

Answers to the point estimation problems in the homework 2

2. Point Estimation

1. parameter p , Bernoulli(p) sample of size n

Likelihood is the probability mass function for a random Variable x

$$f(x; p) = p^x (1-p)^{1-x}$$

$$\text{Likelihood } f(x_1, \dots, x_n; p) = \prod_{i=1}^n p^{x_i} (1-p)^{1-x_i}$$

$$\text{Log Likelihood } \ln f(x_1, \dots, x_n; p) = \sum_{i=1}^n (x_i \ln p + (1-x_i) \ln(1-p))$$

Take derivative and solve for p

$$\frac{\partial \ln f}{\partial p} = \frac{\sum_{i=1}^n x_i}{p} - \frac{\sum_{i=1}^n (1-x_i)}{1-p} = 0$$

$$= \frac{\sum_{i=1}^n x_i}{p} = \frac{(n - \sum_{i=1}^n x_i)}{1-p}$$

$$= \frac{\sum_{i=1}^n x_i}{p(1-p)} = \frac{n}{1-p}$$

$$\boxed{\hat{p} = \frac{1}{n} \sum_{i=1}^n x_i}$$

2. parameter p , Binomial (N, p) sample size n

$$f(x) = \left(\frac{N!}{x!(N-x)!} \right) p^x (1-p)^{N-x}$$

$$f(x_1, \dots, x_n; p) = \prod_{i=1}^n \left(\frac{N!}{x_i!(N-x_i)!} \right) p^{x_i} (1-p)^{N-x_i}$$

$$\ln f(x_1, \dots, x_n; p) = \ln \left(\prod_{i=1}^n \left(\frac{N!}{x_i!(N-x_i)!} \right) \right) + \sum_{i=1}^n x_i \ln p + (Nn - \sum_{i=1}^n x_i) \ln(1-p)$$

$$\frac{d \ln f}{dp} = \sum_{i=1}^n \frac{x_i}{p} - \frac{(Nn - \sum_{i=1}^n x_i)}{1-p} = 0$$

$$\sum_{i=1}^n \frac{x_i}{p} = \frac{Nn - \sum_{i=1}^n x_i}{1-p}$$

$$\frac{\sum_{i=1}^n x_i}{p(1-p)} = \frac{Nn}{1-p}$$

$$\boxed{\hat{p} = \frac{\sum_{i=1}^n x_i}{Nn}}$$

Sample $(3, 6, 2, 0, 0, 3)$, $N = 10$

Sample size $= n = 6$

$$\sum_{i=1}^6 x_i = 3 + 6 + 2 + 0 + 0 + 3$$

$$= 14$$

$$N \times n = 10 \times 6 = 60$$

$$\boxed{\hat{p} = \frac{\sum_{i=1}^n x_i}{Nn} = \frac{14}{60} = 0.2333}$$

3. parameters a, b ; Uniform (a, b) size n

$$f(x; a, b) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b \\ 0 & \text{else} \end{cases}$$

$$f(x_1, \dots, x_n; a, b) = \prod_{i=1}^n \left(\frac{1}{b-a} \right) = \frac{1}{(b-a)^n}$$

$$\ln f(x_1, \dots, x_n; a, b) = n \ln \frac{1}{b-a} = -n \ln(b-a)$$

$$\frac{\partial \ln f}{\partial a} = \frac{n}{b-a} \rightarrow \text{strictly increasing}$$

$$\frac{\partial \ln f}{\partial b} = -\frac{n}{b-a} \rightarrow \text{strictly decreasing}$$

$$\hat{a} = \min(x_1, \dots, x_n), \quad \hat{b} = \max(x_1, \dots, x_n)$$

4. parameter μ , Normal (μ, σ^2) , sample size n
unknown σ^2 , unknown μ

$$f(x; \mu) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$$f(x_1, \dots, x_n; \mu) = \prod_{i=1}^n \left(\frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x_i-\mu)^2}{2\sigma^2}} \right)$$

$$= \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right)^n e^{-\sum_{i=1}^n \frac{(x_i-\mu)^2}{2\sigma^2}}$$

$$\ln f(x_1, \dots, x_n; \mu) = \ln \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right)^n - \sum_{i=1}^n \frac{(x_i-\mu)^2}{2\sigma^2}$$

