

# USING ARIMA MODELS TO MAKE CAUSAL STATEMENTS

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# OUTLINE

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## 2 C-ARIMA APPROACH

- Causal framework
- Inference

## 3 PRACTICAL SESSION

# EMPIRICAL PROBLEM

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**FIGURE:** Store brands (first row) and direct competitor brands (second row).



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Not so easy! Inference of a proper “causal” effect is complicated by several issues, including:

- Autocorrelation
- Seasonality
- Holiday effects
- Interference



Infer the effect attributable to the policy netting out the portion due to other factors e.g., increasing trends started in the past Christmas rush purchases, Sunday effect



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C-ARIMA shares many features with “CausalImpact” [1], but it is based on ARIMA models — frequentist alternative to BSTS models, easy to implement, used for a wide class of processes.

# ASSUMPTIONS

Let  $W_{i,t} \in \{0, 1\}$  be a random variable describing the treatment assignment of unit  $i \in \{1, \dots, N\}$  at time  $t \in \{1, \dots, T\}$ , where 1 denotes that a “treatment” (or “intervention”) has taken place and 0 denotes control. We maintain the following assumptions:

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(A1: SINGLE PERSISTENT INTERVENTION)

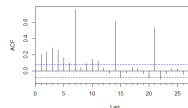
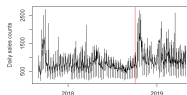
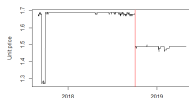
$\exists t^* \in \{1, \dots, T\}$  s.t.  $w_{i,t} = 0 \ \forall t \leq t^*$  and  $\forall t > t^*, w_{i,t} \in \{(1, \dots, 1), (0, \dots, 0)\}$

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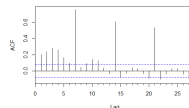
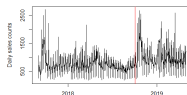
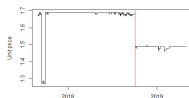


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(A2: TEMPORAL NO-INTERFERENCE)

For all  $i \in \{1, \dots, N\}$ ,  $Y_{i,t}(w_{1:N,t^*+1:T}) = Y_{i,t}(w_{i,t^*+1:T})$

# ASSUMPTIONS

## (A3: COVARIATES-TREATMENT INDEPENDENCE)

$$X_{i,t}(w_{i,t^*+1:T}) = X_{i,t}(w'_{i,t^*+1:T}).$$

## (A4: NON-ANTICIPATING INDIVIDUALISTIC TREATMENT)

$$\begin{aligned} \Pr(W_{1:N,t^*+1} = w_{1:N,t^*+1} \mid W_{1:N,1:t^*}, Y_{1:N,1:T}(w_{1:N,1:T}), X_{1:N,1:T}) = \\ = \prod_{i=1}^N \Pr(W_{i,t^*+1} = w_{i,t^*+1} \mid Y_{i,1:t^*}(w_{i,1:t^*}), X_{i,1:t^*}). \end{aligned}$$

Above assumptions are essential to define, estimate and attribute the causal effect to the intervention. Moreover, they allow us to ease notation: under Assumption 1, for all  $t > t^*$  we can write  $w_{i,t} = w_i$  and if Assumption 2 holds we can also drop the  $i$  subscript. From now on, we can use  $Y_t(w)$  to denote the potential outcome of a generic unit at time  $t > t^*$ .



# CAUSAL ESTIMANDS

## DEFINITION

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The temporal average causal effect is

$$\bar{\tau}_t(1; 0) = \frac{1}{t - t^*} \sum_{s=t^*+1}^t \tau_s(1; 0) = \frac{\Delta_t(1; 0)}{t - t^*}. \quad (3)$$

# C-ARIMA

let us assume  $\{Y_t(w)\}$  evolving as

$$Y_t(w) = \frac{\theta_q(L)}{\phi_p(L)} \varepsilon_t + \tau_t \mathbb{1}_{\{w=1\}} \quad (4)$$

where,

- $\phi_p(L)$  and  $\theta_q(L)$  are lag polynomials having roots all outside the unit circle
- given this representation, the point causal effect at time  $t > t^*$  is  
 $\tau_t \equiv Y_t(w=1) - Y_t(w=0)$
- $\tau_t = 0 \ \forall t \leq t^*$  and  $\mathbb{1}_{\{w=1\}}$  is an indicator function which is one if  $w = 1$
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let us assume  $\{Y_t(w)\}$  evolving as

$$(1 - L^s)^D (1 - L)^d Y_t(w) = \frac{\Theta_Q(L^s) \theta_q(L)}{\Phi_P(L^s) \phi_p(L)} \varepsilon_t + (1 - L^s)^D (1 - L)^d X_t' \beta + \tau_t \mathbb{1}_{\{w=1\}} \quad (4)$$

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- $\varepsilon_t$  is white noise with mean 0 and variance  $\sigma_\varepsilon^2$
- $\Theta_Q(L^s)$ ,  $\Phi_P(L^s)$  are the lag polynomials of the seasonal part of the model having roots all outside the unit circle
- $(1 - L^s)^D$  and  $(1 - L)^d$  are the differencing operators to ensure stationarity

