# **Discrete Probability Distributions**

Modules 2.1 and 2.2 deal with general properties of random variables. Random variables with special probability distributions are encountered in different fields of *science* and *engineering*. Some specific **discrete probability distributions** are discussed in this module and some specific **continuous probability distributions** are discussed in the next module 2.5.

**Discrete Uniform Distribution:** A r.v.X is said to have a **discrete uniform distribution** over the range [1, n], if its p.m.f. is given by

$$p(x) = P(X = x) = \begin{cases} \frac{1}{n} &, & x = 1, 2, ..., n \\ 0 &, & \text{otherwise} \end{cases}$$

**Notation**:  $X \sim U(n)$  ,read as X follows discrete uniform distribution with parameter n.

**Note:** If all possible values of a r.v. are equally likely, then this distribution is used.

**Example 1:** If an unbiased coin is tossed once and X is equal to number of heads, then X = 0.1 and

$$P(X = 0) = P(X = 1) = \frac{1}{2}$$
 and  $X \sim U(2)$ .

**Example 2:** If an unbiased die is thrown once and X is equal to number on the die, then x=1,2,3,4,5,6 and  $P(X=i)=\frac{1}{6}$  for i=1,2,3,4,5,6 and  $X\sim U(6)$ .

**Mean and Variance:** We have  $E(X) = \frac{1}{n} \sum_{i=1}^{n} i = \frac{n+1}{2}$ 

and 
$$E(X^2) = \frac{1}{n} \sum_{i=1}^{n} i^2 = \frac{(n+1)(2n+1)}{6}$$

Thus 
$$V(X) = E(X^2) - (E(X))^2 = \frac{(n+1)(n-1)}{12}$$

**Bernoulli Experiment:** A random experiment whose outcomes are of two types, success (S) and failure (F), occurring with probabilities p and q = 1 - p respectively, is called a **Bernoulli experiment**.

Conducting a Bernoulli experiment once is known as **Bernoulli trial**. Note that p and q are same in each trial and outcomes of different trials are independent.

**Bernoulli distribution:** In a Bernoulli experiment, if a r.v.X is defined such that it takes value 1 with probability p when S occurs and 0 with probability q when F occurs, then we say that X follows Bernoulli distribution and its p.m.f. is given by

$$p(x) = P(X = x) = \begin{cases} p^x q^{1-x} & , & x = 0.1 \\ 0 & , & \text{otherwise} \end{cases}$$

### **Examples:**

- 1) Tossing of a coin (results a head or tail)
- 2) Performance of a student in an examination (results pass or failure)
- 3) Sex of an unborn child (results female or male)

#### Mean and Variance:

Mean = 
$$\mu = E(X) = 0 \times q + 1 \times p = p$$
  
and  $E(X^2) = 0^2 \times q + 1^2 \times p = p$   
$$\therefore \text{ Variance} = \sigma^2 = E(X^2) - \left(E(X)\right)^2 = p - p^2 = p(1-p) = pq$$

**Binomial Distribution:** Suppose we conduct n independent Bernoulli trials and we define

X = number of successes in n trials.

Then X is a discrete random variable and it takes the values 0,1,2,...,n.

**Derivation of** P(X = x)**:** Note that X = x means that there are x successes and (n - x) failures in n trials in a specified order (say) SSFSFFFS ... FSF.

Since outcomes of different trials are independent, by Multiplication Theorem, we have

$$P(SSFSFFFS ...FSF) = P(S) \cdot P(S) \cdot P(F) \cdot P(S) \cdot P(F) \cdot P(F) \cdot P(S) \cdot P(F) \cdot P(S) \cdot P(F) \cdot P(S) \cdot P(F)$$

$$= p p q p q q q p ... q p q$$

$$= \underbrace{p \cdot p \cdot ... \cdot p}_{(x \text{ times})} \cdot \underbrace{q \cdot q \cdot ... \cdot q}_{(n-x \text{ times})} = p^{x} q^{n-x}$$

