

# Random Numbers

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<https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex1/2/main.py>

Run the following command in the terminal to run the code.

```
python3 main.py
```

## 1 UNIFORM RANDOM NUMBERS

Let  $U$  be a uniform random variable between 0 and 1.

- 1.1 Generate  $10^6$  samples of  $U$  using a C program and save into a file called uni.dat .

**Solution:** Download the following files and execute the C program.

```
wget https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex1/1/exrand.c
wget https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex1/1/coeffs.h
```

Use the below command in the terminal to run the code

```
gcc exrand.c -lm
./a.out
```

- 1.2 Load the uni.dat file into python and plot the empirical CDF of  $U$  using the samples in uni.dat. The CDF is defined as

$$F_U(x) = \Pr(U \leq x) \quad (1.1)$$

**Solution:** The graph 1.2 is obtained by running the below code

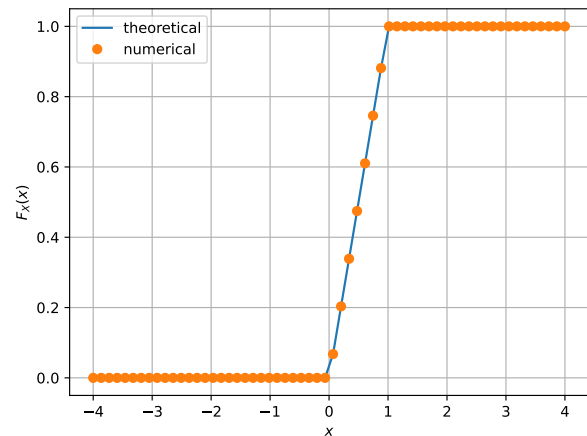


Fig. 1.2: The CDF of  $U$

- 1.3 Find a theoretical expression for  $F_U(x)$ .

**Solution:** Given  $U$  is uniform random variable so

$$f_X(x) = \begin{cases} 1 & 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{aligned} F_X(x) &= \int_{-\infty}^x f_X(x) dx \\ &= 0 + \int_0^x 1 dx \\ &= \begin{cases} 1 & x > 1 \\ x & 0 \leq x \leq 1 \\ 0 & x < 0 \end{cases} \end{aligned}$$

1.4 The mean of  $U$  is defined as

$$E[U] = \frac{1}{N} \sum_{i=1}^N U_i \quad (1.2)$$

and its variance as

$$\text{var}[U] = E[U - E[U]]^2 \quad (1.3)$$

Write a C program to find the mean and variance of  $U$ .

**Solution:**

```
wget https://github.com/NareshBandaru13/
Random-Numbers/blob/main/ex1/4/
exrand.c
wget https://github.com/NareshBandaru13/
Random-Numbers/blob/main/ex1/4/coeffs
.h
```

Use below command to run file,

```
gcc exrand.c -lm
./a.out
```

1.5 Verify your result theoretically given that

$$E[U^k] = \int_{-\infty}^{\infty} x^k dF_X(x)$$

Given

$$E[U^k] = \int_{-\infty}^{\infty} x^k dF_X(x)$$

$$\begin{aligned} E[U] &= \int_{-\infty}^{\infty} x^k f_X(x) dx \\ &= \int_0^1 x \times 1 dx \\ &= \left[ \frac{x^2}{2} \right]_0^1 = \frac{1}{2} \end{aligned}$$

if  $k=2$

$$\begin{aligned} E[U^2] &= \int_0^1 x^2 \times 1 dx \\ &= \left[ \frac{x^3}{3} \right]_0^1 = \frac{1}{3} \end{aligned}$$

$$\begin{aligned} \text{variance} &= E[u - E[u]]^2 \\ &= E[U^2] - E^2[U] \\ &= \frac{1}{3} - \frac{1}{4} = 0.0833 \end{aligned}$$

## 2 CENTRAL LIMIT THEOREM

2.1 Generate  $10^6$  samples of the random variable

$$X = \sum_{i=1}^{12} U_i - 6 \quad (2.1)$$

using a C program, where  $U_i, i = 1, 2, \dots, 12$  are a set of independent uniform random variables between 0 and 1 and save in a file called gau.dat

**Solution:**

```
wget https://github.com/NareshBandaru13/
Random-Numbers/blob/main/ex2/1/
exrand.c
wget https://github.com/NareshBandaru13/
Random-Numbers/blob/main/ex2/1/coeffs
.h
```

Running the above codes generates uni.dat and gau.dat file. Use the command

```
gcc exrand.c -lm
.\a.out
```

2.2 Load gau.dat in python and plot the empirical CDF of  $X$  using the samples in gau.dat. What properties does a CDF have?

**Solution:** The CDF of  $X$  is plotted in plot, Properties of the CDF:

- $F_X(x) = P(X \leq x)$
- $Q_X(x) = P(X > x)$
- $F_X(x) = 1 - Q_X(x)$  This can be used to calculate  $F(x)$ .

