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## 1 Motivation

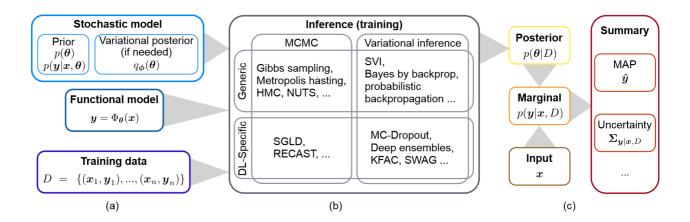
Bayesian inference provides a probabilistic approach for parameter estimation. The Bayesian neural network (BNN), or Bayesian deep learning (BDL), is a neural network with a probability distribution over weights rather than a fixed set of weights. This induces a distribution over outputs, capturing uncertainty in the predictions. It has many theoretical and practical benefits:

- a) Bayesian methods provide an approach to quantify uncertainty[1]. In BNN, the epistemic uncertainty  $p(\theta|D)$  and the aleatoric uncertainty  $p(y|x, \theta)$  can be distinguished. This makes BNNs very data-efficient since they can learn from a small dataset without overfitting[2]. In prediction, out-of-training distribution points will have high epistemic uncertainty instead of blindly giving a wrong prediction.
- **b)** BNNs can encode knowledge or subjective beliefs in priors. For example, in financial problems, Using heavy-tailed priors like Cauchy and LogUniform instead of the common Gaussian prior can improve the calibration and predictive performance[3]. Further, priors can be constructed hierarchically with first-level priors over weights and second-level hyper-priors over weight priors parameters. The hyper-priors are defined to express uncertainty about the hyperparameters themselves[4], which has been proved to have an improved learning ability when the training data is contaminated[5].
- c) BNNs provide probabilistic interpretation for deep learning tools. For example, most regularization methods used for point estimate neural networks can be understood as a prior, with L1 regularization

aligning with a Laplace prior and L2 regularization with a Gaussian prior. In addition, dropout has been shown to be equivalent to a variational posterior in uncertainty estimates[6].

- **d)** BNNs can replace traditional Bayesian approaches like Gaussian processes (GPs) in time series forecasting[7]. First, training large GPs is computationally expensive, and traditional training algorithms scale as the cube of the number of data points in the time series. In contrast, for a fixed width, training a BNN will often be approximately linear in the number of data points. Second, BNNs lend themselves better to GPU and TPU hardware acceleration than GP training operations. Third, compositional BNNs can create "hybrid" architectures, which have the ability to learn the contributions from potentially high-dimensional covariate information.
- **e)** BNNs even can be theoretically described under physical consideration. An intriguing interpretation with statistical mechanics have been derived[8]. The partition function, as a sum over all possible quantum states of a system, can be considered a posterior distribution that integrates over all possible weights[9].

The workflow to design (a), train (b) and use a BNN for predictions (c) are displayed below [1]:



## 2 Literature Review

## 2.1 Related Research

Time series forecasting is a long-standing core challenge in economics, statistics, and machine learning[3,10,11]. As most traditional neural networks (TNNs) BNNs, numerous model combinations have been employed in this field. Many of these studies rely on variational inference[12], using a surrogate model to approximate the posterior distribution. One representative work by H. J. Hortúa[3] combined convolutional neural network architectures, ResNet and AlexNet, with the natural variational inference technique VOGN, achieving great accuracy and uncertainty estimates in predicting the Volatility Index.

For MCMC implementations, F. Liang[11] found that in non-linear time series data, including sunspot numbers, BNNs outperformed traditional Box-Jenkins models, Self-Exciting Threshold Autoregressive Models (SETAR), bilinear models, and TNNs. He also suggested that model weights often become extremely large (in absolute value) to ensure sufficient learning from the training data in a parsimonious model. Furthermore, he emphasized that slight weight variation can be more crucial than a simple model structure.

## 2.2 Acknowledgement

The code implementation primarily draws on the work presented in [13], which was originally designed for the Bayesian Linear model and a Bayesian Neural Network (BNN) with a single hidden layer. Building upon this foundation, this work reconstructs both the Neural Network and the sampling to accommodate a multi-layer feedforward structure while placing greater focus on regression tasks. The Gelman–Rubin convergence diagnostic follows the implementation provided in pt-Bayeslands package [14].

## 2.3 Reference

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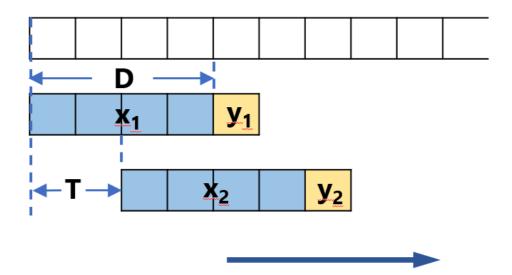
## 3 Dataset

**Background:** Sunspots are temporary phenomena on the Sun's photosphere that appear as spots darker than the surrounding areas. They are regions of reduced surface temperature caused by concentrations of magnetic field flux that inhibit convection. Sunspots usually appear in pairs of opposite magnetic polarity. Their number varies according to the approximately 11-year solar cycle.

**Sunspot dataset** (https://www.swpc.noaa.gov/products/solar-cycle-progression): contains monthly mean total sunspot numbers since 1818. After normalization, the univariate time series are transfromed into a state-space vector through Taken's theorem, which is governed by the embedding dimension (D) and time-lag (T). This is also known as **data windowing**. In this work, D = 4 and T = 2 are used for data reconstruction.

