Tutorial 4 sample answer

Question 1

```
a)
     Command/code:
        TestScore <- read.csv(file.choose())</pre>
        class A <- subset(TestScore, Class=="A")</pre>
        class B <- subset(TestScore, Class=="B")</pre>
        t.test(x=class A$Score, mu=70, alternative="greater")
     Output:
           One Sample t-test
        data: class A$Score
        t = 2.4528, df = 4, p-value = 0.03512
        alternative hypothesis: true mean is greater than 70
        95 percent confidence interval:
         71.96258
                          Inf
        sample estimates:
        mean of x
                 85
     Comment:
     In this hypothesis test, the following hypotheses are used,
                            H_0: \mu_A \le 70 vs H_1: \mu_A > 70
      where \mu_A is the mean score for class A. The R output shows that the test statistic is
     2.4528 and p-value for the test is 0.03512. Since the p-value is less than \alpha (0.05),
      we have enough evidence to reject the null hypothesis, and support the alternative
     hypothesis. We conclude that the mean score for class A is greater than 70.
b)
     Command/code:
        var.test(x=class A$Score, y=class B$Score)
     Output:
           F test to compare two variances
        data: class A$Score and class B$Score
        F = 2.2722, num df = 4, denom df = 4,
        p-value = 0.4462
        alternative hypothesis: true ratio of variances is
        not equal to 1
        95 percent confidence interval:
           0.2365733 21.8231724
        sample estimates:
        ratio of variances
                    2.272175
```

Comment:

The hypotheses are:

$$H_0$$
: $\sigma_A^2 = \sigma_B^2$ vs H_1 : $\sigma_A^2 \neq \sigma_B^2$

where σ_A^2 is the variance score for class A, and σ_B^2 is the variance score for class B. From the R output, the *p*-value is 0.4462, which is greater than the significance level $\alpha = 0.05$. Therefore there is not enough evidence to reject the null hypothesis. We conclude that the two variances are equal.

c) Command/code:

t.test(x=class_A\$GPA, y=class_B\$GPA, var.equal=TRUE, alternative="two.sided")

Output:

Two Sample t-test

Comment:

The hypotheses are:

$$H_0$$
: $\mu_A = \mu_B \text{ vs } H_1$: $\mu_A \neq \mu_B$

where μ_A is the mean GPA for class A, and μ_B is the mean GPA for class B. From the R output, the *p*-value is 0.2262, which is greater than the significance level α . Therefore there is not enough evidence to reject the null hypothesis. We conclude that the mean GPAs of the two classes are equal.

d) Command/code:

summary(aov(GPA ~ Class, data=TestScore))

Output:

Df Sum Sq Mean Sq F value Pr(>F)
Class 1 0.729 0.729 1.719 0.226
Residuals 8 3.392 0.424

Comment:

The hypotheses are:

$$H_0$$
: $\mu_A = \mu_B$ vs H_1 : $\mu_A \neq \mu_B$

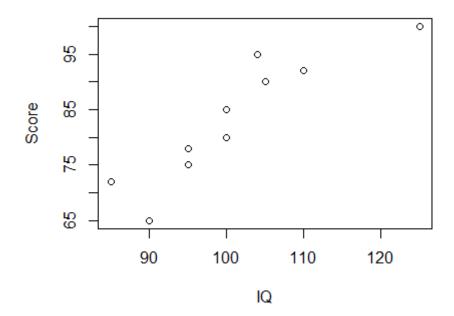
where μ_A is the mean GPA for class A, and μ_B is the mean GPA for class B. From the R output, the *p*-value is 0.2262, which is greater than the significance level α . Therefore there is not enough evidence to reject the null hypothesis. We conclude that the mean GPAs of the two classes are equal.