But x successes in n trials can occur in  $\binom{n}{x}$  orders and the probability for each of these orders is same, viz.,  $p^xq^{n-x}$ . Hence by addition theorem of probability

$$p(x) = P(X = x) = \binom{n}{x} p^x q^{n-x}$$

**Definition:** A r.v. X is said to follow a **binomial distribution** with parameters n and p if its p.m.f. is given by

$$p(x) = P(X = x) = \begin{cases} \binom{n}{x} p^x q^{n-x} & , & x = 0,1,2,...,n, 0$$

**Notation:**  $X \sim B(n, p)$ . Read as X follows binomial distribution with parameters n and p.

## Real life examples:

- 1) Number of heads in n tosses of a coin
- 2) Number of boys in a family of n children
- 3) Number of times hitting a target in n attempts

#### Note:

1.

$$\sum_{x=0}^{n} p(x) = \sum_{x=0}^{n} {n \choose x} p^{x} q^{n-x} = (q+p)^{n} = 1$$

## 2. The c.d.f. of *X* is given by

$$F(x) = P(X \le x) = \sum_{k=0}^{x} {n \choose k} p^k q^{n-k}, x = 0,1,2,...,n$$

# Example 1: Four fair coins are tossed. If the outcomes are assumed to be independent, then find the p.m.f. and c.d.f. of the number of heads obtained.

**Solution:** Let *X* be the no. of heads in tossing 4 coins.

Then 
$$X \sim B\left(4, \frac{1}{2}\right)$$
 where  $p = P(head) = \frac{1}{2}$ .

Thus 
$$p(x) = P(X = x) = {4 \choose x} \left(\frac{1}{2}\right)^x \left(\frac{1}{2}\right)^{4-x}$$
  
=  ${4 \choose x} \left(\frac{1}{2}\right)^4 = {4 \choose x} \left(\frac{1}{16}\right)$  for  $x = 0,1,2,3,4$ .

Then 
$$p(0) = {4 \choose 0} \left(\frac{1}{16}\right) = \frac{1}{16}$$

$$p(1) = {4 \choose 1} \left(\frac{1}{16}\right) = \frac{4}{16}$$

$$p(2) = {4 \choose 2} {1 \choose 16} = {6 \over 16}$$

$$p(3) = {4 \choose 3} \left(\frac{1}{16}\right) = \frac{4}{16}$$

$$p(4) = {4 \choose 4} {1 \choose 16} = {1 \over 16}$$

The p.m.f p(x) and c.d.f F(x) are given in the following table.

x	0	1	2	3	4
p(x)	$\frac{1}{16}$	$\frac{4}{16}$	$\frac{6}{16}$	$\frac{4}{16}$	$\frac{1}{16}$
F(x)	$\frac{1}{16}$	$\frac{5}{16}$	$\frac{11}{16}$	$\frac{15}{16}$	1

Example 2: A and B play a game in which their chances of winning are in the ratio 3: 2. Find A's chance of winning at least three games out of the five games played.

#### **Solution:**

Define X = No. of games A winning out of 5.

Here 
$$p = P(A \ winning) = \frac{3}{5}$$
 and  $n = 5$  and  $X \sim B\left(5, \frac{3}{5}\right)$ . Thus,

$$p(x) = P(X = x) = {5 \choose x} \left(\frac{3}{5}\right)^x \left(\frac{2}{5}\right)^{5-x}$$
 for  $x = 0,1,...,5$ .

Required to find:

 $P(A \text{ winning at least 3 out of 5 games}) = P(X \ge 3)$ 

$$= P(X = 3) + P(X = 4) + P(X = 5)$$

$$= {5 \choose 3} {3 \choose 5}^3 {2 \choose 5}^2 + {5 \choose 4} {3 \choose 5}^4 {2 \choose 5}^1 + {5 \choose 5} {3 \choose 5}^5$$

$$= {3^3 \over 5^5} [10 \times 4 + 5 \times 3 \times 2 + 1 \times 9]$$

$$= {27 \times (40 + 30 + 9) \over 3125} = 0.68$$

Example 3: The probability of a man hitting a target is  $\frac{1}{4}$ .

- (i) If he fires 7 times, what is the probability of his hitting the target at least twice?
- (ii) How many times must he fire so that the probability of his hitting the target at least once is greater than  $\frac{2}{3}$ ?