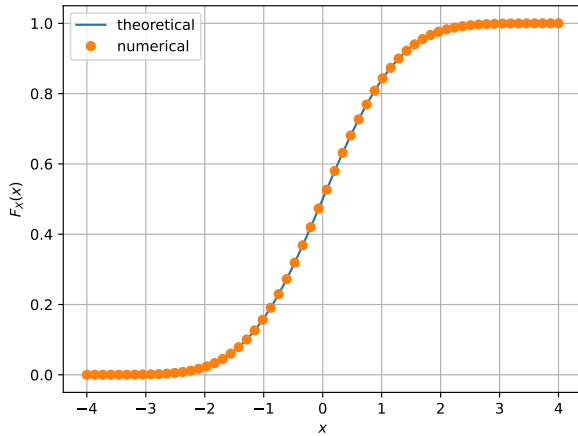


Fig. 2.2: The CDF of  $X$

2.3 Load `gau.dat` in python and plot the empirical PDF of  $X$  using the samples in `gau.dat`. The PDF of  $X$  is defined as

$$p_X(x) = \frac{d}{dx} F_X(x) \quad (2.2)$$

What properties does the PDF have?

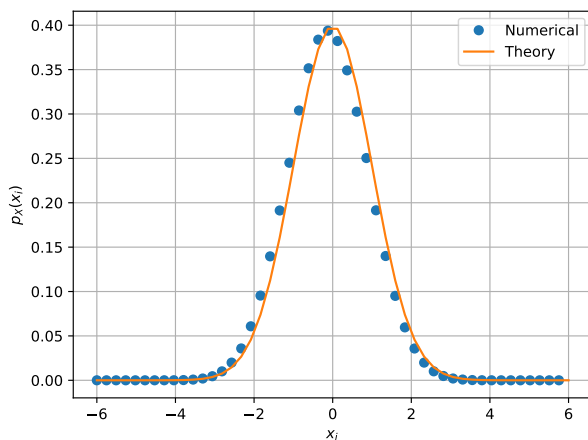


Fig. 2.3: The PDF of  $X$

**Solution:** The PDF of  $X$  is plotted using the code below

<https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex2/3/main.py>

Use the below command to run the code:

`python3 main.py`

Properties of PDF:

- PDF is symmetric about  $x \approx 0$
- graph is similar to bell shaped
- mean of graph is situated at the symmetrical point

2.4 Find the mean and variance of  $X$  by writing a C program.

**Solution:** Running the below code gives  
Mean = 0.000326 Variance= 1.000906

<https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex2/4/exrand.c>

Command used:

```
gcc exrand.c -lm
./a.out
```

2.5 Given that

$$p_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}, -\infty < x < \infty$$

by property of probability

$$\int_{-\infty}^{\infty} p_X(x) dx = 1$$

$$\begin{aligned} F_X(x) &= \int_{-\infty}^x p_X(x) dx \\ &= \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \end{aligned}$$

$$\begin{aligned} E[X] &= \int_{-\infty}^{\infty} x p(x) dx \\ &= \int_{-\infty}^{\infty} \frac{x}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \end{aligned}$$

$\frac{x}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$  is an odd function so integral is zero  
i.e  $E[X] = 0$ .

$$\begin{aligned}
E[X^2] &= \int_{-\infty}^{\infty} x^2 p(x) dx \\
&= x \int_{-\infty}^{\infty} x p(x) dx - \int_{-\infty}^{\infty} \left( \int_{-\infty}^{\infty} x p(x) dx \right) dx \\
&= \left[ -x \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \right]_{-\infty}^{\infty} - \int_{-\infty}^{\infty} -\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\
&= 0 - (-1)
\end{aligned}$$

we know that

$$\int_{-\infty}^{\infty} e^{-\frac{x^2}{2}} = \sqrt{2\pi}$$

by series expansion  $\frac{x}{e^{\frac{x^2}{2}}} = \frac{x}{1 + \frac{x^2}{2} + \frac{x^4}{8} + \dots}$

putting  $x = \infty$ , we get  $\frac{1}{\infty} = 0$

Similarly when  $x = -\infty$  we get 0

$$\text{var}(x) = E[x^2] - E[x]^2 = 1 - 0 = 1$$

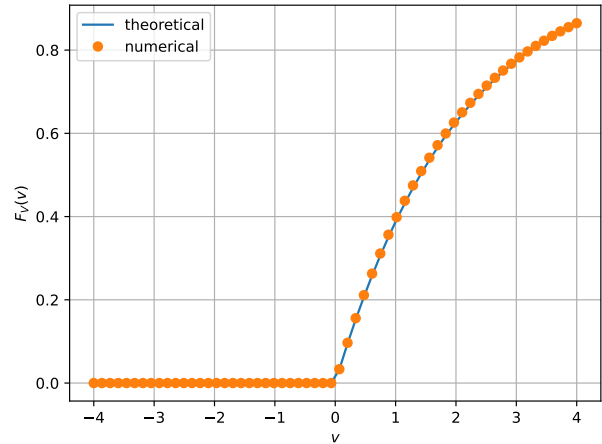


Fig. 3.1: CDF for (3)

