# **BAYESIAN MODEL COMPARISON**

PHD COURSE IN STATISTICAL INFERENCE

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# OVERVIEW OF THE LECTURE

- **Bayesian model comparison**
- Marginal likelihood
- **Log Predictive Score**

## BAYESIAN MODEL COMPARISON

Known parameters, just use the likelihood ratio

$$\frac{p_1(\mathbf{y}|\theta_1)}{p_2(\mathbf{y}|\theta_2)}$$

■ The marginal likelihood for model  $M_k$  with parameters  $\theta_k$ 

$$p_k(\mathbf{y}) = \int p_k(\mathbf{y}|\theta_k) p_k(\theta_k) d\theta_k.$$

- $\blacksquare$   $\theta_k$  is 'removed' by the prior. **Not a silver bullet**. **Priors matter!**
- The Bayes factor

$$B_{12}(y) = \frac{p_1(\mathbf{y})}{p_2(\mathbf{y})}.$$

**■ Posterior model probabilities** 

$$\underbrace{\Pr(M_k|\mathbf{y})}_{\text{posterior model prob.}} \propto \underbrace{p(\mathbf{y}|M_k)}_{\text{marginal likelihood prior model prob.}} \cdot \underbrace{\Pr(M_k)}_{\text{prior model prob.}}$$

# BAYESIAN HYPOTHESIS TESTING - BERNOULLI

■ **Hypothesis testing** is just a special case of model selection:

$$\begin{split} \textit{M}_{O}:&x_{1},...,x_{n} \overset{\textit{iid}}{\sim} \textit{Bernoulli}(\theta_{O}) \\ \textit{M}_{1}:&x_{1},...,x_{n} \overset{\textit{iid}}{\sim} \textit{Bernoulli}(\theta), \theta \sim \textit{Beta}(\alpha,\beta) \\ p(x_{1},...,x_{n}|\textit{M}_{O}) &= \theta_{O}^{s}(1-\theta_{O})^{f}, \\ p(x_{1},...,x_{n}|\textit{M}_{1}) &= \int_{O}^{1} \theta^{s}(1-\theta)^{f}B(\alpha,\beta)^{-1}\theta^{\alpha-1}(1-\theta)^{\beta-1}d\theta \\ &= B(\alpha+s,\beta+f)/B(\alpha,\beta), \end{split}$$

where  $B(\cdot, \cdot)$  is the **Beta function**.

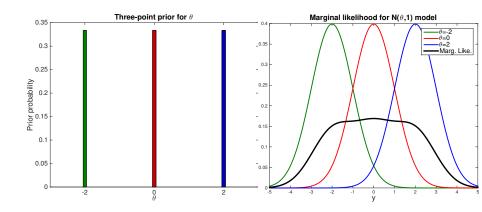
■ Posterior model probabilities

$$Pr(M_k|x_1,...,x_n) \propto p(x_1,...,x_n|M_k)Pr(M_k)$$
, for  $k = 0, 1$ .

■ The Bayes factor

$$BF(M_{O}; M_{1}) = \frac{p(x_{1}, ..., x_{n}|H_{O})}{p(x_{1}, ..., x_{n}|H_{1})} = \frac{\theta_{O}^{s}(1 - \theta_{O})^{f}B(\alpha, \beta)}{B(\alpha + s, \beta + f)}.$$

# PRIORS MATTER



# **EXAMPLE: GEOMETRIC VS POISSON**

- Model 1 **Geometric** with Beta prior:
  - $y_1, ..., y_n | \theta_1 \sim \text{Geo}(\theta_1)$
  - $\theta_1 \sim Beta(\alpha_1, \beta_1)$
- Model 2 Poisson with Gamma prior:
  - $v_1, ..., v_n | \theta_2 \sim Poisson(\theta_2)$
  - $\theta_2 \sim \text{Gamma}(\alpha_2, \beta_2)$
- Marginal likelihood for M<sub>1</sub>

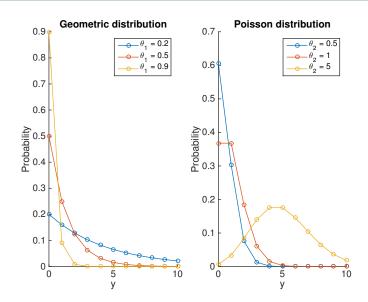
$$p_{1}(y_{1},...,y_{n}) = \int p_{1}(y_{1},...,y_{n}|\theta_{1})p(\theta_{1})d\theta_{1}$$

$$= \frac{\Gamma(\alpha_{1}+\beta_{1})}{\Gamma(\alpha_{1})\Gamma(\beta_{1})} \frac{\Gamma(n+\alpha_{1})\Gamma(n\bar{y}+\beta_{1})}{\Gamma(n+n\bar{y}+\alpha_{1}+\beta_{1})}$$

■ Marginal likelihood for M<sub>2</sub>

$$p_2(y_1, ..., y_n) = \frac{\Gamma(n\bar{y} + \alpha_2)\beta_2^{\alpha_2}}{\Gamma(\alpha_2)(n + \beta_2)^{n\bar{y} + \alpha_2}} \frac{1}{\prod_{i=1}^n y_i!}$$

