

Assignment 7

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T1 (1) Let X be the number of concurrent users

$$\bar{X} = \frac{1}{10} \times (17.2 + 22.1 + 18.5 + 17.2 + 18.6 + 14.8 + 21.7 + 15.8 + 16.3 + 22.8) \\ = 18.5 \text{ k} = 1.85 \times 10^4$$

$$S^2 = \frac{1}{9} \times [(17.2 - 18.5)^2 + (22.1 - 18.5)^2 + (18.5 - 18.5)^2 + (17.2 - 18.5)^2 + \\ (18.6 - 18.5)^2 + (14.8 - 18.5)^2 + (21.7 - 18.5)^2 + (15.8 - 18.5)^2 + (16.3 - 18.5)^2 + \\ (22.8 - 18.5)^2] \approx 7.88 \text{ k}^2 = 7.88 \times 10^6$$

$$S = \sqrt{S^2} = 2.81 \times 10^3$$

(2) sort the sample: 14.8, 15.8, 16.3, 17.2, 17.2, 18.5, 18.6, 21.7, 22.1, 22.8

lower quartile: $0.25 \times 10 = 2.5 \quad \lfloor 2.5 \rfloor + 1 = 3$

$$Q_{0.25} = 16.3 \text{ k} = 1.63 \times 10^4$$

upper quartile: $0.75 \times 10 = 7.5 \quad \lfloor 7.5 \rfloor + 1 = 8$

$$Q_{0.75} = 21.7 \text{ k} = 2.17 \times 10^4$$

interquartile range: $Q_{0.75} - Q_{0.25} = 5.4 \times 10^3$

$$T2. (1) f_{\max}(x) = 3f(x)[F(x)]^2 = \frac{3x^2}{\theta^3}, 0 < x < \theta$$

$$E\left(\frac{4}{3}X_{(3)}\right) = \frac{4}{3} \frac{3}{\theta^3} \int_0^\theta x^3 dx = \frac{4}{\theta^3} \left[\frac{x^4}{4}\right]_0^\theta = 0$$

$\hat{\theta}_1$ is unbiased estimator of θ

$$f_{\min}(x) = 3f(x)[1-F(x)]^2 = \frac{3}{\theta} (1 - \frac{x}{\theta})^2, 0 < x < \theta$$

$$E(4X_{(1)}) = 4 \frac{3}{\theta} \int_0^\theta x (1 - \frac{x}{\theta})^2 dx = \frac{12}{\theta} \left[\frac{\theta^2}{2} - \frac{2\theta^2}{3} + \frac{\theta^2}{4} \right] = 0$$

$\hat{\theta}_2$ is unbiased estimator of θ

$$(2) \text{Var}(\hat{\theta}_1) = \text{Var}\left(\frac{4}{3}X_{(3)}\right) = \left(\frac{4}{3}\right)^2 \text{Var}(X_{(3)}) = \frac{16}{9} \left\{ E(X_{(3)}^2) - [E(X_{(3)})]^2 \right\}$$

$$E(X_{(3)}^2) = \int_0^\theta x^2 \frac{3}{\theta^3} x^2 dx = \frac{3}{\theta^3} \left[\frac{x^5}{5} \right]_0^\theta = \frac{3}{5} \theta^2$$

$$\Rightarrow \text{Var}(\hat{\theta}_1) = \frac{16}{9} \left(\frac{3}{5}\theta^2 - \left(\frac{3}{4}\theta \right)^2 \right) = \frac{1}{15}\theta^2$$

$$\text{Var}(\hat{\theta}_2) = 16 \text{Var}(X_{(1)}) = 16 \{ E(X_{(1)}^2) - E(X_{(1)})^2 \}$$

$$E(X_{(1)}^2) = \frac{3}{\theta} \int_0^\theta x^2 \left(1 - \frac{x}{\theta}\right)^2 dx = \frac{3}{\theta} \left(\frac{x^3}{3} - \frac{x^4}{2\theta} + \frac{x^5}{5\theta^2} \right) \Big|_0^\theta = \frac{1}{10}\theta^2$$

$$\Rightarrow \text{Var}(\hat{\theta}_2) = 16 \left(\frac{1}{10}\theta^2 - \left(\frac{1}{4}\theta \right)^2 \right) = \frac{3}{5}\theta^2$$

$$\Rightarrow \text{Var}(\hat{\theta}_1) < \text{Var}(\hat{\theta}_2)$$

$\hat{\theta}_1$ is more efficient.

