

## MATH 574 Homework 6

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**Collaboration:** I discussed some of the problems with Jackson Ginn, Sam Maloney, and Jack Hyatt.

**Problem 1** In this problem, we will prove a one-sided Chebyshev-type bound in several steps. The goal is to fill in the details of the proof of the theorem.

- (a) Let  $X$  be a random variable and let  $c \in \mathbb{R}$ . Prove that  $V(X) = V(X + c)$ .
- (b) Prove that if  $Y$  is a random variable with  $E(Y) = 0$ , then for any constant  $c \in \mathbb{R}$ ,  $E((Y - c)^2) = V(Y) + c^2$ .
- (c) Prove that for any random variable  $X$ ,  $E(X - E(X)) = 0$ .
- (d) Prove that for any  $c \geq 0$  and random variable  $X$ ,  $p(X \geq c) \leq p(X^2 \geq c^2)$ .
- (e) Now prove the following theorem. A brief outline is given below, but you should write a full proof. You may cite the results proven above.

**Theorem 1.** Let  $X$  be a random variable with variance  $\sigma^2$ . Then for any  $k > 0$ ,

$$p(X - E(X) \geq k) \leq \frac{\sigma^2}{\sigma^2 + k^2}.$$

*Proof outline.* Set  $Y = X - E(X)$ . Then argue that

$$p(X - E(X) \geq k) = p(Y \geq k) \leq p((Y + x)^2 \geq (k + x)^2),$$

for any  $x \geq 0$ .

Use Markov's inequality (see Homework 5, #1) to obtain that

$$p((Y + x)^2 \geq (k + x)^2) \leq \frac{E((Y + x)^2)}{(k + x)^2}.$$

Conclude that

$$\frac{E((Y + x)^2)}{(k + x)^2} = \frac{\sigma^2 + x^2}{(k + x)^2}.$$

Therefore  $p(X - E(X) \geq k) \leq \frac{\sigma^2 + x^2}{(k + x)^2}$ , but this holds for *any*  $x$ . Hence this theorem is most useful when we minimize the function  $\frac{\sigma^2 + x^2}{(k + x)^2}$ .

Optimize (i.e., find a minimum of) the function  $\frac{\sigma^2 + x^2}{(k + x)^2}$  to get a bound matching the theorem. □

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Solution.

(a) We have that

$$\begin{aligned} V(X + c) &= E((X + c)^2) - [E(X + c)]^2 && \text{(definition)} \\ &= E(X^2 + 2Xc + c^2) - [E(X + c)]^2 \\ &= E(X^2) + E(2Xc) + E(c^2) - [E(X) + E(c)]^2 && \text{(linearity of expectation)} \end{aligned}$$

$$\begin{aligned}
&= E(X^2) + E(2Xc) + E(c^2) - (E(X)^2 + 2E(X)E(c) + E(c)^2) \\
&= E(X^2) + 2cE(X) + c^2 - (E(X)^2 + 2cE(X) + c^2) && \text{(linearity of expectation)} \\
&= E(X^2) - E(X)^2 && \text{(combining like terms)} \\
&= V(X). && \text{(definition)}
\end{aligned}$$

(b) We have that

$$\begin{aligned}
E((Y - c)^2) &= E(Y^2 - 2Yc + c^2) \\
&= E(Y^2) - 2cE(Y) + c^2 && \text{(linearity of expectation)} \\
&= E(Y^2) - 2c(0) + c^2 && (E(Y) = 0) \\
&= E(Y^2) - [E(Y)]^2 + c^2 && (0 = 0^2) \\
&= V(Y) + c^2. && (V(Y) = E(Y^2) - [E(Y)]^2)
\end{aligned}$$

It also follows that

$$\begin{aligned}
E((Y + c)^2) &= E((Y - (-c))^2) \\
&= V(Y) + (-c)^2 \\
&= V(Y) + c^2.
\end{aligned}$$

(c) We have that

$$\begin{aligned}
E(X - E(X)) &= E(X) - E(E(X)) && \text{(linearity of expectation)} \\
&= E(X) - E(X) = 0.
\end{aligned}$$

