Introduction to Algorithms for Data Mining and Machine Learning

Chapter 4: Data Fitting and Regression

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For details, please read the book:

Xin-She Yang, Introduction to Algorithms for Data Mining and Machine Learning,
Academic Press/Elsevier, (2019).

Sample Mean and Variance

For a given set of n independent samples/observations $x_1,x_2,...,x_n$ of a random variable X, the sample mean can be calculated by

$$\bar{x} \equiv \langle x \rangle = \frac{1}{n}(x_1 + x_2 + \dots + x_n) = \frac{1}{n}\sum_{i=1}^n x_i,$$

which is essentially the arithmetic average of the values x_i .

Sample Variance

The sample variance S^2 is defined by

$$S^{2} = \frac{1}{n-1} \sum_{i=1}^{n} (x_{i} - \bar{x})^{2}.$$

It is worth pointing out that the factor is 1/(n-1), not 1/n. Loosely speaking, we need at least one sample to estimate the mean, but we need at least two samples to estimate the variance.



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Example

The measurements of a quantity such as the noise level. The readings in dB are :

$$66, \ 73, \ 73, \ 74, \ 83, \ 70, \ 69, \ 77, \ 72, \ 75.$$

From the data, we know that n=10 and the mode is 73 as 73 appears twice (all the rest only appears once). The sample mean is

$$\bar{x} = \frac{1}{10}(x_1 + x_2 + \dots + x_{10})$$

$$= \frac{1}{10}(66 + 73 + 73 + 74 + 83 + 70 + 69 + 77 + 72 + 75) = \frac{732}{10} = 73.2.$$

The corresponding sample variance can be calculated by

$$S^{2} = \frac{1}{n-1} \sum_{i=1}^{n} (x_{i} - \bar{x})^{2}$$

$$= \frac{1}{10-1} \sum_{i=1}^{10} (x_{i} - 73.2)^{2} = \frac{1}{9} [(66 - 73.2)^{2} + (73 - 73.2)^{2} + \dots + (75 - 73.2)^{2}]$$

$$= \frac{1}{9} [(-7.2)^{2} + (-0.2)^{2} + \dots + (1.8)^{2}] = \frac{195.6}{9} \approx 21.73.$$

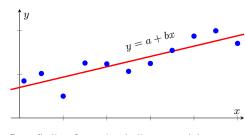
Thus, the standard derivation is $S = \sqrt{S^2} \approx \sqrt{21.73} \approx 4.662$.

Regression

The main idea of regression is to fit the data points (x_i,y_i) to a function y=f(x) (e.g., y=a+bx is a straight line). The best fit should minimize the sum squares of errors $\sum_{i=1}^n [y_i-f(x_i)]^2$. That is

$$\min \Psi = \sum_{i=1}^{n} [y_i - f(x_i)]^2,$$

which is the method of least squares.



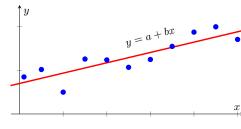
Best fit line for a simple linear model.

Regression

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Best fit line for a simple linear model.

From $\frac{\partial \psi}{\partial a}=0$ and $\frac{\partial \psi}{\partial b}=0,$ we have

$$\sum_{i=1}^{n} [y_i - (a + bx_i)] = 0, \quad \sum_{i=1}^{n} x_i [y_i - (a + bx_i)] = 0,$$

which can determine a and b uniquely. After some algebra, we have

$$a = \frac{1}{n} \left[\sum_{i=1}^{n} y_i - b \sum_{i=1}^{n} x_i \right] = \bar{y} - b\bar{x}, \quad b = \frac{n \sum_{i=1}^{n} x_i y_i - (\sum_{i=1}^{n} x_i)(\sum_{i=1}^{n} y_i)}{n \sum_{i=1}^{n} x_i^2 - (\sum_{i=1}^{n} x_i)^2},$$

where $\bar{x} = \frac{1}{n} \sum_{i=1}^{n} x_i$ and $\bar{y} = \frac{1}{n} \sum_{i=1}^{n} y_i$ are the sample means.

