COMP3206 / COMP6229 Machine Learning

Week 5: Estimation

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Estimation

- We have data \mathbf{x}_k
- We have a model: e.g. the data came from a Gaussian density
- We have parameters relating to the model: e.g. mean of the Gaussian
- Our task is to estimate the parameters given the data
- Frequentist thought
 - The given data is a particular realization of the underlying system
 - Repeated experiments will give different estimates
 - If each experiment uses a lot of data, the variation may be small
 - We define a probabilistic model and maximize likelihood
 - Bias and Variance in estimation
- Bayesian thought
 - We are interested in the uncertainty in parameters
 - We have a prior uncertainty
 - There is some information in the data
 - We combine these to get a posterior uncertainty

Likelihood & Log likelihood

- $p(\mathbf{x} \mid \omega_i)$
- Parametric form $p(\mathbf{x} | \omega_j, \boldsymbol{\theta}_j)$ For example

$$\rho(\mathbf{x} \mid \omega_j, \boldsymbol{\theta}_j) = \mathcal{N}(\mathbf{m}_j, \mathbf{C}_j)$$

- Dataset $\mathcal{D} = \mathbf{x}_1, \mathbf{x}_2, ..., \mathbf{x}_n$ of identical and independently distributed samples (iid)
 - All samples were drawn from this distribution
 - Independent draws (previous value does not affect the next draw)
- Likelihood of an item of data (function of the parameter!)

$$p(\mathcal{D}|\boldsymbol{\theta}) = p(\mathbf{x}_k|\boldsymbol{\theta})$$

Likelihood of the set of data

$$p(\mathcal{D}|\boldsymbol{\theta}) = \prod_{k=1}^{n} p(\boldsymbol{x}_{k}|\boldsymbol{\theta})$$

Log likelihood

$$I(\theta) = \ln p(\mathcal{D}|\theta)$$

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Maximum Likelihood

Maximum likelihood

$$\widehat{m{ heta}} = rg \max_{m{ heta}} I(m{ heta})$$

Maximize by taking derivative

$$oldsymbol{
abla}_{oldsymbol{ heta}} \, = \, \left[egin{array}{c} rac{\partial}{\partial heta_1} \ dots \ rac{\partial}{\partial heta_n} \end{array}
ight]$$

$$abla_{m{ heta}} I = \sum_{k=1}^{n} m{
abla}_{m{ heta}} p\left(\mathcal{D}|m{ heta}\right)$$

... and equating to zero

$$\nabla_{\theta} I = \mathbf{0}.$$

... and solve for the unknown parameter values.

Example: Multivariate Gaussian $\mathcal{N}(\boldsymbol{m},C)$

Mean unknown, (Covariance known)

- ullet ... product of Gaussians; taking log removes exp and turns \prod into \sum
- ... write it out for a single data point

$$\ln p(\mathbf{x}_k|\mathbf{m}) = \frac{1}{2}\ln(2\pi)^d \det C - \frac{1}{2}(\mathbf{x}_k - \mathbf{m})^T C^{-1}(\mathbf{x}_k - \mathbf{m})$$

... the derivative

$$\nabla_{\boldsymbol{m}} \ln p(\boldsymbol{x}_k | \boldsymbol{m}) = C^{-1}(\boldsymbol{x}_k - \boldsymbol{m})$$

• Given n data $x_1, ..., x_n$

$$\sum_{k=1}^{n} C^{-1}(\mathbf{x}_{k} - \widehat{\mathbf{m}}) = \mathbf{0}$$

and the solution is...

$$\widehat{\boldsymbol{m}} = \frac{1}{n} \sum_{k=1}^{n} \boldsymbol{x}_{k}$$

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Example: Univariate Gaussian $\mathcal{N}(m, \sigma^2)$

(Mean and variance unknown)

- Two parameters: $\theta_1 = m$ and $\theta_2 = \sigma^2$
- Log likelihood of a single data:

$$\ln p(x_k|\theta) = \frac{1}{2} \ln 2\pi \theta_2 - \frac{1}{2\theta_2} (x_k - \theta_1)^2$$

Derivative of the log likelihood

$$\nabla_{\boldsymbol{\theta}} p\left(x_k | \boldsymbol{\theta}\right) = \begin{bmatrix} \frac{1}{\theta_2} \left(x_k - \theta_1\right)^2 \\ -\frac{1}{2\theta_2} + \frac{\left(x_k - \theta_1\right)^2}{2\theta_2^2} \end{bmatrix}$$

