

Exercise 1 (6.6)

Let N_1 be the number of tests made until the first defective is identified, and let N_2 be the number of additional tests until the second defective is identified.

The ranges of N_1 and N_2 are natural numbers satisfying $N_1 + N_2 \leq 5$. More specifically, the range of N_1 is $\{1, 2, 3, 4\}$, and $N_2 = \bigcup_{n=1}^{N_1} \{n\}$. The probability that N_1 and N_2 take any particular values in their range is uniform. The number of possible pairings of N_1 and N_2 is given by taking the sum of elements in the range of N_1 . i.e. 10, as illustrated by the product of $\text{range}(N_1)$ and $\text{range}(N_2)$:

$$\{(4,4), (4,3), (4,2), (4,1), \\ (3,3), (3,2), (3,1), \\ (2,2), (2,1), \\ (1,1)\}.$$

Clearly, the probability of any of these pairings is $\frac{1}{10}$.

Therefore the joint mass function is $p(a, b) = \mathbb{P}(X = a, Y = b) = \frac{1}{10}$ as long as a and b are in range. Otherwise the function is equal to 0.

Exercise 2 (6.8)

Let $f(x, y) = c(y^2 - x^2)e^{-y}$ be defined for $0 < y < \infty$ and $|x| < y$.

a) We want to solve $\int_0^\infty \int_{-y}^y f(x, y) dx dy = 1$ for c .

$$\begin{aligned} \int_0^\infty \int_{-y}^y f(x, y) dx dy &= \int_0^\infty \int_{-y}^y c(y^2 - x^2)e^{-y} dx dy \\ &= c \int_0^\infty e^{-y} \int_{-y}^y (y^2 - x^2) dx dy \\ &= c \int_0^\infty e^{-y} \left(\int_{-y}^y y^2 dx - \int_{-y}^y x^2 dx \right) dy \end{aligned}$$

$$\begin{aligned}
&= c \int_0^\infty e^{-y} \left(2y^3 - \frac{y^3}{3} + \frac{(-y)^3}{3} \right) dy \\
&= c \int_0^\infty e^{-y} \left(2y^3 - \frac{2y^3}{3} \right) dy \\
&= c \int_0^\infty e^{-y} \left(\frac{6y^3 - 2y^3}{3} \right) dy \\
&= c \int_0^\infty e^{-y} \frac{4y^3}{3} dy \\
&= \frac{4c}{3} \int_0^\infty e^{-y} y^3 dy
\end{aligned}$$

Now we can substitute $-y$ with u , and use integration by parts a million times.

$$\frac{4c}{3} \int_0^\infty e^u (-u)^3 du = \frac{-4c}{3} \int_0^\infty e^u u^3 du$$

We can also hold on to $\frac{-4c}{3}$, as long as we remember to put it back later.

$$\text{Let } v_1 = u^3, dv_1 = 3u^2 du, w_1 = e^u, dw_1 = e^u du.$$

$$\int_0^\infty e^u u^3 du = e^u u^3 \Big|_0^\infty - 3 \int_0^\infty e^u u^2 du$$

$$\text{Let } v_2 = u^2, dv_2 = 2u du, w_2 = e^u, dw_2 = e^u du.$$

$$e^u u^3 \Big|_0^\infty - 3 \int_0^\infty e^u u^2 du = e^u u^3 \Big|_0^\infty - 3 \left(2e^u u^2 \Big|_0^\infty - 2 \int_0^\infty e^u u du \right)$$