# GEOMETRIC AND POISSON



# GEOMETRIC VS POISSON, CONT.

■ Priors match prior predictive means:

$$E(y_i|M_1) = E(y_i|M_2) \iff \alpha_1\alpha_2 = \beta_1\beta_2$$

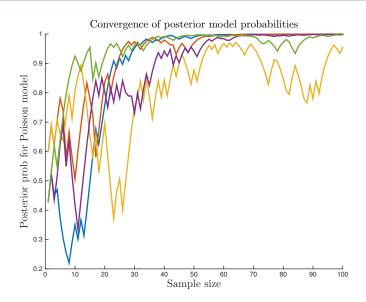
**Data**:  $y_1 = 0$ ,  $y_2 = 0$ .

	$lpha_1=$ 1, $eta_1=$ 2	$lpha_1=$ 10, $eta_1=$ 20	$\alpha_1 = 100, \beta_1 = 200$
	$lpha_2=$ 2, $eta_2=$ 1	$lpha_2=$ 20, $eta_2=$ 10	$\alpha_2 = 200, \beta_2 = 100$
BF <sub>12</sub>	1.5	4.54	5.87
$\Pr(M_1 \mathbf{y})$	0.6	0.82	0.85
$\Pr(M_2 \mathbf{y})$	0.4	0.18	0.15

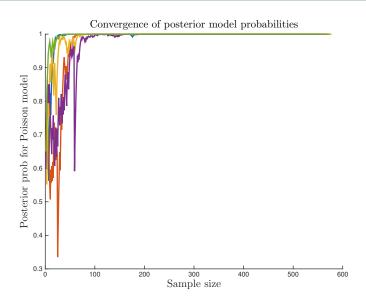
**Data**:  $y_1 = 3$ ,  $y_2 = 3$ .

	$lpha_1=$ 1, $eta_1=$ 2	$lpha_1=$ 10, $eta_1=$ 20	$\alpha_1 = 100, \beta_1 = 200$
	$lpha_2=$ 2, $eta_2=$ 1	$lpha_2=$ 20, $eta_2=$ 10	$lpha_2=$ 200, $eta_2=$ 100
$BF_{12}$	0.26	0.29	0.30
$\Pr(M_1 \mathbf{y})$	0.21	0.22	0.23
$\Pr(M_2 \boldsymbol{y})$	0.79	0.78	0.77

# GEOMETRIC VS POISSON FOR POIS(1) DATA



# GEOMETRIC VS POISSON FOR POIS(1) DATA



# MODEL CHOICE IN MULTIVARIATE TIME SERIES

### **■** Multivariate time series

$$\mathbf{x}_t = \alpha \beta' \mathbf{z}_t + \Phi_1 \mathbf{x}_{t-1} + \dots \Phi_k \mathbf{x}_{t-k} + \Psi_1 + \Psi_2 t + \Psi_3 t^2 + \varepsilon_t$$

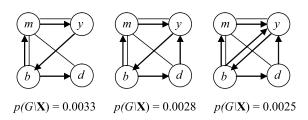
#### Need to choose:

- Lag length, (k = 1, 2..., 4)
- Trend model (s = 1, 2, ..., 5)
- Long-run (cointegration) relations (r = 0, 1, 2, 3, 4).

The most prob	ABLE	(k, r, s	) сом	BINATI	ONS IN	THE	Danisi	H MON	ETARY	DATA.
k	1	1	1	1	1	1	1	1	0	1
r	3	3	2	4	2	1	2	3	4	3
s	3	2	2	2	3	3	4	4	4	5
p(k, r, s y, x, z)	.106	.093	.091	.060	.059	.055	.054	.049	.040	.038

# GRAPHICAL MODELS FOR MULTIVARIATE TIME SERIES

- **Graphical models** for multivariate time series.
- Zero-restrictions on the effect from time series *i* on time series *j*, for all lags. (**Granger Causality**).
- Zero-restrictions on the elements of the inverse covariance matrix of the errors.



# PROPERTIES OF BAYESIAN MODEL COMPARISON

■ Coherence of pair-wise comparisons

$$B_{12} = B_{13} \cdot B_{32}$$

**Consistency** when true model is in  $\mathcal{M} = \{M_1, ..., M_K\}$ 

$$\Pr\left(M = M_{TRUE}|\mathbf{y}\right) \to 1 \text{ as } n \to \infty$$

■ "KL-consistency" when  $M_{TRIJF} \notin \mathcal{M}$ 

$$\Pr\left(M = M^* | \mathbf{y}\right) \to 1 \text{ as } n \to \infty,$$

 $M^*$  minimizes **KL divergence** between  $p_M(\mathbf{y})$  and  $p_{TRUF}(\mathbf{y})$ .

- Smaller models always win when priors are very vague.
- Improper priors cannot be used for model comparison.