<https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex3/cdf.py>

Use the below command to run the code:

```
python3 cdf.py
```

### 3.2 Theoretical expression for $F_V(x)$

$$\begin{aligned}
F_V(x) &= P\{V \leq x\} \\
&= P\{-2 \times \ln(1 - U) \leq x\} \\
&= P\{U \leq 1 - e^{(-\frac{x}{2})}\} \\
&= F_U\{1 - e^{(-\frac{x}{2})}\} \\
&= \begin{cases} 1 - e^{(-\frac{x}{2})} & 0 \leq x < \infty \\ 0 & x < 0 \end{cases}
\end{aligned}$$

## 3 FROM UNIFORM TO OTHER

### 3.1 Generate samples of

$$V = -2 \ln(1 - U) \quad (3.1)$$

and plot its CDF.

**Solution:**

Running the below code generates samples of V from file uni.dat(U).

<https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex3/main.py>

Use the below command in the terminal to run the code:

```
python3 main.py
```

Now these samples are used to plot by running the below code

## 4 TRIANGULAR DISTRIBUTION

### 4.1 Generate

$$T = U_1 + U_2 \quad (4.1)$$

we get the code

<https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex4/1/main.py>

run the command

```
python3 main.py
```

### 4.2 Find the CDF of T. we have code

<https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex4/2/main1.py>

run the command

```
python3 main1.py
```

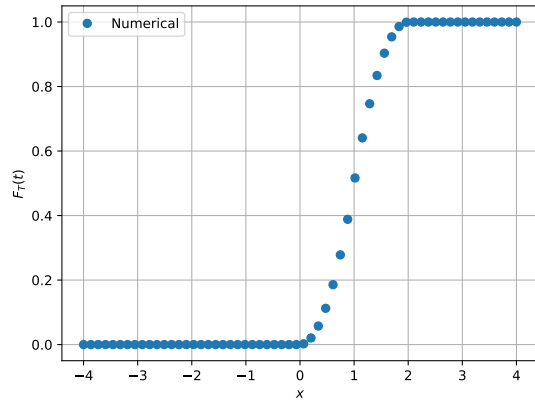


Fig. 4.2: numerical cdf

4.3 Find the PDF of  $T$  we have code

```
https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex4/3/main2.py
```

run the command

```
python3 main2.py
```

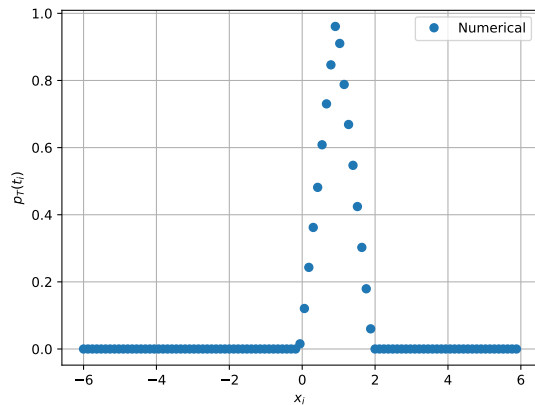


Fig. 4.3: Numerical cdf

4.4 Find the theoretical expressions for the PDF and CDF of  $T$ . **Solution:**

$$\begin{aligned} F_T(t) &= P\{T \leq t\} \\ &= P\{U_1 + U_2 \leq t\} \end{aligned}$$

let us take two cases if  $0 \leq t \leq 1$  and  $1 < t \leq 2$

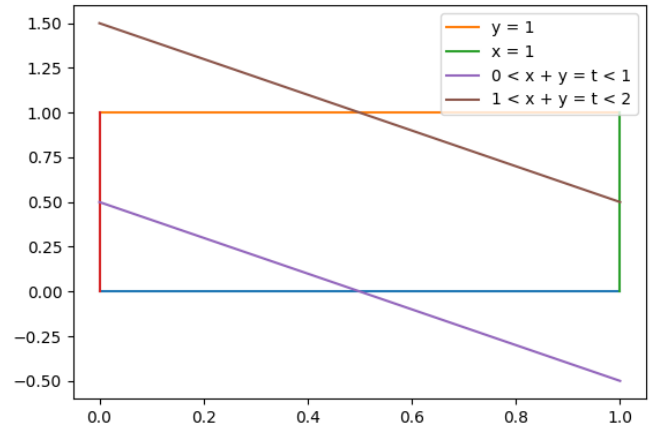


Fig. 4.4: def plot

The above graph is produced by

```
https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex4/2/find.py
```

