**What is learned during statistical learning? Investigating online and offline implicit learning**

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## **Abstract**

## **Introduction**

Statistical learning (SL) is one mechanism by which the brain is able to parse continuous streams of sensory information, based on simple statistical features alone. SL paradigms have shown the brain to be surprisingly effective at detecting statistical regularities in highly impoverished or artificial inputs, leading to a discrimination of meaningful boundaries where previously none were discernable. (Pelucchi, Hay, and Saffran 2009; Richard N. Aslin, Saffran, and Newport 1998; Turk-Browne, Jungé, and Scholl 2005; Turk-Browne et al. 2008)

Statistical learning is canonically measured using an explicit word discrimination task after exposure to the continuous speech stream.

Several online statistical learning measures have been previously shown to be effectively modulated by implicitly-acquired knowledge of the statistical regularity of reoccurring syllables. In a visual statistical learning task, Turk-Browne and colleagues found a graded reaction time (RT) effect to targets as a function of position in the underlying triplets. (Turk-Browne, Jungé, and Scholl 2005) In the auditory domain, Batterink and colleagues have successfully used the target detection task as a measure of learning as both a post-exposure test (Batterink et al. 2015; Batterink, Reber, and Paller 2015; Batterink and Paller 2017) and as an online measure while the participants are being exposed to the structured stimuli (Batterink 2017).

These tests are standard for measuring implicit learning, yet comparatively few studies have examined the relationship between these measures. Is performance on different SL tasks correlated within subjects? Do these tasks measure the same phenomenon (implicit learning)? Do they contribute redundant or distinct information about what the participants learned? Here, we address these questions in a replication with extension of previous work.

% Discuss (Siegelman et al. 2018; Siegelman and Frost 2015; Siegelman, Bogaerts, and Frost 2017; Armstrong, Frost, and Christiansen 2017; Frost et al. 2015)

We asked participants to complete an online target detection task and offline pseudoword recognition task. (Thus, a replication of (Batterink 2017)). We then asked a second cohort of participants to complete the online target detection task with both random and structured syllable streams, to establish a baseline performance level for our stimuli without the embedded pseudoword structure. We analyzed individual and group-level behavior on these measures to gain a better understanding of how these measures may be related and what types of information they can provide.

# **Experiment 1**

## **Method**

#### **Stimuli**

Speech stimuli consisted of 12 consonant-vowel (CV) pairs. We selected 5 unique vowels that are maximally separated in their manner and place of articulation. We ensured that none of these vowels typically occurred in unstressed syllables. We then selected 12 unique consonants, in order to render each syllable phonetically distinct from the others. We used the CELEX database to calculate the frequency of occurrence of each of our syllables in spoken German, as well as the frequency of co-occurrence between each pair of syllables. We eliminated high-frequency CV pairings from our list of possible syllables and formed the final words by combining three syllables (each with distinct vowels) for which no transitions were frequent in spoken German. Final syllables were: be, di, ga, ki, la, mi, nu, po, ro, se, tu, za.

A male native speaker of German was recorded pronouncing each syllable in our set separately and with a flat intonation. Each syllable was repeated several times to ensure we obtained a quality token. The final selection of syllables were then high-pass filtered at 50 Hz and silences before and after syllable were removed using a custom script in Matlab 2017b. The 12 syllables were normalized for pitch and intensity using Praat to ensure relative homogeneity between tokens. Finally, syllables were compressed to 240 ms in duration and a 10 ms silence was added at the end of each syllable, for a total duration of 250 ms.

Syllables were combined into 4 tri-syllabic pseudowords such that each word featured no repeating vowels and similarity between any possible succeeding pairs of syllables was minimized. We also ensured that no pairs were phototactically illegal or shared a resemblance with existing words in German. Pseudowords for our study were: nugadi, rokise, mipola, zabetu. Part-words, used in the word recognition task, were of the form C’AB (word-final syllable from one word followed by word-initial and word-medial syllables from another): dizabe, semipo, lanuga, turoki.

Sequences (24) of ~1 minute in length were created in Matlab by concatenating syllables so that no pseudowords repeated consecutively. As per the design in (Saffran, Aslin, and Newport 1996), the only cue to segmenting the sequence lay in the transitional probabilities between syllables. The transitional probability of word-medial and word-final syllables was 1, while the transitional probability of word-initial syllables (occurring only at word boundaries) was 0.33. Speech streams were ramped up and down in amplitude over a period of 1.5 seconds so that onset and offset syllables were not clearly distinguishable and could not serve as cues to word boundaries.

#### **Procedure**

41 individuals participated in the study (x female, mean age, +- sd). Two participants were removed from the data pool due to technical failure. Of the 39 remaining datasets, 33 were used in analyzing the target detection task (one participant failed to follow instructions, and technical issues caused partial data loss for the other five). Since the design of our experiment was modular, failure in one phase did not necessarily affect data loss in another. Thus, of the 39 remaining datasets, we were able to use 38 for analyzing the word recognition task (data from one participant in this task was overwritten).

Participants were seated in a dimly-lit, sound-attenuated booth, approximately 52 cm from the monitor and listened to the stimuli via headphones connected to a headphone amplifier (Beyerdynamics-DT-770 80 Ohm; Lakepeople G103P1262). Stimulus intensity level was approximately 57 dB (LAF: min 44 dB, max 76 dB), as measured by a NTi Audio device connected to an artificial ear on which the experiment headphones were mounted.

The experiment was designed using Presentation® (Version 20.1 Build 12.04.17) and delivered on two versions of the software (Version 20.0 Build 07.26.17 and Version 21.1 Build 09.05.19, Neurobehavioral Systems, Inc., Berkeley, CA, [www.neurobs.com](http://www.neurobs.com)). The experiment was conducted on a 64-bit Windows machine (Fujitsu Celsius M740B) running Windows 10.

During the exposure phase, where participants performed an online target detection task, participants listened to a total of approximately 24 minutes of continuous speech. Participants were told they would hear brief sequences of sounds from an alien language. Speech sequences of ~1 minute in duration were presented binaurally. Before the start of each stream, one of the 12 syllables was displayed orthographically on the screen and played aurally twice. Participants were instructed to press the spacebar as fast as they could during the subsequent stream whenever they heard this target syllable. Each of the 12 syllables served as a target syllable twice. The presentation order of syllables was pseudo-randomly shuffled for each participant so that a syllable from each ordinal position in the pseudoword (1st, 2nd, or 3rd) was tested before any were repeated. The 24 streams were organized into 8 blocks, were each block consisted of 3 streams with one target syllables from each ordinal position tested. Within each stream, target syllables appeared approximately 18 times. There were no pauses between syllables (aside from the 10 ms silence at the end of each syllable’s offset), and participants could take self-paced breaks between blocks.

In the word recognition task, participants completed 16 trials of a two-alternative forced-choice task. In each trial, a pseudoword and a part-word were presented (counterbalanced across trials), and participants were prompted to determine which of the pair was a word in the alien language they had just heard in the previous section. The inter-stimulus-interval between words was 400 ms, while inter-trial-interval was 1.2 seconds. Each pseudoword was paired with each part-word (see Stimuli) once (4 x 4 trials).

Our experiment also included an additional task, designed to measure perceived speed of the speech stream before versus after the exposure phase. Results from this task will not be discussed here.

#### **Analyses**

All analyses were performed in RStudio (version 1.2.1335; RStudio Team 2018) using R version 3.6.1 (R Core Team 2019). Raw data from Presentation® was transformed into appropriate formats using Matlab R2017b (version 9.3.0.713579). Presentation® scenario files, data wrangling scripts, and R scripts with used packages are available on Github.

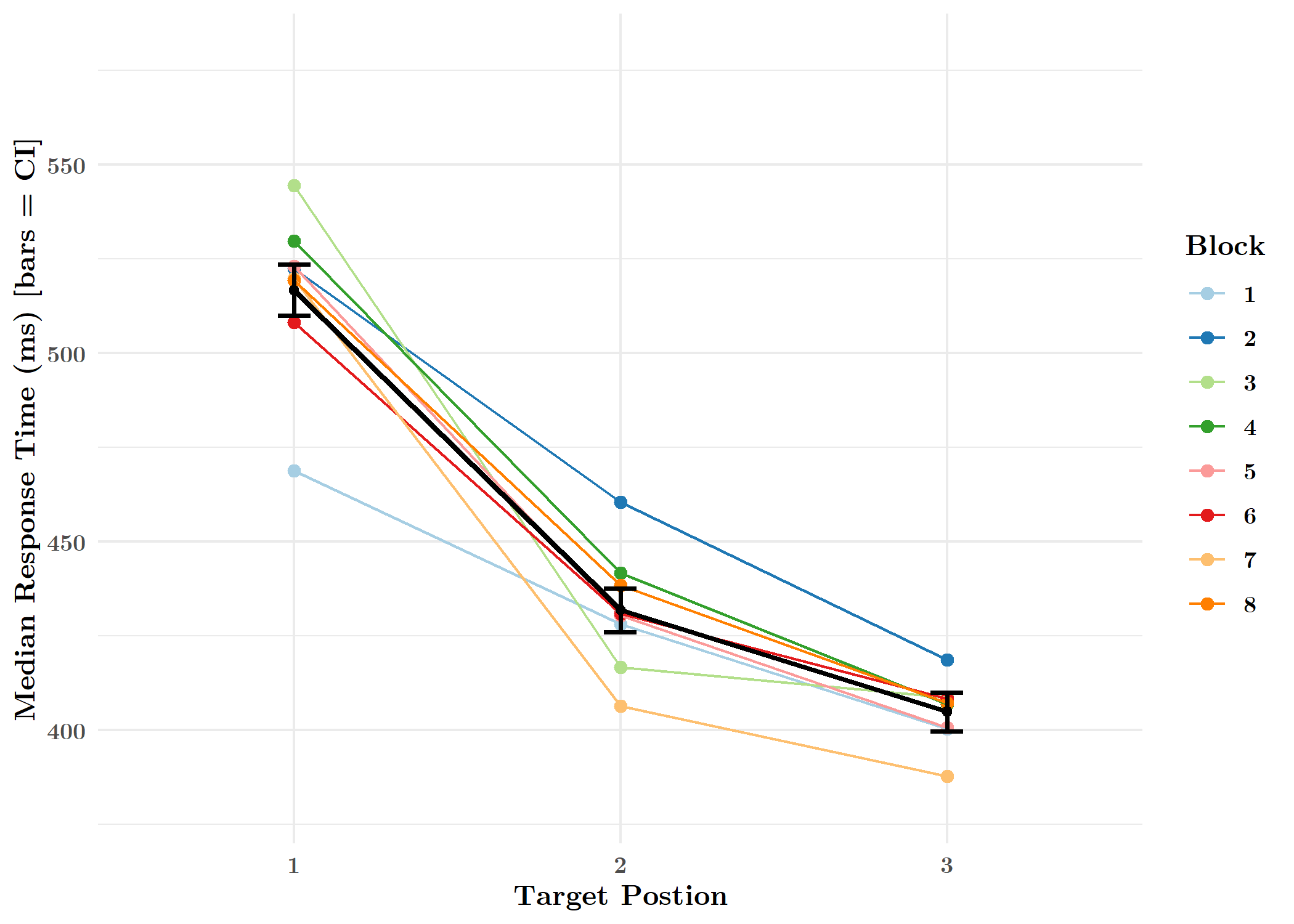
For the target detection task analyses, we considered only those responses that occurred within a boundary of ± 3 times the median absolute deviation. This procedure ensures that RT cutoffs would be based on the distribution of the raw data and not arbitrary limits. At the same time, the use of the median as the centrality metric is arguably more appropriate, given that the mean can be a biased estimator of RT data, which typically follows a gamma, lognormal, or ex-gaussian distribution. This procedure eliminated only 0.034% of the data and resulted in RT that ranged from 0 to 943 msecs (versus the original 0 to 1298 msecs).

