

Lect. 11 Probability

History & Law' ৰাব' কিছুম' relayability কৰ্তৃত'-

Mosteller-Wallace

Fermat, Pascal, Newton - Probability theory ৰে-
establish কৰে গৈছো

Statistics is the logic of uncertainty

Sample Space: The set of all possible outcomes of
an experiment.

Flip a coin twice

HH

HT

TH

TT

Event: A subset of the sample space

Naive definition of probability: $P(A)$ event

$$P(A) = \frac{\# \text{ of favorable outcomes}}{\# \text{ of possible outcomes}}$$

assumes all outcomes are equally likely

Counting' Multiplication rule:

If one experiment has n_1 outcomes and for each outcome of 1st experiment there are n_2 possible outcomes, for the 2nd experiment, n_3 outcomes for the m^{th} experiment, then there are $n_1 n_2 \dots n_m$ overall possible outcomes



$$3 \cdot 2 = 6$$

$$2 \cdot 3 = 6$$

Example ' Probability of a full house in poker
with a 5 card hand.

Full house : Three cards of one kind
Two cards of another kind

suit - 4

kind - 13

$$\binom{n}{k} = \frac{n!}{(n-k)! k!}$$

$$P(A) = \frac{13 \cdot \binom{4}{3} \cdot 12 \cdot \binom{4}{2}}{\binom{52}{5}}$$

Lecture 2. Counting & Story proofs

Sampling table: Choose k objects out of n

Order
matters

Replace(\checkmark)

$$n^k$$

Order
doesn't matter

$$\binom{n+k-1}{k}$$

No replace(\times)

$${}^n p_k$$

$$\binom{n}{k}$$

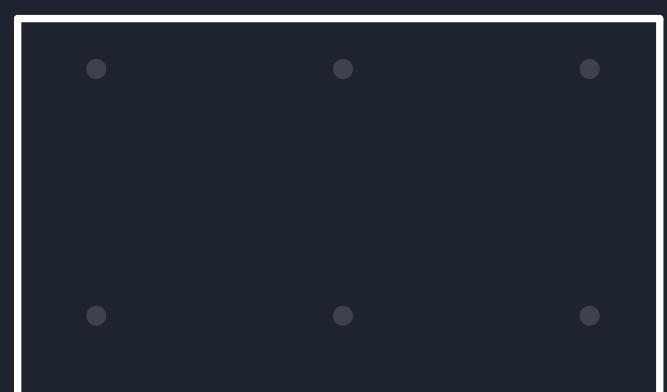
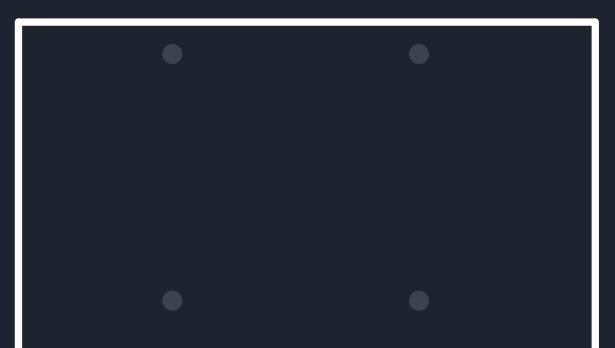
$${}^n p_k = \frac{n!}{(n-k)!}$$

pick k times from n objects where order
doesn't matter

$$k=0 \quad \binom{n-1}{0} = 1$$

$$k=1 \quad \binom{n}{1} = n$$

$$n=2 \quad \binom{k+1}{k} = k+1$$



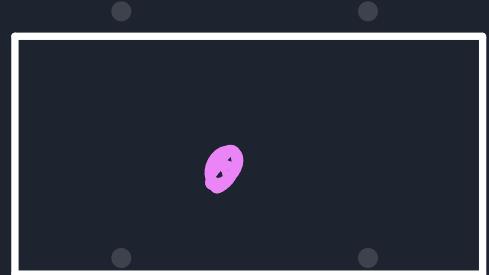
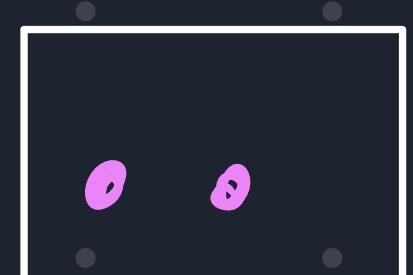
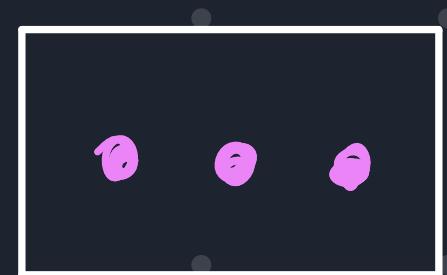
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possible choices

0, 1, ..., 7

of choice $\leq (k+1)$

Equiv: How many ways to put k indistinguishable objects in n distinguishable boxes?



