policy. The published version of the article can be accessed at: Assessing the biological relevance of aquatic connectivity to stream fish communities (doi: 10.1139/cjfas-2013-Title: Assessing the Biological Relevance of Aquatic Connectivity to Stream Fish Communities 1 2 3 4 5 6 7 Authors: Shad Mahlum^a: Department of Biology, Memorial University, St. John's, NL A1B 3X9, Canada. 8 9 Email: skm311@mun.ca Dan Kehler: Parks Canada, 1869 Upper Water St., Halifax, NS, B3J 1S9, Canada. Email: 10 11 dan.kehler@pc.gc.ca 12 David Cote^b: Ocean Sciences Centre, Memorial University of Newfoundland, St. John's NL A1C 5S7, Canada. Email: dave.j.cote@gmail.com 13 Yolanda F. Wiersma: Department of Biology, Memorial University, St. John's, NL A1B 3X9, 14 Canada. Email: ywiersma@mun.ca 15 16 Les Stanfield: Ontario Ministry of Natural Resources, 41 Hatchery Lane, Picton, ON K0K 2T0, Canada. Email: les.stanfield@ontario.ca 17 18 19 20 21 22 ^a Corresponding author

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Abstract:

Recent advances in the ability to quantify longitudinal connectivity of riverine systems is enabling a better understanding of how connectivity affect fish assemblages. However, the role of connectivity relative to other factors such as land use in structuring biological assemblages is just emerging. We assessed the relevance of a structural connectivity index to stream fish communities in five watersheds and examined whether species' sensitivities to connectivity are in accordance with expectations of life history. While controlling for the confounding effect of land use, elevation, and stream topology, we demonstrate that structural connectivity explains significant amounts of variation in community structure (1 to 5.4% as measured by Bray-Curtis similarity) and single species metrics (3 of 7 species abundances). The lower explanatory power of our models compared to studies done at smaller scales suggests that the relevance of connectivity to fish communities is scale dependent and diminishes relative to other environmental factors at larger spatial extents.

Keywords: Fragmentation, Structural Connectivity, Functional Connectivity

Introduction:

The increased awareness of the effects of anthropogenic structures that may act as barriers on aquatic ecosystems has prompted new research to understand, quantify, and mitigate fragmentation impacts (Fullerton et al. 2010). Previous work has focused on individual barriers and how they influence aquatic communities (Coffman 2005, Mahlum et al. 2014, Warren and Pardew 1998). However, recent efforts have extended the spatial scope to consider the effects of multiple potential barriers (Cote et al. 2009, O'Hanley 2011, Padgham and Webb 2010); which theoretically can act in a cumulative fashion at the scales fish operate.

Terrestrial landscape-scale metrics of connectivity have been well studied over the last 30 years, with aquatic environments simply being regarded as a habitat feature embedded within the terrestrial landscape (Wiens 2002). Increasingly, basic principles from landscape ecology have been tailored for river ecosystems (Fausch et al. 2002, Ward 1998, Ward et al. 2002). Following this foundational work, several research efforts have developed ways to measure structural connectivity that are appropriate for the dendritic nature of aquatic systems. These include score and ranking methods (Pess et al. 1998, Poplar-Jeffers et al. 2009, Taylor and Love 2003), optimization techniques (Kemp and O'Hanley 2010, O'Hanley 2011), patch-based graphs (Erős et al. 2012, Erős et al. 2011, Schick and Lindley 2007), and connectivity indices (Cote et al. 2009, Padgham and Webb 2010). These methods are particularly accommodating and valuable in prioritizing restoration efforts, as reconnecting aquatic habitats can be costly (Bernhardt et al. 2005, Januchowski-Hartley et al. 2013). However, the use of structural indices are predicated on being able to efficiently improve ecological integrity by maximizing *assumed* biological gains by increasing structural connectivity (Cote et al. 2009, O'Hanley 2011, Schick and Lindley 2007), from the removal or restoration of particular barriers. Although these indices provide

conceptually simple methods to systematically improve structural connectivity, it is poorly understood whether the recommendations yield biologically meaningful results (see Perkin and Gido 2012 for an exception). It is therefore necessary to understand the limitations (both statistical and ecological) of structural indices at predicting ecological responses in aquatic communities (Kupfer 2012).

One method to assess the ecological relevance of structural indices is to test for relationships between a given structural index and biological community patterns across stream systems with variable degrees of fragmentation. For instance, Perkin and Gido (2012) found a strong relationship between fish community structure, within second and third order stream units, and a structural connectivity index. Understanding the response of structural indices at small spatial extents is an important development, yet it remains unknown whether these relationships will continue to be present at broader spatial extents where confounding variables may have an increased influence on aquatic communities. For example, Branco et al. (2011) found that environmental and human pressures, but not the presence of barriers, were the dominant driver of the distribution of several potamodromous and resident fish species in a 3600 km² watershed. However, Branco et al. (2011) acknowledged that they used a relatively simple index of connectivity and called for a more thorough assessment of connectivity at broader spatial extents.