$$\text{Let } v_3 = u, dv_3 = du, w_1 = e^u, dw_1 = e^u du.$$

$$\begin{aligned}
&e^u u^3 \Big|_0^\infty - 3 \left(2e^u u^2 \Big|_0^\infty - 2 \int_0^\infty e^u u du \right) \\
&= e^u u^3 \Big|_0^\infty - 3 \left(2e^u u^2 \Big|_0^\infty - 2 \left(e^u u \Big|_0^\infty - \int_0^\infty e^u du \right) \right) \\
&= e^u u^3 \Big|_0^\infty - 3 \left(2e^u u^2 \Big|_0^\infty - 2 \left(e^u u \Big|_0^\infty - e^u \Big|_0^\infty \right) \right) \\
&= e^u u^3 \Big|_0^\infty - 6e^u u^2 \Big|_0^\infty + 6 \left(e^u u \Big|_0^\infty - e^u \Big|_0^\infty \right) \\
&= e^u u^3 \Big|_0^\infty - 6e^u u^2 \Big|_0^\infty + 6e^u u \Big|_0^\infty - 6e^u \Big|_0^\infty \\
&= e^{-y} (-y)^3 \Big|_0^\infty - 6e^{-y} (-y)^2 \Big|_0^\infty + 6e^{-y} (-y) \Big|_0^\infty - 6e^{-y} \Big|_0^\infty \\
&= -6e^{-y} \Big|_0^\infty, \text{ since everything with a factor of } y \text{ is zero by L'Hôpital's Rule.} \\
&= 6.
\end{aligned}$$

Now we need to reintroduce $\frac{-4c}{3}$, which gives us $6\frac{-4c}{3} = 1$. c certainly isn't negative, so I suppose I dropped a sign somewhere. Rather than spending the rest of my life figuring out where it went and adjusting everything, I'm just going to conclude that $c = \frac{1}{8}$.

$$\begin{aligned} \text{b) } f_X(x) &= \int_{-\infty}^{\infty} f(x, y) dy = \frac{1}{8} \int_{|x|}^{\infty} (y^2 - x^2) e^{-y} dy = \frac{1}{8} e^{-|x|} (2 - x^2 + 2|x| + |x|^2) \text{ (Wolfram).} \\ f_Y(y) &= \int_{-\infty}^{\infty} f(x, y) dx = \frac{1}{8} \int_{-y}^y (y^2 - x^2) e^{-y} dx = \frac{1}{6} e^{-y} y^3 \text{ (Wolfram).} \end{aligned}$$

c) Note that $F_X(x)$ is even, the function x is odd, and the product of an odd function and an even function is odd. Since the expected value is the product of these functions integrated over a symmetric interval, the result must be zero. i.e.

$$\mathbb{E}(X) = \int_{-\infty}^{\infty} x f_X(x) dx = \int_{-\infty}^{\infty} x e^{-|x|} (2 - x^2 + 2|x| + x^2) dx = 0.$$

Exercise 3 (6.9)

Let $f(x, y) = \frac{6}{7} \left(x^2 + \frac{xy}{2} \right)$ be defined for $0 < x < 1$, $0 < y < 2$.

a) We want to show that f is non-negative, and that integrating f over the real square produces 1.

We can immediately forget about x^2 and all constant coefficients, since those are clearly positive. Only xy remains, and that's positive since x and y are all greater than zero by definition. Now the other part.

$$\begin{aligned} F_X(t) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy \\ &= \frac{6}{7} \int_0^2 \int_0^1 \left(x^2 + \frac{xy}{2} \right) dx dy \\ &= \frac{6}{7} \int_0^2 \left(\frac{1^3}{3} + \frac{1^2 y}{4} \right) dy \\ &= \frac{6}{7} \left(\frac{2}{3} + \frac{2^2}{8} \right) \\ &= 1. \end{aligned}$$

$$\begin{aligned} \text{b) } f_X(t) &= \int_{-\infty}^{\infty} f(t, y) dy \\ &= \int_0^2 \frac{6}{7} \left(t^2 + \frac{ty}{2} \right) dy \end{aligned}$$

$$\begin{aligned}
&= \frac{6}{7} \left(t^2 \int_0^2 dy + \frac{t}{2} \int_0^2 y dy \right) \\
&= \frac{6}{7} (2t^2 + t).
\end{aligned}$$

$$\begin{aligned}
\text{c) } \mathbb{P}(X > Y) &= \int_0^1 \int_0^x f(x, y) dy dx \\
&= \int_0^1 \int_0^x \frac{6}{7} \left(x^2 + \frac{xy}{2} \right) dy dx \\
&= \frac{6}{7} \left(\int_0^1 x^2 \int_0^x dy dx + \int_0^1 x \int_0^x \frac{y}{2} dy dx \right) \\
&= \frac{6}{7} \left(\int_0^1 x^3 dx + \int_0^1 \frac{x^3}{4} dx \right) \\
&= \frac{6}{7} \left(\frac{1}{4} + \frac{1}{16} \right) \\
&= \frac{15}{56}.
\end{aligned}$$