1

2

3

...

n



0 0 0 | | 0 0 | . . | 0

k ଅଂଶ୍ୟ ଜ୍ଞାନ

$n-1$ " box boundary

total object $n+k-1$

$$\binom{n+k-1}{k}$$

Story proofs: ① $\binom{n}{k} = \binom{n}{n-k}$

② $n \binom{n-1}{k-1} = k \binom{n}{k}$

শুধু captain select করে না র mesh of the team টো কৰা এৰ

শুধু team টো কৰে না র captain
choose কৰা ।

Pick k people out of n , designate a captain.

③ $\binom{m+n}{k} = \sum_{j=0}^k \binom{m}{j} \binom{n}{k-j}$ [Vandermonde's identity]

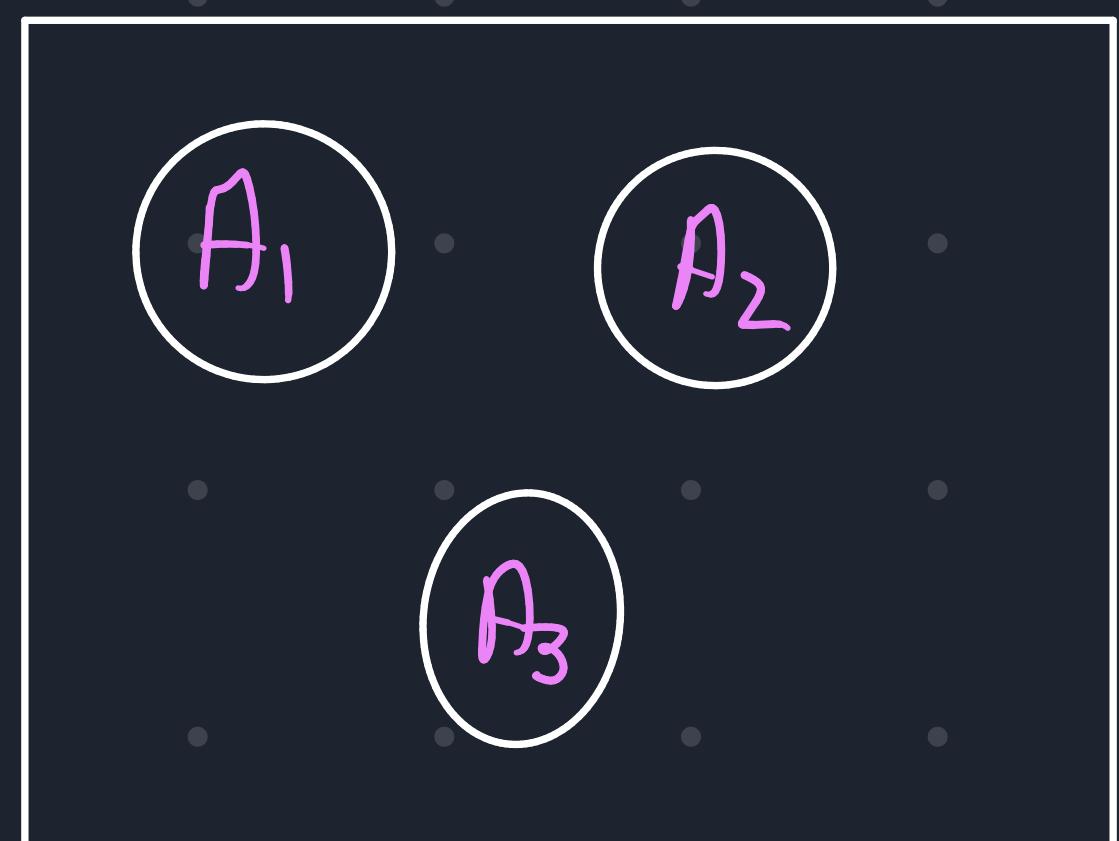
Class \sqcup m মেলে, n ফেমেলে

k শুধুক স্টুন্ডেল চোৱা কৰা

Non naive definition of probability: A probability space consists of s and p , where s is a sample space and p is a function which takes an event $A \subseteq s$ as input and gives $p(A) \in [0, 1]$ such that

$$\textcircled{1} \quad p(\emptyset) = 0, \quad p(s) = 1$$

$$\textcircled{2} \quad p\left(\bigcup_{n=1}^{\infty} A_n\right) = \sum_{n=1}^{\infty} p(A_n) \quad \text{if } A_i \text{ are disjoint}$$



Lecture 3: Birthday Problem

K people find the probability that two people have the same birthday

assume non-leap year

if $K > 365$, $P = 1$ [By pigeonhole principle]

Birthdays are equally likely, assume independent birthday.