The **noise** is assumed to be Gaussian with variance  $\tau^2$ , and  $\mathbf{x}_t = (y_{t-1}, \dots, y_{t-D})$  is a vector of lagged values of  $y_t$ , so that

$$y_t = f(\mathbf{x}_t, heta) + \epsilon_t, ext{ for } t = 1, 2, \dots, n \quad \epsilon_t \sim \mathcal{N}\left(0, au^2
ight) orall t$$



```
In [1]:
        import numpy as np
         # from sklearn.pipeline import Pipeline
         # from sklearn.preprocessing import StandardScaler, MinMaxScaler
         \# D = 4
         \# T = 2
         # def read_data(file_path):
               """Read data from file."""
               with open(file_path, 'r') as file:
         #
         #
                   return [float(line.strip()) for line in file.read().strip().split('\n')]
         # def save data(file path, windows):
               """Save data to file."""
         #
               with open(file_path, 'w') as file:
         #
                   for x, y in windows:
                       formatted_values = ','.join(f' {value:.4f}' for value in x)
         #
                       file.write(f' {formatted_values}, {y:.4f} \n')
         # def windowing(data, window size=D+1):
               """Create data windows."
               return [(data[i:i + 4], data[i + 4]) for i in range(len(data) - window size)]
         # data = read data('sunspot.dat')
```

```
# pipeline = Pipeline([
# ('standardize', StandardScaler()),
# ('normalize', MinMaxScaler(feature_range=(0, 1)))
# ])

# preprocessed_data = pipeline.fit_transform(np.array(data).reshape(-1, 1)).flatten()
# preprocessed_data[preprocessed_data == 0] = 1e-3

# windows = windowing(preprocessed_data)
# windowsT = [window for index, window in enumerate(windows) if (index+T-1) % T == 0]

# split_index = int(0.66 * len(windowsT))
# train = windowsT[:split_index]
# test = windowsT[split_index:]
```

## 4 Model and Training

#### 4.1 Neural Network

In this notebook, we use BNN with 3 hidden layers (10, 5, 3) as an example.

#### 4.1.1 Xaiver Initialization

Proper initialization can help in stable propagation, faster convergence and avoiding the vanishing or exploding gradients problem. **Xavier initialization**[1] is an effective technique to maintain variance in the forward and backward passes, specifically when using certain activation functions like the hyperbolic tangent (tanh) and the logistic sigmoid. In xavier\_init function, the initial weights and biases are drawn from a gaussian distribution with a mean of 0 and a specific standard deviation, determined by the number of neurons in this layer  $d_i$  or the next layer  $d_{i+1}$ .

$$W_i \sim \mathcal{N}\left(0, rac{1}{\sqrt{d_i}}
ight) \quad b_i \sim \mathcal{N}\left(0, rac{1}{\sqrt{d_{i+1}}}
ight)$$

#### 4.1.2 Forward Pass

In this work, a simple architecture of feedforward networks is implemented. Each layer  $\mathbf{I}$  is represented as a linear transformation, followed by a nonlinear operation  $\mathbf{s}$ , known as an activation function:

$$egin{aligned} oldsymbol{l}_0 &= oldsymbol{x} \ oldsymbol{l}_i &= s_i \left( oldsymbol{W}_i oldsymbol{l}_{i-1} + oldsymbol{b}_i 
ight) \quad orall i \in [1,n] \ oldsymbol{y} &= oldsymbol{l}_n. \end{aligned}$$

This process is realized in forward function with sigmoid activation function.

#### 4.1.3 Backward Pass

Gradient-based Langevin dynamics is applied to update the model parameters using in backward pass function.

$$egin{aligned} ar{ heta}_p &= heta_p + r imes 
abla E_{y_{\mathcal{A}_{D,T}}} \left[ heta_p
ight] \ 
abla E_{y_{\mathcal{A}_{D,T}}} \left[ heta_p
ight] &= \left(rac{\delta E}{\delta heta_1}, \ldots, rac{\delta E}{\delta heta_L}
ight) \ 
abla E_{y_{\mathcal{A}_{D,T}}} \left[ heta_p
ight] &= \sum_{t \in \mathcal{A}_{D,T}} \left(y_t - f(x_t)^{[k]}
ight)^2 \end{aligned}$$

where  $\theta = (w_h, w_o, \delta_h, \delta_o)$  comprises the weights and biases of the NN model.

The  $\bar{\theta_p}$  here will serve as proposals in the subsequent MCMC sampling instead of being directly used as new network parameters.