**Solution:** Let X be the no. of times a man hitting the target in 7 fires. Here  $p = P(man\ hitting\ the\ target) = \frac{1}{4}$  and n = 7. Then  $X \sim B\left(7, \frac{1}{4}\right)$  and  $p(x) = {7 \choose x} \left(\frac{1}{4}\right)^x \left(\frac{3}{4}\right)^{7-x}$  for  $x = 0,1,2,\ldots,7$ .

(i) 
$$P(at \ least \ two \ hits) = P(X \ge 2) = 1 - [P(X = 0) + P(X = 1)]$$
  
=  $1 - [p(0) + p(1)]$   
=  $1 - \left\{ {7 \choose 0} \left(\frac{1}{4}\right)^0 \left(\frac{3}{4}\right)^7 + {7 \choose 1} \left(\frac{1}{4}\right) \left(\frac{3}{4}\right)^6 \right\} = \frac{4547}{8192} = 0.55$ 

(ii) Find 
$$n$$
 such that  $P(X \ge 1) > \frac{2}{3}$ 

$$\Rightarrow 1 - P(X = 0) > \frac{2}{3}$$

$$\Rightarrow -1 + P(X = 0) < -\frac{2}{3}$$

$$\Rightarrow P(X = 0) < \frac{1}{3}$$

$$\Rightarrow {n \choose 0} \left(\frac{1}{4}\right)^0 \left(\frac{3}{4}\right)^n < \frac{1}{3}$$

$$\Rightarrow {n \choose 4}^n < \frac{1}{3}$$

$$\Rightarrow n[\log 3 - \log 4] < \log 1 - \log 3$$

$$\Rightarrow n[\log 4 - \log 3] > \log 3$$

$$\Rightarrow n > \frac{\log 3}{\log 4 - \log 3} = 3.8$$
, since  $n$  cannot be fractional, the

required number of shots is 4.

#### **Mean of Binomial Distribution:**

$$\mu = E(X) = \sum_{x=0}^{n} x \, p(x) = \sum_{x=0}^{n} x \binom{n}{x} p^{x} q^{n-x}$$

$$= \sum_{x=1}^{n} x \left(\frac{n}{x}\right) \binom{n-1}{x-1} p^{x} q^{n-x} = np \sum_{x=1}^{n} \binom{n-1}{x-1} p^{x-1} q^{n-x}$$

$$= np(q+p)^{n-1} = np$$

$$\Rightarrow \mu = np$$

#### **Variance of Binomial Distribution:**

$$\sigma^2 = V(X) = npq$$
 (See  $P_1$  for proof)

# Example 4: One hundred balls are tossed into 50 boxes. What is the expected number of balls in the tenth box.

**Solution:** If we think of the balls tossed as Bernoulli trials in which a success is defined as getting a ball in the tenth box, then  $p = \frac{1}{50}$ . If X denotes the number of balls that go into the tenth box.

Then 
$$X \sim B\left(100, \frac{1}{50}\right)$$
 and  $E(X) = np = 100 \times \frac{1}{50} = 2$ .

Example 5: The mean and variance of binomial distribution are 4 and  $\frac{4}{3}$  respectively. Find  $P(X \ge 1)$ .

Solution: Here  $X \sim B(n, p)$ . But np = 4 and  $np(1-p) = \frac{4}{3}$ .

Hence 
$$4(1-p) = \frac{4}{3} \Longrightarrow 1-p = \frac{1}{3} \Longrightarrow p = \frac{2}{3}$$
 and  $n = \frac{4}{p} = 4 \times \frac{3}{2} = 6$ .

Thus,  $X \sim B\left(6, \frac{1}{3}\right)$  and hence  $P(X \ge 1) = 1 - P(X = 0)$ 

$$=1-\binom{6}{0}\binom{2}{3}^0\left(\frac{1}{3}\right)^6=1-\left(\frac{1}{3}\right)^6$$

#### **Poisson Distribution:**

If  $n \to \infty$  and  $p \to 0$  such that  $\lambda = np$  fixed, then  $\binom{n}{x} p^x (1-p)^{n-x} \to \frac{e^{-\lambda} \lambda^x}{x!}$  which is the p.m.f. of Poisson distribution (See  $P_2$ ). Thus the p.m.f. of Poisson distribution is obtained as the limit of p.m.f. of binomial distribution.