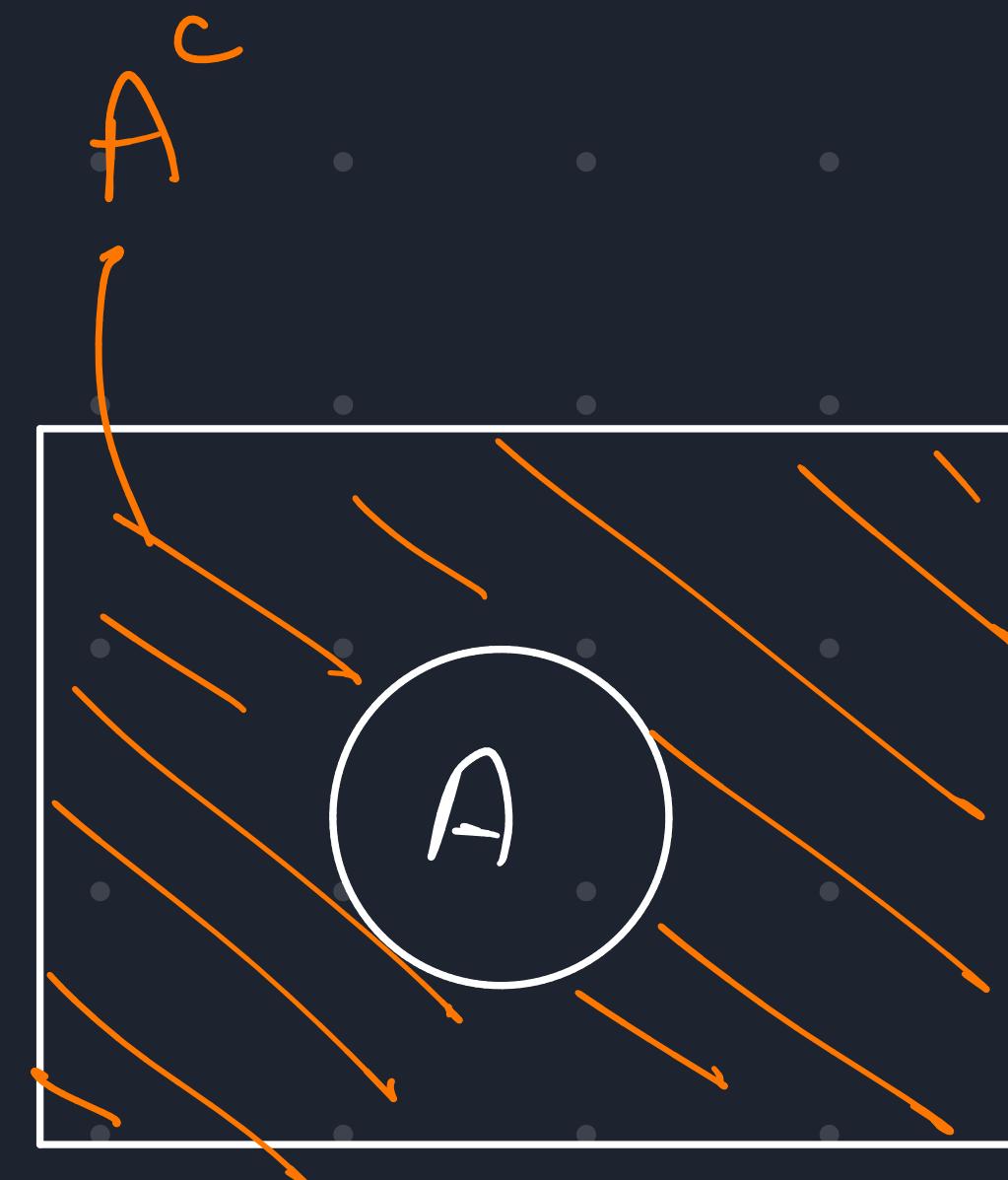
Let $K \leq 365$;

$$P(\text{no match}) = \frac{365 \cdot 364 \cdot \dots \cdot (365 - K + 1)}{365^K}$$

$$P(\text{match}) = \begin{cases} 50.7\% & ; K=23 \\ 97\% & ; K=50 \\ 99.999\% & ; K=100 \end{cases}$$

Property of Probability:

$$\textcircled{1} \quad P(A^c) = 1 - P(A)$$

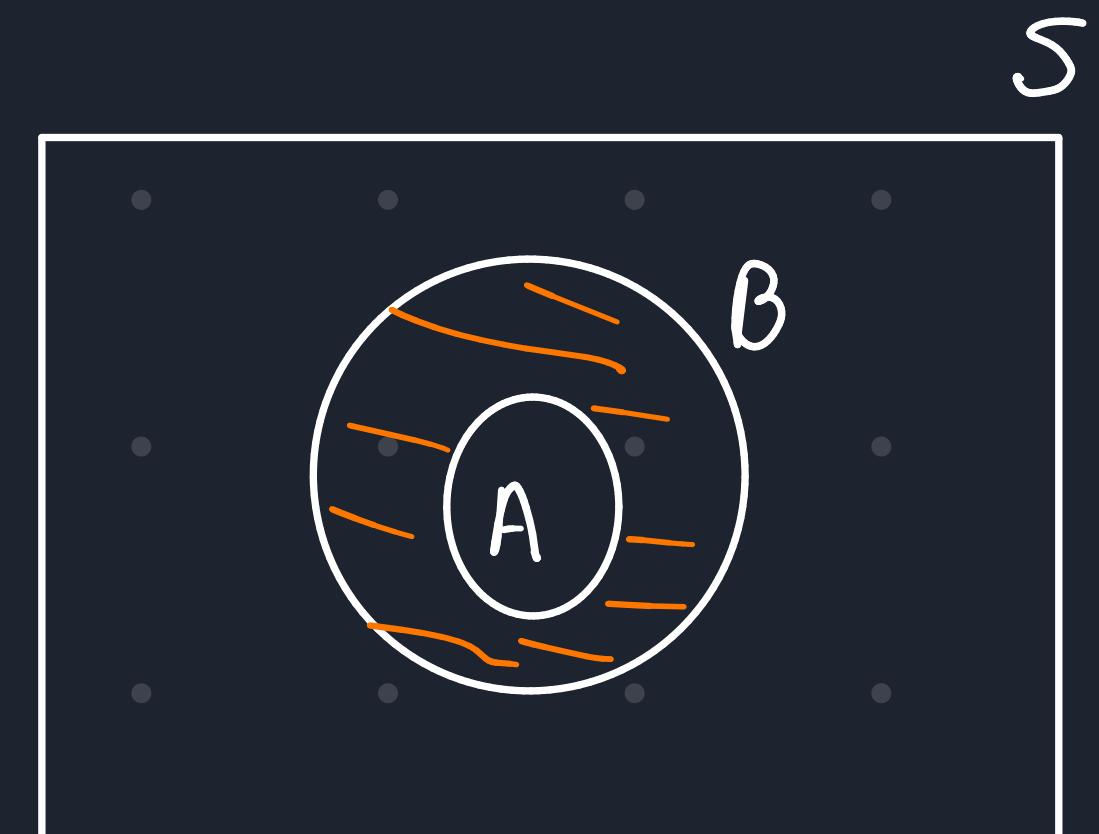


$$1 = P(S) = P(A \cup A^c) = P(A) + P(A^c)$$

$$\Rightarrow P(A^c) = 1 - P(A)$$

⑩ If $A \subseteq B$, then $P(A) \leq P(B)$

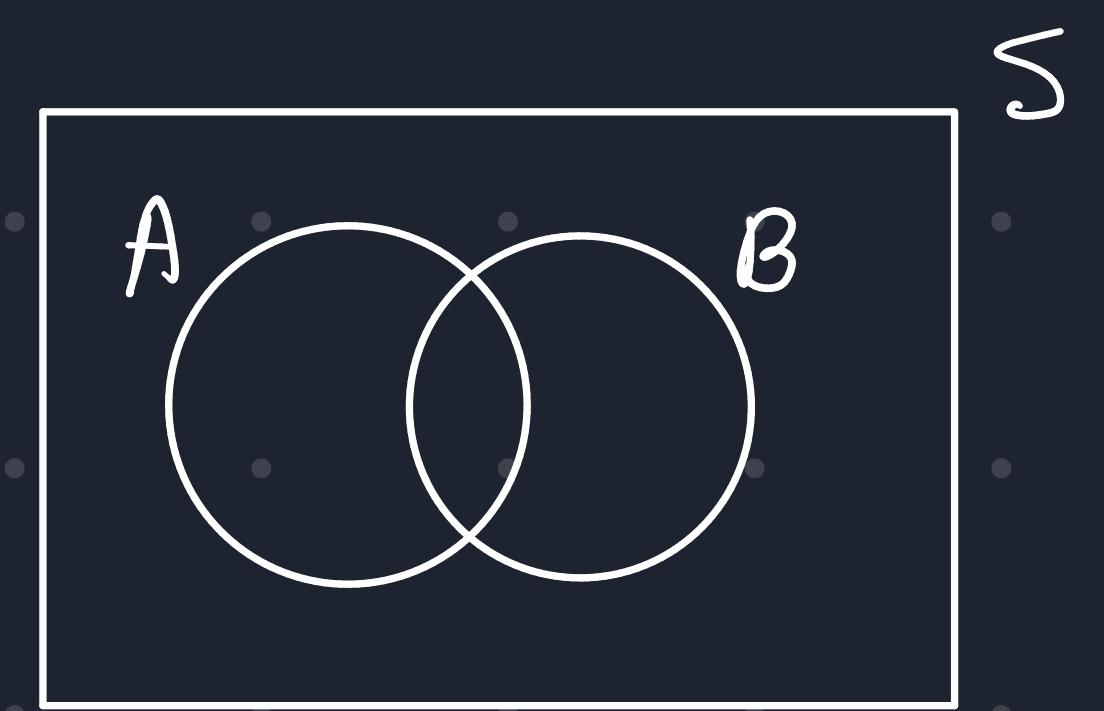
Proof: $B = A \cup (B \cap A^c)$



$$P(B) = P(A) + P(B \cap A^c) \geq P(A)$$

⑪ $P(A \cup B) = P(A) + P(B) - P(A \cap B)$

Proof: $P(A \cup B) = P(A \cup (B \cap A^c))$

