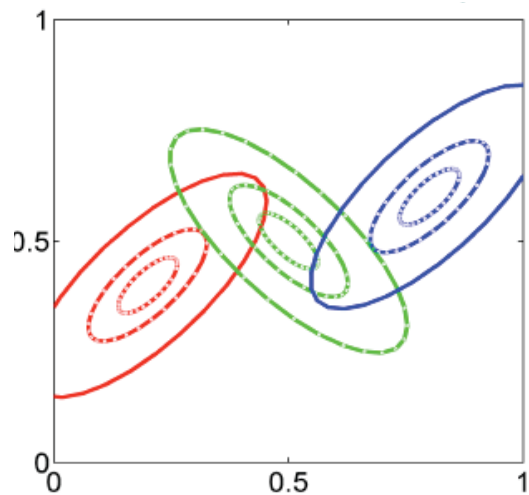


GMM: $p(\mathbf{x}|\theta) = p(\mathbf{x}) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k).$ $0 \leq \pi_k \leq 1$ $\sum_{k=1}^K \pi_k = 1$



GMM: $p(\mathbf{x}|\theta) = p(\mathbf{x}) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k), \quad 0 \leq \pi_k \leq 1 \quad \sum_{k=1}^K \pi_k = 1$

- 1 of K representation

K -dimensional binary random variable \mathbf{z}

$$z_k \in \{0, 1\} \text{ and } \sum_k z_k = 1$$

$$p(z_k = 1) = \pi_k$$

$$p(\mathbf{z}) = \prod_{k=1}^K \pi_k^{z_k}$$

- conditional probability

$$p(\mathbf{x}|z_k = 1) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \quad p(\mathbf{x}|\mathbf{z}) = \prod_{k=1}^K \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)^{z_k}$$

$$p(\mathbf{x}) = \sum_{\mathbf{z}} \underbrace{p(\mathbf{z})p(\mathbf{x}|\mathbf{z})}_{p(\mathbf{x}, \mathbf{z})} = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

GMM: $p(\mathbf{x}|\theta) = p(\mathbf{x}) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k), \quad 0 \leq \pi_k \leq 1 \quad \sum_{k=1}^K \pi_k = 1$

- 1 of k representation

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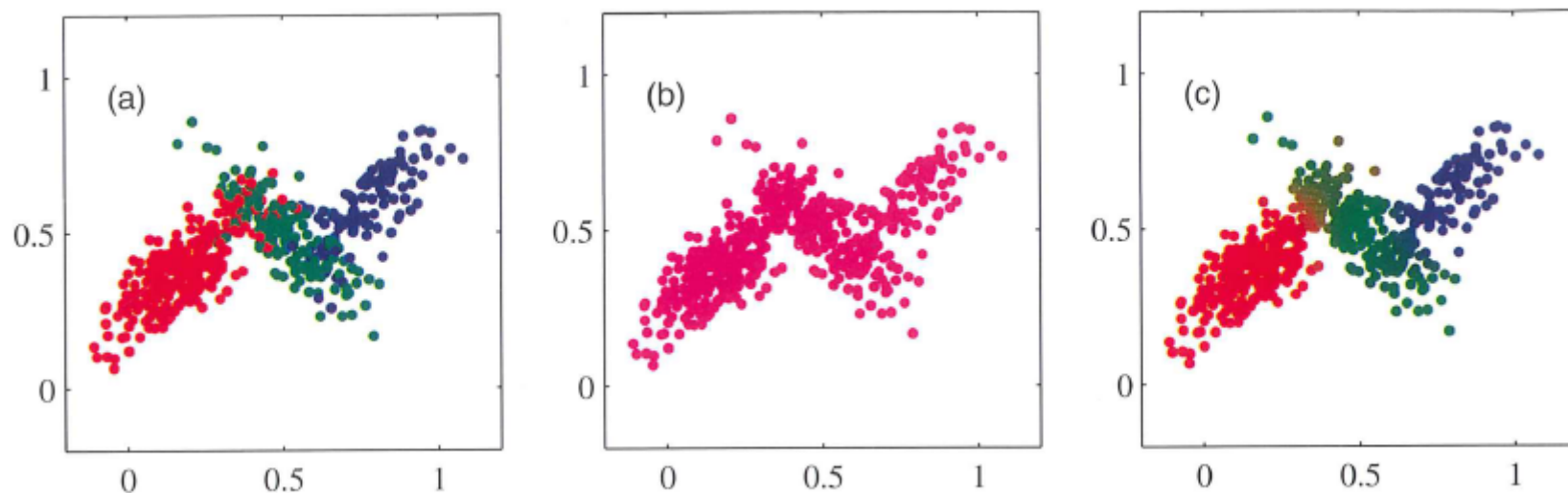
remark: If we have several observations $\mathbf{x}_1, \dots, \mathbf{x}_N$, then, because we have represented the marginal distribution in the form $p(\mathbf{x}) = \sum_{\mathbf{z}} p(\mathbf{x}, \mathbf{z})$, it follows that for every observed data point \mathbf{x}_n there is a corresponding latent variable \mathbf{z}_n .