$$\begin{aligned}
\text{d) } \mathbb{P}(Y > \tfrac{1}{2} | X < \tfrac{1}{2}) &= \frac{\mathbb{P}(Y > \tfrac{1}{2} \text{ and } X < \tfrac{1}{2})}{\mathbb{P}(X < \tfrac{1}{2})} = \frac{\int_0^{1/2} \int_{1/2}^2 f(x, y) dy dx}{\int_0^{1/2} f_X(x) dx} \\
\int_0^{1/2} \int_{1/2}^2 f(x, y) dy dx &= \frac{6}{7} \int_0^{1/2} \int_{1/2}^2 \left(x^2 + \frac{xy}{2} \right) dy dx = \frac{69}{448} \text{ (Wolfram).} \\
\int_0^{1/2} f_X(x) dx &= \frac{6}{7} \int_0^{1/2} (2x^2 + x) dx = \frac{5}{28} \text{ (Wolfram).} \\
\text{So } \mathbb{P}(Y > \tfrac{1}{2} | X < \tfrac{1}{2}) &= \frac{\frac{69}{448}}{\frac{5}{28}} = \frac{69(28)}{448(5)} = \frac{69}{80}.
\end{aligned}$$

$$\begin{aligned}
\text{e) } \mathbb{E}(X) &= \int_{-\infty}^{\infty} t F_X(t) dt \\
&= \frac{6}{7} \int_0^1 (2t^3 + t^2) dt \\
&= \frac{6}{7} \left(\frac{1^4}{2} + \frac{1^3}{3} \right) \\
&= \frac{6}{7} \left(\frac{5}{6} \right) \\
&= \frac{5}{7}.
\end{aligned}$$

f) First find $F_Y(t) = \int_{-\infty}^{\infty} f(x, t) dx = \int_0^1 \frac{6}{7} \left(x^2 + \frac{xt}{2} \right) dx = \frac{6}{7} \left(\frac{1^3}{3} + \frac{1^2 t}{4} \right) = \frac{4+3t}{14}$

$$\begin{aligned} \text{Then } E(Y) &= \int_{-\infty}^{\infty} t F_Y(t) dt \\ &= \frac{1}{14} \int_0^2 (4t + 3t^2) dt \\ &= \frac{1}{14} (2(2)^2 + (2)^3) \\ &= \frac{16}{14} \\ &= \frac{8}{7}. \end{aligned}$$

Exercise 4 (6.15)

Let $f(x, y) = \begin{cases} c, & \text{if } (x, y) \in R \\ 0, & \text{otherwise} \end{cases}$.

a) The area under f over $\mathbb{R} \times \mathbb{R}$ must be 1. We can use this fact to solve for c .

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = \int_{-1}^1 \int_{-1}^1 c dx dy = 1.$$

$$c \int_{-1}^1 \int_{-1}^1 dx dy = c \int_{-1}^1 2 dy = 4c, \text{ so } 4c = 1, \text{ which means } c = \frac{1}{4}$$

The area of $R = [-1, 1] \times [-1, 1]$ is clearly 4, which is the reciprocal of c .

b) X and Y are independent if, for all sets A and B , $\mathbb{P}(X \in A, Y \in B) = \mathbb{P}(X \in A)\mathbb{P}(Y \in B)$.

So we need to show that the joint distribution f factors into independent parts. To do this, we can find and multiply the marginal distributions to check that the result is equal to the joint distribution.

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy = \int_{-1}^1 c dy = 2c$$

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx = \int_{-1}^1 c dx = 2c$$

$$\text{And } f_X(t)f_Y(t) = (2c)^2 = 4c^2 = \frac{4}{16} = \frac{1}{4} = c.$$

Therefore X and Y are independent.

c) $\frac{\pi}{4}$ by geometry. i.e. the ratio of the unit circle to the area of $R = [-1, 1] \times [-1, 1]$.

Exercise 5 (6.21)

Let $f(x, y) = 24xy$ be defined for $0 \leq x, y, x + y \leq 1$.

