EP16: Missing Values in Clinical Research: Multiple Imputation

9. Settings where MICE may have problems

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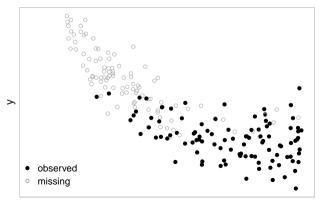


Quadratic effect

Consider the case where the **analysis model** (which we assume to be true) is

$$y = \beta_0 + \beta_1 x + \beta_2 x^2 + \dots,$$

i.e., y has a **quadratic relationship** with x, and x is incomplete.



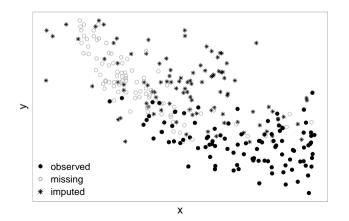
The original data show a curved pattern.

Quadratic effect

The model used to **impute** x when using MICE (naively) is

$$x = \theta_{10} + \theta_{11}y + \dots,$$

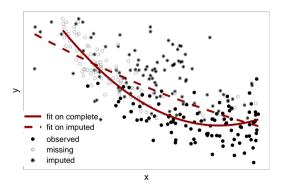
i.e., a **linear relation** between x and y is assumed.



The imputed values **distort the curved pattern** of the original data.

Quadratic effect

The model fitted on the imputed data gives **severely biased results**; the non-linear shape of the curve has almost completely disappeared.



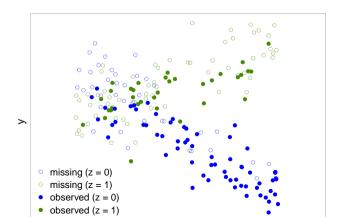
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Interaction effect

Another example occurs when the analysis model (again, assumed to be true) is

$$y = \beta_0 + \beta_x x + \beta_z z + \beta_{xz} xz + \dots,$$

i.e., y has a **non-linear relationship** with x due to the **interaction term**.



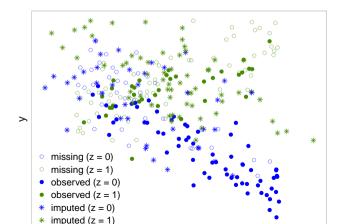
The original data shows a "<" shaped pattern.

Interaction effect

The model used to impute x when using MICE (naively) is

$$x = \theta_{10} + \theta_{11}y + \theta_{12}z + \dots,$$

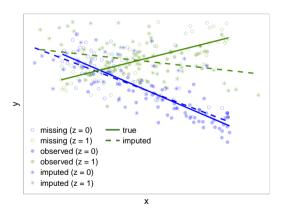
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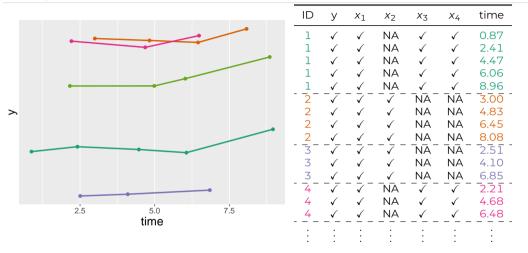


The "<" shaped pattern of the true data is **distorted by the imputed values**.

Interaction effect

And the analysis on these naively imputed values leads to **severely biased estimates**.



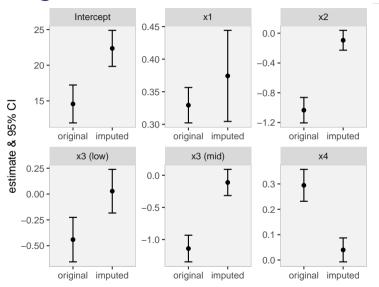


Here, $x_1, ..., x_4$ are baseline covariates, i.e., not measured repeatedly (e.g. age at baseline, gender, education level, ...

If we use MICE in the data in this (long) format, each row would be regarded as independent, which may cause bias and **inconsistent imputations**.

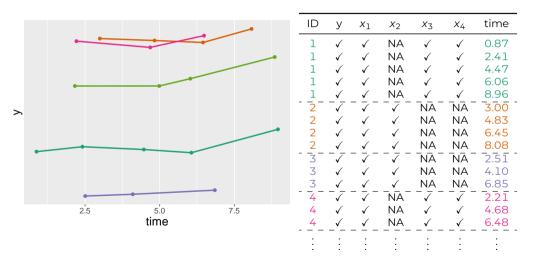
Imputed values of baseline covariates are imputed with different values, creating data that could not have been observed.

ID	У	x_1	<i>x</i> ₂	<i>X</i> ₃	<i>X</i> ₄	time	
1	√	✓	girl	✓	✓	0.87	
1	✓	\checkmark	boy	✓	\checkmark	2.41	
1	\checkmark	\checkmark	girl	✓	\checkmark	4.47	
1	\checkmark	\checkmark	girl	✓	\checkmark	6.06	
1	\checkmark	\checkmark	girl	✓	\checkmark	8.96	
2				mid	38.8	3.00	
2	\checkmark	\checkmark	\checkmark	high	39.9	4.83	
2	\checkmark	\checkmark	\checkmark	mid	40.1	6.45	
2	\checkmark	\checkmark	\checkmark	low	39.7	8.08	
3				high	40.7	2.51	
3	\checkmark	\checkmark	\checkmark	low	40.4	4.10	
3	\checkmark	\checkmark	\checkmark	mid	39.7	6.85	
4			boy			2.21	
4	\checkmark	\checkmark	boy	\checkmark	\checkmark	4.68	
4	_<_		girl	√	✓	6.48	
:	:	:	:	:	:	:	

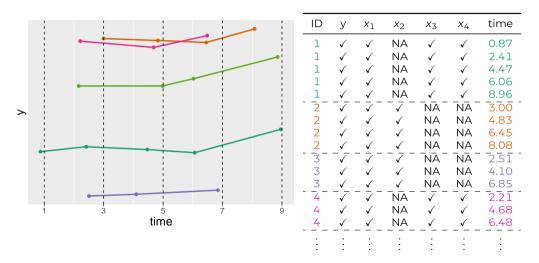


Estimates can be severely biased.