**Definition:** A r.v.X is said to follow a **Poisson distribution** with parameter  $\lambda$  if its p.m.f. is given by

$$p(x) = P(X = x) = \begin{cases} \frac{e^{-\lambda} \lambda^x}{x!} &, & x = 0,1,2,...; \lambda > 0\\ 0 &, & \text{otherwise} \end{cases}$$

**Notation:** Read  $X \sim P(\lambda)$  as: X follows poisson distribution with parameter  $\lambda$ .

### Real life examples

- 1) Number of defectives in a packet of 100 blades.
- 2) Number of telephone calls received at a particular telephone exchange in some unit of time.
- 3) Number of print mistakes in a page of a book.
- 4) The number of fragments received by a surface area A' from a fragment atom bomb.
- 5) The emission of radio active (alpha) particles.
- 6) Number of air accidents in some unit of time.

#### Note:

1. 
$$\sum_{x=0}^{\infty} p(x) = \sum_{x=0}^{\infty} \frac{e^{-\lambda} \lambda^{x}}{x!} = e^{-\lambda} \sum_{x=0}^{\infty} \frac{\lambda^{x}}{x!} = e^{-\lambda} e^{\lambda} = 1 \left( \because e^{x} = \sum_{i=0}^{\infty} \frac{x^{i}}{i!} \right)$$

2. The c.d.f. of *X* is given by

$$F(x) = P(X \le x) = \sum_{k=0}^{x} \frac{e^{-\lambda} \lambda^k}{k!}$$

Example 6: Messages arrive at a switchboard in a Poisson manner at an average rate of six per hour. Find the probability for each of the following events:

- a) Exactly two messages arrive within one hour.
- b) No message arrives within one hour.
- c) At least three messages arrive within one hour.

**Solution:** Let X be the r.v. that denotes the number of messages arriving at the switchboard within a one-hour interval. Then  $X \sim P(6)$  and its p.m.f is given by

$$P(X = x) = p(x) = \frac{e^{-6}6^x}{x!}$$
 for  $x = 0,1,2,...$ 

a) 
$$P(X = 2) = \frac{e^{-6}6^2}{2!} = \frac{36}{2}e^{-6} = 18e^{-6}$$
.

b) 
$$P(X = 0) = \frac{e^{-6}6^0}{0!} = e^{-6}$$
.

c) 
$$P(X \ge 3) = 1 - \{P(X = 0) + P(X = 1) + P(X = 2)\}$$
  
=  $1 - \{\frac{e^{-6}6^0}{0!} + \frac{e^{-6}6^1}{1!} + \frac{e^{-6}6^2}{2!}\} = 1 - e^{-6}\{1 + 6 + 18\} = 1 - 25e^{-6}$ 

Example 7: In a book of 520 pages, 390 typo-graphical errors occur. Assuming Poisson law for the number of errors per page, find the probability that a random sample of 5 pages will contain no error.

#### **Solution:**

 $\lambda = \text{Average number of typo-graphical errors/page} = \frac{390}{520} = 0.75$ 

Let X =Number of errors per page

Then 
$$X \sim P(0.75)$$
 and  $p(x) = P(X = x) = \frac{e^{-0.75}(0.75)^x}{x!}$ 

$$P(\text{No error}) = P(X = 0) = p(0) = e^{-0.75}$$

 $P(A \text{ random sample of 5 pages contain no error}) = [p(0)]^5 = [e^{-0.75}]^5 = e^{-3.75}$ 