After some manipulation, Equation (4) becomes

$$S_t = z_t + \tau_t \mathbb{1}_{\{w=1\}}$$

where,  $S_t = T(Y_t) - T(X_t)' \beta$  and  $T(\cdot)$  is the transformation of  $Y_t$  needed to achieve stationarity, i.e.  $T(Y_t) = (1 - L^s)^D (1 - L)^d Y_t$ ;  $z_t$  includes the stationary part of the model, namely,

$$z_t = \frac{\Theta_Q(L^s) \theta_q(L)}{\Phi_P(L^s) \phi_p(L)} \varepsilon_t$$

If  $w = 0$ , absence of intervention, the  $k$ -step ahead forecast of  $S_t$  conditionally on the information up to time  $t^*$  is

$$\hat{S}_{t^*+k}(0) = E[S_{t^*+k}(0) | \mathcal{I}_{t^*}] = E[z_{t^*+k} | \mathcal{I}_{t^*}] = \hat{z}_{t^*+k|t^*}$$

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Point causal effect estimator!

# INFERENCE

## THEOREM (PART 1)

Let  $\{Y_t\}$  follow the regression model with ARIMA errors defined in Equation (4) and, for any  $k > 0$  let  $H_0 : \tau_{t^*+k}(1; 0) = 0$  the null hypothesis that the intervention has no effect. Then, the estimators of the point, cumulative and temporal average effects under  $H_0$  can be expressed as,

$$\hat{\tau}_{t^*+k}(1; 0)|H_0 = \sum_{i=0}^{k-1} \psi_i \varepsilon_{t^*+k-i} \quad (5)$$

$$\hat{\Delta}_{t^*+k}(1; 0)|H_0 = \sum_{h=1}^k \varepsilon_{t^*+h} \sum_{i=0}^{k-h} \psi_i \quad (6)$$

$$\hat{\bar{\tau}}_{t^*+k}(1; 0)|H_0 = \frac{1}{k} \sum_{h=1}^k \varepsilon_{t^*+h} \sum_{i=0}^{k-h} \psi_i, \quad (7)$$

where, the  $\psi_i$ 's are the coefficients of a moving average of order  $k - 1$  whose values are functions of the ARMA parameters in Equation (4).

# INFERENCE

## THEOREM (PART 2)

*In case the  $\varepsilon_t$  error term is assumed normally distributed, Equations (5)–(7) become*

$$\hat{\tau}_{t^*+k}(1;0)|H_0 \sim N \left[ 0, \sigma_\varepsilon^2 \sum_{i=0}^{k-1} \psi_i^2 \right] \quad (8)$$

$$\hat{\Delta}_{t^*+k}(1;0)|H_0 \sim N \left[ 0, \sigma_\varepsilon^2 \sum_{h=1}^k \left( \sum_{i=0}^{k-h} \psi_i \right)^2 \right] \quad (9)$$

$$\hat{\tau}_{t^*+k}(1;0)|H_0 \sim N \left[ 0, \frac{1}{k^2} \sigma_\varepsilon^2 \sum_{h=1}^k \left( \sum_{i=0}^{k-h} \psi_i \right)^2 \right] \quad (10)$$

If one relies on the normality of the error term, inference can be based on Equations (8)–(10); otherwise, one can compute empirical critical values from Equations (5)–(7) by bootstrapping the errors from the model residuals.

# INFERENCE

Summarizing, to estimate the effect of an intervention with C-ARIMA we need to follow these steps:

- 1 estimate the ARIMA model only in the pre-intervention period, so as to learn the dynamics of the dependent variable and the links with the covariates without being influenced by the treatment
- 2 based on the process learned in the pre-intervention period, perform a prediction step and obtain an estimate of the counterfactual outcome during the post-intervention period
- 3 by comparing the observations with the corresponding forecasts at any time point after the intervention, evaluate the resulting differences, which represent the estimated point causal effects

# CONCLUDING REMARKS

- The proposed C-ARIMA approach can be used for policy evaluation (e.g., new regulation, passage of a law) and, in general, to estimate the causal effect of interventions in time series settings where there are no control units available, as in those cases where all units are treated
- In observational panel studies where a group of control units is also present, difference-in-differences and synthetic control methods may be preferred when the temporal dimension is short compared to the cross-sectional dimension

# PRACTICAL SESSION

[FMenchetti/CausalArima](#)

In the past few days, you were given sales data of one of the store brand products analyzed in the paper. We will now show how to use C-ARIMA to answer the following questions:

- 1 What is the total number of units sold due to the price reduction after 1-week, 1 month and 3 months from the intervention?
- 2 What is the total number of units sold at the end of the analysis period?
- 3 Plot the causal effect and residual diagnostics
- 4 Assume a multiplicative effect, how much more did the product sold due to the premanent price reduction?

[Go to the exercise!](#)

# REFERENCES

## Thanks for your attention!

- [1] Brodersen, K. H., Gallusser, F., Koehler, J., Remy, N. and Scott, S. L. [2015]. Inferring causal impact using Bayesian structural time-series models, *Annals of Applied Statistics* **9**: 247–274.
- [2] Menchetti, F., Cipollini, F. and Mealli, F. [2021]. Estimating the causal effect of an intervention in a time series setting: the C-ARIMA approach, *Preprint*. Available at <https://arxiv.org/abs/2103.06740> .