In some settings imputation in wide format may be possible.

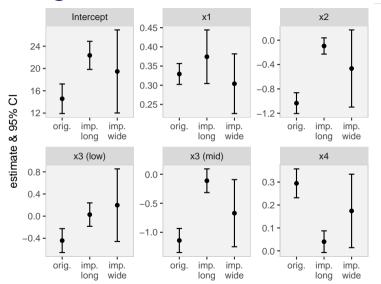


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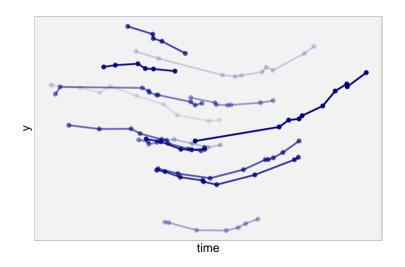


i	d	y.1	time.1	y.3	time.3	y.5	time.5	y.7	time.7	y.9
	1	31.59	0.87	31.79	2.41	31.67	4.47	31.54	6.06	32.5
2	2	NA	NA	36.23	3	36.15	4.83	36.07	6.45	36.63
3	3	NA	NA	29.76	2.51	29.84	4.1	30.01	6.85	NA
4	4	NA	NA	36.12	2.21	35.87	4.68	36.35	6.48	NA
:		:	:	:	:	:	:	:	:	:
•		•	•	•	•	•	•	•	•	

In this **wide format data** frame, missing values in the outcome and measurement times need to be imputed (to be able to use them as predictors to impute covariates), even though we would not need to impute them for the analysis (mixed model is valid when outcome measurements are M(C)AR).

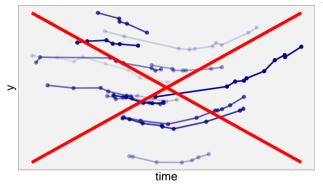


Better, but large confidence intervals.



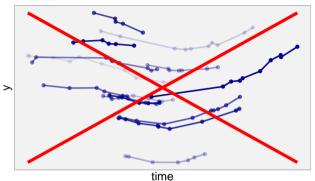
When the data is very **unbalanced**, transformation to wide format is not possible.

(Or at least transformation to wide format leads to variables with high proportions of missing values.)



Naive approaches that are sometimes used are to

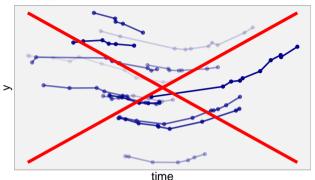
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Naive approaches that are sometimes used are to

- ► ignore the outcome in the imputation, or to
- use only the first/baseline outcome





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However, **important information may be lost**,
resulting in invalid
imputations and biased
results.

In **survival analysis**, the aim is to estimate the effect of covariates on the **time until an event** of interest happens.

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Commonly used method: Cox proportional hazards model

$$h(t) = h_0(t) \exp(x\beta_x + z\beta_z),$$

- \blacktriangleright h(t): hazard = the instantaneous risk of an event at time t, given that the event has not occurred until time t
- \blacktriangleright $h_0(t)$: unspecified baseline hazard
- ▶ x and z: **incomplete** and **complete** covariates, respectively

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Survival outcomes are usually represented by the **observed event time** T and the **event indicator** D (D = 1: event, D = 0: censored).

Naive use of MICE

- T and D are treated just like any other variable.
- ▶ The resulting imputation model for X would have the form

$$p(x \mid T, D, \mathbf{z}) = \theta_0 + \theta_1 T + \theta_2 D + \theta_3 z + \dots$$

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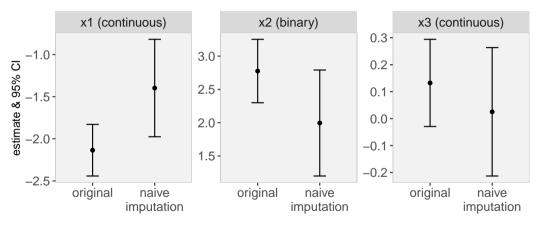
$$p(x \mid T, D, \mathbf{z}) = \theta_0 + \theta_1 T + \theta_2 D + \theta_3 z + \dots$$

The **correct conditional distribution** of *x* given the other variables is, however,

$$\log p(x \mid T, D, z) = \log p(x \mid z) + D(\beta_x x + \beta_z z) - H_0(T) \exp(\beta_x x + \beta_z z) + const.,$$

where $H_0(T)$ is the cumulative baseline hazard. (White & Royston, 2009)

Using the naively assumed imputation model can lead to severe bias:



(Results from MICE imputation with two incomplete normal and one incomplete binary covariate.)

References

White, I. R., & Royston, P. (2009). Imputing missing covariate values for the cox model. *Statistics in Medicine*, 28(15), 1982–1998.