

MTH 102: Probability and Statistics

End Semester Exam

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Sanjit K. Kaul

No books, notes, or devices are allowed. Just a pen and eraser. You have about 120 minutes.

Question 1. 35 marks A toddler is attempting to stand. An attempt fails with probability p . Answer the following questions.

- (5 marks) Calculate the average number of attempts up to and including the n^{th} time the toddler stands successfully, where $n \geq 1$.
- (10 marks) Calculate the variance of the number of attempts up to and including the n^{th} time the toddler stands successfully, where $n \geq 1$.
- (10 marks) Suppose you observe the toddler up to and including the toddler's N^{th} successful attempt at standing. Rewrite the average number of attempts in terms of N . Calculate the expected value of the resulting function for when N is a Poisson RV with PMF $P[N = n] = e^{-1}/n!$, $n \geq 0$.

[Hint: For (a) and (b), you may want to write the number of attempts as a sum of n iid random variables.]

Now suppose that the toddler learns from its attempts to stand. Specifically, the failure probability is halved after a successful attempt to stand. Till the first time the toddler stands, the failure probability is p as before. However, it becomes $p/2$ after the toddler stands successfully the first time and $p/4$ after the toddler stands successfully the second time, and so on.

- (10 marks) Let N_2 be the RV that counts the number of attempts up to and including the attempt in which the toddler stands successfully the second time. Derive the PMF of N_2 .

Question 2. 40 marks Two babies are born at the same time. Baby 1 is born into a chatty family and Baby 2 is born into a family that prefers quiet. Babies that are born in chatty families are expected to learn how to speak in a number of years uniformly distributed over $(0, 3)$. Babies born to quiet families are expected to learn to speak in a number of years uniformly distributed over $(0, 5)$. Answer the following questions.

- (0 marks) For both the babies, name the random variables that count the number of years to speaking.
 - (20 marks) Calculate the probability that the baby born to the quiet family learns to speak sooner than the baby born to the chatty family.
 - (20 marks) Suppose it is known that the baby born to the quiet family learns to speak in greater than 2 years. Calculate the probability that the baby born to the quiet family learns to speak sooner than the baby born to the chatty family.
- [Hint: What is the PDF of interest? Marginal? Joint? Conditional? Some of the above?]

Question 3. 15 marks Consider the estimator $J_n(X)$ of the parameter $E[X]$. We have

$$J_n(X) = \frac{1}{n-1} \sum_{i=1}^n (X_i + I_i W_i).$$

Here the random variables I_i , $i = 1, \dots, n$, are independent and identically distributed and take a value of 0 or 1 with equal probability. We know that $E[W_i] = 0$ and $\text{Var}[W_i] = \sigma^2$, $i = 1, \dots, n$. The random variables W_i , X_i , I_i , $i = 1, \dots, n$ are independent random variables. Answer the following questions

- (2.5 marks) Is the estimator unbiased? Arrive at your answer by calculating the expected value of the estimator.
- (10 marks) Derive the mean squared error of the estimator. [Hint: Approach using first principles. If you need to calculate the second moment, first calculate the variance and use it to calculate the second moment.]
- (2.5 marks) Is the estimator unbiased in the limit as $n \rightarrow \infty$? Calculate the mean squared error in the limit as $n \rightarrow \infty$.

Question 4. 10 marks Let X be a Binomial RV with number of trials $n = 100$ and the probability of success 0.4. Approximate the probability $P[X < 50]$ using the central limit theorem. Leave your answer as a function of the CDF $\Phi_Z(z)$ of the standard Normal.

(Q1)

Let X_1 be the no. of attempts to the first successful attempt to stand.

Let X_2 be the no. of attempts to the second successful attempt after the first succ attempt.

⋮

Let X_i be the no. of attempts to the i^{th} succ attempt after the $(i-1)^{\text{th}}$ succ attempt.

We have X_i , $i=1, 2, \dots$

$X_i \sim \text{Geom}(q)$, where $q = 1-p$.

No of attempts up to \mathbb{I} including the n^{th} succ attempt is

$$\sum_{i=1}^n X_i.$$

(a).

We want

$$E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i]$$

④ for appropriate sum.

If a student recognizes
the Pascal distn ↓
states $\frac{1}{q_i}$, that is fine
too.

$$= \sum_{i=1}^n \left(\frac{1}{q_i}\right)$$

$$= \frac{n}{q_i}$$

Correct values ①.

