Homework 7

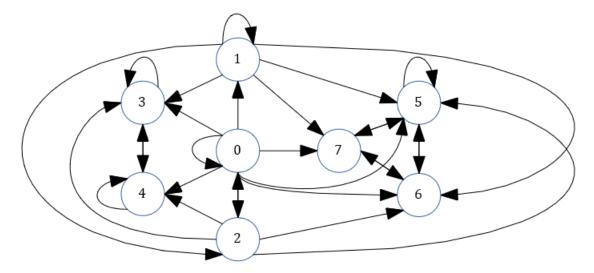
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Problem 1

- X_n is our Markov chain with state space $S = \{0, 1, 2, 3, 4, 5, 6, 7\}$
- The state transition diagram of the given transition probability matrix P is shown below:



- We will apply the following rule to determine the communicating classes:
 - \circ Two states i and j belong to the same communicating class if and only if i \leftrightarrow j
- We will apply the following rule to determine the transient and recurrent states:
 - For any state i, denote $f_i = Pr(\text{"ever reenter i"} \mid X_0 = i)$, where a state i is recurrent if $f_i = 1$, and is transient if $f_i < 1$
- We will apply the following rule to determine if a recurrent state is positive recurrent:
 - o If there is only a finite number of states, then there are no null recurrent states, and not all states can be transient
- We can see that there are three communicating classes:
 - o $A = \{0, 1, 2\}$ which is transient because each of those states are reachable from another, but there is no guarantee that a state $i \in A$ will return to back to state i given that the process started in state i
 - \circ B = {3, 4} which is positive recurrent because each of those states are reachable from another, it is guaranteed that a state i \in B will eventually return to back to state i given that the process started in state i, and there are only a finite number of states in this chain
 - $C = \{5, 6, 7\}$ which is positive recurrent because each of those states are reachable from another, it is guaranteed that a state $i \in C$ will eventually return to back to state i given that the process started in state i, and there are only a finite number of states in this chain
- Let's write out the transition probability sub-matrices P_B and P_C corresponding to classes B and C respectively:

- We can compute the stationary distributions of the states in classes B and C by solving the following system of equations:
 - o For states in class B:

$$\pi_j = \sum_{k \,\in\, B} \, \pi_k P_{kj} \quad \forall \; j \in B$$
 ; where π_j is the initial probability of state $j \in B$
$$1 = \, \sum_{k \,\in\, B} \pi_k$$

For states in class C:

$$\begin{split} \pi_j &= \sum_{k \,\in\, C} \, \pi_k P_{kj} \quad \forall \; j \in C \; ; \text{ where } \pi_j \text{ is the initial probability of state } j \in C \\ 1 &= \, \sum_{k \,\in\, C} \pi_k \end{split}$$

- Applying the above equations, we have the following system of equations:
 - For states in B:

$$\begin{split} &\pi_3 = 0.3\pi_3 + 0.7\pi_4 \\ &\pi_4 = 0.7\pi_3 + 0.3\pi_4 \\ &1 = \pi_3 + \pi_4 \end{split}$$
 Solving the above system will yield: $(\pi_3, \pi_4) = (\frac{1}{2}, \frac{1}{2})$

For states in C:

$$\begin{split} \pi_5 &= 0.3\pi_5 + \ 0.1\pi_6 + 0.8\pi_7 \\ \pi_6 &= 0.4\pi_5 + \ 0.0\pi_6 + 0.2\pi_7 \\ \pi_7 &= 0.3\pi_5 + \ 0.9\pi_6 + 0.0\pi_7 \\ 1 &= \pi_5 + \ \pi_6 + \pi_7 \\ \text{Solving the above system will yield: } (\pi_5, \pi_6, \pi_7) = (\frac{41}{97}, \frac{23}{97}, \frac{33}{97}) \end{split}$$