The encode and decode transform the neural network parameters to and from a flat array format, so that these parameters are optimized in langevin\_gradient over multiple iterations.

```
import pandas as pd
In [2]:
         from tqdm import tqdm
         np. random. seed (seed=6666)
         weights = []
         biases = []
         outputs = []
         def xavier init():
            for i in range(num_layers):
                 weight_matrix = np. random. normal(
                     loc=0, scale=1/np. sqrt(layer_sizes[i]),
                     size=(layer_sizes[i], layer_sizes[i+1]))
                 bias_vector = np. random. normal(
                     loc=0, scale=1/np. sqrt(layer_sizes[i+1]),
                     size=(layer_sizes[i+1],))
                 weights.append(weight_matrix)
                 biases. append (bias vector)
                 outputs.append(np.zeros((1, layer_sizes[i+1])))
         def sigmoid(x):
            return 1 / (1 + np. exp(-x))
         def forward(X):
             current_input = X
             for i in range (num layers):
                 z = np. dot(current_input, weights[i]) + biases[i]
                 current_input = sigmoid(z)
                 outputs[i] = current_input
             return current input
         def backward(X, Y):
             Y transformed = Y
             deltas = [None] * num layers
             deltas[-1] = (Y transformed - outputs[-1]) * (outputs[-1] * (1 - outputs[-1]))
             for i in range (num layers -2, -1, -1):
                 deltas[i] = np. dot(deltas[i + 1], weights[i + 1].T) * (outputs[i] * (1 - outputs[i]))
             inputs = [X] + outputs[:-1]
             for i in range (num layers):
                 weight delta = np. outer(inputs[i], deltas[i])
                 weights[i] += 1rate * weight_delta
                 biases[i] += lrate * deltas[i]
         def encode():
             flat_weights_biases = []
             for w, b in zip(weights, biases):
                 flat weights biases. extend (w. ravel())
                 flat weights biases. extend(b)
             return np. array(flat_weights_biases)
         def decode(theta):
             current index = 0
```

```
for i in range(num_layers):
        input_size, output_size = layer_sizes[i], layer_sizes[i+1]
        weight size = input size * output size
        weights[i] = np. reshape(
            theta[current_index:current_index + weight_size],
            (input_size, output_size)
        current_index += weight_size
        biases[i] = theta[current index:current index + output size]
        current index += output size
def langevin_gradient(x_data, y_data, theta, depth):
    decode (theta)
    size = x_{data}. shape[0]
    for _ in range(depth):
        for ii in range(size):
            forward(x data[ii,])
            backward(x data[ii,], y data[ii])
    theta updated = encode()
    return theta_updated
def rmse(predictions, targets):
    return np. sqrt(((predictions - targets) ** 2). mean())
def R2(predictions, targets):
    ssr = np. sum((predictions - targets)**2)
    sst = np. sum((targets - np. mean(targets))**2)
    return 1 - (ssr / sst)
```

## 4.2 Prior and likelihood

#### **4.2.1 Prior**

To realize a **Conjugate prior**, We assume a multivariate Gaussian prior for  $\theta$  and an **inverse Gamma** distribution (IG) for  $\tau^2$  in prior function .

$$heta \sim \mathcal{N}\left(0,\sigma^2
ight) \quad au^2 \sim IG\left(
u_1,
u_2
ight)$$

That is:

$$p(oldsymbol{ heta}) \propto rac{1}{\left(2\pi\sigma^2
ight)^{L/2}} imes \expigg\{-rac{1}{2\sigma^2}igg(\sum_{h=1}^H\sum_{d=1}^D w_{dh}^2 + \sum_{k=1}^K\sum_{h=1}^Hig(\delta_{hk}^2 + v_{hk}^2ig) + \delta_o^2igg)igg\} imes au^{2(1+v_1)} \expigg(rac{-
u_2}{ au^2}igg)$$

#### 4.2.2 likelihood

The Bayesian approach for the time series problem requires sampling (estimating) the posterior distribution  $p(\theta \mid \mathbf{y})$  that requires the definition of both a likelihood function  $p(\theta \mid \mathbf{x})$  and prior distribution  $p(\theta)$ . In <code>likelihood\_function</code>, We define the likelihood function (Logarithm form in code practice), i.e probability of the data given the model, which is given by the product of the likelihood for every data point in the dataset of S instances:

$$p\left(\mathbf{y}\mid\mathbf{x}, heta, au^{2}
ight)=\prod_{t=1}^{S}p\left(\mathbf{y}_{t}\mid\mathbf{x}_{t}, heta, au^{2}
ight)$$

That is:

$$p\left(\mathbf{y}\midoldsymbol{ heta}
ight) = -rac{1}{(2\pi au^2)^{n/2}} imes \exp\!\left(-rac{1}{2 au^2}\sum_t\left(y_t - f\left(\mathbf{x}_t, heta
ight)
ight)^2
ight)$$

and equivalently,

$$\mathbf{y} = f(\mathbf{x}, heta) + e \quad e \sim \mathcal{N}\left(0, au^2
ight)$$

which is more evident to express the Gaussian noise consideration.

Notably,  $\theta$  is a set of distributions in BNN models rather than a fixed point estimate in conventional linear models.

```
In [3]: def evaluate_proposal(x_data, theta):
             size = x data. shape[0]
             fx = np. zeros(size)
             prob = np. zeros((size, layer sizes[-1]))
             for i in range(size):
                 fx_tmp = forward(x_data[i,])
                 fx[i] = fx_tmp.item()
             return fx, prob
         def prior_likelihood(sigma_squared, nu_1, nu_2, theta, tausq):
             n_params = theta_size
             part1 = -1 * (n_params / 2) * np. log(sigma_squared)
             part2 = 1 / (2 * sigma_squared) * (sum(np. square(theta)))
             log_prior = part1 - part2 - (1 + nu_1) * np. log(tausq) - (nu_2 / tausq)
             return log prior
         def likelihood_function(theta, tausq, test=False):
             if test:
                 x_{data} = x_{test}
                 y data = y test
             else:
                 x_{data} = x_{train}
                 y_data = y_train
             model_prediction, _ = evaluate_proposal(x_data, theta)
             model simulation = model prediction + np. random. normal (0, tausq, size=model prediction. shape
             accuracy = rmse(model_prediction, y_data)
             \log_1 likelihood = np. sum (-0.5 * np. \log(2 * np. pi * tausq) - 0.5 * np. square(y_data - model_1)
             return [log likelihood, model prediction, model simulation, accuracy]
```