(b) We want

$$\text{Var}\left[\sum_{i=1}^n X_i\right]$$

$$= \sum_{i=1}^n \text{Var}[X_i]$$

⑥

②

X_i are independent

$\text{Var}[X_i]$ is the variance of a
geometric RV.

$$\text{Var}[X_i] = \left(\frac{1-q_i}{q_i^2}\right) \therefore$$

$$\begin{cases} \text{Var}\left[\sum_{i=1}^n X_i\right] \\ = \frac{n(1-q_i)}{q_i^2} \end{cases} \quad \begin{array}{l} \text{Recall} \\ q_i = 1-p_i \end{array}$$

(c)

The average no. of attempts for a random N number of successes is

$$\frac{N}{q}.$$

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$$E\left[\frac{N}{q}\right] = \frac{1}{q} E[N],$$

where $E[N]$ is the expected value of the Poisson RV with PMF $P[N=n] = \frac{e^{-\lambda}}{\lambda^n}$,

$$n \geq 0.$$

$$E[N] = \sum_{n=0}^{\infty} n \frac{e^{-\lambda}}{\lambda^n} = e^{-\lambda} + \frac{e^{-\lambda}}{\lambda^2} + \dots = e^{-\lambda} e = 1$$

4

$$\therefore E\left[\frac{N}{q}\right] = \frac{1}{q}. \quad \text{---(1)}$$

(aa) We want the PMF of N_2 .

Consider the RV N_1 . Let N_1 be the no. of attempts to the first successful standing.

$$S_{N_1} = \{1, 2, \dots\}$$

$$S_{N_2} = \{2, 3, \dots\}$$

$$P[N_2=2] = P[N_1=1, N_2=2]$$

$$= (1-p)(p/2)$$

$$P[N_2=3] = P[N_1=1, N_2=3] + P[N_1=2, N_2=3]$$

$$= (1-p)(p/2)(1-p/2)$$

$$+ p(1-p)(1-p/2)$$

$$\vdots \\ P[N_2=k] = \sum_{n_1=1}^{k-1} P[N_1=n_1, N_2=k]$$

Wishing
as
a
sum
4

$$= \sum_{n_1=1}^{k-1} p^{n_1-1}(1-p)(p/2)^{k-n_1-1}(1-p/2)$$

proper substitution 4

So the PMF of N_2 is:

$$P(N_2 = k) = \begin{cases} \sum_{n_1=1}^{k-1} p^{n_1-1}(1-p) \left(\frac{p}{2}\right)^{k-n_1-1} \left(1-\frac{p}{2}\right), & k=2, 3, \dots \\ 0 & \text{otherwise.} \end{cases}$$

Final Answer
2

(Q2)

(1) Let X_1 be the no. of years for speaking for baby born to the chatty family.

Let X_2 be the no. of years to speaking for the baby born to the quiet family.

$$X_1 \sim \text{Unif}(0, 3)$$

$$X_2 \sim \text{Unif}(0, 5)$$

0 marks.

(2) The probability of interest is

$$P[X_2 < X_1].$$
 To calculate this, we

need the joint PDF $f_{X_1, X_2}(x_1, x_2)$ + ??

Assuming $X_1 \perp\!\!\! \perp X_2$ are independent,

We have:

$$f_{X_1, X_2}(x_1, x_2) = f_{X_1}(x_1) f_{X_2}(x_2)$$

$$\begin{aligned}
 &= \begin{cases} \left(\frac{1}{3}\right)\left(\frac{1}{5}\right) & 0 \leq x_1 \leq 3, \\ 0 & \text{otherwise.} \end{cases} \\
 &\quad \text{otherwise.}
 \end{aligned}$$

$$\{P[X_2 < X_1]$$

$$\begin{aligned}
 &= \int_0^3 \int_0^{x_1} f_{X_1, X_2}(x_1, x_2) dx_2 dx_1
 \end{aligned}$$

$$\begin{aligned}
 &= \frac{1}{15} \int_0^3 \left(\int_0^{x_1} dx_2 \right) dx_1
 \end{aligned}$$

$$\begin{aligned}
 &= \frac{1}{15} \int_0^3 x_1 dx_1 = \frac{1}{15} \left(\frac{x_1^2}{2} \right)_0^3
 \end{aligned}$$

$$= \frac{1}{15} \left(\frac{9}{2} \right)$$

$$= \frac{3}{10} //$$

(3) It is given that $X_2 \geq 2$.