Correlation Coefficient

The correlation coefficient measures the degree of correlation between x and y.

$$r = r_{x,y} = \frac{n \sum_{i=1}^{n} x_i y_i - \sum_{i=1}^{n} x_i \sum_{i=1}^{n} y_i}{\sqrt{\left[n \sum_{i=1}^{n} x_i^2 - (\sum_{i=1}^{n} x_i)^2\right] \left[n \sum_{i=1}^{n} y_i^2 - (\sum_{i=1}^{n} y_i)^2\right]}}$$

which can also be written compactly as

$$r = \frac{nK_{xy} - K_x K_y}{\sqrt{(nK_{xx} - K_x^2)(nK_{yy} - K_y^2)}},$$

$$K_x = \sum_{i=1}^n x_i, \quad K_y = \sum_{i=1}^n y_i, \quad K_{xx} = \sum_{i=1}^n x_i^2, \quad K_{yy} = \sum_{i=1}^n y_i^2, \quad K_{xy} = \sum_{i=1}^n x_i y_i.$$

-1 < r < +1

- r > 0, it is positive correlation. That is, y increases as x increases. In case r = 1, it is a perfect, increasing linear relationship.
- r=0 (or near zero), there is no correlation between y and x.
- r < 0, it is negative correlation. r = -1 corresponds to a perfect, decreasing, linear relationship.

However, if |r| > 1, there must be some mistake in calculations!

For given the following given data:

	x_i	1	2	3	4	5
•	y_i	1.6	2.1	2.5	2.9	3.7

We have

$$K_x = \sum_{i=1}^{5} x_i = 1 + 2 + 3 + 4 + 5 = 15, \quad K_y = \sum_{i=1}^{5} y_i = 1.6 + 2.1 + 2.5 + 2.9 + 3.7 = 12.8$$

$$K_{xx} = \sum_{i=1}^{5} x_i^2 = 1^2 + \dots + 5^2 = 55, \quad K_{yy} = \sum_{i=1}^{5} y_i^2 = 1.6^2 + \dots + 3.7^2 = 35.32$$

$$K_{xy} = \sum_{i=1}^{5} x_i y_i = 1 \times 1.6 + \dots + 5 \times 3.7 = 43.4$$

So we have

$$a = \frac{K_{xx}K_y - K_xKxy}{nK_{xx} - K_x^2} = \frac{55 \times 12.8 - 15 \times 43.4}{5 \times 55 - (15)^2} = 1.06,$$

$$b = \frac{nK_{xy} - K_xK_y}{nK_{xx} - K_x^2} = \frac{5 \times 43.4 - 15 \times 12.8}{5 \times 55 - (15)^2} = 0.5.$$

So the best-fit line is

$$y = a + bx = 1.06 + 0.5x.$$

The correlation coefficient is

$$r = \frac{nK_{xy} - K_x K_y}{\sqrt{(nK_{xx} - K_x^2)(nK_{yy} - K_y^2)}} = 0.98976.$$

Exercise: Carry out the above regression in R and compare, the results

Linearization

Some model functions look nonlinear, but they can be converted into a linear form.

Example

Nonlinear function

$$f(x) = \alpha e^{-\beta x},$$

can be transformed into a linear form by taking logarithms of both sides

$$\ln f(x) = \ln(\alpha) - \beta x \Longrightarrow y = a + bx, \quad y = \ln f(x), \quad a = \ln(\alpha), \quad b = -\beta.$$

• Function $f(x) = \alpha x^{\beta}$ can also be transformed into

$$\ln[f(x)] = \ln(\alpha) + \beta \ln(x),$$

which becomes y=a+bv between $y=\ln[f(x)]$ and $v=\ln(x)$ with $a=\ln(\alpha)$ and $b=\beta$.

