

Statistics for high dimensional data

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4 Lectures on Model Selection and high-dimensional statistical analysis

1. Multiple testing issues

- ▶ FWER: Family Wise Error Rate
- ▶ False Discovery Rate

2. Model selection and assessment

- ▶ Problematic, error
- ▶ Criteria for linear model : AIC, BIC, Cp
- ▶ Cross Validation and bootstrap method
- ▶ Variable selection: subset

3. Regularization Methods

- ▶ Ridge Regression
- ▶ Lasso method
- ▶ Elastic-net

4. Reduction method

- ▶ Principal Component Analysis
- ▶ Partial Least Square regression
- ▶ Sparse Methods
- ▶ Discriminant Analysis version

5. Handling missing data

- ▶ Framework, definition
- ▶ Method using maximisation of the likelihood: EM algorithm
- ▶ Imputation methods
- ▶ Method for and with PCA

Introduction

Definition (Little et Rubin [2019]) *Missing data are unobserved values that would be meaningful for analysis if observed ; in other words, a missing value hides a meaningful value.*

Some notations

$$\mathbf{X} = \begin{pmatrix} X_{11} & \dots & X_{1j} & \dots & X_{1p} \\ & & \dots & & \\ X_{i1} & \dots & X_{ij} & \dots & X_{ip} \\ & & \dots & & \\ X_{n1} & \dots & X_{nj} & \dots & X_{np} \end{pmatrix}$$

is associated with

$$\mathbf{R} = \begin{pmatrix} R_{11} & \dots & R_{1j} & \dots & R_{1p} \\ & & \dots & & \\ R_{i1} & \dots & R_{ij} & \dots & R_{ip} \\ & & \dots & & \\ R_{n1} & \dots & R_{nj} & \dots & R_{np} \end{pmatrix}$$

and $R_{ij} = 1$ if X_{ij} is observed, $R_{ij} = 0$ either.

Framework

- ▶ The whole data are (\mathbf{X}, \mathbf{R})
- ▶ $(\mathbf{X}^{\text{obs}}, \mathbf{R})$ are the observed data
- ▶ \mathbf{X}^{miss} are the missing data
- ▶ We say that data for subject i are full if $\prod_{j=1}^p R_{ij} = 1$
- ▶ Number of full subjects is ..

Missing data mechanisms

$f(\mathbf{X}, \mathbf{R})$ is the notation for the distribution of (\mathbf{X}, \mathbf{R})

Definition We say that the mechanism leading to missing data is MCAR (Missing Completely At Random) if

$$f(\mathbf{R}|\mathbf{X}) = f(\mathbf{R})$$

In this case, inference on the complete data is not biased

Definition We say that the mechanism leading to missing data is MAR (Missing At Random) if

$$f(\mathbf{R}|\mathbf{X}) = f(\mathbf{R}|\mathbf{X}^{\text{obs}})$$

In this case, distribution of the observed data differs from the one of the complete data

Main methods for inference in the context of missing data

- ▶ Drop subjects or variables
- ▶ In the context of a survey : ponderation of the responses
- ▶ Inference from the maximisation of the likelihood : EM algorithm
- ▶ Imputation of the missing data
- ▶ Imputation of the missing data for and with PCA
- ▶ Other possibilities
- ▶ Packages

Likelihood methods

Likelihood methods

We suppose the data are MAR and we want make inference on a parameter β . We can define

- ▶ The completed log-likelihood $L_c(\mathbf{X}, \beta) = \sum_{i=1}^n \ln(f(\mathbf{X}_i))$.
- ▶ The observed log-likelihood $L_o(\mathbf{X}^{\text{obs}}, \beta) = \sum_{i=1}^n \int \ln(f(\mathbf{X}_i, \beta)) d(\mathbf{X}_i^{\text{miss}})$
- ▶ We want to tackle $\hat{\beta} = \operatorname{argmax}_{\beta} L_o(\mathbf{X}^{\text{obs}}, \beta)$

Difficulties because the integral in L_o

EM algorithm, principle

Dempster, 1977. The idea is to take the conditional expectation of L_c given \mathbf{X}^{obs} and based on the assumption that distribution of \mathbf{X} is given by the parameter $\tilde{\beta}$:

- ▶ Aim : compute $\mathbb{E}_{\tilde{\beta}}[L_c(\mathbf{X}, \beta) | \mathbf{X}^{\text{obs}}]$
- ▶ EM algorithm alternates between:
 - ▶ Expectation E : Computation of $Q(\beta, \tilde{\beta}^{[r-1]}) = \mathbb{E}_{\tilde{\beta}^{[r-1]}}[L_c(\mathbf{X}, \beta) | \mathbf{X}^{\text{obs}}]$
 - ▶ Maximisation M : $\tilde{\beta}^{[r]} = \operatorname{argmax}_{\beta}(Q(\beta, \tilde{\beta}^{[r-1]}))$
- ▶ Properties : $Q(\beta, \tilde{\beta}^{[r]}) \geq Q(\beta, \tilde{\beta}^{[r-1]})$
- ▶ Convergence towards a local maximum, and to the global maxim in some cases
- ▶ Nice if the steps E and M are explicit