## 4.3 MCMC Sampling

In a Metropolis-Hastings step, the proposal  $\theta^p$  for a position s is accepted with the probability  $\alpha$ :

$$lpha = \min \left( 1, rac{p\left( heta^p \mid y 
ight) q\left( heta^{[s]} \mid heta^p 
ight)}{p\left( heta^{[s} \mid y 
ight) q\left( heta^p \mid heta^{[s]} 
ight)} 
ight)$$

Substitute  $p(\theta \mid y)$  into with  $p(\theta \mid y) \propto p(\theta) \cdot p(y \mid \theta)$ :

$$lpha = \min \left(1, rac{p\left( heta^p
ight) \cdot p\left(y \mid heta^p
ight) \cdot q\left( heta^{[s]} \mid heta^p
ight)}{p\left( heta^{[s]}
ight) \cdot p\left(y \mid heta^{[s]}
ight) \cdot q\left( heta^p \mid heta^{[s]}
ight)}
ight)$$

Simplifying further, we get:

$$\log lpha = \min \left(0, \log \!\left(rac{p\left( heta^p
ight)}{p\left( heta^{[s]}
ight)}
ight) + \log \!\left(rac{p\left(y\mid heta^p
ight)}{p\left(y\mid heta^{[s]}
ight)}
ight) + \log \!\left(rac{q\left( heta^{[s]}\mid heta^p
ight)}{q\left( heta^p\mid heta^{[s]}
ight)}
ight)
ight)$$

The last equation is exactly how it's implemented in the sampler function: mh\_prob = min(1, np.exp(diff\_likelihood + diff\_priorlikelihood + diff\_prop))

To balance the computational effciency and accuracy, a combination of random-walk proposal distribution with Langevin-gradients is applied. The Pseudocode for the MCMC implmentation is provided below:

Algorithm: BNN via Langevin MCMC sampling

**Input:** Sunspot Dataset

Output: Posterior distribution of model parameters (weights and biases)

1.1 Draw initial values  $\theta_0$  from the prior

for each s until S\_max do

1.2 **Use** Langevin-gradient proposal distribution:

$$heta^p \sim \mathscr{M}( heta^{[s]}, \Sigma_ heta)$$

- 1.3 **Evaluate** prior and log-likelihood
- 1.4 **Compute** the acceptance ratio  $\alpha$
- 1.5 **Draw** u from a Uniform-distribution [0,1]

if  $u \le \alpha$  then

**Accept** replica state:  $\theta^{[s+1]} \leftarrow \theta^p$ 

end

else

**Reject** and retain previous state:  $heta^{[s+1]} \leftarrow heta^{[s]}$ 

end

end

```
def sampler():
In [4]:
             step theta = 0.025;
             step eta = 0.2;
             sgd depth = 1
             n_burnin = burn_in
             pos theta = np. ones((n samples, theta size))
             pos_tau = np. ones((n_samples, 1))
             pos eta = np. ones((n samples, 1))
             pred_y = np. zeros((n_samples, x_train.shape[0]))
             sim_y = np. zeros((n_samples, x_train. shape[0]))
             rmse data = np. zeros(n samples)
             test pred y = np. ones((n samples, x test. shape[0]))
             test sim y = np. ones((n samples, x test. shape[0]))
             test_rmse_data = np. zeros(n_samples)
             theta = np. random. randn(theta size)
             pred_y[0,], _ = evaluate_proposal(x_train, theta)
             eta = np. log(np. var(pred_y[0, ] - y_train))
             tau_proposal = np. exp(eta)
             prior = prior likelihood(sigma squared, nu 1, nu 2, theta, tau proposal)
             [likelihood, pred y[0,], sim y[0,], rmse data[0]] = likelihood function(theta, tau proposal
```

```
n \ accept = 0
n langevin = 0
for ii in tqdm(np. arange(1, n_samples), miniters=np. int64(n_samples/20)):
    theta_gd = langevin_gradient(x_train, y_train, theta.copy(), sgd_depth)
    theta_proposal = np. random. normal(theta_gd, step_theta, theta_size)
    theta_proposal_gd = langevin_gradient(x_train, y_train, theta_proposal.copy(), sgd_dept
    wc_delta = (theta - theta_proposal_gd)
    wp delta = (theta proposal - theta gd)
    sigma_sq = step_theta
    first = -0.5 * np. sum(wc_delta * wc_delta) / sigma_sq # this is wc_delta.T * wc_del
    second = -0.5 * np. sum(wp_delta * wp_delta) / sigma_sq
    diff_prop = first - second
    n_{\text{langevin}} += 1
    eta_proposal = eta + np. random. normal(0, step_eta, 1)
    tau_proposal = np. exp(eta_proposal)
    prior_proposal = prior_likelihood(
        sigma squared, nu_1, nu_2, theta_proposal, tau_proposal
    [likelihood_proposal, pred_y[ii,], sim_y[ii,], rmse_data[ii]] = likelihood_function(
        theta_proposal, tau_proposal
    [_, test_pred_y[ii,], test_sim_y[ii,], test_rmse_data[ii]] = likelihood_function(
        theta_proposal, tau_proposal, test=True
    diff_likelihood = likelihood_proposal - likelihood
    diff_priorlikelihood = prior_proposal - prior
    mh_prob = min(1, np. exp(diff_likelihood + diff_priorlikelihood + diff_prop))
   u = np. random. uniform(0, 1)
    if u < mh prob:
        n \ accept += 1
        likelihood = likelihood proposal
        prior = prior_proposal
        theta = theta_proposal
        eta = eta proposal
        # and store
        pos_theta[ii,] = theta_proposal
        pos_tau[ii,] = tau_proposal
        pos_eta[ii,] = eta_proposal
    else:
        # store
        pos_theta[ii,] = pos_theta[ii-1,]
        pos_tau[ii,] = pos_tau[ii-1,]
        pos_eta[ii,] = pos_eta[ii-1,]
accept_ratio = (n_accept / n_samples) * 100
print('{:.3}% was accepted'.format(accept_ratio))
pos theta = pos theta[n burnin:, ]
pos_tau = pos_tau[n_burnin:, ]
pos_eta = pos_eta[n_burnin:, ]
index = 0
results_dict = {}
```

```
bias_counter = 0
           current index = 0
           for i in range(num_layers):
               input_size, output_size = layer_sizes[i], layer_sizes[i+1]
               weight_size = input_size * output_size
               bias_size = output_size
               for j in range(weight_size):
                  results_dict['w{}'. format(weight_counter)] = pos_theta[:, current_index]. squeeze()
                  current index += 1
                  weight\_counter += 1
               for k in range(bias_size):
                  results_dict['b{}'. format(bias_counter)] = pos_theta[:, current_index]. squeeze()
                  current_index += 1
                  bias counter += 1
           results_dict['tau'] = pos_tau. squeeze()
           results dict['eta'] = pos eta. squeeze()
           pred dict = {}
           pred_dict['train_pred'] = pred_y[n_burnin:,:]
           pred_dict['train_sim'] = sim_y[n_burnin:,:]
           pred_dict['test_pred'] = test_pred_y[n_burnin:,:]
           pred_dict['test_sim'] = test_sim_y[n_burnin:,:]
           results_df = pd. DataFrame. from_dict(results_dict)
           return results_df, pred_dict
In [5]:
       train data = np. loadtxt("train. txt")
       test data = np. loadtxt("test.txt")
       x_train = train_data[:,:-1]
       y_train = train_data[:,-1]
       x_{test} = test_{data}[:,:-1]
       y_{test} = test_{data}[:,-1]
       n samples
                     = 3500
       burn in
                     = int(n samples* 0.4)
                     = 0.01
        1rate
       sigma_squared = 25
                     = 0
       nu 1
       nu 2
                      = 0
       theta_size
                     = n_params
       xavier init()
       results, pred = sampler()
       3499/3499 [01:59<00:00, 29.
       29it/s]
       22.1% was accepted
```

