

# Applied Bayesian Analysis : NCSU ST 540

## Homework 5

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let  $Y_i$  be the number of concussions from team  $i = 1, \dots, 32$ . The model is

$$Y_i | \lambda_i \sim \text{Poisson}(\lambda_i)$$

and the prior is

$$\lambda_i | \theta \sim \text{Gamma}(1, \theta)$$

where

$$\theta \sim \text{Gamma}(0.1, 0.1)$$

.

### Derive the full conditional distribution of $\lambda_1$

Assuming independence of  $Y_i$  we can write the full joint distribution as

$$P(Y_1, \dots, Y_n, \lambda_1, \dots, \lambda_n, \theta) \propto \prod_{i=1}^n P(Y_i | \lambda_i) P(\lambda_i | \theta) P(\theta)$$

Now

$$P(\lambda_1 | Y_1, \dots, Y_n, \lambda_2, \dots, \lambda_n, \theta) \propto P(Y_1 | \lambda_1) P(\lambda_1 | \theta) P(\theta) \prod_{i=2}^n P(Y_i | \lambda_i) P(\lambda_i | \theta)$$

Putting the expressions for the densities in here and dropping the product terms on the right hand side unrelated to  $\lambda_1$  we have that

$$P(\lambda_1 | Y_1, \dots, Y_n, \lambda_2, \dots, \lambda_n, \theta) \propto \frac{\lambda_1^{y_1}}{y_1!} e^{-\lambda_1} \theta e^{-\theta \lambda_1} \frac{(0.1)^{0.1}}{\Gamma(0.1)} \theta^{0.1-1} e^{-0.1 \theta} \propto \lambda_1^{(y_1+1)-1} e^{-\lambda_1(1+\theta)}$$

The last expression we recognise as the kernel of a  $\text{Gamma}(y_1 + 1, 1 + \theta)$  distribution.

### Derive the full conditional distribution of $\theta$

From above, we have that

$$P(\theta | Y_1, \dots, Y_n, \lambda_1, \dots, \lambda_n) \propto P(\theta) \prod_{i=1}^n P(\lambda_i | \theta)$$

and putting the expression for the densities in we have

$$P(\theta|Y_1, \dots, Y_n, \lambda_1, \dots, \lambda_n) \propto \prod_{i=1}^n \theta e^{-\theta \lambda_i} \frac{(0.1)^{0.1}}{\Gamma(0.1)} \theta^{0.1-1} e^{-0.1 \theta} \propto \theta^{n-.9} e^{\theta(0.1+\sum \lambda_i)}$$

We recognise this expression as the kernel of a  $\text{Gamma}(n + 0.1, 0.1 + \sum \lambda_i)$  density.