We want to calculate

$$P[X_2 < x_1 | X_2 \geq 2].$$

Identifying the
correct probability
2.5

We need the conditional joint PDF

$$\{ f_{X_1, X_2 | X_2 \geq 2}(x_1, x_2)$$

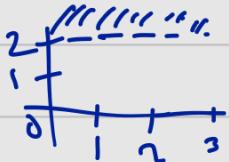
2.5

$$= \begin{cases} \frac{f_{X_1, X_2}(x_1, x_2)}{P[X_2 \geq 2]} & 0 \leq x_1 \leq 3 \\ 0 & 2 \leq x_2 \leq 5, \\ & \text{otherwise.} \end{cases}$$

2.5 $P[X_2 \geq 2] = \int_0^3 \int_2^5 f_{X_1, X_2}(x_1, x_2) dx_2 dx_1$

A simpler way is to $= \frac{1}{15} \int_0^3 3 dx_1 = \frac{9}{15} = \frac{3}{5}$

recognize that $X_2 \sim \text{Unif}(0, 5)$. You could use the marginal PDF.



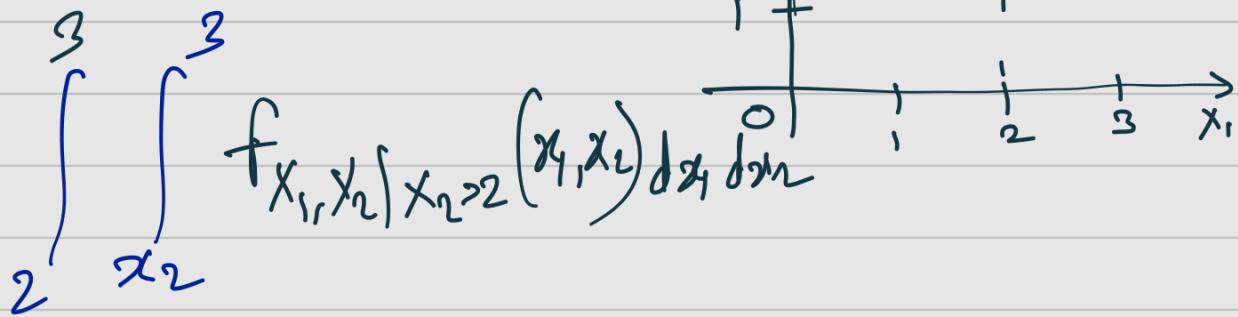
$$f_{X_1, X_2} | X_2 \geq 2 (x_1, x_2)$$

Q.5

$$= \begin{cases} \frac{(1/15)}{(3/5)} = \frac{1}{9} & 0 \leq x_1 \leq 3 \\ 0 & 2 \leq x_2 \leq 5, \\ \text{otherwise.} & \end{cases}$$

$$P[X_2 < x_1 | X_2 \geq 2] =$$

Q.5



$$\begin{aligned} &= \frac{1}{9} \int_2^3 (3 - x_2) dx_2 = \frac{1}{9} \left(3 - \frac{1}{2}(9-4) \right) \\ &= \frac{1}{9} (3 - 2.5) = \frac{1}{18} // \end{aligned}$$

(Q3)

Given estimate

$$J_n(x) = \frac{1}{n-1} \sum_{i=1}^n (x_i + I_i w_i)$$

(1)

$$E[J_n(x)] = E\left[\frac{1}{n-1} \sum_{i=1}^n (x_i + I_i w_i)\right]$$

(2)

$$= \frac{1}{n-1} \sum_{i=1}^n E[x_i + I_i w_i]$$

$$= \frac{1}{n-1} \sum_{i=1}^n E[x_i] + E[I_i w_i]$$

$$= \frac{1}{n-1} \sum_{i=1}^n E[x_i] + E[I_i] E[w_i]$$

$\left\{ \begin{array}{l} \text{Since } I_i \rightarrow \\ w_i \text{ are} \\ \text{independent.} \end{array} \right.$

$$= \frac{1}{n-1} n E[x] \quad \left\{ \begin{array}{l} \text{Since } E[w_i] = 0 \text{ thi} \\ \text{& } x_i \sim X \text{ thi.} \end{array} \right.$$

We have

$$E[J_n(x)] = \frac{n}{n-1} E[x]. \quad \text{⑤}$$

Thus $J_n(x)$ is a ~~not~~ ~~biased~~ estimator.