Run the code in terminal

```
python3 find.py
```

from the figures it is evident that

$$P(U_1 + U_2 < t, 0 \leq t < 1) = \frac{t^2}{2}$$

$$P(U_1 + U_2 < t, 1 \leq t \leq 2) = 1 - \frac{(2-t)^2}{2}$$

$$F_T(t) = \begin{cases} 0 & t < 0 \\ \frac{t^2}{2} & 0 \leq t \leq 1 \\ 1 - \frac{(2-t)^2}{2} & 1 < t \leq 2 \\ 1 & t > 2 \end{cases}$$

$$P_T(t) = \frac{d(F_T(t))}{dt}$$

$$\therefore P_T(t) = \begin{cases} 0 & t < 0 \\ t & 0 \leq t \leq 1 \\ 2 - t & 0 < t \leq 2 \\ 0 & t > 2 \end{cases}$$

4.5 Verify your results through a plot Take the code for cdf

```
https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex4/5/main1.py
```

Run in terminal

```
python3 main1.py
```

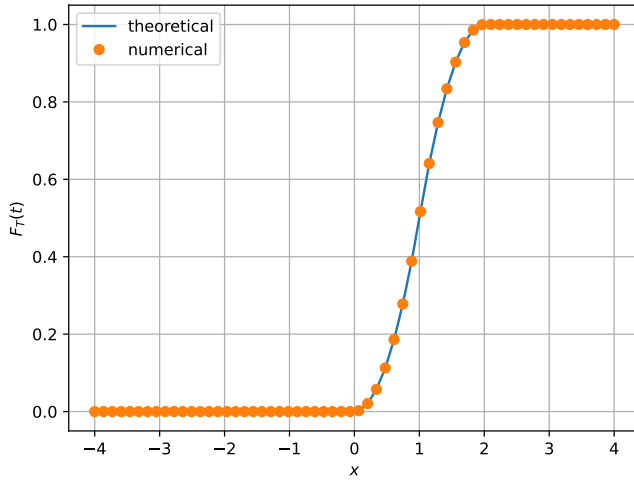


Fig. 4.5: t-cdf

Take the code for pdf

```
https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex4/5/main2.py
```

Run in terminal

```
python3 main2.py
```

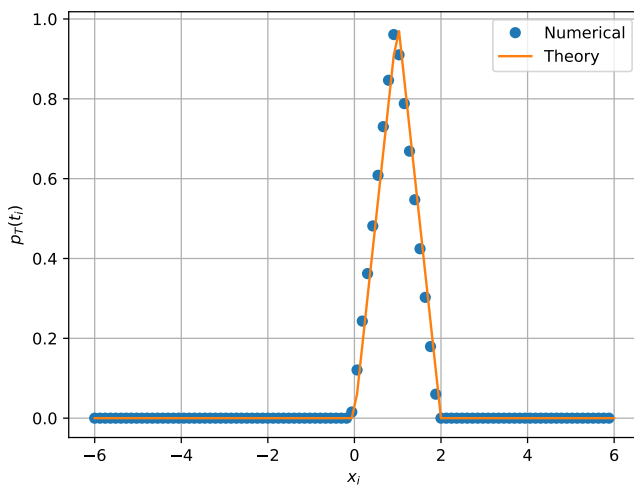


Fig. 4.5: t-pdf

## 5 MAXIMUM LIKELIHOOD

### 5.1 Generate

$$Y = AX + N, \quad (5.1)$$

where  $A = 5$  dB,  $X \in \{1, -1\}$

**Solution:** use bernouli function from exrand.c find the code

```
https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex5/1/exrand.c
```

run the terminal command

```
gcc exrand.c -lm
./a.out
```

### 5.2 Generate

$$Y = AX + N, \quad (5.2)$$

where  $A = 5$  dB, and  $N \sim N(0, 1)$ . find the code

```
https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex5/2/main.py
```

run the command

```
python3 main.py
```

### 5.3 Plot $Y$ using a scatter plot. find the scatter plot

**Solution:**

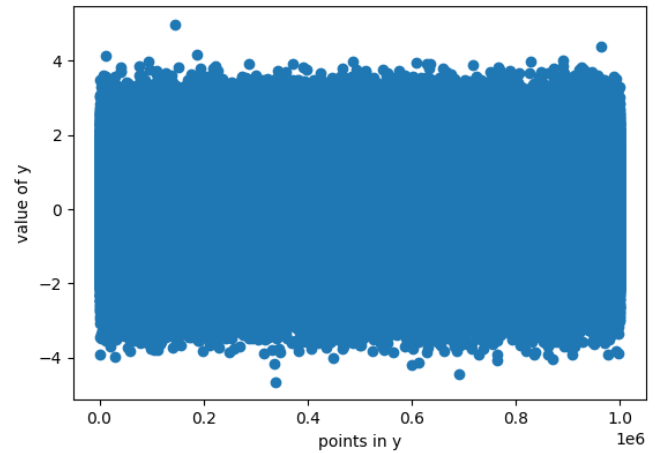


Fig. 5.3: scatter plot

### 5.4 Guess how to estimate $X$ from $Y$ .

**Solution:** To estimate  $X$  from  $Y$ , consider function:

$$\text{sgn}(y) = \begin{cases} -1, & y \in (-\infty, 0] \\ 1, & y \in (0, \infty) \end{cases} \quad (5.3)$$

Using  $\text{sgn}(y)$ , we can operate on  $Y$  to find corresponding values of  $X$ .

