Machine Learning for Data Science (CS 4786)

Lecture 15: EM Algorithm for Gaussian Mixture Models and Mixture of Multinomials

1 Gaussian Mixture Models

Each $\theta \in \Theta$ consist of mixture distribution π which is a distribution over the choices of the K clusters, $\mu_1, \ldots, \mu_K \in \mathbb{R}^d$ the choices of the K means for the corresponding gaussians and $\Sigma_1, \ldots, \Sigma_K$ the choices of the K covariance matrices. The latent variables are c_1, \ldots, c_n the cluster assignments for the n points and x_1, \ldots, x_n are the n observations.

1.1 E-step

On iteration i, for each data point $t \in [n]$, set

$$Q_t^{(i)}(c_t) = P(c_t|x_t, \theta^{(i-1)})$$

Note that

$$Q_t^{(i)}(c_t) = P(c_t|x_t, \theta^{(i-1)})$$

$$\propto \frac{p(x_t|c_t\theta^{(i-1)})}{p(x_t|c_t\theta^{(i-1)})} \times \frac{P(c_t|\theta^{(i-1)})}{p(c_t|\theta^{(i-1)})}$$

$$\propto \frac{1}{\sqrt{(2\pi)^d|\Sigma_{c_t}|}} \exp\left(-(x_t - \mu_{c_t})^{\top} \Sigma_{c_t}(x_t - \mu_{c_t})/2\right) \pi_{c_t}$$

1.2 M-step for GMM

For the M-step (for MLE) we would like to find

$$\theta = \underset{\theta \in \Theta}{\operatorname{argmax}} \sum_{t=1}^{n} \sum_{c_t=1}^{K} Q_t^{(i)}(c_t) \log P(x_t, c_t | \theta)$$

To this end note that

$$\sum_{t=1}^{n} \sum_{c_{t}=1}^{K} Q_{t}^{(i)}(c_{t}) \log P(x_{t}, c_{t}|\theta) = \sum_{t=1}^{n} \sum_{k=1}^{K} Q_{t}^{(i)}(k) \left(\log \phi(x_{t}|\mu_{k}, \Sigma_{k}) + \log \pi_{k}\right)$$

$$= \sum_{t=1}^{n} \sum_{k=1}^{K} Q_{t}^{(i)}(k) \left(\frac{1}{2} \log \left(\frac{1}{(2*3.14)^{d}|\Sigma_{k}|}\right) - \frac{1}{2} (x_{t} - \mu_{k})^{\top} \Sigma_{k}^{-1} (x_{t} - \mu_{k}) + \log \pi_{k}\right)$$

$$= \sum_{t=1}^{n} \sum_{k=1}^{K} Q_{t}^{(i)}(k) \left(-\frac{1}{2} \log \det(\Sigma_{k}) - \frac{1}{2} (x_{t} - \mu_{k})^{\top} \Sigma_{k}^{-1} (x_{t} - \mu_{k}) + \log \pi_{k}\right) + \text{constant terms}$$

For notational convenience define:

$$L(\mu_{1:K}, \Sigma_{1:K}, \pi) = \sum_{t=1}^{n} \sum_{k=1}^{K} Q_t^{(i)}(k) \left(-\frac{1}{2} \log \det(\Sigma_k) - \frac{1}{2} (x_t - \mu_k)^{\top} \Sigma_k^{-1} (x_t - \mu_k) + \log \pi_k \right)$$

Our goal is to find parameters that maximize $L(\mu_{1:K}, \Sigma_{1:K}, \pi)$.

M-step for mean: To optimize with respect to mean we take derivative and equate to 0,

$$\frac{\partial}{\partial \mu_k} L(\mu_{1:K}, \Sigma_{1:K}, \pi) = -\frac{1}{2} \frac{\partial}{\partial \mu_k} \left(\sum_{t=1}^n Q_t^{(i)}(k) (x_t - \mu_k)^\top \Sigma_k^{-1} (x_t - \mu_k) \right)
= -\sum_{t=1}^n Q_t^{(i)}(k) \Sigma_k^{-1} (x_t - \mu_k) = -\Sigma_k^{-1} \left(\sum_{t=1}^n Q_t^{(i)}(k) (x_t - \mu_k) \right)$$

To maximize over μ_k we set derivative equal to 0. Hence

$$\sum_{t=1}^{n} Q_t^{(i)}(k)(x_t - \mu_k) = \sum_{t=1}^{n} Q_t^{(i)}(k)x_t - \mu_k \left(\sum_{t=1}^{n} Q_t^{(i)}(k)\right) = 0$$