$$p(\mathbf{x}) = \sum_{\mathbf{z}} \underbrace{p(\mathbf{z})p(\mathbf{x}|\mathbf{z})}_{p(\mathbf{x}, \mathbf{z})} = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

- Responsibilities

$$\begin{aligned}\gamma(z_k) \equiv p(z_k = 1 | \mathbf{x}) &= \frac{p(z_k = 1)p(\mathbf{x} | z_k = 1)}{\sum_{j=1}^K p(z_j = 1)p(\mathbf{x} | z_j = 1)} \\ &= \frac{\pi_k \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}.\end{aligned}$$

- Example



Maximum likelihood (GMM)

$$\ln p(\mathbf{X}|\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^N \ln \left\{ \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$

Vector of K D -dim. means $\boldsymbol{\mu}_k$

Vector of K $D \times D$ covariances $\boldsymbol{\Sigma}_k$

- maximizing w.r.t $\boldsymbol{\pi}, \boldsymbol{\mu}$ and $\boldsymbol{\Sigma} \rightarrow$

$$\boldsymbol{\mu}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n$$

$$\boldsymbol{\Sigma}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) (\mathbf{x}_n - \boldsymbol{\mu}_k)(\mathbf{x}_n - \boldsymbol{\mu}_k)^T$$

$$\left(N_k = \sum_{n=1}^N \gamma(z_{nk}) \right)$$

$$\pi_k = \frac{N_k}{N}$$

Maximum likelihood (GMM)

$$\ln p(\mathbf{X}|\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^N \ln \left\{ \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$

$$\boldsymbol{\mu}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n$$

$$\boldsymbol{\Sigma}_k = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) (\mathbf{x}_n - \boldsymbol{\mu}_k)(\mathbf{x}_n - \boldsymbol{\mu}_k)^T$$

$$\left(N_k = \sum_{n=1}^N \gamma(z_{nk}) \right)$$

$$\pi_k = \frac{N_k}{N}$$

• so what?! → **Problem**: Expr. depend on $\gamma(z_{nk})$ which depends on $\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}$ which depends on $\gamma(z_{nk})$ which depends on