$$\frac{d \ln f}{d \mu} = \frac{1}{2} \sum_{i=1}^n \frac{(x_i-\mu)}{\sigma^2} = 0$$

$$\sum_{i=1}^n x_i - n\mu = 0$$

$$n\mu = \sum_{i=1}^n x_i$$

$$\boxed{\hat{\mu} = \frac{\sum_{i=1}^n x_i}{n}}$$

5. parameter σ , Normal (μ, σ^2)

$$f(x_i; \sigma) = \frac{1}{\sqrt{2\pi}\sigma^2} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}}$$

$$f(x_1, \dots, x_n; \sigma) = \prod_{i=1}^n \left(\frac{1}{\sqrt{2\pi}\sigma^2} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}} \right) = \frac{1}{(\sqrt{2\pi}\sigma^2)^n} e^{-\sum_{i=1}^n \frac{(x_i - \mu)^2}{2\sigma^2}}$$

$$\ln f(x_1, \dots, x_n; \sigma) = -\frac{n}{2} \ln(2\pi) - n \ln(\sigma) - \sum_{i=1}^n \frac{(x_i - \mu)^2}{2\sigma^2}$$

$$\frac{d \ln f}{d \sigma} = -\frac{n}{\sigma} - \sum_{i=1}^n \frac{(x_i - \mu)^2}{2} (-2\sigma^{-3}) = 0$$

$$\frac{n}{\sigma} = \sum_{i=1}^n \frac{(x_i - \mu)^2}{\sigma^3}$$

$$\boxed{\hat{\sigma} = \frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2}$$

6 parameters (μ, σ^2) Normal (μ, σ^2) sample size n

$$f(x; \mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma^2} e^{-\frac{(x - \mu)^2}{2\sigma^2}}$$

$$f(x_1, \dots, x_n; \mu, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}\sigma^2} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}}$$

$$\ln f(x_1, \dots, x_n; \mu, \sigma^2) = \sum_{i=1}^n \left(-\frac{1}{2} \ln(2\pi\sigma^2) - \frac{(x_i - \mu)^2}{2\sigma^2} \right)$$

$$\frac{\partial \ln f}{\partial \mu} = \sum_{i=1}^n \frac{(x_i - \mu)}{\sigma^2} = 0$$

$$\boxed{\hat{\mu} = \frac{1}{n} \sum_{i=1}^n x_i}$$

$$\frac{\partial \ln f}{\partial \sigma^2} = \sum_{i=1}^n \left(-\frac{1}{2} \frac{1}{\sigma^2} - \frac{(x_i - \mu)^2}{2\sigma^4} \right) = 0$$

$$\frac{n}{2\sigma^2} = \sum_{i=1}^n \frac{(x_i - \mu)^2}{2\sigma^4}$$

$$\boxed{\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2}$$

- ② Coin and Thumbtack, Coin - 60 head, 40 tails
Thumbtack - 70 head, 30 tails

Priors $\text{Beta}(1,1)$, $\text{Beta}(40,60)$, $\text{Beta}(20,70)$

$\text{Beta}(100,100)$, $\text{Beta}(1000,1000)$, $\text{Beta}(100,000,100,000)$

- 1) MLE and MAP for coin and thumbtack

Let θ be parameter to maximize probability for D

$$P(\text{Heads}) = \theta, P(\text{tails}) = 1 - \theta \rightarrow \theta \in [0,1]$$

α_H Heads, α_T tails, Data \rightarrow data of sample observations

Likelihood $P(D|\theta) = \theta^{\alpha_H} (1-\theta)^{\alpha_T}$

Log likelihood $\ln P(D|\theta) = \alpha_H \ln \theta + \alpha_T \ln(1-\theta)$

$$\frac{d \ln P(D|\theta)}{d\theta} = \frac{\alpha_H}{\theta} - \frac{\alpha_T}{1-\theta} = 0 \Rightarrow \alpha_H - \theta \alpha_H = \theta \alpha_T$$