Mean of Poisson Distribution: The mean of poisson distribution is given by

$$\mu = E(X) = \sum_{x=0}^{\infty} x \ p(x) = \sum_{x=1}^{\infty} \frac{x \ e^{-\lambda} \lambda^x}{x!} = \lambda e^{-\lambda} \sum_{x=1}^{\infty} \frac{\lambda^{x-1}}{(x-1)!} = \lambda e^{-\lambda} e^{\lambda} = \lambda$$

$$\Rightarrow \mu = \lambda$$

**Variance of Poisson Distribution:** 

$$E(X^{2}) = \sum_{x=0}^{\infty} x^{2} P(x) = \sum_{x=0}^{\infty} [x(x-1) + x] p(x)$$

$$= \sum_{x=2}^{\infty} x(x-1)p(x) + \sum_{x=1}^{\infty} xp(x) = \sum_{x=2}^{\infty} x(x-1)\frac{e^{-\lambda}\lambda^x}{x!} + \lambda$$
$$= e^{-\lambda}\lambda^2 \sum_{x=2}^{\infty} \frac{\lambda^{x-2}}{(x-2)!} + \lambda = e^{-\lambda}\lambda^2 e^{\lambda} + \lambda$$

$$\implies E(X^2) = \lambda^2 + \lambda$$

The variance of Poisson distribution is given by

$$V(X) = E(X^{2}) - (E(X))^{2} = \lambda^{2} + \lambda - \lambda^{2} = \lambda$$
$$\Rightarrow V(X) = \lambda$$

Note that for Poisson distribution, mean and variance are equal.

Example 8: If X and Y are independent Poisson variates such that

$$P(X=1) = P(X=2)$$
 and  $P(Y=2) = P(Y=3)$ , find the variance of  $X-2Y$ .

**Solution:** Let  $X \sim P(\lambda)$  and  $Y \sim P(\mu)$ . Then

$$P(X = x) = \frac{e^{-\lambda}\lambda^x}{x!}$$
 for  $x = 0,1,2,...$ ;  $\lambda > 0$  and

$$P(Y = y) = \frac{e^{-\mu}\mu^y}{y!}$$
 for  $y = 0,1,2,...; \mu > 0$ .

Since 
$$P(X = 1) = P(X = 2)$$
;  $\lambda e^{-\lambda} = \frac{e^{-\lambda} \lambda^2}{2} \Longrightarrow \lambda^2 - 2\lambda = 0$ 

$$\Rightarrow \lambda(\lambda - 2) = 0 \Rightarrow \lambda = 0, 2 \Rightarrow \lambda = 2 \ (\because \lambda = 0 \text{ is not admissible})$$

Since 
$$P(Y=2) = P(Y=3)$$
, then  $\frac{e^{-\mu}\mu^2}{2} = \frac{e^{-\mu}\mu^3}{6} \Longrightarrow \mu^3 - 3\mu^2 = 0$ 

$$\Rightarrow \mu^2(\mu - 3) = 0 \Rightarrow \mu = 0.3 \Rightarrow \mu = 3.$$

$$V(X-2Y) = 1^2 V(X) + (-2)^2 V(Y) = \lambda + 4\mu = 2 + 4 \times 3 = 2 + 12 = 14$$

## **Negative Binomial (or Pascal) Distribution:**

Let X denote the number of failures before the  $r^{th}$  success in a sequence of Bernoulli trials. Then the number of trials required is X + r.

## Derivation of P(X = x):

In x+r trials, the last trial must be a success whose probability is p. In the remaining (x+r-1) trials, we must have (r-1) successes whose probability is  $\binom{x+r-1}{r-1}p^{r-1}q^x$  (Using binomial distribution).

Thus, by multiplication theorem, we have

$$p(x) = P(X = x) = {x + r - 1 \choose r - 1} p^{r - 1} q^x \cdot p = {x + r - 1 \choose r - 1} p^r q^x$$

**Definition:** A random variable X is said to follow a **Negative binomial distribution**(NBD) with parameters r and p if its p.m.f is given by

$$p(x) = P(X = x) = \begin{cases} \binom{x+r-1}{r-1} p^r q^x & , & x = 0,1,2,... \\ 0 & , & \text{otherwise} \end{cases}$$

Notation:  $X \sim NB(r, p)$ .