### 5.5 Find

$$P_{e|0} = \Pr(\hat{X} = -1 | X = 1) \quad (5.4)$$

and

$$P_{e|1} = \Pr(\hat{X} = 1|X = -1) \quad (5.5)$$

find the code below

<https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex5/5/main.py>

run the code

python3 main.py

we get the values as  $P_{e|0} = 0.310007$   
 $P_{e|1} = 0.310142$

5.6 Find  $P_e$  assuming that  $X$  has equiprobable symbols.

**Solution:** Assume a general value of  $A$ . Our estimation function predicts the data above the  $x$  axis correspond to  $X = 1$ , and the data points below the  $x$ -axis correspond to  $X = -1$ . We have:

$$\begin{aligned} P_{e|0} &= \Pr(\hat{X} = -1|X = 1) \\ &= \Pr(AX + N < 0|X = 1) \\ &= \Pr(N < -A) \\ &= \int_{-\infty}^{-A} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= \int_A^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= Q_N(A) \end{aligned}$$

where  $Q_N$  is the  $Q$ function of the normal distribution. Similarly,

$$\begin{aligned} P_{e|1} &= \Pr(\hat{X} = 1|X = -1) \\ &= \Pr(AX + N > 0|X = -1) \\ &= \Pr(N > A) \\ &= \int_A^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= Q_N(A) \end{aligned}$$

Given  $X$  is equiprobable so we have

$$\begin{aligned} P_e &= P_{e|0} \Pr(X = 1) + P_{e|1} \Pr(X = -1) \\ &= \frac{1}{2} P_{e|0} + \frac{1}{2} P_{e|1} \\ &= \frac{1}{2} Q_N(A) + \frac{1}{2} Q_N(A) \\ &= Q_N(A) \end{aligned}$$

5.7 Verify by plotting the theoretical  $P_e$  with respect to  $A$  from 0 to 10 dB. find the code below

<https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex5/7/main.py>

run the command

python3 main.py

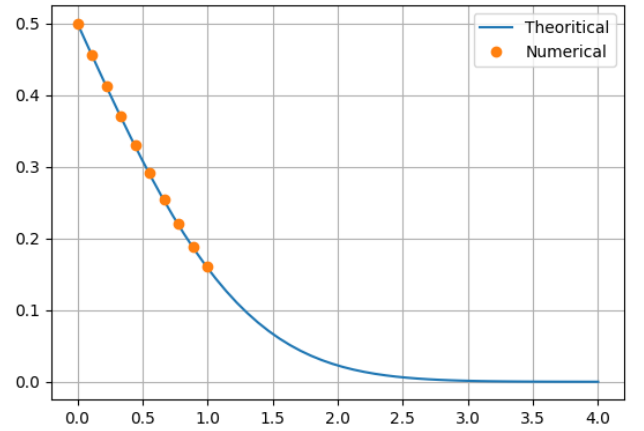


Fig. 5.7: verification of  $p_e$

5.8 Now, consider a threshold  $\delta$  while estimating  $X$  from  $Y$ . Find the value of  $\delta$  that maximizes the theoretical  $P_e$ .

**Solution:** To estimate  $X$  from  $Y$ , we consider

$$X = \begin{cases} 1, & Y > \delta \\ -1, & Y < \delta \end{cases}$$

so we have

$$\begin{aligned} P_{e|0} &= \Pr(\hat{X} = -1|X = 1) \\ &= \Pr(AX + N < \delta|X = 1) \\ &= \Pr(N < \delta - A) \\ &= \int_{-\infty}^{\delta-A} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= \int_{A-\delta}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\ &= Q_N(A - \delta) \end{aligned}$$

$$\begin{aligned}
P_{e|1} &= \Pr(\hat{X} = 1|X = -1) \\
&= \Pr(AX + N > \delta|X = -1) \\
&= \Pr(N > \delta + A) \\
&= \int_{A-\delta}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \\
&= Q_N(A + \delta) \\
P_e &= P_{e|0} \Pr(X = 1) + P_{e|1} \Pr(X = -1) \\
&= \frac{1}{2}(Q_N(A - \delta) + Q_N(A + \delta))
\end{aligned}$$

To minimise  $P_e$ , use differentiate wrt  $\delta$ :

$$\begin{aligned}
\frac{d}{d\delta} \left( \frac{1}{2}(Q_N(A - \delta) + Q_N(A + \delta)) \right) &= 0 \\
\frac{1}{2} \left( -\frac{1}{\sqrt{2\pi}} e^{-\frac{(\delta-A)^2}{2}} - \frac{1}{\sqrt{2\pi}} e^{-\frac{(A+\delta)^2}{2}} \right) &= 0 \\
(\delta - A)^2 &= (A + \delta)^2 \\
\delta &= 0
\end{aligned}$$

so  $\delta = 0$  maximize  $P_e$

5.9 Repeat the above exercise when  $p_X(0) = p$

**Solution:** we have:

$$\begin{aligned}
P_e &= P_{e|0}p + P_{e|1}(1-p) \\
&= pQ_N(A - \delta) + (1-p)Q_N(A + \delta)
\end{aligned}$$