- We can compute the hitting probabilities from the transient states in A to the positive recurrent states in {B, C} with first step analysis.
- Let $u_i^B \forall i \in A$ be the probability of ultimate absorption in class B from state i
- Let $u_i^C \forall i \in A$ be the probability of ultimate absorption in class C from state i
 - Where $u_i^B + u_i^C = 1 \ \forall \ i \in A$ because the Markov process will eventually end up in a recurrent class given it started in a transient state
- Using first step analysis we have the following system of equations:

$$\begin{split} u_0^B &= 0.1 u_0^B + \ 0.1 u_1^B + 0.2 u_2^B + 0.2(1) + 0.1(1) + 0.1(0) + 0.1(0) + 0.1(0) \\ u_1^B &= 0.0 u_0^B + \ 0.1 u_1^B + 0.1 u_2^B + 0.1(1) + 0.0(1) + 0.3(0) + 0.2(0) + 0.2(0) \\ u_2^B &= 0.6 u_0^B + \ 0.0 u_1^B + 0.0 u_2^B + 0.1(1) + 0.1(1) + 0.1(0) + 0.1(0) + 0.0(0) \end{split}$$

- Solving the above system will yield: $\left(u_0^B, u_1^B, u_2^B\right) = \left(\frac{53}{116}, \frac{19}{116}, \frac{55}{116}\right)$ We know that $u_i^B + u_i^C = 1 \ \forall \ i \in A$, therefore: $\left(u_0^C, u_1^C, u_2^C\right) = \left(\frac{63}{116}, \frac{97}{116}, \frac{61}{116}\right)$

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• Using the initial probabilities and hitting probabilities we can now determine the limiting behavior of the Markov chain with the following limiting matrix of P:

$$P^{\infty} = \begin{bmatrix} 0 & 1 & 2 & 3 & 4 & 5 & 6 & 7 \\ 0 & 0 & 0 & u_0^B \pi_3 & u_0^B \pi_4 & u_0^C \pi_5 & u_0^C \pi_6 & u_0^C \pi_7 \\ 1 & 0 & 0 & 0 & u_1^B \pi_3 & u_1^B \pi_4 & u_1^C \pi_5 & u_1^C \pi_6 & u_1^C \pi_7 \\ 2 & 0 & 0 & 0 & u_2^B \pi_3 & u_2^B \pi_4 & u_2^C \pi_5 & u_2^C \pi_6 & u_2^C \pi_7 \\ 3 & 0 & 0 & 0 & \pi_3 & \pi_4 & 0 & 0 & 0 \\ 5 & 6 & 0 & 0 & 0 & 0 & \pi_5 & \pi_6 & \pi_7 \\ 6 & 7 & 0 & 0 & 0 & 0 & \pi_5 & \pi_6 & \pi_7 \\ 7 & 0 & 0 & 0 & 0 & 0 & \pi_5 & \pi_6 & \pi_7 \\ 7 & 0 & 0 & 0 & 0 & 0 & \pi_5 & \pi_6 & \pi_7 \end{bmatrix}$$

	0	1	2	3	4	5	6	7
$P^{\infty} = \begin{array}{c} 0 \\ 1 \\ 2 \\ 3 \\ 4 \\ 5 \\ 6 \end{array}$	L0	0	0	53/232	53/232	2583/11252	1449/11252	2079/11252ן
	0	0	0	19/232	19/232	41/116	23/116	33/116
	0	0	0	55/232	55/232	2501/11252	1403/11252	2013/11252
	0	0	0	1/2	1/2	0	0	0
	0	0	0	1/2	1/2	0	0	0
	10	0	0	0	0	41/97	23/97	33/97
	0	0	0	ő	ŏ	41/97	23/97	33/97
7	Lo	0	0	0	0	41/97	23/97	33/97