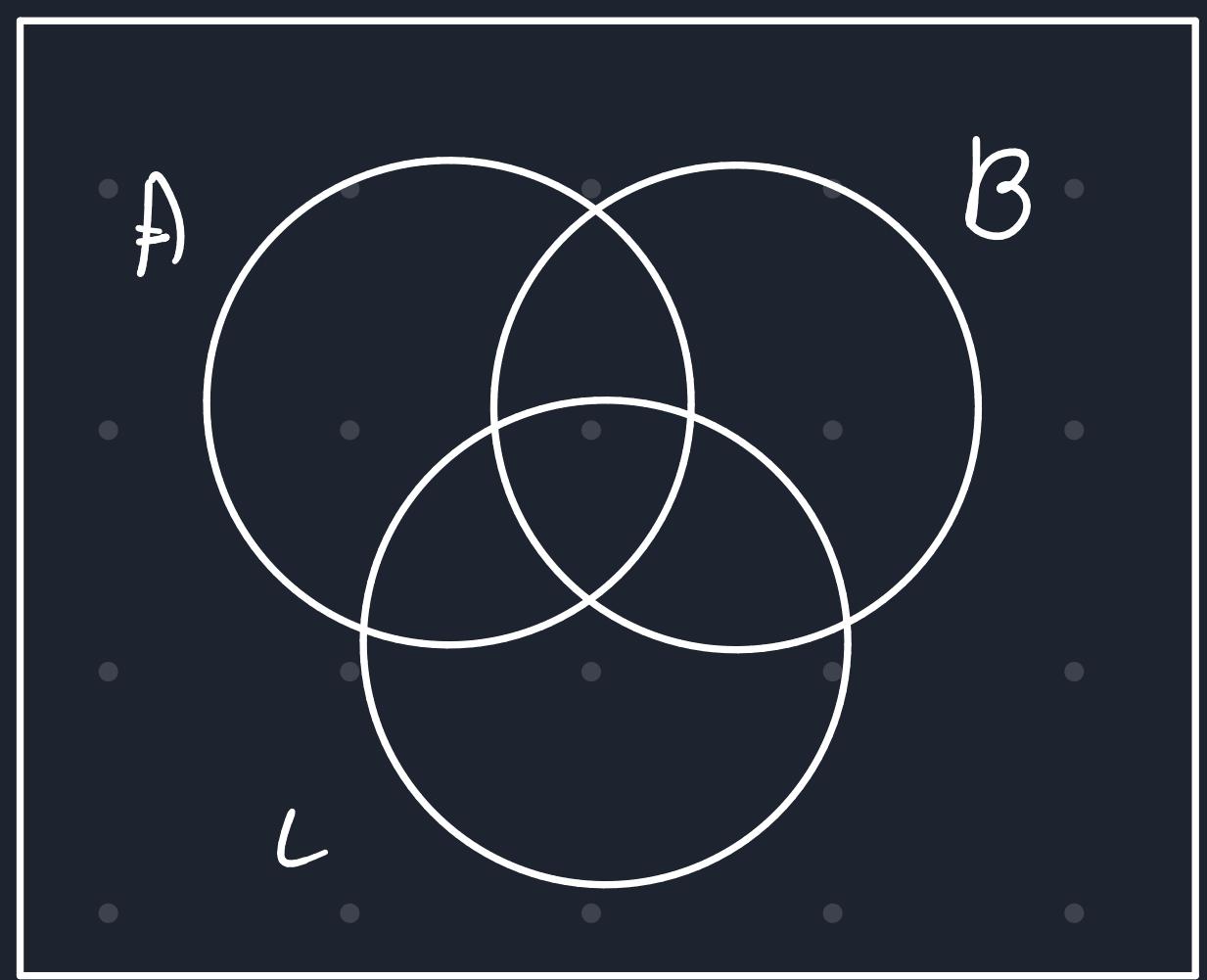


$$= P(A) + P(B \cap A^c)$$

We need to proof $P(B) - P(A \cap B) = P(B \cap A^c)$

$\Rightarrow P(B) = \underbrace{P(B \cap A) + P(B \cap A^c)}_{\text{Disjoint union}}, \text{ which is true.}$

$$\begin{aligned}
 P(A \cup B \cup C) &= P(A) + P(B) + P(C) \\
 &\quad - P(A \cap B) - P(B \cap C) - P(C \cap A) \\
 &\quad + P(A \cap B \cap C)
 \end{aligned}$$



Principle of inclusion & exclusion.