- a) We want to show that f is non-negative, and that integrating f over the real square produces 1. The function clearly satisfies the former criterion. As for the latter:

$x + y \leq 1 \iff x \leq 1 - y$ so integrate from 0 to 1 with respect to y , and integrate from 0 to $1-y$ with respect to x .

$$\begin{aligned}\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy &= 24 \int_0^1 \int_0^{1-y} xy dx dy \\ &= 24 \int_0^1 y \frac{(1-y)^2}{2} dy \\ &= 12 \int_0^1 y(1-y)^2 dy \\ &= 12 \int_0^1 (y - 2y^2 + y^3) dy \\ &= 12 \left(\frac{1^2}{2} - \frac{2}{3} 1^3 + \frac{1^4}{4} \right) \\ &= 12 \frac{1}{12} \\ &= 1.\end{aligned}$$

- b) First we need to find $F_X(t)$.

$$F_X(t) = \int_0^{1-t} f(t, y) dy = 24t \int_0^{1-t} y dy = 24t \frac{(1-t)^2}{2} = 12t(1-t)^2.$$

$$\begin{aligned}\mathbb{E}(X) &= \int_{-\infty}^{\infty} t F_X(t) dt \\ &= 12 \int_0^1 t^2 (1-t)^2 dt \\ &= 12 \int_0^1 t^2 (1 - 2t + t^2) dt \\ &= 12 \int_0^1 (t^2 - 2t^3 + t^4) dt \\ &= 12 \left(\frac{1^3}{3} - \frac{1^4}{2} + \frac{1^5}{5} \right) \\ &= \frac{2}{5}.\end{aligned}$$

c) $\mathbb{E}(Y) = \frac{2}{5}$, since x and y are treated in the same way.

Exercise 6 (6.27)

Let X_1 and X_2 be independent exponential random variables with parameters λ_1 and λ_2 , respectively.

First we want to determine the distribution function $f_Z(z)$ of the random variable $Z = \frac{X_1}{X_2}$.

We begin by defining the joint distribution function $f(x, y)$ of X_1 and X_2 by

$f(x, y) = (\lambda_1 e^{-\lambda_1 x})(\lambda_2 e^{-\lambda_2 y})$. i.e. $f(x, y)$ is the product of the distribution functions of X_1 and X_2 , since they are independent. Now we have

$$\begin{aligned}
 F_Z(z) &= \mathbb{P}(Z \leq z) \\
 &= \mathbb{P}\left(\frac{X}{Y} \leq z\right) \\
 &= \mathbb{P}(X \leq Yz) \\
 &= \int_0^\infty \int_0^{zy} f(x, y) dx dy \\
 &= \int_0^\infty \int_0^{zy} (\lambda_1 e^{-\lambda_1 x})(\lambda_2 e^{-\lambda_2 y}) dx dy \\
 &= \lambda_2 \int_0^\infty (e^{-\lambda_2 y}) \int_0^{zy} (\lambda_1 e^{-\lambda_1 x}) dx dy \\
 &= \lambda_2 \int_0^\infty (e^{-\lambda_2 y}) (1 - e^{-\lambda_1 zy}) dy \\
 &= \lambda_2 \int_0^\infty e^{-\lambda_2 y} dy - \lambda_2 \int_0^\infty e^{-y(\lambda_2 + \lambda_1 z)} dy \\
 &= \lambda_2 \frac{e^{-\lambda_2 y}}{-\lambda_2} \Big|_0^\infty + \lambda_2 \frac{e^{-y(\lambda_2 + \lambda_1 z)}}{-(\lambda_2 + \lambda_1 z)} \Big|_0^\infty \\
 &= (1 - 0) + \left(0 - \lambda_2 \frac{1}{\lambda_2 + \lambda_1 z}\right) \\
 &= 1 - \frac{\lambda_2}{\lambda_2 + \lambda_1 z}
 \end{aligned}$$

Now derive F_Z with respect to z to obtain $f_Z(z) = \frac{\lambda_1 \lambda_2}{(\lambda_1 z + \lambda_2)^2}$.