We analyzed the relationship between structural connectivity and patterns in the fish community using data from five 5th and 6th order watersheds in southern Ontario, Canada, (ranging in extent from 98 km² - 283 km²) which have a high degree of biodiversity (regional species richness of 38). The focus of this study was to determine if a relatively simple structural index, the Dendritic Connectivity Index (DCI), has biological relevance. Although we expect multiple confounding variables (e.g., elevation, watershed land use, stream network topology) to

contribute to the explanation of patterns in community structure; we expected changes in fish community data in response to variation in the DCI. Specifically, once other habitat factors are accounted for, elevated connectivity will reflect habitat attributes of increased patch size and accessible habitat and should support a broader range of stream biota (Bain and Wine 2009, Peterson et al. 2013). Therefore, it is expected that we would see relative increases in species richness and fish abundance with increases of the DCI. We also tested the importance of the DCI for individual fish species for both presence and abundance data. At an individual species level, we expect to see an increase in species presence and abundance as connectivity increases across sites. Primarily, it is anticipated that individual species that have life histories that require broad scale movements (e.g., salmonids) will be more affected by losses in connectivity than species that may not require the same broad scale movements (e.g., cottids).

Methods:

97 Study Area:

Southern Ontario exhibits a high degree of freshwater fish biodiversity (Chu et al. 2003). The diversity is attributed to a combination of postglacial dispersal and the anthropogenic introduction of non-native species (Dextrase and Mandrak 2006). The study was conducted in the watersheds of Wilmot, Oshawa, Ganaraska, Cobourg, and Duffins in southern Ontario, just east of the metropolitan area of Toronto (Figure 1). The five watersheds studied are dominated by developed urban areas at their confluence with Lake Ontario, agricultural landscape in the mid reaches, and a mixture of forest and low intensity agriculture in the headwaters. They range in watershed size of 98 km² for Wilmot to 283 km² for Ganaraska.

106 Data Layers:

Fish community data and habitat variables (including the structural index) were incorporated into the analysis (Table 1). Fish sampling was conducted from 1997 to 2009 by various agencies as part of a collaborative monitoring program (TRCA, 2010) using the Ontario Stream Assessment Protocol (Stanfield 2010). Sample site locations are based on random stratified designs to characterize conditions within stream segments. A handful of long-term monitoring sample sites were initially selected based on their representative conditions which were averaged across sampling periods to eliminate pseudo-replication. Sites were a minimum length of 40 m and were bounded by "crossovers" (where the thalweg crossed to the opposite side of the stream) to ensure adequate sampling of all habitat types (Stanfield 2010). Furthermore, sample site lengths reflect from 5 to 10 bankfull widths and have been shown to provide reliable measures of fish assemblages across time and space for this study area (Stanfield et al., 2012). Single-pass electrofishing was used to capture fish at a targeted effort of 7 to 15 s/m². All fish were measured, weighed, and identified to species with the exception of lampreys (Petromyzontidae), which were identified to family due to inconsistencies in identification to the species level. Finally, we also excluded 16 sites from the analysis that appeared to exhibit difficulties with identification of one or more individuals to the species level.

Connectivity index:

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To measure the structural connectivity across the 5 watersheds, we employed the Dendritic Connectivity Index (Cote et al. 2009). The DCI is calculated based on the probability that an individual can move freely among random points in a dendritic network. This takes into consideration the amount of potential habitat between barriers along with a measure of passability for each barrier. Furthermore, the DCI is flexible in that it can be modified to address the natural connectivity of a stream based on both potamodromous (DCI_p) and diadromous

(DCI_d) life histories. The DCI_p applies to life histories of species that typically live in riverine
systems and do not require diadromous migration. DCI_p is defined as:

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$$D_p = \sum_{i=1}^{n} \frac{1}{i!} *_{ij}$$

where l is the length of the segment i and j, c_{ij} is the connectivity between segments i and j, and L is the total stream length of all stream segments. The DCI_d applies to all life histories that migrate between a fixed point (e.g., estuary) and all upstream areas within a riverine system.

136 DCI_d is calculated as:

where L is the total length of the stream sections, l_i is the length of section i, c_{ij} is the connectivity between segments i and j. While the DCI_p and DCI_d measure the overall connectedness of a system, it could be beneficial to apply a structural connectivity metric at finer spatial scales (e.g., stream reach) to control for local pressures of connectivity on the biotic community. As noted in Cote et al. (2009), the DCI_d can be applied to measure the connectivity from any stream segment to the rest of the watershed. We denote this value as DCI_s, and used this in models for data collected at the scale of the stream segment. We used the Fish Passage Extension (FIPEX v2.2.1) for ArcGIS (v9.3.1) using a hydrological stream network provided by OMNR to calculate connectivity scores (c_{ij}) described above.

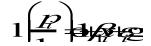
Determining barrier passability:

Identifying all potential barriers in a system is imperative in order to accurately assess connectivity (Cote et al. 2009, Januchowski-Hartley et al. 2013, O'Hanley 2011). A list of barrier locations was provided by OMNR which consisted of 298 locations of dams, perched culverts, and natural barriers across the 5 watersheds used in this study. We also used the

National Hydro Network obtained via GeoBase (http://www.geobase.ca/) to identify dams not in the OMNR dataset. Furthermore, road culverts are thought to outnumber dams by up to 38 times, with as much as 2/3 being designated as complete or partial barriers to fish movement (Januchowski-Hartley et al. 2013). Therefore, to help identify potential barriers not in the OMNR database, we used ArcGIS and files from GeoBase to identify intersections between streams (National Hydro Network) and roads (National Road Network) that would indicate a potential barrier and help create an inclusive barrier database to calculate the DCI. All sources of barrier locations were cross checked to prevent multiple occurrences of the same barrier in the dataset. We calculated and analyzed the DCI with regards to community structure and species richness with only known barriers and then again with the inclusion of potential barriers identified through GIS (stream/road intersections). The intent of this analysis was to provide insight into GIS-derived barrier locations and the potential benefits of modeling all potential barrier locations.

