days, where sunny days correspond to the states (S,S) and (C,S)
- So, we are looking for $\pi_{(S,S)} + \pi_{(C,S)}$, where π_j is the initial probability of state $j \in S$
- We can compute the entire limiting distribution of X_n by solving the following system of equations:

$$\pi_{j} = \sum_{k \in \mathbb{S}} \pi_{k} P_{kj} \quad \forall j \in \mathbb{S}$$

$$1 = \sum_{k \in \mathbb{S}} \pi_{k}$$

Applying the above equations, we have the following system of equations:

$$\begin{split} \pi_{(S,S)} &= 0.7\pi_{(S,S)} + \ 0.0\pi_{(S,C)} + \ 0.5\pi_{(C,S)} + \ 0.0\pi_{(C,C)} \\ \pi_{(S,C)} &= 0.3\pi_{(S,S)} + \ 0.0\pi_{(S,C)} + \ 0.5\pi_{(C,S)} + \ 0.0\pi_{(C,C)} \\ \pi_{(C,S)} &= 0.0\pi_{(S,S)} + \ 0.4\pi_{(S,C)} + \ 0.0\pi_{(C,S)} + \ 0.2\pi_{(C,C)} \\ \pi_{(C,C)} &= 0.0\pi_{(S,S)} + \ 0.6\pi_{(S,C)} + \ 0.0\pi_{(C,S)} + \ 0.8\pi_{(C,C)} \\ 1 &= \pi_{(S,S)} + \ \pi_{(S,C)} + \ \pi_{(C,S)} + \ \pi_{(C,C)} \end{split}$$

- Solving the above system will yield: $(\pi_{(S,S)}, \pi_{(S,C)}, \pi_{(C,S)}, \pi_{(C,C)}) = (0.25, 0.15, 0.15, 0.45)$
- Therefore, the long run fraction of sunny days is $\pi_{(S,S)} + \pi_{(C,S)} = 0.25 + 0.15 = 0.40$

Problem 2

- X_n is our Markov chain with state space $S = \{exam 1, exam 2, exam 3\} = \{1, 2, 3\}$ where X_n represents the type of the n^{th} exam.
 - \circ X_n is a Markov chain because the probability of X_n taking a value in S is dependent only on what the previous exam was, X_{n-1}, and how the class performed on X_{n-1}
- We can build the transition probability matrix P by using conditional probability:

$$P_{ij} = Pr(X_n = j \mid X_{n-1} = i) \forall i, j \in \mathbb{S}$$

• By using the law of total probability and conditioning on the class' performance, we can rewrite Pij:

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P_{ij} = Pr(X_n = j \mid X_{n-1} = i)
   = Pr(X_n = j \mid class \ did \ well \ on \ X_{n-1} = i) + Pr(X_n = j \mid class \ did \ bad \ on \ X_{n-1} = i)
   = Pr(X_n = j \mid class \ did \ well \ on \ X_{n-1} = i) * Pr(class \ did \ well \ on \ X_{n-1} = i) +
       Pr(X_n = j \mid class \ did \ bad \ on \ X_{n-1} = i) * Pr(class \ did \ bad \ on \ X_{n-1} = i)
P_{11} = Pr(X_n = 1 \mid class \ did \ well \ on \ X_{n-1} = 1) * Pr(class \ did \ well \ on \ X_{n-1} = 1) +
       Pr(X_n = 1 \mid class \ did \ bad \ on \ X_{n-1} = 1) * Pr(class \ did \ bad \ on \ X_{n-1} = 1)
    = (1/3)*(0.3) + (1)*(1 - 0.3) = 0.8
P_{12} = Pr(X_n = 2 \mid class \ did \ well \ on \ X_{n-1} = 1) * Pr(class \ did \ well \ on \ X_{n-1} = 1) +
       Pr(X_n = 2 \mid class \ did \ bad \ on \ X_{n-1} = 1) * Pr(class \ did \ bad \ on \ X_{n-1} = 1)
     = (1/3)*(0.3) + (0)*(1 - 0.3) = 0.1
P_{13} = Pr(X_n = 3 \mid class \ did \ well \ on \ X_{n-1} = 1) * Pr(class \ did \ well \ on \ X_{n-1} = 1) +
       Pr(X_n = 3 \mid class \ did \ bad \ on \ X_{n-1} = 1) * Pr(class \ did \ bad \ on \ X_{n-1} = 1)
     = (1/3)*(0.3) + (0)*(1 - 0.3) = 0.1
P_{21} = Pr(X_n = 1 \mid class \ did \ well \ on \ X_{n-1} = 2) * Pr(class \ did \ well \ on \ X_{n-1} = 2) +
       Pr(X_n = 1 \mid class \ did \ bad \ on \ X_{n-1} = 2) * Pr(class \ did \ bad \ on \ X_{n-1} = 2)
     = (1/3)*(0.6) + (1)*(1 - 0.6) = 0.6
P_{22} = Pr(X_n = 2 \mid class \ did \ well \ on \ X_{n-1} = 2) * Pr(class \ did \ well \ on \ X_{n-1} = 2) +
       Pr(X_n = 2 \mid class \ did \ bad \ on \ X_{n-1} = 2) * Pr(class \ did \ bad \ on \ X_{n-1} = 2)
     = (1/3)*(0.6) + (0)*(1 - 0.6) = 0.2
P_{23} = Pr(X_n = 3 \mid class \ did \ well \ on \ X_{n-1} = 2) * Pr(class \ did \ well \ on \ X_{n-1} = 2) +
       Pr(X_n = 3 \mid class \ did \ bad \ on \ X_{n-1} = 2) * Pr(class \ did \ bad \ on \ X_{n-1} = 2)
     = (1/3)*(0.6) + (0)*(1 - 0.6) = 0.2
P_{31} = Pr(X_n = 1 \mid class \ did \ well \ on \ X_{n-1} = 3) * Pr(class \ did \ well \ on \ X_{n-1} = 3) +
       Pr(X_n = 1 \mid class \ did \ bad \ on \ X_{n-1} = 3) * Pr(class \ did \ bad \ on \ X_{n-1} = 3)
     = (1/3)*(0.9) + (1)*(1 - 0.9) = 0.4
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Stochastic Processes

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$$\begin{split} P_{32} &= Pr(X_n = 2 \mid class \ did \ well \ on \ X_{n\text{-}1} = 3) * Pr(class \ did \ well \ on \ X_{n\text{-}1} = 3) + \\ ⪻(X_n = 2 \mid class \ did \ bad \ on \ X_{n\text{-}1} = 3) * Pr(class \ did \ bad \ on \ X_{n\text{-}1} = 3) \\ &= (1/3)^*(0.9) + (0)^*(1 - 0.9) = 0.3 \end{split}$$

$$\begin{split} P_{33} &= \text{Pr}(X_n = 3 \mid \text{class did well on } X_{n\text{-}1} = 3) * \text{Pr}(\text{class did well on } X_{n\text{-}1} = 3) + \\ &\quad \text{Pr}(X_n = 3 \mid \text{class did bad on } X_{n\text{-}1} = 3) * \text{Pr}(\text{class did bad on } X_{n\text{-}1} = 3) \\ &= (1/3)^*(0.9) + (0)^*(1 - 0.9) = 0.3 \end{split}$$

$$P = \begin{bmatrix} 1 & 2 & 3 \\ 0.8 & 0.1 & 0.1 \\ 2 & 0.6 & 0.2 & 0.2 \\ 3 & 0.4 & 0.3 & 0.3 \end{bmatrix}$$

- The proportion of exams being type i in the long run, where $i \in S$, is represented by the entire limiting distribution of X_n
- We can compute the entire limiting distribution of X_n by solving the following system of equations:

$$\begin{split} &\pi_j = \sum_{k \,\in \,\mathbb{S}} \, \pi_k P_{kj} \quad \forall \; j \in \mathbb{S} \; ; \text{where} \; \pi_j \text{ is the initial probability of state} \; j \in \mathbb{S} \\ &1 = \; \sum_{k \,\in \,\mathbb{S}} \pi_k \end{split}$$