## 5 Validation

 $weight\_counter = 0$ 

## **5.1 Regression Performance**

## 5.1.1 RMSE and $R^{2}$

```
sim_y = pred_y + Gaussian noise (\tau^2)
```

```
In [6]: pred_y = pred['train_pred']
    sim_y = pred['test_sim']
    pred_y_test = pred['test_sim']

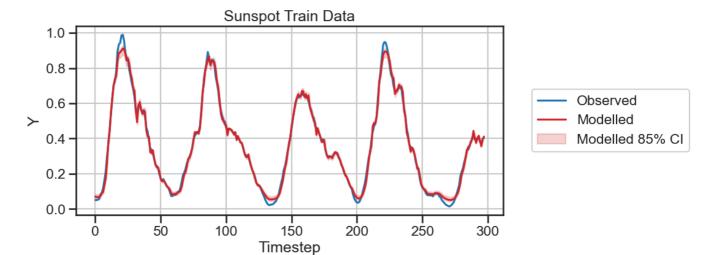
train_RMSE = rmse(sim_y, y_train)
    test_RMSE = rmse(sim_y_test, y_test)
    print('Train RMSE: {:.5f}'. format(train_RMSE))
    print('Test RMSE: {:.5f}'. format(test_RMSE))

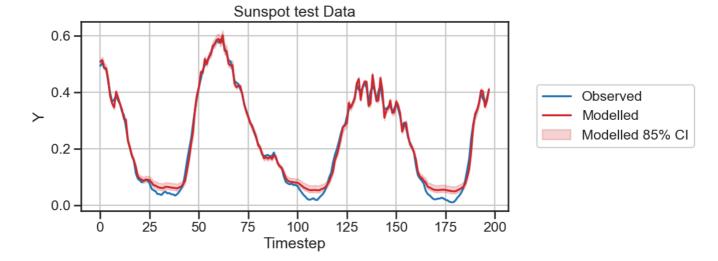
train_R2 = R2(sim_y[:,1:]. mean(axis=0)-y_train[:-1], np. diff(y_train))
    test_R2 = R2(sim_y_test[:,1:]. mean(axis=0)-y_test[:-1], np. diff(y_test))
    print('Train R2 (dy): {:.5f}'. format(train_R2))
    print('Test R2 (dy): {:.5f}'. format(test_R2))
```

Train RMSE: 0.02200 Test RMSE: 0.02022 Train R2 (dy): 0.63701 Test R2 (dy): 0.29024