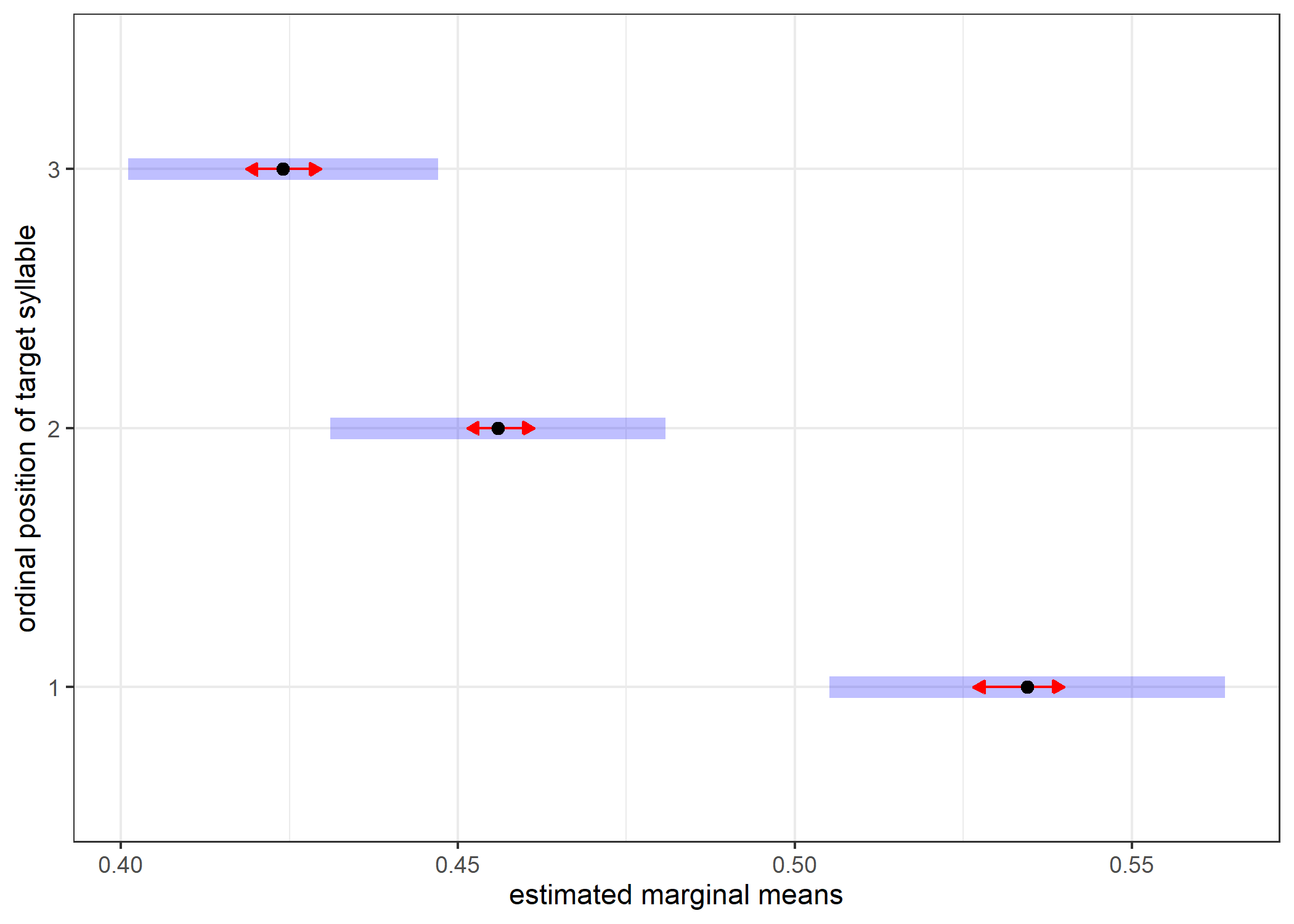
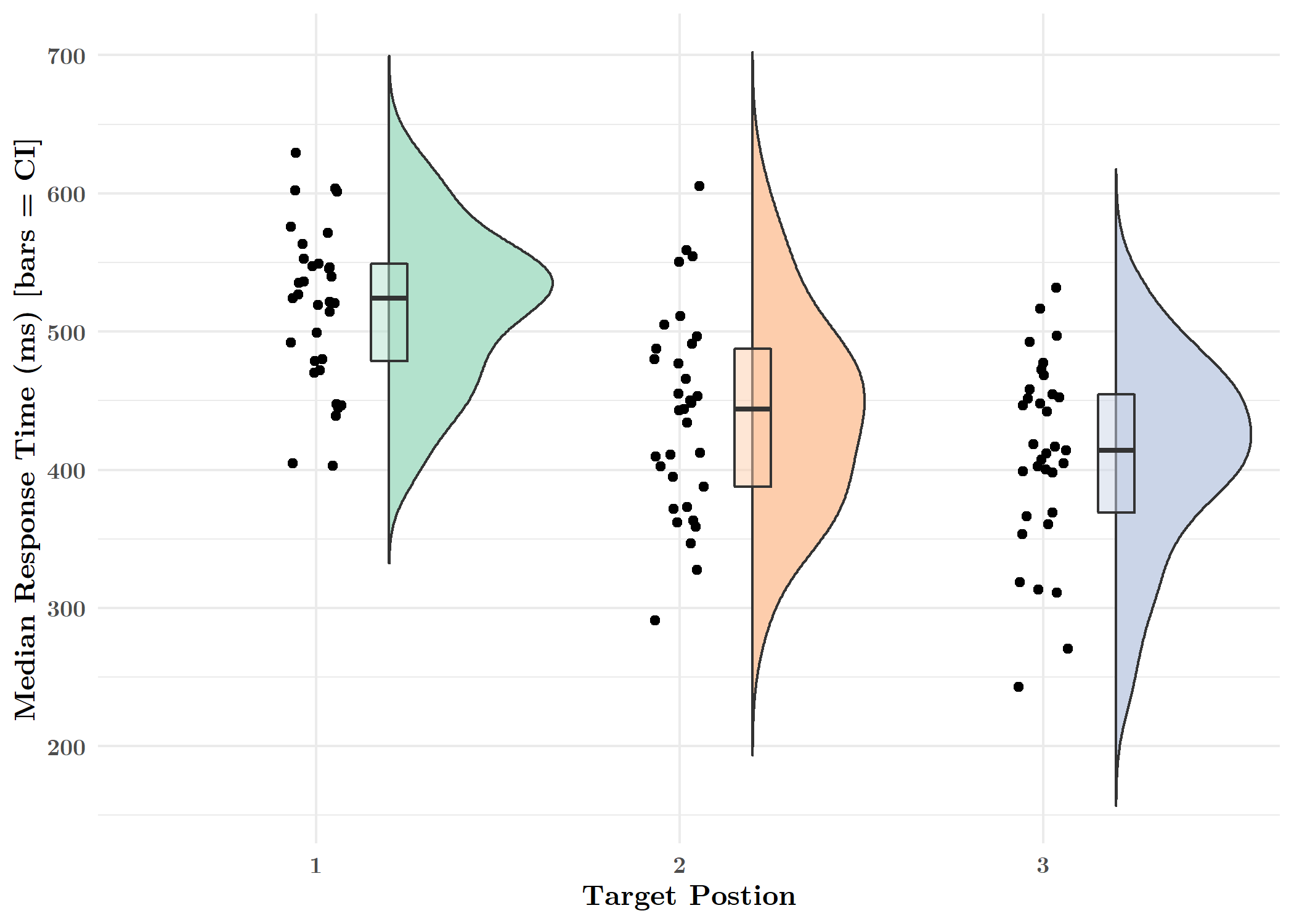
## **Results**

***Ordinal Position Modulates Reaction Time***

To replicate findings that showed graded reaction times in response to syllables in different ordinal positions, we ran a generalized linear model with reaction time (in seconds) as outcome variable, fitted with a gamma function and log link function. Our full model included both ordinal position and block as fixed effects factors, and subject as a random intercept-random effects factor. This model was compared with a lesser model in which only ordinal position was used as a fixed effect. The lesser model provided a better fit of the data, with a lower AIC (-6389.2) value and significantly lower deviance (-6399.2, *Χ2*(21, *N* = 33) = 49.066, *p* = 0.00049). (See **Table S1** for regression results.) We also compared both the fuller and the lesser models with random slopes for levels of ordinal position in the random effects term, but the lesser model with only varying random intercepts in the random effects term still proved a better fit for observed data (see **Table 1.** Model output for the fuller random slopes model not shown; lesser vs. fuller random slopes models deviance was -6624.3 and -6678.3, respectively; *Χ2*(21, *N* = 33) = 54.025, *p* < 0.0001). Thus, we conducted further analysis on results of the lesser model.

We found that reaction times are modulated by ordinal position such that RTs to word-initial syllables are notably slower than those to word-medial and word-final syllables (main effect of ordinal position, *X^2 (2) = 523.49, p = < 0.0001,* ). (**Fig. 1**) We conducted pairwise comparisons on estimated marginal means for levels of the factor position with Tukey adjustment, to explore the drop in reaction times between each ordinal position. (**Table 1**. Note that estimates represent differences in estimated marginal means on the response scale in seconds.) The estimated drop in mean RT between positions 1 and 2 was 78 ms, while the drop between positions 1 and 3 was roughly 110 ms. The difference in mean RT between positions 2 and 3 was smaller, at about 32 msec. All comparisons reached statistical significance at the 5% alpha level.





**Fig. 1.**

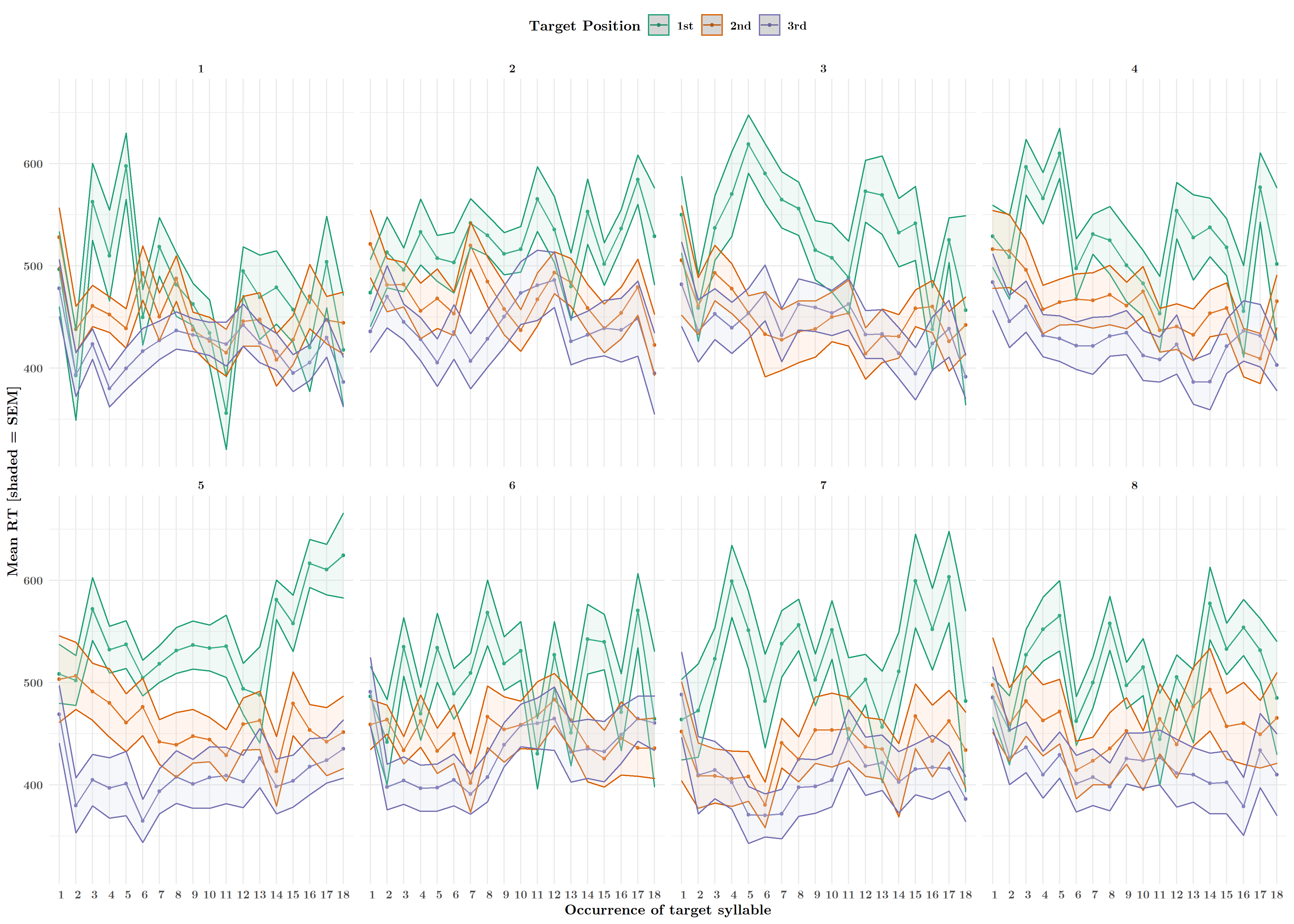
To improve interpretability of the results, we also computed s-values for each p-value in our contrasts. The s-value is a metric of self-information, or surprisal information measure. It can assist in the interpretation of p-values by providing an intuitive transformation of the p-value, taking the negative log of the p-value: . (Greenland et al. 2016) A value of (p = 1) is perfectly unsurprising under the null hypothesis. Surprisal increases exponentially as *s* approaches zero. Thus, *s* becomes a measure of information in bits against the null hypothesis. The difference in means between positions 1 and 3 can be quantified in roughly 139.3 bits. If the null hypothesis of no difference in means is true, this result is as surprising as getting all heads in 140 fair coin tosses (*s* rounded to the nearest integer).