• Applying the above equations, we have the following system of equations:

$$\begin{split} \pi_1 &= 0.8\pi_1 + \ 0.6\pi_2 + \ 0.4\pi_3 \\ \pi_2 &= 0.1\pi_1 + \ 0.2\pi_2 + \ 0.3\pi_3 \\ \pi_3 &= 0.1\pi_1 + \ 0.2\pi_2 + \ 0.3\pi_3 \\ 1 &= \pi_1 + \ \pi_2 + \ \pi_3 \end{split}$$

- Solving the above system will yield: $(\pi_1, \pi_2, \pi_3) = (5/7, 1/7, 1/7)$
- Therefore, the proportion of exams being type 1, 2, and 3 in the long run is 5/7, 1/7, and 1/7 respectively

Problem 3

- X_n is our Markov chain with state space $S = \{sunny, cloudy, rainy\} = \{0, 1, 2\}$ where X_n represents the weather on day n.
 - \circ X_n is a Markov chain because the probability of X_n taking a value in S is dependent only on the value that was taken by X_{n-1}, the weather of the previous day.
- We can build the transition probability matrix P by using conditional probability:

$$\begin{array}{c} P_{ij} = Pr(X_n = j \mid X_{n - 1} = i) \; \forall \; i, j \in \mathbb{S} \\ P_{00} = Pr(X_n = 0 \mid X_{n - 1} = 0) = 0 \\ P_{01} = Pr(X_n = 1 \mid X_{n - 1} = 0) = 1/2 \\ P_{02} = Pr(X_n = 2 \mid X_{n - 1} = 0) = 1/2 \\ P_{10} = Pr(X_n = 0 \mid X_{n - 1} = 1) = 1/4 \\ P_{10} = Pr(X_n = 0 \mid X_{n - 1} = 1) = 1/4 \\ P_{11} = Pr(X_n = 1 \mid X_{n - 1} = 1) = 1/2 \\ P = \begin{pmatrix} 0 & 1 & 2 \\ 1 & 1/4 & 1/2 & 1/4 \\ 2 & 1/4 & 1/4 & 1/2 \end{pmatrix} \\ P = \begin{pmatrix} 0 & 1/2 & 1/2 \\ 1/4 & 1/2 & 1/4 \\ 1/4 & 1/4 & 1/2 \end{pmatrix} \end{array}$$

- The proportion of days being sunny in the long run, and the proportion of days being cloudy in the long run can both be found in entire limiting distribution of X_n
- We are interested in the values of π_0 and π_1 to find the proportions of interest for sunny days and cloudy days respectively, where π_i is the initial probability of state $j \in \mathbb{S}$
- We can compute the entire limiting distribution of X_n by solving the following system of equations:

$$\begin{split} & \pi_j = \sum_{k \in \mathbb{S}} \pi_k P_{kj} & \forall j \in \mathbb{S} \\ & 1 = \sum_{k \in \mathbb{S}} \pi_k \end{split}$$

• Applying the above equations, we have the following system of equations:

$$\begin{split} \pi_0 &= 0\pi_0 + \frac{1}{4}\pi_1 + \frac{1}{4}\pi_2 \\ \pi_1 &= \frac{1}{2}\pi_0 + \frac{1}{2}\pi_1 + \frac{1}{4}\pi_2 \\ \pi_2 &= \frac{1}{2}\pi_0 + \frac{1}{4}\pi_1 + \frac{1}{2}\pi_2 \\ 1 &= \pi_1 + \pi_2 + \pi_3 \end{split}$$

- Solving the above system will yield: $(\pi_0, \pi_1, \pi_2) = (1/5, 2/5, 2/5)$
- Therefore, the proportion of days being sunny and cloudy in the long run is 1/5 and 2/5 respectively