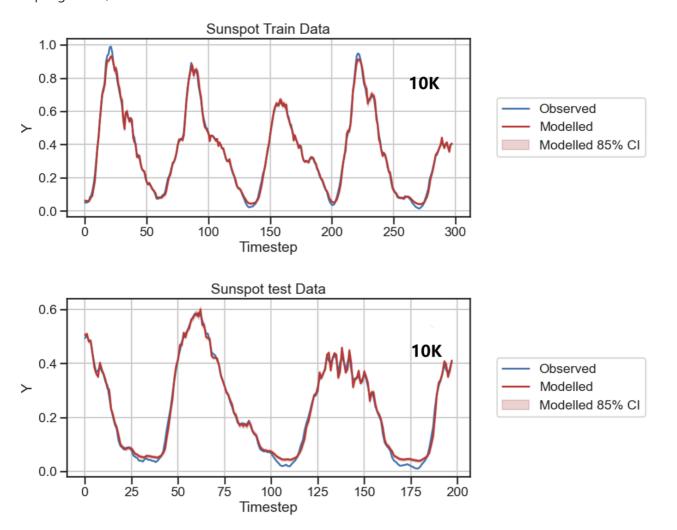
#### 5.1.2 Predictions with confidence interval

```
import matplotlib.pyplot as plt
In [7]:
         import seaborn as sns
         def plot_y_predictions(y_obs, y_mod, type='Train'):
             sns. set_context("talk")
             sns. set style("ticks", {'axes.grid': True})
             fig = plt. figure (figsize= (9, 4))
             ax1 = fig. add subplot (111)
             x = np. arange(y_obs. shape[0])
             sns. lineplot(x=x, y=y_obs. squeeze(), ax=ax1, label='Observed')
             sns. lineplot(x=x, y=np. mean(y mod, axis=0), ax=ax1, color='C3', label='Modelled')
             ax1. fill_between(x, np. percentile(y_mod, 7.5, axis=0), np. percentile(y_mod, 92.5, axis=0), cold
             ax1. set title('Sunspot {} Data'. format(type))
             ax1. set_xlabel('Timestep')
             ax1. set ylabel('Y')
             lgd = plt. legend(loc='center left', bbox_to_anchor=(1.05, 0.5))
         plot_y_predictions(y_train, sim_y)
         plot_y_predictions(y_test, sim_y_test, 'test')
```





Sampling n = 10,000



## 5.1.3 Y-Modelling

A scatter-plot of the observed (Y observed) and predicted values (Y modelled). This gives an indication of model's ability to predict change at each timestep with a skill better than persistence (as observed Yt-1 is given as an input to the model, a model predicting Yt = Yt-1 could have a low RMSE).

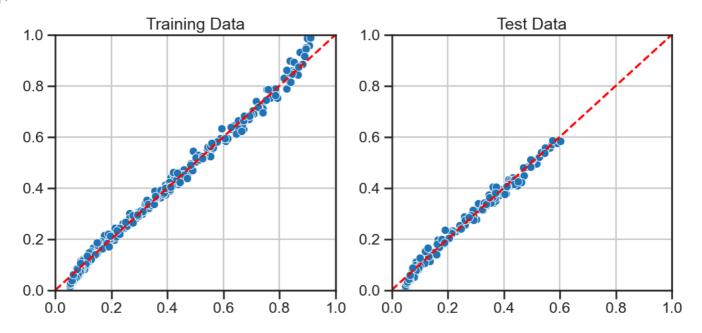
```
In [8]: def plot_ycorr_scatter(ax, y_obs, y_mod, minmax=(0,1), dy=False):
    sns.set_context("talk")
    sns.set_style("ticks", {'axes.grid': True})

ax.set_xlim(minmax[0], minmax[1])
    ax.set_ylim(minmax[0], minmax[1])
    ax.plot(ax.get_xlim(), ax.get_ylim(), '--r')

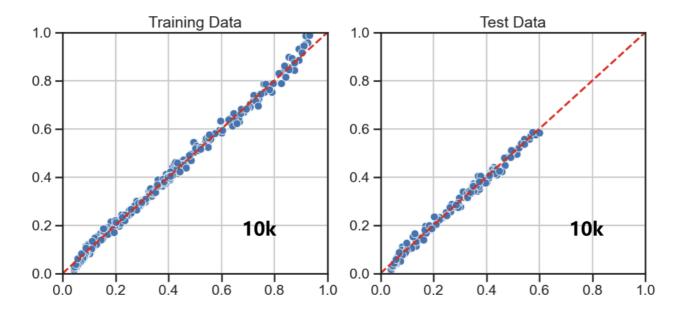
sns.scatterplot(x=np.mean(y_mod, axis=0).squeeze(), y=y_obs.squeeze(), ax=ax)
```

```
fig, (ax1, ax2) = plt.subplots(1, 2, figsize=(12, 5))
plot_ycorr_scatter(ax1, y_train, pred_y, minmax=(0, 1))
plot_ycorr_scatter(ax2, y_test, pred_y_test, minmax=(0, 1))
ax1.set_title('Training Data')
ax2.set_title('Test Data')
```

Out[8]: Text(0.5, 1.0, 'Test Data')



Sampling n = 10,000



## **5.2 MCMC Convergence Diagnose**

## 5.2.1 Acceptance rate

A roughly 23% rate implies that the posterior has been effectively sampled. 22.3% was accepted in this condition.

## 5.2.2 Gelman-Rubin diagnostic

The GelmanRubin (GR) convergence diagnostic is developed by sampling from multiple MCMC chains, whereby the variance of each chain is assessed independently (within-chain variance) and then compared to the variance between the multiple chains (between-chain variance) for each parameter. A

large difference between these two variances would indicate that the chains have not converged on the same stationary distribution.