Or equivallently:

$$\mu_k = \frac{\sum_{t=1}^n Q_t^{(i)}(k) x_t}{\sum_{t=1}^n Q_t^{(i)}(k)} \qquad \begin{array}{c} \text{Qt: n*k} \\ \mu_k = \text{sum of weight*location} \\ \text{divided by} \\ \text{sum of weight} \\ \text{M-step for mixture distribution:} \text{ Since we want to optimize over } \pi \text{ subject to the constraint} \end{array}$$

M-step for mixture distribution: Since we want to optimize over π subject to the constraint $\sum_{k=1}^{K} \pi_k = 1$ (ie. its a distribution), we do so by introducing Lagrange variables. That is we want to optimize the following term w.r.t. π_k

$$L(\mu_{1:K}, \Sigma_{1:K}, \pi) + \frac{\lambda}{k} (\sum_{k=1}^{K} \pi_k - 1)$$

Hence taking derivative of above w.r.t. π we get,

$$\frac{\partial}{\pi_k} L(\mu_{1:K}, \Sigma_{1:K}, \pi) + \lambda (\sum_{k=1}^K \pi_k - 1) = \frac{\partial}{\partial \pi_k} L(\mu_{1:K}, \Sigma_{1:K}, \pi) + \lambda$$

But,

$$\frac{\partial}{\partial \pi_k} L(\mu_{1:K}, \Sigma_{1:K}, \pi) = \frac{\partial}{\partial \pi_k} \sum_{t=1}^n Q_t^{(i)}(k) \log(\pi_k) = \frac{\sum_{t=1}^n Q_t^{(i)}(k)}{\pi_k}$$

Hence,

$$\frac{\partial}{\pi_k} L(\mu_{1:K}, \Sigma_{1:K}, \pi) + \lambda (\sum_{k=1}^K \pi_k - 1) + \sum_{i=1}^K \lambda_i \pi_i = \frac{\sum_{t=1}^n Q_t^{(i)}(k)}{\pi_k} + \lambda$$

?

Setting derivative to 0 we discover that

$$\pi_k \propto \sum_{t=1}^n Q_t^{(i)}(k)$$
 $\pi(\mathbf{k}) = 分子 / \text{lamda}$ $\pi(\mathbf{k}) = 分子 / \text{lamda}$ $\pi(\mathbf{k}) = 分子 / \text{lamda}$

M-step for Covariance: This one needs being able to take derivative w.r.t. matrices and so I will only sketch the proof here. Let us consider optimizing w.r.t. some Σ_k . It makes the problem easier if we instead think of the problem as optimizing over Σ_k^{-1} and then invert the solution.

Here are two facts that come in handy:

$$\frac{\partial}{\partial \mathbf{X}} \log \det(\mathbf{X}^{-1}) = -\mathbf{X}^{-1}$$

and for any vector v,

$$\frac{\partial}{\partial \mathbf{X}} v^{\top} \mathbf{X} v = v v^{\top}$$

Now note that

$$\frac{\partial}{\partial \Sigma_k} L(\mu_{1:K}, \Sigma_{1:K}, \pi) = \frac{\partial}{\partial \Sigma_k} \left(\sum_{t=1}^n Q_t^{(i)}(k) \left(-\frac{1}{2} \log \det(\Sigma_k) - \frac{1}{2} (x_t - \mu_k)^\top \Sigma_k^{-1} (x_t - \mu_k) \right) \right) \\
= \left(\sum_{t=1}^n Q_t^{(i)}(k) \left(\frac{1}{2} (\Sigma_k^{-1})^{-1} - \frac{1}{2} (x_t - \mu_k) (x_t - \mu_k)^\top \right) \right)$$

Hence equating to 0 we get that

$$\Sigma_k = \frac{\sum_{t=1}^n Q_t^{(i)}(k) (x_t - \mu_k) (x_t - \mu_k)^\top}{\sum_{t=1}^n Q_t^{(i)}(k)}$$

that is the weighted sample variance.

2 EM Algorithm: Why it works?

Log likelihood only decreases after one iteration of EM algorithm. Why?

3 Mixture of Multinomials

Each $\theta \in \Theta$ consist of mixture distribution π which is a distribution over the choices of the K clusters or types, p_1, \ldots, p_K the choices of the K multinomial distributions over the d items. The latent variables are c_1, \ldots, c_n the cluster assignments for the n points and x_1, \ldots, x_n are the n observations given as index of one of the d choices for t'th data point (product bought).

3.1 E-step

On iteration i, for each data point $t \in [n]$, set

$$Q_t^{(i)}(c_t) = P(c_t|x_t, \theta^{(i-1)})$$

Note that

$$Q_t^{(i)}(c_t) = P(c_t|x_t, \theta^{(i-1)})$$

$$\propto p(x_t|c_t, \theta^{(i-1)}) \times P(c_t|\theta^{(i-1)})$$

$$\propto p_{c_t}[x_t]\pi_{c_t}$$

3.2 M-step

To be filled in soon ...