**Write Gibbs sampling code to draw samples from the joint distribution of  $(\lambda_1, \dots, \lambda_{32}, \theta)$ .**

```
df <- read.csv("ConcussionsByTeamAndYear.csv",
  header = TRUE)
df$sum <- df$X2012 + df$X2013
n <- nrow(df)
M <- 10000 #Burn in
N <- 10000 #Number of draws

posterior.sample <- matrix(nrow = N, ncol = n +
  1)
# Set the initial values to the means
# of the respective distributions

theta.0 <- 1 #Mean of Gamma(.1,.1)

lambda.0 <- matrix(data = rep(1, n), n,
  1) # Set to mean of Gamma(1,theta.0)

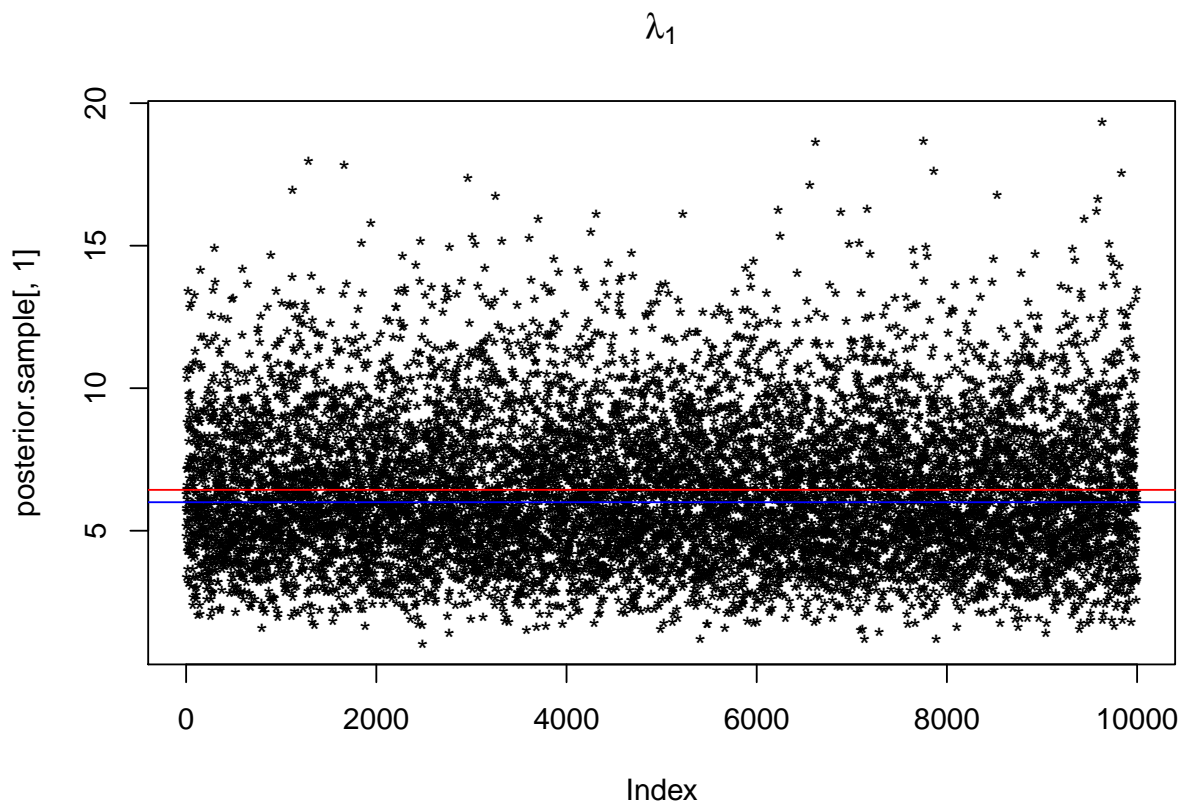
# Initialize outside loop
theta.t <- 1
lambda.t <- matrix(data = rep(1, n), n,
  1)

for (j in 1:M + N) {
  theta.t <- rgamma(1, shape = n + 0.1,
    rate = 0.1 + sum(lambda.t))
  for (i in 1:n) {
    y.i <- df$sum[i]
    lambda.t[i] <- rgamma(1, y.i +
      1, 1 + theta.t)
  }
  if (j > M) {
    sample <- c(lambda.t, theta.t)
    posterior.sample[j - M, ] <- sample
  }
}
```

Show trace plots of the samples for  $\lambda_1$  and  $\theta$ .