(d) Because the event  $X^2 \geq c^2 = X \leq -c \cup X \geq c$ , and the events  $X \leq -c$  and  $X \geq c$  are disjoint, we have

$$P(X \leq -c) + P(X \geq c) = P(X^2 \geq c^2).$$

Since  $P(X \leq -c)$  is nonnegative, we can conclude that

$$P(X \geq c) \leq P(X^2 \geq c^2).$$

(e) Let  $X$  be a random variable with variance  $\sigma^2$ . Then for any  $k > 0$ , we claim that

$$P(X - E(X) \geq k) \leq \frac{\sigma^2}{\sigma^2 + k^2}.$$

Let  $Y = X - E(X)$ . Then, for any  $x > 0$  we have

$$\begin{aligned}
P(X - E(X) \geq k) &= P(Y \geq k) && \text{(by definition)} \\
&= P(Y + x \geq k + x) && \text{(algebra)} \\
&\leq P((Y + x)^2 \geq (k + x)^2). && \text{(as shown in (d))}
\end{aligned}$$

We have from Markov's inequality that for a random variable  $W$  on a sample space  $S$  where  $W(s) > 0$  for every  $s \in S$ ,  $P(W \geq a) \leq \frac{E(W)}{a}$  for all  $a \in \mathbb{R}^+$ . Thus, if we choose  $W = (Y + x)^2$  and  $a = (k + x)^2$ , Markov's inequality applies because every  $w \in (Y + x)^2(S)$  is positive and  $a$  is non-negative (as  $x > 0$  and both results are from squares). So we continue the inequality chain with

$$P((Y + x)^2 \geq (k + x)^2) \leq \frac{E((Y + x)^2)}{(k + x)^2}.$$

We now re-introduce the variable  $X$ . We have

$$\begin{aligned}
 Y &= X - E(X) \\
 \implies E(Y) &= E(X - E(X)) \\
 \implies E(Y) &= 0 && \text{(as shown in part (c))} \\
 \implies E((Y + x)^2) &= V(Y) + x^2. && \text{(as shown in part (b))}
 \end{aligned}$$

Since  $V(Y) = V(X - E(X))$ ,  $V(Y) = V(X) = \sigma^2$  as shown in part (a). So

$$\frac{E(Y + x)^2}{(k + x)^2} = \frac{\sigma^2 + x^2}{(k + x)^2}.$$

Therefore, from our chain of inequalities, we have that  $p(X - E(X) \geq k) \leq \frac{\sigma^2 + x^2}{(k + x)^2}$ , but this holds for any  $x$ . Hence this theorem is most useful when we minimize the function  $f : \mathbb{R} \rightarrow \mathbb{R}$ ,  $f(x) = \frac{\sigma^2 + x^2}{(k + x)^2}$ . We use the first derivative test and solve for  $x$ :

$$\begin{aligned}
 \frac{d}{dx} \left[ \frac{\sigma^2 + x^2}{(k + x)^2} \right] &= 0 \\
 \implies \frac{(k + x)^2(2x) - 2(\sigma^2 + x^2)(k + x)}{(k + x)^4} &= 0 \\
 \implies \frac{(k + x)(2x) - 2(\sigma^2 + x^2)}{(k + x)^3} &= 0 && \text{(provided } x \neq -k) \\
 \implies \frac{2(xk - \sigma^2)}{(x + k)^3} &= 0 \\
 \implies xk - \sigma^2 &= 0 \\
 \implies x &= \frac{\sigma^2}{k}.
 \end{aligned}$$

Since  $k > 0$ ,  $\frac{\sigma^2}{k}$  will always be positive. Thus,  $x = 0 < \frac{\sigma^2}{k}$ , so if  $x = \frac{\sigma^2}{k}$  is our only minimum,  $f'(x)$  will be negative for  $x$ -values less than  $\frac{\sigma^2}{k}$  provided they are greater than the discontinuity at  $x = -k$ . Since  $k > 0$ ,  $x = 0$  satisfies this since  $-k < 0 < \frac{\sigma^2}{k}$ . We have  $f'(x) = \frac{2(xk - \sigma^2)}{(x + k)^3}$ , so

$$f'(0) = -\frac{2\sigma^2}{k^3} < 0.$$

We repeat the same process with  $x = \frac{2\sigma^2}{k}$ , a value guaranteed to be greater than  $x = \frac{\sigma^2}{k}$ . After some simplification, we obtain

$$f' \left( \frac{2\sigma^2}{k} \right) = \frac{2k^3\sigma^2}{(k^2 + 2\sigma^2)^3} > 0.$$