Similarly,

$$y = ax \exp(-x/b), \quad x, y > 0$$

can be converted to

$$\frac{y}{x} = ae^{-x/b}, \quad x \neq 0,$$

and then

$$\ln(\frac{y}{x}) = \ln a - \frac{1}{b}x,$$

which is linear between ln(y/x) and x.

Generalized Linear Regression

Fitting to a polynomial of degree p

$$y(x) = \alpha_0 + \alpha_1 x + \alpha_2 x^2 + ... + \alpha_p x^p,$$

is probably the most widely used. This is equivalent to the regression to the linear combination of the basis functions 1, x, x, ..., and x^p .

For given data points (x_i, y_i) (i = 1, 2, ...n), the minimization of the sum of the errors squared leads to

$$\begin{pmatrix} \sum_{i=1}^{n} 1 & \sum_{i=1}^{n} x_{i} & \dots & \sum_{i=1}^{n} x_{i}^{p} \\ \sum_{i=1}^{n} x_{i} & \sum_{i=1}^{n} x_{i}^{2} & \dots & \sum_{i=1}^{n} x_{i}^{p+1} \\ \vdots & & \ddots & & \\ \sum_{i=1}^{n} x_{i}^{p} & \sum_{i=1}^{n} x_{i}^{p+1} & \dots & \sum_{i=1}^{n} x_{i}^{2p} \end{pmatrix} \begin{pmatrix} \alpha_{0} \\ \alpha_{1} \\ \vdots \\ \alpha_{p} \end{pmatrix} = \begin{pmatrix} \sum_{i=1}^{n} y_{i} \\ \sum_{i=1}^{n} x_{i} y_{i} \\ \vdots \\ \sum_{i=1}^{n} x_{i}^{p} y_{i} \end{pmatrix},$$

which is a matrix equation. In principle, $\alpha = (\alpha_1, ..., \alpha_p)^T$ can be obtained by inversion, though the Gauss-Newton method is more efficient.

Example: Quadratic Regression

$$x: -0.98, 1.00, 2.02, 3.03, 4.00$$

 $y: 2.44, -1.51, -0.47, 2.54, 7.52$

For the formula $y = \alpha_0 + \alpha_1 x + \alpha_2 x^2$, we have

$$\begin{pmatrix} n & \sum_{i=1}^{n} x_{i} & \sum_{i=1}^{n} x_{i}^{2} \\ \sum_{i=1}^{n} x_{i} & \sum_{i=1}^{n} x_{i}^{2} & \sum_{i=1}^{n} x_{i}^{3} \\ \sum_{i=1}^{n} x_{i}^{2} & \sum_{i=1}^{n} x_{i}^{3} & \sum_{i=1}^{n} x_{i}^{4} \end{pmatrix} \begin{pmatrix} \alpha_{0} \\ \alpha_{1} \\ \alpha_{2} \end{pmatrix} \begin{pmatrix} \sum_{i=1}^{n} y_{i} \\ \sum_{i=1}^{n} x_{i}y_{i} \\ \sum_{i=1}^{n} x_{i}y_{i} \end{pmatrix}.$$

Using the data set, we have n=5, $\sum_{i=1}^{n} x_i = 9.07$ and $\sum_{i=1}^{n} y_i = 10.52$. Other quantities can be calculated in a similar way. Therefore, we have

$$\begin{pmatrix} 5.0000 & 9.0700 & 31.2217 \\ 9.0700 & 31.2217 & 100.119 \\ 31.2217 & 100.119 & 358.861 \end{pmatrix} \begin{pmatrix} \alpha_0 \\ \alpha_0 \\ \alpha_2 \end{pmatrix} = \begin{pmatrix} 10.52 \\ 32.9256 \\ 142.5551 \end{pmatrix}.$$

By direct inversion, we have

$$\begin{pmatrix} \alpha_0 \\ \alpha_1 \\ \alpha_2 \end{pmatrix} = \begin{pmatrix} -0.5055 \\ -2.0262 \\ 1.0065 \end{pmatrix}.$$

The best fit equation is

$$y(x) = -0.5055 - 2.0262x + 1.0065x^{2}.$$

Exercise: Implement this using R or Python and compare the results.