Determining passability values for potential barriers in these watersheds was challenging due to their vast number and the limited information available for them. This limitation is not unique to this study and underscores some of the common obstacles to riverscape-scale analyses in larger watersheds (for an example see Meixler et al. 2009). For the DCI, passabilities are defined as a value between 0 (impassable) and 1 (fully passable). Passability scores of zero were first assigned to all dams and perched culverts. Culverts were considered perched when the outlet bottom elevation was greater than the height of the outlet pool (Stanfield 2010). The remaining 75% of potential barriers lacked a passability score. Previous studies have found a relationship between culvert passabilities and channel slope (McCleary and Hassan 2008, Poplar-Jeffers et al. 2009), and we followed this approach to infer values for barriers with unknown passability. We

used an available data set from Terra Nova National Park (TNNP), Newfoundland, Canada that contained both passability scores and channel slopes. Passabilities in TNNP were calculated using FishXing (Furniss et al. 2006) and were based on the percent of time stream flows were within a passable range for brook trout (*Salvelinus fontinalis*). We calculated channel slope for culverts in Newfoundland and Ontario using a 10-m digital elevation model (DEM) by creating a 100 m diameter buffer around the barrier and taking the difference in elevation between the farthest upstream and downstream points and then dividing by the stream length between those points. Finally, we used a nonlinear regression model,



where i=1 to number of culverts (N), p is passability, x is channel slope, and $\epsilon_i \sim N(0,\delta^2)$, to estimate the relationship between culvert passability and channel slope in TNNP. This model fits a sigmoidal curve with a fixed passability of 1, when channel slope is 0. We then applied that relationship to the channel slopes associated with potential barriers in southern Ontario.

Accounting for confounding variables:

It is known that stream process and patterns are continually changing along the longitudinal gradient of the stream (Vannote et al. 1980) and these changes can significantly affect the biotic community (Fausch et al. 2002). Some of these influences can be segregated into habitat variables (e.g., elevation and stream width) and landscape use (e.g., urban and farmland). Several factors were incorporated into our analysis to control for confounding effects that influence community structure (see Table 1). These included elevation (Rahel and Hubert 1991, Stanfield and Kilgour 2006), land cover (Allan et al. 1997, Allan 2004, Stanfield and Kilgour 2006), stream network topology (Betz et al. 2010, Hitt and Angermeier 2008), and stream width (Cote 2007). We extracted elevation (ELE) for each sampling site from a 10-m DEM obtained

from OMNR. Land cover metrics that were thought to influence stream biota were quantified using the Southern Ontario Land Resource Information System (SOLRIS; Ontario Ministry of Natural Resources 2006) by determining the percentage of the watershed in each land cover type (Table 1). Using a metric analogous to stream order, we quantified the hydrological locations of sampling sites within the dendritic network using the Upstream Cell Count (UCC) which consists of the total amount of linear stream habitat above a sampling location (see Betz et al. 2010 for a detailed description). Lastly, stream width (SW) was measured during biological sampling by taking an average of 10 transects measuring SW throughout the sampling site (Stanfield 2010).

To select co-variables (Table 1) for the inclusion in our analysis, we used Akaike's Information Criteria (AIC) to select a candidate model that best explains the data and subsequently can be used for the inclusion of confounding variable in the following analysis of community structure, species richness, and species abundance (Akaike 1973, Burnham and Anderson 2002, Oksanen et al. 2012). Before we identified candidate models, we removed collinear variables (Spearman's rank correlations > 0.7). Next using variables identified in Table 1, *a priori* candidate models were created for the distance-based redundancy analysis (db-RDA, described below) on community similarities ranging from simple (single variable) to more complex (maximum 9 variables in our global model). To assess how well co-variables contributed to explaining the community data, we calculated the Δ AIC (difference in AIC values from the model with the smallest AIC value) and AIC weights (the amount of support that a given model is the best). Only models that were within Δ AIC < 2 of the top model were considered for the inclusion in the analysis (Burnham and Anderson 2002). To maintain consistency between the analyses of community structure, species richness, and species

abundances, we incorporated the same variables identified through the model selection procedure for all levels of analysis.

223 Analysis:

Is fish community similarity related to the DCI metrics?

A multivariate db-RDA was used to analyze how connectivity, as measured by the DCI_s, DCI_p, and DCI_d, affects community structure based on species abundances (Legendre and Anderson 1999). Distance based redundancy analysis is a robust analytical method used to assess the relationship between meaningful measures of species associations (e.g., Bray-Curtis index) and fixed effects within a linear model framework. Furthermore, we chose to use a db-RDA to 1) accommodate for non-Euclidean distance measures used in community similarity metrics; 2) control for confounding variables; and 3) use nonparametric permutation methods which freed us from the assumption of normality (Legendre and Anderson 1999). Prior to the multivariate analysis, a fourth root transformation of the abundance data was employed to emphasize diversity (Clarke and Warwick 2001). Then, we used the Bray-Curtis Index (Bray and Curtis 1957) as a measure of the similarity of the fish communities between sites because of its robustness and appropriateness for ecological community data (Clarke and Warwick 2001, Faith et al. 1987). Finally, a correction factor was not incorporated for the negative-eigenvalues to correct for Type 1 errors based on McArdle and Anderson (2001). Significance was determined by a pseudo-F statistic at alpha = 0.05.

Is DCIs related to fish species richness?