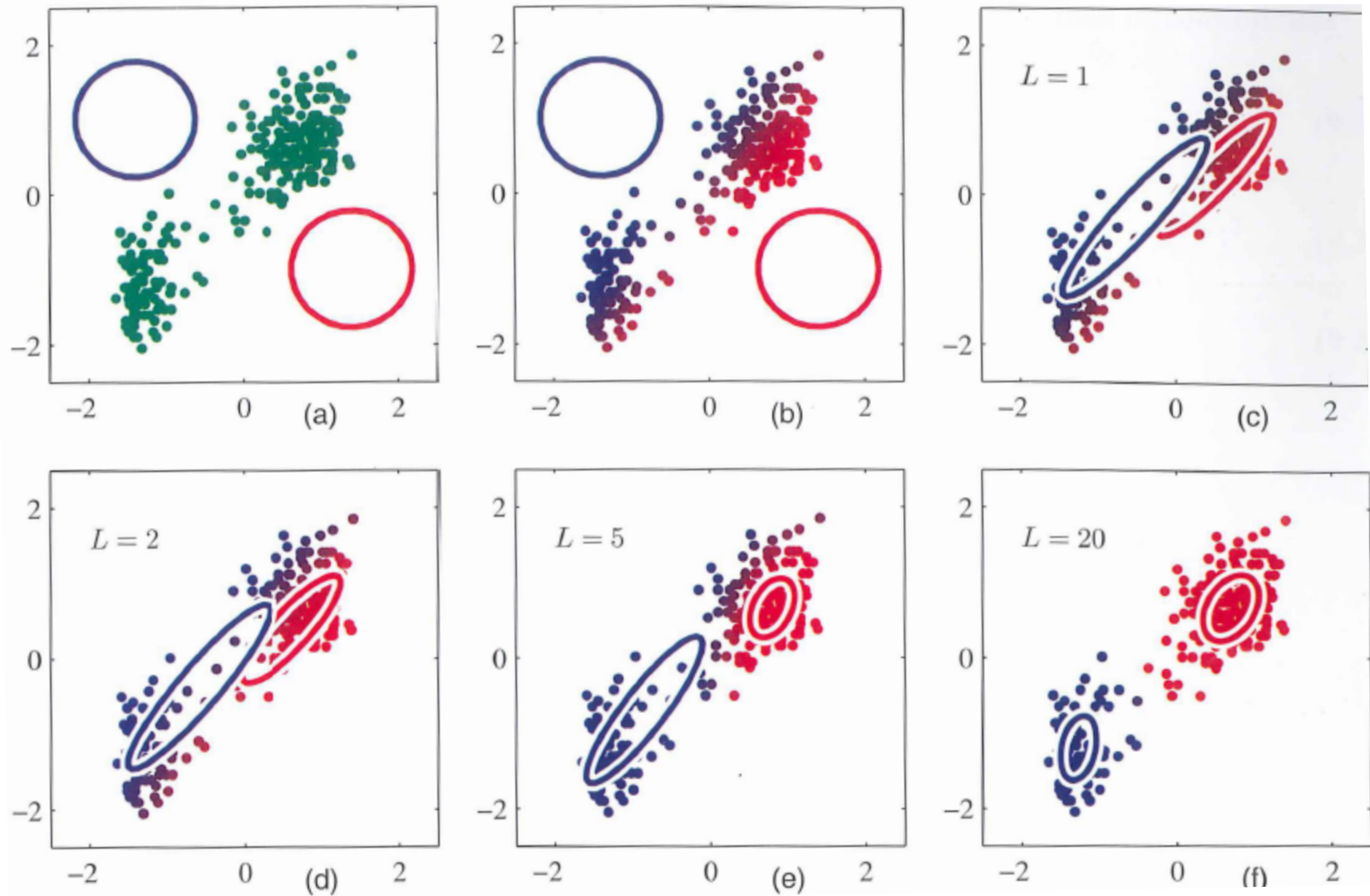
• Idea: Alternating approach (**EM-algorithm**):

Step t: Evaluate $\gamma(z_{nk})_{(t)}$ using $(\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})_{(t-1)}$

Evaluate $(\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})_{(t)}$ using $\gamma(z_{nk})_{(t)}$

GMM-Basics

Maximum likelihood (GMM)



Maximum likelihood (GMM)

EM for Gaussian Mixtures

Given a Gaussian mixture model, the goal is to maximize the likelihood function with respect to the parameters (comprising the means and covariances of the components and the mixing coefficients).

1. Initialize the means μ_k , covariances Σ_k and mixing coefficients π_k , and evaluate the initial value of the log likelihood.
2. **E step.** Evaluate the responsibilities using the current parameter values

$$\gamma(z_{nk}) = \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \mu_k, \Sigma_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_n | \mu_j, \Sigma_j)}. \quad (9.23)$$

Maximum likelihood (GMM)

3. **M step.** Re-estimate the parameters using the current responsibilities

$$\boldsymbol{\mu}_k^{\text{new}} = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n \quad (9.24)$$

$$\boldsymbol{\Sigma}_k^{\text{new}} = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) (\mathbf{x}_n - \boldsymbol{\mu}_k^{\text{new}}) (\mathbf{x}_n - \boldsymbol{\mu}_k^{\text{new}})^T \quad (9.25)$$

$$\pi_k^{\text{new}} = \frac{N_k}{N} \quad (9.26)$$

where

$$N_k = \sum_{n=1}^N \gamma(z_{nk}). \quad (9.27)$$

4. Evaluate the log likelihood

$$\ln p(\mathbf{X} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \sum_{n=1}^N \ln \left\{ \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\} \quad (9.28)$$

and check for convergence of either the parameters or the log likelihood. If the convergence criterion is not satisfied return to step 2.