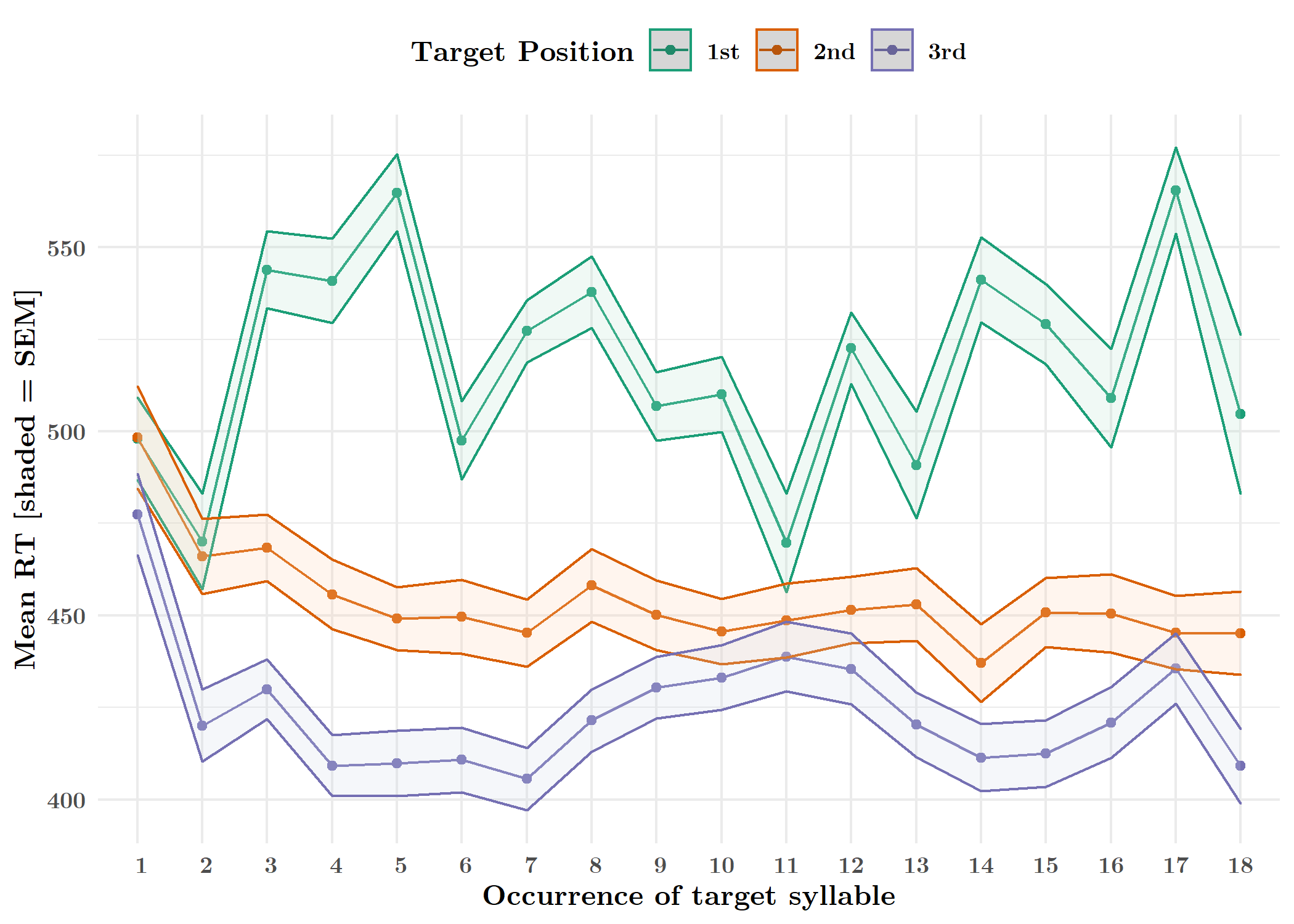
|  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- |
|  | | | | | | | |
| ordinal position | estimate | SE | lower CI | upper CI | z ratio | p value | s value |
|  | | | | | | | |
| 1 - 2 | 0.078 | 0.006 | 0.065 | 0.092 | 13.770 | < .001 | 139.307 |
| 1 - 3 | 0.110 | 0.006 | 0.097 | 0.124 | 18.964 | < .001 | 262.419 |
| 2 - 3 | 0.032 | 0.004 | 0.022 | 0.042 | 7.285 | < .001 | 39.917 |
|  | | | | | | | |

**Table 1. Estimated marginal means contrasts for reaction times to targets in each ordinal position.**

To ensure that participants were able to perform the task, we computed mean detection accuracy across participants and for each target syllable. We observed that certain syllables were detected less often than others, suggesting some unwanted variability in the stimuli (notably for syllables *ro, za, be, and mi).* (See Fig. S2 and following section for discussion.) We therefore sought to validate the results reported above by regressing out the effect of individual syllable as a function of ordinal position. We ran a generalized mixed model with ordinal position and target syllable as fixed effects factors and subject as random effect factor, with the RT (in ms) as outcome variable. We then subtracted the resulting residual values for each data point from the raw RT, and re-ran the lesser model as specified above using the adjusted RT values. We still observed the main effect of ordinal position (*X^2 (2) = 538.53, p < 0.0001, Type II Wald Chisquare Test)* and therefore concluded that slight variations in the acoustics of our stimuli did not significantly affect our results*.*

***Rapid Onset of Graded RT to Predictable Syllables***

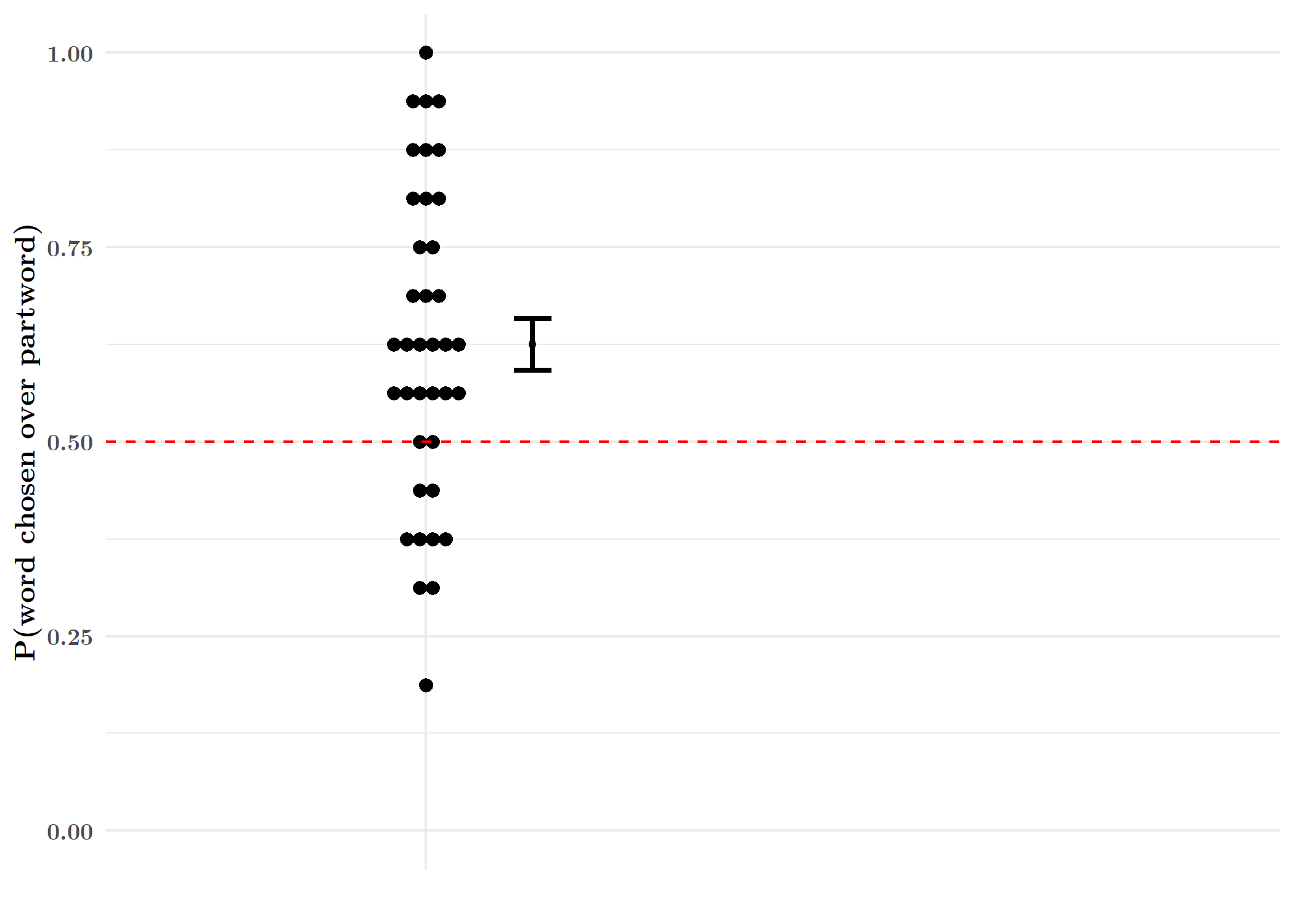
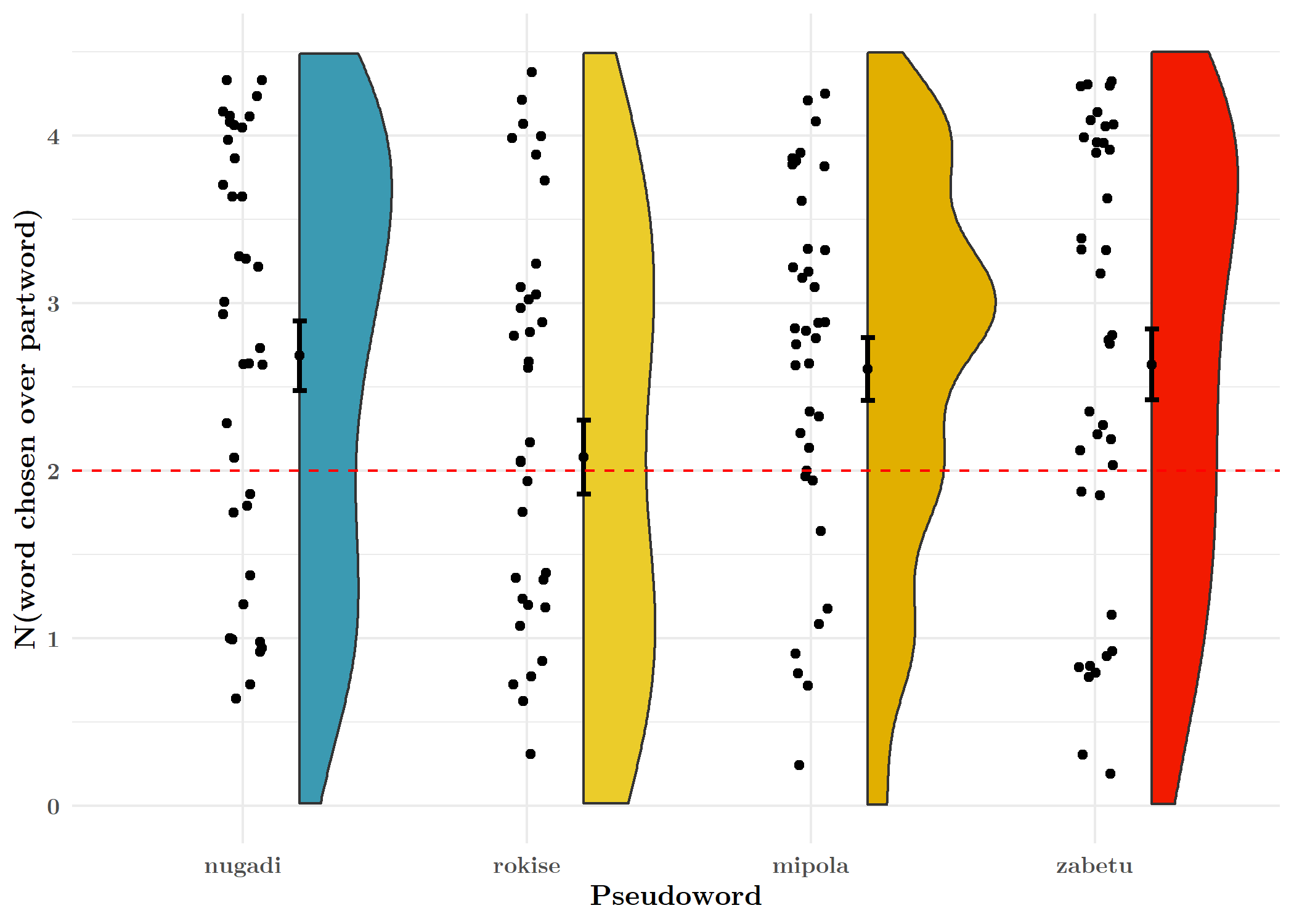
Furthermore, this pattern of reaction times emerged within the first few presentations of the target syllable, and remained stable throughout the remainder of the experimental blocks. (Fig. 2) This observation accounts for why factor block did not contribute significantly to model fit (); reaction times differentiate early on and exhibit little change thereafter.