Thus,  $x = \frac{\sigma^2}{k}$  is an absolute minimum, because  $f'(x)$  is negative for values less than  $\frac{\sigma^2}{k}$  and positive for values that are greater. To conclude, we evaluate

$$\begin{aligned}
 f \left( \frac{\sigma^2}{k} \right) &= \frac{\sigma^2 + \left( \frac{\sigma^2}{k} \right)^2}{\left( k + \frac{\sigma^2}{k} \right)^2} \\
 &= \frac{(\sigma^2 k^2 + \sigma^4)/k^2}{(k^2 + \sigma^2)^2/k^2} \\
 &= \frac{\sigma^2(k^2 + \sigma^2)}{(k^2 + \sigma^2)^2}
 \end{aligned}$$

$$= \frac{\sigma^2}{k^2 + \sigma^2}.$$

Therefore, the minimum value of  $f(x)$  is  $\frac{\sigma^2}{k^2 + \sigma^2}$ , and we have for all  $k > 0$  that

$$p(X - E(X) \geq k) \leq \frac{\sigma^2}{\sigma^2 + k^2}.$$

□

**Problem 2** A biased coin has probability for heads  $p = 0.75$ . Suppose we flip the coin 1,000 times. Give an upper bound for the probability that we flip at least 800 heads

- (a) using Markov's inequality.
- (b) using Chebyshev's inequality.
- (c) using Theorem 1 in the previous question.

Solution.

Let  $X : S \rightarrow \mathbb{R}$  represent the number of heads after flipping the coin 1,000 times. Then,  $X$  follows a binomial distribution with  $n = 1000$  and  $p = \frac{3}{4}$ , and thus  $E(X) = np = 750$  and  $V(X) = np(1 - p) = \frac{375}{2}$ .

(a) We have from Markov's inequality that for a random variable  $W$  on a sample space  $S$  where  $W(s) > 0$  for every  $s \in S$ ,  $P(W \geq a) \leq \frac{E(W)}{a}$  for all  $a \in \mathbb{R}^+$ . So if we choose  $W = X$  and  $a = 800$ , we have

$$P(X \geq 800) \leq \frac{E(X)}{800} = \frac{15}{16}.$$

(b) We have from Chebyshev's Inequality that for a random variable  $W$  on a sample space  $S$ ,

$$P(|W - E(W)| \geq k) \leq \frac{V(W)}{k^2}$$

for all  $k \in \mathbb{R}^+$ . So if we choose  $W = X$  and  $k = 50$ , we have

$$P(|X - E(X)| \geq 50) \leq \frac{V(X)}{2500}.$$

Substituting, we can write  $P(X \leq 700 \cup X \geq 800) \leq \frac{3}{40}$ , so we must have

$$P(X \geq 800) \leq \frac{3}{40}.$$

(c) We have from Theorem 1 that for a random variable  $W$ ,  $P(W - E(W) \geq k) \leq \frac{V(W)}{V(W) + k^2}$ . So if we choose  $W = X$  and  $k = 50$ , we have

$$P(X - E(X) \geq 50) \leq \frac{V(X)}{V(X) + 50^2}.$$

Substituting, it follows that

$$P(X \geq 800) \leq \frac{375/2}{375/2 + 2500} = \frac{3}{43}.$$

**Problem 3** Let  $a_n$  be the number of strings of length  $n$  with digits in  $0-9$  that contain 2 or more consecutive 0s.

- (a) Find a recurrence relation for  $a_n$ .
- (b) What are the initial conditions?
- (c) Determine  $a_7$ .

Solution.

Let  $S_n$  be the set of strings of length  $n$  with digits in  $0 - 9$  that contain 2 or more consecutive 0s (so  $|S_n| = a_n$ ). Then, let  $S'_n$  be the number of strings of length  $n$  with digits in  $0 - 9$  that do not contain 2 or more consecutive 0s, and let  $a'_n = |S'_n|$ . By these definitions, any string of length  $n$  with digits in  $0 - 9$  will either be in  $S_n$  or  $S'_n$ . By the product rule, there are  $10^n$  such possible strings, so  $a_n + a'_n = 10^n$  for all  $n > 0$ .

(a) Let  $n \in \mathbb{N}$  such that  $n \geq 3$ . Then, we consider the strings  $\alpha_1\alpha_2\ldots\alpha_n \in S'_n$ . There are two possibilities:

Case 1:  $\alpha_n \neq 0$ . Then, there are  $a'_{n-1}$  possibilities for  $\alpha_1\alpha_2\ldots\alpha_{n-1}$  because adding a nonzero digit to the end places no additional restrictions on the first  $n - 1$  digits. Since there are 9 possibilities for  $\alpha_n$ , by the product rule there are  $9a'_{n-1}$  such possible strings.