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Problem 2

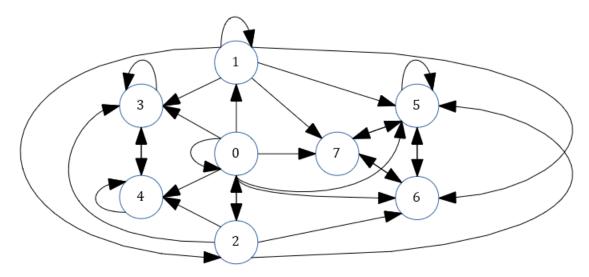
 Let us define a matrix R to denote the expected number of returns to a state j given that the initial state was i

$$R_{ij} = \begin{cases} 0 \text{ ; when } i \text{ is recurrent and } j \text{ is transient} \\ 0 \text{ ; when } i \text{ is recurrent and } j \text{ is in a different irreducible set than } i \\ 0 \text{ ; when } i \text{ is transient and } j \text{ is recurrent and } i \Rightarrow j \\ \infty \text{ ; when } i, j \text{ are in the same irreducible set} \\ \infty \text{ ; when } i \text{ is transient and } j \text{ is recurrent and } i \rightarrow j \\ (I-Q)_{ij}^{-1} \text{ ; when } i, j \text{ are transient} \end{cases}$$

• Where Q is the following matrix, corresponding to the transient states of P from Problem 1 (class A)

$$Q = \begin{bmatrix} 0 & 1 & 2 \\ 0 & 0.1 & 0.1 & 0.2 \\ 1 & 0.0 & 0.1 & 0.1 \\ 2 & 0.6 & 0.0 & 0.0 \end{bmatrix}$$

• Recall the state transition diagram from Problem 1:



- In Problem 1 we defined state 1 as a transient state and state 6 as a positive recurrent state
- The expected number of visits to state 1 given the initial state was 6, would be the entry R₆₁
- By the definition of R_{ij} we can see that the expected number of visits from state 6 to state 1 is 0 because state 6 is recurrent and state 1 is transient
- In the state transition diagram above we can see that state 6 is reachable from state 1 via states 2, 5, or 7
- By the definition of R_{ij} we can see that the expected number of visits from state 1 to state 6 is ∞ because state 1 is transient, state 6 is recurrent, and $1\rightarrow 6$

- In Problem 1 we defined states 0 and 2 as transient states
- By the definition of R_{ij} we can see that the expected number of visits from state 0 to state 2 is the entry $(I-Q)_{02}^{-1}$ because states 0 and 2 are transient
- The value of entry $(I Q)_{02}^{-1}$ is computed below:

$$(I - Q)^{-1} = \begin{bmatrix} 0 & 1 & 2 \\ 75/58 & 25/174 & 95/348 \\ 1 & 5/58 & 65/58 & 15/116 \\ 2 & 45/58 & 5/58 & 135/116 \end{bmatrix}$$

$$(I - Q)_{02}^{-1} = 95/348 \cong 0.273$$
 visits

- In Problem 1 we defined states 2 and 1 as transient states
- By the definition of R_{ij} we can see that the expected number of visits from state 2 to state 1 is the entry $(I-Q)_{21}^{-1}$ because states 2 and 1 are transient • The value of entry $(I-Q)_{21}^{-1}=5/58\cong0.086$ visits

Problem 3

Let us define a matrix F to denote the first passage probability, the probability of eventually reaching a state j given that the initial state was i

$$F_{ij} = \begin{cases} ((I-Q)^{-1}D)_i \text{ ; when } i \text{ is transient and } j \text{ is recurrent} \\ 1-\frac{1}{R_{jj}} \text{ ; when } i,j \text{ are transient and } i=j \\ \frac{R_{ij}}{R_{jj}} \text{ ; when } i,j \text{ are transient and } i\neq j \text{ ; where } \frac{0}{0}=0 \end{cases}$$

Where D is the following matrix, corresponding to the transient states transitioning into the recurrent states of P from Problem 1