EM-algorithm: General View

- Having latent variables \mathbf{Z} , ML becomes

$$\ln p(\mathbf{X}|\theta) = \ln \left\{ \sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z}|\theta) \right\}$$

- Summation inside $\ln \rightarrow$ Problems !
- If we knew the complete dataset $\{\mathbf{X}, \mathbf{Z}\}$ (and thus the distribution $p(\mathbf{X}, \mathbf{Z}|\theta)$), we could use ML to solve for θ with $p(\mathbf{X}, \mathbf{Z}|\theta)$ directly (which is easy, as we will see, because $p(\mathbf{X}, \mathbf{Z}|\theta)$ is of exponential family (the functional form is known!!))
- We only know $p(\mathbf{Z}|\mathbf{X}, \theta)$ (\rightarrow responsibilities, as we will see) \rightarrow compute expectation of (unknown) quantity $p(\mathbf{X}, \mathbf{Z}|\theta)$ or even better of the quantity $\ln p(\mathbf{X}, \mathbf{Z}|\theta)$

EM-algorithm: General View

- alternating EM:

E-Step:

compute

$$Q(\theta, \theta^{\text{old}}) = \sum_{\mathbf{Z}} p(\mathbf{Z}|\mathbf{X}, \theta^{\text{old}}) \ln p(\mathbf{X}, \mathbf{Z}|\theta).$$

M-Step:

compute

$$\theta^{\text{new}} = \arg \max_{\theta} Q(\theta, \theta^{\text{old}}).$$

The General EM Algorithm

Given a joint distribution $p(\mathbf{X}, \mathbf{Z}|\theta)$ over observed variables \mathbf{X} and latent variables \mathbf{Z} , governed by parameters θ , the goal is to maximize the likelihood function $p(\mathbf{X}|\theta)$ with respect to θ .

1. Choose an initial setting for the parameters θ^{old} .

2. **E step** Evaluate $p(\mathbf{Z}|\mathbf{X}, \theta^{\text{old}})$.

3. **M step** Evaluate θ^{new} given by

$$\theta^{\text{new}} = \arg \max_{\theta} Q(\theta, \theta^{\text{old}}) \quad (9.32)$$

where

$$Q(\theta, \theta^{\text{old}}) = \sum_{\mathbf{Z}} p(\mathbf{Z}|\mathbf{X}, \theta^{\text{old}}) \ln p(\mathbf{X}, \mathbf{Z}|\theta). \quad (9.33)$$

4. Check for convergence of either the log likelihood or the parameter values. If the convergence criterion is not satisfied, then let

$$\theta^{\text{old}} \leftarrow \theta^{\text{new}} \quad (9.34)$$

and return to step 2.

EM-algorithm: General View

- applied to GMM:

$$p(\mathbf{z}) = \prod_{k=1}^K \pi_k^{z_k} \quad p(\mathbf{x}|\mathbf{z}) = \prod_{k=1}^K \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)^{z_k} \quad \Rightarrow$$

$$p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \prod_{n=1}^N \prod_{k=1}^K \pi_k^{z_{nk}} \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)^{z_{nk}}$$

$$\ln p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \sum_{n=1}^N \sum_{k=1}^K z_{nk} \{ \ln \pi_k + \ln \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \}.$$

Bayes




$$p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) \propto \prod_{n=1}^N \prod_{k=1}^K [\pi_k \mathcal{N}(\mathbf{x}_n|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)]^{z_{nk}}$$

$$\frac{p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi})}{p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi})} = \frac{p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi})}{\sum_{\mathbf{Z}} p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi})}$$

EM-algorithm: General View

$$p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) \propto \prod_{n=1}^N \prod_{k=1}^K [\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)]^{z_{nk}}$$




$$\begin{aligned} \mathbb{E}[z_{nk}] &= \frac{\sum_{z_{nk} \in \{0,1\}} z_{nk} [\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)]^{z_{nk}}}{\sum_{z_{nj}} [\pi_j \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)]^{z_{nj}}} \\ &= \frac{\pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)} = \gamma(z_{nk}) \end{aligned}$$

EM-algorithm: General View

$$\ln p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi}) = \sum_{n=1}^N \sum_{k=1}^K z_{nk} \{ \ln \pi_k + \ln \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \}.$$

Linearität des
Erwartungswertes



$$\sum_{\mathbf{Z}} p(\mathbf{Z} | \mathbf{X}, \boldsymbol{\theta}^{\text{old}}) \ln p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\theta}) =$$

$$\mathbb{E}_{\mathbf{Z}} [\ln p(\mathbf{X}, \mathbf{Z} | \boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\pi})] = \sum_{n=1}^N \sum_{k=1}^K \gamma(z_{nk}) \{ \ln \pi_k + \ln \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \}.$$

$$= Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{\text{old}})$$

*these are computed
with $\boldsymbol{\theta}^{\text{old}}$*