We used a generalized linear mixed model (GLMM) approach to test the effects of connectivity as determined by the DCI_s on species richness. Treating watershed as a random effect allowed us to account for the potential pseudo-replication within watersheds (Bates et al. 2011). Species richness was quantified by calculating the total number of fish species at each site. For sites with repeated sampling, species richness was averaged across sampling periods. Our approach to calculate species richness was chosen to provide a more accurate measure of this indicator than the single "most recent" observation that was used in the analyses by Stanfield and Kilgour (2006). Averaging richness across sampling events captures temporal variability and minimizes effects of sampling bias/error, but potentially undervalues diversity where sampling effort was lower (Kennard et al. 2006, Stanfield et al. 2013). Finally, using the GLMM, we analyzed the relationship between the DCI_s and the species richness of a site while controlling for confounding variables previously identified. All variables but watershed were treated as fixed effects. Significance was determined by the *z*-statistic at alpha = 0.05.

Is DCIs related to presence and abundance of individual species?

We also tested to see how connectivity, calculated with known barriers and potential barriers, affected the presence and abundance of individual species. Seven relatively abundant species across three families were selected to represent a wide range of life history characteristics (e.g., diadromous) and that were also relatively abundant across sites (Table 2 and 3). We again used a GLMM approach, with presence modeled as binomial and abundance as a Poisson response variable. Watershed was treated as a random effect to account for potential pseudo-replication of observations within watersheds. The same confounding variables identified in the model selection procedures described above were also included as fixed effects. Because the

abundance data exhibited considerable overdispersion, we used a resampling approach (Markov Chain Monte Carlo) to assess significance (Hadfield 2010). All statistical analysis was carried out with the statistical program R (v. 2.15.2, R Development Core Team 2012).

Results:

A total of 273 stream sites were selected across 5 watersheds (range of 27 to 70 sites per watershed). We used the selected sites for all levels of analysis within this study. A total of 38 species were sampled across the study sites with a mean of 25.4 species per watershed (range = 21 to 28). In addition to the 298 barriers identified by OMNR, we identified an additional 85 dams and 1,041 potential barriers. The relationship between stream slope and passability obtained from barriers in Terra Nova National Park was reasonably strong ($r^2 = 0.68$; Figure 2). When applied to potential barriers in southern Ontario, the predicted passabilities of un-surveyed barriers ranged from 0.0 to 0.99 with the passabilities strongly skewed towards the right, which indicates greater passability (Figure 3). Calculated connectivity scores for our study area watersheds in southern Ontario ranged from 0.0 to 41.1 for DCI₈ at the site scale, 14.9 to 22.6 for the DCI_p, and 0.3 to 31.2 for the DCI_d, the latter two versions calculated at the watershed scale (Table 4).

Twenty-two different models were analyzed with AIC scores (Table 5). Results of the Spearman's correlation matrix indicated that SW and UCC were highly correlated (r = 0.8). As a result, we did not include SW and UCC in the same model. The top model for the db-RDA of community similarity (Δ AIC < 2) included ELE, SW, and the land cover metric of built-up areapervious (BUAP), which indicates areas of urban development. All other additional confounding variables did not adequately explain community structure given the dataset and were represented

in models that had $\Delta AIC > 2$. The top model had a weight of evidence of 80 percent in support of the top model, and to maintain consistency between the different analyses, we elected to use ELE, SW, and BUAP to control for confounding effects in subsequent facets of our analysis. Furthermore, while it is likely that we would identify that the selected variables would relate differently to each level of analysis (e.g., community structure vs individual species) and within different univariate analyses (e.g., individual species), we chose to run a single model selection procedure to simplify the analysis and subsequent interpretation of the results between the different levels of the analysis. Moreover, we also found that several variables (e.g., elevation and stream width) remained consistent between this study and other studies within the same geographic area (see Stanfield et al. 2006), indicating that we would gain relatively little from additional model selection procedures.

We explained 21.1, 21.4, and 24.4 percent of the total variation in species composition with the db-RDA models used to analyze the relationships between the DCI_s, DCI_p, and DCI_d, calculations based on known barriers, and community structure for abundance data. Furthermore, we used a type III sum of squares and found all three co-variables significantly related to community structure in all three models (Models 1-3; Table 6). The DCI_s, DCI_p, and DCI_d was significantly related to community structure as well (F = 3.67, df = 1, p < 0.01; F = 4.74, df = 1, p < 0.005; F = 15.64, df = 1, p < 0.005 respectively). A positive correlation was also seen for the DCI_s (r = 0.65) and DCI_d (r = 0.48) for axis 1 and a negative correlation was seen for the DCI_p with axis 2 (r = -0.67).

The db-RDA models used to analyze the relationships between the DCI_s, DCI_p, and DCI_d, calculated with known barriers and potential barriers, and community structure for abundance data, with the co-variables of ELE, SW, and BUAP, explained 21.9, 22.2, and 24.4 percent of the

312 total variation in species composition respectively (Models 4-6; Table 6 and Figure 4). Using 313 additional barrier information derived from GIS data modestly improved our models and the 314 amount of variation explained with our connectivity metric by 1.5, 1.3, and 0.0% respectively. 315 Following the trends with the models which used only known barriers (models 1-3), we found 316 that all confounding variables for models 4-6 significantly explained community structure (Table 317 6). In these models, the DCI_s, DCI_p, and DCI_d were also significantly related to community structure (F = 6.37, df = 1, p < 0.005; F = 7.64, df = 1, p < 0.005; F = 15.52, df = 1, p < 0.005 318 319 respectively). However, the directions of the relationships were confounded between models for elevation, stream width, BUAP and DCI_s (Table 6).