- In Problem 1 we defined states 0, 1, and 2 as transient states
- The probability of ever visiting states 0, 1, and 2 from states 0, 1, and 2 would correspond to a sub-matrix F₁, where F₁ is a square matrix corresponding to all transient states
- We build this sub-matrix F_1 by using the matrix $(I-Q)^{-1}$ and the last two conditions that define F_{ij}

$$F_1 = \begin{bmatrix} 0 & 1 & 2 \\ 17/75 & 5/39 & 19/81 \\ 1/15 & 7/65 & 1/9 \\ 2 & 3/5 & 1/13 & 19/135 \end{bmatrix}$$

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- In Problem 1 we defined states 5 and 6 as positive recurrent states, and being in the same class
- The probability of ever visiting state 6 from 5 is 1 because they are both recurrent and in the same class, they share the same irreducible set
- In Problem 1 we defined state 2 as transient and state 4 as positive recurrent
- By the definition of F_{ij} we can see that the probability of visiting state 4 from state 2 is the entry $((I-Q)^{-1}D)_{24}$ because state 2 is transient and state 4 is recurrent
- The value of entry $((I Q)^{-1}D)_{24}$ is computed below:

$$(I - Q)^{-1}D = \begin{bmatrix} 3 & 4 & 5 & 6 & 7 \\ 209/696 & 109/696 & 139/696 & 43/232 & 55/348 \\ 1 & 33/232 & 5/232 & 83/232 & 57/232 & 27/116 \\ 2 & 65/232 & 45/232 & 51/232 & 49/232 & 11/116 \end{bmatrix}$$

$$((I - Q)^{-1}D)_{24} = 45/232 \approx 0.194$$

Problem 4

• The conditional distribution of X_t given that $X_{t+s} = n$ is shown below:

$$Pr(X_t = k \mid X_{t+s} = n)$$

• Using conditional probability

$$= \frac{\Pr(X_t = k, X_{t+s} = n)}{\Pr(X_{t+s} = n)}$$

$$= \frac{\Pr(X_{t} = k, X_{t+s} - X_{t} = n - k)}{\Pr(X_{t+s} = n)}$$

$$= \frac{\Pr(X_t = k, X_s = n - k)}{\Pr(X_{t+s} = n)}$$

• Since X_t and X_s are independent

$$= \frac{\Pr(X_t = k) \Pr(X_s = n - k)}{\Pr(X_{t+s} = n)}$$

• Substituting the PDF of the Poisson distribution

$$=\frac{\frac{(\mu t)^k e^{-\mu t}}{k!} \frac{(\mu s)^{n-k} e^{-\mu s}}{(n-k)!}}{\frac{(\mu (t+s))^n e^{-\mu (t+s)}}{n!}}$$

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• Simplifying the above fractions

$$= \binom{n}{k} \binom{t}{t+s}^{k} \binom{s}{t+s}^{n-k}$$

• Therefore, X_t conditioned on $X_{t+s} = n$ means X_t is a binomial random variable: binom(n, t / (t + s))

Problem 5

• The first conditional moment is computed below:

$$E[X_T|T = t] = E[X_t] = \lambda t = 4t$$

• The conditional variance is computed below:

$$Var[X_T|T = t] = Var[X_t] = \lambda t = 4t$$

• The second conditional moment is computed below:

$$E[X_T^2|T=t] = Var[X_T|T=t] + (E[X_T|T=t])^2 = \lambda t + (\lambda t)^2 = 4t + 16t^2 = 4t(1+4t)$$

• The unconditional mean is computed below:

$$E[X_T] = \int_{0}^{1} 4t * 1 dt = 2(1)^2 - 0 = 2$$

• The unconditional second moment is computed below:

$$E[X_T^2] = \int_0^1 (4t + 16t^2) * 1 dt = 2(1)^2 + \frac{16}{3}(1)^3 = \frac{22}{3}$$

• The unconditional variance is computed below:

$$Var[X_T] = E[X_T^2] - E[X_T]^2 = \frac{22}{3} - 2^2 = \frac{10}{3}$$