**Fig. 2**

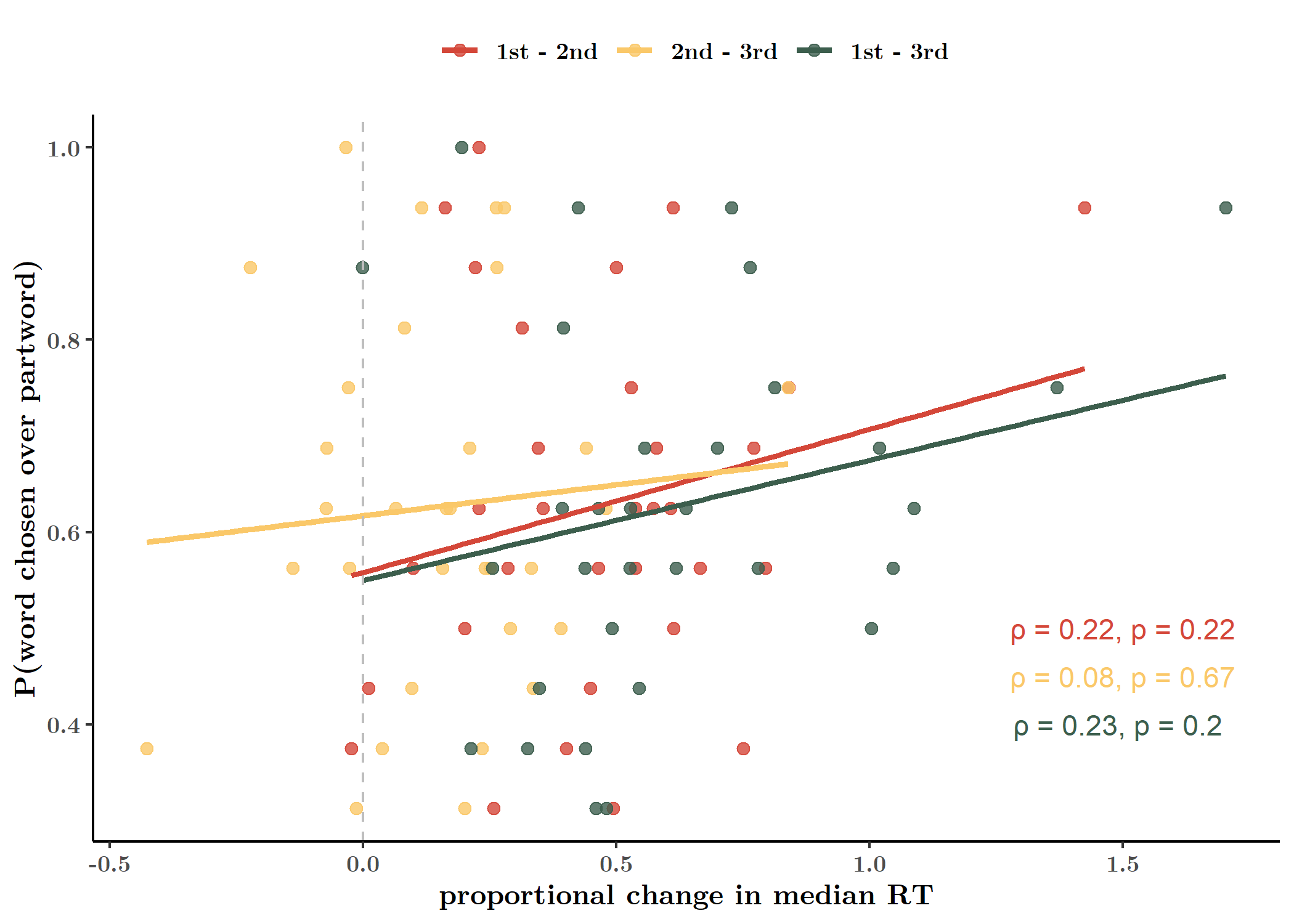
***Psuedowords Can Be Distinguished From Part-words***

In the explicit word recognition task, 71% of participants performed above a 50% chance level (27 out of 38). The proportion of trials on which participants correctly distinguished the pseudoword from the part-word was significantly above a chance level of 0.5 (M = 0.62, SE = 0.2; t(37) = 3.78, p < .001, Cohen’s d = 0.61), indicating that participants were sensitive to the implicit regularities of the syllable stream. **(Fig. 3a)** In an exploratory analysis, we also calculated the proportion correct responses for each psuedoword individually, in order to determine if any particular word was driving the overall word recognition effect. This is because distinguishing only a single word from its paired part-word, would be sufficient to push a participant’s performance above chance level. We found that across participants, 3 out of 4 words were discriminated above chance level of 2 correct discriminations per word. (t\_mipola (37) = 3.24, p = 0.01, t\_nugadi (37) = 3.31, p = 0.008, t\_rokise (37) = 0.36, p = 1.0, t\_zabetu (37) = 2.99, p = 0.02, Bonferroni corrected for four comparisons) (**Fig. 3b**)



**Fig. 3.**

***Online and Offline Measures Are Weakly Correlated***



**Fig. 4. Pearson correlation of online and offline learning measures (median).**

Finally, also asked whether the online and offline measures of statistical learning were correlated, i.e. whether sensitivity to transitional probability online predicts explicit word recognition. Since we had an unequal number of data sets for the two tasks, we used data only from participants with complete data from both tasks (N = 30). For this analysis, we computed the difference in median RTs for each participant and for each ordinal position pair (1-2, 2-3, and 1-3). These values were correlated against the participant’s proportion correct word recognition performance. Surprisingly, word recognition performance (prop. correct responses) and response time change was weakly correlated (Test for association between paired samples of Pearson's product moment correlation coefficient). (**Fig. 4**)

## **Discussion**

Our study replicated two tasks that measure statistical learning in distinct ways. Our offline word recognition task revealed a well-established effect of statistical learning, which is the ability to explicitly discriminate a properly formed pseudoword from a sequence of syllables that was heard but which span a word boundary (a part-word). This pseudoword vs. partword test is more conservative than its sister version, the word vs. nonword test, in which words are tested against random combinations of syllables which never occurred in that particular order during the learning phase. [citation needed]

The online target detection task asks participants to response as soon as they heard a target syllable. Reaction times appeared to be modulated primarily by their transitional probability, which is equivalent to their predictability, such that highly predictable syllables occurring in word-medial and word-final positions elicited much faster responses than less predictable syllables occurring in word-initial positions. The rapid differentiation of response times, in the first block, confirms previous findings that statistical learning is a very fast and robust mechanism. However, it is plausible that segmenting sensory inputs based on transitional probabilities is an easier task than combining those segmented units into a larger whole, such as a word. This explanation may account for the lack of a correlation we observed between our two statistical learning tasks.

Another explanation lies in the fact that the tasks are inherently different by virtue of what they ask of the participant. The online task requires no explicit recall of information learned during the exposure phase, but rather a simple auditory template matching. In contrast, the offline tasks requires participants to take two tokens and determine which is more similar to what they just heard. Since both tokens were indeed heard, success in the task requires not only a tracking of transitional probabilities, but also the inference that these changes in statistics imply natural breaking points in the stream. While much experimental evidence points to a role of statistical structure and surprise in generating implicit event boundaries (Schapiro et al. 2013; Zacks and Swallow 2007), these representations are rather weak and can be easily overridden by subsequent exposure, e.g. to part-words [citation needed].

In a previous studying comparing a similar explicit task of familiarity (of words vs. partwords vs. nonwords) versus change in median RT to word-initial versus word-final syllables, the correlation coefficient was higher, at r = 0.42 (p = 0.044). (Batterink and Paller 2017) This correlation coefficient is more similar to what we observed if we compared word recognition performance with the differences in mean RT to each position (r = 0.34, **Fig. S6**), but this relationship remained weak for our data (p = 0.05). However, in another paper published by the same group, the same analysis yielded no significant correlation. Despite finding a similar reaction time effect, no correlation was observed between the median RT to word-initial versus word-final syllables and accuracy on the familiarity task (r = 0.07, n.s.) (Batterink et al. 2015) An second experiment in the same paper reports a correlation of r = 0.26 (n.s.) for a different cohort of participants.

Taken together, these results provide clear evidence that the online and offline tasks do not measure the same phenomenon.

We wished to investigate why graded response times in the online task failed to adequately predict offline pseudoword recognition. The syllables in the stream are each characterized by several features: ordinal position, transitional probability, within-word duplet pairing, and within-word triplet pairing. Success on the online tasks requires only a tracking of transitional probability, while success in the offline task requires a representation of the within-word triplet pairing (in other words, the representation of the tri-syllablic pseudoword as a single unit). Yet, other representations, as listed above, are also possible.

# **Experiment 2**

## **Method**

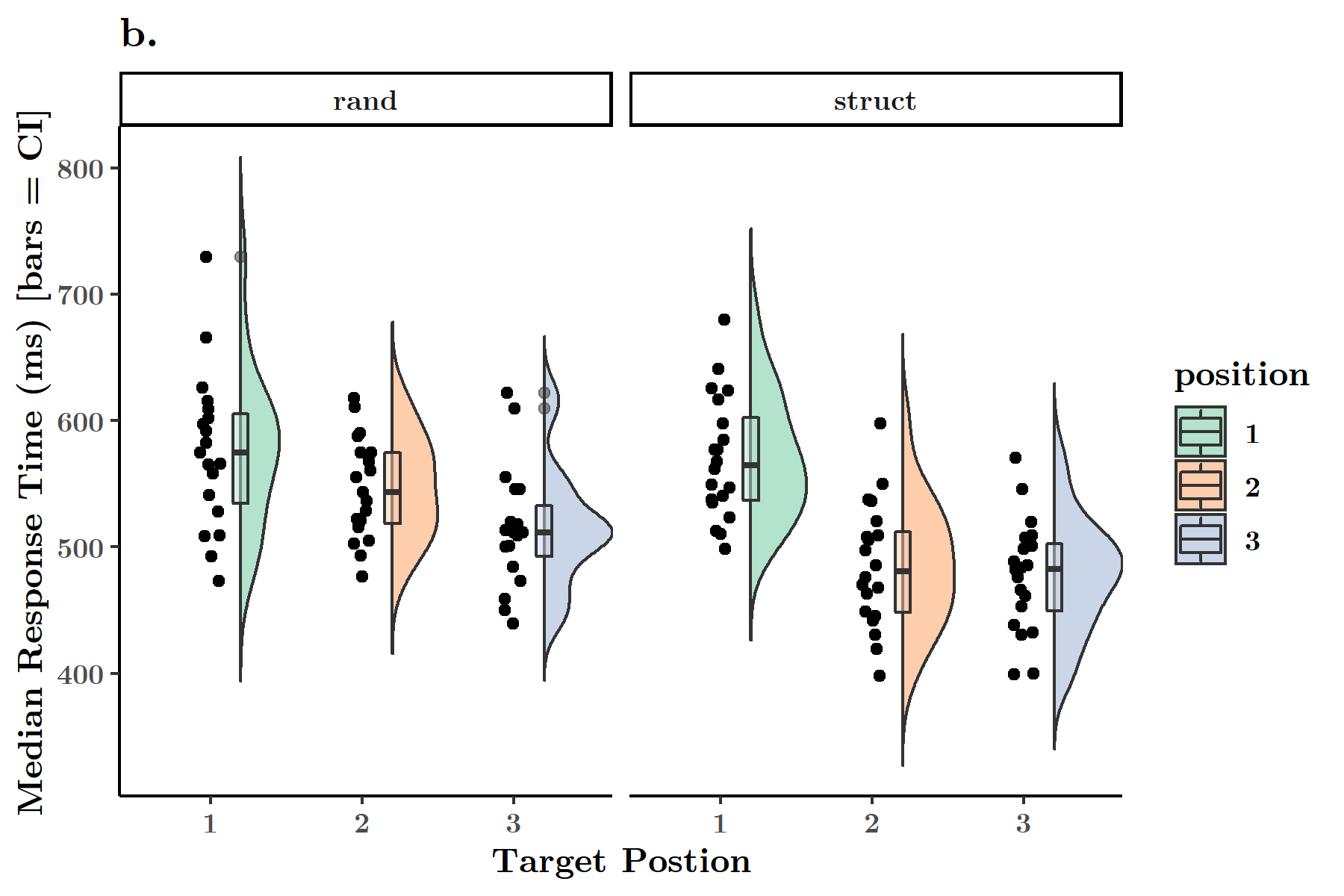
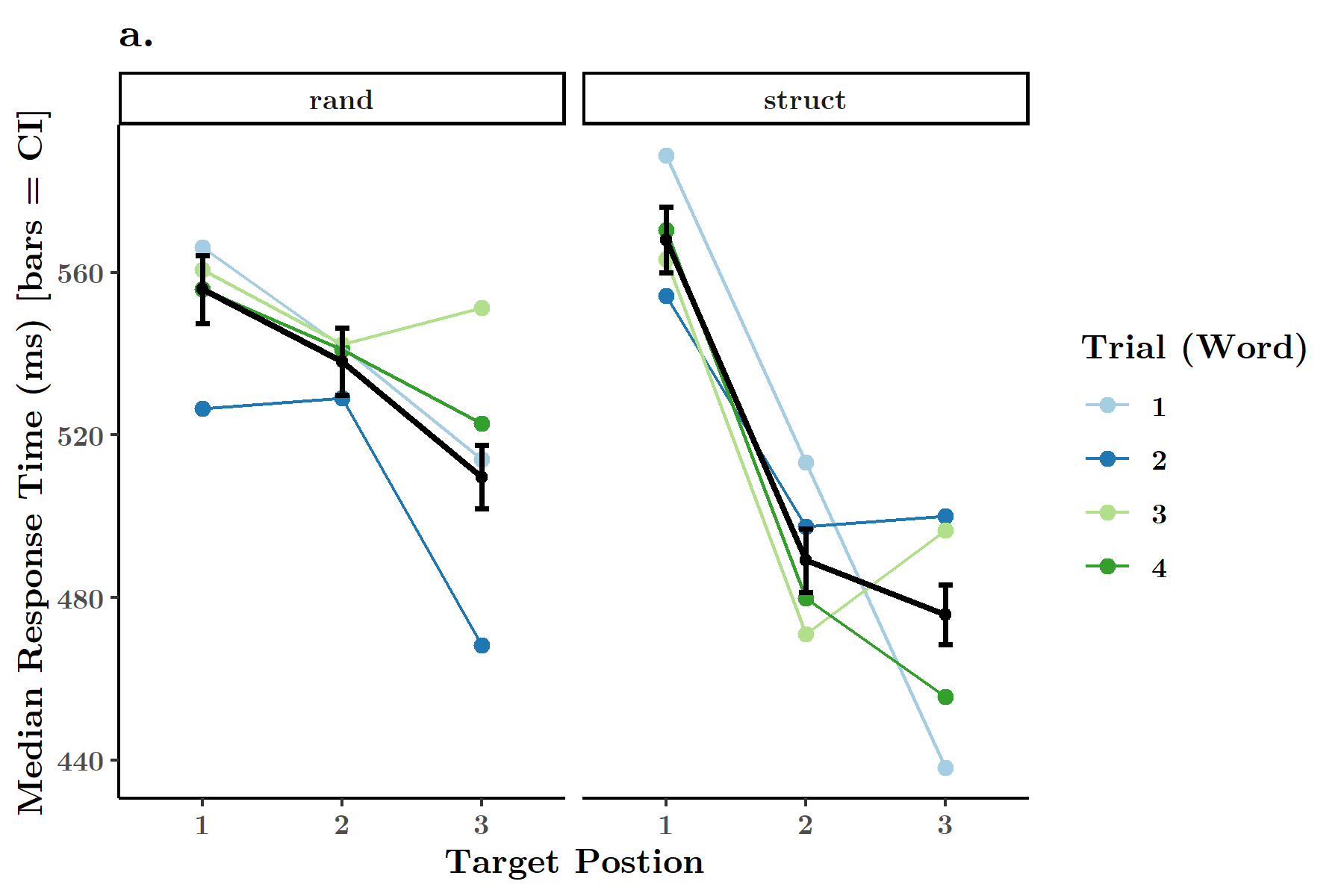
Procedure