Case 2:  $\alpha_n = 0$ . Then,  $\alpha_{n-1} \neq 0$  because if it were, the last two digits would constitute consecutive 0s. So there are  $a'_{n-2}$  possibilities for  $\alpha_1\alpha_2\ldots\alpha_{n-2}$ , 9 possibilities for  $\alpha_{n-1}$ , and 1 possibility for  $\alpha_n$ . By the product rule, there are  $9a'_{n-2}$  such possible strings.

So  $a'_n = 9a'_{n-1} + 9a'_{n-2}$ . Since we have  $a'_n = 10^n - a_n$ , we have  $10^n - a_n = 9(10^{n-1} - a_{n-1}) + 9(10^{n-2} - a_{n-2})$ . Therefore,

$$a_n = 10^n - 9(10^{n-1} - a_{n-1} + 10^{n-2} - a_{n-2}).$$

(b) For  $n = 1$ , there are (rather obviously) no strings of length 1 with 2 or more consecutive 0s. For  $n = 2$ , there is one string of length 2 with 2 or more consecutive 0s, namely 00. So  $a_1 = 0$  and  $a_2 = 1$ .

(c) We calculate each value up to  $a_7$ :

- $a_3 = 10^3 - 9(10^{3-1} - a_{3-1} + 10^{3-2} - a_{3-2}) = 10^3 - 9(10^2 - 1 + 10^1 - 0) = 19$
- $a_4 = 10^4 - 9(10^{4-1} - a_{4-1} + 10^{4-2} - a_{4-2}) = 10^4 - 9(10^3 - 19 + 10^2 - 1) = 280$
- $a_5 = 10^5 - 9(10^{5-1} - a_{5-1} + 10^{5-2} - a_{5-2}) = 10^5 - 9(10^4 - 280 + 10^3 - 19) = 3691$
- $a_6 = 10^6 - 9(10^{6-1} - a_{6-1} + 10^{6-2} - a_{6-2}) = 10^6 - 9(10^5 - 3691 + 10^4 - 280) = 45739$
- $a_7 = 10^7 - 9(10^{7-1} - a_{7-1} + 10^{7-2} - a_{7-2}) = 10^7 - 9(10^6 - 45739 + 10^5 - 3691) = 544870$

**Problem 4** Let  $b_n$  be the number of strings of length  $n$  with digits in  $0 - 9$  that contain no two repeated numbers in a row.

(a) Determine  $b_1$ .

(b) Prove that  $b_n$  has the recurrence relation  $b_n = 9b_{n-1}$  for all  $n \geq 2$ .

(c) Solve the recurrence relation with the initial condition in part (a).

Solution.

Let  $S_n$  be the set of strings of length  $n$  with digits in  $0 - 9$  that contain no two repeated numbers in a row (so  $|S_n| = b_n$ ).

(a) There are clearly no strings of length 1 that have 2 repeated numbers in a row, so any string of length 1 will work. There are 10 choices for the digit, so  $b_1 = 10$ .

(b) We consider the strings  $\beta_1\beta_2\ldots\beta_n \in S_n$ . There are  $b_{n-1}$  choices for  $\beta_1\beta_2\ldots\beta_{n-1}$ , since it is a string of length  $n - 1$  with no two repeated numbers in a row, and 9 choices for  $\beta_n$  since it can equal any digit

other than  $\beta_{n-1}$  (if it did, they would constitute two repeated numbers in a row). So by the product rule,  $b_n = 9b_{n-1}$ .

(c) We claim that  $b_n = 10(9)^{n-1}$  for  $n \in \mathbb{N}$ .

First, let  $n = 1$ . Then,

$$b_1 = 10 = 10(1) = 10(9)^0 = 10(9)^{1-1},$$

so the claim holds for  $n = 1$ .

Next, let  $n \in \mathbb{N}$ ,  $n \geq 1$ . Assume that  $b_n = 10(9)^{n-1}$ . Then, we have

$$\begin{aligned} b_{n+1} &= 9b_n && \text{(recurrence relation)} \\ &= 9(10(9)^{n-1}) && \text{(inductive hypothesis)} \\ &= 10(9)^n && \text{(rewriting product)} \\ &= 10(9)^{(n+1)-1}. \end{aligned}$$

So if the claim holds for  $n$ , it also holds for  $n + 1$ . Therefore, we have that  $b_n = 10(9)^{n-1}$  for all  $n \in \mathbb{N}$ .