

Stochastic Processes

Homework 6

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Problem 1

- X_n is our Markov chain with state space $\mathbb{S} = \{0, 1\}$
- Each time X_n is in a state, it sends one message
- The probability of a good message sent is $\{p_0, p_1\}$ when X_n is in state $\{0, 1\}$ respectively
- The probability of a bad message sent is $\{1 - p_0, 1 - p_1\}$ when X_n is in state $\{0, 1\}$ respectively
- Using the law of total probability and Markovian property, we can write out the probability that a good message is sent on Tuesday given the process is in state 0 on Monday

$$\begin{aligned} & \sum_{i=0}^1 \Pr(X_n = i, \text{"good message in state i"} \mid X_{n-1} = 0) \\ & \sum_{i=0}^1 (\Pr(X_n = i \mid X_{n-1} = 0) * \Pr(\text{"good message in state i"})) \\ & \sum_{i=0}^1 P_{0i} * \Pr(\text{"good message in state i"}) \\ & P_{00}p_0 + P_{01}p_1 \\ & 0.4p_0 + 0.6p_1 \end{aligned}$$

- The proportion of good messages in the long run is computed with the limiting distribution of X_n
- We can compute the limiting distribution of X_n by solving the following system of equations:

$$\begin{aligned} \pi_j &= \sum_{k \in \mathbb{S}} \pi_k P_{kj} \quad \forall j \in \mathbb{S}; \text{ where } \pi_j \text{ is the initial probability of state } j \in \mathbb{S} \\ 1 &= \sum_{k \in \mathbb{S}} \pi_k \end{aligned}$$

- Applying the above equations, we have the following system of equations:

$$\begin{aligned} \pi_0 &= 0.4\pi_0 + 0.2\pi_1 \\ \pi_1 &= 0.6\pi_0 + 0.8\pi_1 \\ 1 &= \pi_0 + \pi_1 \end{aligned}$$

- Solving the above system will yield: $(\pi_0, \pi_1) = (1/4, 3/4)$
- The proportion of good messages in the long run is:

$$\pi_0 p_0 + \pi_1 p_1 = \frac{1}{4} p_0 + \frac{3}{4} p_1$$

- If Y_n were a Markov chain, then it would have the state space $\mathbb{S} = \{1, 2\}$ corresponding to “good message” and “bad message” respectively
- Furthermore, it would have to follow Markovian property where:

$$\Pr(Y_{n+1} = j \mid Y_0 = i_0, \dots, Y_{n-1} = i_{n-1}, Y_n = i) = \Pr(Y_{n+1} = j \mid Y_n = i)$$

- But according to the problem definition, a message being good or bad only depends on the current state of X_n and mentions no dependence on the previous state, X_{n-1}
- Because the current message has no dependence on X_{n-1} it also has no dependence on the previous message in X_{n-1}
- Therefore, by contradiction of the Markovian property, Y_n is not a Markov chain

Problem 2

- X_n is our Markov chain with state space $\mathcal{S} = \{0, 1, 2, 3\}$ with the following transition probability matrix:

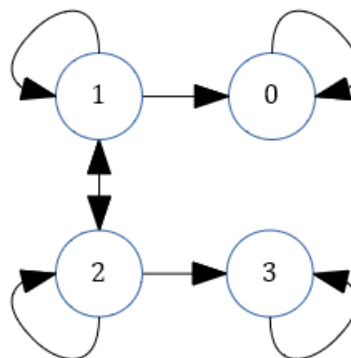
$$P = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \end{matrix} & \begin{bmatrix} 1.00 & 0.00 & 0.00 & 0.00 \\ 0.50 & 0.25 & 0.25 & 0.00 \\ 0.00 & 0.50 & 0.25 & 0.25 \\ 0.00 & 0.00 & 0.00 & 1.00 \end{bmatrix} \end{matrix}$$

- Let A represent all absorbing states where $A = \{0, 3\}$
- Let A^c represent all non-absorbing states where $A^c = \{1, 2\}$
- Let's rewrite P in terms of A and A^c

$$P = \begin{matrix} & \begin{matrix} A^c & A \end{matrix} \\ \begin{matrix} A^c \\ A \end{matrix} & \begin{bmatrix} Q & B \\ 0 & I \end{bmatrix} \end{matrix}$$

$$P = \begin{matrix} & \begin{matrix} 1 & 2 & 0 & 3 \end{matrix} \\ \begin{matrix} 1 \\ 2 \\ 0 \\ 3 \end{matrix} & \begin{bmatrix} 0.25 & 0.25 & 0.50 & 0.00 \\ 0.50 & 0.25 & 0.00 & 0.25 \\ 0.00 & 0.00 & 1.00 & 0.00 \\ 0.00 & 0.00 & 0.00 & 1.00 \end{bmatrix} \end{matrix}$$