Goodness of Fit

In the preceding example, we fit the data using p=2 (quadratic). If we plot out the curve, it is well fit. However, without visualizing the data, how do we know it is well fit by choosing p=2? In fact, p is hyper-parameter (that we have to supply a value).

Using different p = 1, 2, 3, 4, we have $f_p(x)$:

$$f_1(x) = 0.9373x + 0.4038$$
 for $p = 1$.

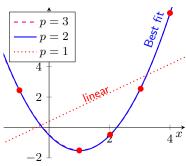
$$f_2(x) = 1.0065x^2 - 2.0262x - 0.5055$$
 (best fit).

$$f_3(x) = 0.0080x^3 + 0.9694x^2 - 2.0131x - 0.4580.$$

$$f_4(x) = 0.0101x^4 - 0.0610x^3 + 1.0778x^2$$
$$-1.9571x - 0.5798.$$

The fit error is the residual sum of squares (RSS)

$$RSS(p) = \sum_{i=1}^{n} [y_i - f_p(x_i)]^2 = \sum_{i=1}^{n} [y_i - \hat{y}_i(x_i)]^2,$$



The best fit is a quadratic function (p = 2).

where $\hat{y}_i(x_i) = f_p(x_i)$ are the predicted values.

Goodness of fit

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	Order	p = 1	p=2	p = 3	p=4
	RSS	36 3485	0.0045	0.0080	6.9×10^{-30}

p=4 leads to overfitting.

Over-fitting

However, higher-order polynomials tend to overfit the data, and oscillations become an issue for higher p. So, RSS is not a good criterion for measuring the goodness of the fit.

To choose the proper order p (hyperparameter) to avoid overfitting, better criteria are Bayesian information criterion (BIC) or Akaike information criterion (AIC).

Information Criteria: Minimization of AIC or BIC

For a model y = f(x) with k parameters to fit n data points, the AIC is defined by

$$AIC = 2k + n \ln(RSS/n).$$

For small number of samples (e.g., n/k < 40), the corrected AIC (i.e., AICc) becomes

AICc =
$$2k + n \ln(RSS/n) + \frac{2k(k+1)}{n-k-1}$$
.

With the same assumptions of errors obeying Gaussian distributions, we have

$$BIC = k \ln n + n \ln \left(\frac{RSS}{n} \right).$$

Both AIC and BIC are useful criteria, and their global minima give the best fit (k value). It is difficult to say which one is better, which may also depend on the types of problems. Let us revisit the previous example for fitting n=5 data points with polynomials.

Using the AIC criterion, and we have n=5 and k=p+1=2 (for p=1)

$$AIC = 2 \times 2 + 5\ln(36.3485/5) = 13.64.$$

Similarly, we have

$$AIC = -29.07, -24.19,$$

for p=2,3, respectively. Among these three values, p=2 has the lowest AIC and AICc. Since the valued start to increase for p=3, we can conclude that the best degree of fit is p=2 with k=3 parameters. The results of AIC and AICc values are summarized in the Table below.

Order	p=1	p=2	p=3			
k = p + 1	2	3	4			
RSS	36.3485	0.0045	0.0080			
AIC	13.92	-29.07	-24.19			
AICc	15.92	-17.07	Inf (division by zero)			
BIC	13 14	-30.24	-25 75			

We also list the BIC values and the conclusion p=2 is consistent with AIC/AICc.

Nonlinear Least Squares

Curve-fitting is optimization

Linear models are rare/special cases. In general, for n data points (x_i, y_i) (i = 1, 2, ..., n)to fit a model $f(x_i, \mathbf{a})$ with $\mathbf{a} = (a_0, a_1, ..., a_m)$ (i.e., m parameters), we have the method of nonlinear least squares to minimize the L_2 -norm of the residuals $R_i = y_i - f(x_i, \boldsymbol{a})$. That is to minimize the sum of the fit errors/residuals:

Minimize
$$E(\mathbf{a}) = \sum_{i=1}^{n} R_i^2(\mathbf{a}) = \sum_{i=1}^{n} [y_i - f(x_i, \mathbf{a})]^2 = ||R_i(\mathbf{a})||_2^2,$$

which can be solved by optimization techniques such as the Gauss-Newton Algorithm.