T3. for Chebychev's inequality, $P(|X-\mu| \geq k\sigma) \leq \frac{1}{k^2}$

$$\mu = 1.3 \times 10^9, \sigma = 0.7 \times 10^9$$

$$9.4 \times 10^9 = \mu + 3\sigma, 5.2 \times 10^9 = \mu - 3\sigma \Rightarrow k=3$$

$$P(|X-\mu| \leq 3\sigma) = 1 - P(|X-\mu| \geq 3\sigma) \geq 1 - \frac{1}{3^2} = 88.9\%$$

\Rightarrow lower bound is 88.9%.

$$\begin{aligned} T4. \hat{M} &= \frac{2}{n(n+1)} \sum_{k=1}^n k X_k = \frac{2}{n(n+1)} \sum_{k=1}^n k \mu \\ &= \frac{2}{n(n+1)} \frac{(1+n)n}{2} \mu = \mu \end{aligned}$$

\Rightarrow unbiasedness

$$\text{Var}(\hat{M}) = \text{Var}\left(\frac{2}{n(n+1)} \sum_{k=1}^n k X_k\right) = \left[\frac{2}{n(n+1)}\right]^2 \text{Var}\left(\sum_{k=1}^n k X_k\right)$$

$$\begin{aligned} \text{for } X_i \text{ i.i.d.} \Rightarrow \text{Var}(\hat{M}) &= \frac{4}{n^2(n+1)^2} \sum_{k=1}^n \text{Var}(k X_k) \\ &= \frac{4}{n^2(n+1)^2} \sum_{k=1}^n \text{Var}(X_k) \cdot k^2 = \frac{4}{n^2(n+1)^2} \sum_{k=1}^n k^2 \sigma^2 \\ &= \frac{4}{n^2(n+1)^2} \frac{n(n+1)(2n+1)}{6} \sigma^2 \\ &= \frac{2}{3} \frac{\sigma^2 (2n+1)}{n(n+1)} \end{aligned}$$

when $n \rightarrow \infty, \text{Var}(\hat{M}) \rightarrow 0$
 \Rightarrow vanishing variance

\Rightarrow consistency

$$T_5. (1) M_1 = \bar{X} = 3\theta + 7(1-\theta) = 7-4\theta$$

$$\Rightarrow \hat{\theta}_1 = \frac{7-M_1}{4}$$

$$M_1 = \frac{1}{8} \times (3+3+3+3+3+7+7+7) = \frac{9}{2}$$

$$\hat{\theta}_1 = \frac{7 - \frac{9}{2}}{4} = \frac{5}{8}$$

$$(2) E(\hat{\theta}_1) = E\left(\frac{7-M_1}{4}\right) = \frac{7-E(\bar{X})}{4} = \frac{7-(7-4\theta)}{4} = \theta$$

$$\text{Var}(\hat{\theta}_1) = \frac{1}{16} \text{Var}(M_1) = \frac{1}{16} \text{Var}(\bar{X}) \Rightarrow \hat{\theta}_1 \text{ is unbiased}$$

$$\begin{aligned} &= \frac{1}{16 \times 8} \text{Var}(X) = \frac{1}{128} (E(X^2) - E^2(X)) = \frac{1}{128} \times (9\theta + 49(1-\theta) - (7-4\theta)^2) \\ &= \frac{1}{128} (16\theta - 16\theta^2) = \frac{\theta - \theta^2}{8} \end{aligned}$$