Exemple : EM algorithm for a linear model

Model

$$Y_i = \mathbf{X}_i^t \beta + \varepsilon_i, \quad \text{where } Y_i \in \mathbb{R}, \quad \beta = (\beta_1, \dots, \beta_p)$$

Assumption on $\mathbf{X} = (X_{ij})_{1 \leq i \leq n, 1 \leq j \leq p} : (\mathbf{X}_i = (X_{i1}, \dots, X_{ip}))_{1 \leq i \leq n} \in \mathbb{R}^p$ are i.i.d. and $\mathbf{X}_i \simeq \mathcal{N}_p(\mathbf{0}_p, \mathbf{Id}_p)$,

Assumption on $\varepsilon = (\varepsilon_1, \dots, \varepsilon_n)^t \in \mathbb{R}^n$ and $\varepsilon \simeq \mathcal{N}_n(\mathbf{0}_n, \mathbf{Id}_n)$.

We want to infer β despite of missing data on some X_{ij} with a ignorable mechanism trough an EM algorithm

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First task : write the log likelihood in the full observed case

$$L(\beta) = -\frac{1}{2} \sum_{i=1}^n (\beta^t \mathbf{X}_i \mathbf{X}_i^t \beta + \beta^t Y_i \mathbf{X}_i + Y_i^2)$$

whose solution is

$$\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y}$$
$$\hat{\beta} = \left(\sum_{i=1}^n \mathbf{X}_i \mathbf{X}_i^t \right)^{-1} * \left(\sum_{i=1}^n Y_i \mathbf{X}_i \right)$$

Conditionnal expectation of the log-likelihood

For a current β^c we have to take the conditional expectation of the complete log-likelihood

$$\mathbb{E}_{\beta^c}[L(\beta)|\mathbf{X}^{\text{obs}}, Y] = -\frac{1}{2} \sum_{i=1}^n \left(\beta^t \mathbb{E}_{\beta^c}[\mathbf{X}_i \mathbf{X}_i^t | \mathbf{X}_i^{\text{obs}}, Y_i] \beta + \beta^t Y_i \mathbb{E}_{\beta^c}[\mathbf{X}_i | \mathbf{X}_i^{\text{obs}}, Y_i] + Y_i^2 \right)$$

And to compute the β that maximise it :: *

$$\hat{\beta}^{c+1} = \left(\sum_{i=1}^n \mathbb{E}_{\beta^c}[\mathbf{X}_i \mathbf{X}_i^t | \mathbf{X}_i^{\text{obs}}, Y_i] \right)^{-1} * \left(\sum_{i=1}^n Y_i \mathbb{E}_{\beta^c}[\mathbf{X}_i | \mathbf{X}_i^{\text{obs}}, Y_i] \right)$$

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There are different kind of terms



$$\mathbb{E}_{\beta^c}[X_{ij} | \mathbf{X}_i^{\text{obs}}, Y_i] = \begin{cases} X_{ij} & \text{if } R_{ij} = 1 \\ \frac{(Y_i - \sum_k \beta_k^c R_{ik} X_{ik}) \beta_j^c}{\sum_{k=1}^p (1 - R_{ik})(\beta_k^c)^2 + \sigma^2} & \text{if } R_{ij} = 0 \end{cases}$$



$$\mathbb{E}_{\beta^c}[X_{ij} X_{ij'} | \mathbf{X}_i^{\text{obs}}, Y_i] = \begin{cases} X_{ij} X_{ij'} & \text{if } R_{ij} R_{ij'} = 1 \\ X_{ij} \mathbb{E}_{\beta^c}[X_{ij'} | \mathbf{X}_i^{\text{obs}}, Y_i] & \text{if } (R_{ij}, R_{ij'}) = (1, 0) \\ -\frac{\beta_j^c \beta_{j'}^c}{\sum_{k=1}^p (1 - R_{ik})(\beta_k^c)^2 + \sigma^2} & \text{if } (R_{ij}, R_{ij'}) = (0, 0) \text{ and } j \neq j' \\ 1 - \frac{(\beta_j^c)^2}{\sum_{k=1}^p (1 - R_{ik})(\beta_k^c)^2 + \sigma^2} & \text{if } (R_{ij}, R_{ij'}) = (0, 0) \text{ and } j = j' \end{cases}$$

EM algorithm

Algorithm

- ▶ Step 0
 - ▶ Initialise the missing values by their expectation (given by the model : 0) \mathbf{X}_0
 - ▶ Compute $\beta^0 = (\mathbf{X}_0' \mathbf{X}_0)^{-1} \mathbf{X}_0' \mathbf{Y}$
- ▶ Step k+1
 - ▶ Step E, calculus of terms of $Q(\beta, \beta^k) = \mathbb{E}_{\beta^k}[L(\beta) | \mathbf{X}^{\text{obs}}, \mathbf{Y}]$
 - ▶ Step M

$$\hat{\beta}^{k+1} = \left(\sum_{i=1}^n \mathbb{E}_{\beta^k} [\mathbf{X}_i \mathbf{X}_i^t | \mathbf{X}_i^{\text{obs}}, Y_i] \right)^{-1} * \left(\sum_{i=1}^n Y_i \mathbb{E}_{\beta^k} [\mathbf{X}_i | \mathbf{X}_i^{\text{obs}}, Y_i] \right)$$