$$\left| \hat{\theta}_{MLE} = \frac{\alpha_H}{\alpha_H + \alpha_T} \right|$$

$$\hat{\theta}_{MLE \text{ Coin}} = \frac{60}{100} = 0.6$$

$$\hat{\theta}_{MLE \text{ thumbtack}} = \frac{70}{100} = 0.7$$

MAP $\rightarrow P(\theta|D)$

$$P(\theta|D) = \frac{P(D|\theta)P(\theta)}{P(D)} \quad \text{Bayes theorem}$$

$$\hat{\theta}_{MAP} \rightarrow \arg \max_{\theta} P(\theta|D)$$

$$= \arg \max_{\theta} P(D|\theta) P(\theta)$$

$P(D)$ doesn't depend on θ

We assume prior $P(\theta)$ is Beta Distribution as $\theta \in [0,1]$

$$P(\theta) = \frac{\theta^{\beta_H-1} (1-\theta)^{\beta_T-1}}{B(\beta_H, \beta_T)} \sim \text{Beta}(\beta_H, \beta_T)$$

$$P(D|\theta) = \theta^{\alpha_H} (1-\theta)^{\alpha_T}$$

$$P(\theta|D) = \theta^{\alpha_H} (1-\theta)^{\alpha_T} \times \frac{\theta^{\beta_H-1} (1-\theta)^{\beta_T-1}}{B(\beta_H, \beta_T)}$$

$$\ln P(\theta|D) = \alpha_H \ln \theta + \alpha_T \ln(1-\theta) + (\beta_H-1) \ln \theta + (\beta_T-1) \ln(1-\theta) + \ln B(\beta_H, \beta_T)$$

$$\frac{d \ln P(\theta|D)}{d\theta} = \frac{\alpha_H}{\theta} + \frac{\alpha_T}{1-\theta} + \frac{\beta_H-1}{\theta} + \frac{\beta_T-1}{1-\theta} = 0$$

$$\alpha_H - \theta \alpha_H + \theta \alpha_T + \beta_H - 1 + \theta \beta_H - \theta + \theta \beta_T - \theta = 0$$

$$\boxed{\hat{\theta}_{MAP} = \frac{\alpha_H + \beta_H - 1}{\alpha_H + \beta_H + \alpha_T + \beta_T - 2}}$$

2) Coin Beta (1,1)
 Head = 60 $\hat{\theta}_{MLE} = 0.6$
 Tails = 40 $\hat{\theta}_{MAP} = \frac{60+1-1}{100+2-2} = 0.6$

$$\hat{\theta}_{MLE} = \frac{\alpha_H}{\alpha_H + \alpha_T} \quad \hat{\theta}_{MAP} = \frac{\alpha_H + \beta_H - 1}{\alpha_H + \beta_H + \alpha_T + \beta_T - 2}$$

Beta (40, 60)

$$\hat{\theta}_{MLE} = 0.6$$

$$\hat{\theta}_{MAP} = \frac{60+40-1}{100+100-2} = \frac{99}{198} = 0.5$$

for Beta (1,1)

$$\hat{\theta}_{MLE} = \hat{\theta}_{MAP}$$

MLE is special case of MAP when prior is uniform

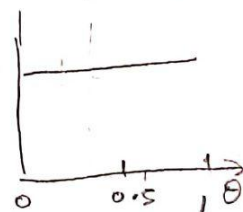
Beta (30, 70)

$$\hat{\theta}_{MLE} = 0.6$$

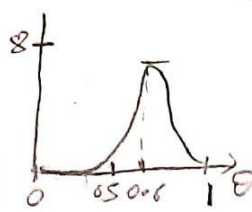
$$\hat{\theta}_{MAP} = \frac{60+30-1}{100+100-2} = \frac{89}{198} = 0.45$$

The stronger or larger the prior is the stronger the effect on curve due to the parameter

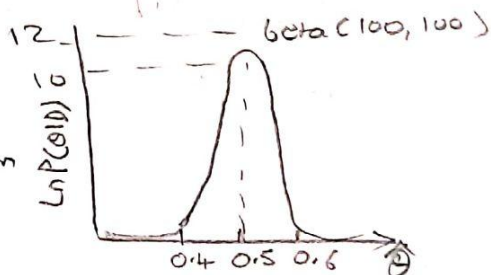
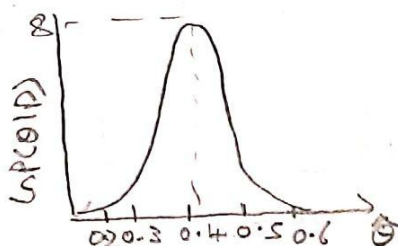
Beta (1,1)