The test using ANOVA gives exactly the same *p*-value as the t-test in (d). ANOVA for two groups are exactly equal as t-test with equal variance. However, ANOVA allows for comparison of mean for more than two groups.

e) Command/code:

```
plot(Score ~ IQ, data=TestScore)
model <- lm(Score ~ IQ, data=TestScore)
summary(model)</pre>
```

Output:



Call:

lm(formula = Score ~ IQ, data = TestScore)

Residuals:

Min 1Q Median 3Q Max -8.5190 -2.8200 0.3789 2.8412 9.0467

Coefficients:

freedom

Residual standard error: 5.195 on 8 degrees of

Multiple R-squared: 0.8054, Adjusted R-squared: 0.7811

F-statistic: 33.11 on 1 and 8 DF, p-value: 0.000427

Comment:

The fitted regression line is:

$$Score = -6.4157 + 0.8882 \times IQ$$

In this case, the *p*-value for variable IQ is 0.00427, which is smaller than α . Therefore, it can be concluded that the parameter related to IQ is non-zero, and that Score and IQ have a significant linear relationship.

Note: The *p*-value is for testing $\beta_1 = 0$ vs $\beta_1 \neq 0$ where β_1 in this case is the parameter for IQ, which is estimated as 0.8882.

Question 2

```
Command/code:
a)
       cement <- read.csv(file.choose())</pre>
       model1 \leftarrow lm(y \sim x1 + x2 + x3 + x4, data=cement)
       summary(model1)
     Output:
       Call:
       lm(formula = y \sim x1 + x2 + x3 + x4, data = cement)
       Residuals:
                     1Q Median 3Q
            Min
                                             Max
       -3.1750 -1.6709 0.2508 1.3783 3.9254
       Coefficients:
                    Estimate Std. Error t value Pr(>|t|)
        (Intercept) 62.4054 70.0710 0.891
                                                     0.3991
                      1.5511 0.7448 2.083
       x1
                                                   0.0708 .
       x2
                      0.5102
                                 0.7238 0.705 0.5009
                     0.1019 0.7547 0.135
-0.1441 0.7091 -0.203
                                 0.7547 0.135
       xЗ
                                                    0.8959
       x4
                                                    0.8441
       Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.'
       0.1 ' ' 1
       Residual standard error: 2.446 on 8 degrees of
       freedom
       Multiple R-squared: 0.9824, Adjusted R-squared:
       0.9736
       F-statistic: 111.5 on 4 and 8 DF, p-value: 4.756e-07
     Based on the p-values, only the variable x_1 has p-value less than \alpha = 0.1.
     Therefore, only x_1 seems to be significant to the model.
```

```
b)
     Command/code:
        model2 <- lm(y \sim x1 + x2, data=cement)
        summary(model2)
     Output:
        Call:
        lm(formula = y \sim x1 + x2, data = cement)
        Residuals:
                    10 Median 30
           Min
                                           Max
        -2.893 -1.574 -1.302 1.363 4.048
        Coefficients:
                      Estimate Std. Error t value Pr(>|t|)
        (Intercept) 52.57735 2.28617 23.00 5.46e-10 ***
                       1.46831 0.12130 12.11 2.69e-07 ***
        x1
                       0.66225
        x2
                                   0.04585 14.44 5.03e-08 ***
        Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.'
        0.1 ' ' 1
        Residual standard error: 2.406 on 10 degrees of
        freedom
        Multiple R-squared: 0.9787, Adjusted R-squared:
        0.9744
        F-statistic: 229.5 on 2 and 10 DF, p-value: 4.407e-
        09
     Comment:
     After removing the variable x_3 and x_4, the p-values for x_1 and x_2 are now very
     small. Therefore, both of these variables (x_1 \text{ and } x_2) are significant to the model.
c)
     The fitted regression line is:
                         \hat{y} = 52.58 + 1.468x_1 + 0.6623x_2
     In this case, x_1 = 8, x_2 = 50, and
d)
                     \hat{y} = 52.58 + 1.468(8) + 0.6623(50) = 97.44
     The estimated heat evolved is 97.44 calories.
```

```
Command:
a)
        property sales <- read.csv(file.choose())</pre>
        cheval sale <- subset(property sales,
                          Neighbourhood=="Cheval") $Sales
        hydepark sale <- subset(property sales,
                            Neighbourhood=="HydePark") $Sales
        var.test(cheval sale, hydepark sale)
     Output:
        > var.test(cheval sale, hydepark sale)
           F test to compare two variances
        data: cheval sale and hydepark sale
        F = 0.24531, num df = 43, denom df = 33,
        p-value = 2.217e-05
        alternative hypothesis: true ratio of variances is
        not equal to 1
        95 percent confidence interval:
         0.1260855 0.4640299
        sample estimates:
        ratio of variances
                   0.2453138
      Comment:
      The hypotheses tested are:
                   H_0: \sigma_{cheval}^2 = \sigma_{hydepark}^2 vs H_1: \sigma_{cheval}^2 \neq \sigma_{hydepark}^2
      where \sigma_{cheval}^2 is the variance for sales in Cheval and \sigma_{hydepark}^2 is the variance for
      sales in Hyde Park.
     Since the p-value is very small (2.2 \times 10^{-5}), we have strong evidence to reject the
     null hypothesis. We conclude that the two variances are not equal
b)
     Command:
        t.test(cheval sale, hydepark sale,
                 alternative="less", var.equal=FALSE)
      Output:
        > t.test(cheval sale, hydepark sale,
        alternative="less", var.equal=FALSE)
           Welch Two Sample t-test
        data: cheval sale and hydepark sale
        t = -2.4857, df = 45.444, p-value = 0.008335
        alternative hypothesis: true difference in means is
        less than 0
        95 percent confidence interval:
```

```
-Inf -68.12958 sample estimates: mean of x mean of y 455.3955 665.3382
```