Differentiating wrt  $\delta$  we get

$$p \frac{1}{\sqrt{2\pi}} e^{-\frac{(\delta-A)^2}{2}} - (1-p) \frac{1}{\sqrt{2\pi}} e^{-\frac{(A+\delta)^2}{2}} = 0$$

Taking ln on both sides and find  $\delta$ :

$$\begin{aligned}
\ln p - \frac{(\delta - A)^2}{2} &= \ln(1-p) - \frac{(A + \delta)^2}{2} \\
\delta &= \frac{1}{2A} \ln \frac{1-p}{p}
\end{aligned}$$

if  $p = \frac{1}{2}$  then  $\delta = 0$  which verifies with above result.

5.10 Repeat the above exercise using the MAP criterion.

**Solution:** Assume that  $\Pr(X = -1) = p$ , and  $\Pr(X = 1) = (1 - p)$ . Then, using the Law of

Total Probability, we have:

$$\begin{aligned}
p_Y(y) &= p_{Y|X=-1}(y| -1) \Pr(X = -1) \\
&\quad + p_{Y|X=1}(y|1) \Pr(X = 1) \\
&= p \times p_{(N-A)}(y) \\
&\quad + (1-p) \times p_{(N+A)}(y)
\end{aligned}$$

where  $p_Y(y)$  is the pdf of  $Y$ . Now,  $p_{(N-A)}$  is the pdf of a shifted normal distribution, so

$$p_Y(y) = p \frac{e^{-\frac{(y+A)^2}{2}}}{\sqrt{2\pi}} + (1-p) \frac{e^{-\frac{(y-A)^2}{2}}}{\sqrt{2\pi}}$$

MAP criterion, find  $p_{X|Y}(x|y)$ . we use the Theorem of Conditional Probability:

$$p_{X|Y}(x|y) = \frac{p_{Y|X}(y|x) \times p_X(x)}{p_Y(y)}$$

When  $X = 1$ , we have:

$$\begin{aligned}
p_{X|Y}(1|y) &= \frac{p_{Y|X}(y|1) \times p_X(1)}{p_Y(y)} \\
&= \frac{(1-p) \frac{e^{-\frac{(y-A)^2}{2}}}{\sqrt{2\pi}}}{p \frac{e^{-\frac{(y+A)^2}{2}}}{\sqrt{2\pi}} + (1-p) \frac{e^{-\frac{(y-A)^2}{2}}}{\sqrt{2\pi}}} \\
&= \frac{(1-p) e^{2yA}}{p + (1-p) e^{2yA}}
\end{aligned}$$

Similarly, when  $X = -1$ , we have:

$$\begin{aligned}
p_{X|Y}(-1|y) &= \frac{p_{Y|X}(y|-1) \times p_X(-1)}{p_Y(y)} \\
&= \frac{(p) \frac{e^{-\frac{(y+A)^2}{2}}}{\sqrt{2\pi}}}{p \frac{e^{-\frac{(y+A)^2}{2}}}{\sqrt{2\pi}} + (1-p) \frac{e^{-\frac{(y-A)^2}{2}}}{\sqrt{2\pi}}} \\
&= \frac{p}{p + (1-p) e^{2yA}}
\end{aligned}$$

Therefore, when  $p_{X|Y}(1|y) > p_{X|Y}(-1|y)$ , we have:

$$\begin{aligned}
\frac{(1-p) e^{2yA}}{p + (1-p) e^{2yA}} &> \frac{p}{p + (1-p) e^{2yA}} \\
e^{2yA} &> \frac{p}{(1-p)} \\
y &> \frac{1}{2A} \ln \frac{p}{(1-p)}
\end{aligned}$$

Therefore, we can assert that  $X = 1$ , and  $X = -1$  otherwise. Now, consider when  $p = \frac{1}{2}$ . We



have:

$$\begin{aligned} y &> \frac{1}{2A} \ln \frac{p}{(1-p)} \\ &= \frac{1}{2A} \ln 1 \\ &= 0 \end{aligned}$$

Therefore, when  $y > 0$ , we choose  $X = 1$ , and we choose  $X = -1$  otherwise.