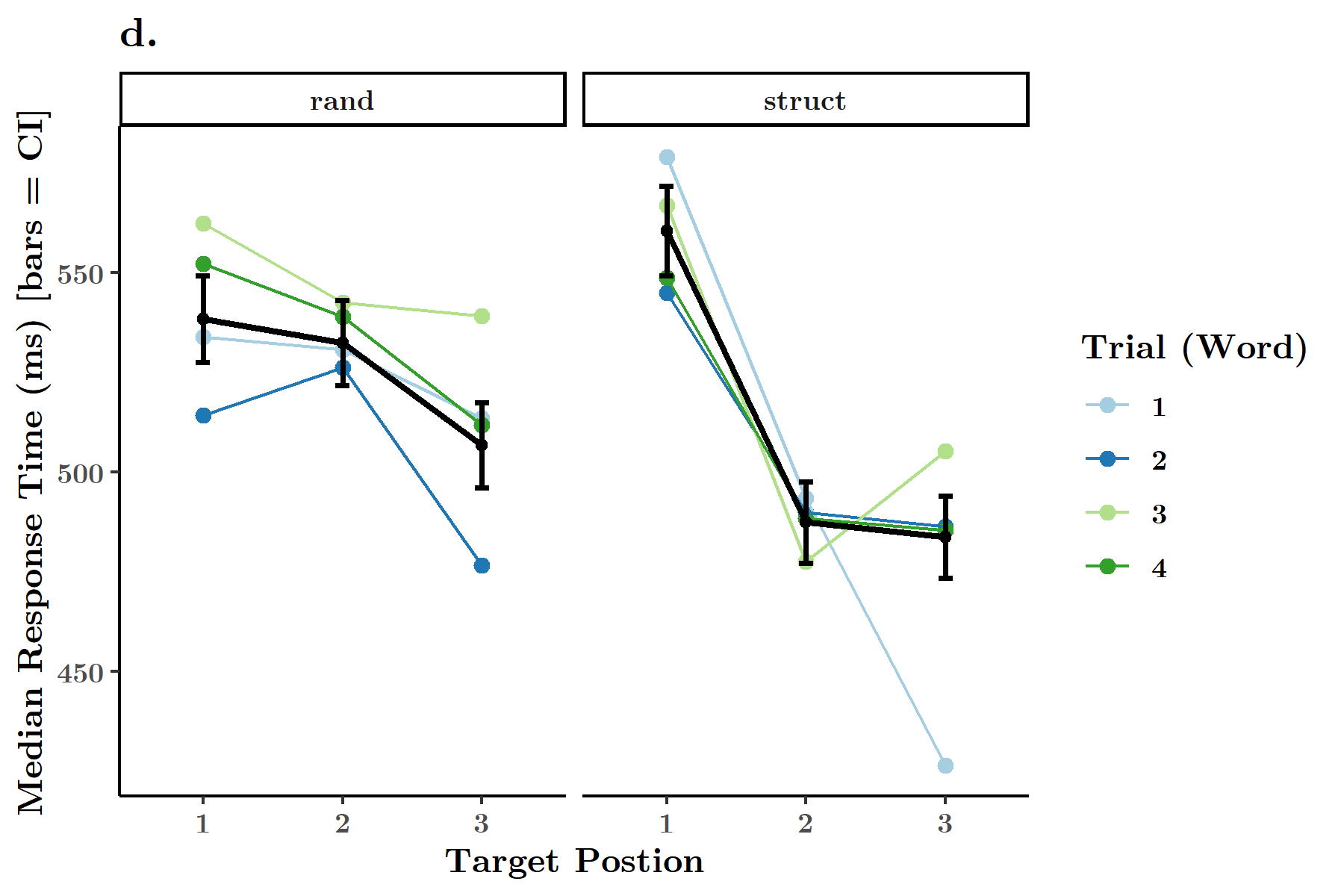
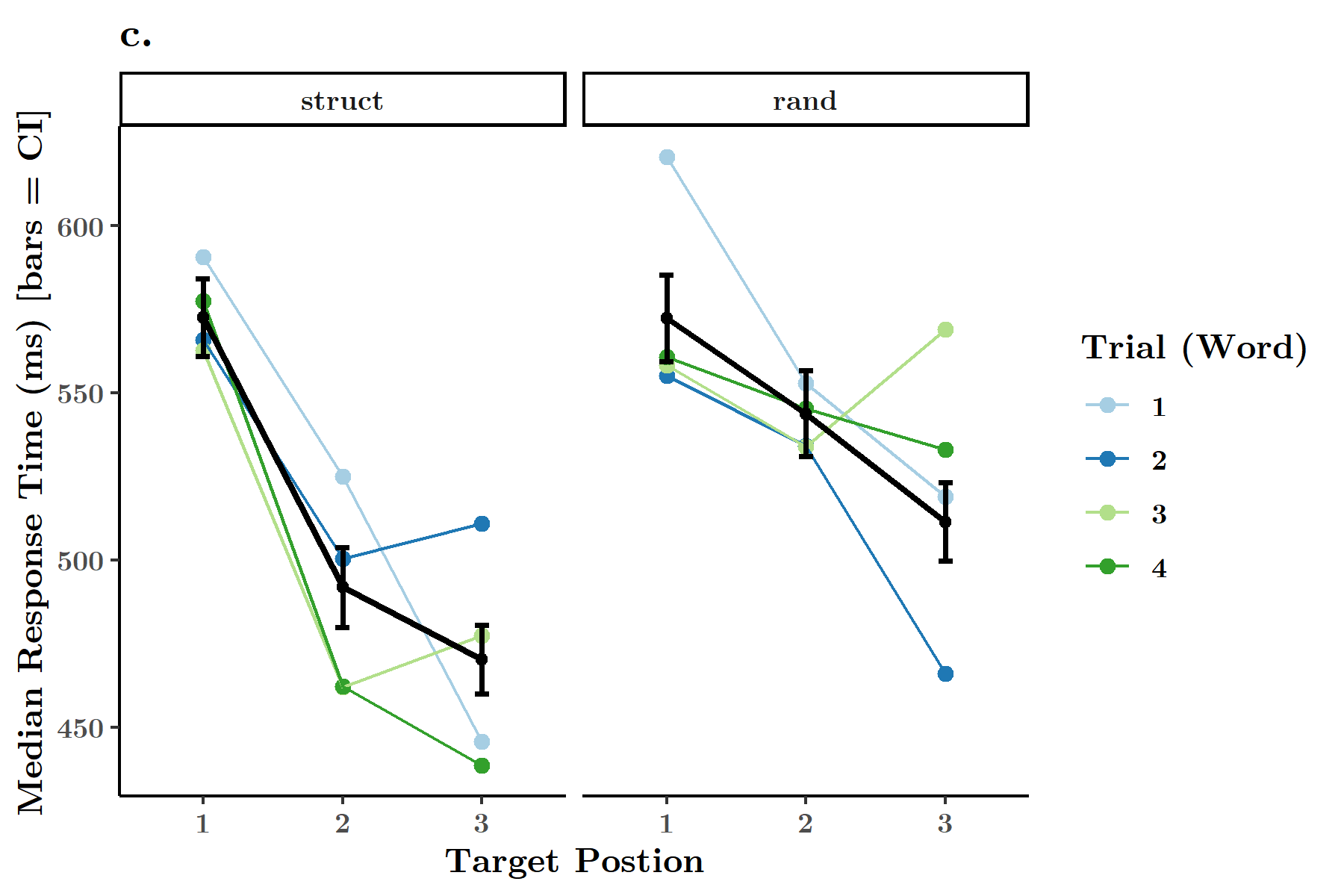
The stimuli used in Experiment 2 were identical to those used in Experiment 1. For this experiment we synthetized 12 structured streams (as per the method above) and 12 random streams (wherein syllables are pseudorandomly ordered).

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## **Results**

***Ordinal Position in Structured Stream Modulates Reaction Time***

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**Fig. 6.**

We performed a modelling procedure similar to that from Experiment 1. Our hypothesis entailed response times (in seconds) would be predicted by a combination of ordinal position and session (structures or random). We included subject as a nested effect within session order (whether participants completed the structured session before the random session, and vice versa), as the latter was our between-subjects variable. We further specified the random effects term by allowing random intercepts and uncorrelated to this, random effects for each level of session. This structure allows the graded RT curve for each participant to vary between sessions, as well as their baseline RT (intercept). (See **Table S2** for regression results.) We observed main effects of both session and ordinal position, as well as an interaction between them ().

Two performed two planned contrasts, the first to evaluate the effect of ordinal position within each level of session (i.e. to determine the modulation of reaction times for each condition). The second evaluated the effect of session for each level of ordinal position (i.e. how much session affected reaction times to targets in each ordinal position). When examining the effect of position within session, we observed a significant drop in estimated means between positions 1 (M = 0.581, SE = 0.011) and 2 (M = 0.565, SE = 0.011), as well as between 1 and 3 (M = 0.534, SE = 0.01) in the structured condition. There was a smaller, but also statistically significant decrease in means between positions 2 - 3 (**Table 2**.). In the random condition, we were surprised to observe a similar pattern, where differences in estimated means for each pair of positions reached still significance, with the smallest change occurring between positions 1 (M = 0.582, SE = 0.011) and 2 (M = 0.510, SE = 0.01); and position 3 syllables having shorter RTs (M = 0.489, SE = 0.009). (**Table 2,** **Fig. 7a**) Given that there were no regularities in the random stream that could bias reaction times to certain tokens more than others, we hypothesized that the modulation observed here is due to variations in the acoustic features of the stimuli, as noted above.