$$\begin{aligned}
 P\left(\bigcup_{i=1}^n A_i\right) &= \sum_{i=1}^n P(A_i) - \sum_{i < j} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j \cap A_k) \\
 &\quad + (-1)^{n+1} P\left(\bigcap_{i=1}^n A_i\right)
 \end{aligned}$$

de Montmort's problem (1718); matching problem

$1 \rightarrow n$ sorted card

shuffle ↓

যেকোনো card previous position \leftrightarrow

n cards : $1, 2, 3, \dots, n$ A_j be the event j th card matches.

$$P\left(\bigcup_{i=1}^n A_i\right)$$

$$P(A_j) = \frac{(n-1)!}{n!}$$

$$= \frac{1}{n}$$

← 1 fix card কে fix কৰে
 বাকিগুলো free কৰে to arrange
 n কাণ্ডাক কাণ্ড কে n! পেমাই
 arrange কৰতে পারি

$$P(A_1 \cap A_2) = \frac{(n-2)!}{n!} = \frac{1}{n(n-1)}$$

$$P(A_1 \cap \dots \cap A_k) = \frac{(n-k)!}{n!} = \frac{1}{n(n-1)(n-2)\dots(n-k+1)}$$

Lecture 4

n cards: $1, 2, 3, \dots, n$ Let A_j be the event that

j -th card matches.

$$\frac{1}{1} \quad \frac{2}{2} \quad \frac{3}{3} \quad \frac{4}{4} \quad \frac{5}{5}$$

Random
shuffle

$$\begin{array}{ccccc} 3 & \xrightarrow{1} & 5 & \xrightarrow{4} & 2 \\ \downarrow & & \downarrow & & \downarrow \\ 1 & 2 & 3 & 4 & 5 \end{array}$$

$$P\left(\bigcup_{i=1}^n A_i\right)$$

$$P(A_j) = \frac{1 \cdot (n-1)(n-2) \dots 1}{n(n-1) \dots 1}$$

$$= \frac{(n-1)!}{n!}$$

$$= \frac{1}{n}$$

$$P(A_1 \cap A_2) = \frac{1 \cdot (n-2) \dots 1}{n!} = \frac{(n-2)!}{n!}$$

$$= \frac{1}{n(n-1)}$$

$$P(A_1 \cap A_2 \cap \dots \cap A_k) = \frac{1 \cdot 2 \cdot \dots \cdot (n-k) \cdot \underbrace{(n-k-1) \cdot \dots \cdot 1}_{n!}}{n!}$$

$$= \frac{(n-k)!}{n!} \text{ ; } \binom{n}{k} \text{ such terms}$$

if all are equal

$$P\left(\bigcup_{i=1}^n A_i\right) = \sum_{i=1}^n P(A_i) - \sum_{i < j} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j \cap A_k)$$

$$+ (-1)^{n+1} P\left(\bigcap_{i=1}^n A_i\right)$$

$$\Rightarrow n \cdot \frac{1}{n} \rightarrow \underbrace{\frac{n(n-1)}{2!}}_{\frac{1}{n(n-1)}} + \underbrace{\frac{n(n-1)(n-2)}{3!}}_{\frac{1}{n(n-1)(n-2)}} \dots$$

$$= \frac{1}{1!} - \frac{1}{2!} + \frac{1}{3!} - \dots (-1)^{n+1} \frac{1}{n!}$$

$$\approx 1 - e^{-1}$$

$$P\left(\overline{A_1}^c \cap \overline{A_2}^c \cap \overline{A_3}^c\right) \approx \frac{1}{e} = 0.3678$$

$$e^{-1} = 1 - \frac{1}{1!} + \frac{1}{2!} - \frac{1}{3!}$$

$$1 - e^{-1} = \frac{1}{1!} - \frac{1}{2!} + \frac{1}{3!}$$

Fast matrix multiplication: Strassen's Algorithm
 $\Theta(n^{2.71})$

Conditional probability

Independence (defⁿ): Events A, B are independent

$$\text{if } P(A \cap B) = P(A) P(B)$$

$$\begin{aligned} A, B, C \text{ are independent if } & \left\{ \begin{aligned} P(A \cap B) &= P(A) P(B) \\ P(B \cap C) &= P(B) P(C) \\ P(C \cap A) &= P(C) P(A) \\ P(A \cap B \cap C) &= P(A) P(B) P(C) \end{aligned} \right. \end{aligned}$$

Newton Peps Problem (1693): Have fair dice.

3 events which is more likely.