• Given n data $x_1, x_2, ..., x_n$, and considering the full log likelihood

$$\sum_{k=1}^{n} \frac{1}{\widehat{\theta}_{2}} \left(x_{k} - \widehat{\theta}_{1} \right) = 0$$

$$- \sum_{k=1}^{n} \frac{1}{\widehat{\theta}_{2}} + \sum_{k=1}^{n} \frac{1}{\widehat{\theta}_{2}^{2}} \left(x_{k} - \widehat{\theta}_{1}^{2} \right)^{2} = 0$$

Example

Univariate Gaussian, unknown mean and variance (cont'd)

... after some algebra

$$\widehat{m} = \frac{1}{n} \sum_{k=1}^{n} x_k$$

$$\widehat{\sigma}^2 = \frac{1}{n} \sum_{k=1}^{n} (x_k - \widehat{m})^2$$

• If we did this estimation from several datasets...

$$E\left[\frac{1}{n}\sum_{k=1}^{n}(x_k-\bar{x})^2\right] = \frac{n-1}{n}\sigma^2 \neq \sigma^2$$

expected value of estimate is not the same as the true value! For the multivariate Gaussian, we give the results (slides of L1):

$$\widehat{\boldsymbol{m}} = \widehat{C} =$$

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Bayesian Estimation

Illustrate the idea through univariate Gaussian, only mean unknown

- Data: \mathcal{D} : $x_1,...x_n$
- Likelihood (as seen before) $p(x|m) \sim \mathcal{N}(m,\sigma^2)$
- Prior uncertainty over parameters: $p(m) \sim \mathcal{N}(m_0, \sigma_0^2)$ m_0 and σ_0^2 are known.
- Posterior via Bayes' formula

$$p(m|\mathcal{D}) = \frac{p(\mathcal{D}|m)p(m)}{\int p(\mathcal{D}|m) p(m) dm}$$

Denominator is a constant, so we deal with

$$p(m|\mathcal{D}) = \alpha \prod_{k=1}^{n} p(x_k|m) p(m)$$

- Two ways forward from here
 - Maximum a posteriori estimation
 - Inference by integrating out parameters

Bayesian Estimation: Univariate Gaussian

• Data: $\mathcal{D}: x_1, ...x_n$

• Likelihood (as seen before) $p(x|m) \sim \mathcal{N}(m, \sigma^2)$

• Prior uncertainty over parameters: $p(m) \sim \mathcal{N}(m_0, \sigma_0^2)$ m_0 and σ_0^2 are known.

• Substituting gives the posterior as a product of Gaussians

$$p(m|\mathcal{D}) = \alpha \prod_{k=1}^{n} \frac{1}{\sigma \sqrt{2\pi}} \exp \left[-\frac{1}{2} \left(\frac{x_k - m}{\sigma} \right)^2 \right] \times \frac{1}{\sigma_0 \sqrt{2\pi}} \exp \left[-\frac{1}{2} \left(\frac{x_k - m_0}{\sigma_0} \right)^2 \right]$$

• Which can be reduced to...

$$p(m|\mathcal{D}) = \alpha_2 \exp \left\{ -\frac{1}{2} \left\{ \left(\frac{n}{\sigma^2} + \frac{1}{\sigma_0^2} \right) m^2 - 2 \left(\frac{1}{\sigma^2} \sum_{k=1}^n x_k + \frac{m_0}{\sigma_0^2} \right) m \right\} \right\}$$

But then...

$$p(m|\mathcal{D}) = \frac{1}{\sigma_n} \exp \left\{ -\frac{1}{2} \left(\frac{m - m_n}{\sigma_n} \right)^2 \right\}$$

Matching terms...

$$\frac{1}{\sigma_n^2} = \frac{n}{\sigma^2} + \frac{1}{\sigma_0^2}$$

$$\frac{m_n}{\sigma_n^2} = \frac{n}{\sigma^2} \widehat{m}_n + \frac{m_0}{\sigma_0^2}, \text{ where, } \widehat{m}_n = \frac{1}{n} \sum_{k=1}^n x_k.$$

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Bayesian Estimation: Univariate Gaussian (cont'd)

Finally...

$$m_n = \left(\frac{n\sigma_0^2}{n\sigma_0^2 + \sigma^2}\right) \widehat{m}_n + \left(\frac{\sigma^2}{n\sigma_0^2 + \sigma^2}\right) m_0$$

$$\sigma_n^2 = \frac{\sigma_0^2 \sigma^2}{n\sigma_0^2 + \sigma^2}$$

- We now have an estimate that combines *prior* information about the parameter $(p(m) = m_0)$ with data $(x_1, ..., x_k)$ to quantify uncertainty about the parameter.
 - Before seeing any data, we have a belief
 - As we see more and more data, our belief is taken over by what the data tells us.