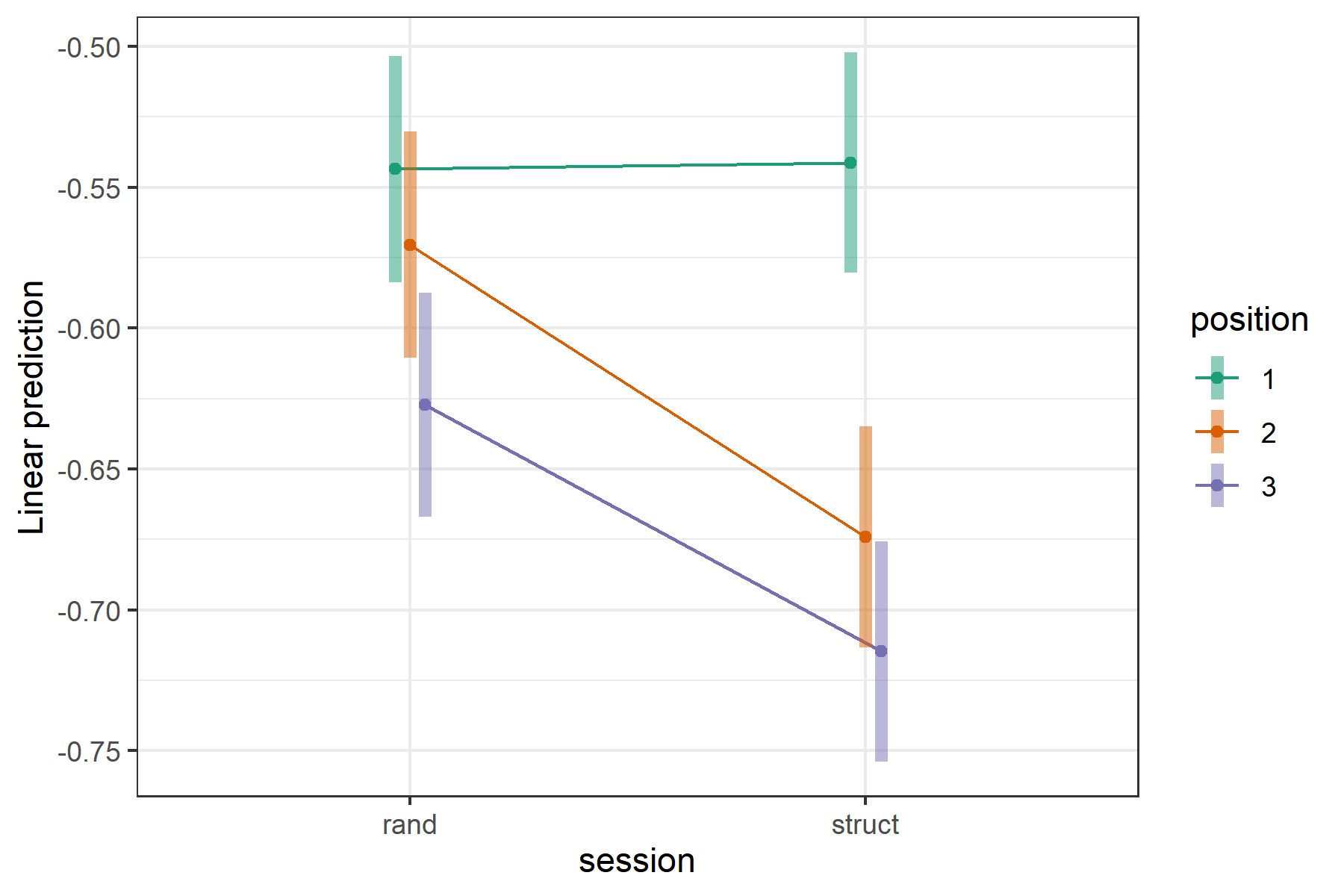
|  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  | | | | | | | | |
| ordinal position | session | estimate | SE | lower CI | upper CI | z ratio | p value | s value |
|  | | | | | | | | |
| 1 - 2 | rand | 0.015 | 0.006 | 0.001 | 0.030 | 2.453 | 0.038 | 4.729 |
| 1 - 3 | rand | 0.047 | 0.006 | 0.032 | 0.061 | 7.731 | < 0.001 | 44.190 |
| 2 - 3 | rand | 0.031 | 0.006 | 0.017 | 0.045 | 5.324 | < 0.001 | 21.650 |
| 1 - 2 | struct | 0.072 | 0.006 | 0.059 | 0.086 | 12.690 | < 0.001 | Inf |
| 1 - 3 | struct | 0.093 | 0.006 | 0.079 | 0.106 | 16.370 | < 0.001 | Inf |
| 2 - 3 | struct | 0.020 | 0.005 | 0.008 | 0.032 | 4.006 | < 0.001 | 12.420 |
|  | | | | | | | | |

**Table 2.**

However, this observation alone only suggests that variability in stimuli caused some variability in response. When examining the effect of session (presence or absence of statistical structure) for each level of ordinal position, we observed that the presence of structure significantly decreased mean reaction times for 2nd (z(Inf) = 5.28, p < 0.0001) and 3rd position targets (z(Inf) = 4.51, p < 0.0001). While RTs to 1st position targets remained largely the same (reaction times for 1st position targets in fact increased, but this change was not statistically significant; z(Inf) = -0.12, p = 0.90). (**Fig. 7**)

We additionally tested a model that included session order as a fixed effect, to ensure that there were no carry-over effects of session. Previous studies have indicated that statistical learning can be easily disrupted by exposure to streams with differing or no statistics. [citation needed] We did not detect a three-way interaction between session order, ordinal position, and session in modulating response times (). (**Fig. 6 c-d**, **Fig. S5**)

**Fig. 7**



## **Discussion**

Here we were able to replicate our findings from Experiment 1, in which the presence of implicit statistical structure modulates reaction times to items embedded in the stream. Notably, we could establish that this modulation is most extreme for those items which become strictly predictable (those with a transitional probability of 1, in word-medial and word-final positions), as compared with the same items tested in a stream of randomly ordered syllables.

# **Coding of Structural Features in Online Target Detection**

In a final analysis, we aimed to explore the reaction time data from the online target detection task further, in order to determine if patterns of RTs could reveal any particular coding strategies which might lend clues to what features of the implicit structure participants picked up.

Specifically, we hypothesizes that one of four possible features could be encoded in reaction time data: transitional probability, ordinal position, word identity, and duplets.

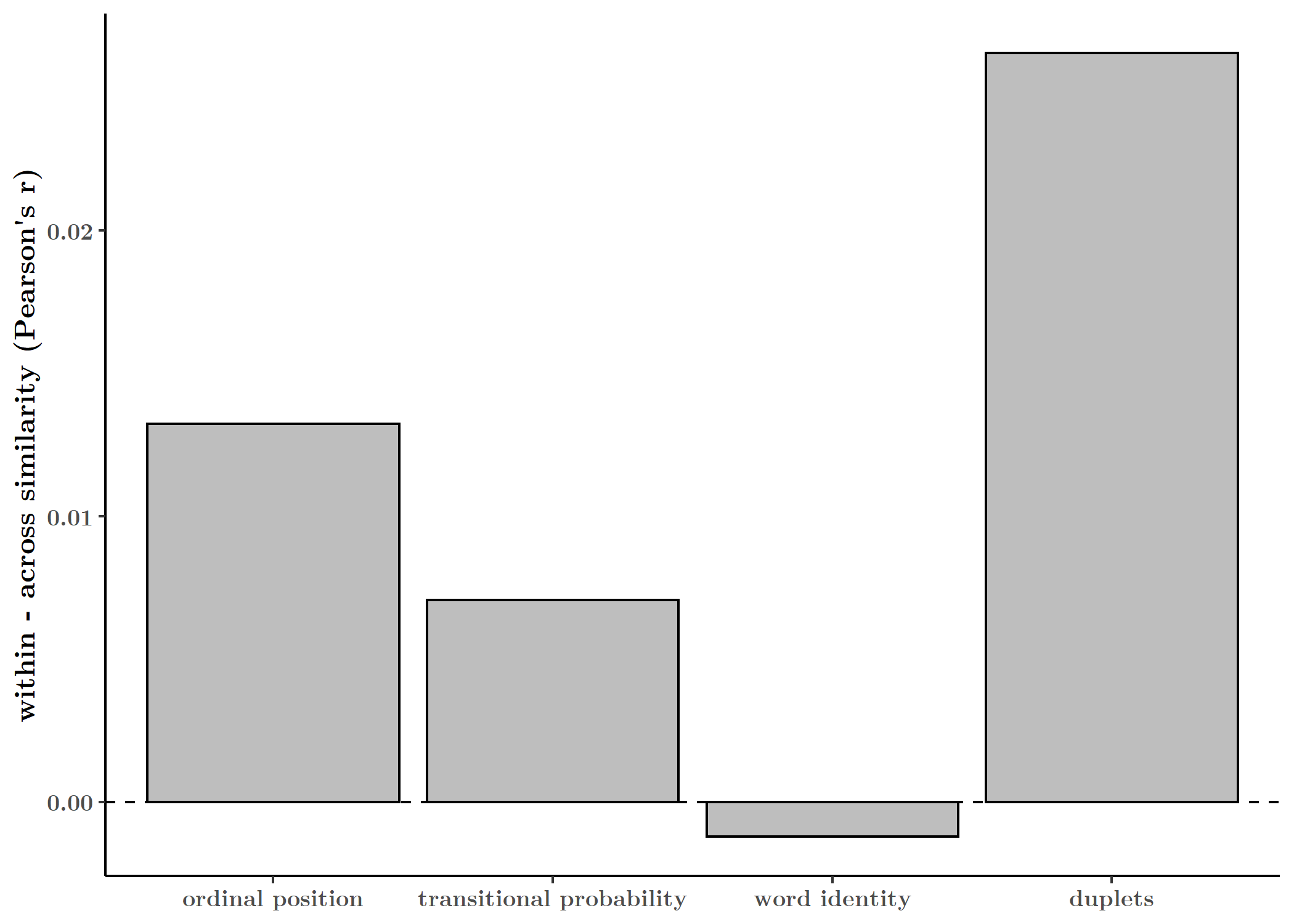
We first combined the data from Experiment 1 (N = 33) with the data from structured sessions in Experiment 2 (N = 20) for a combined data set with greater power (N = 53). For each participant, we computed the Pearson correlation on RTs between each pair of syllables, thus generating a 12-x-12 matrix of correlation values for each participant. For each of the four analyses mentioned above, we identified those cells that would constitute the two data arrays for each analysis. Finally, we performed a random sampling of values from those pre-specified cells from all participants 200 times (with replacement, the N for each sample was equal to 4/5 times the length of the shortest of the two arrays being compared). We then computed a Wilcoxon’s rank sum test on the complete set of sampled values.

For the test of transitional probability, within values included the correlation between all pairs of word-initial syllables (TP = 0.33) versus the correlation between all pairs of word-medial and word-final syllables (TP = 1). Here we observed a significant shift in the true location difference of the two arrays. (W(53) = 1303113221, p-value < 0.0001, Cis: 0.0175, 0.0269; estimates of difference in location 0.022).

For the test of ordinal position, within values included the correlation between all pairs of word-initial syllables (e.g. nu-ro), all pairs of word-medial syllables (e.g. ga-ki), and all pairs of word-final syllables (e.g. di-se) versus correlations between syllables within each word (e.g. nu-ga, ga-di, nu-di). Here we also observed a significant shift in means between the two groups. (W(53) = 4995263642, p-value < 0.0001, Cis: 0.007771155, 0.014582123; estimates of difference in location 0.011).

For the test of word identity, within values included the correlation between syllables within each word (e.g. nu-ga, ga-di, nu-di) versus “phantom” word pairs where each item in the pair is drawn from two different words (e.g. nu-ki, nu-se). Here we observed no significant difference. (W(53) = 4891836441, p-value = 0.4957, Cis: -0.0024, 0.0002; estimates of difference in location: -1.26268e-05).

Finally, to test duplet identity, we compared values from all pairs of duplets within words (e.g. nu-ga, ga-di) versus pairs of word-initial and word-final syllables within words (e.g. nu-di). Our Wilcoxon test suggested the two groups differ in their means. (W(53) = 557892245, p-value = 0.006, Cis: 0.0007, 0.0118; estimates of difference in location: -0.0058).



# **General Discussion**

Tktkt…

# **Conclusion**

Tktkt….

