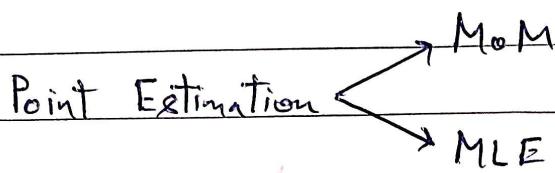


Tutorial 3



MoM → equate population moments & sample moments
 → 'p' population parameters (unknown) \Rightarrow solving 'p'

MLE → maximize likelihood function / log likelihood fn.

$$L(\theta) = \prod_{i=1}^n f_\theta(x_i)$$

Q1. $X_1, X_2, \dots, X_n \sim \text{Poisson}(\lambda)$

a) estimate λ using MoM

First sample moment $m_1 = \frac{1}{n} \sum x_i$

First popn-moment $E[X] = \lambda$

$$\hat{\lambda}_{\text{mom}} = \frac{1}{n} \sum_{i=1}^n x_i = \bar{X}$$

$$\boxed{\hat{\lambda}_{\text{mom}} = \bar{X}}$$

b) Using MLE

$$L(\theta) = \prod_{i=1}^n f_\theta(x_i)$$

$$f(\lambda) = e^{-\lambda} \lambda^{x_i}$$

$$L(\lambda) = \prod_{i=1}^n \frac{e^{-\lambda} \lambda^{x_i}}{(x_i)!}$$

$$L(\lambda) = \frac{e^{-\lambda n} \lambda^{\sum \eta_i}}{\prod_{i=1}^n (\pi_i(\lambda))}$$

$$\begin{aligned}\log L(\lambda) &= \log(e^{-\lambda n} \cdot \lambda^{\sum \eta_i}) - \log\left(\prod_{i=1}^n (\pi_i(\lambda))\right) \\ &= -n\lambda + \sum \eta_i \log \lambda - \sum_{i=1}^n \log(\pi_i(\lambda))\end{aligned}$$

$$\frac{d \log L(\lambda)}{d\lambda} = -n + \sum \eta_i \xrightarrow{\text{Set } \frac{d \log L(\lambda)}{d\lambda} = 0} 0$$

$$\frac{d l}{d \lambda} = 0$$

$$\lambda_{\text{opt}} = \frac{\sum \eta_i}{n}$$

$$\frac{d^2 l}{d \lambda^2} = \frac{-\sum \eta_i}{n^2} \xrightarrow{\text{Set } \frac{d^2 l}{d \lambda^2} = 0} \lambda = \bar{x}$$

$$\bar{x} = \frac{n \bar{\eta}}{\bar{\eta}^2} \xrightarrow{\text{Set } \frac{d^2 l}{d \lambda^2} < 0} 0$$

Q2. x_1, \dots, x_n are samples from the pdf $f_\theta(x)$

$$f_\theta(x) = \theta x^{-2} \quad 0 < \theta \leq x$$

a) Find suff. stats for θ ; we will use the factorization theorem

$$f(x_1, \dots, x_n | \theta) = \prod_{i=1}^n \frac{\theta}{x_i^2} I(x_i \geq \theta)$$

$$= \underbrace{\frac{1}{\pi(x_i)^2}}_{h(x)} \times \underbrace{\theta^n I(x_i \geq \theta)}_{g(T(x)|\theta)}$$

$$h(x) \quad g(T(x)|\theta)$$

$$T(x) = X_{(1)}$$

b) estimate θ using MoM

$$E(X) = \int_0^\infty x f(x) dx$$

$$= \int_0^\infty x \times \frac{\theta}{x^2} dx$$

$$= \theta \log \left| \frac{\infty}{0} \right|$$

$$\text{If } E(X) = \alpha$$

doesn't

MoM estimator doesn't exist here.

So, $\hat{\theta}$ can't be estimated using MoM.

Q3. $X_1, X_2 \dots X_n \sim \text{Binomial}(n, p)$

need to find MoM estimates for n, p

$$E(X) = np \quad \text{Var}(X) = np(1-p) = np(1-p)$$

Now, there are 2 unknown parameters n & p that we need to estimate;

so we will use the first 2 sample moment & form 2 eqns -

$$m_1 = \frac{1}{n} \sum x_i = \bar{x} E(X) \quad \text{--- (1)}$$

$$m_2 = \frac{1}{n} \sum x_i^2 = E(X^2) \quad \text{--- (2)}$$

$$\bar{x} = np ; \frac{1}{n} \sum x_i^2 = \text{Var}(X) + [E(X)]^2 \\ = np(1-p) + (np)^2$$