② MSE is

$$E[(J_n(x) - E[x])^2] \quad \text{②}$$

$$= E[J_n^2(x) + E^2[x] - 2J_n(x)E[x]]$$

$$= E[J_n^2(x)] + E^2[x] - 2E[x]E[J_n(x)]$$

③ of

$$= \boxed{E[J_n^2(x)]} + \boxed{E^2[x] - 2E^2[x]\frac{n}{n-1}}$$

$$= \text{Var}[J_n(x)] + E^2[J_n(x)] + E^2[x] - 2E^2[x]\frac{n}{n-1}.$$

$$= \text{Var}[J_n(x)] + \frac{n^2}{(n-1)^2} E^2[x] + E^2[x]$$

$$- 2E^2[x] \frac{n}{n-1}$$

$$\text{Var}[J_n(x)]$$

$$= \text{Var}\left[\frac{1}{n-1} \sum_{i=1}^n (X_i + I_i W_i) \right]$$

$$= \left(\frac{1}{n-1}\right)^2 \sum_{i=1}^n \text{Var}[X_i + I_i W_i]$$

(Since I_i & W_i
are mutually
independant)

$$= \left(\frac{1}{n-1}\right)^2 \sum_{i=1}^n (\text{Var}[X_i] + \text{Var}[I_i W_i])$$

$$= \frac{n \text{Var}[X_i]}{(n-1)^2} + \underbrace{\frac{1}{(n-1)^2} \text{Var}[I_i W_i]}_{\left(E[I_i^2 W_i^2] - E^2[I_i W_i]\right)}$$

$$= \frac{n \text{Var}[X]}{(n-1)^2} + \frac{1}{(n-1)^2} (0.5) \sigma^2.$$

\therefore MSE is

$$\text{Var}[J_n(x)] + \frac{n^2}{(n-1)^2} E^2[x] + E^2[x]$$

$$- 2 E^2[x] \frac{n}{n-1}$$

$$= \frac{n \text{Var}[x]}{(n-1)^2} + \frac{1}{(n-1)^2} (0.5) \sigma^2.$$

$$+ E^2[x] \left(\frac{n^2 + (n-1)^2 - 2n(n-1)}{(n-1)^2} \right)$$

$$= \frac{n \text{Var}(x)}{(n-1)^2} + \frac{0.5 \sigma^2}{(n-1)^2}$$

$$+ E^2[x] \left(\frac{1}{(n-1)^2} \right)$$

Final
Answer

②

(3) We calculated:

$$E[J_n(x)] = \frac{n}{n-1} E[x].$$

$$\lim_{n \rightarrow \infty} E[J_n(x)] = \lim_{n \rightarrow \infty} \frac{n}{n-1} E[x] = E[x]. \quad ①$$

The estimate is unbiased in the limit as $n \rightarrow \infty$.

The MSE is:

$$\frac{n \text{Var}(x)}{(n-1)^2} + \frac{0.5\sigma^2}{(n-1)^2} + E^2(x) \left(\frac{1}{(n-1)^2} \right)$$

1.5

In the limit as $n \rightarrow \infty$, the MSE $\rightarrow 0$.

Q4

$$E(x) = 40; \quad \sigma_x^2 = (100)(0.4)(0.6) = 24$$

2

$$\begin{aligned} P[x < 50] &= P[x - E(x) < 50 - E(x)] \\ &= P\left[\frac{x - E(x)}{\sigma_x} < \frac{50 - E(x)}{\sigma_x}\right] \quad 5 \\ &\approx \Phi\left(\frac{50 - E(x)}{\sigma_x}\right) \quad 3 \\ &= \Phi\left(\frac{10}{\sqrt{24}}\right) \end{aligned}$$

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Alternatively,

$$P[X < 50] = P[0 \leq X < 50]$$

$$\textcircled{3} = P\left[\frac{-E(X)}{\sigma_X} \leq \frac{X - E(X)}{\sigma_X} \leq \frac{50 - E(X)}{\sigma_X}\right]$$

$$\textcircled{3} = \Phi\left(\frac{50 - E(X)}{\sigma_X}\right) - \Phi\left(\frac{-E(X)}{\sigma_X}\right)$$

\textcircled{8}