- Let's draw out the state transition diagram of P to identify the communicating classes, transient states, and recurrent states
- We will apply the following rule to determine the communicating classes:
 - Two states i and j belong to the same communicating class if and only if $i \leftrightarrow j$
- We will apply the following rule to determine the transient and recurrent states:
 - For any state i , denote $f_i = \Pr(\text{"ever reenter } i" \mid X_0 = i)$, where a state i is recurrent if $f_i = 1$, and is transient if $f_i < 1$



- We can see that for P the communicating classes are $\{1, 2\}$, $\{0\}$, and $\{3\}$ making it not an irreducible chain
- We can see that for P the recurrent states are $\{0, 3\}$ and the transient states are $\{1, 2\}$

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- Let us define a matrix R to denote the expected number of returns to a state j given that the initial state was i

$$R_{ij} = \begin{cases} 0 & \text{; when } i \text{ is recurrent and } j \text{ is transient} \\ 0 & \text{; when } i \text{ is recurrent and } j \text{ is in a different irreducible set than } i \\ 0 & \text{; when } i \text{ is transient and } j \text{ is recurrent and } i \nrightarrow j \\ \infty & \text{; when } i, j \text{ are in the same irreducible set} \\ \infty & \text{; when } i \text{ is transient and } j \text{ is recurrent and } i \rightarrow j \\ (I - Q)_{ij}^{-1} & \text{; when } i, j \text{ are transient} \end{cases}$$

$$R = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \end{matrix} & \begin{bmatrix} \infty & 0.000 & 0.000 & 0.000 \\ \infty & 1.714 & 0.571 & \infty \\ \infty & 1.143 & 1.714 & \infty \\ 0.000 & 0.000 & 0.000 & \infty \end{bmatrix} \end{matrix}$$

- Let us define a matrix F to denote the first passage probability, the probability of eventually reaching a state j at least once given that the initial state was i

$$F_{ij} = \begin{cases} ((I - Q)^{-1}B)_i & \text{; when } i \text{ is transient and } j \text{ is recurrent} \\ 1 - \frac{1}{R_{jj}} & \text{; when } i, j \text{ are transient and } i = j \\ \frac{R_{ij}}{R_{jj}} & \text{; when } i, j \text{ are transient and } i \neq j \text{ ; where } \frac{0}{0} = 0 \end{cases}$$

$$F = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \end{matrix} & \begin{bmatrix} 1.000 & 0.000 & 0.000 & 0.000 \\ 0.857 & 0.417 & 0.333 & 0.143 \\ 0.571 & 0.667 & 0.417 & 0.429 \\ 0.000 & 0.000 & 0.000 & 1.000 \end{bmatrix} \end{matrix}$$

- The probability of the rocket eventually getting on course (ie. state 0) given that the initial state was 2 corresponds to $F_{20} = 0.571$
- The probability of the rocket eventually getting destroyed (ie. state 3) given that the initial state was 2 corresponds to $F_{23} = 0.429$
- The expected number of times a minor correction is made (ie. state 1) given that the initial state was 2 corresponds to $R_{21} = 1.143$
- The expected number of times a major correction is made (ie. state 2) given that the initial state was 2 corresponds to $R_{22} = 1.714$
- When the rocket is launched, 50000 lb. of are used
- When a minor correction is made, 1000 lb. of fuel are used
- When a major correction is made, 5000 lb. of fuel are used
- Therefore, the expected amount of fuel needed for the mission given that the initial state was 2 is:

$$50000 + 1000(1.143) + 5000(1.714) = 59713 \text{ lb.}$$

Problem 3

- X_n is our Markov chain with state space $\mathbb{S} = \{0, 1\}$ corresponding to, the car is where the man is and the car is not where the man is, respectively
- Given that it rains, independent of time, during successive mornings and afternoons with probability p , we can write out the transition probability matrix:

$$P = \begin{matrix} & \begin{matrix} 0 & 1 \end{matrix} \\ \begin{matrix} 0 \\ 1 \end{matrix} & \begin{bmatrix} p & 1-p \\ 1 & 0 \end{bmatrix} \end{matrix}$$

- The long run fraction of days the man walks in the rain is computed with the limiting distribution of X_n
- We can compute the limiting distribution of X_n by solving the following system of equations:

$$\begin{aligned} \pi_j &= \sum_{k \in \mathbb{S}} \pi_k P_{kj} \quad \forall j \in \mathbb{S}; \text{ where } \pi_j \text{ is the initial probability of state } j \in \mathbb{S} \\ 1 &= \sum_{k \in \mathbb{S}} \pi_k \end{aligned}$$

