		Vame: ID:		
Points (Stat 421, Test 3, Fall, Dec. 13, 2016; Mar ONLY a half-size "cheat sheet" is allowed Multiple choice: Circle all the correct answers; there is wrong-For rest, SHOW answer & work; NO CREDIT for correct answer was allowed.	answer penalty		8
1	1. A filter comes in 3 types, and we want to know if the type has an effect Lab we have 3 operators. I give the 3 filters to operator 1 and ask him to of y on each filter in random order. Then, I give the three filters to instructions, and then to operator 3 with the same instructions. What is a) CRD with 1 treatment factor b) RCBD with 1 treatment and 1 block factor c) CRD with 2 treatment factors	o perform one operator 2, wis the design?	measurer with the s	ment
1	In the previous problem, I still have three filter types, but I also we degree of the operator also has an effect. So, I select 3 operators with degree, and 3 with a PhD. Then I randomly assign the 9 treatment comb What kind of design is this? a) CRD with 1 treatment factor b) RCBD with 1 treatment and 1 block factor c) CRD with 2 treatment factors	a BS degree,	3 with a	ı MS
1 hw-85	3. In a factorial experiment with fixed factors, and with n replications the effects, main and interaction, (e.g., by the corresponding parameter decide to treat the replication as a block factor. Circle all of the quantities a) The effects b) The SS of the treatment factors b) The SS of the treatment factors b) In a 2^3 design with 2 blocks, the principal block is $[(1), a, b, ab]$. Which block? a) A b) B c) C d) AB e) AC	rs in the mode ties that may	el). Then change. (c) (deting) founded block =	n, we
1 let?	In a complete 2^k experiment we have arranged for two effects, say X with block. Then (circle all correct statements) a) X is aliased with Y b) X is confounded with Y c) X 6. In a 2-factor random effects model, the ratio MSA/MSE is an approximately a Y 2-sided test of (See last page) a) Y a) Y confound is with block Confound is with Y	Y is confounder ropriate test s	ed with b	olock for a
1	7. A 2 ³ experiment is supposed to be done across 2 days, with half the reknows only a little bit of statistics) decides to do "systematic sampling runs in Yates' standard order ((1), a, b, ab, c, ac, bc, abc) to different thing about this design? Explain your answer. Day 1: (U), b, c, bc] A = low Day 2: [a, ab, ac, abc] A = high	g" by assignin nt days. Wha	g alternates the w	ating vorse
~ 1	8. Consider the $\pm/-$ table for the wine-tasting experiment given at the	ne end of the	test. An	swer 12 tollow

a) Write the +/- table and identify each of the runs in Yates' notation.

b) Write the C effect in Yates' notation; constants of proportionality are not important.

$$> c \sim -def - af - be - abd + cd + ace + bcf + abcdef$$

need 2 cols in the Table

c) Suppose we need to perform the runs in 4 blocks. Give an example of the effect(s) you would ~ 2 choose to confound with block, if you don't care too much about about the D effect. Don't explain.

 ~ 2 ments (in Yates' notation) of each of the blocks. Denote the blocks with Z_i , $i = 1, 2, \cdots$

 ~ 2 e) In terms of the Z_i in part c, write the block effect that is confounded with the C effect? Hint: This does not require a lot of writing; just look at the blocks and ask yourself which ones have some effect at high level and which ones have that effect at low level.

$$C \sim (2_1 + 2_4) - (2_1 + 2_3)$$
 Z_1 and Z_2 have low C Z_2 and Z_4 have high C

 ~ 2 Find the expression for the probability of obtaining a negative value for the estimate of σ_{α}^2 in a problem involving two random factors. See Table on last page.

- 11. We have data y_{ijk} on two factors A and B, with n replications, where $i = 1, \dots a, j = 1, \dots b, k = 1, \dots n$. Suppose we estimate the A effect at each level of the B factor, separately. The model is then $y_{ijk} = \mu + \alpha_i + \epsilon_{ijk}$ for fixed j.
- ~ 2 a) Find the least-squares estimate $\hat{\alpha}_i$, after imposing some natural constraint(s). Hint: Since j is fixed, the expression for SSE will not have a sum over j.

Yijh =
$$M + \alpha_i + \epsilon_{ijh}$$
 for fixed j (suppressing a j index on M , α_{ij} )

SSE = $\frac{8^n}{ik} (Y_{ijk} - M - \alpha_i)^2$ impose $\hat{\alpha}_{i} = 0$ Constraint.