Gauss-Newton Algorithm

Let J denote the Jacobian matrix in the form

$$J = [J_{ij}] = \frac{\partial R_i}{\partial a_i}, \quad i = 1, 2, ..., n, \quad j = 0, 1, 2, ..., m,$$

which is an $n \times (m+1)$ matrix. Then the gradient of the objective (error) function is

$$\frac{\partial E}{\partial a_j} = 2 \sum_{i=1}^n \frac{\partial R_i}{\partial a_j} R_i, \quad j = 0, 1, 2, ..., m.$$

This gradient can be written in the vector form ∇E with m+1 components as

$$\nabla E(\boldsymbol{a}) = 2\boldsymbol{J}^T \boldsymbol{R},$$

where the residual vector \mathbf{R} is given by $\mathbf{R}(\mathbf{a}) = [R_1(\mathbf{a}), R_2(\mathbf{a}), ..., R_n(\mathbf{a})]^T$. The Hessian matrix \boldsymbol{H} [of size $(m+1) \times (m+1)$] of E can be written as

$$\boldsymbol{H} = \nabla^2 E = 2\sum_{i=1}^n [\nabla R_i \nabla R_i^T + R_i \nabla^2 R_i] = 2\boldsymbol{J}^T \boldsymbol{J} + 2\sum_{i=1}^n R_i \nabla^2 R_i,$$

which can be approximated by (after ignoring all the higher-order terms)

$$\boldsymbol{H} \approx 2 \boldsymbol{J}^T \boldsymbol{J}$$
.

The minimization of E(a) can be solved by Newton's method (starting with an initial guess a_0), and we have

$$\boldsymbol{a}_{t+1} = \boldsymbol{a}_t - \frac{\nabla E}{\nabla^2 E} = \boldsymbol{a}_t - \frac{\nabla E}{\boldsymbol{H}} = \boldsymbol{a}_t - \frac{2\boldsymbol{J}^T\boldsymbol{R}}{2\boldsymbol{J}^T\boldsymbol{J}} = \boldsymbol{a}_t - (\boldsymbol{J}^T\boldsymbol{J})^{-1}\boldsymbol{J}^T\boldsymbol{R}(\boldsymbol{a}_t).$$

This is the Gauss-Newton method, which requires J to be full rank [so $(J^TJ)^{-1}$ exists].

Many software packages have implemented this method. 900 As an example, let us use the following data

$$x: 0.10, 0.50, 1.0, 1.5, 2.0, 2.5$$

 $y: 0.10, 0.28, 0.40, 0.40, 0.37, 0.32$

to fit a model

$$y = \frac{x}{a + bx^2},$$

where a and b are the coefficients to be determined by the data. The objective is to minimize the sum of the residual squares

$$S = \sum_{i=1}^{6} R_i^2 = \sum_{i=1}^{6} \left[1 - \frac{x}{a + bx^2} \right]^2, \quad R_i = y_i - \frac{x_i}{a + bx_i^2}.$$

Since

$$\frac{\partial R_i}{\partial a} = \frac{x_i}{(a+bx^2)^2}, \quad \frac{\partial R_i}{\partial b} = \frac{x_i^3}{(a+bx^2)^2}.$$

If we use the initial guess a = 1 and b = 1, then the initial residuals are

$$\mathbf{R} = \begin{pmatrix} 0.0010 & -0.1200 & -0.1000 & -0.0615 & -0.0300 & -0.0248 \end{pmatrix}^T$$
.