(A) at least 1 six with 6 dice $P(A) = 1 - \left(\frac{5}{6}\right)^6 = 0.665$

(B) at least 2 sixes with 12 dice

(C) at least 3 sixes with 18 dice

$$P(B) = 1 - \left\{ \left(\frac{5}{6}\right)^{12} + 12 \times \frac{1}{6} \times \left(\frac{5}{6}\right)^{11} \right\} = 0.61$$

$$P(C) = 1 - \sum_{k=0}^2 \binom{18}{k} \left(\frac{1}{6}\right)^k \left(\frac{5}{6}\right)^{18-k}$$

$$= 0.597$$

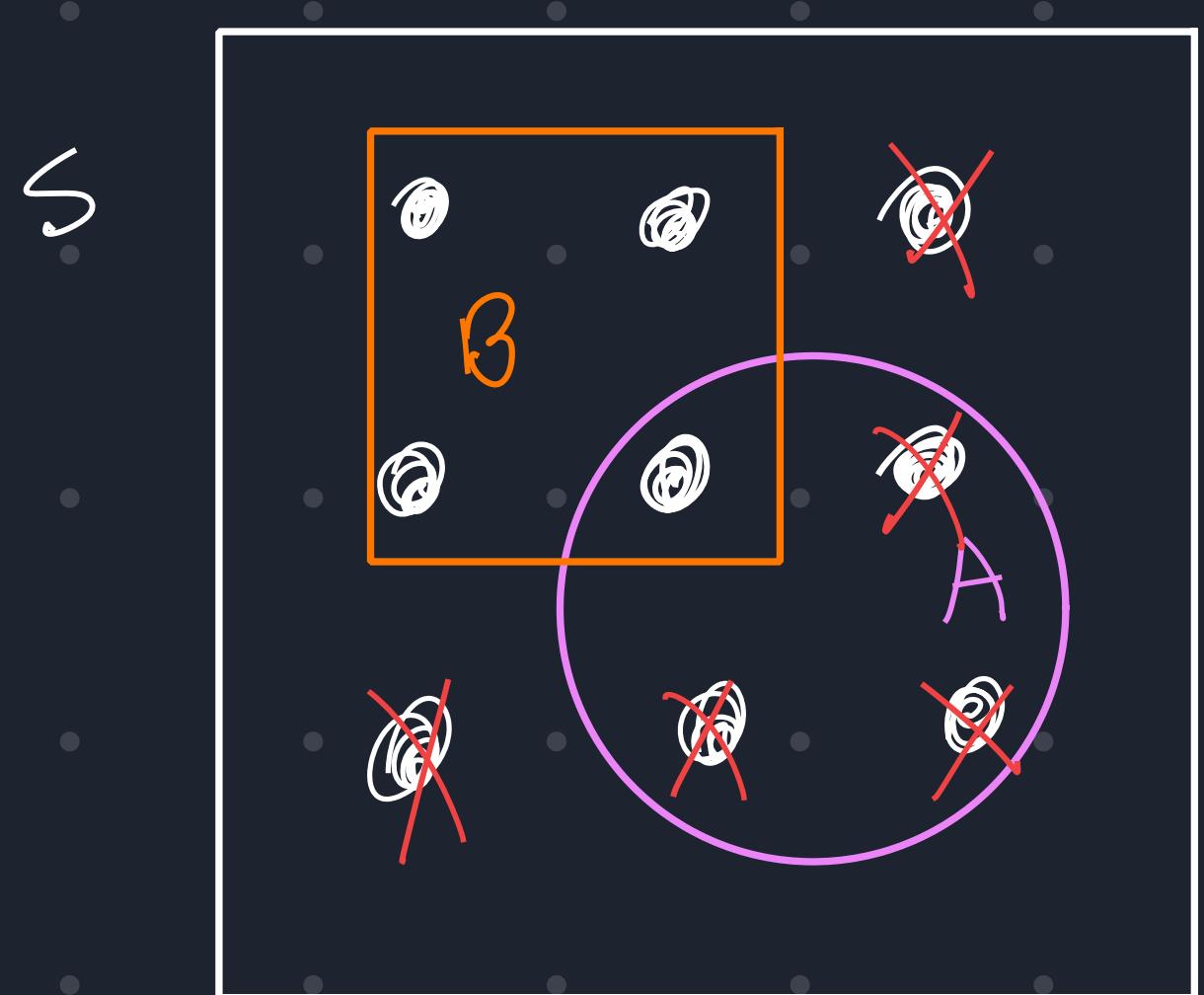
Conditional probability: How should we update probability / belief / uncertainty based on new evidence.

Defn: $P(A|B) = \frac{P(A \cap B)}{P(B)}$, $P(B) > 0$

Lec 5! Conditioning cont.