Problem 4

• Using conditional probability, we can rewrite the given limit as:

$$\lim_{n \to \infty} \Pr(X_{n-1} = 2 \mid X_n = 1) = \\ \lim_{n \to \infty} \frac{\Pr(X_{n} = 1 \mid X_{n-1} = 2) * \Pr(X_{n-1} = 2)}{\Pr(X_{n} = 1)}$$

• Since we are taking the limit of $n \to \infty$ this means we want to know the long run value of those probabilities, where:

$$\begin{array}{l} \lim_{n\to\infty} \Pr(X_n=j\mid X_{n-1}=i) = \ P_{ij} \ ; \ where \ P \ is \ the \ given \ transition \ probability \ matrix \\ \lim_{n\to\infty} \Pr(X_n=j) = \ \pi_j \ ; \ where \ \pi_j \ is \ the \ initial \ probability \ of \ state \ j\in \mathbb{S}=\{0,1,2\} \end{array}$$

- The value of $\pi_i \forall j \in \mathbb{S}$ is the entire limiting distribution of the Markov Chain, X_n
- We can compute the entire limiting distribution of X_n by solving the following system of equations:

$$\begin{split} & \pi_j = \sum_{k \,\in \,\mathbb{S}} \, \pi_k P_{kj} & \ \forall \, j \in \mathbb{S} \\ & 1 = \, \sum_{k \,\in \,\mathbb{S}} \pi_k \end{split}$$

• Applying the above equations, we have the following system of equations:

$$\begin{split} \pi_0 &= 0.4\pi_0 + 0.6\pi_1 + 0.4\pi_2 \\ \pi_1 &= 0.4\pi_0 + 0.2\pi_1 + 0.2\pi_2 \\ \pi_2 &= 0.2\pi_0 + 0.2\pi_1 + 0.4\pi_2 \\ 1 &= \pi_1 + \pi_2 + \pi_3 \end{split}$$

- Solving the above system will yield: $(\pi_0, \pi_1, \pi_2) = (11/24, 7/24, 1/4)$
- Using the given matrix P and the computed values of (π_0, π_1, π_2) we can compute the given limit:

$$\lim_{n \to \infty} \Pr(X_{n-1} = 2 \mid X_n = 1) = \lim_{n \to \infty} \frac{\Pr(X_n = 1 \mid X_{n-1} = 2) * \Pr(X_{n-1} = 2)}{\Pr(X_n = 1)} = \frac{\Pr_{21} * \pi_2}{\pi_1} = \frac{(0.2) * (1/4)}{(7/24)} = 6/35$$

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Problem 5

• When the transition probability matrix has row sums and column sum both equal to one, then this is known as a doubly stochastic transition probability matrix. This means that the given P satisfies the following conditions:

$$\begin{split} P_{ij} \geq 0 \ \forall \ i,j \in \mathbb{S} = \{1,2,...\,,M\} \\ \sum_{k \,\in\, \mathbb{S}} P_{ik} = \ \sum_{k \,\in\, \mathbb{S}} P_{kj} = 1 \ \forall \ i,j \in \mathbb{S} = \{1,2,...\,,M\} \end{split}$$

• When a transition probability matrix P satisfies the above conditions (ie. P is doubly stochastic), then when computing the entire limiting distribution of X_n there is only one solution to the following system of equations:

$$\begin{split} \pi_j &= \sum_{k \,\in \,\mathbb{S}} \, \pi_k P_{kj} & \ \forall \ j \in \mathbb{S} \\ 1 &= \, \sum_{k \,\in \,\mathbb{S}} \pi_k \end{split}$$

• The solution to the above system when P is doubly stochastic is the following:

$$\pi_j = \sum_{k \,\in\, \mathbb{S}} \frac{1}{M} \, P_{kj} = \, \frac{1}{M} \,\, \forall \, j \in \mathbb{S}$$
 ; where M is the size of the state space

• Therefore, the limiting probability vector $\pi = (\pi_1, \pi_2, ..., \pi_M) = (1/M, 1/M, ..., 1/M)$