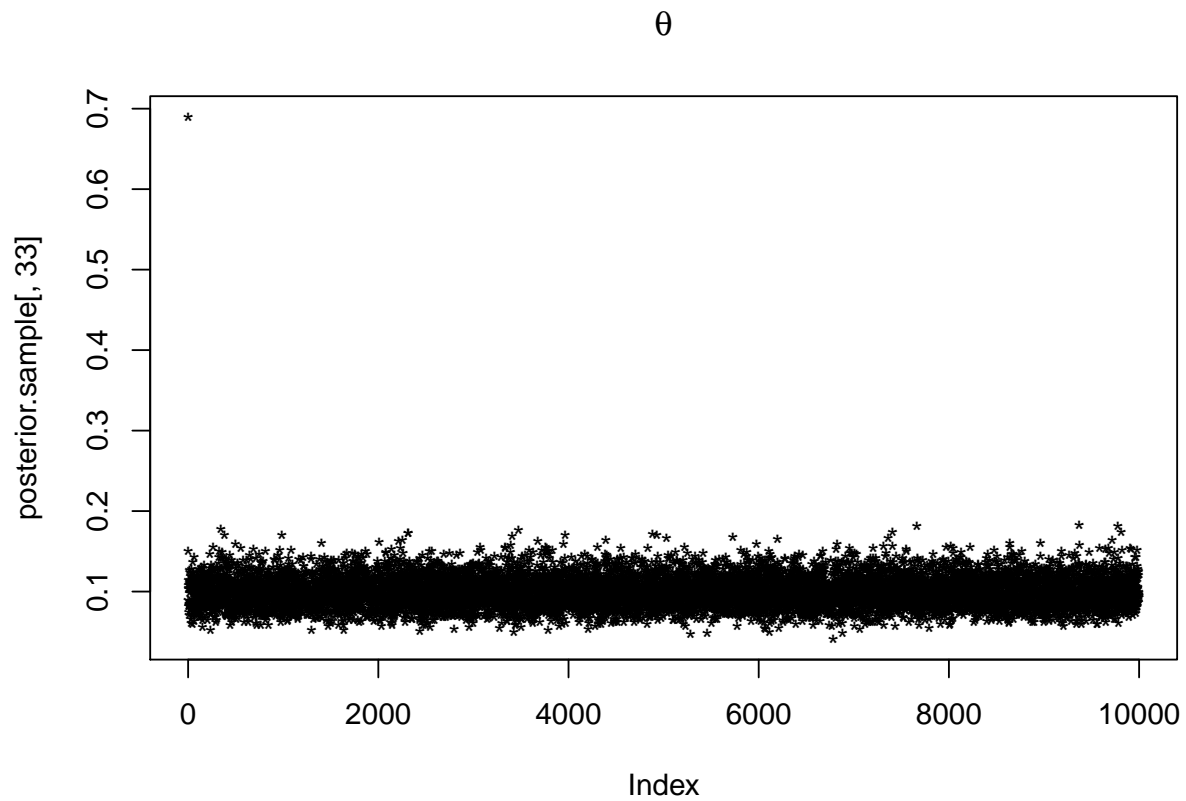
Trace plot for  $\lambda_1$

```
plot(posterior.sample[, 1], main = TeX("$\\lambda_1$"),  
     pch = "*")  
abline(h = mean(posterior.sample[, 1]),  
       col = "red")  
abline(h = mean(df$sum[1]), col = "blue")
```



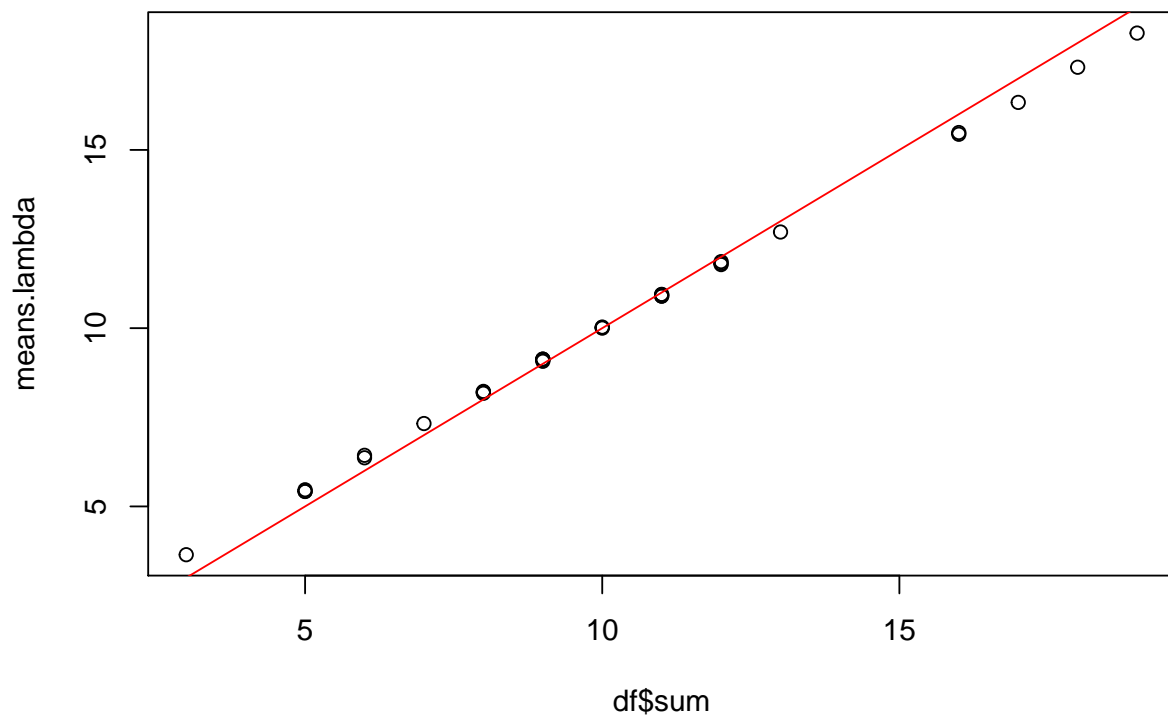
Trace plot for  $\theta$

```
plot(posterior.sample[, 33], main = TeX("$\\theta$"),  
     pch = "*")
```