$$\text{Def'n: } P(A|B) = \frac{P(A \cap B)}{P(B)}, \quad P(B) > 0$$

Intuition:



$$P(A|B)$$

B টাই এওয়ায় শর্কিটুর
probability 0

$$P(B) = \frac{4}{9}$$

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

Normalization

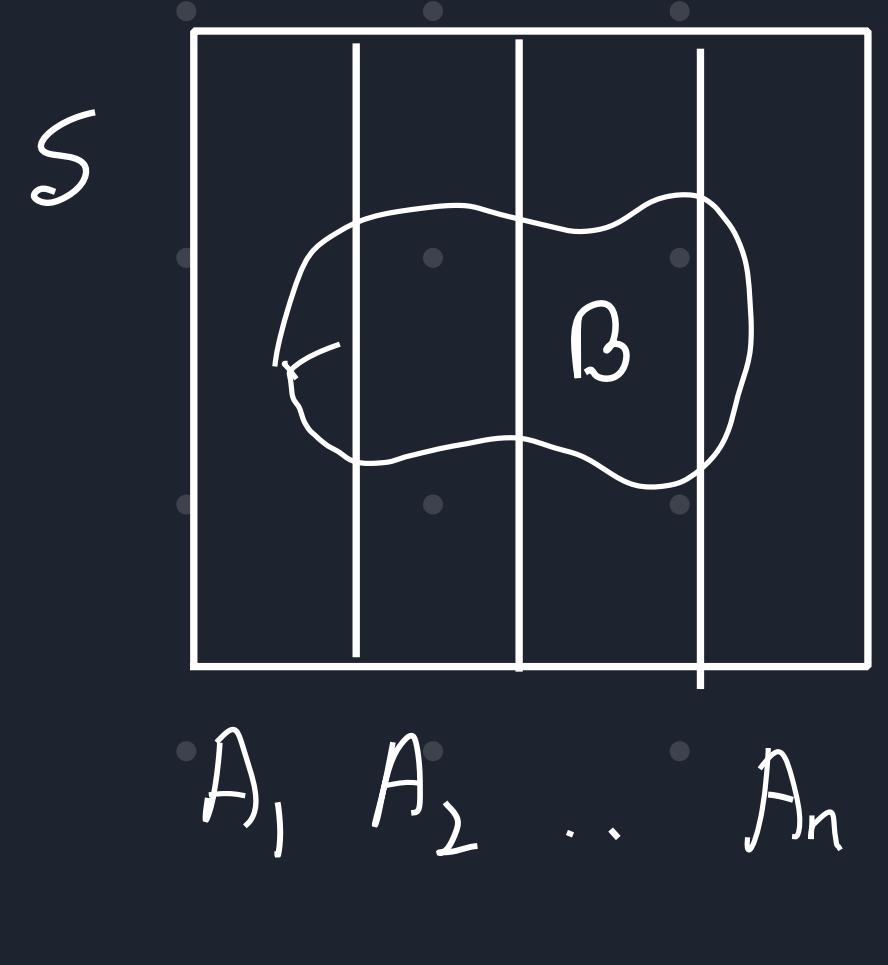
$$\text{Thm 1: } P(A \cap B) = P(B) \cdot P(A|B) = P(A) \cdot P(B|A)$$

$$\text{Thm 2: } P(A_1, A_2, \dots, A_n)$$

$$= P(A_1) P(A_2 | A_1) P(A_3 | A_2, A_1) \dots P(A_n | A_1, \dots, A_{n-1})$$

$$\text{Thm 3: } P(A|B) = \frac{P(A) P(B|A)}{P(B)}$$

Thinking conditionally, how to solve a problem



Let A_1, A_2, \dots, A_n be a partition of S

Then

$$P(B) = P(B \cap A_1) + P(B \cap A_2) + \dots + P(B \cap A_n)$$

$$= P(A_1) P(B|A_1) + P(A_2) P(B|A_2) + \dots$$

$$+ P(A_n) P(B|A_n)$$

→ Law of total probability

Suits

hearts, diamonds,
clubs, spades

Ex. Get random 2-card hand from standard deck Find $P(\text{both aces} \mid \text{have ace})$;

$$P(\text{both aces} \mid \text{have ace of spades})$$

$$\Rightarrow P(\text{both aces} \mid \text{have ace})$$

$$= \frac{P(\text{both aces, have ace})}{P(\text{have ace})}$$

$$= \frac{P(\text{both aces})}{P(\text{have ace})}$$

$$\frac{\binom{4}{2}}{\binom{52}{2}}$$

$$= \frac{\frac{\binom{4}{2}}{\binom{48}{2}}}{\binom{52}{2}}$$

$$= \frac{1}{33} = 3.03\%$$



$P(\text{both aces} \mid \text{have ace of spades})$

$$= \frac{3}{51} = \frac{1}{17} = 5.88\%$$

Ex A disease affect 1% of population

Suppose test is 95% accurate

D' patient has disease

T' test positive

$$P(D) = 0.01$$

$$P(D^c) = 0.99$$

$$P(T \mid D) = 0.95 = P(T^c \mid D^c)$$

$$P(T^c \mid D) = 0.05 = P(T \mid D^c)$$

$$P(D|T) = \frac{P(D \cap T)}{P(T)} = \frac{P(D) P(T|D)}{P(T \cap D) + P(T \cap D^c)}$$

$$= \frac{P(D) P(T|D)}{P(D) P(T|D) + P(D^c) P(T|D^c)}$$

$$= \frac{0.01 \times 0.95}{0.01 \times 0.95 + 0.99 \times 0.05}$$

$$= 0.16$$