Note:

1. 
$$\binom{x+r-1}{r-1} = \binom{x+r-1}{x}$$
  $\left(\because \binom{n}{r} = \binom{n}{n-r}\right)$ 

$$= \frac{(x+r-1)(x+r-2)...(r+1)r}{x!}$$

$$= (-1)^r \frac{(-r)(-r-1)...(-r-x+2)(-r-x+1)}{x!} = (-1)^x \binom{-r}{x}$$

Thus, the p.m.f. of NBD can be written as

$$p(x) = \begin{cases} \binom{-r}{x} p^r (-q)^x & , & x = 0,1,2,... \\ 0 & , & \text{otherwise} \end{cases}$$

Further, it is the  $(x+1)^{th}$  term in the expansion of  $p^r(1-q)^{-r}$ , a binomial expansion with negative index. Therefore, the distribution is known as negative binomial distribution.

2. 
$$\sum_{x=0}^{\infty} p(x) = p^r \sum_{x=0}^{\infty} {\binom{-r}{x}} (-q)^x = p^r (1-q)^{-r} = p^r p^{-r} = 1$$

Mean of NBD: The mean of NBD is given by

$$\mu = E(x) = \sum_{x=0}^{\infty} x \, p(x) = \sum_{x=1}^{\infty} x \, p(x) = \sum_{x=1}^{\infty} x {r \choose x} p^r (-q)^x$$

$$= p^r \sum_{x=1}^{\infty} x \left(\frac{-r}{x}\right) {r+1 \choose x-1} (-q)^x = (-r)(p^r)(-q) \sum_{x=1}^{\infty} {r-r+1 \choose x-1} (-q)^{x-1}$$

$$= (-r)(-q)p^r (1-q)^{-(r+1)}$$

$$= rqp^r p^{-(r+1)} = \frac{rq}{p}$$

Variance of NBD:  $\sigma^2 = \frac{rq}{p^2}$  (See  $P_3$  for proof)

- 3. Notice that  $\frac{\mu}{\sigma^2}=p<1$  and this implies that mean is smaller than variance in NBD.
- 4. If Y =Number of trials required to get  $r^{th}$  success, then Y = X + r and

$$P(Y = y) = P(X + r = y) = P(X = y - r) = {y - 1 \choose r - 1} p^r q^{y - r}$$
 for  $y = r, r + 1, \dots$  and  $E(y) = E(X) + r = \frac{rq}{p} + 1$  and  $V(Y) = V(X) = \frac{rq}{p^2}$ 

# Real life examples

- 1) Number of tails before the third head.
- 2) Number of girls before the second son.
- 3) Number of non-defectives before the third defective.

Example 9: Find the probability that there are two daughters before the second son in a family when probability of a son is 0.5.

**Solution:** Let X = the number of daughters before second son

Then 
$$P(X = x) = {x + 2 - 1 \choose 2 - 1} {1 \over 2}^2 {1 \over 2}^x = {x + 1 \choose 1} {1 \over 2}^{x+2}$$

and 
$$P(X = 2) = {3 \choose 1} \left(\frac{1}{2}\right)^4 = \frac{3}{16}$$

Example 10: An item is produced in large numbers. The machine is known to produce 5% defectives. A quality control inspector is examining the items by taking them at random. What is the probability that at least 4 items are to be examined in order to get 2 defectives?

Solution: Let Y= No. of items to be examined in order to get 2 defectives. Here p= (defective)= $\frac{5}{100}=0.05$ .

Then 
$$P(Y = y) = {y - 1 \choose 2 - 1} (0.05)^2 (0.95)^{y-2}$$

$$\Rightarrow P(Y = y) = (y - 1)(0.05)^{2}(0.95)^{y-2}$$

We want to find

$$P(Y \ge 4) = 1 - \sum_{y=2}^{3} P(Y = y) = 1 - \{P(Y = 2) + P(Y = 3)\}$$
$$= 1 - \{(0.05)^{2} + 2(0.05)^{2}(0.95)\} = 0.9928$$

#### Geometric distribution:

Let X denotes the number of failures before the first success in a sequence of Bernoulli trials. Then the required number of trials is X+1.

Geometric distribution is a particular case of negative binomial distribution with r=1.