# MARGINAL LIKELIHOOD MEASURES OUT-OF-SAMPLE PREDICTIVE PERFORMANCE

■ The marginal likelihood can be decomposed as

$$p(y_1,...,y_n) = p(y_1)p(y_2|y_1)\cdots p(y_n|y_1,y_2,...,y_{n-1})$$

■ Assume that  $y_i$  is independent of  $y_1, ..., y_{i-1}$  conditional on  $\theta$ :

$$p(y_i|y_1,...,y_{i-1}) = \int p(y_i|\theta)p(\theta|y_1,...,y_{i-1})d\theta$$

- **Prediction of**  $y_1$  is based on the prior of  $\theta$ . Sensitive to prior.
- Prediction of  $y_n$  uses almost all the data to infer  $\theta$ . Not sensitive to prior when n is not small.

# NORMAL EXAMPLE

- Model:  $y_1, ..., y_n | \theta \sim N(\theta, \sigma^2)$  with  $\sigma^2$  known.
- Prior:  $\theta \sim N(0, \kappa^2 \sigma^2)$ .
- Intermediate posterior at time i-1

$$\theta|y_1,...,y_{i-1} \sim N\left[w_i(\kappa) \cdot \bar{y}_{i-1}, \frac{\sigma^2}{i-1+\kappa^{-2}}\right]$$

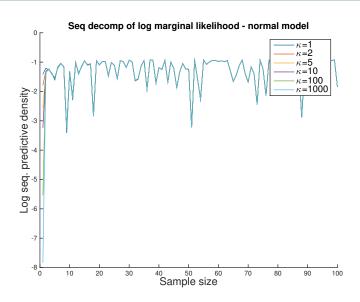
where  $w_i(\kappa) = \frac{i-1}{i-1+\kappa^{-2}}$ .

■ Intermediate predictive density at time i-1

$$y_i|y_1,...,y_{i-1} \sim N\left[w_i(\kappa) \cdot \bar{y}_{i-1}, \sigma^2\left(1 + \frac{1}{i-1+\kappa^{-2}}\right)\right]$$

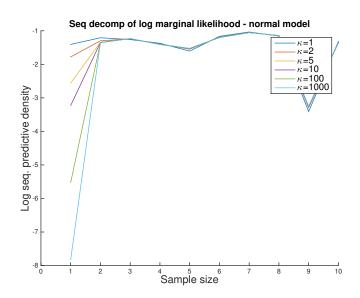
- For i=1,  $y_1 \sim N\left[0, \sigma^2\left(1+\frac{1}{\kappa^{-2}}\right)\right]$  can be very sensitive to  $\kappa$ .
- For large  $i: y_i | y_1, ..., y_{i-1} \stackrel{approx}{\sim} N(\bar{y}_{i-1}, \sigma^2)$ , not sensitive to  $\kappa$ .

# FIRST OBSERVATION IS SENSITIVE TO $\kappa$



15 | 1

# FIRST OBSERVATION IS SENSITIVE TO $\kappa$ - ZOOMED



## LOG PREDICTIVE SCORE - LPS

- Reduce sensitivity to the prior: sacrifice *n*\* observations to train the prior into a posterior.
- **Predictive (Density) Score (PS).** Decompose  $p(y_1, ..., y_n)$  as

$$\underbrace{\frac{p(y_1)p(y_2|y_1)\cdots p(y_{n^*}|y_{1:(n^*-1)})}{training}}_{training}\underbrace{\frac{p(y_{n^*+1}|y_{1:n^*})\cdots p(y_n|y_{1:(n-1)})}{test}}_{test}$$

- Usually report on log scale: Log Predictive Score (LPS).
- Time-series: obvious which data are used for training.
- Cross-sectional data: training-test split by **cross-validation**:

Fold 1	Fold 2	Fold 3	Fold 4	Fold 5
Fold 1	Fold 2	Fold 3	Fold 4	Fold 5
Fold 1	Fold 2	Fold 3	Fold 4	Fold 5
Fold 1	Fold 2	Fold 3	Fold 4	Fold 5
Fold 1	Fold 2	Fold 3	Fold 4	Fold 5

# AND HEY! ... LET'S BE CAREFUL OUT THERE

- Be especially **careful** with Bayesian model comparison when
  - The compared models are
    - very different in structure
    - severly misspecified
    - very complicated (black boxes).
  - The priors for the parameters in the models are
    - not carefully elicited
    - only weakly informative
    - not matched across models.
  - The data
    - has outliers (in all models)
    - has a multivariate response.

# MODEL AVERAGING

- Let  $\gamma$  be a quantity with an interpretation which stays the same across the two models.
- Example: Prediction  $\gamma = (y_{T+1}, ..., y_{T+h})'$ .
- $\blacksquare$  The marginal posterior distribution of  $\gamma$  reads

$$p(\gamma|\mathbf{y}) = p(M_1|\mathbf{y})p_1(\gamma|\mathbf{y}) + p(M_2|\mathbf{y})p_2(\gamma|\mathbf{y}),$$

where  $p_k(\gamma|\mathbf{y})$  is the marginal posterior of  $\gamma$  conditional on model k.

- Predictive distribution includes three sources of uncertainty:
  - **Future errors**/disturbances (e.g. the  $\varepsilon$ 's in a regression)
  - Parameter uncertainty (the predictive distribution has the parameters integrated out by their posteriors)
  - Model uncertainty (by model averaging)