$\hat{\theta}_1$ is unbiased estimator

$$(3) L(\theta; x) = \theta^5 (1-\theta)^3$$

$$l(\theta; x) = 5 \log \theta + 3 \log (1-\theta)$$

$$\frac{dl(\theta; x)}{d\theta} = \frac{5}{\theta} - \frac{3}{1-\theta} = 0$$

$$\Rightarrow \hat{\theta} = \frac{5}{8}$$

the maximum likelihood estimate of θ is $\frac{5}{8}$

$$T_6. (1) M_1 = \bar{X} = \int_0^\infty \frac{1}{\theta} e^{-\frac{x}{\theta}} x dx = \int_0^\infty x \lambda d(e^{-\frac{x}{\theta}})$$

$$= (-x \cdot e^{-\frac{x}{\theta}}) \Big|_0^\infty - \int_0^\infty e^{-\frac{x}{\theta}} d(-x)$$

$$= (-\theta e^{-\frac{x}{\theta}}) \Big|_0^\infty = \theta$$

$\Rightarrow M_1$ is an unbiased estimator of θ

$$\hat{\theta}_1 = \bar{X} \text{ and based on the sample}$$

$$\hat{\theta}_1 = \frac{150}{10} = 15$$

$$(2) SE \quad \text{circled } \hat{\theta}_1 = \sqrt{\frac{\text{Var}(X)}{n}} = \frac{1}{\sqrt{n}} \sqrt{\text{Var}(X)}$$

$$\text{Var}(X) = E(X^2) - E^2(X)$$

$$E(X^2) = \int_0^\infty \frac{1}{\theta} e^{-\frac{x}{\theta}} x^2 dx = \int_0^\infty x^2 d(e^{-\frac{x}{\theta}}) (-x^2) = e^{-\frac{x}{\theta}} (-x^2) \Big|_0^\infty + \int_0^\infty e^{-\frac{x}{\theta}} 2x dx \\ = 2\theta \int_0^\infty \frac{1}{\theta} e^{-\frac{x}{\theta}} x dx = 2\theta \bar{X} = 2\theta^2$$

$$\Rightarrow \text{Var}(X) = 2\theta^2 - \theta^2 = \theta^2$$

$$\hat{\theta}_1 = \frac{\theta}{\sqrt{n}} = \frac{\theta}{\sqrt{10}}$$

$$(3) L(\theta; X) = \prod_{i=1}^n \frac{1}{\theta} e^{-\frac{x_i}{\theta}} = \left(\frac{1}{\theta}\right)^n e^{-\sum_{i=1}^n x_i}$$

$$\Rightarrow l(\theta; X) = -n \log \theta - \frac{\sum x_i}{\theta}$$

$$\frac{d l(\theta; X)}{d \theta} = -\frac{n}{\theta} + \frac{\sum x_i}{\theta^2} = 0 \Rightarrow \hat{\theta}_2 = \frac{\sum x_i}{n} = \bar{x}$$

based on the sample, $\hat{\theta}_2 = 15$

$$T.T. \quad (1) Z_{0.025} = 1.96, \bar{x} = 42, \sigma = 5$$

the 95% CI of μ is $\bar{x} \pm Z_{0.025} \frac{\sigma}{\sqrt{n}} \approx 42 \pm 1.96 \times \frac{5}{\sqrt{64}}$

$$(2) \mu = 40, X \sim N(40, 25) \quad = (40.775, 43.225) \text{ min}$$

$$P(40.775 < X < 43.225) = P\left(\frac{40.775 - 40}{5} < \frac{X - 40}{5} < \frac{43.225 - 40}{5}\right)$$

$$= P(0.155 < Z < 0.645) = \Phi(0.645) - \Phi(0.155)$$

$$= 0.7422 - 0.5636 = 0.1786$$

T8. Let X be the height of woman aged 18 to 25 in region A.

Y be the height of woman aged 18 to 25 in region B.

$$\bar{X} - \bar{Y} \xrightarrow{\text{approx.}} N(\mu_X - \mu_Y, \frac{6x^2}{n} + \frac{6y^2}{m})$$

$$\mu_X - \mu_Y = 0.02 \quad \frac{6x^2}{n} + \frac{6y^2}{m} = \frac{0.2^2}{40} + \frac{0.4^2}{50} = 0.0042$$

$Z_{0.05} = 1.645$ the 90% CI of $\bar{x} - \bar{y}$ is:

$$(\bar{x} - \bar{y} - Z_{0.05} \sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}, \bar{x} - \bar{y} + Z_{0.05} \sqrt{\frac{s_x^2}{n} + \frac{s_y^2}{m}}) = 0.02 \pm 1.645 \times 0.0042$$
$$= (0.013, 0.027) \text{ m}$$