$\hat{\beta}^{k+1}$ converges towards $\hat{\beta}$

Imputation methods

First Ideas

Imputation by mean

- ▶ X^j quantitative : missing X_{ij} imputed by mean on observed values
$$\frac{1}{\sum_{i=1}^n R_{ij}} \sum_{i=1}^n R_{ij} X_{ij}$$
- ▶ X^j qualitative : missing X_{ij} imputed by the mode of observed $(X_{ij})_{1 \leq i \leq n}$
- ▶ X^j qualitative, you can also consider the NA as a level of X^j

Drawback

Does not take other variables into account

Exemple

Model $\mathbf{X}_i \simeq \mathcal{N}_2(0, \Gamma)$ with variances 1 and correlation ρ .

Observation : X^1 is observed but $P(R_{i2} = 0 | X_{i1}) = (1 + e^{1-X_{i1}})^{-1}$

Exemple

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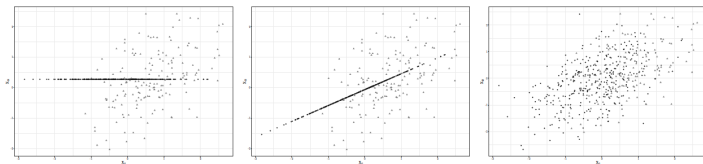


Figure 2.2 : Jeu de données avec trois imputations différentes : imputation par l'espérance (gauche), imputation par l'espérance conditionnelle (centre), imputation stochastique par la loi conditionnelle (droite). Les données complètes sont en gris et les données ayant été imputées en noir.

What about the mean ? the correlation between X^1 and X^2 according to the imputation method ?

Imputation methods by neighbours method

k-N-N imputation

Algorithm

- ▶ Suppose that X^1 is partially observed but all the other $(p - 1)$ variables are observed
- ▶ choose a integer k
- ▶ If $R_{i1} = 0$, compute the distance between the vector $\mathbf{X}_i^{-1} = (X_{i2}, \dots, X_{ip})$ and all the $\mathbf{X}_{i'}^{-1}$ with $R_{i'1} = 1$
- ▶ Select the k -nearest neighbours of $\mathbf{X}_{i_1}^{(-1)}, \dots, \mathbf{X}_{i_k}^{(-1)}$ of $\mathbf{X}_i^{(-1)}$
- ▶ Input X_{i1} by the mean $\frac{1}{k}(\mathbf{X}_{i_1}^{(-1)} + \dots + \mathbf{X}_{i_k}^{(-1)})$

Algorithm

- ▶ Suppose that X^1 is partially observed but all the other $p - 1$ variables are observed
- ▶ Regress $(X^1)^{\text{obs}}$ on the $p - 1$ other corresponding variables by the model $(X^1)^{\text{obs}} = \mathbf{X}^{(-1), \text{obs}} \beta + \varepsilon$, compute $\hat{\beta}$
- ▶ Imput X_{i1} by the $\hat{X}_{i1} = \mathbf{X}_i^{(-1)} \hat{\beta}$

Some times, it is only a "local regression" : evaluation of β is done on the k-NN of \mathbf{X}_i^{-1}

Random forest Imputation

Algorithm

- ▶ Suppose that X^1 is partially observed but all the other $p - 1$ variables are observed
- ▶ Fit $(X^1)^{\text{obs}}$ on the $p - 1$ other corresponding variables to have $(x^1)^{\text{obs}} \simeq RF(\mathbf{x}^{(-1), \text{obs}})$
- ▶ Imput x_{i1} by $\hat{x}_{i1} \simeq RF(\mathbf{x}_i^{(-1)})$

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Algorithm

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- ▶ Imput x_{i1} by $\hat{x}_{i1} \simeq RF(\mathbf{x}_i^{(-1)})$

Miss Forest Algorithm

- ▶ Step 0: Complete X^1 by a simple method
- ▶ (while criterion), do
 - ▶ New matrix X replaced by imputed matrix
 - ▶ Fit $(X^1)^{\text{imput}}$ with $\mathbf{X}^{(-1)}$ by RF
 - ▶ If $R_{i1} = 1$, imput X_{i1} by $\hat{X}_{i1} \simeq RF(\mathbf{X}_i^{(-1)})$
- ▶ evaluate criterion

package Miss Forest

Imputation methods by PCA

Multiple imputation

- ▶ For instance, in regression add the noise to reduce the over-fitting
- ▶ Do it several times and infer parameters
- ▶ Give an idea of the variability

R packages

- ▶ VIM to visualize the missing data
- ▶ mice
- ▶ Miss Forest
- ▶ missMDA based on FactoMineR

References



Wiki stat Toulouse

<https://www.math.univ-toulouse.fr/besse/Wikistat/pdf/st-m-app-idm.pdf>