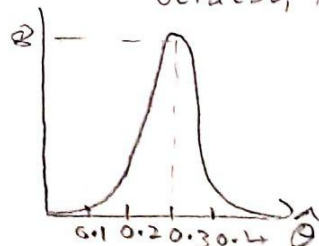
Beta (61, 41)



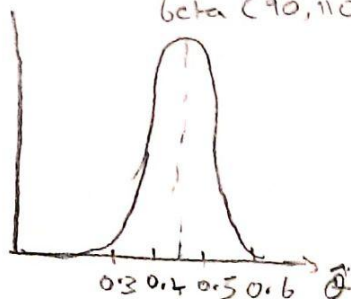
Beta (40, 60)



Beta (30, 70)



Beta (90, 110)



Thumbtack

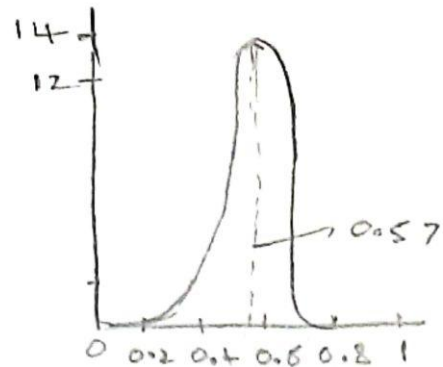
Heads = 70

Tails = 30

beta(100, 100)

$$\theta_{MLE} = 0.7$$

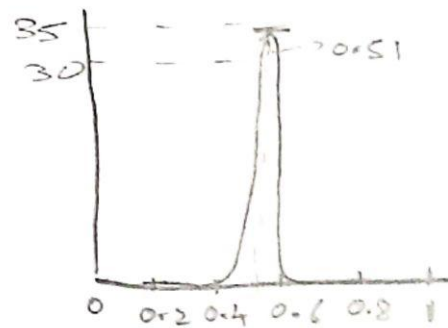
$$\begin{aligned}\theta_{MAP} &= \frac{70 + 100 - 1}{100 + 200 - 2} \\ &= \frac{169}{298} = 0.57\end{aligned}$$



beta(1000, 1000)

$$\theta_{MLE} = 0.7$$

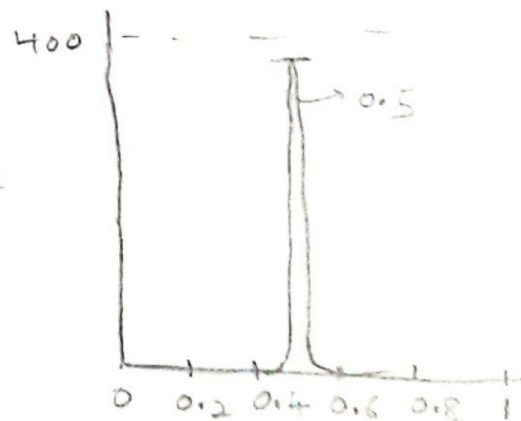
$$\begin{aligned}\theta_{MAP} &= \frac{1000 + 70 - 1}{1000 + 100 - 2} \\ &= \frac{1069}{2098} = 0.51\end{aligned}$$



beta(100,000, 100,000)

$$\theta_{MLE} = 0.7$$

$$\begin{aligned}\theta_{MAP} &= \frac{100,000 + 70 - 1}{200,000 + 100 - 2} \\ &= \frac{100,069}{200,098} = 0.5\end{aligned}$$



The answer to the above problem is that the larger the prior will be the more it will effect the value of the parameter. As seen in the above curves as the prior reaches the value $\beta(100,000, 100,000)$ the parameter ap

3) False, the MLE estimate will approach the MAP estimate just once when for $B(\alpha, \beta)$
 $\alpha=1, \beta=1$.

4) True, when we use a stronger prior the effect of the prior will be weak for small sample size
eg Thumbtack and coin the θ_{MAP} value for prior $B(100,000, 100,000)$ is same
coin $\theta_{MLE} = 60$ $\hat{\theta}_{MAP} = 0.5$
Thumbtack $\theta_{MLE} = 70$ $\hat{\theta}_{MAP} = 0.5$