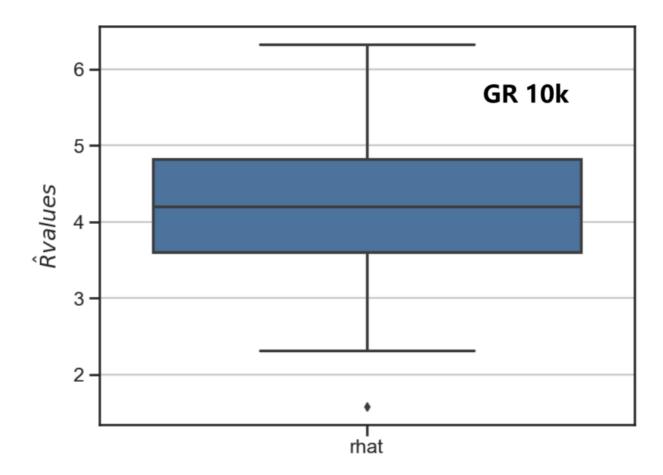
Gelman-Rubin method needs to run m  $\geq$  2 independent Markov chains in parallel and compute GR statistic  $\hat{R} \approx \sqrt{1+\frac{m}{\mathrm{ESS}}}$ , where m is the number of chains. When  $\hat{R}$  equals to 1, the sampleing is considered convergent. Here, five chains are evaluated, each with 10,000 samples of weights (excluding 40% burn-in samples) and computing the  $\hat{R}$  values for each parameter.

```
In [9]:
         import xarray as xr
         from collections import ChainMap
         n_posterior_samples = results.shape[0]
         mcmc_run = xr. Dataset(
             data_vars=ChainMap(*[{
                 _: (('chain', 'samples'), results.loc[:,_].values.reshape((1, n_posterior_samples))) for
                 _: (('samples', 'train_idx',), pred[_]) for _ in pred if 'train' in _
            }, {
            _: (('samples','test_idx',), pred[_]) for _ in pred if 'test' in _ }]),
             coords={
                 'chain': np. array([0]),
                 'samples': np. arange(n_posterior_samples),
                 'train_idx': np.arange(x_train.shape[0]),
                 'test idx': np. arange(x test. shape[0])
         # mcmc run. to netcdf('6666-10k.nc')
```

```
In [10]:
         import glob
          from convergence import gelman_rubin
          results files = glob. glob('*.nc')
          tmp_data = xr.open_mfdataset(results_files, combine='nested', concat_dim='chain')
          tmp_data. coords['chain'] = np. arange(len(results_files))
          results5 = tmp data
          model r hat = []
          r hat collect = []
          data collect = []
          train dim = results5. coords['train idx']. shape[0]
          test_dim = results5. coords['test_idx']. shape[0]
          train_pred = results5. train_pred. values. reshape((-1, train_dim))
          test pred = results5. test pred. values. reshape((-1, test dim))
          params = results5. drop_vars(['train_pred', 'test_pred', 'train_sim', 'test_sim']). copy()
          params = params.drop_vars(['train_idx', 'test_idx'])
          params = params. to_stacked_array(new_dim='param', sample_dims=['chain', 'samples'])
          gr = np. log(gelman rubin(params. values))
          param_names = np. array([_[0] for _ in params. param. values])
          good_params = param_names != 'tau'
          param_names = param_names[good_params]
          gr = gr[good params]
          results5 = results5. assign coords (param=param names)
          results5 = results5. assign(rhat=(['param'], gr))
          r hat collect. append(gr)
          data_collect.extend(gr.shape[0] * ['Sunspot'])
          r hat collect = pd. DataFrame(
              np. concatenate (r hat collect, axis=0),
              columns=['rhat']
```

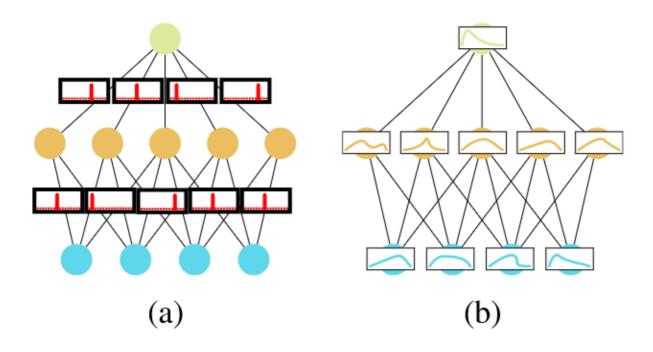
```
model_r_hat = pd. concat([r_hat_collect])
sns. set_context('talk')
fig, ax = plt. subplots(1, 1, figsize=(8, 6))
sns. boxplot(data = model_r_hat, ax=ax)
ax. set_ylabel('$\hat{R}\ values$', labelpad=10)
```

no enough space to store data of 5 chains



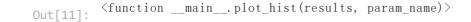
# 6 White-boxing

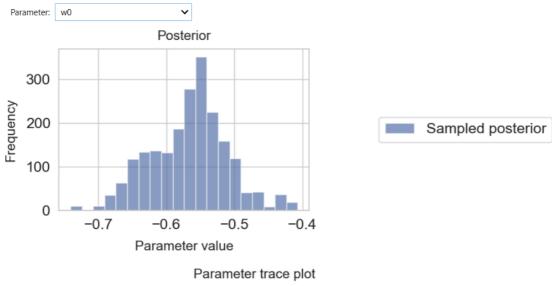
A notable advantage of BNNs is that we can observe each parameter's posterior distribution to understand what has been learned from the data.