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Species richness was not associated with changes in connectivity based on known barriers (DCI_s; z = 1.26, n = 273, p-value = 0.204; Figure 5a). However, when we included potential barriers into the DCI calculation, species richness became weakly correlated with the DCI_s (z = 1.99, n = 273, p-value = 0.047; Figure 5b) as were ELE and SW (z = -0.003, n = 273, p-value < 0.001; z = 0.05, n = 273, p-value < 0.001 respectively). However, the land cover variable used (BUAP) did not show a significant relationship with species richness (z = 0.068, n = 273, p-value = 0.058).

The presence of only two species had a positive relationship with the DCI_s: rainbow trout (Oncorhynchus mykiss) and mottled sculpin (Cottus bairdii; z = 0.07, n = 273, p-value = <0.001and z = 0.017, n = 273, p-value = <0.001 respectively; Table 2). Furthermore, abundance increased for rainbow trout (mean = 0.07, n = 273, p-value = 0.001), mottled sculpin (mean = 0.09, n = 273, p-value = 0.001), and longnose dace (mean = 0.05, n = 273, p-value = 0.014; Rhinichthys cataractae) with an increase in the DCI_s (Table 3; Figure 6). At least one

confounding variable had a significant relationship in the individual species analysis, where ELE was the dominant predictor variable most commonly seen between the species.

Discussion:

The use of connectivity indices as a tool to assess the fragmentation of a system and assist in prioritizing restoration efforts can be a valuable asset in reconnecting aquatic habitat patches. While minimal, we demonstrated that the DCI has biological relevance with regards to understanding fish communities and individual species distribution and abundance, even in the presence of confounding variables such as elevation, stream width, and land cover. Although it is necessary to address alternate pressures simultaneously when improving biological connectivity, selecting barriers to restore based on structural gains in connectivity can contribute to recovery and persistence of the aquatic community.

This conclusion is also consistent with findings by Perkin and Gido (2012) who found a significant relationship between the same connectivity index analyzed here and community structure within relatively fine scale study units consisting of second and third order streams. However, the fine spatial extents examined in that study likely minimized confounding variables and showed a much stronger relationship between connectivity and fish communities ($r^2 = 0.66$). Since the importance of environmental factors to stream biota is often scale-dependent (Fausch et al. 2002, Poff 1997, Wiens 2002), it remains unknown whether links between structural connectivity and communities will persist at spatial extents broader than the present study. However, it has been shown that increases in interpatch distance significantly decrease landscape connectivity (Goodwin and Fahrig 2003) and it could be expected that the same trends would persist in aquatic environments. Structural indices have been increasingly used to determine the

degree of connectivity across watersheds but interpretation of these results are hampered by the lack of demonstrations of biological relevance to aquatic ecosystems (Tischendorf and Fahrig 2000). Understanding these relationships is important to provide context into the appropriateness and limitations of simple structural indices, such as the DCI, and their use in aquatic ecosystems.

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The biology of the species in this study likely impacted the sensitivity of the species to structural connectivity. This study found relationships between the DCI_s and the abundance of several species. As expected, we found species that require extensive movements during their life history (e.g., rainbow trout) were significantly influenced by a lack of longitudinal connectivity (DCI_s). In contrast, other species (mottled sculpin and longnose dace), less known for extensive migration (Johnston 2003), were also influenced by the presence of anthropogenic barriers. Past studies have found local scale effects of barriers on small stream fishes (Coffman 2005, Norman et al. 2009, Warren and Pardew 1998). However, as documented by Meixler et al. (2009), it appears that local scale effects of barriers can translate into population wide impacts on the persistence of at least some small stream fishes. Furthermore, some of our species-specific expectations with regards to connectivity did not bear out. For example, we expected brook trout, a native species to the study area, would be more affected by losses in connectivity than other species because they require a variety of habitats throughout their life cycle, which could result in long migrations (Gowan and Fausch 1996). However, the presence of anthropogenic barriers did not seem to have a significant relationship with brook trout abundance. This may be attributed to low abundance or confounding variables not modeled in this study. For instance, brown trout (Salmo trutta) impact brook trout through competition of important habitat (e.g., spawning habitat and refugia) and predation (Fausch and White 1981). Similarly, others (e.g., Stanfield et al. 2006) have found that brook trout distribution and abundance in this area are affected by the cumulative effects of competition from multiple salmonids and land use. Supporting Fausch and White (1981) and Stanfield et al (2006), we found a strong elevation influence between these two species implying that brook trout are being pushed into the headwaters where competition is lessened. Although fragmentation may be a factor in the eventual recovery of brook trout and other salmonids, it appears that other confounding variables currently have a greater impact on the persistence of this species. Continuing to improve our understanding of the role of fragmentation in species distributions will assist managers in the recovery of imperiled species and how to mitigate the effects of anthropogenic disturbances.

In the absence of anthropogenic barriers, alternate pressures can influence ecological processes and patterns (Fagan 2002, Hargis et al. 1999). In addition to the modest effects of the DCI, elevation, stream width, and land cover had a strong relationship with community structure as well as with individual species (as observed by Stanfield et al. 2006). This supports previous connectivity studies that found environmental factors affected metapopulations (e.g., land cover and water quality; Branco et al. 2011, Meixler et al. 2009). Confounding variables such as the ones modeled here are an important aspect associated with stream communities and controlling for these environmental variables will help assist in determining how structural indices influence stream biota.

Presenting connectivity at watershed scales is useful to estimate watershed health or to prioritize restoration actions, but can be limiting for analyses aimed at local scales (e.g., studies targeting site-specific relationships between fish communities and habitat variables; Cote 2007). To address the need for locally-focused studies, we modified this watershed scale index into a local habitat variable (DCI_s) and matched it to corresponding biotic information. We consider this a useful addition to typical quantification methods of connectivity that either focused

primarily on barrier prioritization (Kemp and O'Hanley 2010, O'Hanley 2011, Poplar-Jeffers et al. 2009) or are overly simplistic (e.g., count of the number of barriers; Branco et al. 2011), and therefore miss important aspects of fragmentation (for a review see Kindlmann and Burel 2008, Padgham and Webb 2010). Measuring connectivity at a scale coincident with other aquatic community variables will expand the understanding of how connectivity processes relate to biota and will be useful in theoretical and management applications.