- Applying the above equations, we have the following system of equations:

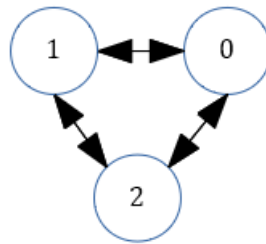
$$\begin{aligned} \pi_0 &= p\pi_0 + \pi_1 \\ \pi_1 &= (1-p)\pi_0 + 0\pi_1 \\ 1 &= \pi_0 + \pi_1 \end{aligned}$$

- Solving the above system will yield: $(\pi_0, \pi_1) = (\frac{1}{2-p}, \frac{1-p}{2-p})$
- The fraction of days that the man walks in the rain is when either of these two scenarios occur:
 - he doesn't have his car and it rains
 - he has his car and it doesn't rain but then it rains on his way back
- We can compute, in the long run, the fraction of days that the man walks in the rain using the limiting distribution, the transition probability matrix, and the two previous scenarios:

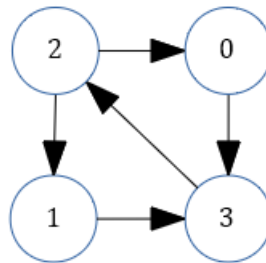
$$\begin{aligned} &\Pr((\text{"no car"} \cap \text{"rain at time } i\text{"}) \cup (\text{"has car"} \cap \text{"no rain at time } i\text{"} \cap \text{"rain at time } i+1\text{"})) \\ &= \Pr(\text{"no car"}) * \Pr(\text{"rain at time } i\text{"}) + \Pr(\text{"has car"}) * \Pr(\text{"no rain at time } i\text{"}) * \Pr(\text{"rain at time } i+1\text{"}) \\ &= \pi_1 * p + \pi_0 * (1-p) * p \\ &= \frac{1-p}{2-p} * p + \frac{1}{2-p} * (1-p) * p \\ &= 2 \frac{(1-p)p}{2-p} \end{aligned}$$

Problem 4

- Let P_1 have the following state space $\mathbb{S}_1 = \{0, 1, 2\}$ where the rows and columns of P_1 are labeled in the same order as the elements of \mathbb{S}_1
- Let P_2 have the following state space $\mathbb{S}_2 = \{0, 1, 2, 3\}$ where the rows and columns of P_2 are labeled in the same order as the elements of \mathbb{S}_2
- Let P_3 have the following state space $\mathbb{S}_3 = \{0, 1, 2, 3, 4\}$ where the rows and columns of P_3 are labeled in the same order as the elements of \mathbb{S}_3
- Let P_4 have the following state space $\mathbb{S}_4 = \{0, 1, 2, 3, 4\}$ where the rows and columns of P_4 are labeled in the same order as the elements of \mathbb{S}_4
- We will apply the following rule to determine the communicating classes:
 - Two states i and j belong to the same communicating class if and only if $i \leftrightarrow j$
- We will apply the following rule to determine the transient and recurrent states:
 - For any state i , denote $f_i = \Pr(\text{"ever reenter } i" \mid X_0 = i)$, where a state i is recurrent if $f_i = 1$, and is transient if $f_i < 1$
- Let's draw out the state transition diagram of P_1 to identify the communicating classes, transient states, and recurrent states



- We can see that for P_1 the one communicating class is $\{0, 1, 2\}$ making it an irreducible chain
- We can see that for P_1 the recurrent states are $\{0, 1, 2\}$ and there are no transient states
- Let's draw out the state transition diagram of P_2 to identify the communicating classes, transient states, and recurrent states

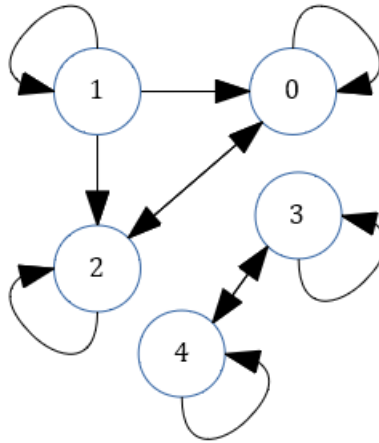


- We can see that for P_2 the one communicating class is $\{0, 1, 2, 3\}$ making it an irreducible chain
- We can see that for P_2 the recurrent states are $\{0, 1, 2, 3\}$ and there are no transient states