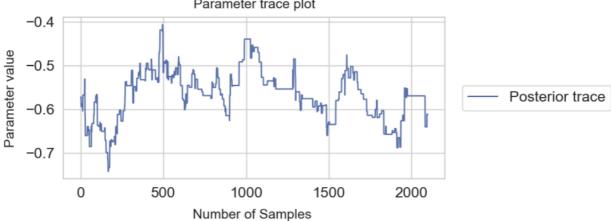


```
In [11]:
          from ipywidgets import interact, fixed, widgets
          def histogram_trace(pos_points, **kwargs):
              size = 15
              sns. set (font scale=1.5)
              sns. set_style("whitegrid")
              fig = plt. figure(figsize=(10, 4))
              ax1 = fig. add subplot (111)
              ax1. hist(pos_points, bins = 20, color='CO', alpha=0.7, label='Sampled posterior')
              x_1ims = ax1. get_xlim()
              ax1. set_xlim(x_lims)
              ax1. set_ylabel('Frequency', fontsize = size, labelpad=10)
              axl. set_xlabel(kwargs.get('param_name', 'Parameter value'), fontsize = size, labelpad=10)
              axl. set_title(kwargs. get('title', 'Posterior'), fontsize = size, pad=10)
              lgd=plt. legend (bbox to anchor= (1.25, 0.5), loc='center left')
              fig. tight layout()
              fig = plt. figure (figsize= (10, 4))
              ax1 = fig. add_subplot(111)
              axl. plot (pos_points, label='Posterior trace')
              lgd = plt. legend(loc='center left', bbox_to_anchor=(1.025, 0.5))
              plt.title("Parameter trace plot", fontsize = size, pad=10)
              plt. xlabel(' Number of Samples ', fontsize = size, labelpad=10)
              plt.ylabel(' Parameter value ', fontsize = size, labelpad=10)
              plt. tight_layout()
          def plot_hist(results, param_name):
              posterior_mean = results[param_name]. mean()
              histogram_trace(results[param_name].values)
          interact(
              plot hist,
              results=fixed(results),
              param name=widgets.Dropdown(
                  options=results.columns,
                  value='w0',
                  description='Parameter:',
```

interactive (children=(Dropdown (description='Parameter:', options=('w0', 'w1', 'w2', 'w4', 'w5', 'w6', 'w···







```
In [12]: def violinplot_weights(results, width=20, skip=6):
    fig = plt.figure(figsize=(width, width * 0.4))
    ax1 = fig.add_subplot(111)

df = pd.melt(results.drop(columns=['rmse'], errors='ignore'))

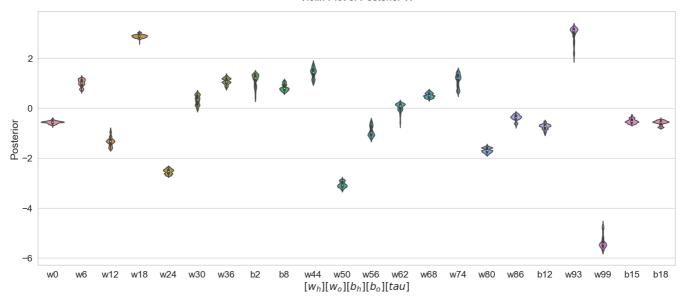
filtered_variables = df['variable'].unique()[::skip]
    filtered_df = df[df['variable'].isin(filtered_variables)]

sns.violinplot(data=filtered_df, x='variable', y='value', ax=ax1)

axl.set_ylabel('Posterior')
    axl.set_title("Violin Plot of Posterior W", pad=20)
    axl.set_xlabel('$[w_h][w_o][b_h][b_o][tau]$')
    plt.show()

violinplot_weights(results, width=20)
```





# 7 Physical/ML insight: MCMC Convergence and Prediction Accuracy

In experiments involving 6,000, 10,000, 50,000, and 150,000 samples on superconductors, Here are several observations:

**Acceptance Rate:** As the sample size increased, the acceptance rate decreased, reaching approximately 2% at 150,000 samples.

**Gelman-Rubin**  $\hat{R}$  **Value:** The Gelman-Rubin  $\hat{R}$  value gradually approached 1, indicating convergence at 150,000 samples.

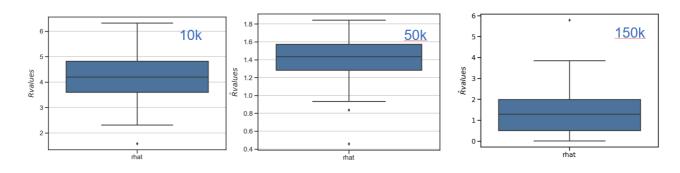
**R-squared** ( $R^2$ ) **Score:** The coefficient of determination ( $R^2$ ) significantly improved, particularly on the test set, increasing from 0.46 to 0.89.

Despite less-than-ideal convergence metrics, the model exhibited commendable accuracy and satisfactory RMSE across all sample sizes, even as low as 6,000, when compared to the Bayesian linear model in [13]. This performance can be attributed to several factors:

**Periodic Dataset:** The sunspot data is highly periodic, with an approximate 11-year cycle.

**Ineffective Convergence Indicators:** The chosen MCMC convergence metrics might not effectively reflect model performance. The 23 % acceptance rate standard is typically based on statistical and linear models, which may not apply to BNNs. Therefore, more research needs to be done to determine a good acceptance rate that aligns with convergence and ergodicity. The Gelman-Rubin diagnostics does not necessarily imply a poor performance in prediction tasks.

**Superior Performance of BNNs:** The strong learning ability of Bayesian Neural Networks contributes to their excellent performance.



In [ ]:			