Comment:

The hypotheses tested are:

 H_0 : $\mu_{cheval} \ge \mu_{hydepark}$ vs H_1 : $\mu_{cheval} < \mu_{hydepark}$

where μ_{cheval} is the mean for sales in Cheval and $\mu_{hydepark}$ is the mean for sales in Hyde Park.

Since the p-value is smaller than $\alpha = 0.05$, we reject the null hypothesis. Therefore, we have evidence to support that the mean for sales in Cheval is smaller than the mean for sales in Hyde Park.

Note that the "var.equal=FALSE" in the argument is not needed, as it is the default value for the t-test function.

c) Command:

Output:

Comment:

The null hypothesis, H_0 states that the mean land values are equal for all four neighbourhoods, whereas the alternative hypothesis states that at least two means are different.

The p-value from the ANOVA is very small. Therefore we reject the null hypothesis and conclude that at least a pair of neighbourhood have different mean land values.

d) Command:

Output:

Residuals 172 6591254 38321

Comment:

The null hypothesis, H_0 states that the mean improvement values are equal for all four neighbourhoods, whereas the alternative hypothesis states that at least two means are different.

The p-value from the ANOVA is large, larger than $\alpha = 0.05$. Therefore we do not have enough evidence to reject the null hypothesis and conclude that the mean improvement values for all four neighbourhoods are equal.

e) Command:

```
fit <- lm(Sales~Land.value+Improvement.value,
data=property_sales)
summary(fit)</pre>
```

Output:

```
> summary(fit)
```

Call:

lm(formula = Sales ~ Land.value + Improvement.value,
data = property sales)

Residuals:

```
Min 1Q Median 3Q Max -382.46 -49.30 0.79 40.42 540.15
```

Coefficients:

```
Signif. codes:
```

```
0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Residual standard error: 112.9 on 173 degrees of freedom
```

Multiple R-squared: 0.9242, Adjusted R-squared: 0.9233

F-statistic: 1055 on 2 and 173 DF, p-value: < 2.2e-16

Comment:

The fitted regression line is:

 $Sales = -16.176 + 1.393 \times Land. value + 1.330 \times Improvement. value$

The land value is significant in the regression model since the p-value for the parameter related to the land value is very small.

If the land value is \$100 000 and the improvement value is \$200 000, the estimated sales price is

 $Sales = -16.176 + 1.393 \times 100 + 1.330 \times 200 = 389.124 = 389.124