 $\frac{\partial}{\partial M} : \sum_{ik} (Y_{ijk} - M - \alpha_i)^2 \Rightarrow Y_{ij}, -\hat{M} - \hat{\alpha}_{i} = 0 \Rightarrow \hat{\alpha}_{i} = Y_{ij}, -\hat{Y}_{ij}$
 $\frac{\partial}{\partial \alpha_{i}} : \sum_{k} (Y_{ijk} - M - \alpha_{i})^2 \Rightarrow Y_{ij}, -\hat{M} - \hat{\alpha}_{i} = 0 \Rightarrow \hat{\alpha}_{i} = Y_{ij}, -\hat{Y}_{ij}$

b) Now, consider the full model $y_{ijk} = \mu + \alpha_i + \beta_j + (\alpha\beta)_{ij} + \epsilon_{ijk}$, and recall the following estimates (The symbol has been neglected for convenience): $\mu = \overline{y_{...}}, \alpha_i = \overline{y_{i...}} - \overline{y_{...}}, \beta_j = \overline{y_{.j.}} - \overline{y_{...}}, (\alpha\beta)_{ij} = \overline{y_{ij.}} - \overline{y_{i...}} - \overline{y_{...}} + \overline{y_{...}}$. Is there a relationship between the A effect estimated in part a, and the estimates in the full model? If so, what is it? If not, explain why not?

There are infact 2 relationships (either one is sufficient)

1)
$$di$$
 (full model) = Sample arg (over j) of $\hat{\alpha}_i$; (part α_i)

2) $(\hat{\alpha}_i\beta_i)_{ij} = \overline{Y_{ij}}_{ij} - \overline{Y_{ij}}_{ij} - \overline{Y_{ii}}_{ii} + \overline{Y_{iii}}_{ii}$
 $(\hat{\alpha}_i\beta_i)_{ij} + \hat{\alpha}_i = \overline{Y_{ij}}_{ij} - \overline{Y_{ij}}_{ij} = \hat{\alpha}_i \text{ above}_i$

Since $(\hat{\alpha}_i\beta_i)_{ij} = 0$, taking arg of both sides gives

Run	Variable									
	A	В	C -	D	E	F	G	H		
1	-	-	-	-	-	-	-	_		
2	+	-	1	_	-	+	+	+		
3	-	+	-	-	+	01-	+	+		
4	+	+	-	-	+	+	-	-		
5	111-	-	+	-	+	+	+	-		
6	+	-	+	-	+	-	-	+		
7	1 44	+	+	-	-	+	-	+		
8	+	+	+	-	-	-	+	-		
9	-	-	11	+	+	+	-	+		
10	+	-	-	+	+	-	+	-		
11	-	+	-	+	-	+	+	-		
12	+	+		+	-	-	-	+		
13	-	-	+	+	-	-	+	+		
14	+	-	+	+	-	+	-	-		
15	_	+	+	+	+	-	_	-		
16	+	+	+	+	+	+	+	- +		

(e)
$$2^{6-3}$$
; 1/8 fraction of 6 factors in 8 runs

Design Generators

 $D = AB \quad E = AC \quad F = BC$

Defining relation: $I = ABD = ACE = BCDE = BCF = ACDF = ABEF = DEF$

Aliases

 $A = BD = CE = CDF = BEF \quad E = AC = DF = BCD = ABF$
 $B = AD = CF = CDE = AEF \quad F = BC = DE = ACD = ABE$
 $C = AE = BF = BDE = ADF \quad CD = BE = AF = ABC = ADE = BDF = CEF$
 $D = AB = EF = BCE = ACE$

Table of expected values for 2 random factors:

$$E[MSA] = \sigma_{\epsilon}^{2} + nb\sigma_{\alpha}^{2} + n\sigma_{\alpha\beta}^{2}$$

$$E[MSB] = \sigma_{\epsilon}^{2} + na\sigma_{\beta}^{2} + n\sigma_{\alpha\beta}^{2}$$

$$E[MSAB] = \sigma_{\epsilon}^{2} + n\sigma_{\alpha\beta}^{2}$$

$$E[MSE] = \sigma_{\epsilon}^{2}$$

ACF ~ Block.

9c): We need to pick 2 words/effects and confound each one with block. Given The alias structure, one good choice is a high-order word/effect from the CO = --- line, e-g (CEF) we should not take the other word from the same line, e-g. CD, because then CDXCEF = DEF will also be confounded with block. But DEF = 1 is one of The defining relations (ie. it cannot be confounded with block, because it's not even estimable). For the 2nd one, we could take a high-order effect from the D line, e.g. ACF because according to The problem, we don't cave if The D effect is aliased with something which is itself confounded with block. Of conver, Then CEFX ACF = AE is also confounded with black. But note That AE is aliased with C, and so C gets confounded with block, too. So that's bad! We can try confounding Ditself with block. In That case CEFXD = CDEF is also confounded with block; but Then CDEF is aliased with C again. So, again C will be aliased with something That is confounded with black. At the end, it looks like choosing CEF as one of the effects to confound with blocks leads to undesirable vesults!

So, This was just a bad question

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