$$\frac{1}{n} \sum x_i^2 = np(1-p) + (\bar{x})^2$$

$$\frac{1}{n} \sum x_i^2 - (\bar{x})^2 = np(1-p)$$

$$\text{pop. variance } S^2 = \frac{1}{n} \sum (x_i - \bar{x})^2$$

$$= \frac{1}{n} \left(\sum x_i^2 - 2\bar{x} \sum x_i + n(\bar{x})^2 \right) \quad ; \quad n\bar{x} = \sum x_i \\ = \frac{1}{n} \left(\sum x_i^2 - 2n(\bar{x})^2 + n(\bar{x})^2 \right)$$

$$\frac{1}{n} \sum x_i^2 - (\bar{x})^2 = S^2$$

$$S^2 = np(1-p)$$

$$\text{Now } \bar{x} = np \hat{p} = (x) \bar{x}$$

$$S^2 = \bar{x}(1-\bar{x})$$

$$1 - \frac{S^2}{\bar{x}} = \hat{p}(1-\hat{p})$$

$$\boxed{\hat{p}_{\text{max}} (\hat{p}_{\text{max}}) = \frac{\bar{x} - S^2}{\bar{x}}} \quad \boxed{(1-\bar{x}) = \frac{\bar{x} - S^2}{\bar{x}}}$$

$$\bar{x} = \hat{p}_{\text{max}}$$

$$n_{\text{max}} = \frac{\bar{x}}{(\bar{x} - S^2)}$$

$$\boxed{\frac{\bar{x}}{(\bar{x} - S^2)} = \hat{n}_{\text{max}}}$$

$$Q5 \quad x_1, x_2, \dots, x_n \sim U(\theta, \theta+1)$$

$$f_\theta(x) = \frac{1}{b-a} = \frac{1}{(\theta+1)-\theta} = \frac{1}{1} = 1$$

$$L(\theta) = \begin{cases} 1 & \text{if } \theta \leq x_i \leq \theta + 1 \\ 0 & \text{otherwise} \end{cases}$$

Use MLE to estimate θ

$$\theta \leq x_i \Rightarrow x_i - \theta \geq 0 \quad \text{for all } i \quad \text{①}$$

$$\Rightarrow \theta \leq x_{(1)} < x_{(2)} < \dots < x_{(n)} \text{ and } \min(x_1, \dots, x_n) - \theta \geq 0 \quad \text{②}$$

$$\Rightarrow \min(x_1, x_2, \dots, x_n) - 1 \leq \theta$$

∴ from ① & ②

$$\min(x_1, \dots, x_n) - 1 \leq \hat{\theta}_{MLE} \leq \min(x_1, x_2, \dots, x_n)$$

Q6. a) $X_i \sim \text{Binomial}(3, \theta)$

$$(x_1=1, x_2=3, x_3=2, x_4=2) \quad n=4$$

need to estimate θ using MLE based on the above observations

$$L(\theta) = \prod_{i=1}^4 \binom{3}{x_i} \theta^{x_i} (1-\theta)^{3-x_i} \quad \text{constant (not depending on } \theta \text{)}$$

$$\log L(\theta) = \sum_{i=1}^4 (x_i \log \theta + (3-x_i) \log(1-\theta) + \log C)$$

$$\frac{d \log L(\theta)}{d\theta} = \sum_{i=1}^4 \frac{x_i}{\theta} - \frac{3-x_i}{1-\theta} + 0$$

$$= \frac{1}{\theta} - \frac{2}{1-\theta} + \frac{3}{\theta} - 0 + \frac{2}{\theta} - \frac{1}{1-\theta} + \frac{2}{\theta} - \frac{1}{1-\theta}$$

$$= \frac{8}{\theta} - \frac{4}{1-\theta}$$

$$\frac{d \log(L(\theta))}{d\theta} = 0$$

$$\Rightarrow \frac{8}{\theta} = \frac{4}{1-\theta} \Rightarrow 2 - 2\theta = 0$$

$$\boxed{\theta = 2/3}$$

$$\text{check } \frac{d^2 L}{d\theta^2} < 0$$

b) $x_i \sim \text{exponential}(\theta)$

$$(x_1 = 1.23, x_2 = 3.32, x_3 = 1.98, x_4 = 2.12)$$

$$f_\theta(x_i) = \theta e^{-\theta x_i}$$

$$L(\theta) = \prod_{i=1}^4 \theta e^{-\theta x_i} = \theta^4 e^{-\theta \sum_{i=1}^4 x_i}$$

$$\log L(\theta) = \sum_{i=1}^4 4 \log \theta - \theta \sum_{i=1}^4 x_i$$

$$\frac{d \log L(\theta)}{d \theta} = \frac{4}{\theta} - \sum_{i=1}^4 x_i$$

$$\frac{d \log L(\theta)}{d \theta} = 0 = \frac{4}{\theta} - (1.23 + 3.32 + 1.98 + 2.12)$$
$$\boxed{\theta = \frac{4}{8.65} = 0.46}$$

$$\text{check } \frac{d^2 l}{d \theta^2} < 0$$