# **Acknowledgements**

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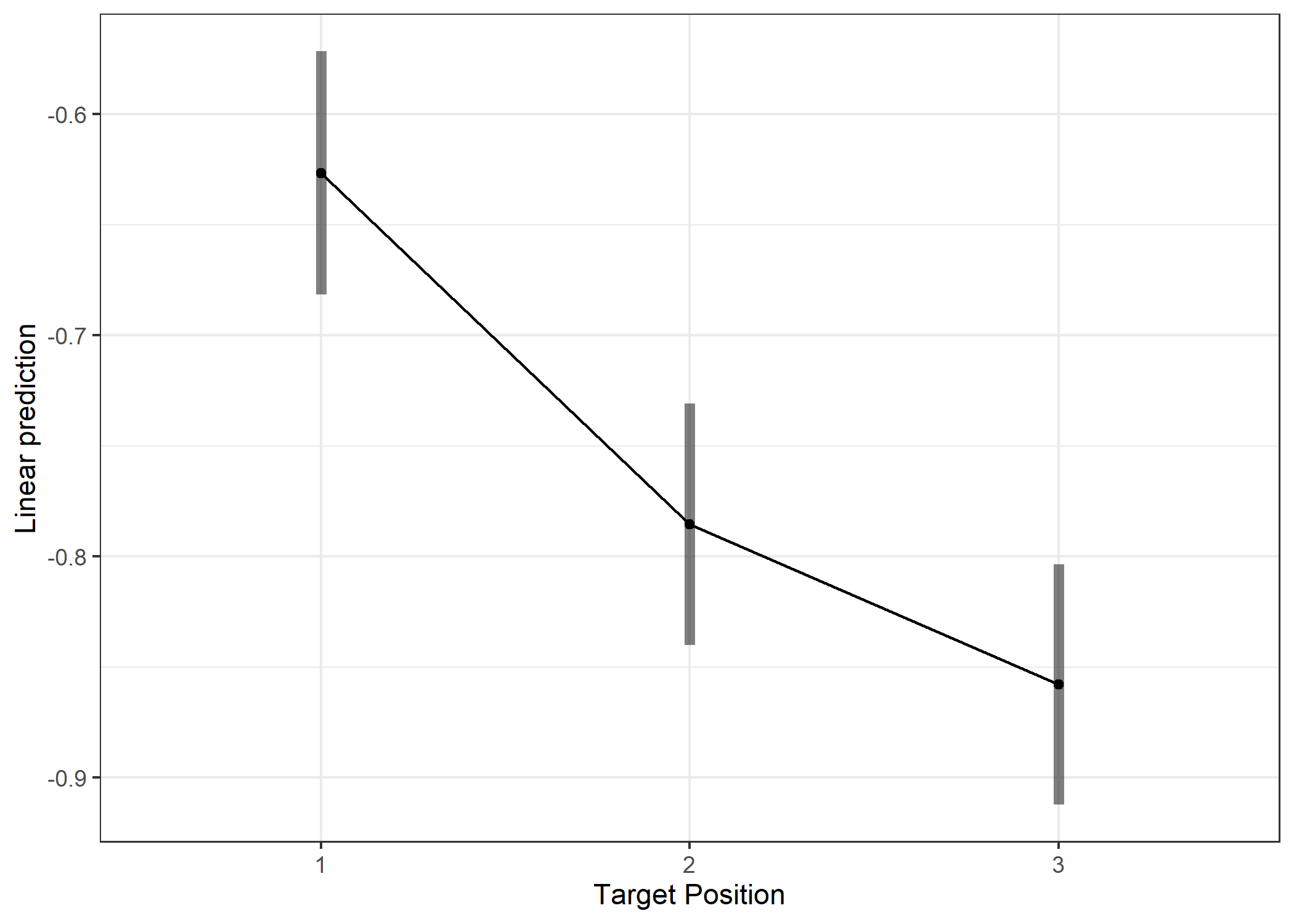
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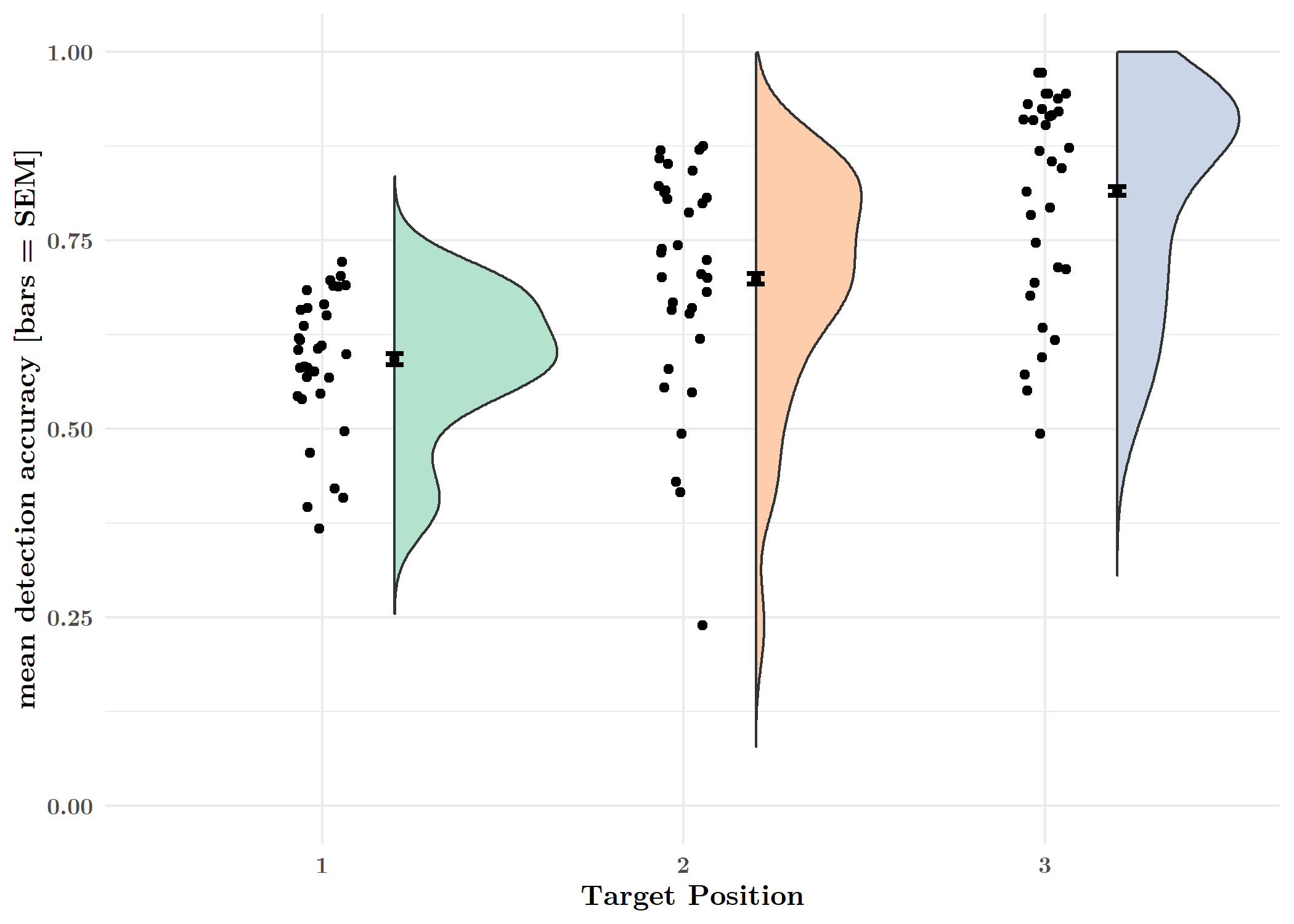
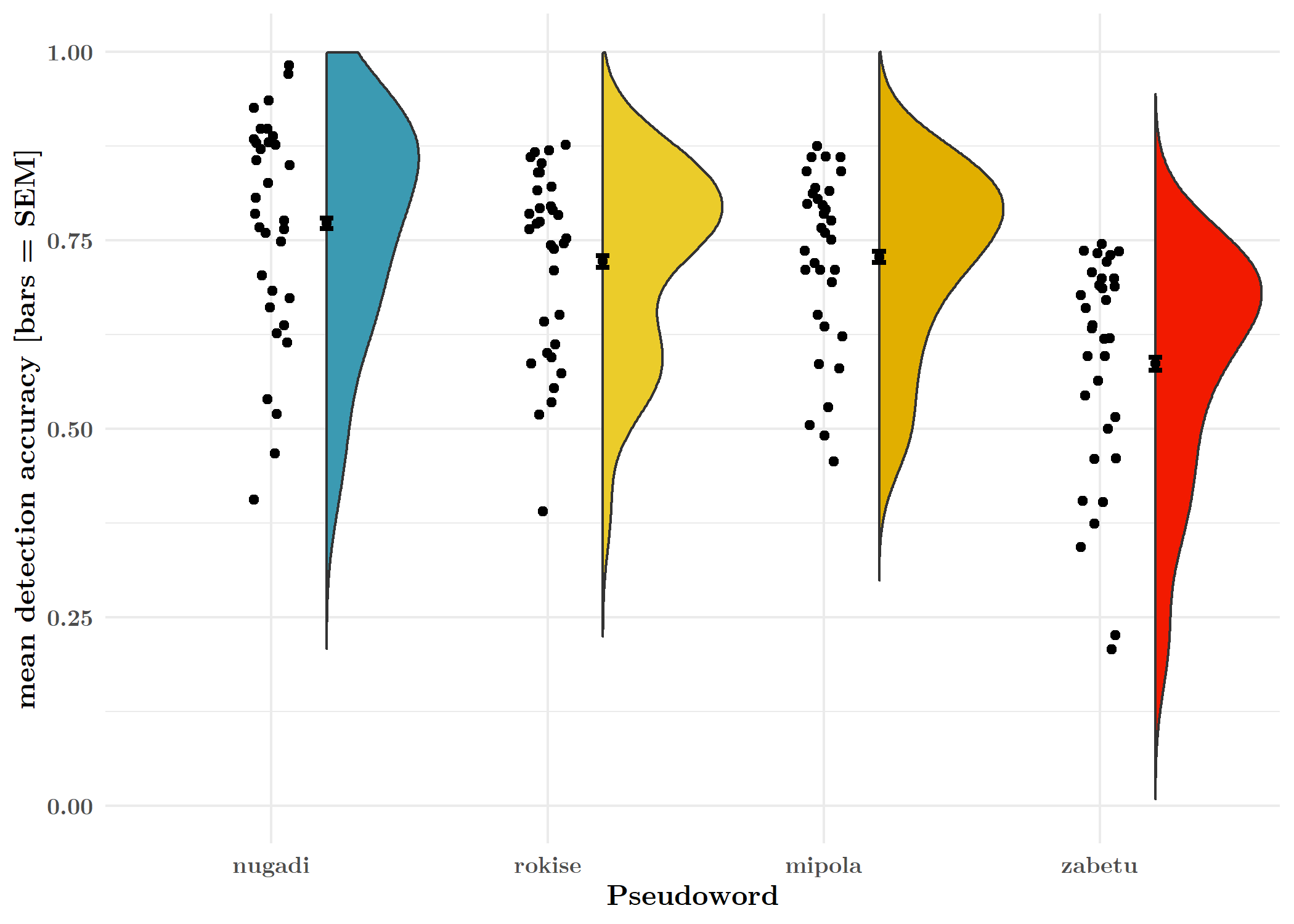
# **Supplementary Materials**