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The initial Jacobian matrix is

$$\boldsymbol{J} = \left(\begin{array}{ccc} 0.0980 & 0.0010 \\ 0.3200 & 0.0800 \\ 0.2500 & 0.2500 \\ 0.1420 & 0.3195 \\ 0.0800 & 0.3200 \\ 0.0476 & 0.2973 \end{array} \right).$$

Thus, the first iteration using the Gauss-Newton algorithm gives

$$\left(\begin{array}{c} a \\ b \end{array}\right)_1 = \left(\begin{array}{c} 1 \\ 1 \end{array}\right) - (\boldsymbol{J}^T\boldsymbol{J})^{-1}\boldsymbol{J}\boldsymbol{R} = \left(\begin{array}{c} 1.3449 \\ 1.0317 \end{array}\right).$$

Then, by updating the new Jacobian and residuals, we have

$$a = 1.4742, \quad b = 1.0059,$$

after the second iteration. Similarly, we have

$$a = 1.4852, b = 1.0022$$
 (third iteration),

$$a = 1.4854, b = 1.0021$$
 (fourth iteration).

It converges quickly (parameters almost remain the same values after 10 iterations).

Levenberg-Marquardt Algorithm

The Levenberg-Marquardt algorithm is more robust by using a damping term in the approximation of the Hessian, that is

$$\boldsymbol{H} \approx 2[\boldsymbol{J}^T \boldsymbol{J} + \mu \boldsymbol{I}],$$

where $\mu>0$ is the damping coefficient, also called Marquardt parameter, and ${\pmb I}$ is the identity matrix of the same size as ${\pmb H}.$ Thus, the iteration formula becomes

$$m{a}_{t+1} = m{a}_t - rac{m{J}^T m{R}(m{a}_t)}{m{J}^T m{J} + \mu m{I}} = m{a}_t - (m{J}^T m{J} + \mu m{I})^{-1} m{J}^T m{R}(m{a}_t),$$

which is equivalent to the step size s_t in the line search given by

$$(\boldsymbol{J}^T\boldsymbol{J} + \mu \boldsymbol{I})\boldsymbol{s}_t = -\boldsymbol{J}^T\boldsymbol{R}(\boldsymbol{a}_t), \quad \boldsymbol{a}_{t+1} = \boldsymbol{a}_t + \boldsymbol{s}_t.$$

Mathematically speaking, a large μ will effectively reduce the step size (in comparison with those in the Gauss-Newton algorithm) and damps the moves with the right amount.

- If the reduction in E is sufficient, we can either keep this value of μ or reduce it. However, if the reduction E is not sufficient, we can increase μ .
- ullet If $\mu=0$, this algorithm reduces to the standard Gauss-Newton algorithm.
- Since $\mu \neq 0$, $(\boldsymbol{J}^T\boldsymbol{J} + \mu \boldsymbol{I})$ is always full rank (thus invertible).

Weighted Least Squares

The methods so far implicitly assume that all data points have equal variance σ^2 . To consider the possible different variances σ_i^2 , we have the weighted least squares.

Weighted Least Squares

Minimize
$$\sum_{i=1}^{n} w_i R_i^2 = \sum_{i=1}^{n} \frac{R_i^2}{\sigma_i^2} = ||\frac{R_i}{\sigma_i}||_2^2$$
, $R_i = y_i - f(x_i, \boldsymbol{a})$.

Here, $w_i = 1/\sigma_i^2$ and σ_i^2 is the variance associated with data point (x_i, y_i) . By defining a weight matrix W as

$$\mathbf{W} = \operatorname{diag}(w_i) = \begin{pmatrix} w_1 & 0 & \dots & 0 \\ 0 & w_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & w_n \end{pmatrix},$$

and following the similar derivations as above, steps/increments will become

$$\boldsymbol{a}_{t+1} = \boldsymbol{a}_t - (\boldsymbol{J}^T \boldsymbol{W} \boldsymbol{J})^{-1} (\boldsymbol{J}^T \boldsymbol{W} \boldsymbol{R}),$$

which is equivalent to approximating the Hessian by $\boldsymbol{H} = 2\boldsymbol{J}^T\boldsymbol{W}\boldsymbol{J}$ and the gradient by $\nabla E = 2\boldsymbol{J}^T\boldsymbol{W}\boldsymbol{R}$.