## 6 GAUSSIAN TO OTHER

6.1 Let  $X_1 \sim N(0, 1)$  and  $X_2 \sim N(0, 1)$ . Plot the CDF and PDF of

$$V = X_1^2 + X_2^2 \quad (6.1)$$

**Solution:** The sum of squares of  $n$  independent standard random normal variables is  $\chi^2$  distribution with  $n$  degrees of freedom.

$$P_{\chi^2}(x|n) = \frac{x^{\frac{n}{2}-1} e^{-\frac{x}{2}}}{2^{\frac{n}{2}} \Gamma(\frac{n}{2})}, \forall x \geq 0$$

Here  $k=2$ ,

$$P_{\chi^2}(x|2) = P_V(v) = \frac{e^{-\frac{v}{2}}}{2}$$

For the cumulative distribution

$$\begin{aligned} F_V(v) &= \int_0^v \frac{e^{-\frac{v}{2}}}{2} dv \\ &= 1 - e^{-\frac{v}{2}} \end{aligned}$$

To generate data for  $V$ , run the following code,

```
https://github.com/NareshBandaru13/Random
-Numbers/blob/main/ex6/1/coeffs.h
https://github.com/NareshBandaru13/Random
-Numbers/blob/main/ex6/1/exrand.c
```

Run the below command in terminal,

```
gcc exrand.c -lm
```

The PDF plot of the  $\chi^2(2)$  can be obtained by running the code below,

```
https://github.com/NareshBandaru13/Random
-Numbers/blob/main/ex6/1/main1.py
```

Use the following command in the terminal to run the code

```
python3 main1.py
```

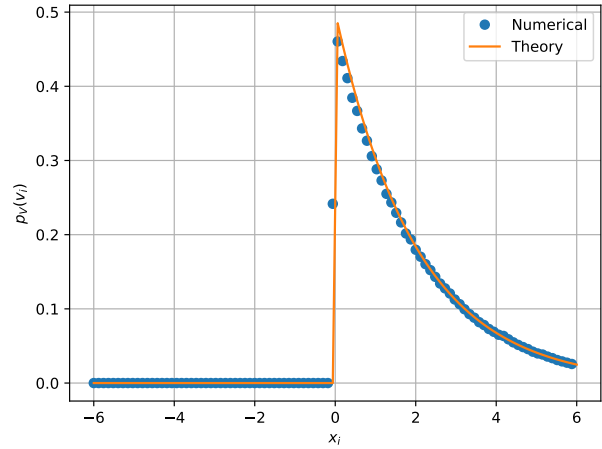


Fig. 6.1: PDF plot

The CDF plot of the  $\chi^2(2)$  can be obtained by running the code below,

```
https://github.com/NareshBandaru13/Random
-Numbers/blob/main/ex6/1/main2.py
```

Use the following command in the terminal to run the code

```
python3 main2.py
```

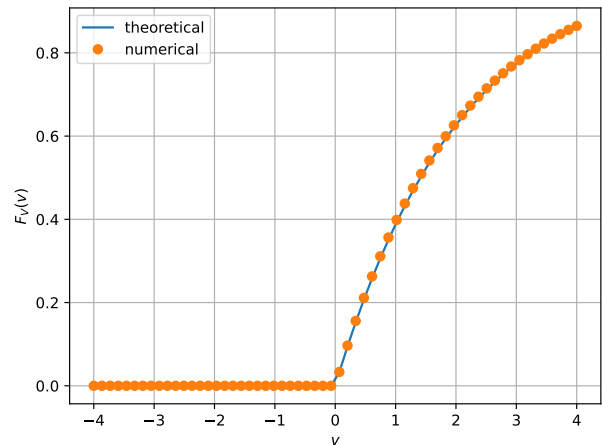


Fig. 6.1: CDF plot

6.2 If

$$F_V(x) = \begin{cases} 1 - e^{-\alpha x} & x \geq 0 \\ 0 & x < 0, \end{cases}$$

find  $\alpha$ .

**Solution:** We will assume that  $X_1$  and  $X_2$  are i.i.d. Let

$$X_1 = r \cos \theta$$

$$X_2 = r \sin \theta$$

The Jacobian Matrix is then defined as:

$$J(r, \theta) = \begin{pmatrix} \frac{\partial x_1}{\partial r} & \frac{\partial x_1}{\partial \theta} \\ \frac{\partial x_2}{\partial r} & \frac{\partial x_2}{\partial \theta} \end{pmatrix}$$

$$J = \begin{pmatrix} \frac{\partial r \cos \theta}{\partial r} & \frac{\partial r \cos \theta}{\partial \theta} \\ \frac{\partial r \sin \theta}{\partial r} & \frac{\partial r \sin \theta}{\partial \theta} \end{pmatrix}$$

$$J = \begin{pmatrix} \cos \theta & -r \sin \theta \\ \sin \theta & r \cos \theta \end{pmatrix}$$

$$|J(r, \theta)| = R$$

Then as  $X_1, X_2$  are independent we have,

$$p_{X_1, X_2}(x_1, x_2) = p_{X_1}(x_1)p_{X_2}(x_2)$$

$$= \frac{1}{2\pi} e^{-\frac{(x_1^2 + x_2^2)}{2}}$$

$$= \frac{1}{2\pi} e^{-\frac{r^2}{2}}$$

Now, since

$$p_{r, \theta}(r, \theta) = |J(r, \theta)| p_{X_1, X_2}(x_1, x_2)$$

we have:

$$p_{R, \theta}(r, \theta) = \frac{r}{2\pi} e^{-\frac{r^2}{2}}$$

$$p_R(r) = \int_0^{2\pi} p_{R, \theta}(r, \theta) d\theta$$

$$= \int_0^{2\pi} \frac{r}{2\pi} e^{-\frac{r^2}{2}} d\theta$$

$$= r e^{-\frac{r^2}{2}}$$

$$F_R(r) = \Pr(R \leq r)$$

$$= \int_0^r f_R(r) dr$$

$$= 1 - e^{-\frac{r^2}{2}}$$

$F_V(x)$  is given by:

$$F_V(x) = F_{X_1^2 + X_2^2}(x)$$

$$= F_{R^2}(x)$$

$$= \Pr(R^2 \leq x)$$

$$= \Pr(R \leq \sqrt{x})$$

Therefore,

$$F_V(x) = \begin{cases} 0, & x < 0 \\ 1 - e^{-\frac{x}{2}}, & x \geq 0 \end{cases}$$

by Comparing we get  $\alpha = \frac{1}{2}$

### 6.3 Plot the CDF and PDF of

$$A = \sqrt{V}$$

**Solution:** To generate data for A , run the following code,

```
https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex6/3/coeffs.h
https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex6/3/exrand.c
```

Run the below command in terminal,

```
gcc exrand.c -lm
```

The PDF plot of A can be obtained by running the code below,

```
https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex6/3/main1.py
```

Use the following command in the terminal to run the code

```
python3 main1.py
```

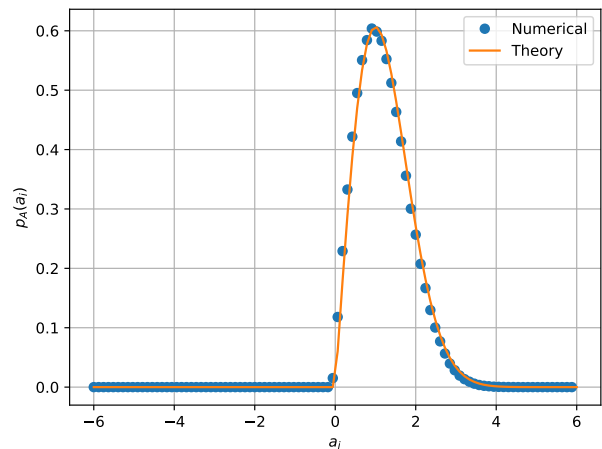


Fig. 6.3: PDF

The CDF plot of the A can be obtained by running the code below,

<https://github.com/NareshBandaru13/Random-Numbers/blob/main/ex6/3/main2.py>

Use the following command in the terminal to run the code

python3 main2.py

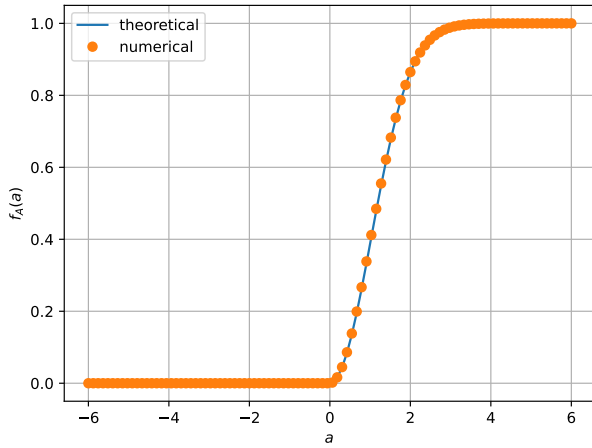


Fig. 6.3: CDF

The CDF of  $A$  for  $x \geq 0$  is given by

$$\begin{aligned}
 F_A(x) &= \Pr(A \leq x) \\
 &= \Pr(\sqrt{V} \leq x) \\
 &= \Pr(V \leq x^2) \\
 &= F_V(x^2) \\
 &= 1 - \exp\left(-\frac{x^2}{2}\right)
 \end{aligned}$$

The PDF of  $A$  is given by

$$\begin{aligned}
 p_A(x) &= \frac{d}{dx} F_A(x) \\
 &= \frac{d}{dx} \left(1 - \exp\left(-\frac{x^2}{2}\right)\right) \\
 &= x \exp\left(-\frac{x^2}{2}\right)
 \end{aligned}$$

Therefore,

$$\begin{aligned}
 F_A(x) &= \begin{cases} 1 - \exp\left(-\frac{x^2}{2}\right) & x \geq 0 \\ 0 & \text{otherwise} \end{cases} \\
 p_A(x) &= \begin{cases} x \exp\left(-\frac{x^2}{2}\right) & x \geq 0 \\ 0 & \text{otherwise} \end{cases}
 \end{aligned}$$