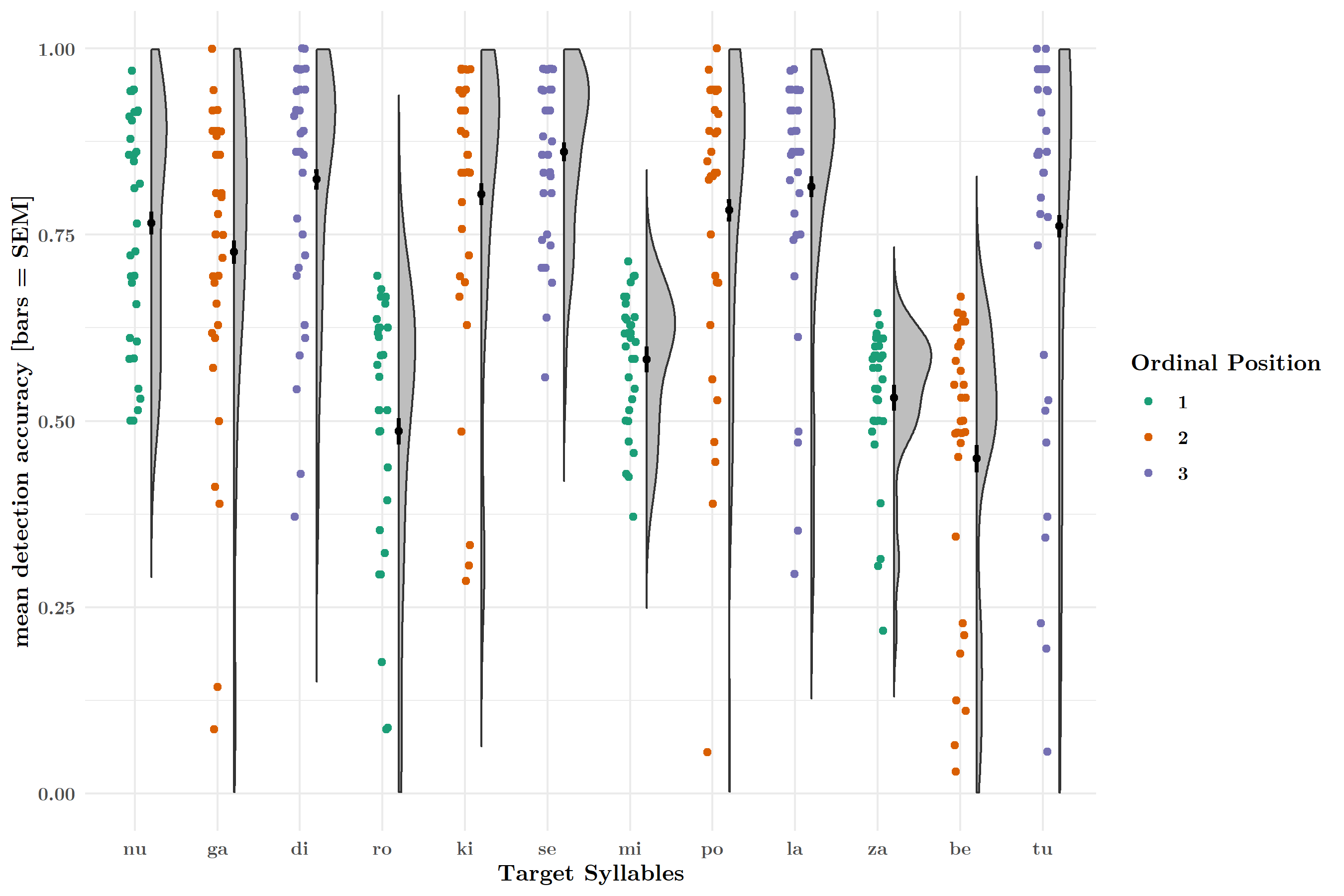
## **Stimuli**

## **Supplementary Figures, Tables**



**Fig. S1. Predicted coefficients from GLM for each ordinal position.**

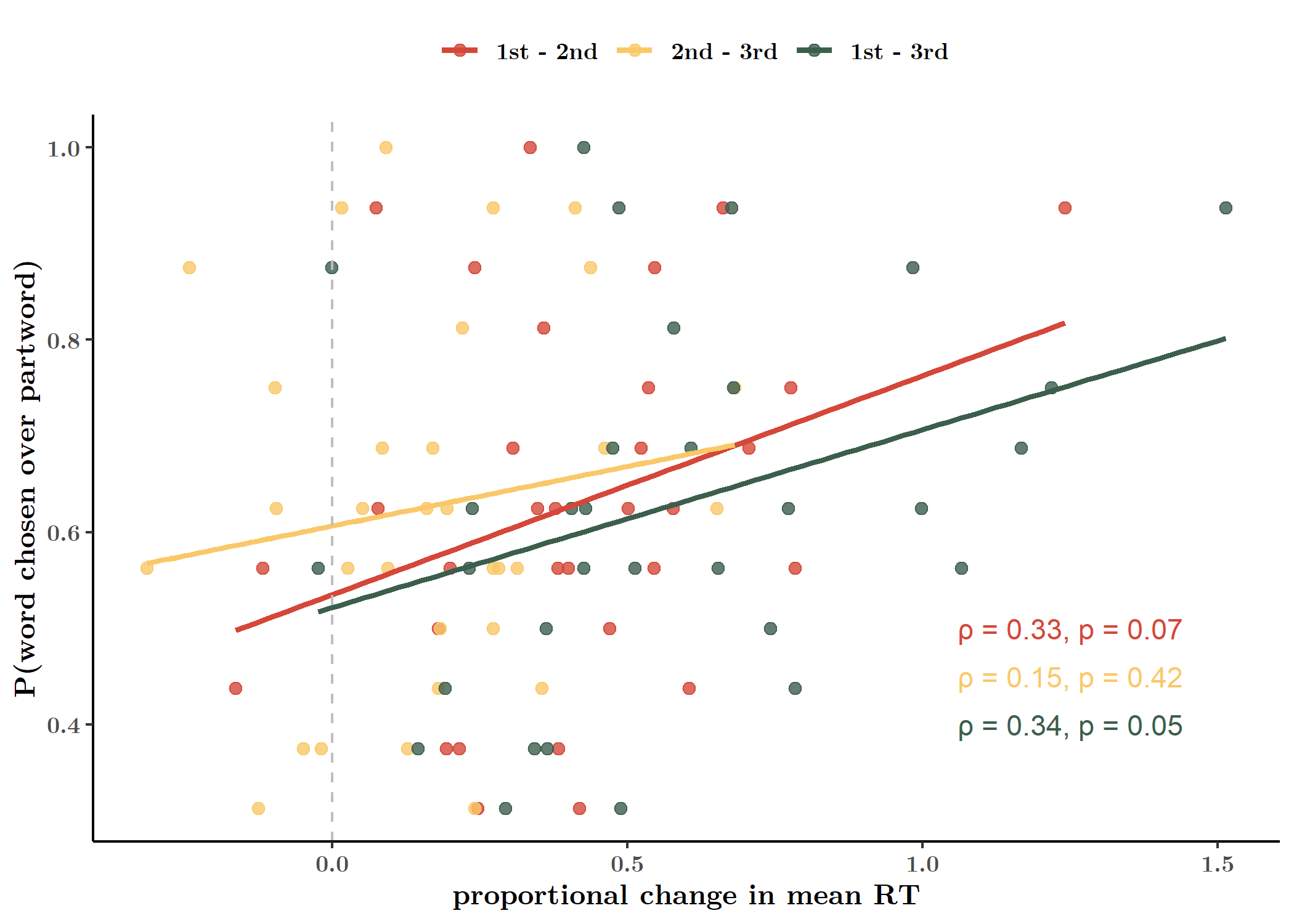
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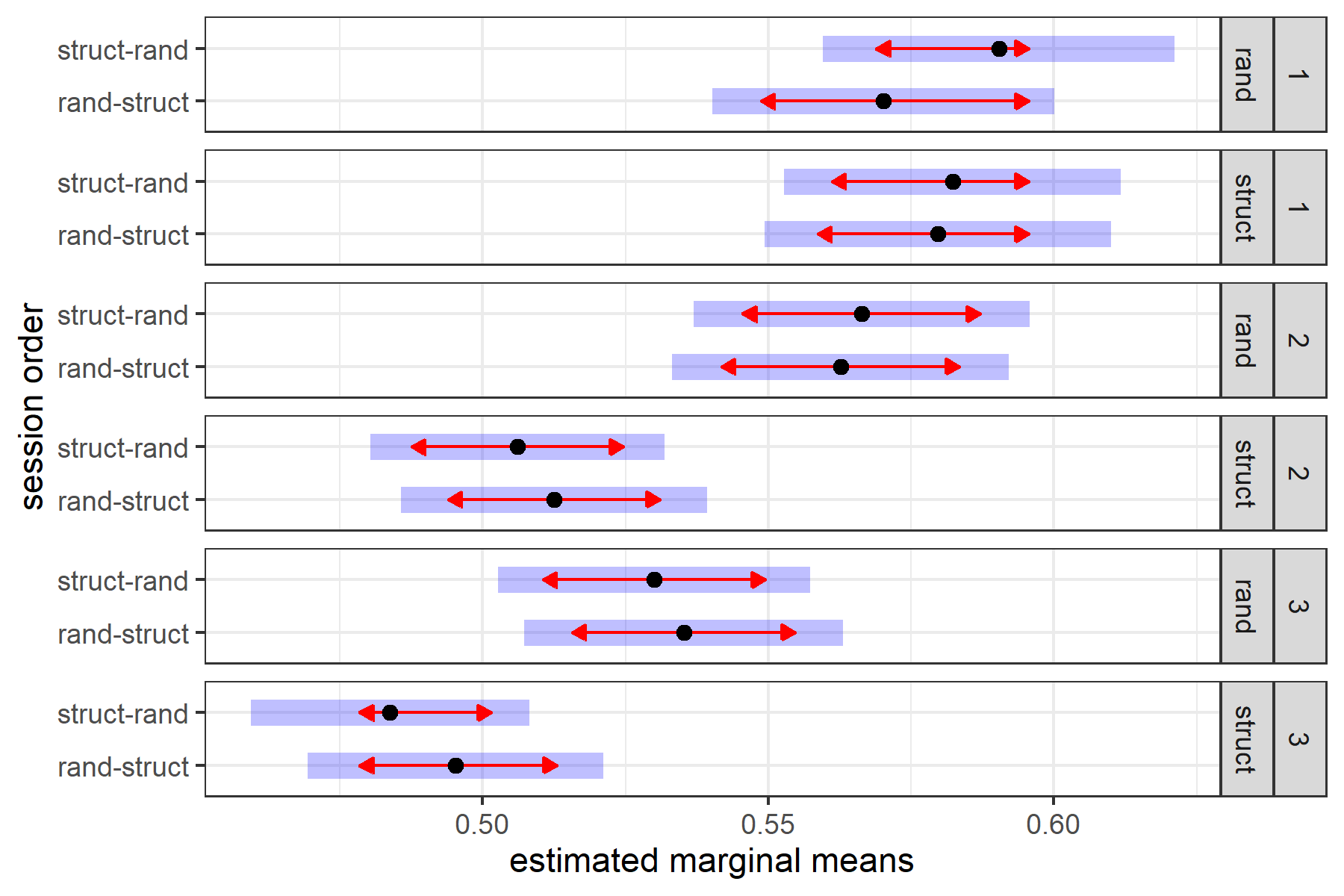
**Fig. S2. Mean detection accuracy for each ordinal position, syllable and pseudoword.**

We quantified accuracy in the target detection task to ensure participants complied with task instructions. The hit rate in Experiment 1 was 0.70 (sd = 0.45). The hit rate was also modulated by ordinal position, with each successive position having a higher mean accuracy (MPos 1 = 0.59, sdPos 1 = 0.49; MPos 2 = 0.70, sdPos 2 = 0.46; MPos 3 = 0.82, sdPos 3 = 0.39). (Fig. S2a) When averaging across all syllables in a pseudoword, accuracy varied between the four words. (Fig. S2b) This effect appears to be driven by differences in recognizability of individual syllables; certain CV syllable pairs may have been easier to detect than others, due to minor variations in stimuli acoustics. (Fig. S2c)

|  |  |  |  |
| --- | --- | --- | --- |
| **Table S1. GLM Results** | | | |
|  | | | |
|  | *Dependent variable:* | | |
|  |  | | |
|  | reaction time (s) | | |
|  | lesser | lesser (random slopes) | fuller |
|  | (1) | (2) | (3) |
|  | | | |
| Intercept(Pos 1) | -0.63\*\*\* (-0.68, -0.57) | -0.64\*\*\* (-0.68, -0.61) | -0.69\*\*\* (-0.76, -0.62) |
| Pos 2 | -0.16\*\*\* (-0.18, -0.14) | -0.14\*\*\* (-0.19, -0.10) | -0.11\*\*\* (-0.17, -0.05) |
| Pos 3 | -0.23\*\*\* (-0.25, -0.21) | -0.22\*\*\* (-0.28, -0.15) | -0.18\*\*\* (-0.24, -0.12) |
| Block 2 |  |  | 0.07\*\* (0.01, 0.13) |
| Block 3 |  |  | 0.08\*\*\* (0.02, 0.15) |
| Block 4 |  |  | 0.08\*\* (0.02, 0.14) |
| Block 5 |  |  | 0.11\*\*\* (0.05, 0.17) |
| Block 6 |  |  | 0.03 (-0.03, 0.09) |
| Block 7 |  |  | 0.05 (-0.01, 0.11) |
| Block 8 |  |  | 0.06\* (-0.005, 0.12) |
| Pos 2:Block 2 |  |  | -0.04 (-0.12, 0.04) |
| Pos 3:Block 2 |  |  | -0.03 (-0.11, 0.05) |
| Pos 2:Block 3 |  |  | -0.07\* (-0.16, 0.01) |
| Pos 3:Block 3 |  |  | -0.03 (-0.11, 0.05) |
| Pos 2:Block 4 |  |  | -0.05 (-0.13, 0.04) |
| Pos 3:Block 4 |  |  | -0.07 (-0.15, 0.01) |
| Pos 2:Block 5 |  |  | -0.06 (-0.14, 0.03) |
| Pos 3:Block 5 |  |  | -0.14\*\*\* (-0.22, -0.06) |
| Pos 2:Block 6 |  |  | -0.01 (-0.10, 0.07) |
| Pos 3:Block 6 |  |  | 0.01 (-0.06, 0.09) |
| Pos 2:Block 7 |  |  | -0.06 (-0.15, 0.02) |
| Pos 3:Block 7 |  |  | -0.09\*\* (-0.17, -0.004) |
| Pos 2:Block 8 |  |  | -0.06 (-0.14, 0.03) |
| Pos 3:Block 8 |  |  | -0.06 (-0.14, 0.02) |
|  | | | |
| Fixed Effects | Subject | Position | Subject | Subject |
| Fixed Effects Struct. | Rand. Int. | Rand. Int., Slope | Rand Int. |
| Observations | 9,531 | 9,531 | 9,531 |
| Log Likelihood | 3,199.61 | 3,312.13 | 3,224.15 |
| Akaike Inf. Crit. | -6,389.23 | -6,604.26 | -6,396.29 |
| Bayesian Inf. Crit. | -6,353.42 | -6,532.64 | -6,210.07 |
|  | | | |
| *Note:* | \*p\*\*p\*\*\*p<0.01 | | |
|  | Fitted using Gamma distribution and log link function. | | |



**Fig. S3. Pearson correlation of online and offline learning measures (mean).**

****

**Fig. S4. No interaction of session order with session and ordinal position**

|  |  |
| --- | --- |
| **Table S2. GLM Results** | |
|  | |
|  | *Dependent variable:* |
|  |  |
|  | reaction time (s) |
|  | |
| Intercept(Pos 1/Rand) | -0.54\*\*\* (-0.58, -0.50) |
| Struct | 0.002 (-0.04, 0.04) |
| Pos 2 | -0.03\*\* (-0.05, -0.01) |
| Pos 3 | -0.08\*\*\* (-0.10, -0.06) |
| Struct:Pos 2 | -0.11\*\*\* (-0.14, -0.08) |
| Struct:Pos 3 | -0.09\*\*\* (-0.12, -0.06) |
|  | |
| Fixed Effects | Session Order/Subject + Session |
| Fixed Effects Struct. | Rand. Int. + Rand. Slope |
| Observations | 6,524 |
| Log Likelihood | 4,140.28 |
| Akaike Inf. Crit. | -8,250.56 |
| Bayesian Inf. Crit. | -8,148.81 |
|  | |
| *Note:* | \*p\*\*p\*\*\*p<0.01 |
|  | Fitted using Gamma distribution and log link function. |

**Fig. S5.**