Identifying barrier locations is an important aspect in the management of aquatic systems. The failure to account for all barriers may result in costly management actions that produce negligible ecological benefits if the analysis fails to identify limiting factors (Bernhardt et al. 2005, Januchowski-Hartley et al. 2013). Although minimal barrier information (known barriers) significantly explained community structure, we saw an improvement with the inclusion of potential barriers (stream/road intersections) both in explaining community structure and species richness. This conclusion lends support to Januchowski-Hartley et al. (2013) who advocate for the incorporation of all potential barriers into current barrier databases.

We had relatively low explanatory power to explain community structure and species richness and we were unable to predict abundance of several species (4 of 7 species) with aquatic connectivity. One explanation could be in our methodology for calculating passability. Identifying the passability of barriers was the largest obstacle in assessing connectivity over the relatively large study area. While direct site evaluations of all known and potential barriers in a system is recommended and could potentially improve our predictive power, the large number of barriers within this study required us to identify an alternate method to assess passability. A priority for future work in these watersheds should be a more comprehensive inventory of dams on private lands (e.g., ponds). The use of GIS allowed us to identify potential barriers based on

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locations where streams and roadways intersected. However, assigning passability values required estimates based on known relationships with channel slope in another well studied area. Furthermore, our passabilities were based on brook trout movements. This is not appropriate for all species and likely overestimates passage for many species (e.g., Cyprinidae; Coffman 2005, McLaughlin et al. 2006). Thus functional connectivity for these species may actually be lower in these five watersheds than predicted by our model. Similarly, for species (e.g., *Salmo salar*) thought to have higher swimming/jumping ability than brook trout, these watersheds may actually have higher functional connectivity than predicted here. While the relationship between channel slope and passability allowed us to identify potential barrier passabilities, it is recommended that managers accurately inventory and assess the passability of all barriers across study areas to allow them to maximize habitat gains with current connectivity models.

Based on organisms' response to fragmentation in terrestrial systems, it is reasonable to expect that thresholds of aquatic connectivity also exist and are associated with the biology of the focal organism or community. Within our five watersheds, only the lower end of the connectivity spectrum were captured and thus critical thresholds may exist outside the range studied here. Capturing the full spectrum of possible connectivity scores at watershed scales may be difficult as pristine and highly fragmented stream systems will likely differ from one another in many other ways. However, identifying ecological thresholds for connectivity will assist with setting management goals for protection and recovery of focal species..

As in terrestrial landscape ecology, where work has been done to link structural connectivity metrics with ecological response (i.e., functional connectivity, Kindlmann and Burel 2008, Tischendorf and Fahrig 2000), we have shown that aquatic structural connectivity indices can do the same. The structural indices, derived from relatively straightforward physical

parameters (e.g., stream length, barrier properties), help to explain biologically relevant phenomena such as habitat quality and observed fish movement across barriers. It remains necessary to further incorporate the organisms' perceptions of its landscape into structural indices to achieve meaningful measures of connectivity (Kindlmann and Burel 2008), but doing so comes with tradeoffs such as increased data requirements, computational complexity, and decreased ease of interpretation (Kupfer 2012). Moreover, incorporating more functional metrics without understanding their limitations may not necessarily increase their validity (Kupfer 2012). Recent work by Bourne (2013) found that incorporating a more functional habitat variable into structural indices influenced the magnitude of fragmentation of a system but not necessarily the qualitative conclusions (i.e., prioritization of the restoration action) when compared to physical properties of habitat. This indicates that, at least in some cases, simple physical measurements may be appropriate, and can save considerable time and resources.

Considerable work remains to understand how processes associated with aquatic connectivity relates to faunal communities. The availability of structural connectivity metrics and indices that have been evaluated for their ecological relevance and an understanding of their limitations will prove useful in future research and management efforts in this field.

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641 Tables:

Table 1: Categories of variables used in the analysis with the associated symbol used within the text. Predictions of the db-RDA for abundance is included in the table with (+) indicating a predicted change in community structure and (-) indicating no predicted change in community structure.

,			
Category	Variable	Symbol	Units
Fish Community	Abundance	A	Count
Structural Index	DCI_d	DCI_d	Percentage of natural connectivity
	DCI_p	DCI_p	Percentage of natural connectivity
	DCI_s	$\mathrm{DCI}_{\mathrm{s}}$	Percentage of natural connectivity
Stream Position	Up-Stream Cell Count	UCC	Count
	Elevation	ELE	Meters
	Stream Width	sw	Meters
Land Cover	Build-up area Pervious	BUAP	Proportion of watershed
	Build-up area Impervious	BUAI	Proportion of watershed
	Cropland	CR	Proportion of watershed
	Pasture and Abandoned Fields	PAF	Proportion of watershed
	Mixed forest	MF	Proportion of watershed
	Deciduous forest	DF	Proportion of watershed

Table 2. The results of the single species presence analysis. Predictions represent the expected relationship between the species and variable. Positive values indicate that species presence is predicted to increase with increases in the corresponding variable while negative values indicate that species presence is predicted to decrease with increases in the corresponding variable.

