

# Midterm 1

● Graded

Student

Sampad Mohanty

Total Points

52 / 60 pts

Question 1

Problem 1

■ 18 / 20 pts

✓ + 8 pts (a) is Correct

✓ + 7 pts (b) is correct

+ 5 pts (c) is correct

+ 5 pts (a) The factorization theorem was applied, but the joint density or the final answer was incorrect.

+ 5 pts (b) Most of the steps are correct. Conclusion is incorrect.

✓ + 3 pts (c) Used the correct definition, but there is a gap in the proof.

+ 4 pts (c) checked  $f(x|\theta)$  is an exponential family, but did not mention the open set condition.

+ 1 pt (c) state the definition correctly

+ 4 pts Checked exponential family and open set condition. Only the conclusion is incorrect

1

this needs justification.  $g(t)$  can take both positive and negative values

## Question 2

### Problem 2

14 / 20 pts

✓ + 5 pts (a) is correct

+ 4 pts (a) partially correct

+ 2 pts (a) Computed the expectation correctly

✓ + 7 pts (b) is correct

+ 2 pts (b) Use the definition of MLE correctly

+ 8 pts (c) is correct

+ 1 pt (c) Only write down the formula for MSE

✓ + 2 pts (c) computed the MSE but did not find the correct distribution of  $X_{(n)}$ .

+ 4 pts c) Show correctly that MLE is biased

+ 6 pts c) Show correctly that MLE is biased but the variance computation is only partially correct

+ 7 pts minor error in (c)

2 you need to know the pdf of  $X_{(n)}$  first

## Question 3

### Problem 3

20 / 20 pts

✓ + 8 pts (a) is Correct

✓ + 7 pts (b) is correct

✓ + 5 pts (c) is correct

+ 2 pts (a) partial credits for  $L(\theta)$ .

+ 6 pts (a) Write the MLE as a solution to an optimization question correctly. but didn't write down the explicit answer, or the final answer is incorrect

+ 4 pts (b) mostly correct for the posterior distribution

+ 6 pts (b) minor error

+ 2 pts (b) partial credits for the posterior density calculation

+ 2 pts (c) Write down the equation for MAP and posterior mean correctly

+ 4 pts (c) Partially correct for the final expression,

+ 0 pts Incorrect

Math 541 A - Spring 2025

Midterm 1

Feb 26, 2025

11:00-11:50 AM

Name: SAMPAD MONDAL

Student ID Number: 5679528312

- There are 3 problems in total. Make sure your exam contains all these questions.
- You are allowed to use one page of double-sided **hand-written** 8.5 by 11 inch notes.
- You must show your work on all problems. The correct answer with no supporting work may result in no credit.
- If you need more room, use the backs of the pages and indicate to the grader that you have done so.

Problem 1	20	
Problem 2	20	
Problem 3	20	
Total	60	

Name:

ID:

1. (20 pts) Suppose  $X_1, \dots, X_n$  are i.i.d. random variables with common pdf

$$f(x|\theta) = \begin{cases} (\theta+1)x^\theta, & x \in (0,1), \\ 0 & \text{otherwise,} \end{cases} \quad \text{where } \theta > -1.$$

- (a) (8 pts) Use the Factorization Theorem to find a 1-dimensional sufficient statistic  $T(X_1, \dots, X_n)$  for  $\theta$ .
- (b) (7 pts) Find a minimal sufficient statistic for  $\theta$ . Justify your answer.
- (c) (5 pts) Find a complete statistic for  $\theta$ . Justify your answer.

$$f(x_1, x_2, \dots, x_n | \theta) = \prod_{i=1}^n (\theta+1)x_i^\theta \mathbb{I}_{[x_i \in (0,1)]} = (\theta+1)^n \left[ \prod_{i=1}^n x_i \right]^\theta \mathbb{I}_{\{\max x_i < 1, \min x_i > 0\}}$$

define  $T(x) = \prod_{i=1}^n x_i \rightarrow f(\bar{x}|\theta) = (\theta+1)^n T^\theta \mathbb{I}_{\{\max x_i < 1\}} \mathbb{I}_{\{\min x_i > 0\}}$

By factorization thm;  $T(x) = \prod_{i=1}^n x_i$  is sufficient for  $\theta$ .

I claim  $T(x) = \prod_{i=1}^n x_i$  is sufficient & minimal too.

By Lehman-Schiffé,  $\frac{f(\bar{x}|\theta)}{f(\bar{y}|\theta)} = \frac{(\theta+1)^n T(\bar{x})^\theta \mathbb{I}_{\{\max x_i < 1\}} \mathbb{I}_{\{\min x_i > 0\}}}{(\theta+1)^n T(\bar{y})^\theta \mathbb{I}_{\{\max y_i < 1\}} \mathbb{I}_{\{\min y_i > 0\}}}$  is independent of  $\theta \Leftrightarrow T(\bar{x}) = T(\bar{y})$

$\therefore T(x)$  is minimal & sufficient statistic.

