Appendix

Q1-a

```
prob1_dt <- read.table('problem1.txt', header=T)</pre>
cal_dt <- prob1_dt$calcium</pre>
n <- length(cal_dt)</pre>
cal_resid <- (cal_dt-mean(cal_dt)) / (500 * sqrt(1 - 1/n))
plot(cal_resid, ylim=c(-4, 4), ylab='Residual', main='Calcium Residual Plot')
abline(h=0)
qqnorm(cal_resid)
abline(a=0, b=1)
Q1-b
qqnorm(cal_resid)
abline(a=0, b=1)
Q1-c
cal_r <- cal_dt - mean(cal_dt)</pre>
cal_df <- length(cal_r) - 1</pre>
D \leftarrow sum(cal_r^2) / 500^2
p_value <- pchisq(D, cal_df)</pre>
Q1-d
count <- c(length(cal_dt[cal_dt <= 600]),</pre>
            length(cal_dt[cal_dt > 600 & cal_dt <= 1200]),</pre>
            length(cal_dt[cal_dt > 1200 & cal_dt <= 1800]),</pre>
            length(cal_dt[cal_dt > 1800]))
sigma <- 500
likelifun <- function(x) {</pre>
  return(pnorm((600-x)/sigma)^9 * (pnorm((1200-x)/sigma) - pnorm((600-x)/sigma))^20
         * (pnorm((1800-x)/sigma) - pnorm((1200-x)/sigma))^7
         * (1 - pnorm((1800-x)/sigma))^2)
}
theta_seq <- seq(400, 2000)
likeli_value <- likelifun(theta_seq)</pre>
plot(x=theta_seq, y=likeli_value, xlab='Theta', ylab='Likelihood')
# find mle theta
```

```
LL <- function(theta) {
 p1 <- pnorm((600-theta)/sigma)
 p2 <- pnorm((1200-theta)/sigma) - pnorm((600-theta)/sigma)
 p3 <- pnorm((1800-theta)/sigma) - pnorm((1200-theta)/sigma)
 p4 <- 1 - pnorm((1800-theta)/sigma)
 x1 <- 9
 x2 <- 20
 x3 <- 7
 x4 < -2
 mloglik \leftarrow suppressWarnings(-x1 * log(p1) - x2 * log(p2) - x3 * log(p3) - x4 * log(p4))
}
library(stats4)
mle(LL, start=list(theta=1000))
# chi-square test
theta <- 914.2893
p1 <- pnorm((600-theta)/sigma)
p2 <- pnorm((1200-theta)/sigma) - pnorm((600-theta)/sigma)
p3 <- pnorm((1800-theta)/sigma) - pnorm((1200-theta)/sigma)
p4 <- 1 - pnorm((1800-theta)/sigma)
probs <- c(p1, p2, p3, p4)
chisq.test(count, probs)
expected <- probs * 38
D <- sum((count - expected)^2 / expected)</pre>
p_value <- 1- pchisq(D, df=2)</pre>
\mathbf{Q2}
# find mle of sigma theta
LL <- function(theta, sigma) {
 p1 <- pnorm((600-theta)/sigma)
 p2 <- pnorm((1200-theta)/sigma) - pnorm((600-theta)/sigma)
 p3 <- pnorm((1800-theta)/sigma) - pnorm((1200-theta)/sigma)
 p4 <- 1 - pnorm((1800-theta)/sigma)
 x1 <- 9
 x2 <- 20
 x3 <- 7
 x4 < -2
 mloglik \leftarrow suppressWarnings(-x1 * log(p1) - x2 * log(p2) - x3 * log(p3) - x4 * log(p4))
}
mle(LL, start=list(theta=1000, sigma=500))
# chi-square test
```

```
theta <- 917.6799
sigma <- 477.9240
p1 <- pnorm((600-theta)/sigma)
p2 <- pnorm((1200-theta)/sigma) - pnorm((600-theta)/sigma)
p3 <- pnorm((1800-theta)/sigma) - pnorm((1200-theta)/sigma)
p4 <- 1 - pnorm((1800-theta)/sigma)
probs <- c(p1, p2, p3, p4)
expected <- probs * 38
D <- sum((count - expected)^2 / expected)</pre>
p_value <- 1- pchisq(D, df=1)</pre>
Q_5
# problem 5-a
diet_pvalue <- read.table('problem4.txt', header=T, sep=',')</pre>
bf_pvalue <- p.adjust(diet_pvalue$P.value, 'bonferroni')</pre>
data.frame(Variable=diet_pvalue$Dietary.Variable, Bonferroni.P = bf_pvalue)
# problem 5-b
m <- nrow(diet_pvalue)</pre>
ordered_pvalue <- diet_pvalue[order(diet_pvalue$P.value),]</pre>
ben_pvalue <- rep(NA, m)
k_reject <- 0
for (ii in 1:m) {
  adj_pvalue <- ordered_pvalue$P.value[ii] * m / ii
  ben_pvalue[ii] <- adj_pvalue</pre>
  if (adj_pvalue < 0.05) {
    k_reject <- ii
  }
ordered_pvalue[1:k_reject,]
plot(x=1:m, y=ben_pvalue, xlab='K', ylab='Benjamini Adjusted', pch=16)
abline(h=0.05, lty=3, lwd=2)
legend('topleft', legend='Test Size', lty=3, lwd=2, cex=0.8)
```