Species	Prediction	Variable	n	Estimate	SE	z-value	p-value
Oncorhynchus mykiss	+	ELE	273	-0.018	0.004	-4.417	< 0.001*
	-	BUAP		-0.518	0.467	-1.111	0.267
	+	SW		0.111	0.057	1.936	0.053
	+	DCI _s		0.058	0.016	3.757	< 0.001*
Salmo trutta	+	ELE	273	0.003	0.004	0.718	0.473
	-	BUAP		-0.837	0.217	-3.854	< 0.001*
	+	SW		0.349	0.062	5.641	< 0.001*
	+	DCI_s		0.020	0.016	1.255	0.209
Salvelinus fontinalis	+	ELE	273	0.030	0.005	6.471	< 0.001*
	-	BUAP		-0.674	0.308	-2.191	0.028*
	+	SW		0.109	0.063	1.723	0.085
	+	DCI _s		0.002	0.016	0.121	0.903
Rhinichthys obtusus	-	ELE	273	-0.018	0.004	-4.485	< 0.001*
	-	BUAP		0.557	0.282	1.977	0.048*
	-	SW		0.014	0.063	0.214	0.830
	+	DCI_s		-0.019	0.015	-1.253	0.210
Rhinichthys cataractae	- ·	ELE	273	-0.018	0.005	-3.721	< 0.001*
	-	BUAP	273	0.375	0.538	0.696	0.486
		SW		0.760	0.112	6.758	< 0.001*
	+	DCI _s		0.019	0.021	0.883	0.377
Semotilus atromaculatus	-	ELE	273	-0.013	0.003	-4.070	< 0.001*
	-	BUAP		0.531	0.210	2.531	0.011*
	-	SW		-0.051	0.051	-0.990	0.322
	+	DCI_s		0.008	0.013	0.639	0.523
Cottus bairdii	-	ELE	273	-0.005	0.004	-1.310	0.190
	-	BUAP		0.021	0.560	0.037	0.971
	+	SW		0.196	0.059	3.289	0.001*
	+	DCI_s		0.081	0.017	4.917	< 0.001*

^{*} indicates significance at $\alpha = 0.05$

Table 3. The results of the single species abundance analysis. Predictions represent the expected relationship between the species and variable. Positive values indicate that species abundance is predicted to increase with increases in the corresponding variable while negative values indicate that species abundance is predicted to decrease with increases in the corresponding variable.

Species	Prediction	Variable	n	Estimat	e SE	p-value
Oncorhynchus mykiss	+	ELE BUAP	273	-0.02 -0.50	0.00 0.02	0.001* 0.310
	+	SW		0.16	0.02	0.001*
	+	DCIs		0.07	0.00	0.001*
Salmo trutta	+	ELE	273	0.02	0.00	0.001*
	-	BUAP		-0.96	0.02	0.082
	+	SW		0.42	0.00	0.001*
	+	DCIs		0.02	0.00	0.126
Salvelinus fontinalis	+	ELE	273	0.03	0.00	0.001*
,	-	BUAP		-0.62	0.02	0.084
	+	SW	4	0.02	0.00	0.792
	+	DCIs		-0.01	0.00	0.722
Rhinichthys obtusus	-	ELE	273	-0.02	0.00	0.001*
	-	BUAP		0.58	0.02	0.154
	-	SW		-0.07	0.00	0.212
	+	DCIs		0.00	0.00	0.756
Rhinichthys cataractae	-	ELE	273	-0.02	0.00	0.001*
	-	BUAP		0.06	0.04	0.920
	-	SW		0.65	0.00	0.001*
	+	DCIs		0.05	0.00	0.014*
Semotilus atromaculatus	-	ELE	273	-0.02	0.00	0.001*
		BUAP		0.74	0.03	0.262
	-	SW		-0.18	0.00	0.004*
	+	DCIs		0.00	0.00	0.898
Cottus bairdii	-	ELE	273	-0.01	0.00	0.060
	-	BUAP			0.06	
	+	SW			18 0.00	
	+	DCIs		0.	09 0.00	0.001*

^{*} indicates significance at $\alpha = 0.05$

Table 4. Dendritic Connectivity Index scores for each watershed.

		Known B	arriers	Known Barriers with Stream/River				
				Intersects				
Watershed	DCI_p	DCI_d	DCI _s Range	DCI_p	DCI_d	DCI _s Range		
Duffins	35.4	2.3	0.0 - 58.52	16.1	1.7	0.0 - 35.0		
Oshawa	24.2	42.0	0.0 - 46.63	16.8	24.8	0.4 - 33.7		
Cobourg	20.4	32.4	0.0 - 32.35	14.9	22.1	0.0 - 26.2		
Ganaraska	24.4	0.4	0.0 - 46.63	18.4	0.3	0.5 - 39.1		
Wilmot	51.3	67.0	0.0 - 67.02	22.6	31.2	14.9 - 41.1		

Table 5. The results of co-variable selection based on the Akiake's Information Criterion for 22 combinations of predictor variables against the db-RDA of community similarity using abundance data (CS).