**Definition:** A random variable X is said to follow a **Geometric distribution** (GD) with parameter p if its p.m.f. is given by

$$p(x) = P(X = x) = \begin{cases} p \cdot q^x & \text{for } x = 0,1,2, \dots \\ 0, & \text{otherwise} \end{cases}$$

Notation:  $X \sim GD(r)$ 

#### Mean and variance of GD

$$\mu = \frac{q}{p}$$
 and  $\sigma^2 = \frac{q}{p^2}$  (take  $r = 1$  in  $\mu$  and  $\sigma^2$  of NBD)

**Note:** Let  $Y = \text{Number of trials required to get first success, then <math>Y = X + 1$  and

$$P(Y = y) = P(X + 1 = y) = P(X = y - 1) = \begin{cases} p \ q^{y-1} & \text{for } y = 1,2,3, \dots \\ 0 & \text{otherwise} \end{cases}$$

Further, 
$$E(Y) = E(X) + 1 = \frac{1}{p}$$
 and  $V(Y) = V(X) = \frac{q}{p^2}$ .

### **Real life examples**

- 1) Number of tails before the third head
- 2) Number of girls before the second son
- 3) Number of non-defectives before the first defective

# Example 11: find the probability that there are two daughters before the first son in a family where probability of a son is 0.5.

**Solution:** Let X = Number of daughters before the first son.

Then 
$$P(X = x) = \left(\frac{1}{2}\right) \cdot \left(\frac{1}{2}\right)^x = \left(\frac{1}{2}\right)^{x+1}$$
 for  $x = 0,1,2,...$ 

and 
$$P(X = 2) = \left(\frac{1}{2}\right)^{2+1} = \left(\frac{1}{2}\right)^3 = \frac{1}{8}$$

# **Hyper geometric Distribution:**

Consider an urn with N balls, M of which are white and N-M are red. Suppose we draw a sample of n balls at random with replacement. Let X denote the

number of white balls in the sample. Then  $X \sim B(n,p)$  where  $p = \frac{M}{N}$  which remains same for all trials and outcomes of different trials are independent. The p.m.f of X is given by  $P(X = x) = \binom{n}{x} \left(\frac{M}{N}\right)^x \left(1 - \frac{M}{N}\right)^{n-x}$  for i = 1,2,...,n.

If the sample is selected without replacement, p is not same for all trials and outcomes of different trials are not independent and hence binomial distribution cannot be applied.

**Derivation of** P(X = x)**:** The number of all possible samples without replacement  $= \binom{N}{x}$ .

The number of samples in which there are x white balls and

$$(n-x)$$
 red balls  $= {M \choose x} {N-M \choose n-x}$ .

Thus 
$$p(x) = P(X = x) = \frac{\binom{M}{x}\binom{N-M}{n-x}}{\binom{N}{n}}$$
.

**Definition:** A random variable X is said to follow the **hyper geometric distribution** if its p.m.f is given by

$$p(x) = P(X = x) = \begin{cases} \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}} & , x = 0,1, \dots \min(n, M) \\ 0 & , otherwise \end{cases}$$

Example 12: A bag contains 4 white balls and 3 green balls. Three balls are drawn. What is the probability that 2 are white.

**Solution:** N = 4 + 3 = 7, M = 4, n = 3

X = Number of white balls and

$$P(X=2) = \frac{\binom{4}{2}\binom{3}{1}}{\binom{7}{3}} = \frac{4 \times 3 \times 3 \times 6}{2 \times 7 \times 6 \times 5} = \frac{18}{35}$$

Note: 
$$\sum_{x=0}^{\min(n,M)} p(x) = \begin{cases} \sum_{x=0}^{n} \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}} = \frac{\binom{N}{n}}{\binom{N}{n}} = 1 & if \min(n,M) = n \\ \sum_{x=0}^{n} \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{M}} = \frac{\binom{N}{m}}{\binom{N}{m}} = 1 & if \min(n,M) = M \end{cases}$$

# Mean and variance of Hyper geometric distribution:

The mean is given by  $\mu = \frac{nM}{N}$  and variance is given by  $\sigma^2 = \frac{NM(N-M)(N-n)}{N^2(N-1)}$  if  $\min(n,M) = n$  (For proof, see  $P_4$ ).