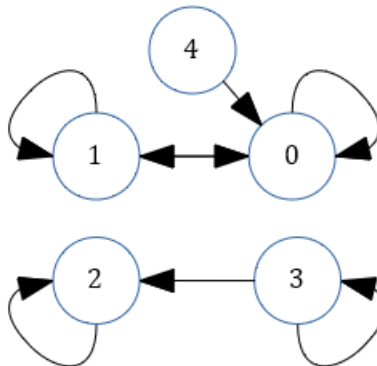
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- Let's draw out the state transition diagram of P_3 to identify the communicating classes, transient states, and recurrent states



- We can see that for P_3 the communicating classes are $\{0, 2\}$, $\{3, 4\}$, and $\{1\}$ making it not an irreducible chain
- We can see that for P_3 the recurrent states are $\{0, 2, 3, 4\}$ and the transient states are $\{1\}$
- Let's draw out the state transition diagram of P_4 to identify the communicating classes, transient states, and recurrent states



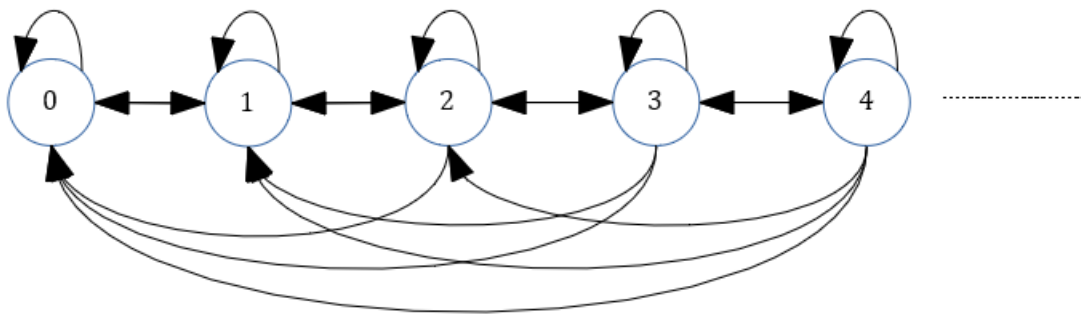
- We can see that for P_4 the communicating classes are $\{0, 1\}$, $\{2\}$, $\{3\}$, and $\{4\}$
- We can see that for P_4 the recurrent states are $\{0, 1, 2\}$ and the transient states are $\{3, 4\}$

Problem 5

- The transition probability matrix of this Markov chain is built with the following pattern:

$$P = \begin{matrix} & \begin{matrix} 0 & 1 & 2 & 3 & 4 & 5 & \dots \end{matrix} \\ \begin{matrix} 0 \\ 1 \\ 2 \\ 3 \\ \vdots \end{matrix} & \begin{bmatrix} 1/2 & 1/2 & 0/2 & 0/2 & 0/2 & 0/2 & \dots \\ 1/3 & 1/3 & 1/3 & 0/3 & 0/3 & 0/3 & \dots \\ 1/4 & 1/4 & 1/4 & 1/4 & 0/4 & 0/4 & \dots \\ 1/5 & 1/5 & 1/5 & 1/5 & 1/5 & 0/5 & \dots \\ \vdots & \vdots & \vdots & \vdots & \vdots & \ddots & \ddots \end{bmatrix} \end{matrix}$$

- The state transition diagram of P is built with the following pattern:



- We can see that there is one communicating class of all states, making it an irreducible chain
- We can see that all states are recurrent
- We can compute the stationary distribution of this Markov chain X_n by solving the following system of equations:

$$\begin{aligned} \pi_j &= \sum_{k \in \mathbb{S}} \pi_k P_{kj} \quad \forall j \in \mathbb{S}; \text{ where } \pi_j \text{ is the initial probability of state } j \in \mathbb{S} \\ 1 &= \sum_{k \in \mathbb{S}} \pi_k \end{aligned}$$

- Applying the above equations, we can build a system of equations with the following pattern:

$$\begin{aligned} 1 &= \pi_0 + \pi_1 + \pi_2 + \pi_3 + \pi_4 \\ \pi_0 &= \pi_1 = \frac{1}{2}\pi_0 + \frac{1}{3}\pi_1 + \frac{1}{4}\pi_2 + \frac{1}{5}\pi_3 \\ \pi_2 &= \frac{1}{3}\pi_1 + \frac{1}{4}\pi_2 + \frac{1}{5}\pi_3 \\ \pi_3 &= \frac{1}{4}\pi_2 + \frac{1}{5}\pi_3 \\ \pi_4 &= \frac{1}{5}\pi_3 \end{aligned}$$

- The above system has no solution, this is because the first two columns of P are identical and therefore linearly dependent
- This means that there is no stationary distribution for this Markov chain