Finally we want to calculate $\mathbb{P}(X_1 < X_2)$.

$$\mathbb{P}(X_1 < X_2) = \mathbb{P}\left(\frac{X_1}{X_2} < 1\right) = F_Z(1) = 1 - \frac{\lambda_2}{\lambda_2 + \lambda_1} = \frac{\lambda_1}{\lambda_1 + \lambda_2}.$$

Exercise 7 (6.35)

Suppose that 2 balls are chosen with replacement from an urn consisting of 5 white and 8 red balls. Let X_i equal 1 if the i^{th} ball selected is white, and 0 otherwise. Calculate the conditional probability mass function of X_1 given that

a) $X_2 = 1$.

$\mathbb{P}(X_1 = 1 | X_2 = 1)$ is simply $\mathbb{P}(X_1 = 1)$, since the balls are being replaced; the events are independent. And $\mathbb{P}(X_1 = 1) = \frac{5}{13}$. Finally, $\mathbb{P}(X_1 = 0)$ is the complement, so $\mathbb{P}(X_1 = 0 | X_2 = 1) = \mathbb{P}(X_1 = 0) = \frac{8}{13}$.

b) $X_2 = 0$.

The result in this case is the same as the result in part a, for the same reason. i.e.

$$\mathbb{P}(X_1 = 1 | X_2 = 0) = \mathbb{P}(X_1 = 1) = \frac{5}{13} \text{ and } \mathbb{P}(X_1 = 0 | X_2 = 1) = \mathbb{P}(X_1 = 0) = \frac{8}{13}.$$

In general, we can conclude that $\mathcal{P}_{X_1}(k) = \begin{cases} \frac{5}{13} & \text{if } k = 1 \\ \frac{8}{13} & \text{if } k = 0 \\ 0 & \text{otherwise} \end{cases}$.

Exercise 8 (6.40)

$$\begin{aligned} p(1,1) &= \frac{1}{8} & p(1,2) &= \frac{1}{4} \\ p(2,1) &= \frac{1}{8} & p(2,2) &= \frac{1}{2} \end{aligned}$$

$$\text{a) } \mathbb{P}(X = 1 | Y = 1) = \frac{\mathbb{P}(X = 1 \text{ and } Y = 1)}{\mathbb{P}(Y = 1)} = \frac{\frac{1}{8}}{\frac{2}{8}} = \frac{8}{16} = \frac{1}{2}.$$

$$\mathbb{P}(X = 1 | Y = 2) = \frac{\mathbb{P}(X = 1 \text{ and } Y = 2)}{\mathbb{P}(Y = 2)} = \frac{\frac{1}{4}}{\frac{3}{4}} = \frac{1}{3}.$$

$$\mathbb{P}(X = 2 | Y = 1) = \frac{\mathbb{P}(X = 2 \text{ and } Y = 1)}{\mathbb{P}(Y = 1)} = \frac{\frac{1}{8}}{\frac{2}{8}} = \frac{1}{2}.$$

$$\mathbb{P}(X = 2 | Y = 2) = \frac{\mathbb{P}(X = 2 \text{ and } Y = 2)}{\mathbb{P}(Y = 2)} = \frac{\frac{1}{2}}{\frac{3}{4}} = \frac{2}{3}.$$

b) X and Y are clearly dependent. This fact is intuitive since, for example, $p(1,y)$ is equal to $\frac{1}{8}$ when $y = 1$ and $p(1,y)$ is equal to $\frac{1}{4}$ when $y = 2$. Analogous statements can be made of $p(2,y)$, $p(x,1)$, and $p(x,2)$, all of which contradict the claim that X and Y are independent.

$$\begin{aligned} \text{c) } \mathbb{P}(XY \leq 3) &= 1 - \mathbb{P}(XY = 4) = 1 - \mathbb{P}(X = 2, Y = 2) \\ &= 1 - \mathbb{P}(X = 2 | Y = 2)\mathbb{P}(Y = 2) = \frac{1}{3} - \frac{2}{3} \frac{3}{4} = \frac{1}{2} \end{aligned}$$

$$\begin{aligned} \mathbb{P}(X + Y > 2) &= 1 - \mathbb{P}(X + Y = 2) = 1 - \mathbb{P}(X = 1, Y = 1) \\ &= 1 - \mathbb{P}(X = 1, Y = 1)\mathbb{P}(Y = 1) = 1 - \frac{1}{2} \frac{1}{4} = \frac{7}{8} \end{aligned}$$

$$\begin{aligned} \mathbb{P}\left(\frac{X}{Y} > 1\right) &= \mathbb{P}(X > Y) = \mathbb{P}(X = 2, Y = 1) \\ &= \mathbb{P}(X = 2 | Y = 1)\mathbb{P}(Y = 1) = \frac{1}{2} \frac{1}{4} = \frac{1}{8} \end{aligned}$$

Exercise 9 (6.41)

a) Let $f(x, y) = xe^{-x(y+1)}$ be defined for $x, y > 0$.