- (5) Plot the estimated posterior mean of the  $\lambda_i$  versus  $Y_i$  and comment on whether the code is returning reasonable estimates. Turn in your solution on one piece of paper. Also, turn in MCMC code on a second piece of paper stapled to the solution.

```
means.lambda <- as.list(colMeans(posterior.sample))
means.lambda[[33]] <- NULL
plot(df$sum, means.lambda)
abline(a = 0, b = 1, col = "red")
```



We see good alignment between the values  $Y_i$  and  $\lambda_i$  - remember  $Y \sim \text{Poisson}(\lambda) \implies E[Y] = \lambda$

**Test - Let's validate with Jags**

```
library(rjags)
library(coda)
a      <- .1
b      <- .1
model_string <- "model{

  for(i in 1:n)
  {
    lambda[i] ~ dgamma(1,theta)
    # Likelihood
    Y[i] ~ dpois(lambda[i])
  }

  # Prior
  theta ~ dgamma(a, b)
}"
Y <- df$sum
```

```

model.concuss <- jags.model(textConnection(model_string), data = list(Y=Y,n=n,a=a,b=b))

## Compiling model graph
##   Resolving undeclared variables
##   Allocating nodes
## Graph information:
##   Observed stochastic nodes: 32
##   Unobserved stochastic nodes: 33
##   Total graph size: 69
##
## Initializing model

update(model.concuss, 10000, progress.bar="none"); # Burnin for 10000 samples

samp.theta <- coda.samples(model.concuss, variable.names=c("theta"),
                           n.iter=20000, progress.bar="none")

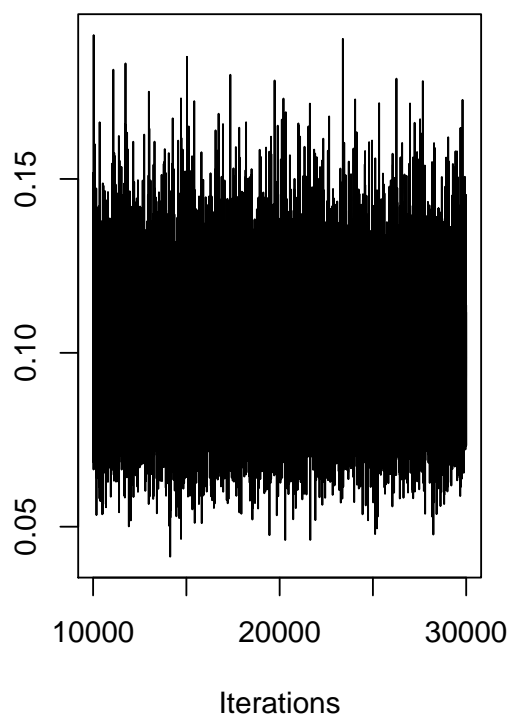
summary(samp.theta)

##
## Iterations = 10001:30000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 20000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##           Mean           SD       Naive SE Time-series SE
##    0.0995230    0.0183314    0.0001296    0.0001415
##
## 2. Quantiles for each variable:
##
##    2.5%    25%    50%    75%    97.5%
## 0.06757 0.08655 0.09819 0.11100 0.13876

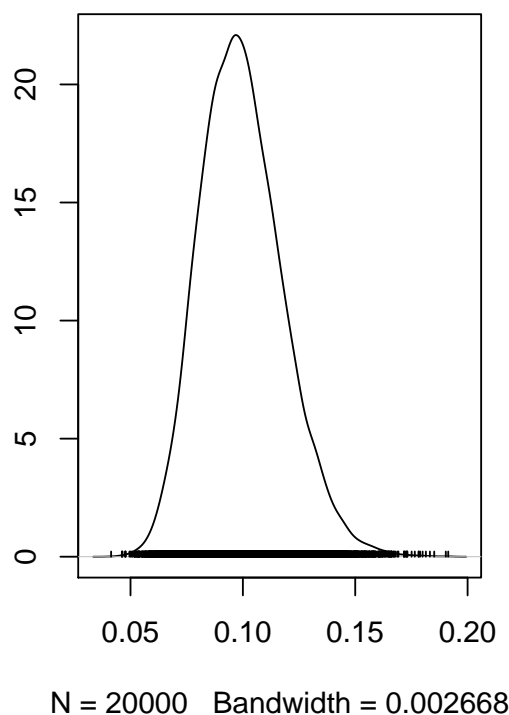
plot(samp.theta)