Regularization is another approach to deal with overfitting. In case of multi-variate cases with p components $\mathbf{Z}=(Z_1,Z_2,...,Z_p)^T$, we have the linear regression model

$$y = \beta_0 + \beta_1 Z_1 + \beta_2 Z_2 + ... + \beta_p Z_p = \mathbf{Z}^T \boldsymbol{\beta} + \beta_0,$$

where β_0 is the bias, while $\boldsymbol{\beta} = (\beta_1, \ \beta_2, ..., \beta_p)^T$ is the coefficient vector.

The standard method of least squares is to minimize the residual sum of squares (RSS)

Minimize
$$\sum_{i=1}^{n} (y_i - \boldsymbol{Z}_i^T \boldsymbol{\beta} - \beta_0)^2.$$

Ridge Regression

The Ridge regression uses a penalized RSS in the form

Minimize
$$\sum_{i=1}^{n} (y_i - \mathbf{Z}_i^T \boldsymbol{\beta} - \beta_0)^2 + \lambda \sum_{j=1}^{p} \beta_j^2,$$

where λ is the penalty coefficient. The above formula can be written compactly as

Minimize
$$||y - \beta_0 - \mathbf{Z}^T \boldsymbol{\beta}||_2^2 + \lambda ||\boldsymbol{\beta}||_2^2$$
,

where $||.||_2$ is the L_2 -norm, similar to Tikhonov regularization.

Here the bias β_0 is not part of the penalty or regularization term [we can always pre-process the data y_i (by subtracting their mean) so that the bias β_0 becomes zero].

Lasso Method

Lasso

The Lasso method uses an L_1 -norm in the regularization term

Minimize
$$||y - \beta_0 - \mathbf{Z}^T \boldsymbol{\beta}||_2^2 + \lambda ||\boldsymbol{\beta}||_1$$
,

which is equivalent to the following minimization problem:

Minimize
$$\sum_{i=1}^{n} (y_i - \beta_0 - \boldsymbol{Z}_i^T \boldsymbol{\beta})^2,$$

subject to $||\beta||_1 = |\beta_1| + |\beta_2| + ... + |\beta_p| \le \delta$ where $\delta > 0$ is a predefined hyper-parameter. [Again, β_0 is not penalized.]

Hybrid Regularization

A hybrid method is the elastic net regularization or regression, which combines the Ridge and Lasso methods into a hybrid as

Minimize
$$||y - \beta_0 - \mathbf{Z}^T \boldsymbol{\beta}||_2^2 + \lambda_1 ||\boldsymbol{\beta}||_1 + \lambda_2 ||\boldsymbol{b}||_2^2$$
.

Both the L_1 -norm and L_2 -norm are used with two hyper-parameters λ_1 and λ_2 .

These regularization methods have been implemented in many software packages.

References

References

- Xin-She Yang, Introduction to Algorithms for Data Mining and Machine Learning, Academic Press/Elsevier, (2019).
- Xin-She Yang, Optimization Techniques and Applications with Examples, John Wiley & Sons, (2018).

Notes on Software

- Matlab: Matlab has some well-tested curve-fitting tools fit and polyfit, nonlinear least squares lsqnonlin (including Levenberg-Marquardt method), generalized least squares fitnlm and lsqcurvefit, Lasso method lassglm, and many others.
- Octave: Octave has linear least squares lsqlin, exponential fit expfit, Levenberg-Marquardt nonlinear regression leqsqr, polynomial fit polyfitinf, and nonlinear least squares lsqnonlin.
- R: R has many functions for processing and visualizing data, including linear regression lm(), nonlinear least squares fit nls and many other statistical functionalities.
- Python: Python has at least two modules for regression statsmodels and Scikit-learn.
 The module Scikit learn sklearn is mainly for data mining and machine learning, which can do k-means clustering and basic clustering as well as various statistical analysis.

Any questions?