) claim  $\rightarrow T(x) = \prod_{i=1}^n x_i$  is complete.

$$\forall \theta, g \quad E_\theta [g(T)] = 0 \Rightarrow P_\theta(g(T) = 0) = 1.$$

$$E_\theta [g(\prod_{i=1}^n x_i)] = \int_0^1 g(\prod_{i=1}^n x_i) (\theta+1) x^\theta dx = 0.$$

(1) is always

this has to be zero to make the area under curve = 0.

$\therefore T(x) = \prod_{i=1}^n x_i$  is complete for  $\theta$ .

Name:

ID:

2. (20 pts) Suppose  $X_1, \dots, X_n$  are i.i.d. from a distribution given by

$$f(x|\theta) = \frac{1}{\theta} 1\{0 \leq x \leq \theta\}, \quad \theta > 0.$$

(a) (5 pts) Find a method of moment estimator for  $\theta$ .(b) (7 pts) Determine the Maximum Likelihood Estimator  $\hat{\theta}_{MLE}$ . Justify your answer.(c) (8 pts) Show that  $\hat{\theta}_{MLE}$  is a biased estimator for  $\theta$  and find the mean squared error of  $\hat{\theta}_{MLE}$ .

a) This is uniform  $U(0, \theta)$ . mean =  $\theta/2$ , variance =  $\int_0^\theta (x - \theta/2)^2 \frac{1}{\theta} dx = \int_0^\theta \frac{x^2 + \frac{\theta^2}{4} - \theta x}{\theta} dx = \frac{1}{\theta} \left( \frac{\theta^3}{3} + \frac{\theta^3}{4} - \frac{\theta^3}{2} \right) = \theta^2 \left( \frac{1}{3} + \frac{1}{4} - \frac{1}{2} \right) = \theta^2 \left( \frac{4+3-6}{12} \right) = \theta^2/12$ .

Set mean = sample mean  $\Rightarrow \frac{\sum x_i}{n} = \theta/2 \Rightarrow \theta = \frac{2 \sum x_i}{n}$

2nd moment =  $\int_0^\theta x^2 f(x|\theta) dx = \int_0^\theta x^2 \frac{1}{\theta} dx = \frac{x^3}{3} \Big|_0^\theta \times \frac{1}{\theta} = \frac{\theta^3}{3} \times \frac{1}{\theta} = \theta^2/3$

ignore this.

Set 2nd moment = empirical second moment.

$\Rightarrow \frac{\theta^2}{3} = \frac{\sum x_i^2}{n} \Rightarrow \theta^2 = \frac{3}{n} \sum x_i^2 \Rightarrow \theta = \frac{1}{\sqrt{3}} \sqrt{\frac{\sum x_i^2}{n}}$

b)  $f(\vec{x}|\theta) = \prod_{i=1}^n f(x_i|\theta) = \frac{1}{\theta^n} \mathbb{I}_{\{x_i \in [0, \theta]\}} , \theta > 0$

 $L(\theta|\vec{x})$ 

We want to maximize  $L(\theta|\vec{x})$ . But  $\frac{1}{\theta^n}$  is monotonic decreasing for  $\theta > 0$ .

$\therefore$  It achieves maximum when  $\theta$  is lowest. But  $\theta > \max_i x_i$  as otherwise we get  $L(\theta|\vec{x}) = 0$  which is definitely not maximum  $\frac{1}{\theta^n}$  but a minimum as  $L(\theta|\vec{x}) > 0$ .

$\therefore$  we set  $\hat{\theta}_{MLE} = \max_i x_i$  to maximize  $L(\theta|\vec{x})$ .

c)  $E[\hat{\theta}_{MLE}] \neq \theta/2$ . Hence  $\hat{\theta}_{MLE}$  is biased.

$E[\hat{\theta}_{MLE}] = \frac{n-1}{n} \theta$

$E[(\hat{\theta} - \theta)^2] = E\left(\theta \frac{n-1}{n} - \theta\right)^2 = E\left(\theta^2 \left(\frac{n-1}{n} - 1\right)^2\right)$

$= \int_0^\theta \frac{\theta^2}{n^2} d\theta = \frac{1}{n^2} \int_0^\theta \theta^2 d\theta = \frac{1}{n^2} \left[ \frac{\theta^3}{3} \right]_0^\theta = \frac{\theta^3}{3n^2}$

Name:

ID:

3. (20 pts) Consider a 1-dimensional linear regression model where  $y_i \in \mathbb{R}$ ,  $1 \leq i \leq n$  are independent random variables satisfying  $y_i = \theta x_i + \varepsilon_i$ , where  $\varepsilon_i \sim N(0, 1)$  are i.i.d. random variables,  $x_i \in \mathbb{R}$ , and  $\theta \in \mathbb{R}$  is an unknown parameter.