```

Trace of theta



Density of theta

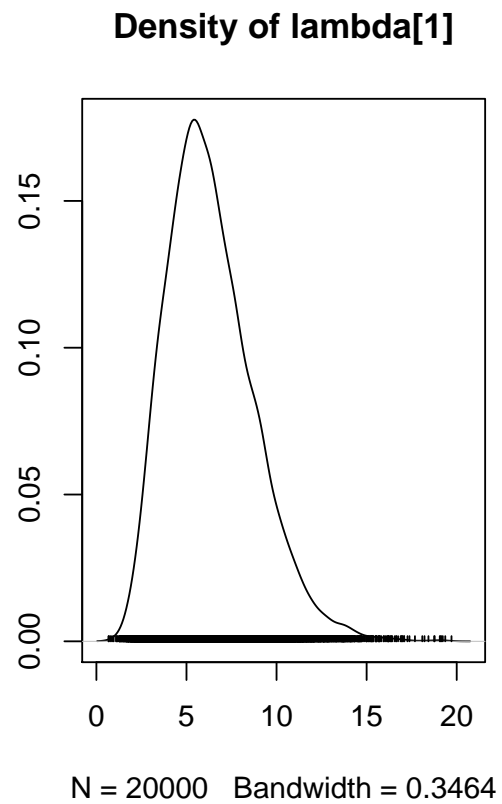
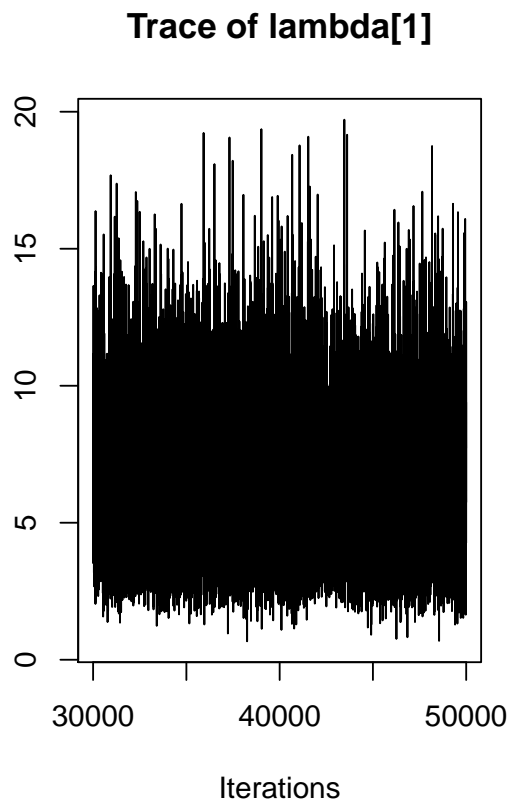


```
samp.lambda1 <- coda.samples(model.concuss, variable.names=c("lambda[1]"),
                             n.iter=20000, progress.bar="none")
```

```
summary(samp.lambda1)
```

```
##
## Iterations = 30001:50000
## Thinning interval = 1
## Number of chains = 1
## Sample size per chain = 20000
##
## 1. Empirical mean and standard deviation for each variable,
##    plus standard error of the mean:
##
##           Mean           SD      Naive SE Time-series SE
##      6.35249      2.41543      0.01708      0.01708
##
## 2. Quantiles for each variable:
##
##   2.5%   25%   50%   75%  97.5%
##   2.545  4.610  6.055  7.784 11.808
```

```
plot(samp.lambda1)
```



```
mean(unlist(samp.lambda1))
```

```
## [1] 6.352491
```