Model	K	AIC	ΔΑΙС	Exp	Weight
CS ~ ELE + SW + BUAP	4	1181.69	0.00	1.000	0.805
$CS \sim ELE + UCC + BUAP$	4	1184.79	3.09	0.213	0.171
$CS \sim ELE + SW + BUAI$	4	1189.69	8.00	0.018	0.015
$CS \sim ELE + UCC + BUAI$	4	1192.71	11.01	0.004	0.003
$CS \sim ELE + SW + PAF$	4	1194.27	12.58	0.002	0.001
$CS \sim ELE + SW + FAP$	4	1194.27	12.58	0.002	0.001
$CS \sim ELE + SW + MF$	4	1194.54	12.85	0.002	0.001
$CS \sim ELE + SW + DF$	4	1197.75	16.05	0.000	0.000
$CS \sim ELE + UCC + PAF$	4	1197.76	16.06	0.000	0.000
$CS \sim ELE + UCC + FAP$	4	1197.76	16.06	0.000	0.000
$CS \sim ELE + SW + CR$	4	1197.90	16.21	0.000	0.000
$CS \sim ELE + SW + CR$	4	1197.90	16.21	0.000	0.000
$CS \sim ELE + UCC + MF$	4	1198.08	16.39	0.000	0.000
$CS \sim ELE + UCC + MF$	4	1198.08	16.39	0.000	0.000
$CS \sim ELE + SW$	3	1201.23	19.54	0.000	0.000
$CS \sim ELE + UCC + DF$	4	1201.56	19.87	0.000	0.000
CS ~ ELE + UCC + CR	4	1202.14	20.44	0.000	0.000
CS ~ ELE + UCC	3	1205.64	23.95	0.000	0.000
CS ~ ELE	2	1211.63	29.93	0.000	0.000
CS ~ SW	2	1216.68	34.98	0.000	0.000
CS ~ UCC	2	1219.02	37.32	0.000	0.000
a CS \sim ELE + UCC + SW + BUAP + BUAI + CR + PAF + MF + DF	8	1226.36	44.66	0.000	0.000

^a Represents the global model (model that includes all variables) used in the model selection.

Table 6. The output of 6 different models for abundance to determine the relationship between longitudinal connectivity as measured by the Dendritic Connectivity Index (Cote et al. 2009) and community structure as measured by the Bray-Curtis similarity. Abundance 1 models used DCI values calculated with only known barriers whereas Abundance 2 models used DCI values calculated with known barriers and potential barriers.

Model	df	% Variation Explained	Pseudo-F	p-value	Axis 1	Axis 2
Abundance 1:						
Full Model 1	4	21.1	17.93	0.005		
ELE	1	8	17.83	0.005	0.91	-0.16
BUIP	1	8.7	21.79	0.005	0.11	0.88
SW	1	3.8	12.6	0.005	-0.77	-0.28
DCIs	1	1.2	3.76	0.01	-0.49	-0.14
Residuals	268	78.3				
Full Model 2	4	21.4	18.23	0.005		
ELE	1	9.5	20.06	0.005	-0.92	0.17
BUIP	1	6.5	17.11	0.005	-0.13	-0.87
SW	1	3.9	12.82	0.005	0.79	0.25
DCIp	1	1	4.74	0.005	0.31	0.46
Residuals	268	79				
Full Model 3	4	24.4	21.64	0.005		
ELE	1	9.4	20.6	0.005	-0.77	0.54
BUIP	1	6.5	18.01	0.005	-0.43	-0.68
SW	1	4.3	13.65	0.005	0.74	-0.19
DCId	1	5.4	15.64	0.005	0.54	0.41
Residuals	268	74.4				
Abundance 2:						
Full Model 4		21.0	10.74	0.005		
	4	21.9	18.74	0.005	0.00	0.20
ELE	1	7.1	16.63	0.005	-0.88	-0.28
BUIP	1	8.4	21.33	0.005	-0.2	0.85
SW	1	3.6	12.39	0.005	0.77	-0.16
DCIs	1	2.7	6.37	0.005	0.65	-0.16
Residuals	268	78.1				
Full Model 5	4	22.2	19.14	0.005		
ELE	1	9.8	20.69	0.005	0.93	-0.16
BUIP	1	4.4	12.35	0.005	0.13	0.86
SW	1	4	13.01	0.005	-0.79	-0.23
DCIp	1	2.3	7.64	0.005	-0.22	-0.67
Residuals	268	79.6	,.01	0.005	0.22	0.07
residudis	200	17.0				

Full Model 6	4	24.4	21.6	0.005			
ELE	1	9.3	20.64	0.005		-0.78	0.52
BUIP	1	7.5	20.24	0.005		-0.41	-0.69
SW	1	4.4	13.66	0.005	(0.75	-0.17
DCId	1	5.4	15.52	0.005	(0.48	0.37
Residuals	268	73.4					

678 **Figures** 679 Figure 1. The study area in southern Ontario with the barrier locations. The insert illustrates an 680 example area of the Duffins. 681 682 Figure 2. Relationship between channel slope and passability in Terra Nova National Park, 683 Newfoundland and Labrador, Canada. We applied this relationship to barriers in Southern 684 Ontario to determine the passability of unidentified barriers. 685 686 Figure 3. Histogram of barrier passabilities in the study watersheds based on the relationship 687 between channel slope and culvert passability in Terra Nova National Park, Newfoundland and 688 Labrador, Canada. 689 690 Figure 4. The distance based redundancy analysis comparing the DCI_s, DCI_p and DCI_d (panels A, 691 B, and C respectively) calculated with known barriers and potential barriers; and associated co-692 variables (ELE = Elevation, SW = Stream Width, and BUAP = Built-up area-pervious) for 693 abundance data in southern Ontario. 694 695 Figure 5. Relationship between species richness and the DCI_s in 5 southern Ontario streams 696 while controlling for elevation, stream width, and built-up area-pervious. The DCIs in panel A is 697 calculated using only known barriers and the DCIs in panel B is calculated using known barriers 698 and potential barriers.

Figure 6. Relationship of the DCI and species abundances (solid line) and 95% confidence intervals (dashed line) for rainbow trout, longnose dace, and mottled sculpin.