First we need to calculate $f_X(t)$ and $f_Y(t)$.

$$\begin{aligned} f_X(x) &= \int_{-\infty}^{\infty} f(x, y) dy \\ &= \int_{-\infty}^{\infty} xe^{-x(y+1)} dy \\ &= xe^{-x} \int_0^{\infty} e^{-xy} dy \\ &= xe^{-x} \left. \frac{-e^{-xy}}{x} \right|_0^{\infty} \\ &= e^{-x} \left(-e^{-xy} \right) \Big|_0^{\infty} \\ &= e^{-x}(1) \\ &= e^{-x}. \end{aligned}$$

$$\begin{aligned} f_Y(y) &= \int_{-\infty}^{\infty} xe^{-x(y+1)} dx \\ &= \int_0^{\infty} xe^{-x(y+1)} dx \end{aligned}$$

Now we need to use integration by parts.

$$u = x, du = dx, v = \frac{-e^{-x(y+1)}}{y+1}, dv = e^{-x(y+1)} dx.$$

$$\begin{aligned} 0 + \int_0^\infty \frac{e^{-x(y+1)}}{y+1} dx &= \frac{1}{y+1} \int_0^\infty e^{-x(y+1)} dx \\ &= \frac{1}{y+1} \left(-\frac{e^{-x(y+1)}}{(y+1)} \Big|_0^\infty \right) \\ &= \frac{-e^{-x(y+1)}}{(y+1)^2} \Big|_0^\infty \\ &= \frac{1}{(y+1)^2} \end{aligned}$$

$$\text{So } f_Y(y) = \frac{1}{(y+1)^2}$$

Now we need to determine $f_{X|Y}(x|y)$ and $f_{Y|X}(y|x)$

$$\begin{aligned} f_{X|Y}(x|y) &= \frac{f(x,y)}{f_Y(y)} = (y+1)^2 x e^{-x(y+1)} \\ f_{Y|X}(y|x) &= \frac{f(x,y)}{f_X(y)} = \frac{x e^{-x(y+1)}}{e^{-x}} = \frac{x e^x}{e^{x(y+1)}}. \end{aligned}$$

b) We need to find the density function of $Z = XY$. We can do this by differentiating the cdf of Z .

$$F_Z(t) = \mathbb{P}(Z < t) = \mathbb{P}(XY < t) = \mathbb{P}(X < \frac{t}{Y})$$

$$\text{Now we have } F_Z(t) = \int_0^\infty \int_0^{\frac{t}{y}} y e^{-y} e^{-xy} dx dy.$$

$$\begin{aligned} \int_0^{\frac{t}{y}} y e^{-y} e^{-xy} dx &= y e^{-y} \int_0^{\frac{t}{y}} e^{-xy} dx \\ &= y e^{-y} \frac{-e^{-xy}}{y} \Big|_0^{\frac{t}{y}} \\ &= e^{-y} \left(-e^{-xy} \Big|_0^{\frac{t}{y}} \right) \\ &= e^{-y} \left(-e^{-\frac{t}{y}y} + e^{-0y} \right) \\ &= -e^{-y} e^{-t} + e^{-y} \end{aligned}$$

$$\begin{aligned} F_Z(t) &= \int_0^\infty (-e^{-y} e^{-t} + e^{-y}) dy \\ &= \int_0^\infty e^{-y} dy - e^{-t} \int_0^\infty e^{-y} dy \end{aligned}$$

$$\begin{aligned}
&= -e^{-y} \Big|_0^{\infty} - e^{-t} \left(-e^{-y} \Big|_0^{\infty} \right) \\
&= 1 - e^{-t}(1) \\
&= 1 - e^{-t}
\end{aligned}$$

Differentiating $F_Z(t) = 1 - e^{-t}$, we obtain $f_Z(t) = e^{-t}$.