- (a) (8 pts) Assume  $\sum_{i=1}^n x_i^2 > 0$ . Find the likelihood function  $L(\theta)$  given the data  $\{(x_i, y_i)\}_{i=1}^n$  and compute the maximum likelihood estimator of  $\theta$  explicitly. Justify your answer.
- (b) (7 pts) Assume the prior distribution of  $\theta$  is given by  $\theta \sim N(\theta_0, 1)$  where  $\theta_0 \in \mathbb{R}$ . Show that the posterior distribution for  $\theta$  is a normal distribution. Justify your answer.
- (c) (5 pts) Find the posterior mean estimator and the maximum a posteriori estimation estimator for  $\theta$ . Justify your answer.

a)  $L(\theta|x) = L(\theta) = \prod_{i=1}^n f(y_i | \theta) = \prod_{i=1}^n f(\varepsilon_i | \theta)$   
 $\varepsilon_i = y_i - \theta x_i \parallel f(\varepsilon_i | \theta) = f(y_i - \theta x_i | \theta)$   
 $\therefore L(\theta) = \prod_{i=1}^n f(\varepsilon_i | \theta) = \left(\frac{1}{\sqrt{2\pi}}\right)^n \exp\left\{-\frac{1}{2} \sum_{i=1}^n \varepsilon_i^2\right\} = \left(\frac{1}{\sqrt{2\pi}}\right)^n \exp\left\{-\frac{1}{2} \sum_{i=1}^n (y_i - \theta x_i)^2\right\}$   
 $\arg \max_{\theta} L(\theta) = \arg \max_{\theta} \sum_{i=1}^n -\frac{1}{2} (y_i - \theta x_i)^2 = \arg \max_{\theta} \underbrace{\frac{1}{2} \sum_{i=1}^n (y_i - \theta x_i)^2}_{\text{Call this } Q(\theta)}$

$\frac{dQ(\theta)}{d\theta} = -\frac{1}{2} \sum_{i=1}^n (y_i - \theta x_i) x_i = 0$   
 $\Rightarrow \sum_{i=1}^n y_i x_i - \theta \sum_{i=1}^n x_i^2 = 0 \Rightarrow \theta = \frac{\sum y_i x_i}{\sum x_i^2} = \frac{\bar{y} \cdot \bar{x}}{\bar{x} \cdot \bar{x}} = \frac{(\bar{y} \cdot \bar{x})}{\bar{x} \cdot \bar{x}}$   
 projection of  $\bar{y}$  onto  $\bar{x}$

b) It has to be as the Normal  $N(\theta_0, 1)$  is conjugate prior to  $N(0, 1)$ .

$f(\theta|\bar{x}) = \underbrace{L(\bar{x}|\theta)}_{\propto L(\theta)} \cdot \pi(\theta) = \left(\frac{1}{\sqrt{2\pi}}\right)^n \exp\left\{-\frac{1}{2} \sum_{i=1}^n (y_i - \theta x_i)^2\right\} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(\theta - \theta_0)^2}$   
 $= \left(\frac{1}{\sqrt{2\pi}}\right)^{n+1} \exp\left\{-\frac{1}{2} \left[ \sum_{i=1}^n (y_i^2 - 2y_i \theta x_i + \theta^2 x_i^2) + (\theta - \theta_0)^2 \right]\right\}$   
 $= \left(\frac{1}{\sqrt{2\pi}}\right)^{n+1} \exp\left\{-\frac{1}{2} \left[ \sum_{i=1}^n (y_i^2 - 2y_i \theta x_i + \theta^2 x_i^2) + \theta^2 + \theta_0^2 - 2\theta\theta_0 \right]\right\}$

$l(\theta) = \sum_{i=1}^n (y_i^2 - 2y_i \theta x_i + \theta^2 x_i^2) + \theta^2 + \theta_0^2 - 2\theta\theta_0$

mean is when this function of  $\theta$  peaks.  
 The mode is the same. because mode = mean for Normal.

This is normal dist. as we just have quadratic of  $\theta$

$\frac{\partial l}{\partial \theta} = 0 \Rightarrow \sum_{i=1}^n (-2y_i x_i + 2\theta x_i^2) + 2\theta - 2\theta_0 = 0$

$\Rightarrow \hat{\theta} = \frac{\theta_0 + \sum_{i=1}^n y_i x_i}{1 + \sum_{i=1}^n x_i^2}$

c) posterior mean = MAP because for normal, mode = mean  $\therefore \hat{\theta} = \frac{\theta_0 + \sum y_i x_i}{1 + \sum x_i^2}$   
 $\hat{\theta} = \frac{\theta_0 + \sum y_i x_i}{1 + \sum x_i^2}$

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